

# vbTPM user guide

Martin Lindén, bmelinden@gmail.com, October 21, 2013.

vbTPM is an acronym for variational Bayes for Tethered Particle Motion. The latest version of this manual and the corresponding software is available as open source via <http://sourceforge.net/projects/vbtpm/>.

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## 1 Getting started with vbTPM

### 1.1 Installation

Get the source code, and make sure VB7, HMMcore, and tools are in your matlab path, for example by running the script `vbTPMstart` from the matlab command prompt. (To add these paths permanently to your matlab path, see the matlab documentation).

Make sure that `HMMcore/` contains binaries for your systems. If not, a simple matlab compilation script can be found in `HMMcore/`.

### 1.2 Hardware requirements

Tethered particle motion often produces large data sets of many long trajectories, which makes the HMM analysis computer intensive. As an example, one parameter point in our test data sets, about 90 trajectories averaging 45 min, downsampled to 10 Hz, took about 24 h to go through on two 6 core Intel Xeon E5645 2.40GHz processors. The analysis time increases sharply with

the number of states (including spurious ones, like transient sticking events).

### 1.3 A small test problem

A small test problem, which runs on a (fast) laptop in about one hour, with examples of data and runinput files, can be found in `example1/`, with the actual data in `example1/lacdata/`. The data set has one calibration (cal) and one production (trj) trajectory for each bead.

#### Runinput files

Runinput files contain all parameters to run the analysis and access the results. The meaning of the parameters are documented in the help text of `VB7_batch_run.m`, and commented in those files. `runinput1.m` refers to an already completed analysis (results in `example1/HMMresults1/`), while `runinput2.m` has not yet run.

#### Run basic analysis

To start analyzing the test data set, type `VB7_batch_run('runinput2')` in the Matlab command prompt. Since the runinput file has `one_at_a_time=true`; this will analyze one trajectory in the data set. Several calls are needed to complete the analysis.

In our experience, matlab tends to hoard memory if several large data sets are analyzed consecutively. To work around that, one can use scripts that starts several matlab sessions with a single trajectory in each. An example for the bash shell is `runscript1.sh`, which calls `runinput1.m`. To parallelize, run several instances of the script at once. `VB7_batch_run` keeps track of which trajectories have already been 'checked out', so it is also possible to run on

several computers, if the results folder is synced regularly (if the same trajectory is checked out multiple times, old results are simply overwritten).

#### Manage the analysis

`VB7_batch_manage` is a tool for managing the basic analysis. It can collect the results and write them to a file, count how many trajectories in a data set has been analyzed, and also clean up temporary files from unfinished trajectories, which is useful if an analysis run is interrupted.

#### Access the results

The GUI for manual state classification is called `VB7_batch_postprocess()`. The GUI can be used to inspect the analysis results in detail, and can also convert the simple HMM models to factorial models for further analysis. To try it out, use the runinput file `runinput1.m`, which is already analyzed. To access the fitted models directly, use `VB7_batch_manage` with the 'collect' option. The results are returned as cell vectors for calibration and production trajectories, with the same index structure as the filenames in the runinput file.

For details on how vbTPM represents the models etc., we refer to section 5.

### 1.4 Other useful scripts

#### 1.4.1 Data and options

`VB7_getOptions` reads a runinput file and return all variables in a struct.

`VB7_preprocess` Converts trajectory data to a format that the analysis code uses. Input trajectory should be drift-corrected.

**BWdriftcorrect** applies driftcorrection to a position trajectory using a Butterworth-filter.

**RMSKBgaussfilter** computes running averages of RMS and other things, using a Gaussian kernel filter for smoothing.

**VB7\_getTrjData** returns the data for a single trajectory in a runinput file in various formats.

#### 1.4.2 Models

**VB7\_priorParent** A tool to initiate models of various sizes with consistent prior distributions.

**VB7\_initialGuess\_KBregion** is a rather complicated function to generate an initial guess for a model struct (e.g., fill out the M and Mc fields) based on analyzing the data.

**VB7\_GSconversion** is a tool to create factorial models, by converting genuine states into spurious ones.

**VB7\_removeState** removes states from a model object.

**VB7\_findGenuine** applies a simple set of rules to determine which states in a given model are genuine and spurious. An analyzed model for the corresponding calibration trace is also needed to provide a baseline.

**VB7\_inspectStates** is a simple tool to navigate in the raw data with the help of a converged model, for example to take a closer look at hard-to-classify states.

#### 1.4.3 VB-EM iterations

**VB7\_VBEMiter** is the computational core of vbTPM, and runs a single VB-EM iteration.

**VB7iterator** runs VB-EM iterations of a model until convergence.

**VB7\_greedySearch** runs a model search on a single trace from a given initial guess. Briefly, the search strategy is to systematically remove low-occupancy states until the lower bound  $F$  stops increasing.

**VB7\_analyzeTrace** runs several greedy model searches on single traces.

## 2 Diffusive model for TPM

### 2.1 Diffusive hidden Markov model

We model the looping dynamics by a discrete Markov process  $s_t$  with  $N$  states, a transition probability matrix  $\mathbf{A}$ , and initial state distribution  $\vec{\pi}$ ,

$$p(s_t|s_{t-1}, \mathbf{A}) = A_{s_{t-1}s_t}, \quad p(s_1) = \pi_{s_1}. \quad (1)$$

This is the standard hidden part of an HMM, and the physics of TPM goes into the emission model, that describes the restricted Brownian motion of the bead. We use a discrete time model of over-damped 2D diffusion in a harmonic potential, that has been suggested as a simplified model for TPM (1, 2),

$$\vec{x}_t = K_{s_t}\vec{x}_{t-1} + \vec{w}_t/(2B_{s_t})^{1/2}, \quad (2)$$

where the index  $s_t$  indicate parameters that depend on the hidden state. Thermal noise enters through the uncorrelated Gaussian random

vectors  $\vec{w}_t$  with unit variance. The unintuitive parametrization is chosen for computational convenience;  $K_j$  and  $B_j$  are related to the spring and diffusion constant of the bead, and some insight into their physical meaning can be gained by noting that with a single hidden state, Eq. 2 describes a Gaussian process with zero mean and

$$\begin{aligned} RMS &= \sqrt{\langle \vec{x}^2 \rangle} = (B(1 - K^2))^{-1/2}, \\ \frac{\langle \vec{x}_{t+m} \cdot \vec{x}_t \rangle}{\langle \vec{x}^2 \rangle} &= K^m \equiv e^{-m\Delta t/\tau}, \end{aligned} \quad (3)$$

where  $\Delta t$  is the sampling time, and  $\tau$  is a bead correlation time. This model thus captures the diffusive character of the bead motion, while still retaining enough simplicity to allow efficient variational algorithms (3, 4).

## 2.2 Factorial model

To separate genuine looping dynamics from artifacts such as transient sticking events and tracking errors, we introduce a second hidden state  $c_t$  that works like a dynamic indicator:  $c_t = 1$  indicates genuine TPM, with bead motion described by  $K_{s_t}, B_{s_t}$ , while  $c_t > 1$  indicates some experimental artifact in action. The bead motion is then described by parameters  $\hat{K}_{c_t}, \hat{B}_{c_t}$  that depend on  $c_t$ , while  $s_t$  evolves without influencing the data.

One might imagine using different types of models for genuine and spurious states (e.g., transient sticking events might have a different tethering point than genuine states). We do not explore this option, but instead classify states based on parameter values, as described in the main text (?).

We also allow the transition probabilities from  $c_t = 1$  to  $c_t > 1$  to depend on  $s_t$ , to allow for the possibility that sticking events happen more easily in looped than unlooped states for example.

Thus, the hidden states evolve according to

$$p(s_{t+1}, c_{t+1} | s_t, c_t) = p(s_{t+1} | s_t) p(c_{t+1} | s_t, c_t), \quad (4)$$

with  $p(s_{t+1} | s_t) = A_{s_t s_{t+1}}$  as usual, and

$$p(c_{t+1} | s_t, c_t) = \begin{cases} \hat{A}_{s_t c_{t+1}}, & \text{if } c_t = 1, \\ \hat{R}_{c_t c_{t+1}}, & \text{if } c_t > 1, \end{cases} \quad (5)$$

and initial probabilities are independent,

$$p(s_1 = j, c_1 = k) = \pi_j \hat{\pi}_k. \quad (6)$$

## 3 The VB-algorithm

We will start by discussing the statistical analysis of the simple HMM in detail, and then discuss generalization to the factorial model.

### 3.1 Model selection by maximum evidence

Our analysis aims not only to extract parameter values from TPM data, but also to learn the number of hidden states  $N$ , corresponding to different DNA-protein conformations. This means that we need to compare models with different number of unknown parameters. We take a Bayesian approach to this problem.

A distinguishing feature of Bayesian data analysis is the treatment of random variables and unknown parameters on an equal footing (5, 6). Hence, given some data  $\vec{x}_{1:T}$  and a set of competing models with different number of states  $N = 1, 2, \dots$  (and  $1:T$  is a compact way to denote a whole time series), we can use the laws of probability to express our confidence about those models in terms of conditional probabilities,

$$p(N | \vec{x}_{1:T}) = p(\vec{x}_{1:T} | N) p(N) / p(\vec{x}_{1:T}), \quad (7)$$

where  $p(N)$  expresses our beliefs about the different models prior to seeing the data, and  $p(\vec{x}_{1:T})$  is a normalization constant. A Bayesian rule for model selection is therefore to prefer the model that maximizes  $p(\vec{x}_{1:T}|N)$ , a quantity known as the evidence. For our more complex model, parameters and hidden states will have to be integrated out,

$$p(\vec{x}_{1:T}|N) = \int d\theta \sum_{s_{1:T}} p(\vec{x}_{1:T}, s_{1:T}|\theta) p(\theta|N), \quad (8)$$

where the first factor in the integrand describes the model, and the second expresses our prior beliefs about the parameters (see below).

The integrand in the evidence, Eq. (8), requires an explicit expression for the probability of a sequence of bead positions and hidden states. This expression can be written down based on the above model, and factorizes in the usual HMM fashion, as

$$\begin{aligned} p(\vec{x}_{1:T}, s_{1:T}|\theta) p(\theta|N) &= p(\vec{x}_1) p(s_1|\vec{\pi}) \\ &\times \prod_{t=2}^T p(\vec{x}_t|\vec{x}_{t-1}, s_t, \vec{K}, \vec{B}) p(s_t|s_{t-1}, \mathbf{A}) \\ &\times p(\vec{\pi}|N) \prod_{j=1}^N p(K_j, B_j|N) p(A_{j,:}|N), \end{aligned} \quad (9)$$

where  $A_{j,:}$  denote row  $j$  of the matrix  $\mathbf{A}$ . The first right hand side line in Eq. (9) describes the initial state and bead position. We will neglect the factor  $p(\vec{x}_1)$  from now on, but the initial state  $p(s_1|\vec{\pi})$  and transition probabilities  $p(s_t|s_{t-1}, \mathbf{A})$  are given by Eq. (1), and the bead motion follows from Eq. (2),

$$p(\vec{x}_t|\vec{x}_{t-1}, s_t, \vec{K}, \vec{B}) = \frac{B_{s_t}}{\pi} e^{-B_{s_t}(\vec{x}_t - K_{s_t}\vec{x}_{t-1})^2}. \quad (10)$$

Finally, the last line of Eq. (9) contains prior distributions over parameters conditional on the number of states. We use conjugate priors, parameterized to have minimal impact on the inference results (see SI).

### 3.2 The variational approximation

An exact computation of the Bayesian evidence is impractical or impossible for most interesting models, and clever approximations are needed. The approximation we use here is variously known as ensemble learning, variational Bayes, or (in statistical physics jargon) mean field theory (4, 6), has previously been applied to biophysical time-series of FRET data (7–9) and *in vivo* single particle tracking (10). The idea is to approximate the log evidence by a lower bound,  $\ln p(x|N) \geq F_N$ , with

$$F_N = \int d\theta \sum_s q(s) q(\theta) \ln \frac{p(x, s|\theta) p(\theta|N)}{q(s) q(\theta)}, \quad (11)$$

where  $q(s)$  and  $q(\theta)$  are arbitrary probability distributions over the hidden states and parameters respectively. These are optimized to make the bound as tight as possible for each model, the model that achieves the highest lower bound wins, and the corresponding optimal distributions  $q(s)q(\theta)$  can be used for approximate inference about parameter values and hidden states. In particular, optimizing  $F_N$  with respect to the variational distributions leads to

$$\ln q(\theta) = -\ln Z_\theta + \ln p(\theta|N) + \langle \ln p(x, s|\theta) \rangle_{q(s)}, \quad (12)$$

$$\ln q(s) = -\ln Z_s + \langle \ln p(x, s|\theta) \rangle_{q(\theta)}, \quad (13)$$

where the  $Z$ 's are Lagrange multipliers to enforce normalization, and  $\langle \cdot \rangle_{q(\cdot)}$  denotes an aver-

age over  $q(\cdot)$ . We solve these equations iteratively until the lower bound converges, repeating the analysis many times with independent initial conditions in order to find a global maximum. The iterative solution approach results in an EM-type variational algorithm, detailed below. We refer to Refs. (3, 10, 11) for details on how to derive variational algorithms for HMMs, and Refs. (4, 6, 11) for more general discussion of variational inference methods.

### 3.2.1 Parameter distributions

The results of plugging our diffusive HMM into the parameter update equation (12) are as follows. The initial state probability vector, and each row in the transition matrix (denoted  $A_{j,:}$ ), are Dirichlet distributed,

$$q(\vec{\pi}) = \text{Dir}(\vec{\pi} | \vec{w}^{(\vec{\pi})}), \quad (14)$$

$$w_j^{(\vec{\pi})} = \tilde{w}_j^{(\vec{\pi})} + \langle \delta_{j,s_1} \rangle_{q(s_{1:T})}, \quad (15)$$

$$q(A_{i,:}) = \text{Dir}(A_{i,:} | \vec{w}^{(\mathbf{A})}), \quad (16)$$

$$w_{ij}^{(\mathbf{A})} = \tilde{w}_{ij}^{(\mathbf{A})} + \sum_{t=2}^T \langle \delta_{i,s_{t-1}} \delta_{j,s_t} \rangle_{q(s_{1:T})}. \quad (17)$$

Here, variables under tilde's ( $\tilde{\cdot}$ ) are hyperparameters that parameterize the prior distributions, and can be interpreted as pseudo-observations. The Dirichlet density function is

$$\text{Dir}(\vec{\pi} | \vec{u}) = \frac{\Gamma(u_0)}{\prod_j \Gamma(u_j)} \prod_j \pi_j^{u_j-1}, \quad u_j > 1, \quad (18)$$

where  $u_0 = \sum_j u_j$  is called the strength, and the density is non-zero in the region  $0 \leq \pi_j \leq 1$ ,  $\sum_j \pi_j = 1$ . Before moving on, we quote some useful expectation values for future reference,

$$\langle \ln \pi_i \rangle_{q(\vec{\pi})} = \psi(w_i^{(\vec{\pi})}) - \psi(w_0^{(\vec{\pi})}), \quad (19)$$

$$\langle \ln A_{ij} \rangle_{q(\mathbf{A})} = \psi(w_{ij}^{(\mathbf{A})}) - \psi(w_{i0}^{(\mathbf{A})}), \quad (20)$$

where  $\psi(x)$  is the digamma function, and

$$\langle \pi_i \rangle_{q(\vec{\pi})} = \frac{w_i^{(\vec{\pi})}}{w_0^{(\vec{\pi})}}, \quad (21)$$

$$\text{Var}[\pi_i]_{q(\vec{\pi})} = \frac{w_i^{(\vec{\pi})}((1 - w_i^{(\vec{\pi})}))}{(w_0^{(\vec{\pi})})^2(1 + w_0^{(\vec{\pi})})}, \quad (22)$$

$$\langle A_{ij} \rangle_{q(\mathbf{A})} = \frac{w_{ij}^{(\mathbf{A})}}{w_{i0}^{(\mathbf{A})}}, \quad (23)$$

$$\text{Var}[A_{ij}]_{q(\mathbf{A})} = \frac{w_{ij}^{(\mathbf{A})}(1 - w_{ij}^{(\mathbf{A})})}{(w_{i0}^{(\mathbf{A})})^2(1 + w_{i0}^{(\mathbf{A})})}. \quad (24)$$

The bead motion parameters have the following variational distributions

$$q(K_j, B_j) = \frac{B_j^{n_j}}{W_j} e^{-B_j(v_j(K_j - \mu_j)^2 + c_j)}, \quad (25)$$

$$W_j = \frac{c^{-(n_j + \frac{1}{2})} \Gamma(n_j + \frac{1}{2})}{\sqrt{v_j/\pi}}, \quad (26)$$

with the range  $B_j \geq 0$ ,  $-\infty < K_j < \infty$ . Physically, we might rather expect  $0 < K_j < 1$ , but the extended range for  $K_j$  simplifies the calculations a lot. The VBM equations are

$$n_j = \tilde{n}_j + M_j, \quad (27)$$

$$c_j = \tilde{c}_j + C_j + \tilde{v}_j \tilde{\mu}_j^2 - \frac{(\tilde{v}_j \tilde{\mu}_j + U_j)^2}{\tilde{v}_j + V_j}, \quad (28)$$

$$v_j = \tilde{v}_j + V_j, \quad (29)$$

$$\mu_j = \frac{\tilde{v}_j \tilde{\mu}_j + U_j}{\tilde{v}_j + V_j}, \quad (30)$$

$$(31)$$

where hyperparameters describing the prior distribution are again indicated by  $\tilde{\cdot}$ , and the (conjugate) prior distributions are recovered by setting  $M_j = C_j = U_j = V_j = 0$ . These data-

dependent terms are given by

$$M_j = \sum_{t=2}^T \langle \delta_{s_t, j} \rangle, \quad (32)$$

$$C_j = \sum_{t=2}^T \langle \delta_{s_t, j} \rangle \vec{x}_t^2, \quad (33)$$

$$V_j = \sum_{t=2}^T \langle \delta_{s_t, j} \rangle \vec{x}_{t-1}^2. \quad (34)$$

$$U_j = \sum_{t=2}^T \langle \delta_{s_t, j} \rangle \vec{x}_t \cdot \vec{x}_{t-1}. \quad (35)$$

Some useful expectation values for future reference are

$$\langle K_j \rangle_{q(\vec{B}, \vec{K})} = \mu_j, \quad (36)$$

$$\text{Var}[K_j]_{q(\vec{B}, \vec{K})} = \frac{c_j}{2v_j(n_j - \frac{1}{2})}. \quad (37)$$

$$\langle B_j \rangle_{q(\vec{B}, \vec{K})} = \frac{n_j + \frac{1}{2}}{c_j}, \quad (38)$$

$$\text{Var}[B_j]_{q(\vec{B}, \vec{K})} = \frac{n_j + \frac{1}{2}}{c_j^2}, \quad (39)$$

$$\langle \ln B_j \rangle_{q(\vec{B}, \vec{K})} = \psi(n_j + \frac{1}{2}) - \ln c_j, \quad (40)$$

$$\langle B_j K_j^2 \rangle_{q(\vec{B}, \vec{K})} = \frac{1}{2v_j} + \mu_j^2 \frac{n_j + \frac{1}{2}}{c_j}, \quad (41)$$

$$\langle B_j K_j \rangle_{q(\vec{B}, \vec{K})} = \mu_j \frac{n_j + \frac{1}{2}}{c_j}, \quad (42)$$

In addition to being needed during the algorithm, these averages can be used to translate prior knowledge in terms of means and standard deviations of  $K_j, B_j$  into prior parameters  $\tilde{n}_j, \tilde{c}_j, \tilde{\mu}_j, \tilde{v}_j$

### 3.2.2 Hidden state distribution

The variational distribution has a simple form,

$$\ln q(s_{1:T}) = -\ln Z + \sum_{t=1}^T \ln h_{s_t}(t) + \sum_{t=2}^T \ln J_{s_{t-1}, s_t}, \quad (43)$$

i.e., an initial state distribution, a point-wise term that depends on the initial conditions and the data, and a transition probability. The point-wise

$$\ln q(s_{1:T}) = -\ln Z + \sum_{t=1}^T \ln h_{s_t}(t) + \sum_{t=2}^T \ln J_{s_{t-1}, s_t}, \quad (44)$$

i.e., an initial state distribution, point-wise terms that depends on the initial conditions and the data, and transition terms. The mathematical form of this expression is the same as encountered in maximum-likelihood optimization of hidden Markov Models, and hence the normalization constant and expectation values needed for the parameter update equations can be computed by the Baum-Welch algorithm (12), which resembles the transfer matrix solution for spin models in statistical physics.

Similarly, and the most likely sequence of hidden states can be computed by the Viterbi algorithm (13).

Specifically, the initial term is given by

$$\ln h_j(1) = \langle p(s_1 = j | \vec{\pi}) \rangle_{q(\vec{\pi})} = \psi(w_j^{(\vec{\pi})}) - \psi(w_0^{(\vec{\pi})}), \quad (45)$$

the point-wise contributions for  $t > 1$  are

$$\begin{aligned} \ln h_j(t) = & \psi\left(n_j + \frac{1}{2}\right) - \ln(\pi c_j) - \frac{\vec{x}_{t-1}^2}{2v_j} \\ & - \frac{n_j + \frac{1}{2}}{c_j} \left( \vec{x}_{t-1}^2 \left( \mu_j - \frac{\vec{x}_t \cdot \vec{x}_{t-1}}{\vec{x}_{t-1}^2} \right)^2 \right. \\ & \left. + \vec{x}_t^2 - \frac{(\vec{x}_t \cdot \vec{x}_{t-1})^2}{\vec{x}_{t-1}^2} \right), \end{aligned} \quad (46)$$

and the transition terms are given by

$$\ln J_{ji} = \psi(w_{j,i}^{(\mathbf{A})}) - \psi\left(\sum_{k=1}^N w_{j,k}^{(\mathbf{A})}\right). \quad (47)$$

### 3.3 VBEM iterations and model search

The iterative optimization of the variational distributions are done as follows. To start with, an initial guess for the variational parameter distributions are generated. We then alternate between VBE step, in which we construct the hidden state distribution and compute the averages  $\langle \delta_{j,s_t} \rangle_{q(s)}$  and  $\langle \delta_{j,s_t} \delta_{k,s_{t+1}} \rangle_{q(s)}$  in a Baum-Welch forward-backward sweep, and a VBM step, in which we use these averages to update the parameter variational distributions, until the lower bound converges.

The variational approach has the additional useful tendency to penalizing overfitting already during the VBEM iterations, by depopulating superfluous states (3, 10, 11). We exploit this property by using a greedy search algorithm to explore the model space. The basic strategy is to start by fitting a model with many states from random initial conditions, and then exploring less complex models by gradually removing the least populated states. This saves computing time by supplying good initial guesses for the

low complexity models (which therefore converge quickly), and by lowering the number of independent restarts, since it is easier to construct a good initial guess for a model with many states.

### 3.4 The lower bound

has an especially simple form just after the VBE step (3, 10, 11), given by the normalization constant  $\ln Z$  of the variational hidden state distribution, minus the Kullback-Leibler divergences between the variational and prior parameter distributions,

$$\begin{aligned} F = \ln Z - & \int d\vec{\pi} q(\vec{\pi}) \ln \frac{q(\vec{\pi})}{p(\vec{\pi})} \\ & - \sum_{j=1}^N \left[ \int d^N A_{j,:} q(A_{j,:}) \ln \frac{q(A_{j,:})}{p_0(A_{j,:})} \right. \\ & \left. + \int dB_j dK_j q(B_j, K_j) \ln \frac{q(B_j, K_j)}{p_0(B_j, K_j)} \right]. \end{aligned} \quad (48)$$

The Kullback-Leibler terms can be expressed in terms of the expectation values computed above. For the initial state distribution, we get

$$\begin{aligned} \int d\vec{\pi} q(\vec{\pi}) \ln \frac{q(\vec{\pi})}{p_0(\vec{\pi})} = & \ln \tilde{w}_0^{(\vec{\pi})} - \psi(\tilde{w}_0^{(\vec{\pi})}) - \frac{1}{\tilde{w}_0^{(\vec{\pi})}} \\ & + \sum_{j=1}^N \left[ (w_j^{(\vec{\pi})} - \tilde{w}_j^{(\vec{\pi})}) \psi(w_j^{(\vec{\pi})}) - \ln \frac{\Gamma(w_j^{(\vec{\pi})})}{\Gamma(\tilde{w}_j^{(\vec{\pi})})} \right]. \end{aligned} \quad (49)$$

To get this simple form, we used that  $w_0^{(\vec{\pi})} = 1 + \tilde{w}_0^{(\vec{\pi})}$  (since  $\sum_j \langle \delta_{j,s_1} \rangle = 1$ ), and the identities  $\Gamma(x+1) = x\Gamma(x)$  and  $\psi(x+1) = \psi(x) + \frac{1}{x}$ . Furthermore, each row of the transition probability



matrix contributes

$$\begin{aligned}
& \int d^N A_{j,:} q(A_{j,:}) \ln \frac{q(A_{j,:})}{p_0(A_{j,:})} \\
&= \ln \frac{\Gamma(w_{j0}^{(\mathbf{A})})}{\Gamma(\tilde{w}_{j0}^{(\mathbf{A})})} - (w_{j0}^{(\mathbf{A})} - \tilde{w}_{j0}^{(\mathbf{A})}) \psi(w_{j0}^{(\mathbf{A})}) \\
&- \sum_{k=1}^N \left[ \ln \frac{\Gamma(w_{jk}^{(\mathbf{A})})}{\Gamma(\tilde{w}_{jk}^{(\mathbf{A})})} - (w_{jk}^{(\mathbf{A})} - \tilde{w}_{jk}^{(\mathbf{A})}) \psi(w_{jk}^{(\mathbf{A})}) \right].
\end{aligned} \tag{50}$$

Finally, the emission parameter of each state contributes

$$\begin{aligned}
& \int dB_j \int d^N K_j q(B_j, K_j) \ln \frac{q(B_j, K_j)}{p(B_j, K_j)} = \dots \\
&= -\frac{n_j + \frac{1}{2}}{c_j} \left( c_j - \tilde{c}_j - \tilde{v}_j (\mu_j - \tilde{\mu}_j)^2 \right) \\
&+ \frac{1}{2} \ln \frac{v_j}{\tilde{v}_j} + \left( \tilde{n}_j + \frac{1}{2} \right) \ln \frac{c_j}{\tilde{c}_j} - \ln \frac{\Gamma(n_j + \frac{1}{2})}{\Gamma(\tilde{n}_j + \frac{1}{2})} \\
&+ (n_j - \tilde{n}_j) \psi(n_j + \frac{1}{2}) + \frac{\tilde{v}_j}{2v_j} - \frac{1}{2}.
\end{aligned} \tag{51}$$

### 3.5 Specification of prior distributions

**Emission parameters** We specify priors for the emission parameters  $K, B$  in terms of the mean values  $\langle K_j \rangle, \langle B_j \rangle$ , the standard deviation of  $K_j$ , and the number of pseudo-counts  $\tilde{n}_j$ , and then solve for the remaining hyper-parameters  $\tilde{c}_j, \tilde{\mu}_j, \tilde{v}_j$  using Eqs. (36-39). We take the same hyperparameters for all states, and independent of the number of states  $N$ . Priors for spurious states are specified analogously.

**Initial state** The initial state probability vector has a Dirichlet prior with weights  $\tilde{w}^{(\vec{\pi})}$  (see Eq. (15)). We choose a constant total strength  $f^{(\vec{\pi})}$ , i.e.,

$$\tilde{w}_j^{(\vec{\pi})} = f^{(\vec{\pi})} / N. \tag{52}$$

**Transition matrix** The prior for the transition matrix is independent Dirichlet distributions for each row (see Eq. (17)), with pseudo-count matrix  $\tilde{w}_{ij}^{(\mathbf{A})}$ . Following Persson et al. (10), we parameterize this prior in terms of an expected mean dwell time and an overall number of pseudocounts (prior strength) for each hidden state. To make the definition invariant under changes of sampling time, we specify the strength  $t_A$  and prior mean dwell time  $t_D$  in time units. We then construct a transition *rate* matrix  $Q$  with mean dwell time  $t_D$ ,

$$Q_{ij} = \frac{1}{t_D} \left( -\delta_{ij} + \frac{1 - \delta_{ij}}{N - 1} \right), \tag{53}$$

and construct the pseudo-counts based on the transition probability propagator per unit time step,

$$\tilde{w}_{ij}^{(\mathbf{A})} = \frac{t_A f_{\text{sample}}}{n_{\text{downsample}}} e^{\Delta t Q}. \tag{54}$$

This expression uses the downsampled timestep  $\Delta t = n_{\text{downsample}} / f_{\text{sample}}$ , where  $f_{\text{sample}}$  is the sampling frequency (30 Hz in our case), and  $n_{\text{downsample}}$  is the downsampling factor. Numerical experiments by Persson et al. (10) show that choosing the strength too low compared to the mean dwell time produces a bias towards sparse transition matrices. This is often not desirable, and we use  $t_D = 1$  s and  $t_A = 5$  s as default parameters, and also throughout our work (14).

Details on how to specify parameters for prior distributions are given in table 1.

### 3.6 Model search

To find the best number of states, we adopt a greedy search approach with multiple restarts, similar to the one used in vbSPT (10). The

variable	comment	Eq.
B0	$\langle B_j \rangle$	(38)
fB	$\tilde{n}_j$	(27)
K0	$\tilde{\mu}_j$	(30)
Kstd	$\sqrt{\text{Var} K_j}$	(37)
KBscaling	Possibility to specify an obsolete state-dependence of the prior parameters for $K_j$ , $B_j$ . We recommend the default value 2.	
	Bc0, fBc, Kc0, and KcStd are spurious state parameters, with the same interpretation as for genuine states.	
fPi	$f^{(\vec{\pi})}$	(52)
tD	Mean prior dwell time $t_D$ . Default 1 s.	(53)
tA	$\mathbf{A}$ prior strength $t_A$ . Default 5 s.	(54)
tDc0	Mean prior dwell time of the genuine indicator state $c_t = 1$ . Default 10 s.	

Table 1: Specifying prior parameters in runinput files. See also the documentation of VB7\_priorParent, and the example runinput files.

basic idea is to exploit the fact that the variational Bayes EM iterations tend to sense overfitting by making superfluous states unpopulated. Each search sweep thus starts by converging the largest model to be considered, and then gradually removes low-occupancy states until the lower bound  $F$  does not improve any more. We also attempt to reconverge each size with some extra pseudo-counts added to the transition matrix to avoid getting stuck in non-interconverting models.

Initial conditions for the largest model are generated randomly, partly based on user-specified intervals for the  $K$  and  $B$  parameters, and partly using the actual data to find high-occupancy values of  $K, B$ . For the transition matrices, we initialize using a mean dwell time, similar as for the prior distributions, but with higher total strength. Details are given in the documentation of VB7\_initialGuess\_KBregion.m, and in the example runinput file.

## 4 Extensions

### 4.1 Factorial model

The above algorithm is readily extended to treat the factorial model where genuine ( $s_t$ ) and spurious ( $c_t$ ) states are separated into two different hidden processes, with  $c_t = 1$  indicating a genuine state (bead motion described by genuine parameters  $K_{s_t}, B_{s_t}$ ), and  $c_t > 1$  indicating a spurious state (bead motion described by genuine parameters  $\hat{K}_{c_t}, \hat{B}_{c_t}$ ).

The bead motion/emission parameters are exactly analogous to genuine state parameters in the simple HMM above. For the transitions, we assume that genuine states evolve independently, and likewise for spurious states ( $c_t > 1$ ), but that the transitions from genuine to spurious

$c_t = 1 \rightarrow c_{t+1} > 1$  depends on the genuine state (see main text (14)). This leads to the following model:

$$p(s_{t+1}, c_{t+1} | s_t, c_t) = p(s_{t+1} | s_t) p(c_{t+1} | s_t, c_t), \quad (55)$$

with  $p(s_{t+1} | s_t) = A_{s_t s_{t+1}}$  as earlier, and

$$p(c_{t+1} | s_t, c_t) = \begin{cases} \hat{A}_{s_t c_{t+1}}, & \text{if } c_t = 1, \\ \hat{R}_{c_t c_{t+1}}, & \text{if } c_t > 1, \end{cases} \quad (56)$$

We implemented a brute force variational treatment of this extended model, where we define new composite hidden states  $\hat{s}_t = (s_t, c_t)$  and modify the simple vb algorithm to run this composite model, which mainly involves book-keeping to update the transition count matrices for  $\hat{A}$ ,  $\hat{R}$ . This has a significant computational cost, since a simple model with  $N_{gen.}$  genuine states and  $N_{sp.}$  spurious ones gets  $N_{gen.} \times (1 + N_{sp.})$  states after conversion. However, since we do not perform exhaustive model search in this representation and can utilize the simpler model to make good initial guesses, this is not a significant problem.

**Priors** We take the priors for the spurious states to be as similar to the genuine state prior as possible. Thus, we parameterize the priors for  $\hat{K}_j, \hat{B}_j$  in the same way as for the genuine state parameters. For the spurious-spurious transitions  $\hat{\mathbf{A}}$ , we use the same prior mean dwell time and strength as for the genuine transition matrix  $\mathbf{A}$ . For the genuine $\rightarrow$ spurious transition matrix  $\hat{\mathbf{R}}$ , we use the same strength, but a different (longer) mean life-time for the state  $c_t = 1$ , since spurious events are presumably rare. Further details: table 1, the example runinput files, and VB7\_priorParent.

## 4.2 Empirical Bayes update equations

The empirical Bayes update equations optimizes the lower bound with respect to the hyperparameters in the prior distribution. This means optimizing sums of Kullback-Leibler divergence terms.

The initial state probability, and the rows of the transition probability matrix, are both Dirichlet distributed. Thus, for  $M$  trajectories with Dirichlet parameters  $u_j^{(i)}$ ,  $i = 1, 2, \dots, M$ , and hyperparameters  $\tilde{u}_j$  ( $u = w^{(\tilde{\pi})}, u^{(\mathbf{A})}$ ), we need to solve

$$\frac{d}{d\tilde{u}_j} \sum_i \left( \ln \frac{\Gamma(u_0^{(i)})}{\Gamma(\tilde{u}_0)} - (u_0^{(i)} - \tilde{u}_0) \psi(u_0^{(i)}) \right) - \sum_{k=1}^N \left[ \ln \frac{\Gamma(u_k^{(i)})}{\Gamma(\tilde{u}_k^{(i)})} - (u_k^{(i)} - \tilde{u}_k) \psi(u_k^{(i)}) \right] = 0, \quad (57)$$

where  $u_0^{(i)} = \sum_k u_k^{(i)}$  and similar for  $\tilde{u}_0$ . This leads to the update equations

$$\psi(\tilde{u}_0) - \psi(\tilde{u}_j) = \frac{1}{M} \sum_i \left( \psi(u_0^{(i)}) - \psi(u_j^{(i)}) \right). \quad (58)$$

A numerical solution turned out to be easier using the variables  $\tilde{U}_j = \ln \tilde{u}_j$  (to numerically enforce  $\tilde{u}_j > 0$ ).

For the emission parameters, the update equations are instead derived from minimizing

Eq. (51) summed over  $M$  trajectories,

$$f_{KB} = \sum_i \left( -\frac{n^{(i)} + \frac{1}{2}}{c^{(i)}} \left( c^{(i)} - \tilde{c} - \tilde{v}(\mu^{(i)} - \tilde{\mu})^2 \right) + \frac{1}{2} \ln \frac{v^{(i)}}{\tilde{v}} + \left( \tilde{n} + \frac{1}{2} \right) \ln \frac{c^{(i)}}{\tilde{c}} - \ln \frac{\Gamma(n^{(i)} + \frac{1}{2})}{\Gamma(\tilde{n} + \frac{1}{2})} + (n^{(i)} - \tilde{n})\psi\left(n^{(i)} + \frac{1}{2}\right) + \frac{1}{2} \left( \frac{\tilde{v}}{v^{(i)}} - 1 \right) \right). \quad (59)$$

Minimizing with respect to  $\tilde{\mu}$  and  $\tilde{v}$  leads to

$$\tilde{\mu} = \frac{1}{M} \sum_i \mu^{(i)}, \quad (60)$$

$$\frac{1}{\tilde{v}} = \frac{1}{M} \sum_i \left( \frac{1}{v^{(i)}} + 2(\tilde{\mu} - \mu^{(i)})^2 \right). \quad (61)$$

The remaining  $\tilde{c}$  and  $\tilde{n}$  lead to

$$\frac{\tilde{n} + \frac{1}{2}}{\tilde{c}} = \frac{1}{M} \sum_i \frac{n^{(i)} + \frac{1}{2}}{c^{(i)}}, \quad (62)$$

$$\ln \tilde{c} - \psi\left(\tilde{n} + \frac{1}{2}\right) = \frac{1}{M} \sum_i \left( \ln c^{(i)} - \psi\left(n^{(i)} + \frac{1}{2}\right) \right), \quad (63)$$

which we solve numerically. This gets easier by defining  $\alpha = \frac{\tilde{n} + \frac{1}{2}}{\tilde{c}}$ , then solve the second equation for  $\tilde{c}$  numerically, and finally compute  $\tilde{n} = \alpha \tilde{c} - \frac{1}{2}$ .

### 4.3 Doing empirical Bayes

Our empirical Bayes analysis of multiple models was not implemented as part of our analysis pipeline, but instead ran using tailored scripts. These are not included with vbTPM, but the optimization procedures for the individual prior distributions are included, in `VB7_EBupdate_dirichlet`, and `VB7_EBupdate_KB`.

Table 2: Fields in a model object  $W$ .

field	
<b>W.N</b>	$N$ , number of (genuine) states.
<b>W.Nc</b>	Number of indicator states $\hat{N}+1$ . <code>W.Nc=1</code> means no spurious states, i.e., the simple HMM.
<b>W.F</b>	Lower bound $F$ .

## 5 Notation and symbols

vbSPT stores mathematical objects in matlab structures that contain parameters for the variational and prior distributions, and various other things. In this section we list some of them.

First, `VB7_batch_manage` with the `collect` option returns a filename, and cell vectors of structs that contain results of the model search for each trajectory, indexed as the filenames in the run-input file, e.g. `trj{k}{j}` contains the analysis result for `looping_filename{k}{j}` etc.

Most importantly, the `Wtrj` and `Wcal` fields are the converged model structs, whose content are detailed below (using `W` as the generic model name). In addition the columns of the arrays `NFtrj/NFcal` contain  $N, \hat{N}, F, iter$  for each model that was converged during the model search, and `iter` is the restart number that produced it. In `NFitrj/NFical`, the best model for each size is listed in the same way, with the last column indicating the rows in `NFtrj, NFcal` where these optimal models can be found.

Further details about the model structs are given in tables 2-5.

Table 3: Representation of variational distributions for genuine states in a model object named  $W$ .

field	symbol	Eq.
W.M.wPi(j)	$w_j^{(\vec{\pi})}$	(15)
W.PM.wPi(j)	$\tilde{w}_j^{(\vec{\pi})}$	
W.E.ds_1(j)	$\langle \delta_{j,s_1} \rangle$	
W.PM.wA(i,j)	$\tilde{w}_{ij}^{(\mathbf{A})}$	(17)
W.M.wA(i,j)	$w_{ij}^{(\mathbf{A})}$	
W.E.wA(i,j)	$\sum_{t=2}^T \langle \delta_{i,s_{t-1}} \delta_{j,s_t} \rangle$	
W.PM.n(j)	$\tilde{n}_j$	(27)
W.M.n(j)	$n_j$	(27)
W.E.M(j)	$M_j$	(32)
W.PM.c(j)	$\tilde{c}_j$	(28)
W.M.c(j)	$c_j$	(28)
W.E.C(j)	$C_j$	(33)
W.PM.v(j)	$\tilde{v}_j$	(29)
W.M.v(j)	$v_j$	(29)
W.E.V(j)	$V_j$	(34)
W.PM.mu(j)	$\tilde{\mu}_j$	(30)
W.M.mu(j)	$\mu_j$	(30)
W.E.U(j)	$U_j$	(35)

Table 4: Representation of variational distributions for indicator states  $c_t$  in a model object named  $W$ . Fields relating to the emission model (.n, .c, .v, .mu, etc.) have the same meaning as for the genuine states  $s_t$ , except that their first element is not used, since  $c_t = 1$  indicate a genuine state.

field	symbol	Eq.
W.Mc.wPi(j)	$w_j^{(\vec{\pi})}$	(??)
W.PMc.wPi(j)	$\tilde{w}_j^{(\vec{\pi})}$	
W.Ec.ds_1(j)	$\langle \delta_{j,c_1} \rangle$	
W.PMc.wA(j)		(??)
W.Mc.wA(j)		
W.Ec.wA(j)		
W.PMc.wR(i,j)		(??)
W.Mc.wR(i,j)		
W.Ec.wR(i,j)		

Table 5: Selected fields that characterize converged models (named W). The fields W.est and W.est2 are constructed by VB7\_VBEMiter.m (although W.est2 must be specifically requested), and fields not mentioned here can be looked up there. Averages are w.r.t. variational parameter distributions unless stated otherwise.

field	comment
W.est.sAverage	Occupation probability of genuine states $s_t$ , computed by classification, i.e., sAverage(j) proportional to $\sum_t \langle \delta_{j,s_t} \rangle$ .
W.est.cAverage	Occupation probability of indicator states $c_t$ , by classification.
W.est.sVisible	Occupation probability of genuine states $s_t$ , by classification that excludes spurious states, i.e., sAverage(j) proportional to $\sum_t \langle \delta_{j,s_t} \delta_{1,c_t} \rangle$ .
W.est.A	Mean transition probabilities for genuine states, $\langle \mathbf{A} \rangle_{q(\mathbf{A})}$ .
W.est.dA	Standard deviation of transition probability matrix $\mathbf{A}$ .
W.est.tD	Mean dwell times of genuine states in units of time, computed from the elements of $\langle \mathbf{A} \rangle$ .
W.est.lnQss	Log average transition probabilities, goes into $q(s_{1:T})$ .
	Corresponding averages for spurious state distributions are also computed, Ac, dAc, Rc, dRc, tStick, lnQcc, lnQsc, tStick, tUnstick.
W.est.sKaverage(j)	$\langle K_j \rangle = \mu_j$ .
W.est.sBaverage(j)	$\langle B_j \rangle$ .
W.est.sRMS(j)	$RMS_j = \sqrt{\langle \vec{x}_t^2   s_t = j \rangle} \approx (\langle B_j \rangle (1 - \langle K_j \rangle^2))^{-\frac{1}{2}}$ .
W.est.sTC(j)	Approx. correlation time $\tau_j \approx -\Delta t / \log \langle K_j \rangle$ , units of time.
W.est.sKstd(j)	Standard deviation of $K_j$ .
W.est.sBstd(j)	Standard deviation of $B_j$ .
W.est.cXXX	Corresponding properties of spurious states are named with $s \rightarrow c$ .
W.est2.qt	State occupancy probability for combined states $(s_t, c_t)$ . Use sMap and cMap to extract genuine/spurious occupancies, e.g., $p(s_t = j) = \text{sum}(\text{W.est2.qt}(t,:).*(\text{W.est.sMap}==j))$ .
W.est2.sMaxP(t)	Most likely genuine state at time $t$ .
W.est2.cMaxP(t)	Most likely indicator state at time $t$ .
W.est2.sViterbi	Viterbi path (most likely sequence of states) for $s_t$ .
W.est2.cViterbi	Viterbi path (most likely sequence of states) for $c_t$ .
	W.est2 also contains a few other intermediate fields from the VBEM iteration that are mainly good for debugging. This substructure is thus very bulky and somewhat expensive to compute, which is the reason computing it is optional.

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