Homework 6

Benjamin Noland

1. The following code generates 100 observations from the MA(1) model $Y_t = e_t + 0.7e_{t-1}$:

```
set.seed(1000)
y <- arima.sim(model = list(ma = c(0.7)), n = 100)</pre>
```

Let Y_1, \ldots, Y_{100} denote the observed time series, and let Σ denote the covariance matrix of $Y = (Y_1, \ldots, Y_{100})$.

a. The covariance matrix Σ is given by $\Sigma_{ij} = (\gamma_{|i-j|})$. Explicitly,

$$\Sigma_{ij} = \begin{cases} (1 + \theta_1^2 + \theta_2^2)\sigma_e^2 & \text{if } i = j\\ (-\theta_1 + \theta_1\theta_2)\sigma_e^2 & \text{if } |i - j| = 1\\ -\theta_2\sigma_e^2 & \text{if } |i - j| = 2\\ 0 & \text{otherwise} \end{cases}.$$

b. Assume the elements of the white noise process $\{e_t\}$ are drawn independently from a $N(0, \sigma_e^2)$ distribution (this is the case for the simulated data generated above). Then the sample $Y = (Y_1, \ldots, Y_{100})$ has a mean zero multivariate normal distribution with covariance matrix Σ . Let $y = (y_1, \ldots, y_{100})$ a denote realization of the sample Y. Then the likelihood function is given by

$$L(\theta_1, \theta_2 | y) = \frac{1}{\sqrt{(2\pi)^{100}|\Sigma|}} \exp\left(-\frac{1}{2}y^T \Sigma^{-1}y\right).$$

Maximizing $L(\theta_1, \theta_2|y)$ (or equivalently, the log-likelihood log $L(\theta_1, \theta_2|y)$) with respect to (θ_1, θ_2) yields a maximum likelihood estimate (MLE) $(\hat{\theta}_1, \hat{\theta}_2)$ of (θ_1, θ_2) .

c. The following code fits an MA(2) model to the data using maximum likelihood:

```
arima(y, order = c(0, 0, 2), method = "ML")
```

```
##
## Call:
## arima(x = y, order = c(0, 0, 2), method = "ML")
##
##
  Coefficients:
##
           ma1
                    ma2
                         intercept
                -0.184
                             0.011
##
         0.525
## s.e.
         0.107
                 0.110
                             0.134
##
## sigma^2 estimated as 0.995: log likelihood = -142, aic = 290
```

From the output, we see that the maximum likelihood estimates for θ_1 and θ_2 are $\hat{\theta}_1 = -0.5246$ and $\hat{\theta}_2 = 0.1835$, respectively (taking into account the differences in convention between R and the book with regard to the signs of the parameters).

- d. From the output in part (c), we see that $\hat{\theta}_1$ and $\hat{\theta}_2$ have approximate standard errors of 0.1072 and 0.1101, respectively.
- 2. (Cryer & Chan, Exercise 7.1) We can compute method of moments estimates $\hat{\phi}_1$ and $\hat{\phi}_2$ of ϕ_1 and ϕ_2 , respectively, by solving for $\hat{\phi}_1$ and $\hat{\phi}_2$ in the sample Yule-Walker equations:

$$r_1 = \hat{\phi}_1 + r_1 \hat{\phi}_2 \\ r_2 = r_1 \hat{\phi}_1 + \hat{\phi}_2.$$

We get

$$\hat{\phi}_1 = \frac{r_1(1-r_2)}{1-r_1^2} = \frac{r_2-r_1^2}{1-r_1^2} = \frac{r_2-r_1^2}{1-r_1^2} = \frac{r_2-r_1^2}{1-r_1^2}$$

Using these estimates, we can get estimates $\hat{\theta}_0$ and $\hat{\sigma}_e^2$ of θ_0 and σ_e^2 , respectively:

$$\hat{\theta}_0 = \bar{Y}(1 - \hat{\phi}_1 - \hat{\phi}_2) =$$

$$\hat{\sigma}_e^2 = (1 - \hat{\phi}_1 r_1 - \hat{\phi}_2 r_2) s^2 =$$

3. (Cryer & Chan, Exercise 7.11) The following code simulates the MA(1) process in question:

```
set.seed(1000)
n <- 48
theta <- -0.6
y <- arima.sim(model = list(ma = c(-theta)), n = n)</pre>
```

a. The following code fits an MA(1) model to the data simulated above using maximum likelihood:

```
##
## Call:
   arima(x = y, order = c(0, 0, 1), method = "ML")
##
## Coefficients:
##
           ma1
                intercept
                   -0.228
##
         0.542
##
         0.124
                    0.206
  s.e.
##
## sigma^2 estimated as 0.868: log likelihood = -64.9, aic = 134
```

arima(y, order = c(0, 0, 1), method = "ML")

From the output we see that the maximum likelihood estimate of θ is $\hat{\theta} = -0.5422$.

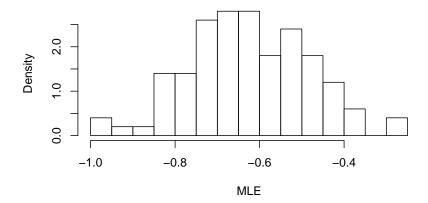
b. The following code repeatedly simulates the same series and collects the maximum likelihood estimate of θ on each trial:

```
N <- 100
mle <- numeric(N)
for (i in 1:N) {
    y <- arima.sim(model = list(ma = c(-theta)), n = n)
    fit <- arima(y, order = c(0, 0, 1), method = "ML")
    mle[[i]] <- -fit$coef[["ma1"]]
}</pre>
```

c. The following displays the approximate sampling distribution of the MLE $\hat{\theta}$ based on the simulation in part (b):

```
hist(mle, freq = FALSE, breaks = 20,
    main = "Sampling distribution of MLE", xlab = "MLE")
```

Sampling distribution of MLE



- d. The true parameter value is $\theta = -0.6$. The approximate mean of the MLE sampling distribution based on the simulation in part (b) is $\bar{\hat{\theta}} = -0.627$, with approximate variance $\widehat{\text{Var}(\hat{\theta})} = 0.021$. Since the mean is close to the true value, and the variance is small, the estimates appear to be approximately unbiased (i.e., approximately centered around the true value θ).
- e. The approximate variance of the sampling distribution is $\text{Var}(\hat{\theta}) = 0.021$. Large sample theory predicts that for large n, $\text{Var}(\hat{\theta}) = (1 \theta^2)/n = 0.013$. These two values are relatively close.
- 4. (Cryer & Chan, Exercise 7.31) The following code simulates the time series in question:

```
n <- 48
phi <- 0.7
y <- arima.sim(model = list(ar = c(phi)), n = n)</pre>
```

5. (Cryer & Chan, Exercise 8.9)