

Homework 4

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Ergodic Theorem

Exercise 1

Since $P(A) = P(B) = 1$, $P(A^c) = P(B^c) = 0$, and hence

$$P(A^c \cup B^c) \leq P(A^c) + P(B^c) = 0.$$

Thus, upon taking complements, we find

$$P(A \cap B) = 1 - P(A^c \cup B^c) \geq 1 - 0 = 1,$$

so that $P(A^c \cup B^c) = 0$.

Exercise 2

Define a map $h : \cup_{m=0}^{\infty} \mathcal{X}^{m+1} \rightarrow \mathbb{R}$ by

$$h(x_0, \dots, x_t) = \sum_{i=0}^t f(x_i) \quad \text{for any } x_0, \dots, x_t \in \cup_{m=0}^{\infty} \mathcal{X}^{m+1}. \quad (1)$$

Then for any $k \geq 0$, we can write $S_k = h(X_{\tau_x^k}, X_{\tau_x^k+1}, \dots, X_{\tau_x^{k+1}-1})$. Roughly speaking, the fact that the S_k 's are IID follows from the fact that the chain starts afresh at each return time τ_x^k . This can probably be formalized using Exercises 7 and 8 of the Strong Markov Property question, along with the map defined in (1) (but I don't have time to figure this out).

Exercise 3

First, note that $E_x(|S_0|) < \infty$. To see this, note that since the state space \mathcal{X} is finite, we can set

$$M = \max_{x \in \mathcal{X}} |f(x)|.$$

Thus,

$$\mathbb{E}_x(|S_0|) = \mathbb{E}_x \left(\sum_{t=0}^{\tau_x^1-1} f(X_t) \right) \leq \mathbb{E}_x \left(\sum_{t=0}^{\tau_x^1-1} M \right) = \mathbb{E}_x(M\tau)x^1 = M\mathbb{E}_x(\tau_x^1) < \infty$$

where the fact that $\mathbb{E}_x(\tau_x^1) < \infty$ follows from Lemma 1.13, since the chain is irreducible. Note that for any $n \geq 0$ we have the relation

$$\sum_{k=0}^{n-1} S_k = \sum_{k=0}^{n-1} \sum_{t=\tau_x^k}^{\tau_x^{k+1}-1} f(X_t) = \sum_{t=0}^{\tau_x^n-1} f(X_t).$$

Thus, since the S_k 's are IID, the Strong Law of Large Numbers yields

$$\mathbb{P}_x \left\{ \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{t=0}^{\tau_x^n-1} f(X_t) = \mathbb{E}_x(S_0) \right\} = 1.$$

In addition, note that for any $k \geq 0$,

$$\tau_x^{k+1} - \tau_x^k = \inf\{t - \tau_x^k > 0 : X_{t-\tau_x^k} = x\},$$

so that $(\tau_x^{k+1} - \tau_x^k)$ has the same distribution as τ_x^1 , since the chain starts afresh at time τ_x^k . In addition, this implies that the differences $(\tau_x^{k+1} - \tau_x^k)$ are mutually independent, and hence IID. Moreover, for any $n \geq 1$,

$$\tau_x^n = \sum_{k=0}^{n-1} (\tau_x^{k+1} - \tau_x^k),$$

and that $\mathbb{E}_x(\tau_x^1 - \tau_x^0) = \mathbb{E}_x(\tau_x^1) < \infty$ by Lemma 1.13, since the chain is irreducible. So the Strong Law of Large Numbers yields

$$\mathbb{P}_x \left\{ \lim_{n \rightarrow \infty} \frac{1}{n} \tau_x^n = \mathbb{E}_x(\tau_x^1) \right\} = 1.$$

Thus, applying Exercise 1 of this problem, we get

$$\begin{aligned} 1 &= \mathbb{P}_x \left\{ \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{t=0}^{\tau_x^n-1} f(X_t) = \mathbb{E}_x(S_0), \lim_{n \rightarrow \infty} \frac{1}{n} \tau_x^n = \mathbb{E}_x(\tau_x^1) \right\} \\ &\leq \mathbb{P}_x \left\{ \lim_{n \rightarrow \infty} \frac{1}{\tau_x^n} \sum_{t=0}^{\tau_x^n-1} f(X_t) = \frac{\mathbb{E}_x(S_0)}{\mathbb{E}_x(\tau_x^1)} \right\}, \end{aligned}$$

so that

$$\mathbb{P}_x \left\{ \lim_{n \rightarrow \infty} \frac{1}{\tau_x^n} \sum_{t=0}^{\tau_x^n-1} f(X_t) = L \right\} = 1 \quad \text{where} \quad L = \frac{\mathbb{E}_x(S_0)}{\mathbb{E}_x(\tau_x^1)}$$

Exercise 4

First, note that

$$\begin{aligned} \mathbb{E}_x(S_0) &= \mathbb{E}_x \left(\sum_{t=0}^{\tau_x^1-1} f(X_t) \right) = \mathbb{E}_x \left(\sum_{t=0}^{\tau_x^1-1} \sum_{y \in \mathcal{X}} I\{X_t = y\} f(y) \right) \\ &= \mathbb{E}_x \left(\sum_{y \in \mathcal{X}} f(y) \sum_{t=0}^{\tau_x^1-1} I\{X_t = y\} \right) = \sum_{y \in \mathcal{X}} f(y) \mathbb{E}_x \left(\sum_{t=0}^{\tau_x^1-1} I\{X_t = y\} \right). \end{aligned}$$

But note that $\sum_{t=0}^{\tau_x^1-1} I\{X_t = y\}$ is the number of visits to state y before returning to state x . Thus,

$$\tilde{\pi}(y) = \mathbb{E}_x \left(\sum_{t=0}^{\tau_x^1-1} I\{X_t = y\} \right),$$

where $\tilde{\pi}$ is the distribution defined on page 11 of the book. Therefore, since $\mathbb{E}_x(\tau_x^1) < \infty$ by Lemma 1.13, and since π is the unique stationary distribution of the chain by Corollary 1.17, Proposition 1.14(ii) implies that

$$\pi = \frac{\tilde{\pi}}{\mathbb{E}_x(\tau_x^1)}.$$

Thus,

$$\mathbb{E}_x(S_0) = \sum_{y \in \mathcal{X}} f(y) \tilde{\pi}(y) = \sum_{y \in \mathcal{X}} f(y) \pi(y) \mathbb{E}_x(\tau_x^1),$$

so that

$$L = \frac{\mathbb{E}_x(S_0)}{\mathbb{E}_x(\tau_x^1)} = \sum_{y \in \mathcal{X}} f(y) \pi(y) = E_\pi(f).$$

Exercise 5

For every integer $T \geq 1$, define

$$A_T = \frac{1}{T} \sum_{i=1}^T a_i.$$

Then for any choice of $T \geq 1$, since $T_k \rightarrow \infty$ as $k \rightarrow \infty$, there exists $k \geq 1$ satisfying $T_k \leq T \leq T_{k+1}$. Thus we can write

$$A_T = \frac{1}{T} \sum_{i=1}^T a_i = \frac{T_k}{T} \left(\frac{1}{T_k} \sum_{i=1}^{T_k} a_i \right) + \frac{1}{T} \sum_{i=T_k+1}^T a_i = \frac{T_k}{T} A_{T_k} + \frac{1}{T} \sum_{i=T_k+1}^T a_i.$$

I will consider each term in this expression separately. Since $T_k \leq T \leq T_{k+1}$,

$$A_{T_k} = \frac{T_k}{T} A_{T_k} \leq \frac{T_k}{T} A_{T_k} \leq \frac{T_k}{T_{k+1}} A_{T_k}.$$

Thus, since $A_{T_k} \rightarrow L$ and $T_k/T_{k+1} \rightarrow 1$ as $k \rightarrow \infty$, we see that

$$\lim_{k \rightarrow \infty} \frac{T_k}{T} A_{T_k} = L,$$

or equivalently, since $T_k \leq T \leq T_{k+1}$,

$$\lim_{T \rightarrow \infty} \frac{T_k}{T} A_{T_k} = L.$$

As for the second term, since $a_i \in [-M, M]$ for every $i \geq 1$,

$$\begin{aligned} \left| \frac{1}{T} \sum_{i=T_k+1}^T a_i \right| &\leq \frac{1}{T} \sum_{i=T_k+1}^T |a_i| \leq \frac{1}{T} \sum_{i=T_k+1}^T M = \frac{M}{T} (T - T_k) \leq \frac{M}{T} (T_{k+1} - T_k) \\ &\leq \frac{M}{T_k} (T_{k+1} - T_k) = M \left(\frac{T_{k+1}}{T_k} - 1 \right) \rightarrow 0 \quad \text{as } T \rightarrow \infty \end{aligned}$$

since because $T_k/T_{k+1} \rightarrow 1$ as $T \rightarrow \infty$, $T_{k+1}/T_k \rightarrow 1$ as $T \rightarrow \infty$ as well. Therefore

$$\lim_{T \rightarrow \infty} \frac{1}{T} \sum_{i=T_k+1}^T a_i = 0,$$

so that

$$\lim_{T \rightarrow \infty} \frac{1}{T} \sum_{i=1}^T a_i = L.$$

Exercise 6

To see that $\tau_x^n \geq n$ for every $n \geq 0$, I will proceed by induction on n . When $n = 0$, we have

$$\tau_x^0 = 0 \geq 0$$

by definition. Now let $n > 0$, and suppose that $\tau_x^n \geq n$. Then

$$\tau_x^{n+1} = \inf\{t > \tau_x^n : X_t = x\} = \inf\{t \geq \tau_x^n + 1 : X_t = x\} \geq \tau_x^n + 1 \geq n + 1$$

by the inductive hypotheses. Hence $\tau_x^n \geq n$ for every $n \geq 0$ by induction. In particular, this result implies that

$$P_x \left\{ \lim_{n \rightarrow \infty} \tau_x^n = \infty \right\} = 1.$$

Moreover, note that for every $n \geq 0$,

$$\frac{\tau_x^n}{\tau_x^{n+1}} < \frac{\tau_x^n + \tau_x^1}{\tau_x^{n+1} + \tau_x^1}, \quad (1)$$

where the ratio $(\tau_x^n + \tau_x^1)/(\tau_x^{n+1} + \tau_x^1)$ has the same distribution as $\tau_x^{n+1}/\tau_x^{n+2}$, since the chain starts afresh at each of the times τ_x^n and τ_x^{n+1} . Thus, since $\tau_x^n/\tau_x^{n+1} \leq 1$ by definition, it follows from the inequality (1) that the ratio τ_x^n/τ_x^{n+1} can be made arbitrarily close to 1 if n is chosen large enough. Therefore,

$$\mathbb{P}_x \left\{ \lim_{n \rightarrow \infty} \frac{\tau_x^n}{\tau_x^{n+1}} = 1 \right\} = 1.$$

Since the function f is bounded (specifically, $f : \mathcal{X} \rightarrow [-M, M]$ for some $M > 0$), it follows from Exercise 5 of this problem that

$$1 = \mathbb{P}_x \left\{ \lim_{n \rightarrow \infty} \frac{1}{\tau_x^n} \sum_{t=0}^{\tau_x^n - 1} f(X_t) = E_\pi(f) \right\} \leq \mathbb{P}_x \left\{ \lim_{n \rightarrow \infty} \frac{1}{T} \sum_{t=0}^{T-1} f(X_t) = E_\pi(f) \right\},$$

and the result follows.

Exercise 7

By Exercise 6 of this problem,

$$\begin{aligned} \mathbb{P}_\mu \left\{ \lim_{n \rightarrow \infty} \frac{1}{T} \sum_{t=0}^{T-1} f(X_t) = E_\pi(f) \right\} &= \sum_{x \in \mathcal{X}} \mathbb{P} \left\{ \lim_{n \rightarrow \infty} \frac{1}{T} \sum_{t=0}^{T-1} f(X_t) = E_\pi(f) \middle| X_0 = x \right\} \mathbb{P}\{X_0 = x\} \\ &= \sum_{x \in \mathcal{X}} \mathbb{P}_x \left\{ \lim_{n \rightarrow \infty} \frac{1}{T} \sum_{t=0}^{T-1} f(X_t) = E_\pi(f) \right\} \mu(x) \\ &= \sum_{x \in \mathcal{X}} \mu(x) = 1. \end{aligned}$$

Exercise 8

The map $f : \mathcal{X} \rightarrow \{0, 1\}$ is given by

$$f(y) = I\{y = x_0\} \quad \text{for any } y \in \mathcal{X}.$$

Thus,

$$E_\pi(f) = \sum_{y \in \mathcal{X}} \pi(y) f(y) = \sum_{y \in \mathcal{X}} \pi(y) I\{y = x_0\} = \pi(x_0).$$

So in this special case the result becomes

$$\mathbb{P}_x \left\{ \lim_{n \rightarrow \infty} \frac{1}{T} \sum_{t=0}^{T-1} I\{X_t = x_0\} = \pi(x_0) \right\} = 1$$

(i.e., we expect the long run proportion of time spent in state x_0 to be $\pi(x_0)$).

Random mapping representation

Exercises from Homework 2

Exercise 5

Using the given random mapping representation, we get

$$\begin{aligned} & P_\mu\{f(X_t, \dots, X_{t+k}) = 1 | X_t = x_t, p(X_0, \dots, X_t) = 1\} \\ &= P_\mu\{f(x_t, \dots, X_{t+k}) = 1 | p(X_0, \dots, x_t) = 1\} \\ &= P_\mu\{f(x_t, r(x_t, Z_{t+1}), \dots, r(X_{t+k-1}, Z_{t+k}) = 1 | p(r_0(Z_0), r(X_0, Z_0), \dots, x_t) = 1\}. \end{aligned}$$

But notice that the event $\{f(x_t, r(x_t, Z_{t+1}), \dots, r(X_{t+k-1}, Z_{t+k}) = 1\}$ depends solely upon Z_{t+1}, \dots, Z_{t+k} , and that the event $\{p(r_0(Z_0), r(X_0, Z_0), \dots, x_t) = 1\}$ depends solely upon Z_0, \dots, Z_{t-1} . Hence these two events are independent, and so we can write

$$\begin{aligned} & P_\mu\{f(x_t, r(x_t, Z_{t+1}), \dots, r(X_{t+k-1}, Z_{t+k}) = 1 | p(r_0(Z_0), r(X_0, Z_0), \dots, x_t) = 1\} \\ &= P_\mu\{f(x_t, r(x_t, Z_{t+1}), \dots, r(X_{t+k-1}, Z_{t+k}) = 1\} \\ &= P_\mu\{f(X_t, r(X_t, Z_{t+1}), \dots, r(X_{t+k-1}, Z_{t+k}) = 1 | X_t = x_t\} \\ &= P_\mu\{f(X_t, \dots, X_{t+k}) = 1 | X_t = x_t\} \\ &= P_{x_t}\{f(X_0, \dots, X_k) = 1\}, \end{aligned}$$

where the final equality follows from the time-homogeneity of the chain.

Exercise 6

Using the given random mapping representation, we get

$$\begin{aligned} & P_\mu\{f(X_t, \dots, X_{t+k}) = 1, p(X_0, \dots, X_t) = 1 | X_t = x_t\} \\ &= P_\mu\{f(x_t, \dots, X_{t+k}) = 1, p(X_0, \dots, x_t) = 1\} \\ &= P_\mu\{f(x_t, r(x_t, Z_{t+1}), \dots, r(X_{t+k-1}, Z_{t+k})) = 1, p(r_0(Z_0), r(X_0, Z_1), \dots, x_t) = 1\}. \end{aligned}$$

But notice that the event $\{f(x_t, r(x_t, Z_{t+1}), \dots, r(X_{t+k-1}, Z_{t+k})) = 1\}$ depends solely upon Z_{t+1}, \dots, Z_{t+k} , and that the event $\{p(r_0(Z_0), r(X_0, Z_1), \dots, x_t) = 1\}$ depends solely upon Z_0, \dots, Z_{t-1} . Hence these two events are independent, and so we can write

$$\begin{aligned} & P_\mu\{f(x_t, r(x_t, Z_{t+1}), \dots, r(X_{t+k-1}, Z_{t+k})) = 1, p(r_0(Z_0), r(X_0, Z_1), \dots, x_t) = 1\} \\ &= P_\mu\{f(x_t, r(x_t, Z_{t+1}), \dots, r(X_{t+k-1}, Z_{t+k})) = 1\} P_\mu\{p(r_0(Z_0), r(X_0, Z_1), \dots, x_t) = 1\} \\ &= P_\mu\{f(x_t, \dots, X_{t+k}) = 1\} P_\mu\{p(X_0, \dots, x_t) = 1\} \\ &= P_\mu\{f(X_t, \dots, X_{t+k}) = 1 | X_t = x_t\} P_\mu\{p(X_0, \dots, X_t) = 1 | X_t = x_t\} \\ &= P_{x_t}\{f(X_0, \dots, X_k) = 1\} P_\mu\{p(X_0, \dots, X_t) = 1 | X_t = x_t\}, \end{aligned}$$

where the final equality follows from the time-homogeneity of the chain.

Exercise 7

Using the given random mapping representation, we get

$$\begin{aligned} & P_\mu\{f(X_{t+m}, \dots, X_{t+m+k}) = 1, p(X_0, \dots, X_t) = 1|E\} \\ &= P_\mu\{f(x_{t+m}, r(x_{t+m}, Z_{t+m+1}), \dots, r(X_{t+m+k-1}, Z_{t+m+k})) = 1, \\ & \quad p(r_0(Z_0), r(X_0, Z_1), \dots, x_t) = 1|E\}. \end{aligned}$$

But notice that the event $\{f(x_{t+m}, r(x_{t+m}, Z_{t+m+1}), \dots, r(X_{t+m+k-1}, Z_{t+m+k})) = 1\}$ depends solely upon $Z_{t+m+1}, \dots, Z_{t+m+k}$, and that the event $\{p(r_0(Z_0), r(X_0, Z_1), \dots, x_t) = 1\}$ depends solely upon Z_0, \dots, Z_{t-1} . Hence these two events are independent under conditioning on E , and so we can write

$$\begin{aligned} & P_\mu\{f(x_{t+m}, r(x_{t+m}, Z_{t+m+1}), \dots, r(X_{t+m+k-1}, Z_{t+m+k})) = 1, \\ & \quad p(r_0(Z_0), r(X_0, Z_1), \dots, x_t) = 1|E\} \\ &= P_\mu\{f(x_{t+m}, r(x_{t+m}, Z_{t+m+1}), \dots, r(X_{t+m+k-1}, Z_{t+m+k})) = 1|E\} \\ & \quad \times P_\mu\{p(r_0(Z_0), r(X_0, Z_1), \dots, x_t) = 1|E\} \\ &= P_\mu\{f(X_{t+m}, \dots, X_{t+m+k}) = 1|E\} P_\mu\{p(X_0, \dots, X_t) = 1|E\} \\ &= P_\mu\{f(X_{t+m}, \dots, X_{t+m+k}) = 1|X_{t+m} = x_{t+m}\} P_\mu\{p(X_0, \dots, X_t) = 1|X_t = x_t\} \\ &= P_{x_{t+m}}\{f(X_0, \dots, X_k) = 1\} P_\mu\{p(X_0, \dots, X_t) = 1|X_t = x_t\}, \end{aligned}$$

where the final equality follows from the time-homogeneity of the chain. As before, the result still holds if E is replaced by $\{X_t = x_t, X_{t+m} = x_{t+m}\}$.

Exercises from Homework 3

Exercise 5

(Ran out of time).

Exercise 7

(Ran out of time).

Exercises from the book

Exercise 3.1

We want to show that

$$\pi(x)P(x, y) = \pi(y)P(y, x) \quad \text{for every } x, y \in \mathcal{X}.$$

Let $x, y \in \mathcal{X}$. When $x = y$ the relation holds trivially, so assume $x \neq y$. Then

$$\begin{aligned}\pi(x)P(x, y) &= \pi(x)\Psi(x, y) \left[\frac{\pi(y)\Psi(y, x)}{\pi(x)\Psi(x, y)} \wedge 1 \right] = [\pi(y)\Psi(y, x) \wedge \pi(x)\Psi(x, y)] \\ &= \pi(y)\Psi(y, x) \left[1 \wedge \frac{\pi(x)\Psi(x, y)}{\pi(y)\Psi(y, x)} \right] = \pi(y)P(y, x).\end{aligned}$$

Hence the chain is reversible, and π is stationary for P by Proposition 1.20.

Exercise 3.2

Let $x, y \in \mathcal{X}$. We want to show that

$$\pi(x)P(x, y) = \pi(y)P(y, x) \quad \text{for every } x, y \in \mathcal{X}. \quad (1)$$

Note that $P(x, y) > 0$ if and only if there exists a unique $v \in V$ with $y \in \mathcal{X}(x, v)$. In addition, note that $y \in \mathcal{X}(x, v)$ if and only if $x \in \mathcal{X}(y, v)$. Thus if such a unique $v \in V$ does not exist, then $P(x, y) = P(y, x) = 0$, and the relation (1) holds trivially. On the other hand, if such a unique $v \in V$ does exist, then

$$P(x, y) = \frac{1}{|V|} \pi^{x,v}(y) = \frac{1}{|V|} \frac{\pi(y)}{k_x},$$

where $k_x = \sum_{z \in \mathcal{X}(x,v)} \pi(z)$. In addition, noting that $\mathcal{X}(x, v) = \mathcal{X}(y, v)$, we see that

$$k_y = \sum_{z \in \mathcal{X}(y,v)} \pi(z) = \sum_{z \in \mathcal{X}(x,v)} \pi(z) = k_x,$$

and thus

$$P(y, x) = \frac{1}{|V|} \pi^{y,v}(x) = \frac{1}{|V|} \frac{\pi(x)}{k_y} = \frac{1}{|V|} \frac{\pi(x)}{k_x}.$$

Therefore,

$$\pi(x)P(x, y) = \pi(x) \left[\frac{1}{|V|} \frac{\pi(y)}{k_x} \right] = \pi(y) \left[\frac{1}{|V|} \frac{\pi(x)}{k_x} \right] = \pi(y)P(y, x).$$

Hence the chain is reversible, and π is stationary for P by Proposition 1.20.

Metropolis on graphs

Exercises 1-2

See the file `Metropolis_on_graphs_1.ipynb`.

Exercise 3

See the file `Metropolis_on_graphs_3.ipynb`.

Glauber dynamics on hardcore configurations

Exercises 1-3

See the file `Glauber_hardcore.ipynb`.

Exercise 4

I do this in the code. See the file `Glauber_hardcore.ipynb` for details of the implementation. To justify this theoretically, let $f : \mathcal{X} \rightarrow \mathbb{R}$ be a function defined as follows:

$$f(x) = \text{the number of particles in the configuration } x$$

for any $x \in \mathcal{X}$. Then f is a bounded function; specifically, $0 \leq f(x) \leq |V|$ for any $x \in \mathcal{X}$, where V denotes the vertex set of the graph. Since the uniform distribution π , defined by

$$\pi(x) = \frac{1}{|\mathcal{X}|} \quad \text{for any } x \in \mathcal{X}$$

is stationary for the chain, the Ergodic Theorem implies that

$$\mathbb{P}_{x_0} \left\{ \lim_{T \rightarrow \infty} \frac{1}{T} \sum_{t=0}^{T-1} f(X_t) = E_\pi(f) \right\} = 1, \quad (1)$$

where x_0 denotes the starting configuration of the chain. But note that

$$E_\pi(f) = \sum_{x \in \mathcal{X}} \pi(x) f(x) = \sum_{x \in \mathcal{X}} \frac{1}{|\mathcal{X}|} f(x),$$

the average number of particles over all hardcore configurations. Thus if we use the code to compute an observed value of the average $\frac{1}{T} \sum_{t=0}^{T-1} f(X_t)$ in (1) for a large value of T , then we expect this value to be close to the true average $E_\pi(f)$.