Homework 3

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1. Since $a \leq X \leq b$, we have the inequality

$$a + \frac{-b-a}{2} \le X + \frac{-b-a}{2} \le b + \frac{-b-a}{2}$$

which upon simplification becomes

$$\frac{a-b}{2} \le X + \frac{-b-a}{2} \le \frac{b-a}{2}.$$

Thus,

$$\begin{aligned} \operatorname{Var}(X) &= \operatorname{Var}\left(X + \frac{-b - a}{2}\right) \\ &= \operatorname{E}\left[\left(X + \frac{-b - a}{2}\right)^{2}\right] - \operatorname{E}\left[\left(X + \frac{-b - a}{2}\right)\right]^{2} \\ &\leq \operatorname{E}\left[\left(X + \frac{-b - a}{2}\right)^{2}\right] \leq \left(\frac{b - a}{2}\right)^{2} = \frac{(b - a)^{2}}{4}. \end{aligned}$$

However, this inequality is not tight. To see this, let $X \sim \mathrm{unif}(0,1)$. Then $0 \le X \le 1$, but

$$Var(X) = \frac{1}{12} < \frac{1}{4}.$$

2. (i) First, suppose $p \ge 1$ is even. Then the statement to prove reduces to

$$\Gamma(k) \le 3k^k$$
 for every integer $k \ge 1$. (1)

I will proceed by induction on k. Since $\Gamma(1) = 1 \le 3$, we see that (1) holds when k = 1. Now suppose that (1) holds for some k > 1. Then

$$\Gamma(k+1) = k\Gamma(k) \le k \cdot 3k^k = 3k^{k+1} \le 3(k+1)^{k+1}.$$

Thus (1) holds for every $k \geq 1$ by induction.

Next, suppose that $p \geq 1$ is odd. Then the statement to prove reduces to

$$\Gamma(k+1/2) \le 3(k+1/2)^{k+1/2} \quad \text{for every integer } k \ge 0.$$
 (2)

As before, I will proceed by induction on k. Since $\Gamma(1/2) = \sqrt{\pi} \le 3(1/2)^{1/2}$, we see that (2) holds when k = 0. Now suppose that (2) holds for some k > 0. Then

$$\Gamma((k+1)+1/2) = (k+1/2)\Gamma(k+1/2)$$

$$\leq (k+1/2) \cdot 3(k+1/2)^{k+1/2}$$

$$= 3(k+1/2)^{(k+1)+1/2}$$

$$\leq 3((k+1)+1/2)^{(k+1)+1/2}.$$

Thus (2) holds for every $k \ge 0$ by induction.

Combining these two cases, we therefore see that

$$\Gamma(p/2) \le 3(p/2)^{p/2}$$
 for every integer $p \ge 1$.

- (ii) Define a map $f:[0,\infty)\to\mathbb{R}$ by $f(x)=(2\sqrt{2})^x-x$. Note that $f(1)=2\sqrt{2}-1>0$. Moreover, since $\ln(2\sqrt{2})>1$, $f'(x)=\ln(2\sqrt{2})(2\sqrt{2})^x-1>0$ for every $x\geq 1$. Thus f is strictly increasing on $[1,\infty)$. In particular, since f(1)>0 as noted, f(x)>0 for every $x\geq 1$. Equivalently, $(2\sqrt{2})^x>x$ for every $x\geq 1$. Thus, for every integer $p\geq 1$, $(2\sqrt{2})^p>p$, so that $p^{1/p}\leq 2\sqrt{2}$.
- 3. (b) \implies (c). Let $p \ge 1$. Then

$$E(|X|^{p}) = \int_{0}^{\infty} P\{|X|^{p} > u\} du$$

$$= \int_{0}^{\infty} P\{|X|^{p} > t^{p}\} p t^{p-1} dt \quad \text{(substitute } u = t^{p})$$

$$= \int_{0}^{\infty} P\{|X| > t\} p t^{p-1} dt$$

$$\leq \int_{0}^{\infty} 2e^{-t^{2}/K_{1}} p t^{p-1} dt \quad \text{(by (b))}$$

$$= \int_{0}^{\infty} e^{-s} p(K_{1}s)^{p/2-1} K_{1} ds \quad \text{(substitute } s = t^{2}/K_{1})$$

$$= pK_{1}^{p/2} \int_{0}^{\infty} e^{-s} s^{p/2-1} ds$$

$$= pK_{1}^{p/2} \Gamma(p/2)$$

$$\leq 3K_{1}^{p/2} p(p/2)^{p/2} \quad \text{(by problem 2(i))}.$$

Thus,

$$\begin{split} [\mathrm{E}(|X|^p)]^{1/p} & \leq 3^{1/p} K_1^{1/2} p^{1/p} (p/2)^{1/2} \\ & \leq 3 K_1^{1/2} p^{1/p} (p/2)^{1/2} \quad \text{(since } 3^{1/p} \leq 3) \\ & \leq 6 \sqrt{2} K_1^{1/2} (p/2)^{1/2} \quad \text{(by problem 2(ii))} \\ & = 6 K_1^{1/2} \sqrt{p}. \end{split}$$

Thus, letting $K_2 = 6K_1^{1/2} > 0$, we have

$$[\mathrm{E}(|X|^p)]^{1/p} \le K_2 \sqrt{p}.$$

(c) \implies (d). Let $K_3 = 4eK_2^2 > 0$. We have the following:

$$E(e^{X^2/K_3}) = E\left[\sum_{p=0}^{\infty} \frac{(X^2/K_3)^p}{p!}\right] = 1 + E\left[\sum_{p=1}^{\infty} \frac{(X^2/K_3)^p}{p!}\right]$$
$$= 1 + E\left[\sum_{p=1}^{\infty} \frac{X^{2p}/K_3^p}{p!}\right] = 1 + \sum_{p=1}^{\infty} \frac{E[X^{2p}]/K_3^p}{p!}.$$

By (c), $E(X^{2p}) \leq K_2^{2p}(2p)^p$ for every $p \geq 1$. In addition, by Stirling's approximation,

$$p! \ge \sqrt{2\pi} p^{p+1/2} e^{-p} \ge p^p e^{-p} = \left(\frac{p}{e}\right)^p$$

for every $p \ge 1$. Thus,

$$E(e^{X^2/K_3}) = 1 + \sum_{p=1}^{\infty} \frac{E[X^{2p}]/K_3^p}{p!} \le 1 + \sum_{p=1}^{\infty} \frac{K_2^{2p}(2p)^p/K_3^p}{(p/e)^p}$$

$$= 1 + \sum_{p=1}^{\infty} \frac{K_2^{2p}(2p)^p/(4eK_2^2)^p}{(p/e)^p} = 1 + \sum_{p=1}^{\infty} \frac{(2p)^p/(4e)^p}{(p/e)^p}$$

$$= 1 + \sum_{p=1}^{\infty} \left(\frac{1}{2}\right)^p = \sum_{p=0}^{\infty} \left(\frac{1}{2}\right)^p = 2.$$

(d) \Longrightarrow (b). Let t>0. Then since $\mathrm{E}(e^{X^2/K_3})\leq 2$ by (d), an application of Markov's Inequality yields

$$P\{|X| > t\} = P\{e^{X^2/K_3} > e^{t^2/K_3}\} \le e^{-t^2/K_3} E(e^{X^2/K_3}) \le 2e^{-t^2/K_3}.$$

Thus, taking $K_1 = K_3$, we have

$$P\{|X| > t\} < 2e^{-t^2/K_1}.$$

(c) \implies (a). Let $p \ge 1$. Then since

$$[\mathrm{E}(|X|^p)]^{1/p} \le K_2 \sqrt{p}$$

by (c), an application of Minkowski's Inequality yields

$$[E(|X - \mu|^p)]^{1/p} \le [E(|X|^p)]^{1/p} + [E(|-\mu|^p)]^{1/p} = [E(|X|^p)]^{1/p} + |\mu|$$

$$\le K_2 \sqrt{p} + |\mu| \le K_2 \sqrt{p} + |\mu| \sqrt{p}$$

$$= (K_2 + |\mu|) \sqrt{p} = K_2' \sqrt{p},$$

where $K_2' = K_2 + |\mu| > 0$. Thus, since $E(X - \mu) = 0$, the result proven in the lectures implies that

$$E(e^{\lambda(X-\mu)}) \le e^{\lambda^2/K_0}$$
 for any $\lambda \in \mathbb{R}$

for some $K_0 > 0$.

(a) \Longrightarrow (c). Let $p \ge 1$. Since $\mathrm{E}(X-\mu)=0$, the result proven in the lectures implies that

$$[E(|X - \mu|^p)]^{1/p} \le K_2' \sqrt{p}$$

for some $K'_2 > 0$. Thus by Minkowski's Inequality,

$$[E(|X|^p)]^{1/p} = [E(|(X - \mu) + \mu|^p)]^{1/p} \le [E(|X - \mu|^p)]^{1/p} + [E(|\mu|^p)]^{1/p}$$

$$= [E(|X - \mu|^p)]^{1/p} + |\mu| \le K_2' \sqrt{p} + |\mu|$$

$$\le K_2' \sqrt{p} + |\mu| \sqrt{p} = (K_2' + \mu) \sqrt{p} = K_2 \sqrt{p},$$

where $K_2 = K_2' + |\mu| > 0$.

4. Let $S \subseteq \{1, 2, ..., \beta\}$, and let $\nu_0 > 0$ and $\xi > 1$. Assume the restricted eigenvalue condition holds; that is, for any $b \in \mathbb{R}^p$ satisfying

$$\sum_{j \notin S} |b_j| \le \xi \sum_{j \in S} |b_j|,\tag{1}$$

we have

$$\nu_0^2 \left(\sum_{j \in S} b_j^2 \right) \le b^T \widetilde{\Sigma} b. \tag{2}$$

So let $b \in \mathbb{R}^p$ satisfy (1). Then b satisfies (2). Then by Cauchy-Schwarz,

$$\left(\sum_{j\in S} |b_j|\right)^2 \le \left(\sum_{j\in S} b_j^2\right) |S|.$$

Thus, by (2),

$$\nu_0^2 \left(\sum_{j \in S} |b_j| \right)^2 \le \nu_0^2 \left(\sum_{j \in S} b_j^2 \right) |S| \le (b^T \widetilde{\Sigma} b) |S|.$$

We therefore see that the compatibility condition holds. Hence the restricted eigenvalue condition implies the compatibility condition.

5. By definition, $\hat{\beta}$ minimizes the Lasso objective function. Thus, for any $0 \le t \le 1$,

$$\frac{1}{2}\|Y - X\hat{\beta}\|_n^2 + \lambda \|\beta\|_1 \le \frac{1}{2}\|Y - X[(1-t)\hat{\beta} + t\beta^*]\|_n^2 + \lambda \|(1-t)\hat{\beta} + t\beta^*\|_1. \tag{1}$$

Note that we have

$$\begin{split} &\frac{1}{2}\|Y - X[(1-t)\hat{\beta} + t\beta^*]\|_n^2 \\ &= \frac{1}{2}\|Y - X\hat{\beta} - tX(\hat{\beta} - \beta^*)\|_n^2 \\ &= \frac{1}{2}\|Y - X\hat{\beta}\|_n^2 + \frac{1}{2}t^2\|X(\hat{\beta} - \beta^*)\|_n^2 - t\langle Y - X\hat{\beta}, X(\hat{\beta} - \beta^*)\rangle_n. \end{split}$$

Thus we can write (1) as

$$\lambda \|\hat{\beta}\|_{1} \leq \frac{1}{2} t^{2} \|X(\hat{\beta} - \beta^{*})\|_{n}^{2} - t \langle Y - X\hat{\beta}, X(\hat{\beta} - \beta^{*})_{n} + \lambda \|(1 - t)\hat{\beta} + t\beta^{*})\|_{1}.$$

[This is as far as I got – all attempts hereafter were fruitless. Is there a trick, or is this just a lot of elaborate algebra?]