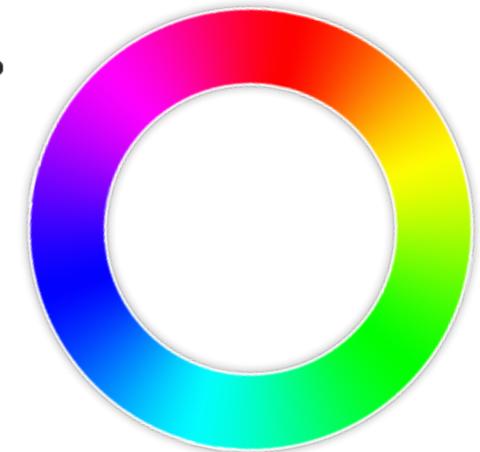


## Continuous Random Variables

- What if the range of a random variable consists of one or more intervals?

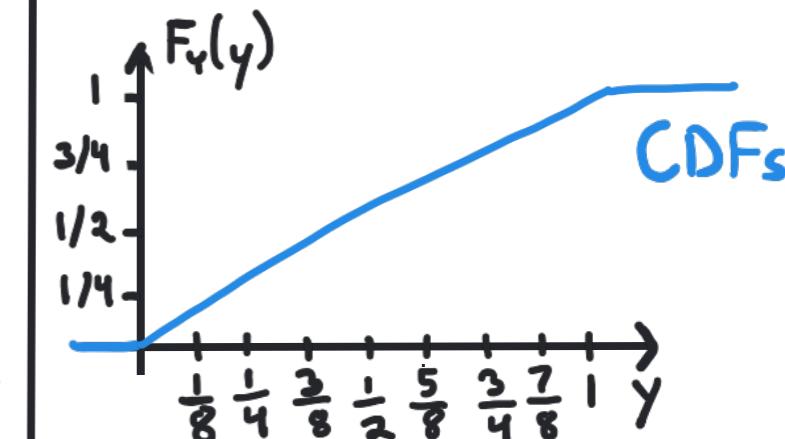
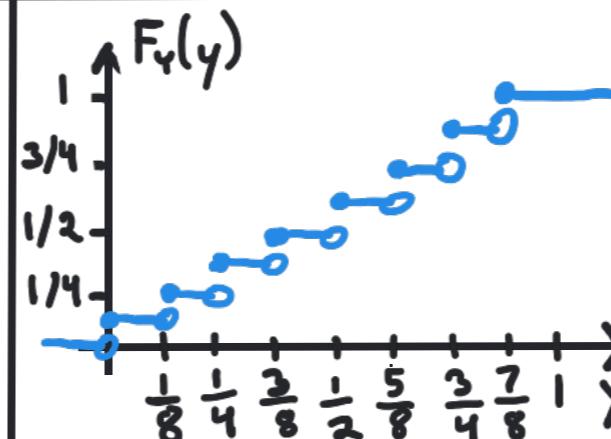
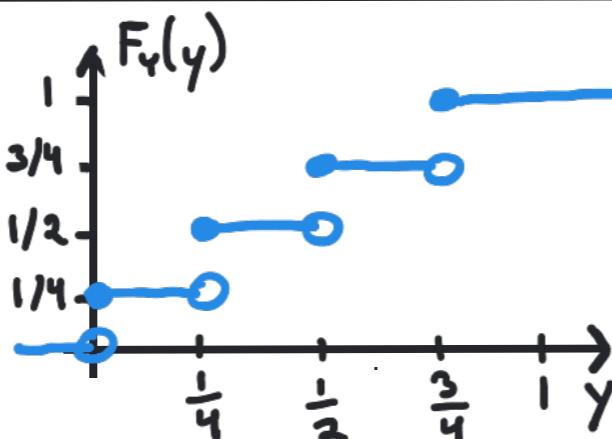
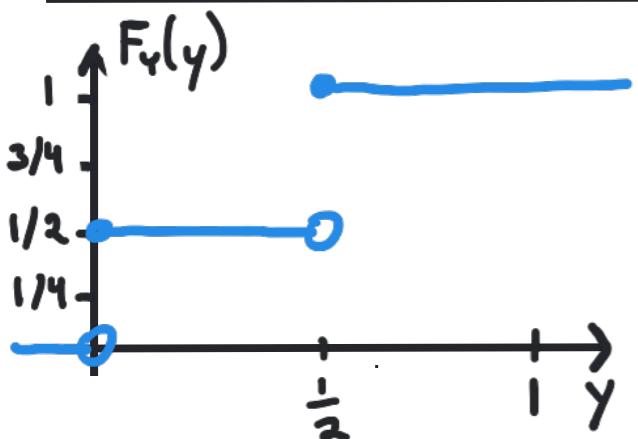
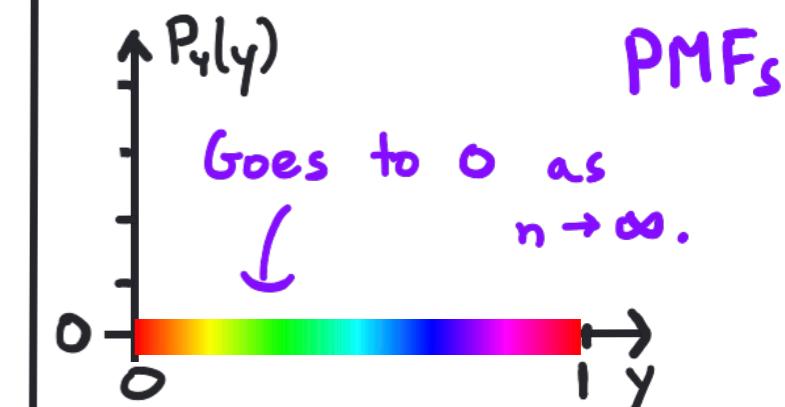
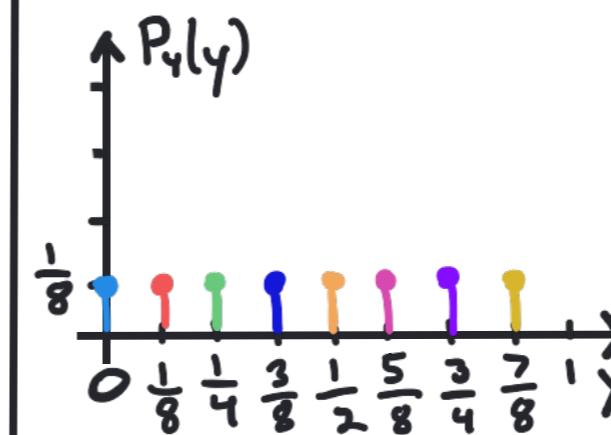
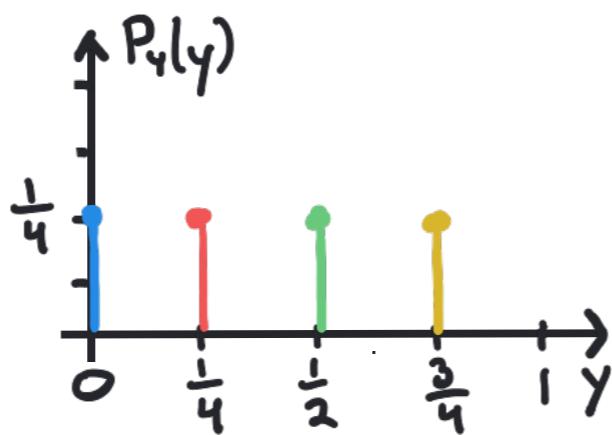
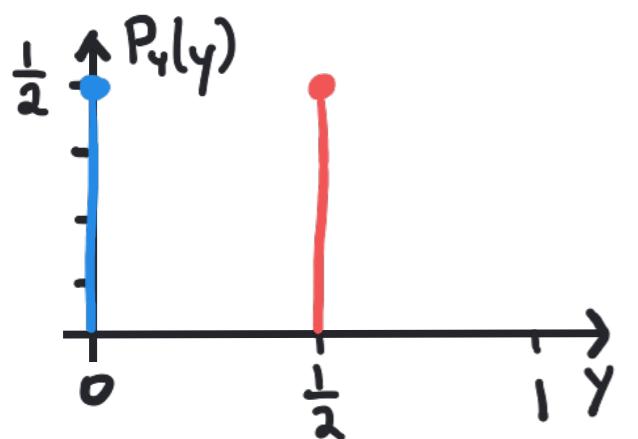
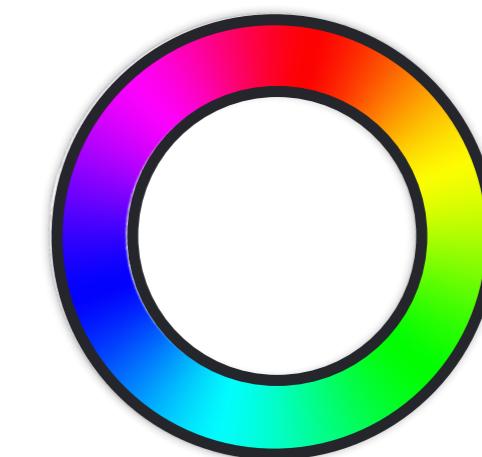
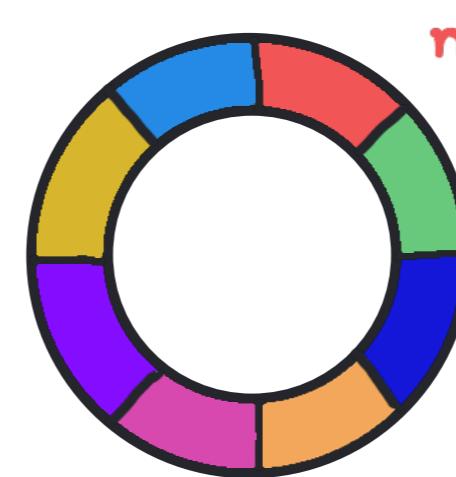
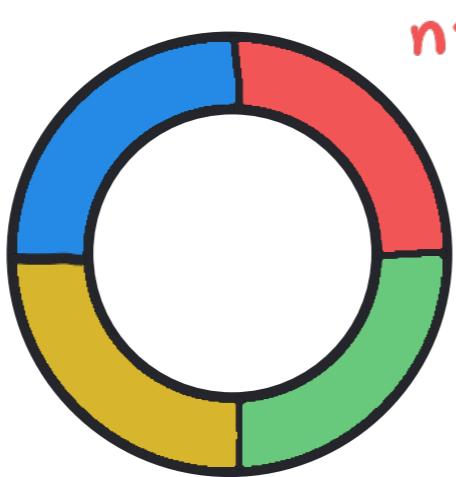
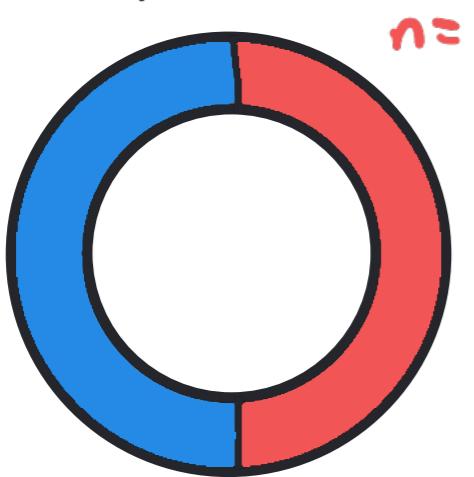
→ Ex: Measure a voltage between -5 to +5 Volts.

→ Range is no longer discrete so we need a new type of random variable.



- Motivating Example: → Wheel (of fortune) of diameter 1 meter.
  - Spin the wheel very fast and record where it stops.
  - Reasonable to model all stopping positions as equally likely.
  - Overall, we have a random variable  $X$  that is equally likely to take any value  $x \in [0, 1]$ .
  - What is the probability that  $X = x$  for some value  $x \in [0, 1]$ ?  
*Intuitively, this should be 0.*
  - Let's try to model this scenario using discrete random variables.

- Divide wheel into  $n$  equal-sized pieces. Let  $\gamma$  be the slice you land in. (Formally, let  $\gamma = \frac{1}{n} \cdot \text{floor}(nX)$ .)  $R_\gamma = \{0, \frac{1}{n}, \frac{2}{n}, \dots, \frac{n-1}{n}\}$ .

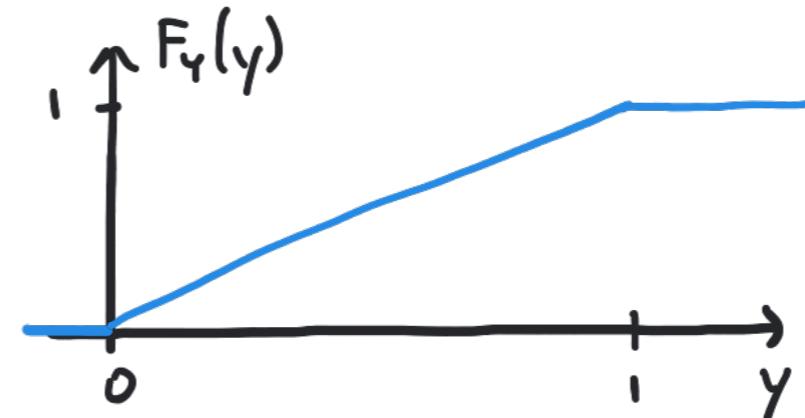


$$F_y(y) = \mathbb{P}[\{\gamma \leq y\}]$$

Makes sense even as  $n \rightarrow \infty$ !

- As  $n$  increases, the PMF values tend to 0, while the CDF remains well-behaved. Specifically, as  $n \rightarrow \infty$ ,

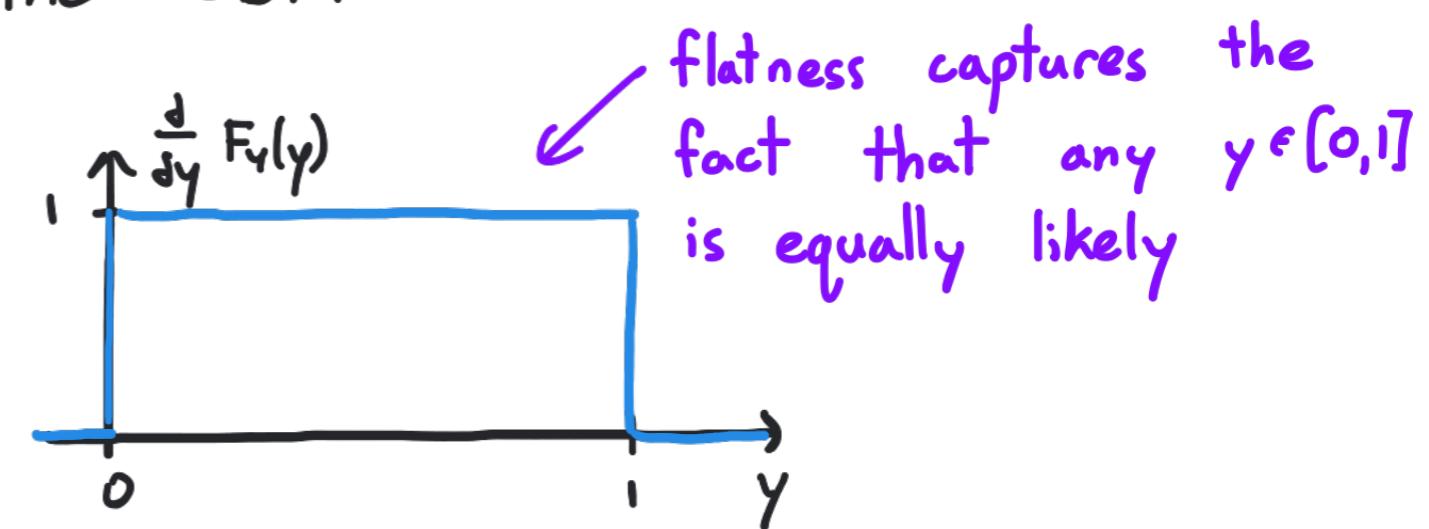
$$F_Y(y) = \begin{cases} 0 & y < 0 \\ y & 0 \leq y \leq 1 \\ 1 & y > 1 \end{cases}$$



→ It would be nice to have a way to represent the fact that all positions  $y \in [0, 1]$  are equally likely.

- Idea: Take the derivative of the CDF.

$$\frac{d}{dy} F_Y(y) = \begin{cases} 0 & y < 0 \\ 1 & 0 \leq y < 1 \\ 0 & y > 1 \end{cases}$$



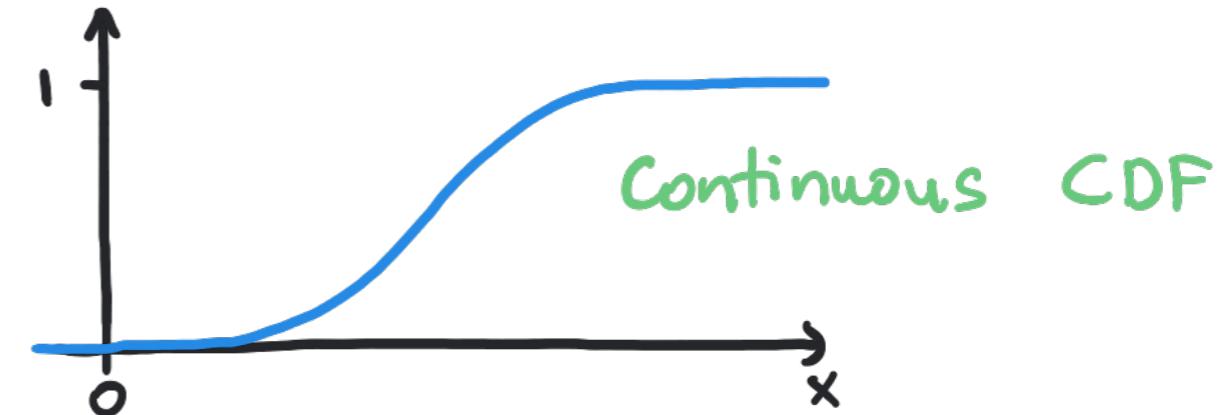
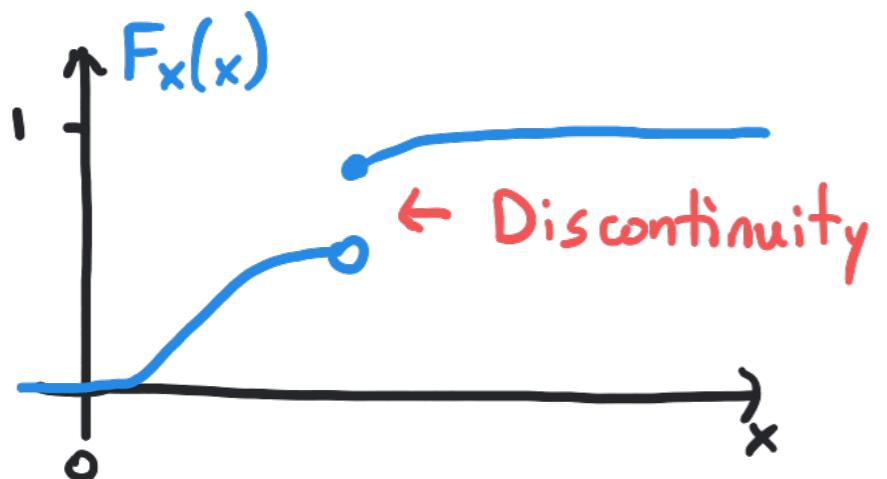
- Note that  $P[\{Y = y\}] = 0$  for any choice of  $y \in [0, 1]$ .

## Cumulative Distribution Function

- Recall that we defined the cumulative distribution function (CDF)  $F_x(x)$  as the probability that a random variable  $X$  is less than or equal to the value  $x$ ,

$$\begin{aligned} F_x(x) &= \mathbb{P}\left[\{\omega \in \Omega : X(\omega) \leq x\}\right] \\ &= \mathbb{P}[X \leq x] \quad \text{shorthand notation} \\ &= \mathbb{P}[X \leq x] \quad \text{shorthand notation} \end{aligned}$$

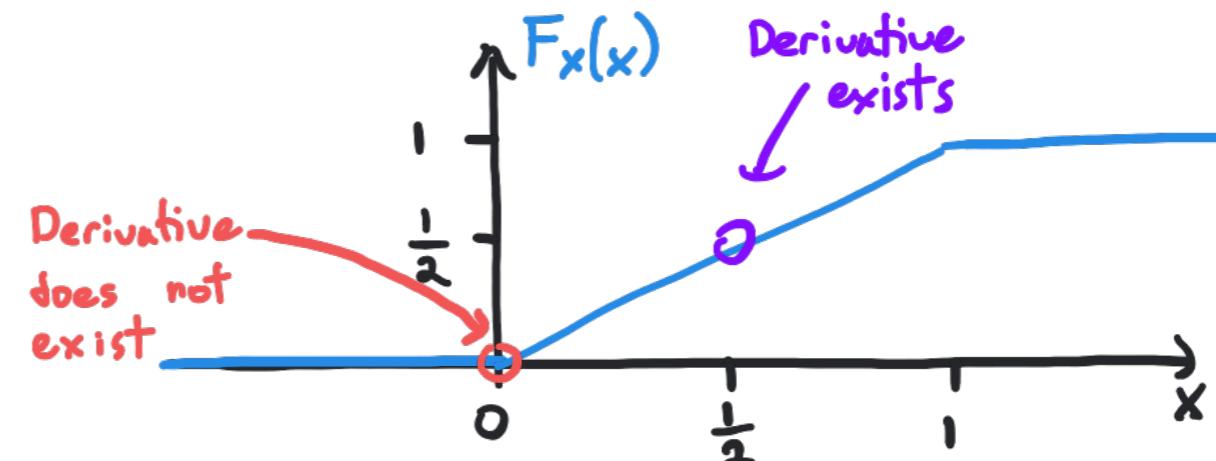
- The CDF is continuous if it has no abrupt changes,  
 $\lim_{\epsilon \rightarrow 0} F_x(x + \epsilon) = F_x(x)$  for all  $x$ .



- A random variable  $X$  is continuous if it has a continuous CDF  $F_X(x)$  and this CDF is differentiable almost everywhere. Everywhere except for a countable number of values where the derivative does not exist.

- Recall that a function  $g(x)$  is differentiable at  $x$  if the derivative  $g'(x) = \lim_{\epsilon \rightarrow 0} \frac{g(x+\epsilon) - g(x)}{\epsilon}$  exists.

$$\rightarrow \underline{\text{Ex:}} \quad F_X(x) = \begin{cases} 0 & x < 0 \\ x & 0 \leq x \leq 1 \\ 1 & x > 1 \end{cases}$$



$$\lim_{\epsilon \downarrow 0} \frac{F_X(\frac{1}{2} + \epsilon) - F_X(\frac{1}{2})}{\epsilon} = \lim_{\epsilon \downarrow 0} \frac{\frac{1}{2} + \epsilon - \frac{1}{2}}{\epsilon} = 1 = \lim_{\epsilon \uparrow 0} \frac{\frac{1}{2} + \epsilon - \frac{1}{2}}{\epsilon} \quad \text{Limit from the right} \quad \text{Limit from the left} \quad \text{Derivative exists.}$$

$$\lim_{\epsilon \downarrow 0} \frac{F_X(0 + \epsilon) - F_X(0)}{\epsilon} = \lim_{\epsilon \downarrow 0} \frac{0 + \epsilon - 0}{\epsilon} = 1$$

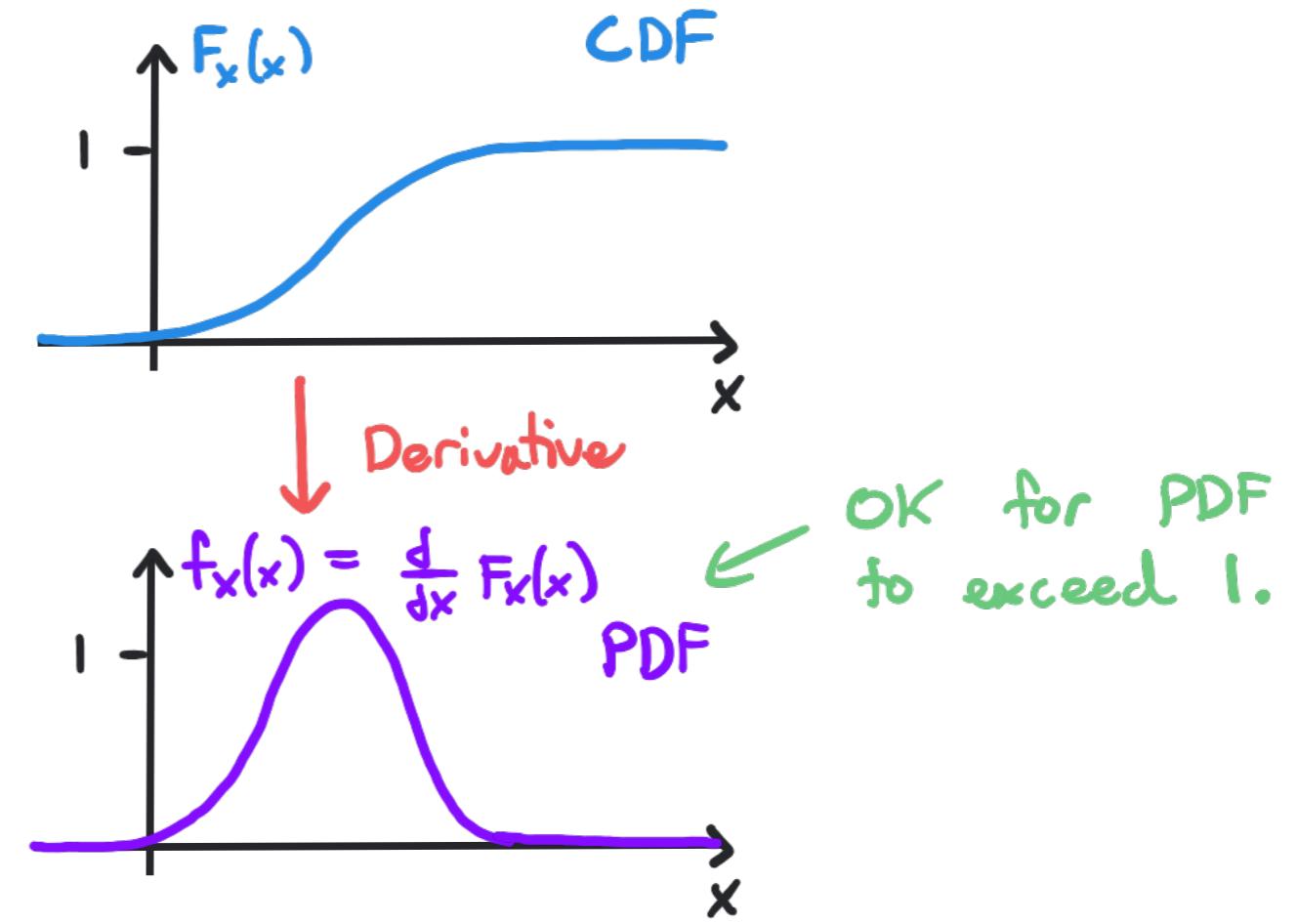
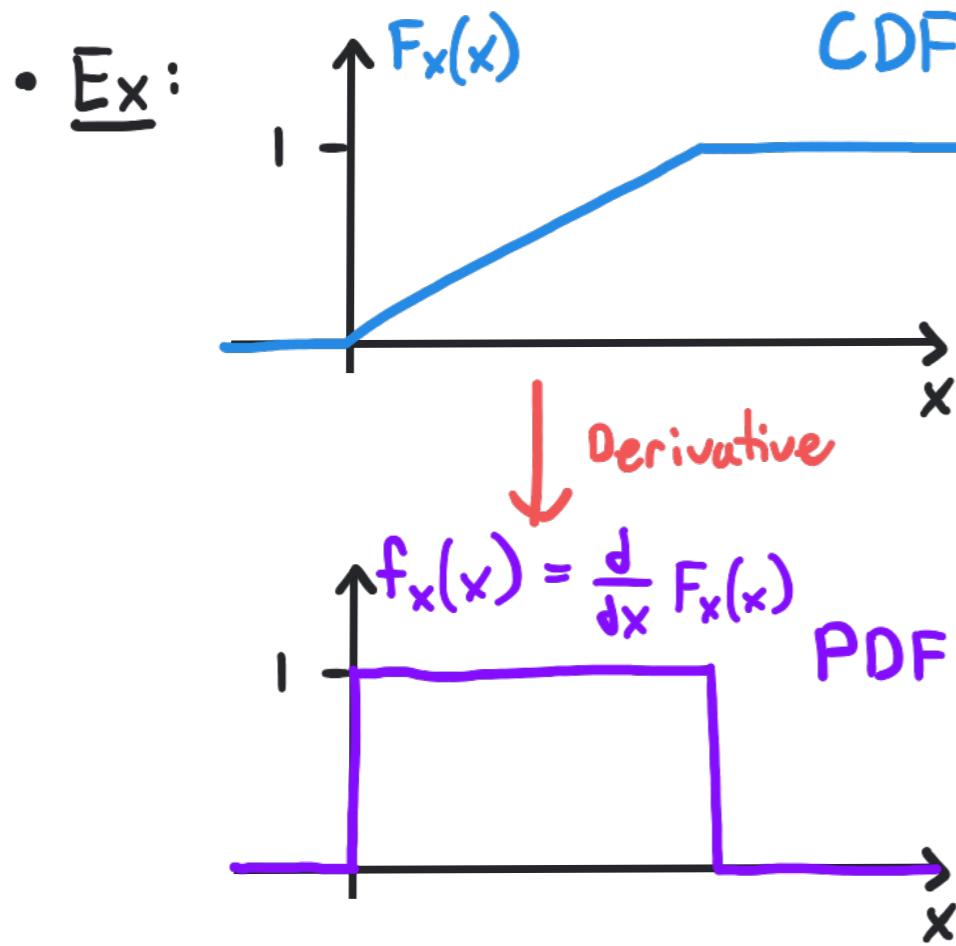
Limit from the right and from the left do not agree so the derivative does not exist at  $x=0$

$$\lim_{\epsilon \uparrow 0} \frac{F_X(0 + \epsilon) - F_X(0)}{\epsilon} = \lim_{\epsilon \uparrow 0} \frac{0 - 0}{\epsilon} = 0$$

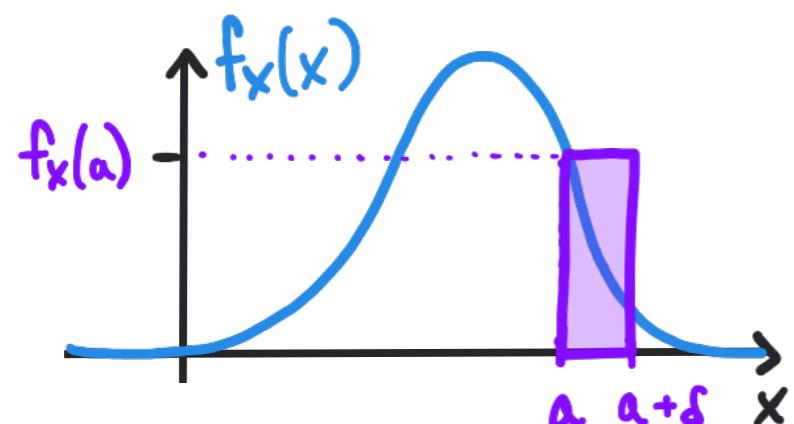
## Probability Density Function

- The probability density function (PDF)  $f_x(x)$  of a continuous random variable  $X$  is the derivative of its CDF  $F_x(x)$ .

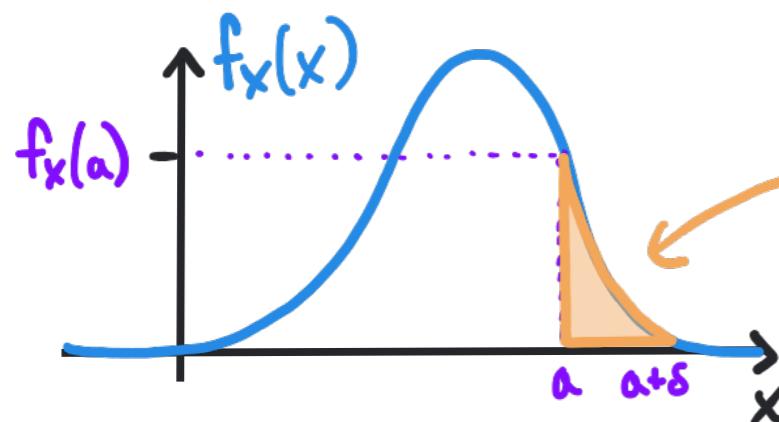
$$f_x(x) = \begin{cases} \frac{d}{dx} F_x(x) & \text{if } F_x(x) \text{ is differentiable at } x \\ \text{any non-negative number} & \text{otherwise} \end{cases}$$



- The PDF  $f_x(x)$  does not tell us the probability  $\mathbb{P}\{X=x\}$ , which is always equal to 0 for a continuous random variable.
- Instead, the PDF  $f_x(x)$  tells us the "density" of probability around  $x$ .



Say we want to know  $\mathbb{P}\{X \in [a, a+\delta]\}$ . We can approximate this by the area  $\delta f_x(a)$ . The approximation quality improves as  $\delta \rightarrow 0$ . This sounds like integration!



The exact probability is the area under the curve,

$$\mathbb{P}\{X \in [a, a+\delta]\} = \int_a^{a+\delta} f_x(x) dx .$$

- Intuition: The PDF can play the role of the PMF if we replace sums with integrals.

## PDF and CDF Properties

PDF Property	Property Name	CDF Property
$f_x(x) \geq 0$	Non-Negativity	$F_x(x)$ is non-decreasing
$\int_{-\infty}^{\infty} f_x(x) dx = 1$	Normalization	$F_x(+\infty) = 1, F_x(-\infty) = 0$
$\int_{-\infty}^x f_x(u) du = F_x(x)$ <small>integration variable</small>	PDF $\longleftrightarrow$ CDF	$\frac{d}{dx} F_x(x) = f_x(x)$ <small>for <math>x</math> where derivative exists</small>
$P[\{a < X \leq b\}]$ $= \int_a^b f_x(x) dx$	Probability of an Interval	$P[\{a < X \leq b\}]$ $= F_x(b) - F_x(a)$

• Since for a continuous random variable  $P[\{X = x\}] = 0$ , we have

$$P[\{a < X < b\}] = P[\{a \leq X < b\}] = P[\{a < X \leq b\}] = P[\{a \leq X \leq b\}].$$

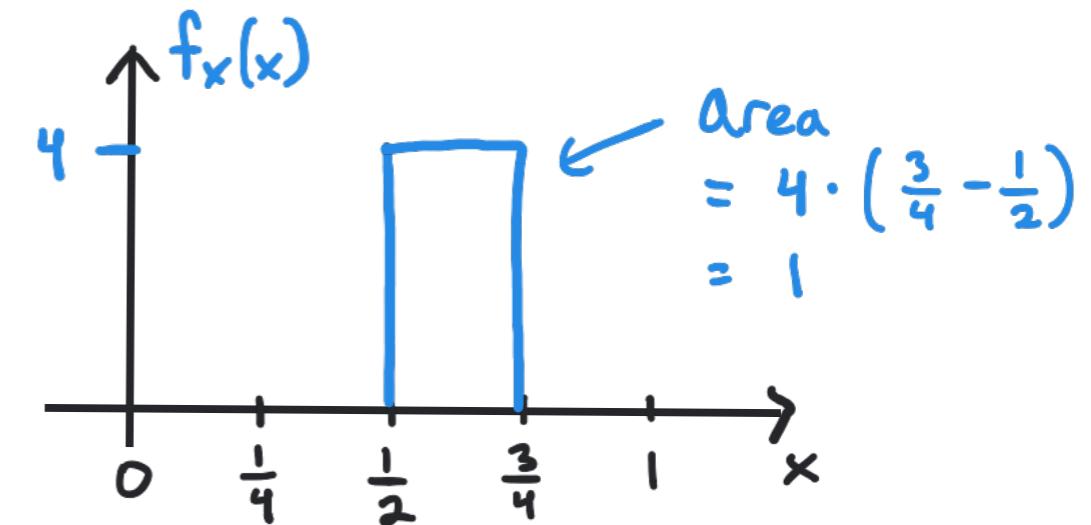
- Example:  $X$  is equally likely to take any value in  $[\frac{1}{2}, \frac{3}{4}]$ .

→ Determine the PDF.

\* Easier to start with a plot.

Set height so that area = 1.

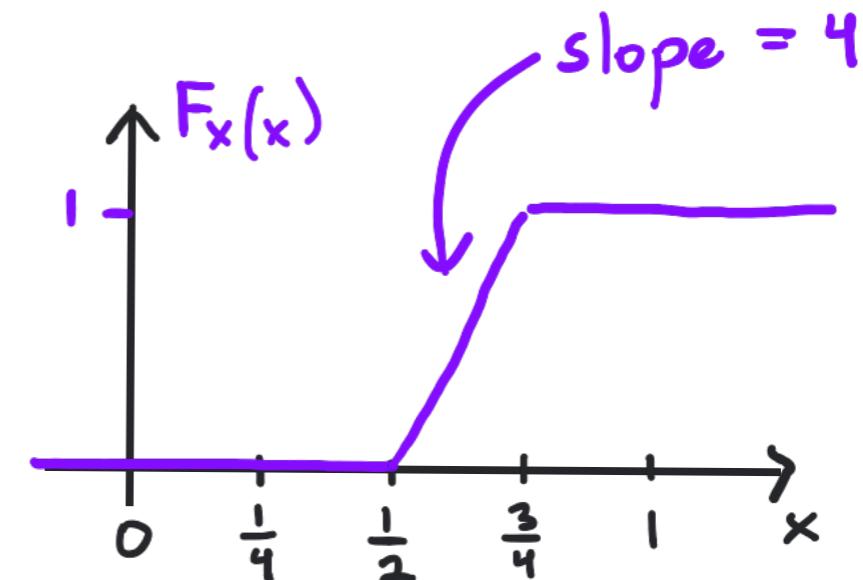
$$f_x(x) = \begin{cases} 0 & x < \frac{1}{2} \\ 4 & \frac{1}{2} \leq x \leq \frac{3}{4} \\ 0 & \frac{3}{4} < x \end{cases}$$



→ Determine the CDF.

\* Can guess integral from plot or compute directly.

$$F_x(x) = \int_{-\infty}^x f_x(u) du = \begin{cases} \int_{-\infty}^x 0 du & x < \frac{1}{2} \\ \int_{-\infty}^{1/2} 0 du + \int_{1/2}^x 4 du & \frac{1}{2} \leq x \leq \frac{3}{4} \\ \int_{-\infty}^{1/2} 0 du + \int_{1/2}^{3/4} 4 du + \int_{3/4}^x 0 du & \frac{3}{4} < x \end{cases}$$



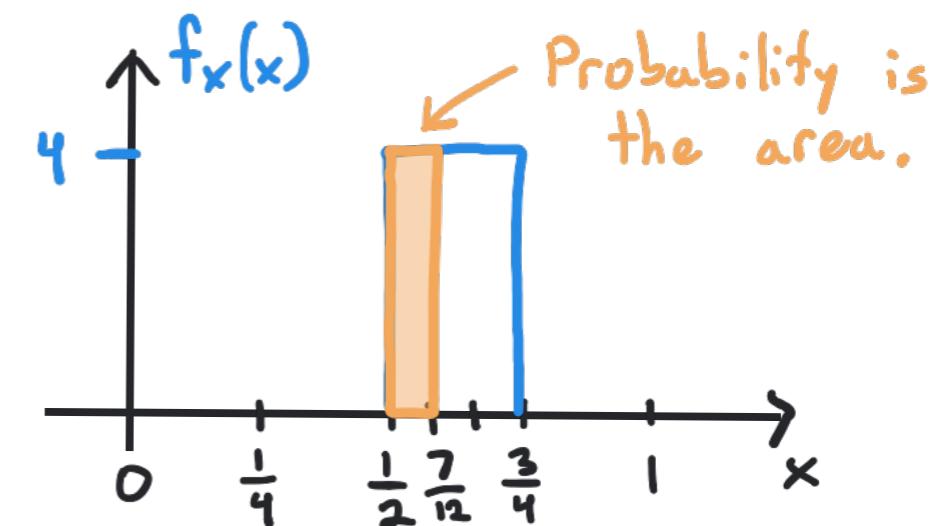
$$= \begin{cases} 0 & x < \frac{1}{2} \\ 4x - 2 & \frac{1}{2} \leq x \leq \frac{3}{4} \\ 1 & \frac{3}{4} < x \end{cases}$$

- Example:  $X$  is equally likely to take any value in  $\left[\frac{1}{2}, \frac{3}{4}\right]$ .

→ Determine  $\text{IP}\left[\left\{\frac{1}{2} \leq X \leq \frac{7}{12}\right\}\right]$ .

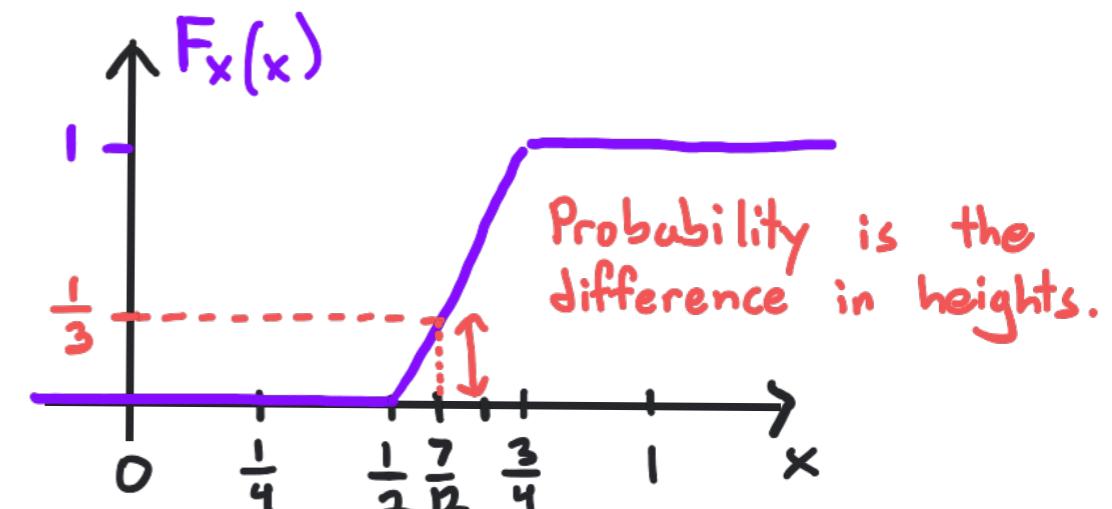
\* PDF Calculation:

$$\begin{aligned} \int_{1/2}^{7/12} f_x(x) dx &= \int_{1/2}^{7/12} 4 dx = (4x) \Big|_{1/2}^{7/12} \\ &= 4 \left( \frac{7}{12} - \frac{1}{2} \right) \\ &= 4 \cdot \frac{1}{12} = \frac{1}{3} \end{aligned}$$



\* CDF Calculation:

$$\begin{aligned} F_x\left(\frac{7}{12}\right) - F_x\left(\frac{1}{2}\right) \\ &= \left(4 \cdot \frac{7}{12} - 2\right) - \left(4 \cdot \frac{1}{2} - 2\right) \\ &= \frac{7}{3} - 2 = \frac{1}{3} \end{aligned}$$



- Example:  $X$  is equally likely to take any value in  $[\frac{1}{2}, \frac{3}{4}]$ .

→ Determine  $\text{IP}[\{\frac{5}{8} \leq X \leq 1\}]$ .

\* PDF Calculation:

$$\int_{5/8}^1 f_x(x) dx = \int_{5/8}^1 4 dx = 4 \left(1 - \frac{5}{8}\right) = \frac{12}{8}$$

This is wrong.

$$\int_{5/8}^{3/4} 4 dx + \int_{3/4}^1 0 dx = 4 \left(\frac{3}{4} - \frac{5}{8}\right) = \frac{1}{2}$$

Correctly split into two cases.

\* CDF Calculation:

$$\begin{aligned} F_x(1) - F_x\left(\frac{5}{8}\right) \\ = 1 - \left(4 \cdot \frac{5}{8} - 2\right) \\ = 1 - \frac{5}{2} + 2 = \frac{1}{2} \end{aligned}$$

We don't need to be as careful in the CDF calculation since we already handled the different integration regions.

