$$H_0: \hat{\beta}_1 + \hat{\beta}_2 = 0 H_1: \hat{\beta}_1 + \hat{\beta}_2 \neq 0$$
 (1)

(b.)

To test single linear hypothesis, we could do t-test by constructing test statistic:

$$T = rac{c'\hat{eta}}{s\sqrt{c'(x'x)^{-1}c}} \sim t(6)$$

where c' is $[0 \quad 1 \quad 1 \quad 0]$, and $s=\frac{\sum \hat{\epsilon}^2}{6}=\frac{\hat{\epsilon}'\hat{\epsilon}}{6}$

For |T|>2.45, null hypothesis is rejected. Thus, effects of increased expenditure of candidate A or B do not cancel out.

(c.)

If null hypothesis rejected, then at least one candidate's expenditure will have effect on the outcome.

$$H_0: \hat{\beta}_1 = \hat{\beta}_2 = 0$$

 $H_1: \hat{\beta}_1 \neq 0 \text{ or } \hat{\beta}_2 \neq 0$ (2)

(d.)

To perform an joint hypotheses F-test, construct test statistics:

$$F = \frac{(R\hat{\beta} - r)'(R(X'X)^{-1}R')^{-1}(R\hat{\beta} - r)/2}{\hat{\epsilon}'\hat{\epsilon}/6} \sim F(2, 6)$$
 (3)

where
$$R=egin{bmatrix} 0 & 1 & 0 & 0 \ 0 & 0 & 1 & 0 \end{bmatrix}$$
 , $r=egin{bmatrix} 0 \ 0 \end{bmatrix}$.

For F>5.14, null hypothesis is rejected. At least one candidate's expenditure have effect on the outcome.

- 2. (a.) t-statistic of branchWest = -1.29 with degree of freedom = 1261. p-value = 0.20. Null hypothesis is not rejected under $\alpha=0.05$
 - (b.) t-statistic of branchWest = -1.29 with degree of freedom = 1261. Since t distribution is symmetric, p-value of one-sided test is half of two-sided test = 0.10. Null hypothesis is still not rejected under $\alpha=0.05$
 - (c.) This is an overall F-test. F-statistic = 97.9 follows F distribution with degree of freedom (4, 1261). p-value $< 2 \times 10^{-16}$. Null hypothesis is rejected under $\alpha = 0.05$
- 3. (a.) Let schyr = X, cog = Z

$$\begin{split} \hat{\beta}_1 &= \frac{\sum (X_i - \bar{X})(Y_i - \bar{Y})}{\sum (X_i - \bar{X})^2} \\ &\xrightarrow{prob.} \frac{\text{Cov}(X, Y)}{\text{Var}(X)} \\ &= \frac{\text{Cov}(X, \beta_0 + \beta_1 X + \beta_2 Z + \epsilon)}{\sigma_1^2} \\ &= \frac{\beta_1 \text{Var}(X) + \beta_2 \text{Cov}(X, Z) + \text{Cov}(X, \epsilon))}{\sigma_1^2} \\ &= \frac{\beta_1 \sigma_1^2 + \beta_2 \rho \sigma_1 \sigma_2}{\sigma_1^2} \\ &= \beta_1 + \beta_2 \rho \frac{\sigma_2}{\sigma_1} \end{split}$$

Thus, $\hat{\beta}_1$ is inconsistent estimator for β_1 .

(b.)

$$\hat{eta}_1 - eta_1 \xrightarrow{prob.} eta_2
ho(\sigma_2/\sigma_1) > 0, \, ext{for } eta_2 > 0$$
 (4)

 $\hat{\beta}_1$ is upward biased.

(c.)

$$\sqrt{n}(\hat{\beta}_1 - \beta_1) \xrightarrow{p} \sqrt{n}(\beta_2 \rho(\sigma_2/\sigma_1)) \to \infty \text{ as } n \to \infty$$
 (5)