Math 170S

6.8

Bayesian Estimation

Coin Toss

Example 6.1

Suppose there are three coins in a pocket with success (head) probabilities of 0.2, 0.5, and 0.8 respectively. I pick one coin uniformly at random and flip it twice. Calculate the probability of getting heads on both tosses.

$$P(\theta = 0.2) = \frac{1}{3}$$

$$P(X_{1}, X_{2} = | \theta = 0.2) = 0.2^{2}$$

$$P(\theta = 0.8) = \frac{1}{3}$$

$$P(X_{1}, X_{2} = | \theta = 0.5) = (0.5)^{2}$$

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 $=(0.2^{2})\cdot\frac{1}{2}+(0.5)^{2}\cdot\frac{1}{1}+(0.8)^{2}\cdot\frac{1}{5}=0.31$

$$\Pi_{\text{prior}}(\Theta=0.2) = \Pi_{\text{prior}}(\Theta=0.5) = \Pi_{\text{prior}}(\theta=0.8) = \frac{1}{3}$$

Coin Toss

Example 6.2

Compute the probability

$$P[\Theta = 0.2 \mid X_1, X_2 = 1];$$
 $P[\Theta = 0.5 \mid X_1, X_2 = 1];$ $P[\Theta = 0.8 \mid X_1, X_2 = 1].$

$$\frac{\text{TIpost } (\theta = 0.1)}{\text{P}(X_1, X_2 = 1 \mid 0 = 0.2) \cdot \text{P}(\theta = 0.2)} = \frac{P(X_1, X_2 = 1 \mid 0 = 0.2) \cdot P(\theta = 0.2)}{P(X_1, X_2 = 1)} = \frac{(0.2)^{2} \cdot \frac{1}{5}}{0.51} = 0.044$$

$$= \frac{P(X_1, X_2=1)}{P(X_1, X_2=1)} = \frac{P(X_1, X_2=1|0=0.5) \cdot P(0=0.5)}{P(X_1, X_2=1)} = \frac{(0.5)^2 \cdot \frac{1}{5}}{0.55}$$

$$= P(0 = 0.8 \mid X_{11} X_{2} = 1) = \frac{P(X_{11} \mid X_{2} = 1 \mid 0 = 0.8) \cdot P(0 = 0.8)}{P(X_{11} \mid X_{2} = 1)} = \frac{0.8^{2} \cdot \frac{1}{3}}{0.31} = 0.688$$

posterior pmf.

Bayesian Inference; Discrete Case

- 1. Let X be a random variable with distribution $f(\cdot|\theta) =: f_{\theta}(\cdot)$ for some parameter Θ .
- 2. Θ is a **discrete** random variable on Ω with an unknown pmf π .
- 3. **Problem:** Estimate the unknown pmf π .
- 4. Input:
 - ightharpoonup Sample values x_1, \ldots, x_n from n experiments.
 - A **prior pmf** π_{prior} which we think is the best estimate for π before we run the experiments.
- 5. Output: A posterior pmf π_{post} that we think is the best estimate for π after we observe the experiments.
- 6. Method:
 - 6.1 Compute the quantity

$$\mathcal{K} := \sum_{\theta \in \Omega} f_{\theta}(x_1) \dots f_{\theta}(x_n) \pi_{prior}(\theta).$$

6.2 Compute the posterior pmf π_{post} by

$$\pi_{post}(\theta) = \frac{f_{\theta}(x_1) \dots f_{\theta}(x_n) \pi_{prior}(\theta)}{K}.$$

Back to coin toss once again

Example 6.3

In the coin toss example. I pick one coin at random following an unknown pmf. In the language of Bayesian inference,

- \triangleright X is a Bernoulli random variable with success probability Θ ;
- $ightharpoonup \Theta$ is randomly picked from the set $\{0.2, 0.5, 0.8\}$ following some unknown pmf π .
- Since we have no information regarding π , our best guess would be π_{prior} is the uniform distribution,

$$\pi_{prior}[\Theta = 0.2] = \pi_{prior}[\Theta = 0.5] = \pi_{prior}[\Theta = 0.8] = \frac{1}{3}.$$

Now the chosen coin is flipped twice, and both outcomes are equal to head, so $x_1 = x_2 = 1$. **After** observing these experiments, we update our prediction on π by π_{post} using the given method.

Bayesian Inference(continuous)

Bayesian inference for the continuous case works the same way with discrete case, except that sum in the formula for K is replaced with integrals.

$$K := \int_{-\infty}^{\infty} f_{\theta}(x_1) \dots f_{\theta}(x_n) \pi_{prior}(\theta) d\theta.$$

Bayesian Inference(continuous)

Example 6.4

Let X be binomial distribution with parameters n (given) and θ (unknown),

$$f(\cdot|\theta) = f_{\theta}(x) = \binom{n}{x} \theta^{x} (1-\theta)^{n-x} \qquad x = 0, 1, \dots, n.$$

Suppose that π_{prior} is the beta pdf with parameter α, β (given),

$$\pi_{prior}(\theta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \, \theta^{\alpha - 1} (1 - \theta)^{\beta - 1} \qquad 0 < \theta < 1,$$

where Γ is the gamma function. Suppose that we have performed one experiment with outcome equal to x. Compute π_{post} .

$$T(n) = (n-1)!$$
, n is an integer.

$$\begin{bmatrix}
E[X] = \frac{\lambda}{\lambda+\beta} \\
Median = \frac{\lambda^{-\frac{1}{2}}}{\lambda+\beta^{-\frac{1}{2}}}
\end{bmatrix}$$

$$\begin{aligned} & = \int_{-\infty}^{\infty} \int_{0}^{\infty} (x) \, \pi_{prix}(\theta) \, d\theta \\ & = \int_{0}^{1} \left(\frac{n}{x} \right) \, \theta^{x} (1-\theta)^{n-x} \, \frac{T(x+\mu)}{T(x)T(\mu)} \, \theta^{x+\mu-1} \, d\theta \\ & = \left(\frac{n}{x} \right) \, \frac{\Gamma(x+\mu)}{\Gamma(x)T(\mu)} \, \int_{0}^{1} \, \frac{T(x+\mu)}{\Gamma(x+\mu)} \, \int_{0}^{1} \, \frac{\Gamma(x+\mu)}{\Gamma(x+\mu)} \, \theta^{x+\mu-1} \, d\theta \\ & = \left(\frac{n}{x} \right) \, \frac{\Gamma(x+\mu)}{\Gamma(x)T(\mu)} \, \frac{\Gamma(x+\mu)}{\Gamma(x+\mu)} \, \int_{0}^{1} \, \frac{\Gamma(x+\mu)}{\Gamma(x+\mu)} \, \theta^{x+\mu-1} \, d\theta \\ & = \left(\frac{n}{x} \right) \, \frac{\Gamma(x+\mu)}{\Gamma(x)T(\mu)} \, \frac{\Gamma(x+\mu)}{\Gamma(x+\mu)} \, \int_{0}^{1} \, \frac{\Gamma(x+\mu)}{\Gamma(x+\mu)} \, \theta^{x+\mu-1} \, d\theta \\ & = \left(\frac{n}{x} \right) \, \frac{\Gamma(x+\mu)}{\Gamma(x+\mu)} \, \frac{\Gamma(x+\mu)}{\Gamma(x+\mu)} \, \frac{\theta^{x+\mu-1}}{\Gamma(x+\mu)} \, \frac{\theta$$

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Bayesian Estimator

In the scenario of Bayes inference, the estimate for the unknown parameter Θ is not a *fixed number*, but a *random variable*. However, there are situations in real life where we are asked to give a fixed number $\hat{\theta}$ as our estimate. $\min \mathbb{E}[\Theta - b]^2 \implies b = \mathbb{E}[\Theta]$

Then **Bayesian estimator** $\hat{\theta}$ would depend on the penalty for errors created by incorrect guesses:

- 1. The loss function is $\mathbb{E}[(\Theta \hat{\theta})^2]$, the **mean square error**. Then best guess $\hat{\theta}$ would be the mean of the posterior pdf.
- 2. The loss function is $\mathbb{E}[|\Theta \hat{\theta}|]$, the **mean absolute error**. Then best guess $\hat{\theta}$ would be the median of the posterior pdf.

Bayesian Estimator

Example 6.5

Let X be the binomial random variable with parameters n and θ . Let π_{prior} be the beta pdf with parameters α and β . Suppose that we have one sample with value x. Compute the Bayesian estimator $\hat{\theta}$ that minimizes

- ► mean square error; $\hat{\theta} = \frac{(\alpha + \alpha) + (n + \beta \alpha)}{(\alpha + \alpha)}$
- ightharpoonup mean absolute error, if $\alpha + x = n + \beta x = 1$.

Median =
$$\frac{2 + x - \frac{1}{5}}{(x+x) + (n+(5-x)) - \frac{2}{5}} = \frac{1 - \frac{1}{5}}{2 - \frac{2}{5}} = \frac{\frac{2}{3}}{\frac{4}{5}} = \frac{1}{2}$$

For an insurance Company, there are two risk classes. prior 35% are smokers S
65% are non-smokers. S' Annual claim count for each policy followes a Poisson distribution . For smokers, & has value of s. For non-smokers, I has a value of 1 data / ols N=1. N=0 A randomly selected insured has 3 claims in Year I and no Claim in Year 2. Calculate the expected number of claims by this insured in Years. $\mathbb{P}(\text{data}|S) = \mathbb{P}(N_1 = 3|S) \cdot \mathbb{P}(N_2 = 0|S) = \frac{e^{-3}3^{\circ}}{3!} \cdot \frac{e^{-5}3^{\circ}}{0!} = 0.011$ P(S) = 0.45 $P(S') = 0.65 \in priors.$ $P(data|S') = P(N_1=5|S')P(N_2=0|S') = \frac{e^{-1}1^3}{51} \frac{e^{-1}1^0}{0!} = 0.01256$ K = total prob = P(data) = 0.011.0.35 + 0.02256.0.65 = 0.01857 Posterior = P(S|data) = 0.2| P(5'1 data) = 0.0256.0.65 = 0.79.

 $E[N_3|data] = E[N_3|S] \cdot P(S|data) + E[N_3|S'] - P(S'|data)$ = 3.0.21 + 1.0.78 = 1.4206