Likelihood of Heirarchical Matern Gaussian Random Field

We model trait values of individuals as independent normally distributed variables around a mean \bar{z} with variance v. In turn the mean \bar{z} is a function of location $\mathbf{x} = (x_1, x_2)^{\top}$ and follows a stationary Gaussian random field with spatially homogeneous mean \tilde{z} and Matern covariance C(d) with collocated variance V = C(0), characteristic length ξ and smoothness parameter $\nu = 1$. Hence, observed individual trait values z_1 and z_2 at locations $\mathbf{x}_1, \mathbf{x}_2$ separated at a distance d will be bivariate normally distributed with the following statistical properties:

$$\mathbb{E}z_1 = \mathbb{E}z_2 = \tilde{z}, \ \mathbb{V}z_1 = \mathbb{V}z_2 = v + V, \mathbb{C}(z_1, z_2) = C(d),$$

$$\mathbb{E}(z_i|\bar{z}) = \bar{z}(\boldsymbol{x}_i), \ \mathbb{V}(z_i|\bar{z}) = v, \ \mathbb{C}(z_1, z_2|\bar{z}) = 0.$$

Since we only get a single sample of this random field, the likelihood function is just

$$L(\tilde{z}, v, V, \xi | \boldsymbol{z}) = \frac{1}{\sqrt{(2\pi)^n |\Sigma|}} \exp\left(-\frac{1}{2}(\boldsymbol{z} - \tilde{\boldsymbol{z}}) \Sigma^{-1} (\boldsymbol{z} - \tilde{\boldsymbol{z}})^{\top}\right),$$

where $\mathbf{z} = (z_1, \dots, z_n)^{\top}$ is the vector of observed individual trait values (here there are n individuals) observed at locations $\mathbf{z}_1, \dots, \mathbf{z}_n$, $\tilde{\mathbf{z}}$ is the vector of expected trait values $(\tilde{z}, \dots, \tilde{z})^{\top}$ and Σ is a covariance matrix with v + V on the diagonal and $C(\|\mathbf{z}_i - \mathbf{z}_i\|)$ on the i, jth off-diagonal entry.

To maximize likelihood, we first compute gradients of the log-likelihood $\ell = \ln L$ with respect to model parameters. We then solve for the parameter values when those gradients are zero:

$$\begin{split} \frac{\partial \ell}{\partial \tilde{z}} &= -\frac{1}{2} \frac{\partial}{\partial \tilde{z}} \left[(\boldsymbol{z} - \tilde{\boldsymbol{z}}) \boldsymbol{\Sigma}^{-1} (\boldsymbol{z} - \tilde{\boldsymbol{z}})^{\top} \right] = -\frac{1}{2} \frac{\partial}{\partial \tilde{z}} \sum_{ij} (\sigma^{-1})_{ij} (z_i - \tilde{z}) (z_j - \tilde{z}) \\ &= \sum_{ij} (\sigma^{-1})_{ij} \left(\frac{z_i + z_j}{2} - \tilde{z} \right) = \sum_{ij} \frac{z_i + z_j}{2} (\sigma^{-1})_{ij} - \tilde{z} \sum_{ij} (\sigma^{-1})_{ij}, \\ &\implies \hat{\tilde{z}} = \frac{\sum_{ij} \frac{z_i + z_j}{2} (\sigma^{-1})_{ij}}{\sum_{ij} (\sigma^{-1})_{ij}}. \end{split}$$

Unfortunately, the rest is analytically intractable. So we proceed via numerical optimization to maximize likelihood... in julia... then import the results below.





