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# THE PHENOTYPIC SIGNATURE OF HOST-PARASITE COEVOLUTION IN CONTINUOUS SPACE

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A PREPRINT

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January 30, 2021

## Abstract

Here we identify the phenotypic signature of host-parasite coevolution in continuous space.

**Keywords** blah · blee · bloo · these are optional and can be removed

## 1 Introduction

Space is hypothesized to play a fundamental role in the coevolutionary process. In particular, the combined effects of local co-adaptation and spatial movement are thought to produce complex spatial patterns of matched and mismatched traits. However, theoretical studies investigating the spatial signal of coevolution have for the most part restricted themselves to understanding patterns of spatial correlation using spatially implicit models. Hence, the fine-grained details that can be produced in spatially explicit settings have yet to be understood. Here we close this gap in the context of host-parasite coevolution using continuous-space quantitative genetic models. In particular, we study the autocorrelations of mean traits within species and the cross-correlation of mean traits between species when trait dynamics are driven by (1) random genetic drift, gene-flow and spatially homogeneous abiotic stabilizing selection, (2) random genetic drift, gene-flow, spatially homogeneous stabilizing selection and unilateral adaptation of the host species, (3) random genetic drift, gene-flow and unilateral adaptation of the parasite species and (4) random genetic drift, gene-flow, stabilizing selection and host-parasite coevolution. By comparing patterns of spatial autocorrelation and cross-correlation in these four scenarios we are able to elucidate the signature of coevolution in comparison to underlying non-coevolutionary processes.

## 2 The Model

### 2.1 Fitness

The host is denoted  $H$  and the parasite  $P$ . Individual trait values of hosts and parasites are respectively denoted by  $z_H, z_P \in \mathbb{R}$ . Assuming a trait-matching model, where host fitness is minimized and parasite fitness is maximized when  $z_H = z_P$ , we have the following two fitness functions:

$$w_H(z_H, z_P) \propto \exp\left(-\frac{A_H}{2}(\theta_H - z_H)^2 + \frac{B_H}{2}(z_P - z_H)^2\right), \quad (1a)$$

$$w_P(z_P, z_H) \propto \exp\left(-\frac{A_P}{2}(\theta_P - z_P)^2 - \frac{B_P}{2}(z_H - z_P)^2\right), \quad (1b)$$

where  $A_H, A_P > 0$  capture the strengths of abiotic stabilizing selection,  $\theta_H, \theta_P \in \mathbb{R}$  are the abiotic optimal phenotypes and  $B_H, B_P > 0$  capture the strengths of biotic selection experienced by each species.

## 2.2 Non-spatial Dynamics

Using a standard approach, we use the above fitness functions to derive the following non-spatial coevolutionary model:

$$\frac{d\bar{z}_H}{dt} = G_H A_H(\theta_H - \bar{z}_H) - G_H B_H(\bar{z}_P - \bar{z}_H), \quad (2a)$$

$$\frac{d\bar{z}_P}{dt} = G_P A_P(\theta_P - \bar{z}_P) + G_P B_P(\bar{z}_H - \bar{z}_P), \quad (2b)$$

where  $G_H, G_P > 0$  denote the additive genetic variances of each species. In this linear set of equations, when abiotic stabilizing selection is absent ( $A_H = A_P = 0$ ), the parasite mean trait  $\bar{z}_P$  evolves to match the host mean trait  $\bar{z}_H$ . Simultaneously  $\bar{z}_H$  evolves away from  $\bar{z}_P$ . In particular, when  $\bar{z}_H > \bar{z}_P$ , the host mean trait will evolve upwards. Similarly, when  $\bar{z}_H < \bar{z}_P$ , the host mean trait will evolve downwards. This and related models have been thoroughly studied by Sergey Gavrillets and others. For the case when  $A_H = A_P = 0$  two possible outcomes have been demonstrated. First, when  $G_P B_P > G_H B_H$  so that the parasite evolves faster than the host, the system asymptotically evolves to the equilibrium  $\bar{z}_H = \bar{z}_P$ . Second, when  $G_P B_P < G_H B_H$  so that the host evolves faster than the parasite, there is no stable equilibrium. This implies the host evolves to escape the interaction.

## 2.3 Space

In this case mean traits become functions of spatial variables (eg.,  $\bar{z}_H(x), \bar{z}_P(x)$  for  $x \in \mathbb{R}$  and  $\bar{z}_H(x_1, x_2), \bar{z}_P(x_1, x_2)$  for  $(x_1, x_2) \in \mathbb{R}^2$ ). As a first step, we assume all model parameters are spatially homogeneous. If we assume interactions occur locally (individuals interact only when they ‘collide’) and abundance densities are spatially homogeneous, then we can obtain the following continuous space model:

$$\frac{\partial \bar{z}_H}{\partial t} = G_H A_H(\theta_H - \bar{z}_H) - G_H B_H(\bar{z}_P - \bar{z}_H) + \frac{D_H}{2} \Delta \bar{z}_H, \quad (3a)$$

$$\frac{\partial \bar{z}_P}{\partial t} = G_P A_P(\theta_P - \bar{z}_P) + G_P B_P(\bar{z}_H - \bar{z}_P) + \frac{D_P}{2} \Delta \bar{z}_P, \quad (3b)$$

where  $D_H, D_P > 0$  are dispersal rates for each species and  $\Delta$  is the diffusion operator representing spatial movement. In particular,  $\Delta = \frac{\partial^2}{\partial x^2}$  in one-dimensional space and  $\Delta = \frac{\partial^2}{\partial x_1^2} + \frac{\partial^2}{\partial x_2^2}$  in two-dimensional space.

## 2.4 Random Genetic Drift

To incorporate the effects of random genetic drift, we should be able to justify the following SPDE model:

$$\frac{\partial \bar{z}_H}{\partial t} = G_H A_H(\theta_H - \bar{z}_H) - G_H B_H(\bar{z}_P - \bar{z}_H) + \frac{D_H}{2} \Delta \bar{z}_H + \sqrt{\frac{G_H}{N_H}} \dot{W}_H, \quad (4a)$$

$$\frac{\partial \bar{z}_P}{\partial t} = G_P A_P(\theta_P - \bar{z}_P) + G_P B_P(\bar{z}_H - \bar{z}_P) + \frac{D_P}{2} \Delta \bar{z}_P + \sqrt{\frac{G_P}{N_P}} \dot{W}_P, \quad (4b)$$

where  $N_P, N_H$  are the local abundance densities (here assumed to be constant in space and time) and  $\dot{W}_H, \dot{W}_P$  are space-time white noise processes. This SPDE model may possess some major benefits over more traditional continuous-space models of population genetics. In particular, it is a simple extension of the stochastic heat equation, which is known to admit solutions in dimensions one, two and three.

### 3 Methods

To study the continuous-space signature of coevolution, we can start with an analytical approach. However, if the corresponding calculations prove to be analytically intractable or time costly, we can employ simulation-based methods to describe the shape of the autocorrelations and cross-correlation in relevant regions of parameters space. The analytical approach would begin by identifying stable attractors of the deterministic system using classical PDE theory. Following this, we may possibly apply theory of SPDE to describe stationary distributions associated with each attractor. This may lend analytical solutions for the expectation and variance of intraspecific autocorrelations and the interspecific cross-correlation, which are defined below.

#### 3.1 Definitions of Spatial Autocorrelation and Cross-correlation

##### 3.1.1 Autocorrelation

In one-dimensional continuous space, the autocorrelation of a function  $f$  is defined by

$$\rho_f(x) = \int_{-\infty}^{+\infty} f(y)f(y-x)dy. \quad (5)$$

Sometimes the autocorrelation  $\rho_f(x)$  is written as  $(f \star f)(x)$ . To extend this definition to two-dimensional space, we can calculate a bivariate autocorrelation function of  $f$ ,

$$\rho_f(x_1, x_2) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(y_1, y_2)f(y_1 - x_1, y_2 - x_2)dy_1dy_2, \quad (6)$$

and then average this quantity across distances  $r = \sqrt{x_1^2 + x_2^2}$ :

$$\bar{\rho}_f(r) = \frac{1}{2\pi r} \int_0^{2\pi} \rho_f(r \cos \theta, r \sin \theta)d\theta. \quad (7)$$

##### 3.1.2 Cross-correlation

The cross-correlation of two functions  $f, g$  is defined by

$$\rho_{fg}(x) = \int_{-\infty}^{+\infty} f(y)g(y-x)dy. \quad (8)$$

Sometimes this is written as  $f \star g$ . Just as with the bivariate autocorrelation above, we can extend this definition to two-dimensional space. In particular, the bivariate cross-correlation of two functions  $f, g$  can be written

$$\rho_{fg}(x_1, x_2) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(y_1, y_2)g(y_1 - x_1, y_2 - x_2)dy_1dy_2. \quad (9)$$

We can also reduce this quantity to a function of distance using the average

$$\bar{\rho}_{fg}(r) = \frac{1}{2\pi r} \int_0^{2\pi} \rho_{fg}(r \cos \theta, r \sin \theta)d\theta. \quad (10)$$