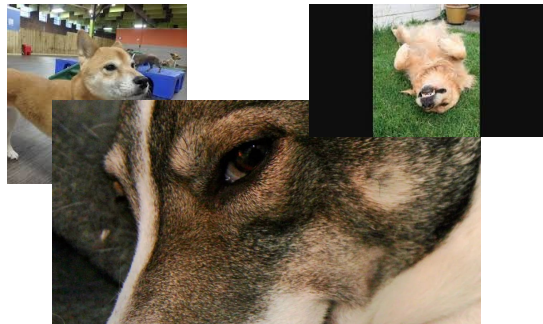
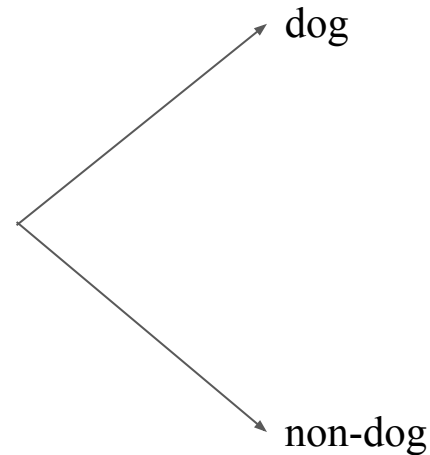
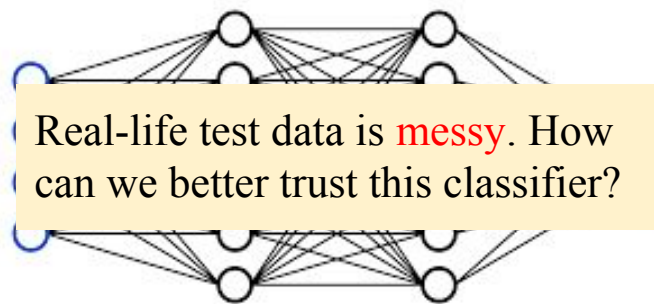




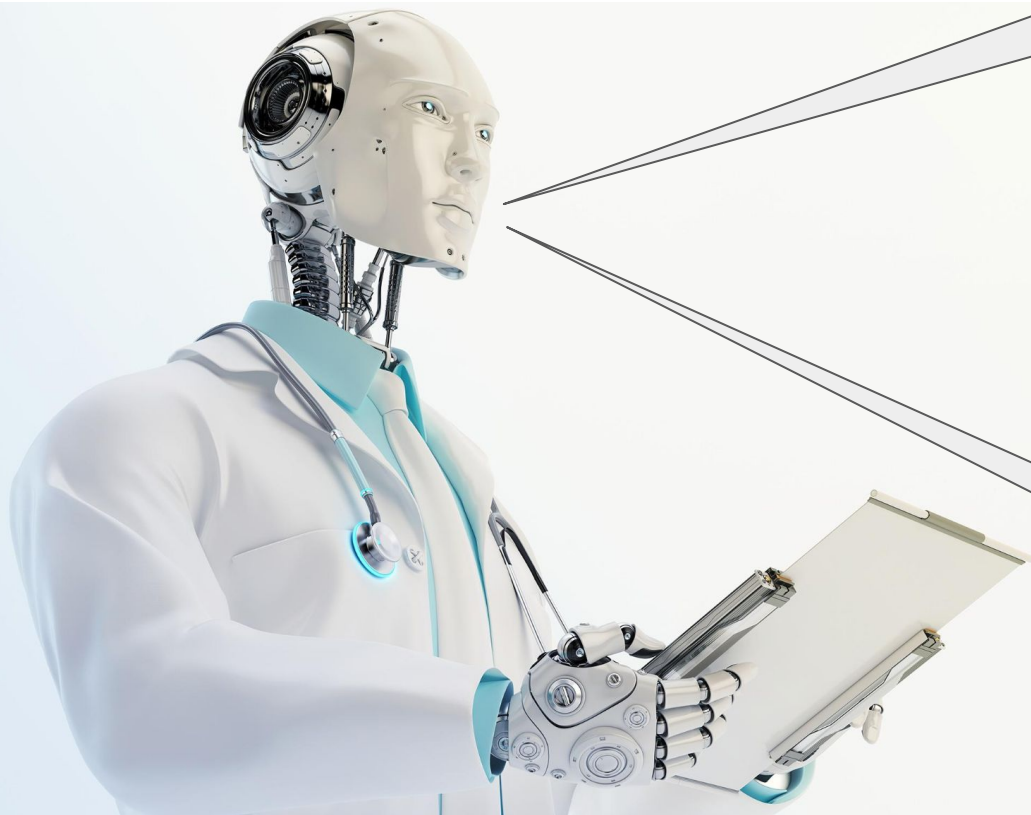
Lecture 04:

“Uncertainty, Bootstrapping”

An Uncertain World



An Uncertain World



You have the
disease.

You have the
disease. I have
30% confidence
that this is true.

Uncertainty at 3 Points

Training data

$$\mathcal{D} = \left\{ (x_1, y_1), (x_2, y_2), \dots, (x_n, y_n) \right\}$$

Model 1

$$\hat{y}_i = f(x_i, \theta)$$
$$L(y, f(x, \theta))$$

Parameter estimate

$$\hat{\theta}$$

Prediction

$$\hat{y}_i = f(x_i, \hat{\theta})$$

Test Loss

$$L(y, f(x_i, \hat{\theta}))$$

Model 2

$$\hat{y}_i = g(x_i, \varphi)$$
$$L(y, g(x, \varphi))$$

Parameter estimate

$$\hat{\varphi}$$

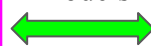
Prediction

$$\hat{y}_i = g(x_i, \hat{\varphi})$$

Test Loss

$$L(y, g(x_i, \hat{\varphi}))$$

Compare
Models



We want to know
uncertainty here

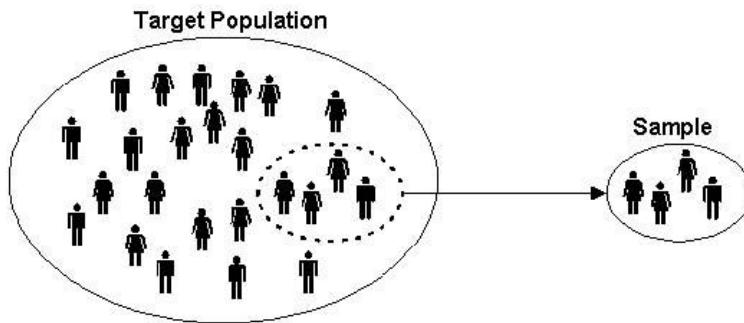
Test data

$$\mathcal{T} = \left\{ (x_1, y_1), \dots \right\}$$

Parameter Uncertainty

- Parameter = value which summarizes data for a population; these can be expectations (*mean*) or values which describe an input-output relationship (*slope of a linear model*)
- Statistic = value which summarizes data from a particular sample (*i.e. sample mean*).
- Estimation = use a *statistic* to estimate a *parameter* of the distribution of a random variable, where
 - Estimator ($\hat{\theta}$): function used to compute estimate
 - Estimand (θ): parameter of interest

μ_X = parameter

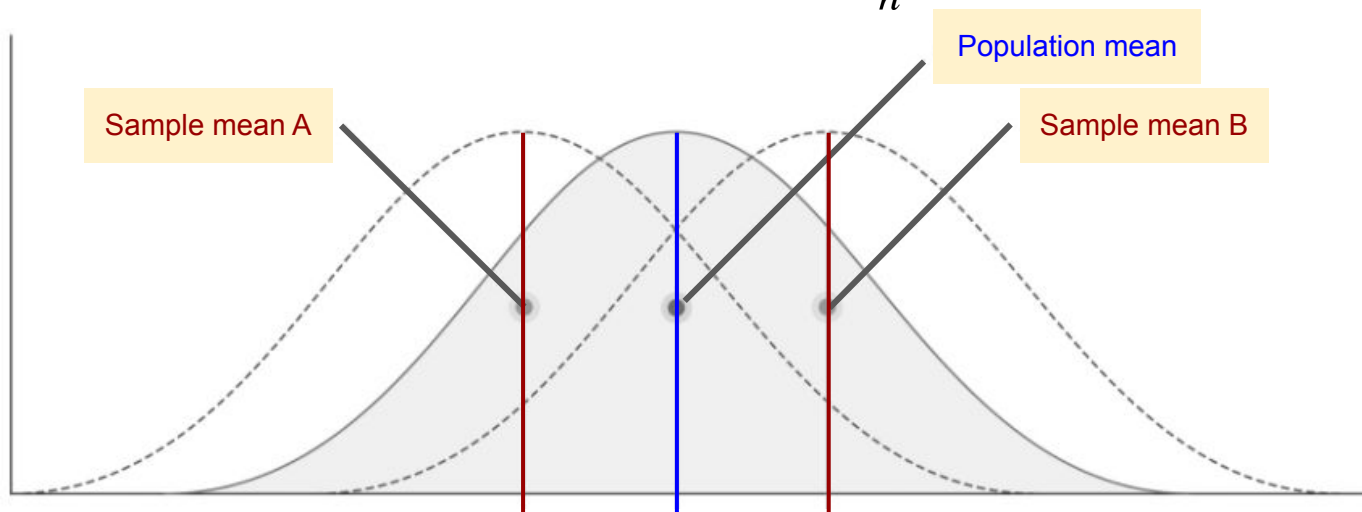


\bar{x}_n = statistic

Example of a parameter: mean

- Consider a model which predicts the mean... i.e. $\hat{y} = \theta$
- Given a dataset $\{x_1, x_2, \dots, x_n\}$, the estimate for this parameter is the sample mean:

$$\hat{\theta} = \frac{\sum_{i=1}^n x_i}{n}$$



The distribution of an estimator is called its **sampling distribution**.

Bias and Variance

- Bias = expected difference between estimator ($\hat{\theta}$) and parameter (θ)

In general:
$$\text{Bias}(\hat{\theta}) = E[\hat{\theta} - \theta]$$

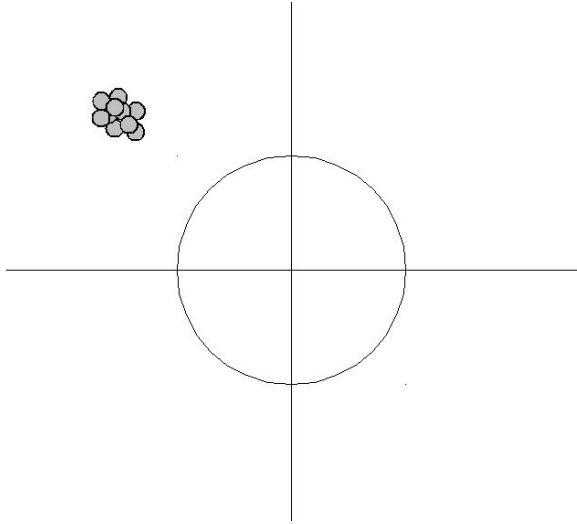
For example:
$$E[\bar{X}_n - \mu_X]$$

- Variance = expected squared difference between estimator ($\hat{\theta}$) and $E[\text{estimator}]$ (mean)

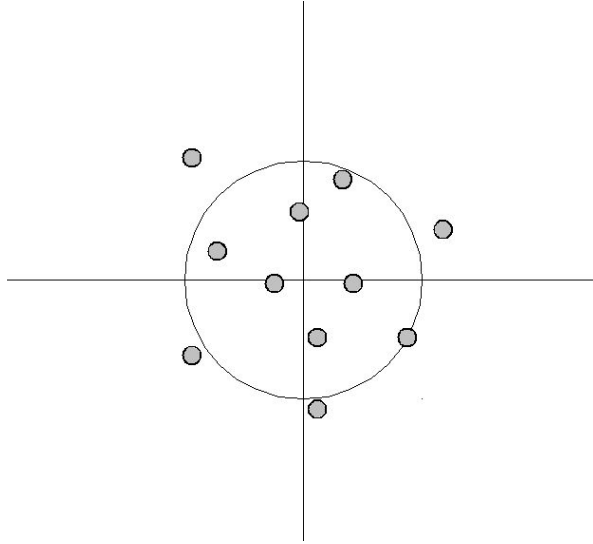
In general:
$$E[(\hat{\theta} - E[\hat{\theta}])^2]$$

For example:
$$E[(\bar{X}_n - E[\bar{X}_n])^2]$$

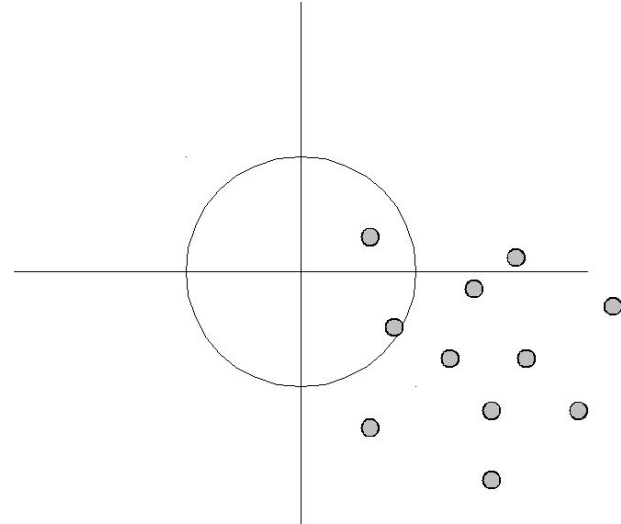
Bias and Variance



High bias, low variance



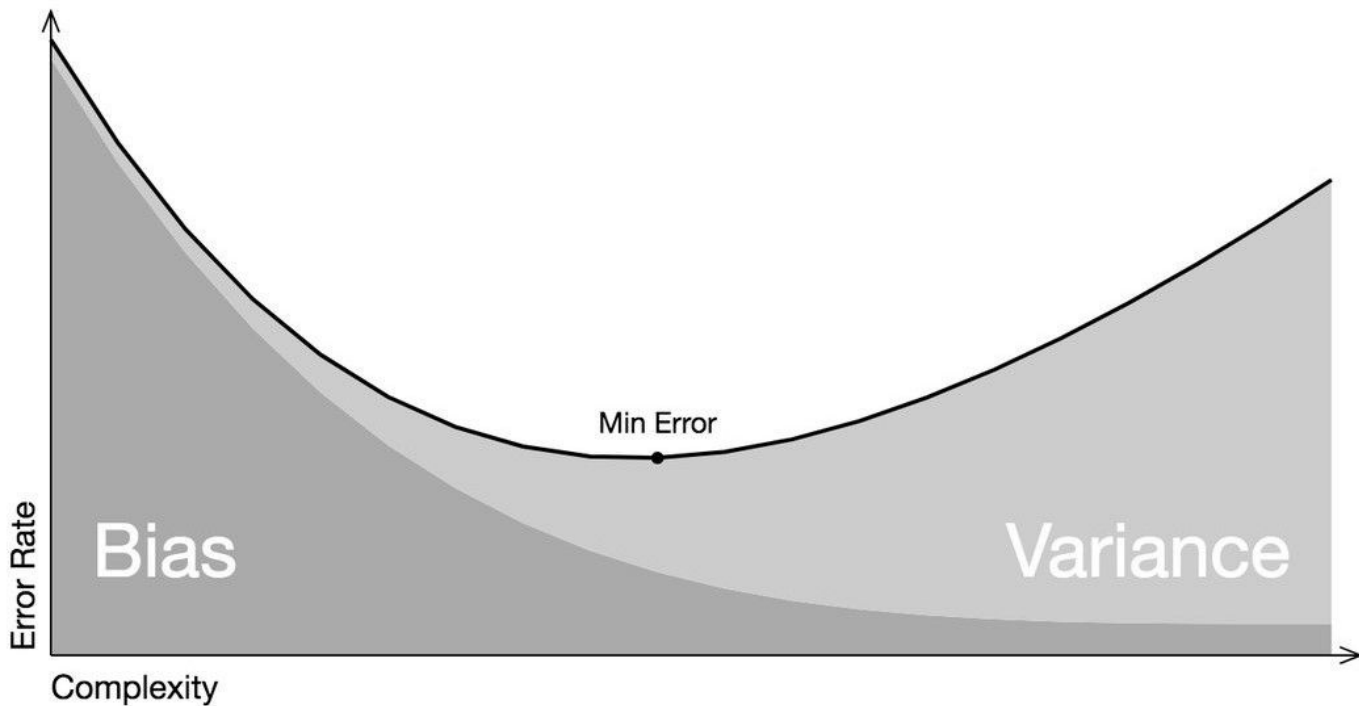
Low bias, high variance



High bias, high variance

Bias-Variance Tradeoff

Bias-Variance Tradeoff

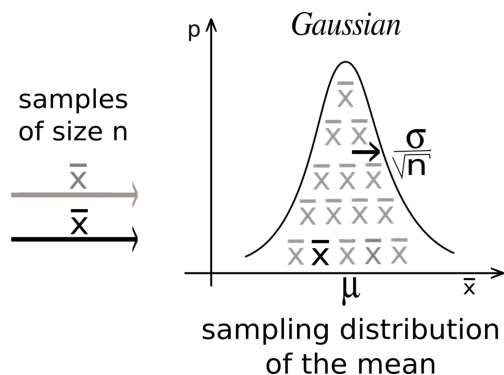
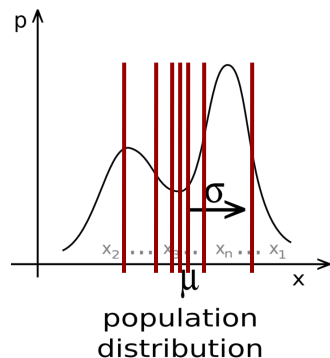


Central Limit Theorem (CLT)

- For large n , the sampling distribution of \bar{X}_n is approximately normal.
- Formally, we can write:

$$\bar{x}_n \sim N\left(\mu, \sigma^2_{\bar{X}_n}\right), \text{ where } \sigma_{\bar{X}_n} = \frac{\sigma_X}{\sqrt{n}}$$

Variance
Standard error



Whatever the form of the population distribution, the sampling distribution tends to a Gaussian, and its dispersion is given by the central limit theorem [1]

Central Limit Theorem (CLT)

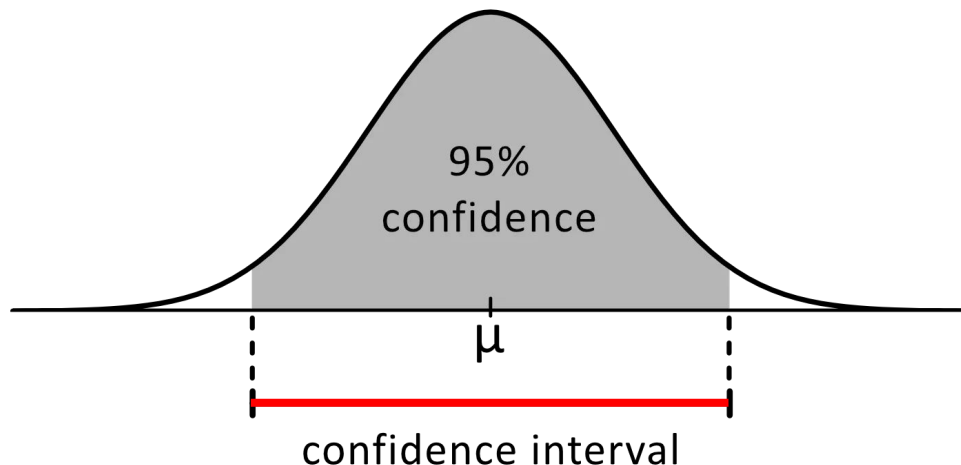
- We can use the CLT to construct **Confidence Intervals**

Question: What's a 95% confidence interval?

Answer: An interval which includes 95% of the sample means.

Another Answer: If we constructed this interval 100 times, it would contain the true mean in 95 of those instances.

Distribution of sample means (\bar{x})
around population mean (μ)

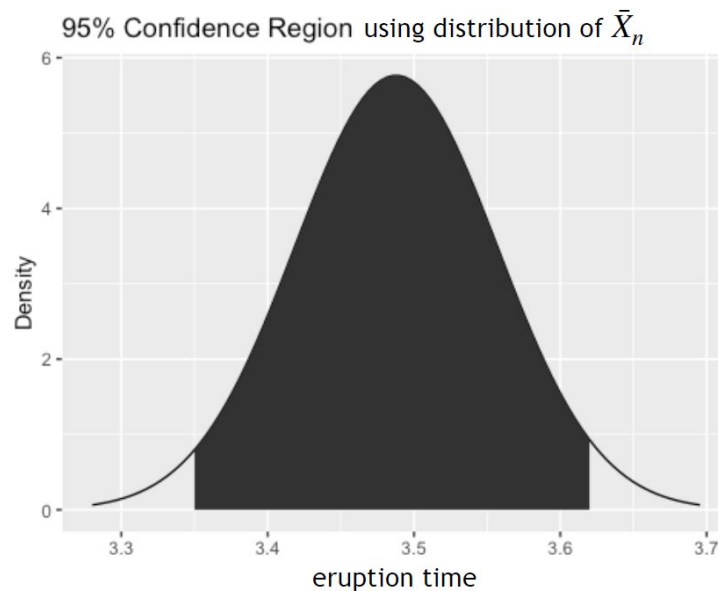


Central Limit Theorem (CLT)

- We can also say that 95% of the sample means are between $\mu - 1.96\sigma$ and $\mu + 1.96\sigma$
- Alternatively, 95% of the time the true mean μ will be between $\bar{x}_n - 1.96\sigma$ and $\bar{x}_n + 1.96\sigma$

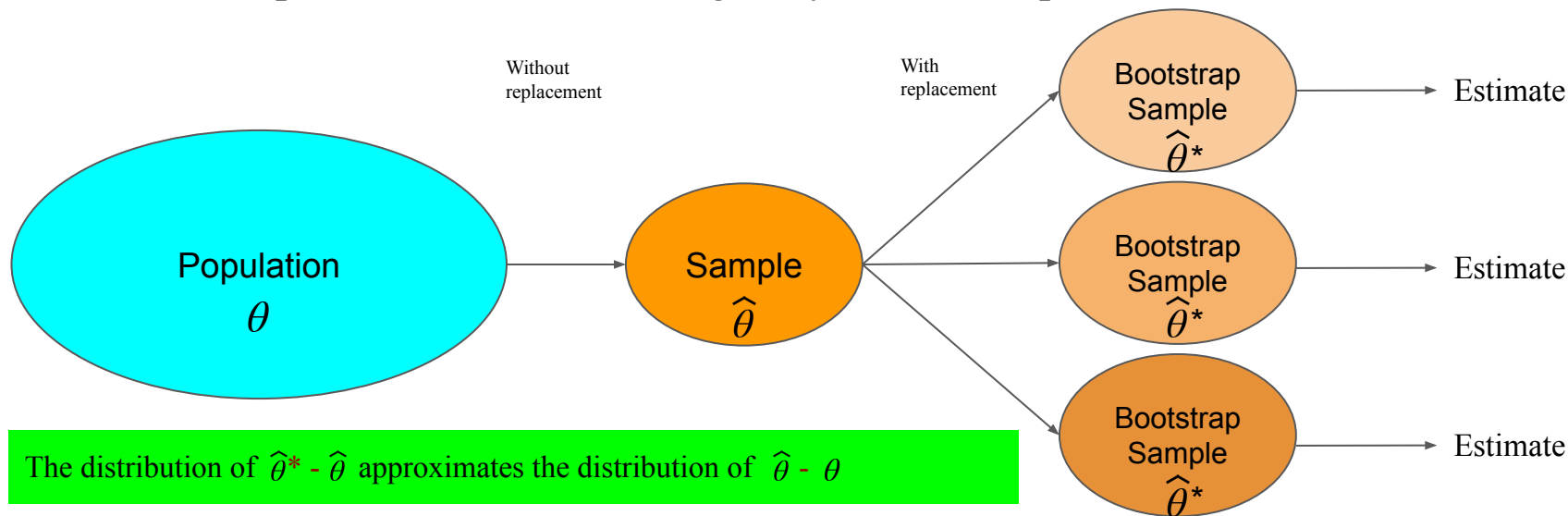
Example: Old Faithful

- Say we estimate that the mean value of eruption times is 3.4877831 (with $n=272$ observations)
- Is this a good estimate? How good is it?
- Mean = 3.49, Stdev = 1.14, SE = 0.07
- CI is therefore $3.49 \pm 1.96*(0.07) = 3.49 \pm 0.14$



The Bootstrap

- CLT excellent for datasets with approximately gaussian noise, and does a good job getting a distribution of parameter estimates. What if standard errors not normal?
- Bootstrap = a powerful technique to construct confidence intervals using artificially drawn samples in addition to an originally-drawn sample



The Bootstrap

```
Your Sample S has N observations
```

```
For b in 1:numBootstrap:
```

```
    resample N from S with replacement -> S*
```

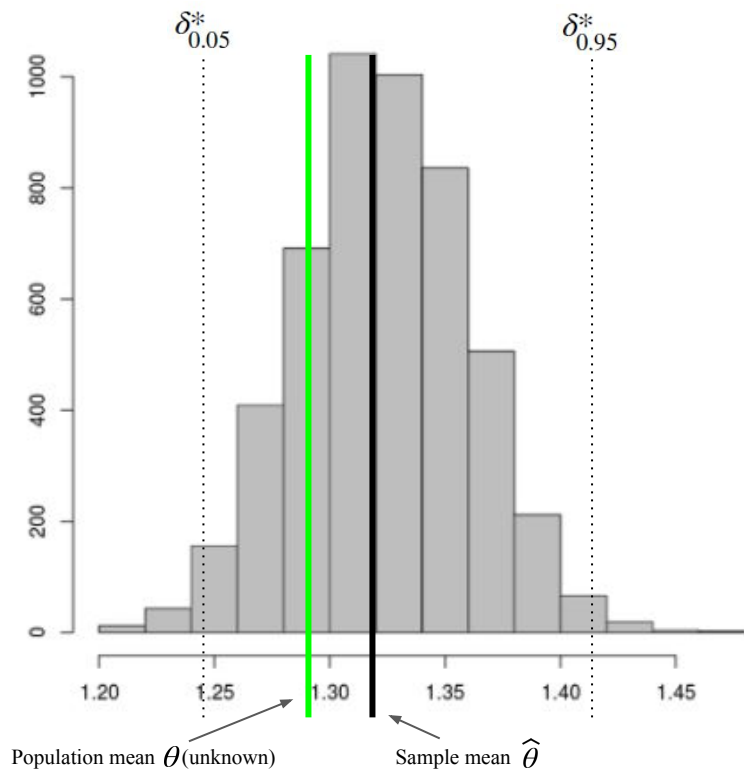
```
    Fit model to S* ->  $\hat{\theta}^*$  (bootstrap statistics)
```

```
    Record your bootstrap statistics
```

```
Return the distribution of bootstrap statistics|
```

The Bootstrap

- Now that we have a distribution of bootstrap statistics, we can construct a CI



- For example, a 90% confidence interval centred at the sample mean would be

$$CI = \left[\hat{\theta} - \delta_{0.95}^*, \hat{\theta} - \delta_{0.05}^* \right]$$

where $\delta^* = \hat{\theta}^* - \hat{\theta}$

and where $\delta_{0.95}^*$ is the 95% percentile of the bootstrap distribution

Prediction Uncertainty

Training data

$$\mathcal{D} = \left\{ \left(x_1, y_1 \right), \left(x_2, y_2 \right), \dots, \left(x_n, y_n \right) \right\}$$



Model

$$\hat{y}_i = f \left(x_i, \theta \right)$$

$$L \left(y, f \left(x, \theta \right) \right)$$



Parameter estimate

$$\hat{\theta}$$



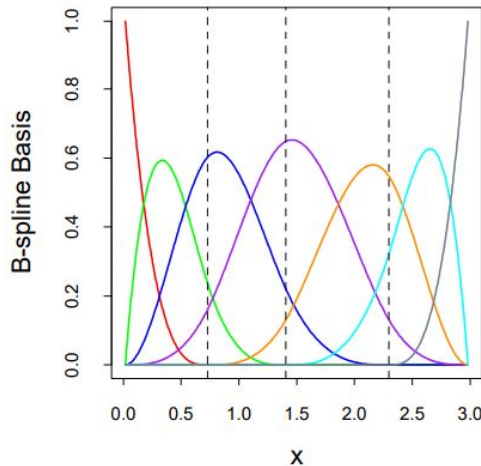
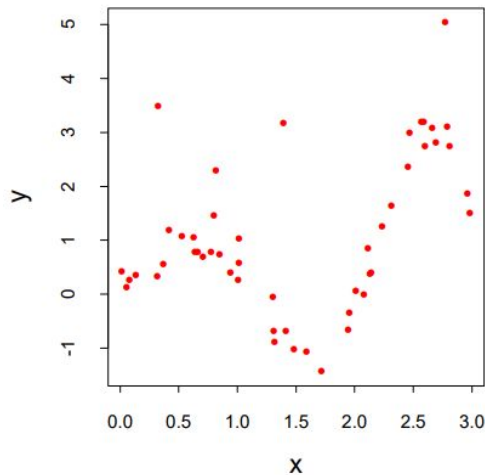
Prediction

$$\hat{y}_i = f \left(x_i, \hat{\theta} \right)$$

How does uncertainty in our parameter estimate influence the uncertainty of our prediction?

How much would the prediction change if we had used a different set of training data?

Prediction Uncertainty (Bootstrap)

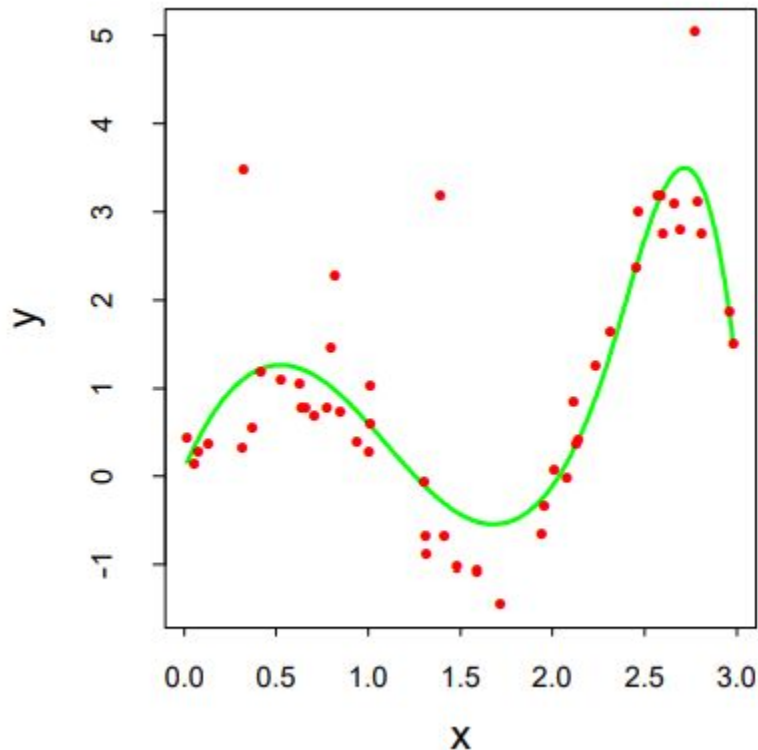


- We store the B coefficients of these basis functions into a vector θ , and fit $\hat{y} = f(x) = X\theta$.

$$\begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{pmatrix}_{n \times 1} = \begin{pmatrix} h_1(x_1) & h_2(x_1) & \cdots & h_p(x_1) \\ h_1(x_2) & h_2(x_2) & \cdots & h_p(x_2) \\ \vdots & \vdots & \ddots & \vdots \\ h_1(x_n) & h_2(x_n) & \cdots & h_p(x_n) \end{pmatrix}_{n \times p} \begin{pmatrix} \beta_1 \\ \vdots \\ \beta_p \end{pmatrix}_{p \times 1}$$

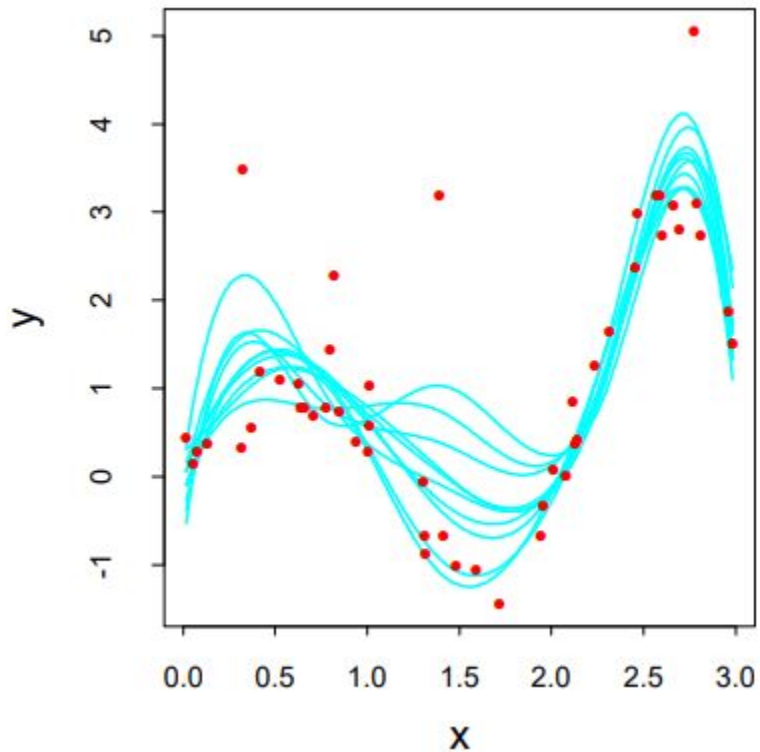
- Example: say we want to fit a cubic spline to this data. We can use a linear expansion of B-spline basis functions $h_i(x)$.

Prediction Uncertainty (Bootstrap)



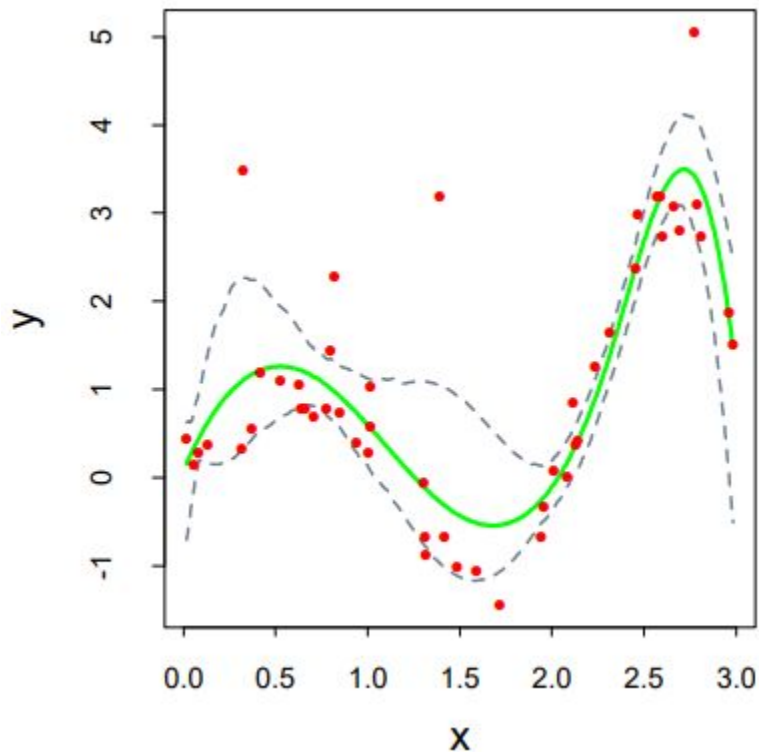
- Here is our fit $\hat{y} = \hat{f}(x) = X\hat{\theta}$
- Is it any good? **Yes**, but how confident can we be of this?
- **Let's use bootstrap:**
 - From our original sample, generate a new sample (with replacement)
 - For this new sample, get a new parameter estimate $\hat{\theta}_b^*$
 - Do this as many times as you can

Prediction Uncertainty (Bootstrap)



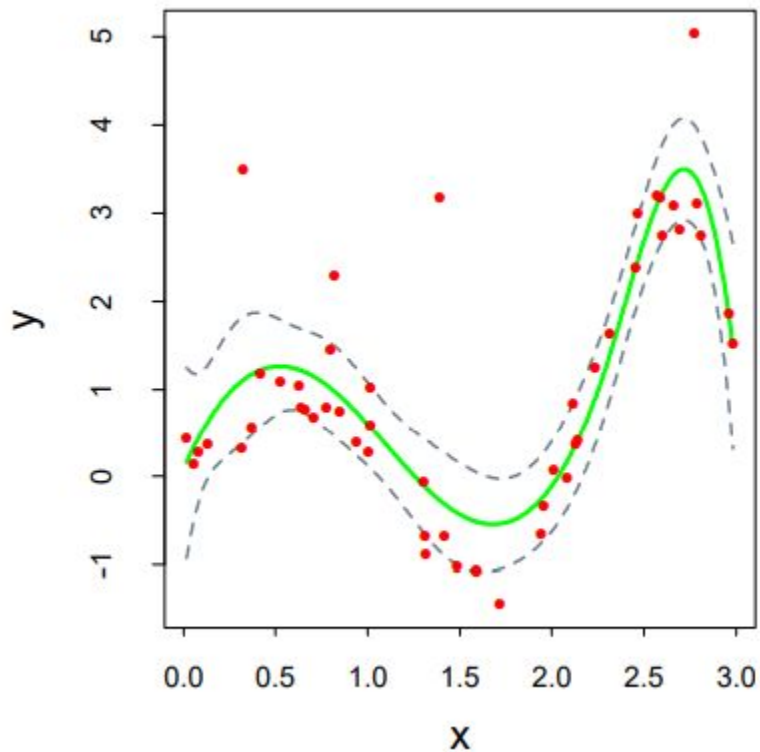
- We can plot each new prediction $\hat{y}_b^* = \hat{f}_b^*(x) = X\hat{\theta}_b^*$

Prediction Uncertainty (Bootstrap)



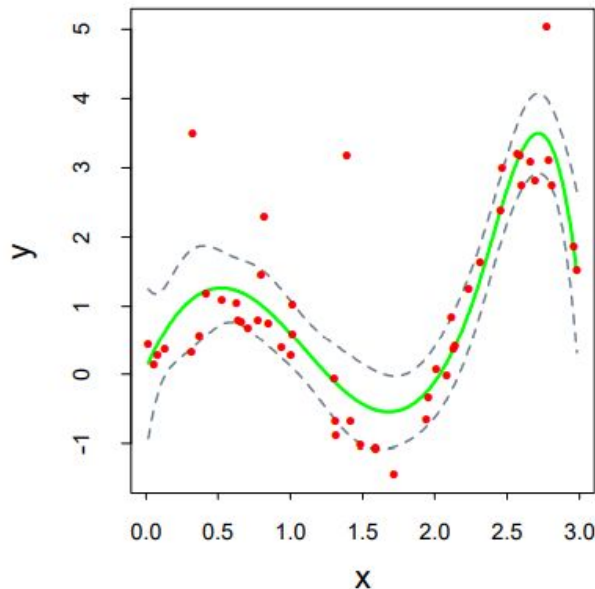
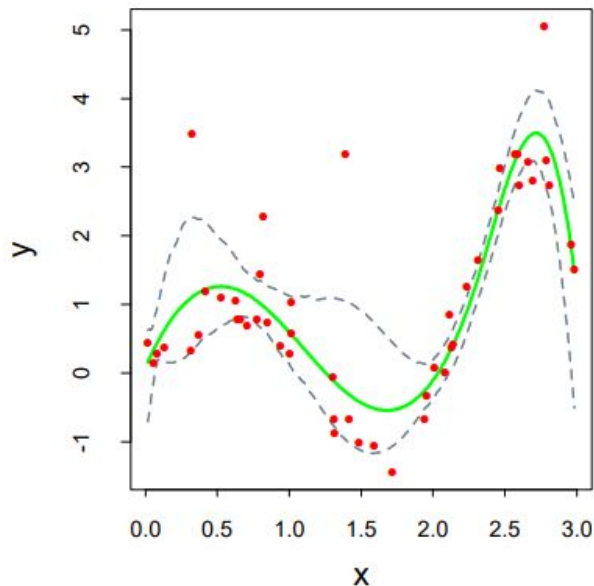
- And since we now have a distribution of samples for each x , we can compute a 95% Confidence Interval (CI)

Prediction Uncertainty (Bootstrap)



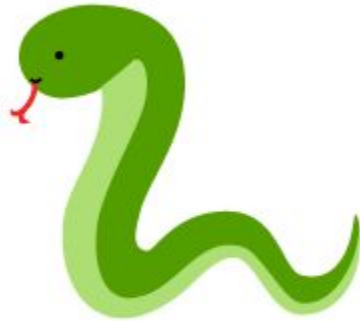
- Note that we could have also used CLT to get the CIs

Prediction Uncertainty (Bootstrap)



- Warning: our Confidence Interval (via bootstrap [left] or CLT [right]) is for the true value of $f(x)$, not for new observations ($x_{\text{new}}, y_{\text{new}}$)
- Why? A CI for new data would need to also consider random variability (σ^2) between $f(x_n)$ and y_n .

Let'sss try it in Python...



Summary

- Parameter Uncertainty
 - Parameters, Statistics, Estimation
 - Example using Population/Sample Mean
 - Bias and variance
 - The Central Limit Theorem (CLT)
 - Constructing a Confidence Interval (CI)
 - Bootstrap
- Prediction Uncertainty
 - B-spline example (Bootstrap)
- Coding examples of CLT, Bootstrap