Effect size and power

Session 8

MATH 80667A: Experimental Design and Statistical Methods
for Quantitative Research in Management
HEC Montréal

Outline

Effect sizes

Power

Effect size

Motivating example

Quote from the OSC psychology replication

> The key statistics provided in the paper to test the “depletion” hypothesis is the main effect of a one-way ANOVA with three experimental conditions and confirmatory information processing as the dependent variable; $F(2, 82) = 4.05$, $p = 0.02$, $\eta^2 = 0.09$. Considering the original effect size and an alpha of 0.05 the sample size needed to achieve 90% power is 132 subjects.

Replication report of Fischer, Greitemeyer, and Frey (2008, JPSP, Study 2) by E.M. Galliani

Translating statement into science

Q: How many observations should I gather to reliably detect an effect?

Q: How big is this effect?

Does it matter?

Statistical significance $\neq$ practical relevance

With large enough sample size, **any** sized difference between treatments becomes statistically significant.

But whether this is important depends on the scientific question.

Example

- What is the minimum difference between two treatments that would be large enough to justify commercialization of a drug?
- Tradeoff between efficacy of new treatment vs status quo, cost of drug, etc.

Using statistics to measure effects

Statistics and p-values are not good summaries of magnitude of an effect:

- the larger the sample size, the bigger the statistic, the smaller the p-value

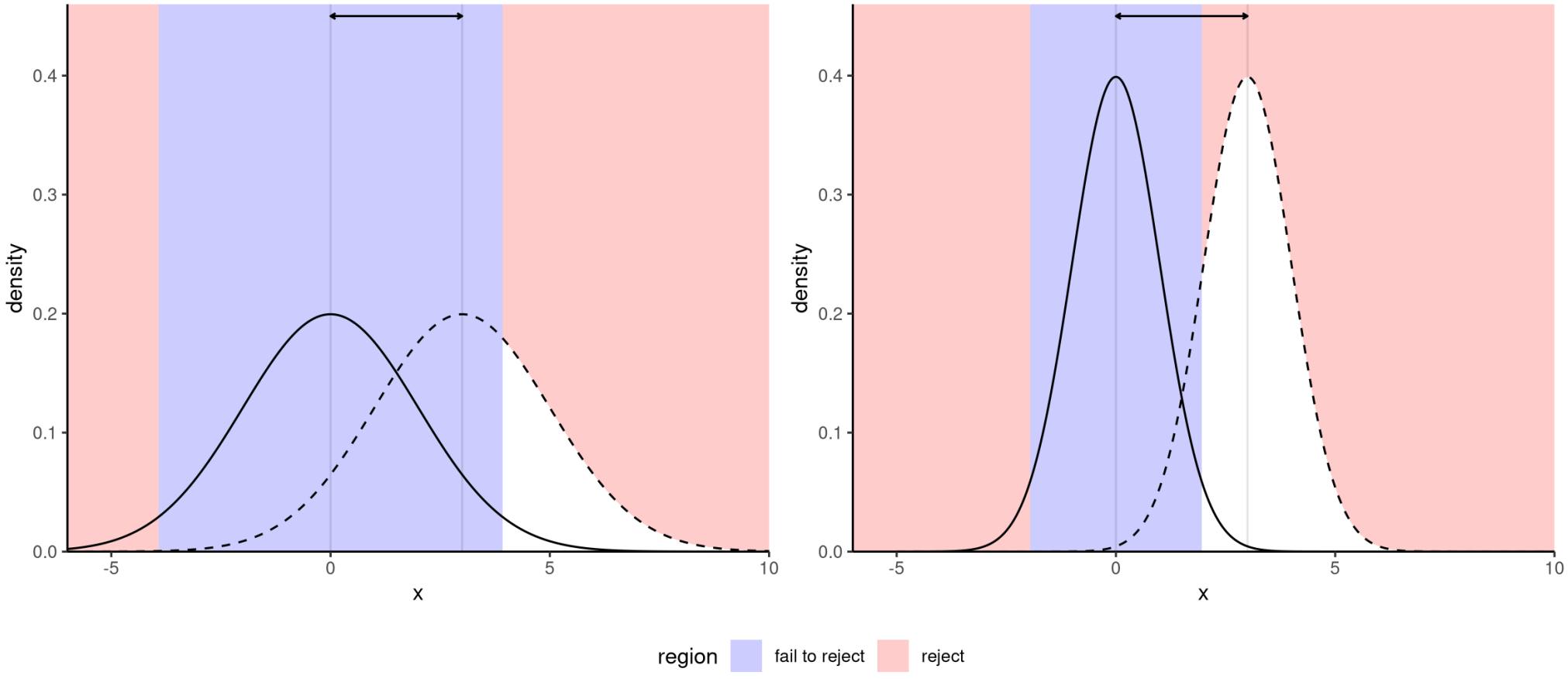
Instead use

standardized differences **percentage of variability explained**

Estimators popularized in the handbook

> Cohen, Jacob. Statistical Power Analysis for the Behavioral Sciences, 2nd ed., Routhledge, 1988.

Illustrating effect size (differences)



The plot shows null (thick) and true sampling distributions (dashed) for sample mean with small (left) and large (right) samples.

Estimands, estimators, estimates

- μ_i is the (unknown) population mean of group i (parameter, or estimand)
- $\widehat{\mu}_i$ is a formula (an estimator) that takes data as input and returns a numerical value (an estimate).
- throughout, use hats to denote estimated quantities:

Left to right: parameter μ (target), estimator $\widehat{\mu}$ (recipe) and estimate $\widehat{\mu} = 10$ (numerical value, proxy)

Cohen's *d*

Standardized measure of effect (dimensionless=no units):

Assuming equal variance σ^2, compare mean of two groups i and j:

$$d = \frac{\mu_i - \mu_j}{\sigma}$$

- Cohen's d's usual estimator $\hat{d}$ uses sample average of groups and the pooled variance estimator $\hat{\sigma}$.
- Hedge's g includes a finite sample correction and is less biased (*prefered in practice*)

Cohen's classification of effect sizes: small (*d=0.2*), medium (*d=0.5*) or large (*d=0.8*).

Cohen's f

For a one-way ANOVA (equal variance σ^2) with more than two groups,

$$f^2 = \frac{1}{\sigma^2} \sum_{j=1}^{k} \frac{n_j}{n} (\mu_j - \mu)^2,$$

a weighted sum of squared difference relative to the overall mean μ.

For $k = 2$ groups, Cohen's f and Cohen's d are related via $f = d/2$.

Effect size: proportion of variance

If there is a single experimental factor, use **total** effect size.

Break down the variability

$$\sigma^2_{\text{total}} = \sigma^2_{\text{resid}} + \sigma^2_{\text{effect}}$$

and define the percentage of variability explained by the $_{\text{effect}}$.

$$\eta^2 = \frac{\text{explained variability}}{\text{total variability}} = \frac{\sigma^2_{\text{effect}}}{\sigma^2_{\text{total}}}.$$

Coefficient of determination estimator

For the balanced one-way ANOVA, typical estimator is the **coefficient of determination**

$$\widehat{R}^2 = \frac{F\nu_1}{F\nu_1 + \nu_2}$$

where $\nu_1 = K - 1$ and $\nu_2 = n - K$ are the degrees of freedom for the one-way ANOVA with n observations and K groups.

- $\widehat{R}^2$ is an upward biased estimator (too large on average).
- People frequently write η^2 when they mean $\widehat{R}^2$
- for the replication, $\widehat{R}^2 = (4.05 \times 2)/(4.05 \times 2 + 82) = 0.09$

ω^2 square estimator

Another estimator of η^2 that is recommended in Keppel & Wickens (2004) for power calculations is $\widehat{\omega}^2$.

For one-way ANOVA, the latter is obtained from the F-statistic as

$$\widehat{\omega}^2 = \frac{\nu_1(F-1)}{\nu_1(F-1)+n}$$

- for the replication, $\widehat{\omega}^2 = (3.05 \times 2)/(3.05 \times 2 + 84) = 0.0677.$
- if the value returned is negative, report zero.

Converting η^2 to Cohen's f

Software usually take Cohen's f (or f^2) as input for the effect size.

Convert from η to f via the relationship

$$f^2 = \frac{\eta^2}{1 - \eta^2}.$$

If we plug-in estimated values

- with $\widehat{R}^2$, we get $\widehat{f} = 0.314$
- with $\widehat{\omega}^2$, we get $\tilde{f} = 0.27$.

Effect sizes for multiway ANOVA

With a completely randomized design with only experimental factors, use **partial** effect size

$$\eta^2_{\langle \text{effect} \rangle} = \sigma^2_{\text{effect}} / (\sigma^2_{\text{effect}} + \sigma^2_{\text{resid}})$$

In **R**, use `effectsize::omega_squared(model, partial = TRUE)`.

Partial effects and variance decomposition

Consider a completely randomized balanced design with two factors A, B and their interaction AB. We can decompose the total variance as

$$\sigma^2_{\text{total}} = \sigma^2_A + \sigma^2_B + \sigma^2_{AB} + \sigma^2_{\text{resid}}.$$

Cohen's partial f measures the proportion of variability that is explained by a main effect or an interaction, e.g.,

$$f_{\langle A \rangle} = \frac{\sigma^2_A}{\sigma^2_{\text{resid}}}, \qquad f_{\langle AB \rangle} = \frac{\sigma^2_{AB}}{\sigma^2_{\text{resid}}}.$$

Partial effect size (variance)

Effect size are often reported in terms of variability via the ratio

$$\eta^2_{\langle \text{effect} \rangle} = \frac{\sigma^2_{\text{effect}}}{\sigma^2_{\text{effect}} + \sigma^2_{\text{resid}}}.$$

- Both $\hat{\eta}^2_{\langle \text{effect} \rangle}$ (aka $\widehat{R}^2_{\langle \text{effect} \rangle}$) and $\widehat{\omega}^2_{\langle \text{effect} \rangle}$ are **estimators** of this quantity and obtained from the F statistic and degrees of freedom of the effect.

Estimation of partial ω^2

Similar formulae as the one-way case, with

$$\widehat{\omega}^2_{\langle \text{effect} \rangle} = \frac{\text{df}_{\text{effect}}(F_{\text{effect}} - 1)}{\text{df}_{\text{effect}}(F_{\text{effect}} - 1) + n},$$

where n is the overall sample size.

In **R**, `effectsize::omega_squared` reports these estimates with one-sided confidence intervals.

Reference for confidence intervals: Steiger (2004), Psychological Methods

Converting ω^2 to Cohen's f

Given an estimate of $\eta^2_{\langle \text{effect} \rangle}$, convert it into an estimate of Cohen's partial $f^2_{\langle \text{effect} \rangle}$, e.g.,

$$\widehat{f}^2_{\langle \text{effect} \rangle} = \frac{\widehat{\omega}^2_{\langle \text{effect} \rangle}}{1 - \widehat{\omega}^2_{\langle \text{effect} \rangle}}.$$

The package `effectsize::cohens_f` returns $\tilde{f}^2 = n^{-1} F_{\text{effect}} \text{df}_{\text{effect}}$, a transformation of $\hat{\eta}^2_{\langle \text{effect} \rangle}$.

Semipartial effect sizes

If there is a mix of experimental and blocking factor...

Include the variance of all blocking factors and interactions (only with the effect!) in denominator.

- e.g., if A is effect of interest, B is a blocking factor and C is another experimental factor, use

$$\eta^2_{\langle A \rangle} = \frac{\sigma^2_A}{\sigma^2_A + \sigma^2_B + \sigma^2_{AB} + \sigma^2_{\text{resid}}}.$$

In **R**, use `effectsize::omega_squared(model, partial = TRUE, generalized = "blocking")` where `blocking` gets replaced with a vector containing the name of the blocking factors.

Summary

- Effect sizes can be recovered using information found in the ANOVA table.
- Multiple estimators for the same quantity
  - report the one used along with confidence or tolerance intervals.
- The correct measure may depend on the design
  - partial vs total effects,
  - different formulas for within-subjects (repeated measures) designs!
- Include blocking as part of the variability considered.

Power

Power and sample size calculations

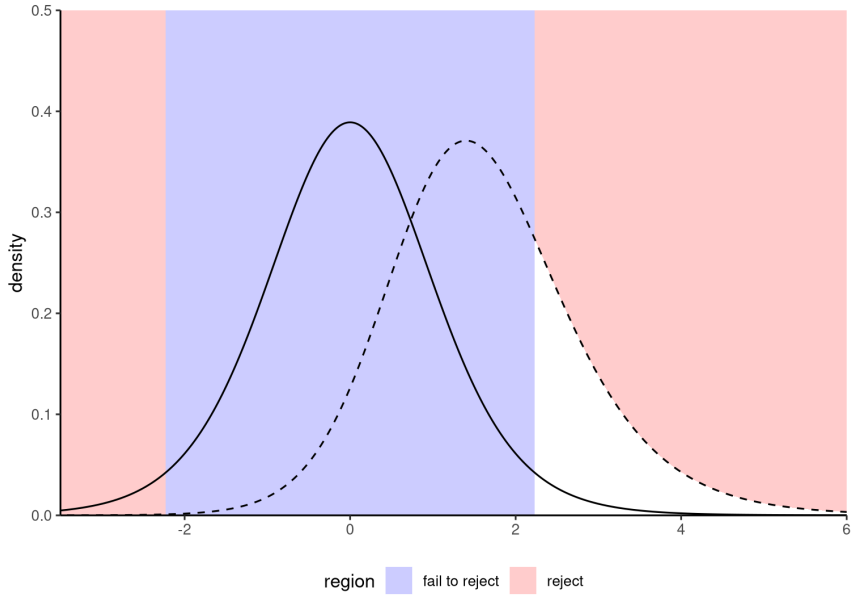
Journals and grant agencies oftentimes require an estimate of the sample size needed for a study.

- large enough to pick-up effects of scientific interest (good signal-to-noise)
- efficient allocation of resources (don't waste time/money)

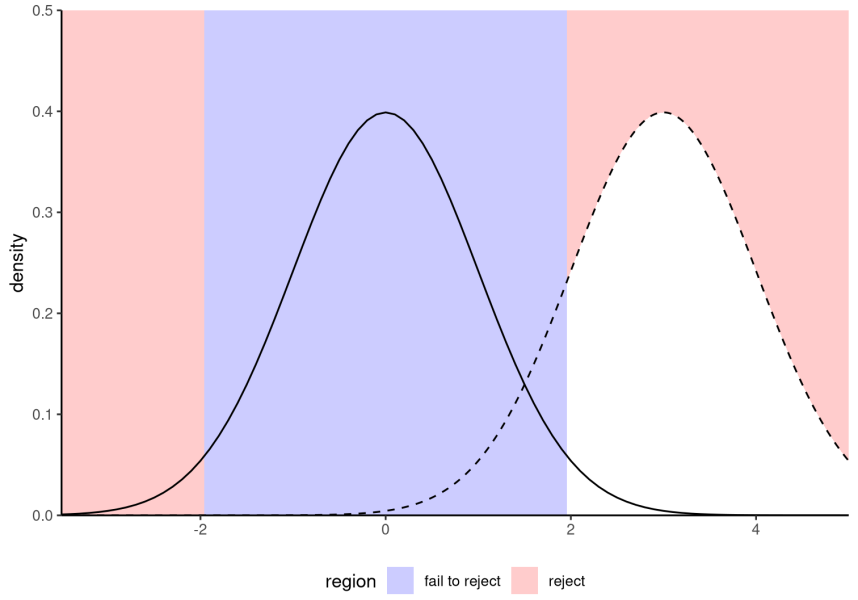
Same for replication studies: how many participants needed?

I cried power!

The null alternative corresponds to a single value (equality in mean), whereas there are infinitely many alternatives...



Power is the ability to detect when the null is false, for a given alternative (dashed).



Power is the area in white under the dashed curved, beyond the cutoff.

What determines power?

Think in your head of potential factors impact power for a factorial design.

1. The size of the effects, $\delta_1 = \mu_1 - \mu$, ..., $\delta_K = \mu_K - \mu$
2. The background noise (intrinsic variability, σ^2)
3. The level of the test, α
4. The sample size in each group, n_j
5. The choice of experimental design
6. The choice of test statistic

We focus on the interplay between

effect size | power | sample size

Living in an alternative world

In a one-way ANOVA, the alternative distribution of the F test has an additional parameter Δ, which depends on both the sample and the effect sizes.

$$\Delta = \frac{\sum_{j=1}^{K} n_j(\mu_j - \mu)^2}{\sigma^2} = nf^2.$$

Under the null hypothesis, $\mu_j = \mu$ for $j = 1, \ldots, K$ and $\Delta = 0$.

The greater Δ, the further the mode (peak of the distribution) is from unity.

Noncentrality parameter and power

$$\Delta = \frac{\sum_{j=1}^{K} n_j (\mu_j - \mu)^2}{\sigma^2}.$$

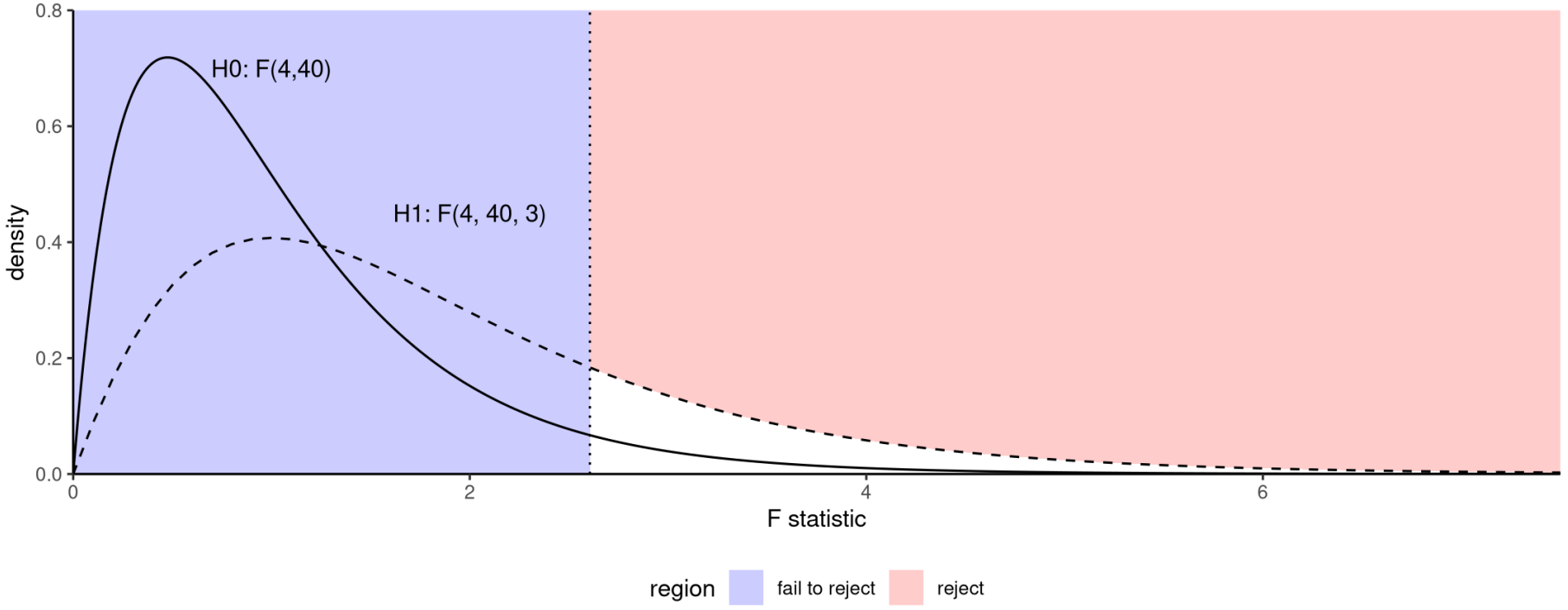
When does power increase?

What is the effect of an increase of the

- group sample size $n_1, \ldots, n_K$.
- variability σ^2.
- true mean difference $\mu_j - \mu$.

Noncentrality parameter

The alternative distribution is $F(\nu_1, \nu_2, \Delta)$ distribution with degrees of freedom ν_1 and ν_2 and noncentrality parameter Δ.



Power for factorial experiments

- $\mathrm{G^{\star}Power}$ and **R** packages take Cohen's f (or f^2) as inputs.
- Calculation based on F distribution with
  - $\nu_1 = \mathrm{df}_{\mathrm{effect}}$ degrees of freedom
  - $\nu_2 = n - n_g$, where n_g is the number of mean parameters estimated.
  - noncentrality parameter $\phi = n f^2_{\langle \mathrm{effect} \rangle}$

Example

Consider a completely randomized design with two crossed factors A and B.

We are interested by the interaction, $\eta^2_{\langle AB \rangle}$, and we want 80% power:

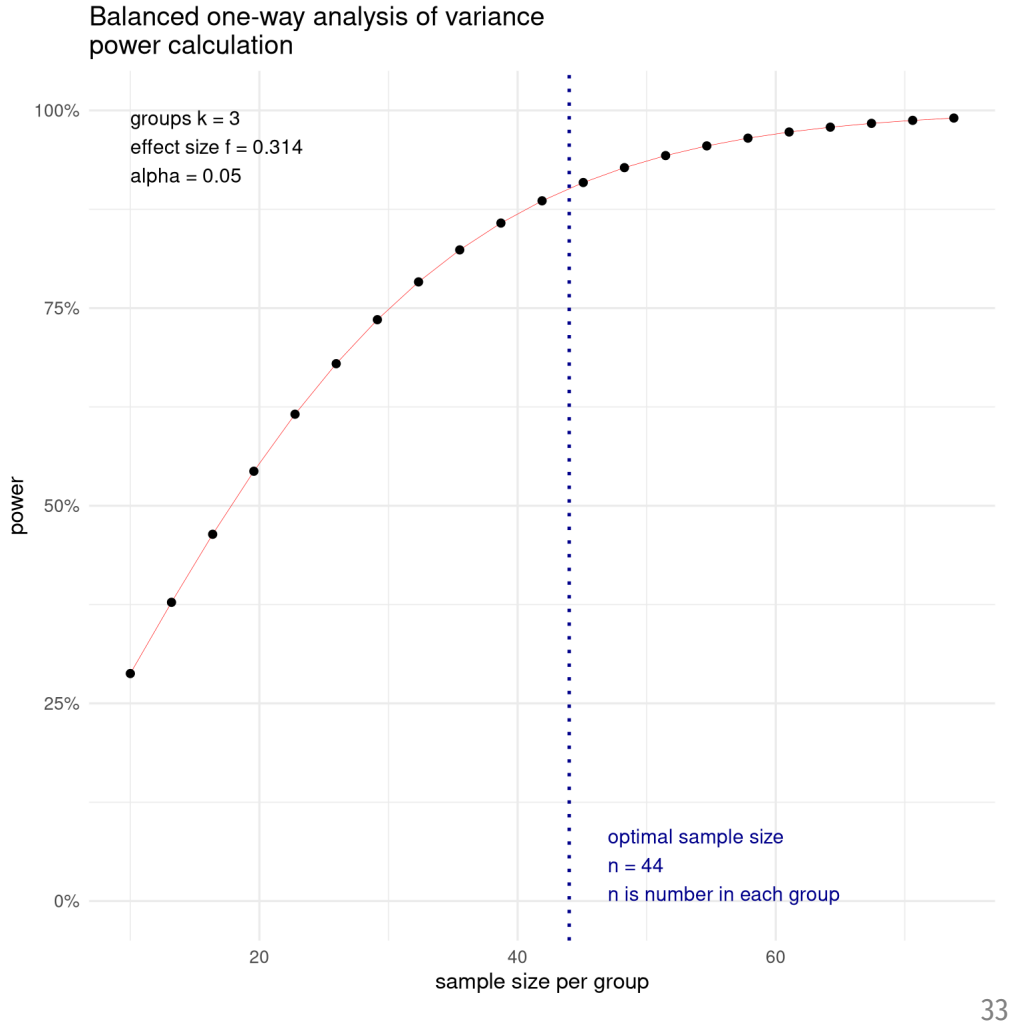
```
# Estimate Cohen's f from omega.sq.partial
fhat <- sqrt(omega.sq.part/(1-omega.sq.part))
# na and nb are number of levels of factors
WebPower::wp.kanova(power = 0.8,
                    f = fhat,
                    ndf = (na-1)*(nb-1),
                    ng = na*nb)
```

Power curves

```
library(pwr)
power_curve <-
 pwr.anova.test(
  f = 0.314, #from R-squared
  k = 3,
  power = 0.9,
  sig.level = 0.05)
plot(power_curve)
```

Recall: convert η^2 to Cohen's f (the effect size reported in `pwr`) via $f^2 = \eta^2/(1-\eta^2)$

Using $\tilde{f}$ instead (from $\widehat{\omega}^2$) yields $n = 59$ observations per group!



Effect size estimates

WARNING!

Most effects reported in the literature are severely inflated.

Publication bias & the file drawer problem

- Estimates reported in meta-analysis, etc. are not reliable.
- Run pilot study, provide educated guesses.
- Estimated effects size are uncertain (report confidence intervals).

Beware of small samples

Better to do a large replication than multiple small studies.

Otherwise, you risk being in this situation:

Observed (post-hoc) power

Sometimes, the estimated values of the effect size, etc. are used as plug-in.

- The (estimated) effect size in studies are noisy!
- The post-hoc power estimate is also noisy and typically overoptimistic.
- Not recommended, but useful pointer if the observed difference seems important (large), but there isn't enough evidence (too low signal-to-noise).

Statistical fallacy

Because we reject a null doesn't mean the alternative is true!