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#### **Classical tests**

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5-9th of August 2019

### Quiz 1

- A parametric test assumes:
  - no distribution of observed data
  - quantiles of observed data
  - standard error of observed data
  - normal distribution of observed data

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- If you want to know whether a female lion runs faster than male lion, and you know that the data collected is normally distributed, you would use:
  - Kruskall Wallis test
  - t-test
  - ANCOVA test
  - chi-squared test

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- If you want to know whether a grouper fish caught in Sigli, Tj Luar, and Kepulauan Seribu are different in body length, and you do not make any assumption of the data distribution, you would use:
  - Kruskall Wallis test
  - t-test
  - ANCOVA test
  - chi-squared test

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  - Kruskall Wallis test
  - t-test
  - ANCOVA test
  - chi-squared test

- If you want to know whether a lovebird is sold due to its different sizes and or due to the area the shop is located, and you assume normal distribution of the observed data you would use:
  - Kruskall Wallis test
  - t-test
  - ANCOVA test
  - chi-squared test

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  - Kruskall Wallis test
  - t-test
  - ANCOVA test
  - chi-squared test

### Quiz 5

- If you want to know whether a survey on whether female or male eat more sharks depends on the education level of female/male, then we could use:
  - Kruskall Wallis test
  - t-test
  - ANCOVA test
  - chi-squared test

- If you want to know whether a survey on whether female or male eat more sharks depends on the education level of female/male, then we could use:
  - Kruskall Wallis test
  - t-test
  - ANCOVA test
  - chi-squared test

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#### Classical tests

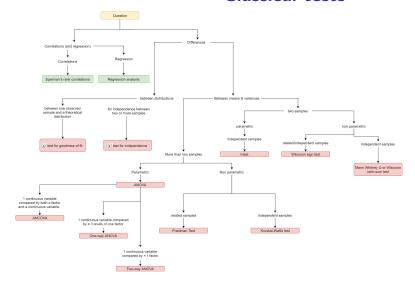


Figure 1: Classical tests diagram

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# Classical tests: Parametric vs non- parametric

- A parametric test assumes normal distribution of observed variables.
- A non-parametric test assumes no distribution of observed variables.

#### Parametric test

#### t-test

- t-test assumes normal distribution.
- One-sample t-test is used to compare a known mean value with a certain distribution.

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### t-test case study

We assume the daily energy intake in kJ for 11 gorillas to be:

```
daily.intake<-c(5260,5470,5640,6180,6390,6515,6805,7515,7515,8230,8770)
summary(daily.intake) #summary statistics
```

```
## Min. 1st Qu. Median Mean 3rd Qu. Max.
## 5260 5910 6515 6754 7515 8770
```

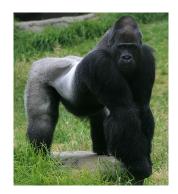


Figure 2: Gorilla

```
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```

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#### t-test case study

Is the intake value deviates from a recommended value of 7725 kJ?

```
t.test(daily.intake,mu=7725)
```

```
## ## One Sample t-test
## ## data: daily.intake
## t = -2.8208, df = 10, p-value = 0.01814
## alternative hypothesis: true mean is not equal to 7725
## 5986.348 7520.925
## sample estimates:
## mean of x
## 6753.636
```

Conclusion: mean daily intake is not equal to the recommended value ( $H_0$  is rejected).

#### Two-sample t-test

- A **two-sample t-test** tests: if two samples mean (  $barx_1$  and  $\bar{x}_2$  ) may come from the same distribution or not (  $\mu_1 = \mu_2$  ).
- We can then calculate:

$$t = \frac{\bar{x}_1 - \bar{x}_2}{SEDM}$$

• Where the standard error of difference of means is:

$$SEDM = \sqrt{SEM_1^2 + SEM_2^2}$$

.

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### Two-sample t-test case study

We want to calculate whether there are differences between the energy expend by obese gorilla vs lean gorilla.

```
library(ISwR)
attach(energy)
head(energy,4)
     expend stature
      9.21
             obese
     7.53 lean
## 3 7.48 lean
## 4 8.08 lean
t.test(expend-stature,var.equal=T) #expend (energy spent) is described by stature (height)
## Two Sample t-test
## data: expend by stature
## t = -3.9456, df = 20, p-value = 0.000799
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -3.411451 -1.051796
## sample estimates:
## mean in group lean mean in group obese
             8.066154
                         10.297778
##
                                  #we assume the variance is the same in two groups
detach(energy)
```

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#### Two-sample t-test case study

- We found that the  $H_0$  is rejected (p-value <0.05) and the alternative hypothesis is accepted.
- Conclusion: energy spent by obese vs lean gorillas are different.

#### Paired t-test

- Paired t-test: two measurements on the same experimental unit.
- **Pre- and post-** measurements/experiments.

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attach(intake)

### Paired t-test case study

 Difference between the energy intake pre and post-release into the wild (after captured by hunters) in a group of gorillas.

```
head(intake,3)
      pre post
## 1 5260 3910
## 2 5470 4220
## 3 5640 3885
t.test(pre,post, paired=T)
## Paired t-test
## data: pre and post
## t = 11.941, df = 10, p-value = 3.059e-07
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## 1074.072 1566.838
## sample estimates:
## mean of the differences
                  1320.455
    #the paired=T is important because it indicates that the sample
    #comes from the same experiment
detach(intake)
```

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#### Paired t-test case study

• There is a significant difference between the energy intake before and after release into the wild (p-value < 0.05).

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# **Analysis of variance (ANOVA)**

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### **One-way ANOVA**

- With more than two groups: use ANOVA (parametric).
- One-way analysis of variance : total variation  $(SSD_T)$  is a sum of variaton within groups  $(SSD_W)$  and variation between groups  $(SSD_B)$ .
- Total variation is split into: describing between group means and between individual measurements within the groups.

### **One-way ANOVA**

- The sum of the two variations can only be positive
- If there is no systematic differences between the groups, the sum of squares should be partitioned by degrees of freedom
  - k-1 for  $SSD_B$  ( k = number of groups)
  - n k for  $SSD_W$  ( N = total number of observations).

### **One-way ANOVA**

Mean squares or pooled variance:

$$MS_W = SSD_W/(N-k)$$
  
 $MS_B = SSD_B/(k-1)$ 

- With no true group effect,  $MS_B$  will be an estimate of  $\sigma^2$
- If there is a group effect,  $MS_B$  will tend to be large.

#### **One-way ANOVA**

- So, a test for differences between a two group **means** can be done by comparing the **variance** estimates.
- The approach uses *F* distribution:

$$F = MS_B/MS_W$$

• *F* is ideally 1, but some variation around the value is expected.

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# One-way ANOVA case study

- Simple analyses of variance can be performed using function 1m (similar to regression analysis).
- Example data of "red cell folate" (Altman 1991, p. 208).
   But assume the data is about the number of Pangolin scales confiscated in year 2013, 2014, and 2015.



Figure 3: Pangolin

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### One-way ANOVA case study

- One-way analysis of variance is analogous to regression analysis
- But the explanatory variables are factors not numeric variable.

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# One-way ANOVA case study

head(pangolin1,3)

```
## scales year
## 1 243 year2013
## 2 251 year2013
## 3 275 year2013
anova(lm(pangolin1$scales-pangolin1$year))
```

- We get  $SSD_B$  and  $MS_B$  (top line)
- $SSD_W$  and  $MS_W$  (second line).
- If the F test shows difference between groups (p=value <0.05), we need to know where the difference lies.</li>

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# One-way ANOVA case study

We should then compare the individual groups:

summary(lm(pangolin1\$scales~pangolin1\$year))

```
##
## Call:
## lm(formula = pangolin1$scales ~ pangolin1$year)
## Residuals:
             10 Median 30
      Min
                                    Max
## -73.625 -35.361 -4.444 35.625 75.375
##
## Coefficients:
##
                      Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                        316.62
                                    16 16 19 588 4 65e-14 ***
## pangolin1$yearyear2014 -60.18 22.22 -2.709 0.0139 *
## pangolin1$yearyear2015 -38.62 26.06 -1.482 0.1548
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 45.72 on 19 degrees of freedom
## Multiple R-squared: 0.2809, Adjusted R-squared: 0.2052
## F-statistic: 3.711 on 2 and 19 DF, p-value: 0.04359
```

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### One-way ANOVA case study

- We found: the intercept is the mean in the first group (year2013) (baseline)
- Two others describe the difference between year2014 and year2015 to year2013.
- We haven't known the comparison between the second and the third group?

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### One-way ANOVA case study

- Use multiple testing.
- Conducting many tests can increase the probability of finding one of them to be significant
- Use an adjustment method, the Bonferroni correction

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### One-way ANOVA case study

pairwise.t.test(pangolin1\$scales, pangolin1\$year, p.adj="bonferroni")

```
## Pairwise comparisons using t tests with pooled SD
## data: pangolin1$scales and pangolin1$year
## year2013 year2014
## year2014 0.042 -
## year2015 0.464 1.000
##
## P value adjustment method: bonferroni
```

- We found there is difference between year2013 and year2014
- No difference between year2015 with year2013 nor year2014.

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### One-way ANOVA case study

 We can relax the assumption of equal variances for all groups by using the oneway.test

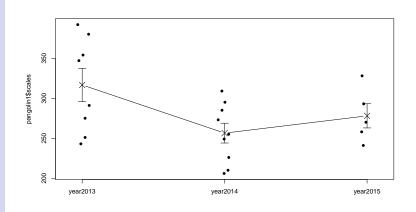
### One-way ANOVA case study

```
oneway.test(pangolin1$scales~pangolin1$year)
```

```
## ## One-way analysis of means (not assuming equal variances)
## ## data: pangolin1$scales and pangolin1$year
## F = 2.9704, num df = 2.000, denom df = 11.065, p-value = 0.09277
```

- We now see that the p-value has increased to a non-significant values
- Maybe, the group that seems to differ from the two others also has the largest variance.

# One-way ANOVA case study



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# Plotting: which error bar to use?

- Should we draw using 1 SEM (done here) or using CIs ( $\sim$  2 SEM?)
- Or maybe SD instead of SEM?
- SEM = not useful for describing the distributions in the groups; only say how precise the mean is.
- SDs is difficult to see significant difference.
- Traditionally, people use 1 SEM (more dramatic)

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# Bartless test for checking variance

 Bartlett's test: to test whether the distribution of a variable has the same variance in all groups.

bartlett.test(pangolin1\$scales~pangolin1\$year)

```
##
## Bartlett test of homogeneity of variances
##
## data: pangolini$scales by pangolini$year
## Bartlett's K-squared = 2.0951, df = 2, p-value = 0.3508
```

 No violation of equal variances assumptions in the three groups.

## Two-way ANOVA

- When we have more than two factors to test on a dependent variable: two-way ANOVA.
- Single observation per cell of data: **residual variation**  $SSD_res = \text{substraction of variation between rows}$   $(SSD_R)$  and variation between columns  $(SSD_C)$ .

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## Two-way ANOVA case study

- A fake dataset: mean daily catch of Silky shark (*Carcharhinus falciformis*) in January 2017.
- We want to see the difference of catch by area (Meulaboh, Sigli and Biereun) (factor 1) depending on the whether the shark is female/male (factor 2. \* The real dataset comes from here.



Figure 4: Silky shark

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# Two-way ANOVA case study

```
## [1] "sex" "area" "catch.per.day"
```

We can then conduct the two-way ANOVA as:

```
calci.aov<-aov(catch.per.day-sex+area,data=calci)
summary(calci.aov)</pre>
```

```
## Sex Df Sum Sq Mean Sq F value Pr(>F)  
## sex 1 0.067 0.0666 0.087 0.770  
## area 2 0.278 0.1392 0.181 0.835  
## Residuals 32 24.575 0.7680
```

 We found: both species sex and area are not statistically significant to the catch.

## Two-way ANOVA case study

To see if there is an interaction effect:

calci2.aov<-aov(catch.per.day~sex+area+sex:area,data=calci)
summary(calci2.aov)</pre>

• We do not find any significant interaction.

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## Two-way ANOVA case study

We can compute some summary statistics (mean, sd, se):

```
model.tables(calci.aov, type="means", se=TRUE)
```

```
## Tables of means
## Grand mean
##
## 3.170833
##
## sex
## female male
## 3.201 3.11
## rep 24.000 12.00
##
## area
## bireuen meulaboh sigli
## 3.291 3.083 3.139
## rep 12.000 12.000 12.000
```

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# Two-way ANOVA case study

```
model.tables(calci2.aov, type="means", se=TRUE)
```

```
## Tables of means
## Grand mean
##
## 3.170833
## sex
## female male
    3.201 3.11
## rep 24.000 12.00
##
## area
     bireuen meulaboh sigli
       3.291 3.083 3.139
## rep 12.000 12.000 12.000
##
## sex:area
##
         area
## sex
        bireuen meulaboh sigli
## female 3.235 3.050 3.319
## rep 8.000 8.000 8.000
## male 3.403 3.148 2.780
## rep 4.000 4.000 4.000
```

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## Two-way ANOVA case study

- We compute Tukey HSD (Honest Significant Differences) to perform multiple pairwise-comparison between the means of groups.
- We have seen that no difference between the groups, but let's give it a try:

TukevHSD(calci.aov, which = "area")

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# Two-way ANOVA case study

```
TukeyHSD(calci2.aov, which = "area")
```

```
## Tukey multiple comparisons of means
## 95% family-wise confidence level
##
## Fit: aov(formula = catch.per.day ~ sex + area + sex:area, data = calci)
##
## $area
## upr p adj
## meulaboh-bireuen -0.20833333 -1.1041435 0.6874768 0.8352605
## sigli-bireuen -0.151666667 -1.0474768 0.7441435 0.9087323
## sigli-meulaboh 0.056666667 -0.8391435 0.9524768 0.9666881
```

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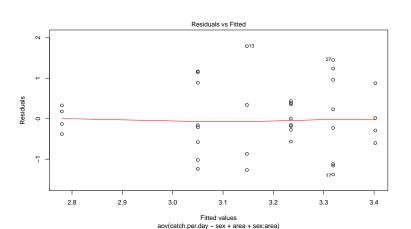
# Checking assumptions of ANOVA

- Assumptions to check before running ANOVA (or any test assuming normal distribution):
- ① Checking the homogeneity of variance assumptions

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# Checking assumptions of

plot(calci2.aov, 1)



Points 13, 27 and 17 are detected as outliers, which can severely affect normality and homogeneity of variance. It can be useful to remove outliers to meet the test assumptions

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# Checking assumptions of ANOVA

- Use Levene's test to check the homogeneity of variances.
- leveneTest() in car package will be used:

```
library(car)
leveneTest(catch.per.day-sex*area,data=calci)
```

```
## Levene's Test for Homogeneity of Variance (center = median)
## Df F value Pr(>F)
## group 5 3.4628 0.01374 *
## 30
## ---
## 5ignif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

- p-value < 0.05 = evidence to suggest group variances are different.
- We cannot assume homogeneity of variances.

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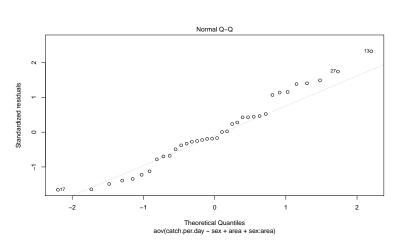
# Checking assumptions of ANOVA

- 2 Check the normality assumption
- Use normal probability plot of residuals
- Should approximately follow a straight line.

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# Checking assumptions of

plot(calci2.aov, 2)



 All points fall along the reference line: can assume normality

```
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```

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# Checking assumptions of ANOVA

• Run a Shapiro-Wilk test on the ANOVA residuals:

```
# Extract the residuals
aov_residuals <- residuals(object = calci2.aov)
# Run Shapiro-Wilk test
shapiro.test(x = aov_residuals )</pre>
```

```
##
## Shapiro-Wilk normality test
##
## data: aov_residuals
## W = 0.97011, p-value = 0.4284
```

p-value is >0.05: do not find violation of normality.

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# **Unbalanced design**

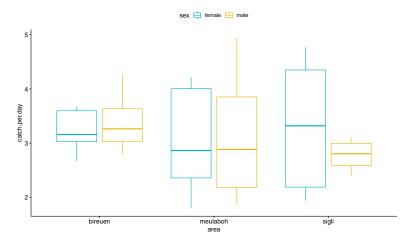
- Our example above: deals with unbalanced design/not having equal numbers of subjects in each group.
- Three different ways to run an ANOVA: Type-I, Type-II and Type-III (recommended) sums of squares.

 We find similar results, but already acknowledge the unbalanced design.

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## Plotting result

Overview of the data set, separated into the different sex and area by:



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# **ANCOVA** (analysis of covariance)

- Used to test main and interaction effects of categorical variables on a continuous dependent variable
- Controlling for other continuous variables: co-vary with the dependent.
- The control variables: "covariates."
- No covariates: use ANOVA. Yes covariates: use ANCOVA.

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## **ANCOVA** case study

Fake data: how long does a criminal got prison sentence?
 Variables: weight (kg skin) of confiscated evidence (by species: tiger or clouded leopard).



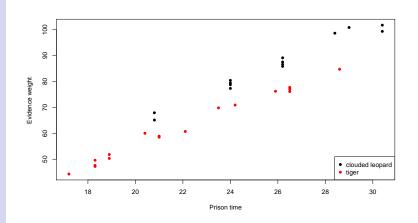


Clouded leopard & Indian Tiger.

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# **ANCOVA** case study

```
## 1 clouded leopard 20.8 67.9
## 2 clouded leopard 20.8 65.1
## 3 clouded leopard 24.0 77.3
```



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## species

## Residuals ## ---

## **ANCOVA** case study

```
model.1 = lm (prison.time ~ evidence.weight + species + evidence.weight:species,
             data = confs)
Anova(model.1, type="II")
## Anova Table (Type II tests)
##
## Response: prison.time
                         Sum Sq Df F value Pr(>F)
##
## evidence.weight 4376.1 1 1388.839 < 2.2e-16 ***
```

598.0 1 189.789 9.907e-14 \*\*\*

## evidence.weight:species 4.3 1 1.357 0.2542 85.1 27

## Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.05 '.' 0.1 ' ' 1

 Interaction = not significant; slope across groups = not different.

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# **ANCOVA** case study

 Re-run the model: is there difference among different species (without interaction):

```
## Anova Table (Type II tests)
##
## Response: prison.time
## Sum Sq Df F value Pr(>F)
## evidence.weight 4376.1 1 1371.4 < 2.2e-16 ***
## species 598.0 1 187.4 6.272e-14 ***
## Residuals 89.3 28
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

 Species is significant; intercepts among groups are different.

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# **ANCOVA** case study

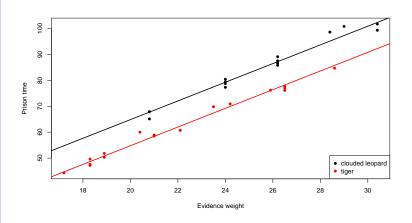
summary(model.2)

```
##
## Call:
## lm(formula = prison.time ~ evidence.weight + species, data = confs)
##
## Residuals:
      Min 10 Median 30
                                    Max
## -3.0128 -1.1296 -0.3912 0.9650 3.7800
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept) -7.21091
                             2.55094 -2.827 0.00858 **
## evidence.weight 3.60275 0.09729 37.032 < 2e-16 ***
## speciestiger -10.06529 0.73526 -13.689 6.27e-14 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.786 on 28 degrees of freedom
## Multiple R-squared: 0.9896, Adjusted R-squared: 0.9888
## F-statistic: 1331 on 2 and 28 DF, p-value: < 2.2e-16
```

- Slope estimate is the same.
- Intercept for species 1 (Clouded leopard) is (Intercept).
- Intercept for species 2 (Tiger) is (intercept) + Speciestiger.
- Determined from the fact that Speciestiger is shown in coefficient table above.

# **ANCOVA** case study

Add the fitted lines:



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# Some notes on Type I, II, and III sums of squares:

- Type I = anova function.
- ullet Type II and Type III = Anova in the car package
- Type I = "sequential." The factors tested in the order they are listed in the model.
- Type III = "partial." Every term in the model is tested after every other term in the model (interaction and other main effects)
- Type II are similar to Type III, but uses principle of marginality: test for main effects but not interaction.
- When data are balanced and the design is simple, types I,
   II, and III = same results.

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## **Conclusions on ANOVA**

- One way ANOVA = one continuous variable (e.g., test score) compared by  $\geq 3$  levels of one factor variable (e.g., level of education)
- Two-way ANOVA = one continous variable (e.g., test score) compared by \*>1 \* factor variable (e.g., level of education and zodiac sign)
- ANCOVA = one continuous variable, but it compares a response variable by both a factor and a continuous independent variable (e.g., comparing test score by 'level of education' and 'number of hours spent studying'). The continuous independent variable is 'covariate'.

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## One-way ANOVA

#### One-way ANOVA Example

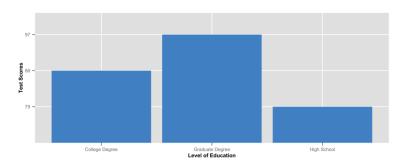


Figure 5: One-way ANOVA

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# Two way ANOVA



Figure 6: Two-way ANOVA

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## **ANCOVA**

#### **ANCOVA EXAMPLE**

# Independent Variables (Factor) Level of Education (High School, College Degree, or Graduate Degree) (Covariate) Number of Hours Spent Studying

Figure 7: ANCOVA

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# Non-parametric test

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# Wilcoxon test : One-sample Wilcoxon

- Non-parametric version of t-test.
- Data are replaced with their ranks without regard to the grouping.
- One-sample Wilcoxon signed-rank test = done similarly to one-sample t-test.

```
#daily energy intake in kJ for 11 gorillas to be:
t.test(daily.intake,mu=7725) #parametric
```

```
##
## One Sample t-test
##
## data: daily.intake
## t = -2.8208, df = 10, p-value = 0.01814
## alternative hypothesis: true mean is not equal to 7725
## 5986.348 7520.925
## sample estimates:
## mean of x
## 6753.636
```

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# Wilcoxon test : One-sample Wilcoxon

```
wilcox.test(daily.intake,mu=7725) #non-parametric
```

```
##
## Wilcoxon signed rank test with continuity correction
##
## data: daily.intake
## V = 8, p-value = 0.0293
## alternative hypothesis: true location is not equal to 7725
```

 Both tests show significant difference between the daily intake vs recommended intake (7725 kJ).

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# Two-sample Wilcoxon signed-rank test

### Similar to two-sample t-test

```
attach(energy)

#Differentiate the energy expend difference for obese gorilla vs lean gorilla.

t.test(expend-stature)#parametric
```

```
##
## Welch Two Sample t-test
##
data: expend by stature
## t = -3.8555, df = 15.919, p-value = 0.001411
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -3.459167 -1.004081
## sample estimates:
## mean in group lean mean in group obese
## 8.066154 10.297778
```

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# Two-sample Wilcoxon signed-rank test

wilcox.test(expend~stature)#non parametric

```
##
## Wilcoxon rank sum test with continuity correction
##
## data: expend by stature
## W = 12, p-value = 0.002122
## alternative hypothesis: true location shift is not equal to 0
```

 Both tests show significant difference between obese vs lean gorillas energy expend.

```
Classical tests
```

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## Paired Wilcoxon test

Similar to the paired t.test:

```
#Difference between the energy intake pre and post-release
#into the wild (after captured by hunters) in a group of gorillas.
attach(intake)
t.test(pre,post, paired=T) #parametric
```

```
## ## Paired t-test
## data: pre and post
## t = 11.941, df = 10, p-value = 3.059e-07
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## 1074.072 1566.838
## sample estimates:
## mean of the differences
## 1320.455
```

 Both showed significant difference between the energy intake of pre and post release of gorillas into the wild.

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### Paired Wilcoxon test

```
##
## Wilcoxon signed rank test with continuity correction
##
## data: pre and post
## V = 66, p-value = 0.00384
## alternative hypothesis: true location shift is not equal to 0

detach(intake)
```

wilcox.test(pre,post, paired=T)#non parametric

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### Kruskal Wallis test

- 'One way ANOVA' for non-parametric data.
- Compares differences between > 2 means.
- Similar to Wilcoxon-two sample test: data are replaced with their ranks, no regard to grouping.

oneway.test(pangolin1\$scales~pangolin1\$year) #parametric

```
##
## One-way analysis of means (not assuming equal variances)
##
## data: pangolin1$scales and pangolin1$year
## F = 2.9704, num df = 2.000, denom df = 11.065, p-value = 0.09277
```

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### Kruskal Wallis test

kruskal.test(pangolin1\$scales~pangolin1\$year) #non-parametric

```
##
## Kruskal-Wallis rank sum test
##
## data: pangolin1$scales by pangolin1$year
## Kruskal-Wallis chi-squared = 4.1852, df = 2, p-value = 0.1234
```

 For both, we found no difference of pangolin scales confiscated evidence along the years.

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- Used to analyze the frequency table (i.e. contingency table)
- Two categorical variables.
- Test significant association between two categorical variables.

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# Chi-square test of independence

### A boring contingency table

 Here is an exceptionally boring contingency table, showing how a deck of cards is divided into suits and colors:

	Diamond	Heart	Club	Spade	Total
Red	13	13	0	0	26
Black	0	0	13	13	26
Total	13	13	13	13	52

• Note that the row totals and column totals both sum to the grand total of 52.

Figure 8: Contingency table example

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## Chi-square test of independence

 Are these songbirds associated with certain provinces? (ownership)









Cucak rowo, murai batu, lovebird and cucak hijau respectively.

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```
songbird<-read.csv("fake_songbird.csv",row.names = 1)</pre>
```

- Can be visualized using balloonplot() (gplots package).
- Draws graphical matrix, each cell contains a dot whose size reflects the relative magnitude of the corresponding component.

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# Chi-square test of independence

### **Provinces**



cucak.rowo	murai.batu	love.bird	cucak.hijau
	•	•	•
	•	•	•
	•	•	•
	•	•	•
	•		
•	•	•	
•	•	•	
•	•	•	•
•	•		•
•	•	•	
•			
	•		•
		•	

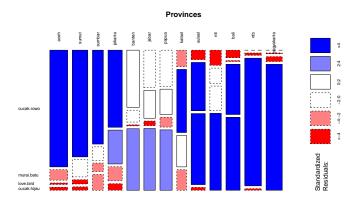
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- We found: cucak rowo corresponding strongly to Aceh, Sumut, Sumbar and Jakarta.
- Need further tests to see whether this association is significant or not.

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- Visualize using mosaic plot.
- Use mosaicplot()

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## Chi-square test of independence

### Conclusions:

- Blue color = observed value > expected value; if data were random.
- Red color = specifies that observed value < expected value; if data were random
- Cucak Rowo is mainly kept as a pet in Aceh, Sumut, Sumbar and Jakarta.

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- Examines whether rows and columns of a contingency table are associated.
  - H<sub>0</sub>: row and the column variables of the contingency table are independent.
  - $H_1$ : row and column variables are dependent

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## Chi-square test of independence

• For each cell, expected value is calculated:

$$e = \frac{row.sum \times col.sum}{grand.total}$$

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# Chi-square test of independence

• Chi-square statistic is calculated as:

$$\chi^2 = \sum \frac{(o-e)^2}{e}$$

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- Chi-square statistic is compared to the critical value (obtained from statistical tables)
- df = (r-1)(C-1) degrees of freedom and p = 0.05.
- Calculated Chi-square statistic > critical value = row and the column variables are not independent of each other (associated)

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# Chi-square test of independence case study

```
chisq<-chisq.test(songbird)
chisq</pre>
```

```
## Pearson's Chi-squared test
## ## data: songbird
## X-squared = 1944.5, df = 36, p-value < 2.2e-16
```

 Row and the column variables are statistically significantly associated (p-value = 0).

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# Chi-square test of independence case study

Observed and the expected counts can be extracted:

chisq\$observed #observed counts

##		cucak rowo	murai hatu	love bird	cucak.hijau
	,				
##	aceh	156	14	2	4
##	sumut	124	20	5	4
##	sumbar	77	11	7	13
##	jakarta	82	36	15	7
##	banten	53	11	1	57
##	jabar	32	24	4	53
##	papua	33	23	9	55
##	kalsel	12	46	23	15
##	sulsel	10	51	75	3
##	ntt	13	13	21	66
##	bali	8	1	53	77
##	ntb	0	3	160	2
##	yogyakarta	0	1	6	153

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# Chi-square test of independence case study

chisq\$expected #expected counts

```
cucak.rowo murai.batu love.bird cucak.hijau
## aceh
             60.55046
                       25.63303 38.44954
                                          51.36697
## sumut
            52.63761 22.28326 33.42489 44.65424
## sumbar
           37.15596 15.72936 23.59404 31.52064
## jakarta
          48.16514 20.38991 30.58486 40.86009
## banten
          41.97248 17.76835 26.65252
                                         35.60665
## jabar 38.87615 16.45757 24.68635
                                         32 97993
## papua 41.28440 17.47706 26.21560
                                          35.02294
## kalsel
            33.02752 13.98165 20.97248
                                          28.01835
## sulsel
            47.82110
                       20.24427 30.36640
                                          40.56823
## ntt
            38.87615 16.45757 24.68635
                                          32.97993
## bali
            47.82110 20.24427 30.36640
                                         40.56823
## nth
            56.76606 24.03096 36.04644
                                          48 15654
## vogvakarta
            55.04587
                       23.30275 34.95413
                                          46 69725
```

# Chi-square test of independence case study

 To know the most contributing cells to the total Chi-square score:

$$r = \frac{o - e}{\sqrt{e}}$$

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# Chi-square test of independence case study

- Previous formula returns Pearson residuals (r) for each cell (or standardized residuals).
- Cells with highest absolute standardized residuals contribute most to total Chi-square score.
- Pearson residuals can be easily extracted from the output of the function chisq.test():

### chisq\$residuals

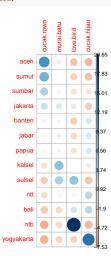
```
cucak.rowo murai.batu love.bird cucak.hijau
## aceh
            12 266344 -2 2976972 -5 8782288
                                            -6.608968
                                           -6.083794
## sumut
            9.836062 -0.4836881 -4.9165873
## sumbar
             6 536557 -1 1924678 -3 4162600
                                           -3.298819
          4.875263 3.4569876 -2.8180566
## iakarta
                                           -5.297104
## banten
                                           3.585198
            1.702142 -1.6056807 -4.9689076
          -1.102817 1.8592103 -4.1634703 3.486107
## jabar
          -1.289342 1.3211002 -3.3623444 3.375634
## papua
## kalsel -3.658896 8.5628768 0.4427320
                                           -2.459431
## sulsel -5.469207 6.8355754 8.0996318
                                           -5.898312
         -4.150095 -0.8522912 -0.7419395
                                           5.749805
## ntt
## bali -5.758422 -4.2771092 4.1073055 5.719883
## ntb
       -7.534325 -4.2901609 20.6456123
                                            -6.651286
## vogvakarta -7.419291 -4.6201370 -4.8973513
                                            15 556033
```

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# Chi-square test of independence case study

Visualize Pearson residuals

library(corrplot)
corrplot(chisg\$residuals, is.cor=FALSE)



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# Chi-square test of independence case study

- Sign of the standardized residuals = very important to interpret association
  - Positive residuals (blue): attraction (positive association)
  - Results: association between Cucak Rowo and Aceh, Sumut, Sumbar, Jakarta
  - Strong positive association between murai baru and Kalses
  - Negative residuals (red) = repulsion (negative association).
  - Cucak Rowo are negatively associated (~ "not associated") with the row Yogyakarta.
  - Repulsion between Love bird and with Aceh and Sumut.

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### Chi square goodness of fit test in R

- Used to compare the observed distribution to an expected distribution
- Compares multiple observed proportions to expected probabilities.

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### Chi square goodness of fit test in R case study

• Stingray species (*Dasyatis americana*) in Tj. Luar were found in January 2017: 81 juveniles, 50 adults and 27 old.

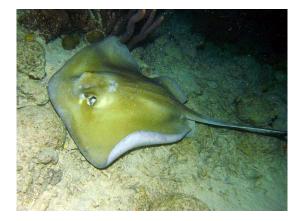


Figure 9: Stingray

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### Chi square goodness of fit test in R case study

- Question 1:
  - Are these age group equally common?
  - If these age groups were equally distributed, the expected proportion would be 1/3 for each age group.

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## Chi square goodness of fit test in R case study

- Question 2:
- Suppose that, in Indonesia ratio of juvenile, adults and old is 3:2:1 (3+2+1=6).
- Expected proportion is:
  - 3/6 (= 1/2) for juvenile
  - 2/6 ( = 1/3) for adults
  - 1/6 for old
- Is there significant difference between observed and expected proportions?

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### Chi square goodness of fit test in R case study

- H<sub>0</sub>: No significant difference between observed and expected value.
- H<sub>a</sub>: There is significant difference between observed and expected value.

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# Chi square goodness of fit test in R case study

• Q1: are the age groups equally common?

```
age<-c(81,50,27)
res<-chisq.test(age, p = c(1/3, 1/3,1/3)) #expected
```

- p-value is  $8.80310^{-7}$  (p-value < 0.05).
- Age groups are significantly not commonly distributed.

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# Chi square goodness of fit test in R case study

Access to the expected values:

res\$expected

## [1] 52.66667 52.66667 52.66667

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# Chi square goodness of fit test in R case study

Q2: comparing observed to expected proportions

```
age <- c(81, 50, 27)
res <- chisq.test(age, p = c(1/2, 1/3, 1/6))
res

##
## Chi-squared test for given probabilities
##
## data: age
## X-squared = 0.20253, df = 2, p-value = 0.9037</pre>
```

- p-value = 0.9037
- Observed proportions are not significantly different from the expected proportions.

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### Correlation does not mean causation!

- Inability to deduce a cause-and-effect relationship between two variables solely on the basis of an **observed** association or correlation between them.
- Logical fallacy: how two events occurring together are considered to be a cause-and-effect relationship.

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### Correlation does not mean causation!

Children that watch a lot of TV are the most violent. Clearly, TV makes children more violent.

This could easily be the other way round; that is, violent children like watching more TV than less violent ones.

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### Correlation does not mean causation!

- Europeans in the Middle Ages believed that lice were beneficial to health -> rarely any lice on sick people -> people got sick because the lice left.
- Lice are extremely sensitive to body temperature
- Impression that lice left before the person got sick.

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### Correlation does not mean causation!

As ice cream sales increase, the rate of drowning deaths increases sharply. Therefore, ice cream consumption causes drowning.

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### Correlation does not mean causation!

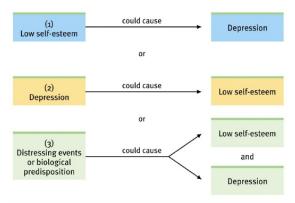
Since the 1950s, both the atmospheric CO2 level and obesity levels have increased sharply. Hence, atmospheric CO2 causes obesity. Richer populations tend to eat more food and produce more CO2.

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# Correlation does not mean causation!

### Correlation and Causation

Correlation does not mean causation!



#### Practice!

- What test to use on this dataset?
  - "fake\_deforestation.csv"
  - % soil organic carbon content (numeric) & deforestation rate (3 levels)
  - check the assumption, then decide which test!

Original data from see reference list.



#### Practice!

- What test to use on this dataset?
  - "fake.pangolin.csv"
  - Years of experience of the criminal, origin (province) of the criminal, weight of pangolin scale confiscated
  - check the assumption, then decide which test!
  - This is originally dataset ToothGrowth in R.

#### Practice!

- What test to use on this dataset?
  - "fake.manta.catch.csv"
  - Manta ray's total body length (TL in cm) is measured a year before MPA, and a year after MPA. Is there any difference?
  - check the assumption, then decide which test!
  - Original dataset, check reference list.

# Quiz 1

- A parametric test assumes:
  - no distribution of observed data
  - quantiles of observed data
  - standard error of observed data
  - normal distribution of observed data

- A parametric test assumes:
  - no distribution of observed data
  - quantiles of observed data
  - standard error of observed data
  - normal distribution of observed data

- If you want to know whether a female lion runs faster than male lion, and you know that the data collected is normally distributed, you would use:
  - Kruskall Wallis test
  - t-test
  - ANCOVA test
  - chi-squared test

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- If you want to know whether a grouper fish caught in Sigli, Tj Luar, and Kepulauan Seribu are different in body length, and you do not make any assumption of the data distribution, you would use:
  - Kruskall Wallis test
  - t-test
  - ANCOVA test
  - chi-squared test

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- If you want to know whether a lovebird is sold due to its different sizes and or due to the area the shop is located, and you assume normal distribution of the observed data you would use:
  - Kruskall Wallis test
  - t-test
  - ANCOVA test
  - chi-squared test

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- If you want to know whether a survey on whether female or male eat more sharks depends on the education level of female/male, then we could use:
  - Kruskall Wallis test
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kruskal two way anova t-test