Homework 3, DATA 556: Due Tuesday, 10/16/2018

Alexander Van Roijen

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Please complete the following:

1. Problem 1

(a) Find the mean and the variance of a Discrete Uniform random variable on $1, 2, \ldots$, n.

Let $X \sim \text{DUnif}$

$$P(X = x) = \frac{1}{n}$$
 by the def of uniform distrib (1)

$$E[X] = \sum_{j=0}^{n} P(X=j) * j = \sum_{j=0}^{n} \frac{1}{n} * j = \frac{1}{n} * \sum_{j=0}^{n} j = \frac{1}{n} * \frac{n(n+1)}{2} = \frac{n+1}{2}$$
 (2)

It is a similar proof with a different substitution for the expected value of X squared

(3)

$$E[X^2] = \sum_{j=0}^{n} P(X=j) * (j^2) \text{ LOTUS}$$
 (4)

$$= \sum_{j=0}^{n} \frac{1}{n} * (j^2) = \frac{1}{n} * \sum_{j=0}^{n} (j^2) = \frac{1}{n} * \frac{n(n+1)(2n+1)}{6} = \frac{(n+1)(2n+1)}{6}$$
 (5)

$$E[X^{2}] - (E[X])^{2} = \frac{(n+1)(2n+1)}{6} - (\frac{n+1}{2})^{2}$$
 (6)

$$\frac{4*(n+1)(2n+1)}{24} - \left(\frac{6*(n+1)^2}{24}\right) = \frac{n^2 - 1}{12}$$
 (7)

(b) Let X be a discrete random variable that satisfies the properties listed in the homework. Find E[X]

We have a symmetry such that
$$P(X = x) = P(X = -x)$$
 (8)

$$E[X] = \sum_{x = -n}^{n} P(X = x) * x \tag{9}$$

$$= P(X=-n)*-n + P(X=-(n-1))*(-(n-1)) + \ldots + P(X=n)*n \qquad (10)$$

$$= -1 * (P(X = -n) * n - P(X = n) * n) + ...P(X = 0) * 0$$
 (11)

$$= -1 * (P(X = n) * n - P(X = n) * n) + ...P(X = 0) * 0 = 0 + 0 + ... + 0$$
 (12)

$$=>E[X]=0 \tag{14}$$

2. We have X with PMF $P(X = k) = \frac{-1*p^k}{\log(1-p)*k}$ for k = 1, 2, . . . Here p is a parameter with 0 ; p ; 1. Find the mean and the variance of X

First, note that the value $\frac{-1}{\log(1-p)}$ is a positive constant

$$E[X] = \sum_{k=1}^{\infty} P(X=k) * k = \sum_{k=1}^{\infty} \frac{-1 * p^k}{\log(1-p) * k} * k = \sum_{k=1}^{\infty} \frac{-1 * p^k}{\log(1-p)}$$
 (15)

$$= \frac{-1}{\log(1-p)} \sum_{r=1}^{\infty} p^k$$
 (16)

$$= \sum_{x=1}^{\infty} p^k = S = S * p = \sum_{x=2}^{\infty} p^k = S - Sp = p = S = \frac{p}{(1-p)}$$
 (17)

 $=> E[X] = \frac{-1}{\log(1-p)} * \frac{p}{(1-p)}$ similarly, the $E[X^2]$ is proved in a similar manner except

(18)

$$S - Sp = \sum_{x=1}^{\infty} p^k \tag{19}$$

which we already proved so by substituting we get (20)

$$S - Sp = \frac{p}{(1-p)} \Longrightarrow S = \frac{p}{(1-p)^2} \tag{21}$$

$$=> E[X^2] = \frac{-1}{\log(1-p)} * \frac{p}{(1-p)^2}$$
 (22)

$$= Var(X) = E[X^{2}] - E[X]^{2} = \frac{p(p - \log(1 - p))}{\log(1 - p)^{2} * (1 - p)^{2}}$$
(23)

3. (a) Use LOTUS to show that for X $\sim \text{Pois}(\lambda)$ and any function g(.), $E[Xg(X)] = \lambda E[g(X+1)]$

$$E[Xg(X)] = \sum_{x=0}^{\infty} x * g(x) \frac{e^{-\lambda} * \lambda^x}{x!} = \sum_{x=1}^{\infty} g(x) * \lambda \frac{e^{-\lambda} * \lambda^{x-1}}{(x-1)!}$$
(24)

now let there be a substitution such that j = x-1 (25)

$$= \sum_{j=0}^{\infty} g(j+1) * \lambda \frac{e^{-\lambda} * \lambda^{j}}{j!} = \lambda * \sum_{j=0}^{\infty} g(j+1) \frac{e^{-\lambda} * \lambda^{j}}{j!}$$
 (26)

considering g(x) is any valid function, we conclude (27)

$$E[Xg(X)] = \lambda * E[g(X+1)]$$
(28)

We bring this back to this form using LOTUS and our lack of knowledge on g(x)

(29)

(b) Find the third moment $E(X^3)$ for X Pois(λ)

let
$$g(x) = x^2 = E(Xg(X)) = E[g(X+1)] = \lambda * E[(X+1)^2]$$
 (30)

$$= \lambda * E[X^2 + 2X + 1] = \lambda * (E[X^2] + 2E[X] + 1)$$
(31)

using the known properties of the variance and mean of the uniform dist we get

(32)

$$= \lambda * (\lambda^2 + \lambda + 2\lambda + 1) = \lambda^3 + 3\lambda^2 + \lambda \tag{33}$$

4. Show that for any events $A_1...A_n$, $P(A_1 \cap ... \cap A_n) \ge \sum_{j=1}^n P(A_j) - n + 1$ lets prove this through induction

Base case,
$$n=1:P(A) \ge P(A) - 1 + 1 => P(A) \ge P(A)$$
 (34)

Now we assume this is true for
$$1 \dots n$$
, NTS for $n+1$ (35)

n+1 case: we have
$$\sum_{j=1}^{n+1} P(A_j) - (n+1) + 1 = \sum_{j=1}^{n} P(A_j) + P(A_{n+1}) - (n+1) + 1 \quad (36)$$

Now notice we can rewrite $P(A_1 \cap ... \cap A_{n+1}) = P(B \cap A_{n+1})$ where $B = A_1 \cap ... \cap A_n$

(37)

$$=> P(B \cap A + n + 1) = P(B) + P(A_{n+1}) - P(A_{n+1} \cup B)$$
 by definition (38)

By the induction hypothesis, we know
$$P(B) \ge \sum_{j=1}^{n} P(A_j) - n + 1$$
 (39)

That means all that is left to show $isP(A_{n+1}) - P(A_{n+1} \cup B) \ge P(A_{n+1}) - 1$ (40)

$$=> P(A_{n+1}) - P(A_{n+1} \cup B) \ge P(A_{n+1}) - 1 => -P(A_{n+1} \cup B) \ge -1 => P(A_{n+1} \cup B) \le 1$$

$$(41)$$

We know this inequality is true as
$$0 \le P(X) \le 1$$
 (42)

Thus, we have proved the induction step. \Box (43)

5. For $X \sim Pois(\lambda)$, find $E[2^X]$, if it is finite.

$$E[2^X] = \sum_{n=0}^{\infty} \frac{2^n e^{-\lambda} \lambda^n}{n!} = \sum_{n=0}^{\infty} \frac{e^{-\lambda} (2 * \lambda)^n}{n!} = e^{-\lambda} \sum_{n=0}^{\infty} \frac{(2 * \lambda)^n}{n!}$$
(44)

now let
$$\lambda' = 2 * \lambda = e^{-\lambda} \sum_{n=0}^{\infty} \frac{(\lambda')^n}{n!}$$
 (45)

we recognize this is the taylor expansion of the exponential function (46)

we get
$$\sum_{n=0}^{\infty} \frac{(\lambda')^n}{n!} = e^{\lambda'} = e^{2\lambda} = E[2^X] = e^{-\lambda}e^{2\lambda} = e^{\lambda} \square$$
 (47)

6. For $X \sim \text{Geom}(p)$, find $E[2^X]$ and $E[2^{-X}]$ (if it is finite).

$$E[2^X] = \sum_{n=1}^{\infty} 2^n (1-p)^{n-1} * pletS = \sum_{n=1}^{\infty} 2^n (1-p)^{n-1} = 2 + 4(1-p) + \dots$$
 (48)

$$=> S * 2(1-p) = 4(1-p) + 8(1-p)^{2} + \dots => S - S * 2(1-p) = 2$$
(49)

$$S - 2S + 2Sp = -S + 2Sp = S(-1 + 2p) = 2 \implies S = \frac{2}{2p - 1} \implies E[2^X] = p * S = \frac{2p}{2p - 1}$$
(50)

Similarly
$$E[2^{-X}] = \sum_{n=1}^{\infty} 2^{-n} (1-p)^{n-1} * p \text{we let} S = \sum_{n=1}^{\infty} 2^{-n} (1-p)^{n-1} = \sum_{n=1}^{\infty} \frac{(1-p)^{n-1}}{2^n}$$
(51)

$$= \frac{(1-p)^0}{2^1} + \frac{(1-p)^1}{2^2} + \dots = S * \frac{1-p}{2} = \frac{(1-p)^1}{2^2} + \frac{(1-p)^2}{2^3}$$
 (52)

$$=> S - S * \frac{1-p}{2} = S(1 - \frac{1-p}{2}) = \frac{1+p}{2} = \frac{1}{2} => S = \frac{1}{1+p}$$
 (53)

$$=> E[2^{-X}] = \frac{p}{1+p} \square$$
 (54)