

Markov Chain Monte Carlo: Metropolis-Hastings & Hamiltonian Monte Carlo

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1 Introduction

All the code for this project (including code to generate the report and its graphs) can be found at <https://github.com/bollu/sampleraytracer>.

1.1 MCMC sampling: the fundamental algorithm

markov-chain-monte-carlo methods are a class of algorithms which blend together two or three properties; these properties blend together in subtle ways, and permit us to solve very different-looking-problems using the same toolkit, by leveraging different portions of the algorithm's power. Broadly, they all solve the following problem:

Given an **arbitrary** function $\mathfrak{P} : X \rightarrow \mathbb{R}$, **create a sampler** for the **normalized** probability distributed from \mathfrak{P} , $\mathfrak{P}(x) \equiv \frac{\mathfrak{P}(x)}{\int \mathfrak{P}(x)dx}$. Furthermore, allow the sampling process to be **guided by** a proposal distribution: $\text{Prop} : X \rightarrow (X \rightarrow [0, 1])$. This proposal, for each location $x_0 \in X$, permits the specification of a '**likely direction** of higher likelihood', $\text{Prop}(x_0) : X \rightarrow \mathbb{R}$.

1.2 Sampling use case 1. Simulation And Sampling

Often, we do really want samples from a particular distribution. For example, we might often want to apply Bayes' rule and sample from the posterior distribution. Assume that we have a computational procedure to compute $P(X = x_0|Y = y_0)$. We now want to $P(Y|X = x_0)$.

$$P(Y|X = x_0) = \frac{P(X = x_0|Y)P(Y)}{\sum_y P(X = x_0|Y = y)}$$

There are two problems in sampling the above distribution:

- 1 The expression $P(Y = y|X = x)$ has no a-priori structure. Hence, we will need an algorithm that can sample from an arbitrary distribution.
- 2 The denominator term, which is a normalization factor can be extremely expensive to compute due to the summation involved

Hence, we need to use MCMC to sample from $P(Y|X = x_0)$, and we can expedite this by sampling from $\mathfrak{P}(Y|X = x_0) = P(X = x_0|Y)P(Y) \propto P(Y|X = x_0)$

1.3 Sampling use case 2. Gradient Free Optimisation

We want to maximize a function $f : X \rightarrow \mathbb{R}$. However, we lack gradients for f , hence we cannot use techniques such as gradient descent, or other techniques from convex optimisation. In such a case, we can consider f as some sort of unnormalized probability distribution, and use MCMC to sample from f .

That is, we set $\mathfrak{P} \equiv f$. Now, samples from \mathfrak{P} will be 'more likely' to come from those regions where f is large, since a sampler will be more likely to sample from regions of high probability.

1.4 Sampling use case 3. Numerical Integration

We wish to calculate $\int_l^u f(x)dx$ where f is some complicated function with no closed form for the definite integral of f . In this case, we know that the average value of a function, denoted by $\mathbb{E}[f]$, can be computed as:

$$\mathbb{E}[f] \equiv \frac{1}{|u-l|} \int_l^u f(x)dx$$

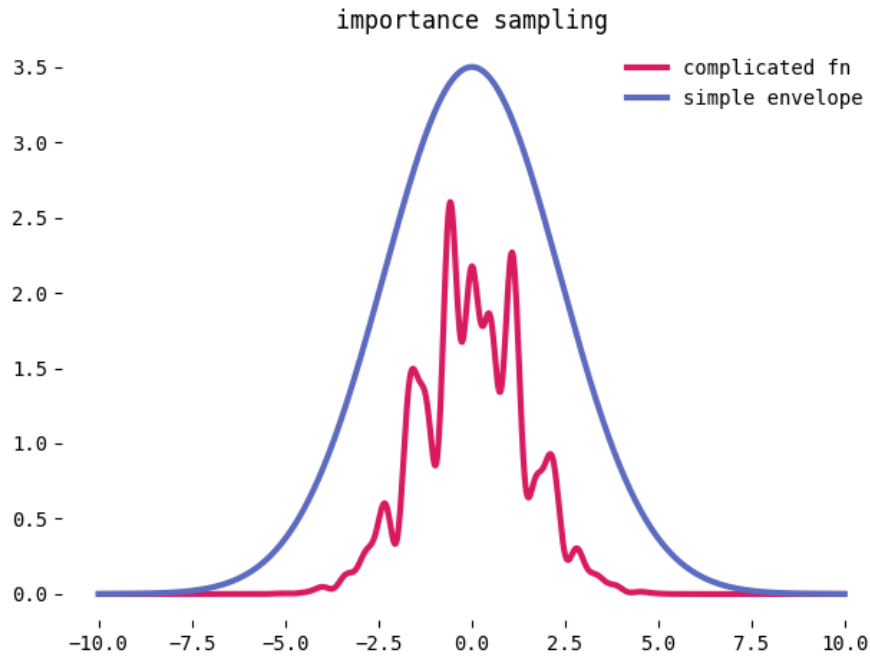
On rearrangement, we get:

$$\int_l^u f(x)dx = |u-l|\mathbb{E}[f]$$

Hence, if we have a good approximation of $\mathbb{E}[f]$, we have a good approximation for the integration. The naive approximation of $\mathbb{E}[f]$ can be found by using:

$$\mathbb{E}[f] \simeq \frac{1}{n} \sum_{i=1}^n f(x_i)$$

However, we often know *something* about f , just not its *precise* shape:



In this case, we can *guide* the sampling process, by setting the proposal **Prop** as the envelope of the function f . That way, we will be more likely to pick up *useful samples*, which rapidly sample from the region of space that has mass for the integration.

2 Where it all begins: The Metropolis Hastings sampler

2.1 The big idea

We wish to sample from a distribution \mathfrak{P} , but we do not know how to do so. The idea is that we build a markov chain $M[\mathfrak{P}, \text{Prop}]$ where $\mathfrak{P}, \text{Prop}$ are supplied by the user. We will show that the *stationary distribution* of $M[\mathfrak{P}, \text{Prop}]$ is going to be \mathfrak{P} . This will ensure that if we interpret *states of M* as samples, these samples will be distributed according to \mathfrak{P} .

2.2 Detail balance: A tool for proofs

We will first require a condition that will enable to rapidly establish that some distribution $\mathfrak{P} : X \rightarrow \mathbb{R}$ is the stationary distribution of a markov chain M . If the transition kernel $K : X \times X \rightarrow \mathbb{R}$ of $M \equiv (X, K)$ is such that:

$$\forall x, x' \in X, \mathfrak{P}(x)K(x, x') = \mathfrak{P}(x')K(x', x)$$

Then \mathfrak{P} is said to be **detail balanced** with respect to M .

Theorem 1 *If \mathfrak{P} is detail balanced to $M \equiv (X, K : X \rightarrow (X \rightarrow [0, 1]))$, then \mathfrak{P} is the stationary distribution of M .*

Proof 1.1 *Let \mathfrak{P} be detail balanced to M . This means that:*

$$\forall x, x' \in X, \mathfrak{P}(x)K(x)(x') = \mathfrak{P}(x')K(x')(x)$$

Let us say that we are in state \mathfrak{P} . We wish to find the probability distribution after one step of transition. The probability of being in some state x'_0 is going to be:

$$P_{\text{next}}(x'_0) \equiv \sum_x \mathfrak{P}(x)K(x)(x'_0)$$

since we have $\mathfrak{P}(x)$ probability to be at a given x , and $K(x, x'_0)$ probability to go from x to x'_0 . If we add over all possible $x \in X$, we get the probability of all states to enter in x'_0 . Manipulating P_{next} , we get:

$$\begin{aligned}
P_{next}(x'_0) &\equiv \sum_x \mathfrak{P}(x)K(x)(x'_0) \\
&= \sum_x \mathfrak{P}(x'_0)K(x'_0)(x) \quad (\text{by detail balance}) \\
&= \mathfrak{P}(x'_0) \sum_x K(x'_0)(x) \quad (P(x_0) \text{ is constant}) \\
&= \mathfrak{P}(x'_0) \cdot 1 \quad (K(x'_0) \text{ is a distribution which is being summed over}) \\
&= \mathfrak{P}(x'_0) \quad (\text{eliminate multiplication with 1})
\end{aligned}$$

Hence, $P_{next}(n'_0) = \mathfrak{P}(x'_0)$ if \mathfrak{P} is the current state, and \mathfrak{P} is in detail balance with the kernel K . This means that \mathfrak{P} is the stationary distribution of M . \triangle

2.3 Metropolis Hastings

There are three key players in the metropolis hastig sampler:

- 1 $\mathfrak{P} : X \rightarrow \mathbb{R}$: the probability distribution we wish to sample from.
- 2 $\mathbf{Prop} : X \rightarrow (X \rightarrow \mathbb{R})$. For each $x_0 \in X$, provide a distribution $\mathbf{Prop}(x) : X \rightarrow \mathbb{R}$ that is used to sample points around x_0 . \mathbf{Prop} for *proposal*.
- 3 $M[\mathfrak{P}, \mathbf{Prop}] \equiv (X, K[\mathfrak{P}, \mathbf{Prop}] : X \rightarrow \mathbb{R})$: The Metropolis Markov chain we will sample from, whose stationary distribution is $\mathfrak{P} — M$ for *Markov*.

We want the stationary distribution of $M[\mathfrak{P}, \mathbf{Prop}]$ to be \mathfrak{P} . We also wish for $K[\mathfrak{P}, \mathbf{Prop}](x_0) \sim \mathbf{Prop}(x_0)$: That is, at a point x_0 , we want to choose new points in a way that is 'controlled' by the proposal distribution $\mathbf{Prop}(x_0)$, since this will allow us to 'guide' the markov chain towards regions where \mathfrak{P} is high. If we had a gradient, then we could use \mathfrak{P}' to 'move' from the current point x_0 to a new point. Since we lack a gradient, we will provide a custom $\mathbf{Prop}(x_0)$ for each x_0 that will tell us how to pick a new x' , in a way that will improve \mathfrak{P} . So, we tentatively define

$$K[\mathfrak{P}, \mathbf{Prop}](x)(x') \stackrel{?}{=} \mathbf{Prop}(x)(x').$$

Recall that for \mathfrak{P} to be a stationary distribution of K , it is sufficient for \mathfrak{P} to be in detail balance for K . So, we write:

$$\begin{aligned}
\mathfrak{P}(x)K(x, x') &\stackrel{?}{=} \mathfrak{P}(x')K(x', x) \quad (\text{going forward equally likely as coming back}) \\
\mathfrak{P}(x)\mathbf{Prop}(x)(x') &\stackrel{?}{=} \mathfrak{P}(x')\mathbf{Prop}(x')(x) \quad (\text{this is a hard condition to satisfy})
\end{aligned}$$

This is far too complicated a condition to impose on \mathbf{Prop} and \mathfrak{P} , and there is no reason for this condition to be satisfied in general. Hence, we add a custom

"fudge factor" $\alpha \in X \rightarrow (X \rightarrow \mathbb{R})$ that tells us how often to transition from x to x' . We redefine the kernel as:

$$K[\mathfrak{P}, \text{Prop}](x)(x') \equiv \text{Prop}(x)(x')\alpha(x)(x')$$

Redoing detail balance with this new K , we get:

$$\begin{aligned} \mathfrak{P}(x)K(x, x') &\stackrel{?}{=} \mathfrak{P}(x')K(x', x) \quad (\text{going forward equally likely as coming back}) \\ \mathfrak{P}(x)\text{Prop}(x)(x')\alpha(x)(x') &= \mathfrak{P}(x')\text{Prop}(x')(x)\alpha(x')(x) \quad (\text{Fudge a hard condition with } \alpha) \\ \frac{\alpha(x)(x')}{\alpha(x')(x)} &= \frac{\mathfrak{P}(x')\text{Prop}(x)(x')}{\mathfrak{P}(x)\text{Prop}(x')(x)} \quad (\text{Find conditions for } \alpha) \end{aligned}$$

What we have above is a *constraint* for α . We now need to *pick* an α that satisfies this. A reasonable choice is:

$$\alpha(x)(x') \equiv \min \left(1, \frac{\mathfrak{P}(x')\text{Prop}(x')(x)}{\mathfrak{P}(x)\text{Prop}(x)(x')} \right)$$

since K is a transition kernel, we cannot have its entries be greater than 1. Hence, we choose to clamp it with a $\min(1, \cdot)$. This finally gives us the kernel as:

$$\begin{aligned} K[\mathfrak{P}, \text{Prop}](x)(x') &\equiv \text{Prop}(x)(x')\alpha(x)(x') \\ K[\mathfrak{P}, \text{Prop}](x)(x') &= \text{Prop}(x)(x') \min \left(1, \frac{\mathfrak{P}(x')\text{Prop}(x')(x)}{\mathfrak{P}(x)\text{Prop}(x)(x')} \right) \\ K[\mathfrak{P}, \text{Prop}](x)(x') &= \min \left(\text{Prop}(x)(x'), \frac{\mathfrak{P}(x')\text{Prop}(x')(x)}{\mathfrak{P}(x)} \right) \end{aligned}$$

We can make sure that detail balance is satisfied:

$$\begin{aligned} \mathfrak{P}(x)K[\mathfrak{P}, \text{Prop}](x)(x') &= \mathfrak{P}(x) \min \left(\text{Prop}(x)(x'), \frac{\mathfrak{P}(x')\text{Prop}(x')(x)}{\mathfrak{P}(x)} \right) \\ &= \min \left(\mathfrak{P}(x)\text{Prop}(x)(x'), \mathfrak{P}(x) \frac{\mathfrak{P}(x')\text{Prop}(x')(x)}{\mathfrak{P}(x)} \right) \\ &= \min (\mathfrak{P}(x)\text{Prop}(x)(x'), \mathfrak{P}(x')\text{Prop}(x')(x)) \end{aligned}$$

Note that the above right-hand-side is symmetric in x and x' , and hence we can state that:

$$\mathfrak{P}(x)K[\mathfrak{P}, \text{Prop}](x)(x') = \min (\mathfrak{P}(x)\text{Prop}(x)(x'), \mathfrak{P}(x')\text{Prop}(x')(x)) = \mathfrak{P}(x')K[\mathfrak{P}, \text{Prop}](x')(x)$$

Hence, we can wrap up, stating that our design of K does indeed give us a markov chain whose stationary distribution is \mathfrak{P} , since \mathfrak{P} is detail balanced with $K[\mathfrak{P}, \text{Prop}]$. As an upshot, we also gained a level of control with Prop , where we are able to provide "good" samples for a given point.

2.4 Simplification when proposal is symmetric

If our function Prop is symmetric: $\forall x, x' \text{Prop}(x)(x') = \text{Prop}(x')(x)$, then a lot of the above derivation becomes much simpler. We will perform those simplifications here for pedagogy.

When we have $\text{Prop}(x)(x') = \text{Prop}(x')(x)$, we can simply α :

$$\begin{aligned}\alpha(x)(x') &\equiv \min \left(1, \frac{\mathfrak{P}(x')\text{Prop}(x')(x)}{\mathfrak{P}(x)\text{Prop}(x)(x')} \right) \\ \alpha(x)(x') &= \min \left(1, \frac{\mathfrak{P}(x')}{\mathfrak{P}(x)} \right) \quad [\text{cancelling: } \text{Prop}(x)(x') = \text{Prop}(x')(x)]\end{aligned}$$

This also makes the kernel look a lot more pleasing:

$$\begin{aligned}K[\mathfrak{P}, \text{Prop}](x)(x') &\equiv \text{Prop}(x)(x')\alpha(x)(x') \\ K[\mathfrak{P}, \text{Prop}](x)(x') &= \text{Prop}(x)(x') \min \left(1, \frac{\mathfrak{P}(x')}{\mathfrak{P}(x)} \right)\end{aligned}$$

```

# prob is the distribution to sample from;
# symproposol is the *symmetric* proposal function
# symproposol: X -> X; produces a new 'X' from a
#             given 'X' with some distribution.
# prob: X -> |R: gives probability of point 'xi'.
# N: number of markov chain walks before returning a new sample.
def metropolis_hastings(prob, symproposol, x0, N):
    x = x0
    while True:
        for i in range(N):
            # xnext chosen with Prop(x)(x') prob.
            xnext = symproposol(x); px = prob(x); pxnext = prob(xnext);
            # x' chosen with Prop(x)(x') * alpha prob.
            r = uniform01(); alpha = min(1, pxnext / px); if r < alpha: x = xnext
        yield x

```

2.5 Implementing MH: Devil in the Details

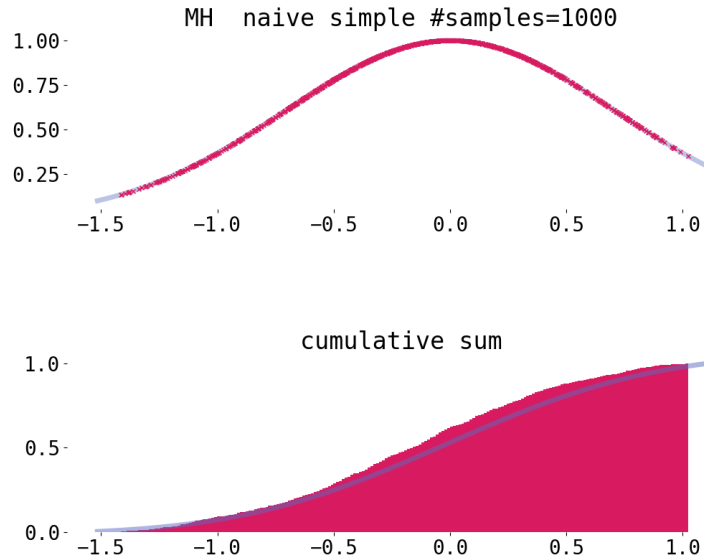
2.5.1 Naive implementation

Let us try to transcribe the equations we have derived for symmetric proposal into code, and see what happens. Let us choose $\mathfrak{P}(x) \equiv \text{normal}(0, 1) = e^{-x*x}$ and the proposal function to be $P(x_0)(x) = \text{normal}(x_0, 1e-2)$ is, a Gaussian centered around x_0 with standard deviation $1e-2$. This gives us the code:

```

# mcmc1d.py
def mhsimple(x0, prob, prop):
    yield x0; x = x0;
    while true:
        xnext = prop(x); p = prob(x); pnext = prob(xnext)
        r = np.random.uniform() + 1e-5;
        if r < pnext/p: x = xnext
        yield xnext
    ...
def exp(x): return np.exp(-x*x)
def expprop(x): return np.random.normal(loc=x, scale=1e-1)
...
nsamples = 1000
xs = list(itertools.islice(mhsimple(0, exp, expprop), nsamples))

```

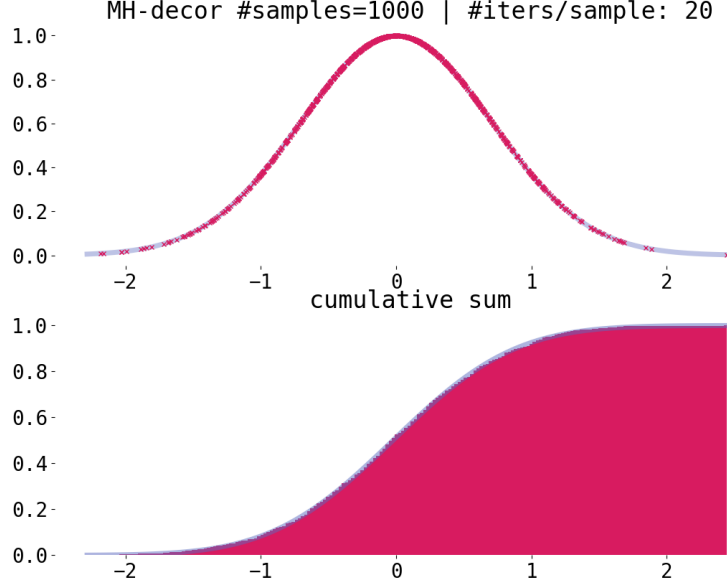



We can see from the plots of the raw samples and the cumulative distribution that we wind up overshooting. This is because the markov chain, by definition, has correlations between samples; However, we are supposed to be drawing *independent* samples from \mathfrak{P} for plotting/downstream use.

2.5.2 Sampling per `iters_per_sample` steps

The solution is to take samples that are *spaced out* — that is, we do not consider each step from the markov chain as a sample. Rather, we consider the state we are in after some `iters_per_sample` steps to be a sample. In code, this is now:

```
# mcmc1d.py
def mh_uncorr(x0, prob, prop, iters_per_sample):
    yield x0; x = x0;
    while True:
        for i in range(iters_per_sample):
            xnext = prop(x); p = prob(x); pnext = prob(xnext)
            r = np.random.uniform() + 1e-5;
            if r < min(1, pnext/p): x = xnext
        yield xnext
...
def exp(x): return np.exp(-x*x)
def expprop(x): return np.random.normal(loc=x, scale=1e-1)
...
NSAMPLES = 1000; ITERS_PER_SAMPLE = 20
xs = list(itertools.islice(mhsimple(0, exp, expprop, ITERS_PER_SAMPLE), NSAMPLES))
```



3 Hamiltonian Monte Carlo

If our target probability density $\mathfrak{P} : X \rightarrow [0, 1]$ is *differentiable*, then we can use the derivative of $\mathfrak{P}(x)$ to provide better proposals. The idea is as follows:

- Interpret the probability landscape as a potential, with points of high probability being "valleys" and points of low probability being "peaks" (ie, invert the probability density with a transform such as $U[\mathfrak{P}](x) = -\log \mathfrak{P}(x)$. This way, a ball rolling on this terrain will try to move towards the valleys — which are the locations of high probability \mathfrak{P} .
- For a proposal at a position $x_0 \in X$, keep a ball at x_0 , *randomly choose its velocity*, simulate the ball according to classical mechanics (Newton's laws of motion) for a fixed duration $D \in \mathbb{R}$ and propose the final position as the final position of the ball. This is reversible and detail balanced because *classical mechanics is reversible and detail balanced*.

We wish to sample from a new distribution P_H , which is defined as:

$$\begin{aligned}
P_H[\mathfrak{P}](\mathbf{p}, \mathbf{q}) &\equiv \exp(-H[\mathfrak{P}](\mathbf{p}, \mathbf{q})) && \text{(Sample states with low energy: Boltzmann/Ising model)} \\
&= \exp(-(U[\mathfrak{P}](\mathbf{q}) + K(\mathbf{p}))) && \text{(by defn. of hamiltonian, } H[\mathfrak{P}] = U[\mathfrak{P}] + K) \\
&= \exp(-(-\log \mathfrak{P}(\mathbf{q}) + \mathbf{p}^T \mathbf{p}/2)) && \text{(by choice of } U \text{ as NLL, } K \text{ as classical KE)} \\
&= \exp(\log \mathfrak{P}(\mathbf{q}) - \mathbf{p}^T \mathbf{p}/2) \\
&= \exp(\log \mathfrak{P}(\mathbf{q})) \cdot \exp(-\mathbf{p}^T \mathbf{p}/2) \\
&= \mathfrak{P}(\mathbf{q}) \cdot \exp(-\mathbf{p}^T \mathbf{p}/2)
\end{aligned}$$

Note that the probability distribution for $P_H[\mathfrak{P}]$ has split into two independent distributions: $\mathfrak{P}(\mathbf{q})$, and $\exp(-\mathbf{p}^T \mathbf{p}/2)$. Hence, sampling $(\mathbf{p}, \mathbf{q}) \sim P_H[\mathfrak{P}]$ correctly is *equivalent* to sampling $\mathbf{q} \sim \mathfrak{P}$, and $\mathbf{p} \sim \exp(-\mathbf{p}^T \mathbf{p}/2)$. Therefore, if we can sample from P_H effectively, taking only $(\mathbf{q}, -) \sim P_H[\mathfrak{P}]$ will correctly give us the samples we are looking for.

We will show that using MCMC to sample from the $P_H[\mathfrak{P}]$ above, with a particular proposal function Prop_H that we shall define momentarily will provide us with samples (\mathbf{p}, \mathbf{q}) such that the \mathbf{q} is sampled from \mathfrak{P} . This begs the question: "why do we need \mathbf{p} anyway?" We will show how introducing this \mathbf{p} allows for rapid exploration of the space of possible \mathbf{q} s.

3.1 The proposal function for HMC

In classical mechanics, once we have a hamiltonian $H(\mathbf{q}, \mathbf{p})$, the dynamics of the system are simulated according to the equations (called Hamilton's equations):

$$\frac{\partial \mathbf{q}}{\partial t} = \frac{\partial H[\mathfrak{P}]}{\partial \mathbf{p}} \quad \frac{\partial \mathbf{p}}{\partial t} = -\frac{\partial H[\mathfrak{P}]}{\partial \mathbf{q}}$$

These equations, specialized to our situation become:

$$\begin{aligned}
\frac{\partial \mathbf{q}}{\partial t} &= \frac{\partial H[\mathfrak{P}]}{\partial \mathbf{p}} = \frac{\partial (U[\mathfrak{P}](\mathbf{q}) + K(\mathbf{p}))}{\partial \mathbf{p}} = \frac{\partial K(\mathbf{p})}{\partial \mathbf{p}} = \frac{\partial (\mathbf{p}^T \mathbf{p})}{\partial \mathbf{p}} = 2\mathbf{p} \\
\frac{\partial \mathbf{p}}{\partial t} &= -\frac{\partial H[\mathfrak{P}]}{\partial \mathbf{q}} = -\frac{\partial (U[\mathfrak{P}](\mathbf{q}) + K(\mathbf{p}))}{\partial \mathbf{q}} = -\frac{\partial U[\mathfrak{P}](\mathbf{q})}{\partial \mathbf{q}} = \frac{-\partial -\log \mathfrak{P}(\mathbf{q})}{\partial \mathbf{q}} = \frac{\partial \log \mathfrak{P}(\mathbf{q})}{\partial \mathbf{q}}
\end{aligned}$$

We will denote a solution to the above system of equations as $(\mathbf{q}_{final}, \mathbf{p}_{final}) \equiv X_H[\mathfrak{P}](\mathbf{q}_0, \mathbf{p}_0, T_0)$. That is, starting from the position $(\mathbf{p}_0, \mathbf{q}_0)$, simulate the system for $T_0 \in \mathbb{R}$ time, according to Hamilton's equations. The final (position, momentum) is $(\mathbf{q}_{final}, \mathbf{p}_{final})$.

Now, our proposal function will be:

$$\text{Prop}(\mathbf{q}_0, -) \equiv \text{let } \mathbf{p}_0 \sim \text{normal}(0, 1) \text{ in } (X_H[\mathfrak{P}](\mathbf{q}_0, \mathbf{p}_0, T_0))$$

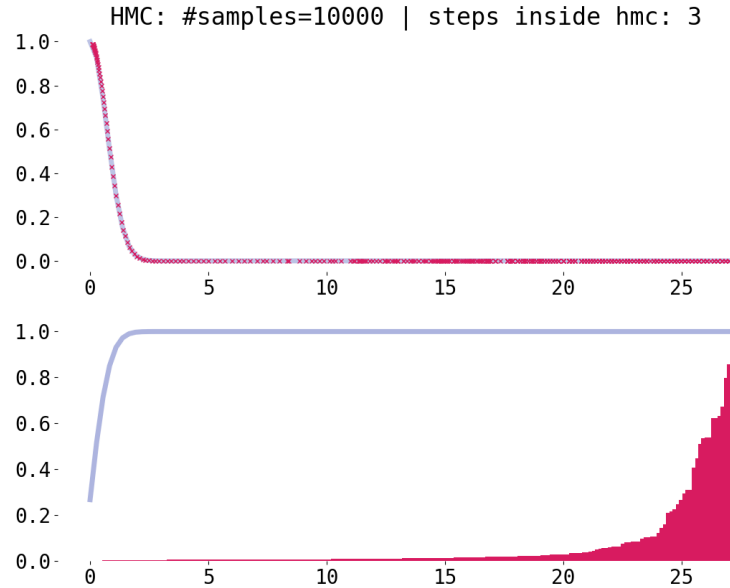
That is, we will propose the new point as one obtained from time evolution of the solutions to the hamiltonian; (a) we start from the point \mathbf{q}_0 ; (b) pick a *random velocity* \mathbf{p}_0 ; (c) simulate the system for T_0 time using hamilton's equations; (d) return the final (position, momentum) as our proposal.

3.2 Simulating the above regime

```
def euler(dhdp, dhdq, q, p, dt): % f(x + dx) = f(x) + f'(x) * dx
    pnew = p + -dhdp(q, p) * dt; qnew = q + dhdp(q, p) * dt
    return (qnew, pnew)

def hmc(q0, U, dU, nsteps, dt):
    # hamiltonian definition
    def h(q, p): return U(q) + 0.5*p*p
    def hdp(q, p): return p
    def hdq(q, p): return dU(q)
    def nextsample(q, p): # run 'euler' for n steps
        for _ in range(nsteps):
            (q, p) = euler(hdq, hdp, q, p, dt)
        return (q, p)

    yield q0; q = q0
    while True:
        p = np.random.normal(0, 1) # pick random momentum
        (qnext, pnext) = nextsample(q, p) # simulate next sample
        pnext = -p # reverse momentum so our process is reversible
        r = np.random.uniform(); # if accept according to MH, accept.
        if np.log(r) < h(q, p) - h(qnext, pnext): q = qnext
        yield q # return point
```

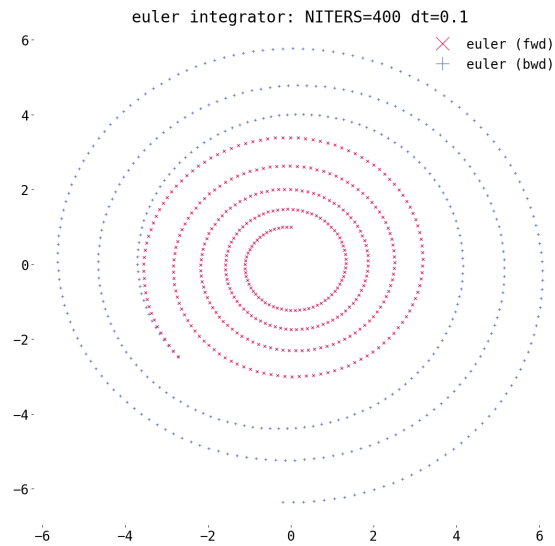


This has sampled disastrously: what is going wrong?

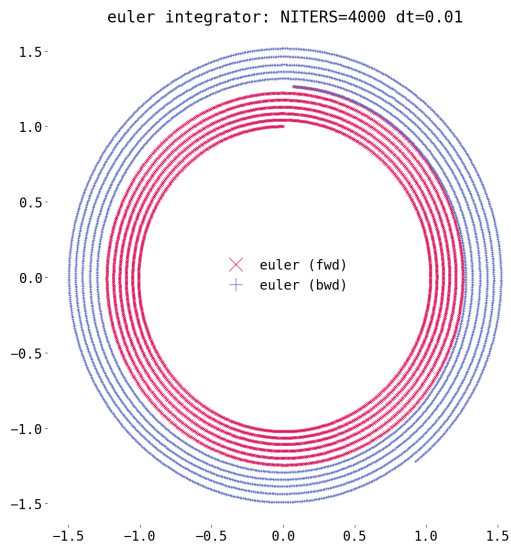
3.3 Simulation: Euler integration

Unfortunately, it turns out that running this simulation is in fact numerically unstable on using a naive simulation schemes. For example, let us say that we wish to simulate the orbit of a planet. Recall that we want the proposal to be symmetric: so, if we simulate the trajectory of the planet for N timesteps, each timestep of time δt , and then *reverse* the momentum of the planet, run the next phase of the simulation for N timesteps with timestep Δt , we should begin where we started. However, if we try to use the euler integration equations, here is what we see:

```
def euler(dhdp, dhdq, q, p, dt):
    pnew = p + -dhdq(q, p) * dt
    qnew = q + dhdp(q, p) * dt
    return (qnew, pnew)
```



We can clearly see the forward trajectory (in pink) spiralling out, and the backward trajectory (in blue), spiralling out even more. We can attempt to fix this by making Δt smaller:

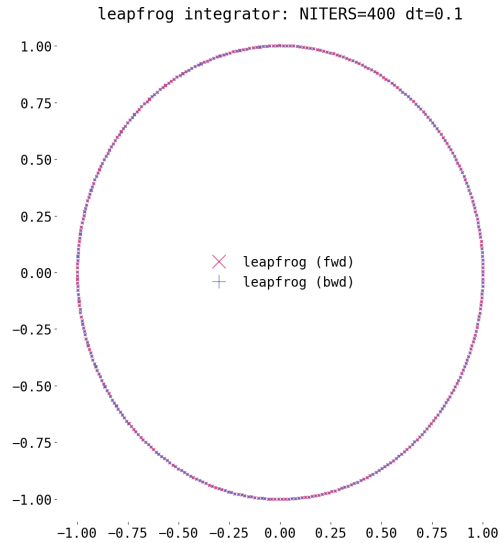


to no avail. Indeed, this is a **fundamental limitation of euler integration**. Hence, we will need to explore more refined integration schemes.

3.4 Simulation: Symplectic integrator

The type of integrators that will allow us to get 'reasonable orbits' that do not decay with time are known as *symplectic integrators*. (An aside: the word *symplectic* comes from Weyl, who substituted the latin root in the word *complex* by the corresponding greek root. It is a branch of differential geometry which formalizes the constructs needed to carry out hamiltonian mechanics on spaces that are more complicated than Euclidian \mathbb{R}^n).

```
## dq/dt = dH/dp|_{p0, q0}, dp/dt = -dH/dq|_{p0, q0}
def leapfrog(dhdp, dhdq, q0, p0, dt):
    p0 += -dhdq(q0, p0) * 0.5 * dt # kick: half step momentum
    q0 += dhdp(q0, p0) * dt # drift: full step position
    p0 += -dhdq(q0, p0) * 0.5 * dt # kick: half step momentum
    return (q0, p0)
```



It is clear from the plots that the leapfrog integrator is stable: orbits stay as orbits. Indeed, we can prove the symplecticity/energy-preserving property of this integrator.

3.5 Proof that leapfrog is symplectic

We will show that in phase space, the leapfrog integrator preserves infinitesimal volumes of the form $dV \equiv d\mathbf{p} \times d\mathbf{q}$.

3.6 Using HMC on the 1D gaussian

Now that we have the math to setup a symmetric proposal distribution, let's code up HMC on a 1D gaussian:

```
def leapfrog(dhdq, dhdp, q0, p0, dt):
    p0 += -dhdq(q0, p0) * 0.5 * dt # kick: half step momentum
    q0 += dhdp(q0, p0) * dt # drift: full step position
    p0 += -dhdq(q0, p0) * 0.5 * dt # kick: half step momentum
    return (q0, p0)

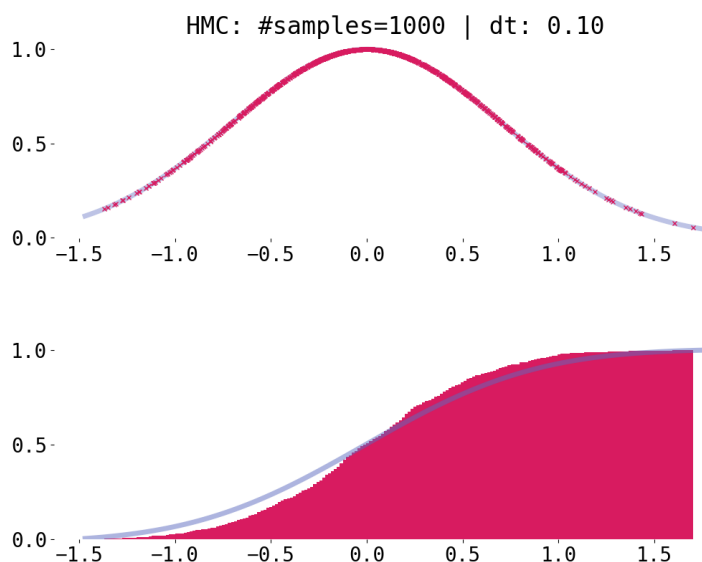
def hmc(q0, U, dU, nsteps, dt):
    def h(q, p): return U(q) + 0.5 * p*p
    def nextsample(q, p):
        for _ in range(nsteps):
            def hdp(q, p): return p
            def hdq(q, p): return dU(q)
            (q, p) = leapfrog(hdq, hdp, q, p, dt)
        return (q, p)
```



```

yield q0; q = q0
while True:
    p = np.random.normal(0, 1)
    (qnext, pnext) = nextsample(q, p)
    pnext = -p # reverse momentum so our process is reversible
    r = np.random.uniform();
    if np.log(r) < h(q, p) - h(qnext, pnext): q = qnext
    yield q
...
def neglogexp(x): return -1 * logexp(x)
def neglogexpgrad(x): return -1 * logexpgrad(x)
xs = list(take_every_nth(DECORRELATE_STEPS,
    itertools.islice(hmc(1, neglogexp, neglogexpgrad, NSTEPS, DT), NSAMPLES*DECORRELATE_STEPS)))

```

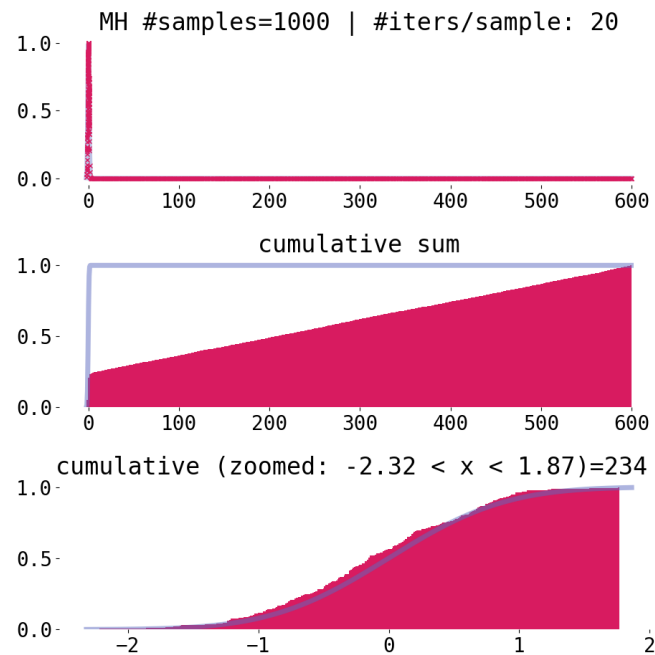


This looks far worse than the results we got from vanilla metropolis hastings: why would anybody use this technique?

3.7 HMC v/s MH when the proposal is atypical

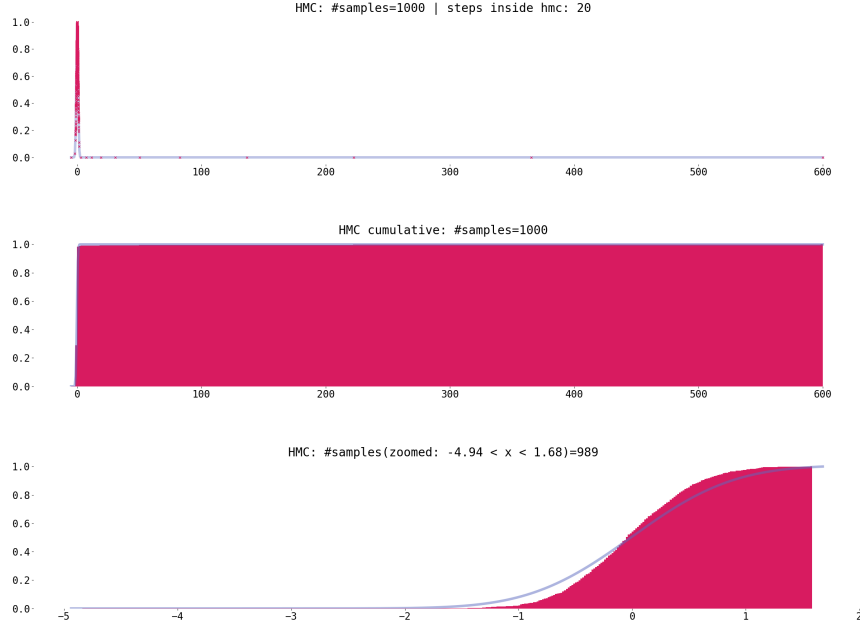
So far, I've hidden one thing under the rug: the *initialization* of the distribution. We've been starting at $x = 1$ — let's now change that, and start at $x = 600$. Note that this is very far from the "region of interest" in the case of a standard normal distribution: 99% of the probability mass is in $[-3, 3]$ (the 3σ rule).

3.7.1 MH starting at $x = 600$



Note that out of the 1000 samples we started from, only 234 are present in the region of $-2 \leq x \leq 2$. We have wasted around 80% of the samples we have *moving through the space* to get to the typical set.

3.7.2 HMC starting at $x = 600$



Note that out of the 1000 samples we started from, only 987 are present in the region of $-2 \leq x \leq 2$. Notice the distribution of samples (pink crosses) in the hamiltonian monte-carlo case: most samples clustered around the gaussian; our initial samples that are far away are powered by a strong potential energy to move towards the center.

This is in stark contrast to the MH case, where the entire x -axis is pink, due to the 'current point' having to move, proposal-by-proposal (which allows at most a movement of distance 1), from 600 to 0.