A Network Model for Dynamic Textual Communications with Application to Government Email Corpora

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Abstract

We introduce the interaction-partitioned topic model (IPTM)—a probabilistic model for who communicates with whom about what, and when. Broadly speaking, the IPTM partitions time-stamped textual communications, according to both the network dynamics that they reflect and their content. To define the IPTM, we integrate a dynamic version of the exponential random graph model—a generative model for ties that tend toward structural features such as triangles—and latent Dirichlet allocation—a generative model for topic-based content. The IPTM assigns each topic to an "interaction pattern"—a generative process for ties that is governed by a set of dynamic network features. Each communication is then modeled as a mixture of topics and their corresponding interaction patterns. We use the IPTM to analyze emails sent between department managers in Dare county government in North Carolina, and demonstrate that the model is effective at predicting and explaining continuous-time textual communications.

1. Introduction

In recent decades, real-time digitized textual communication has developed into a ubiquitous form of social and professional interaction (Kanungo & Jain, 2008; Szóstek, 2011; Burgess et al., 2004; Pew, 2016). From the perspective of the computational social scientist, this has lead to a growing need for methods of modeling interactions that manifest as text exchanged in continuous time. A number of models that build upon topic modeling through Latent Dirichlet Allocation (Blei et al., 2003) to incorporate link data as well as textual content have been developed recently (McCallum et al., 2005; Lim et al., 2013; Krafft et al., 2012). These

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models are innovative in their extensions that incorporate network tie information. However, none of the models that are currently available in the literature integrate the rich random-graph structure offered by state of the art models for network structure—in particular, the exponential random graph model (ERGM) (Robins et al., 2007; Chatterjee et al., 2013; Hunter et al., 2008). The ERGM is the canonical model for network structure, as it is flexible enough to specify a generative model that accounts for nearly any pattern of tie formation (e.g., reciprocity, clustering, popularity effects) (Desmarais & Cranmer, 2017). We build upon recent extensions of ERGM that model time-stamped ties (Perry & Wolfe, 2013; Butts, 2008), and develop the interaction-partitioned topic model (IPTM) which simultaneously models the network structural patterns that govern tie formation, and the content in the communications.

ERGM, and models based on ERGM, provide a framework for explaining or predicting ties between nodes using the network sub-structures in which the two nodes are embedded (e.g., an ERGM specification may predict ties between two nodes that have many shared partners). ERGM-style models have been used for many applications in which the ties between nodes are annotated with text. The text, despite providing rich information regarding the strength, scope, and character of the ties, has been largely excluded from these analyses, due to the inability of ERGM-style models to incorporate textual attributes of ties. These application domains include, among other applicaitons, the study of legislative networks in which networks reflect legislators' co-support of bills, but exclude bill text (Bratton & Rouse, 2011; Alemán & Calvo, 2013); the study of alliance networks in which networks reflect countries' co-signing of treaties, but exclude treaty text (Camber Warren, 2010; Cranmer et al., 2012b;a; Kinne, 2016); the study of scientific co-authorship networks that exclude the text of the co-authored papers (Kronegger et al., 2011; Liang, 2015; Fahmy & Young, 2016); and the study of text-based interaction on social media (e.g., users tied via 'mentions' on twitter) (Yoon & Park, 2014; Peng et al., 2016; Lai et al., 2017).

In defining and testing the IPTM we embed core conceptual property—interaction pattern—to link the content compo-
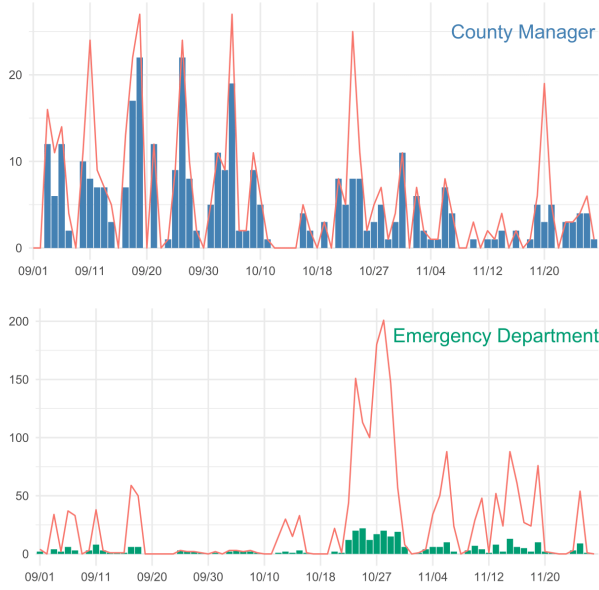
Figure 1. Sending behavior of two most active nodes in Dare County email data between 09/01/2012 and 11/30/2012. *Top*: the number of emails per day sent by County manager (blue bar) and the number of recipients per day (red line). *Bottom*: the number of emails per day sent by emergency department official (green bar) and the number of recipients per day (red line).

nent of the model, and network component of the model such that knowing who is communicating with whom at what time (i.e., the network component) provides information about the content of communication, and vice versa (Section **??**). Figure 1 Plot needs to be replaced illustrates this structure. IPTM leads to an efficient MCMC inference algoritmIn (Section 3) and acheives good predictive peformance (Section 5). Finally, the IPTM discovers interesting and interpretable latent structure through application to email corpora of internal communications by government officials in Dare County, NC (Section 6).

2. Interaction-partitioned Topic Model

Data generated under the IPTM consists of D unique documents. A single document, indexed by $d \in [D]$, is represented by the four components: the author $a_d \in [A]$, an indicator vector of recipients $\boldsymbol{r}_d = \{u_{dj}\}_{j=1}^A$, the timestamp $t_d \in (0, \infty)$, and a set of tokens $\boldsymbol{w}_d = \{w_{dn}\}_{n=1}^{N_d}$ that comprise the text of the document, where N_d denotes the total number of tokens in a document. For simplicity, we assume that documents are ordered by time such that $t_d < t_{d+1}$.

2.1. Content Generating Process

The words $\boldsymbol{w}_d$ are generated according to latent Dirichlet allocation (LDA) (Blei et al., 2003), where we generate

the corpus-wide global variables that describe the content via topics. As in LDA, we model each topic $k \in [K]$ as a discrete distribution over V unique word types

$$\boldsymbol{\phi}_k \sim \text{Dirichlet}\Big(\beta, (\frac{1}{V}, \ldots, \frac{1}{V})\Big), \qquad (1)$$

where β is the concentration parameter. Next, we assume a document-topic distribution over K topics

$$\boldsymbol{\theta}_d \sim \text{Dirichlet}(\alpha, \boldsymbol{m}), \qquad (2)$$

where α is the concentration parameter and $\boldsymbol{m} = (m_1, \ldots, m_K)$ is the probability vector. Given that N_d is known, a topic z_{dn} is drawn from the document-topic distribution and then a word w_{dn} is drawn from the chosen topic for each $n \in [N_d]$—i.e.,

$$\begin{aligned} z_{dn} &\sim \text{Multinomial}(\boldsymbol{\theta}_d), \\ w_{dn} &\sim \text{Multinomial}(\phi_{z_{dn}}). \end{aligned} \qquad (3)$$

2.2. Interaction Patterns

They key idea that combines the IPTM component modeling "what" with the component modeling "who," "whom," and "when" is that different topics are associated with different interaction patterns. Each interaction pattern $c \in [C]$ is characterized by a set of dynamic network features—such as the number of messages sent from i to j in some time interval—and corresponding coefficients. We associate each topic with the interaction pattern that best describes how people interact when talking about that topic.

The topic-interaction pattern assignments are discrete-uniform distributed,

$$l_k \sim \text{Uniform}(1, C). \qquad (4)$$

The content of each document is summarized as a distribution over interaction patterns:

$$\pi_{dc} = \frac{\sum_{k:l_k=c} N_{dk}}{N_d}, \qquad (5)$$

where N_{dk} is the number of times topic k appears in document d. In other words, for each document and interaction pattern, we compute the fraction of tokens that were generated using a topic corresponding to that interaction pattern. We then use this to generate the tie components, which are discussed in the next section.

2.3. Tie Generating Process

We generate ties—author a_d, recipients $\boldsymbol{r}_d$, and timestamp t_d—using a continuous-time process that depends on the interaction patterns' various features. Conditioned on the content (Section 2.1), we assume the following steps of tie generating process.

2.3.1. LATENT RECIPIENTS

For every possible author–recipient pair $(i, j)_{i \neq j}$, we define the "interaction-pattern-specific recipient intensity":

$$\nu_{idjc} = \boldsymbol{b}_c^\top \boldsymbol{x}_{idjc}, \qquad (6)$$

where $\boldsymbol{b}_c$ is P–dimensional vector of coefficients and $\boldsymbol{x}_{idjc}$ is a set of network features which vary depending on the hypotheses regarding canonical processes relevant to network theory such as popularity, reciprocity, and transitivity. We place a Normal prior $\boldsymbol{b}_c \sim N(\boldsymbol{\mu}_b, \Sigma_b)$.

We then compute the weighted average of $\{\nu_{idjc}\}_{c=1}^C$ and obtain the "recipient intensity"—the likelihood of document d being sent from i to j— using the the document's distribution over interaction patterns as mixture weights:

$$\lambda_{idj} = \sum_{c=1}^C \pi_{dc} \nu_{idjc}. \qquad (7)$$

Next, we hypothesize "If i were the author of document d, who would be the recipent/recipients?" To do this, we draw each author's set of recipients from a non-empty Gibbs measure (Fellows & Handcock, 2017)—a probability measure we defined in order to 1) allow multiple recipients or "multicast", 2) prevent from obtaining zero recipient, and 3) ensure tractable normalizing constant.

Because the IPTM allows multicast, we draw a binary (0/1) vector $\boldsymbol{u}_{id} = (u_{id1}, \dots, u_{idA})$

$$\boldsymbol{u}_{id} \sim \text{Gibbs}(\delta, \boldsymbol{\lambda}_{id}), \qquad (8)$$

where δ is a real number controlling the average number of recipients and $\boldsymbol{\lambda}_{id} = \{\lambda_{idj}\}_{j=1}^A$. We place a Normal prior $\delta \sim N(\mu_\delta, \sigma_\delta^2)$. In particular, we define $\text{Gibbs}(\delta, \boldsymbol{\lambda}_{id})$ as

$$p(\boldsymbol{u}_{id}|\delta, \boldsymbol{\lambda}_{id})$$
$$= \frac{\exp\left\{\log\left(\text{I}(\|\boldsymbol{u}_{id}\|_1 > 0)\right) + \sum_{j \neq i}(\delta + \lambda_{idj})u_{idj}\right\}}{Z(\delta, \boldsymbol{\lambda}_{id})}, \qquad (9)$$

where $Z(\delta, \boldsymbol{\lambda}_{id})$ is the normalizing constant and $\|\cdot\|_1$ is the l_1–norm. We provide the derivation of the normalizing constant as a tractable form in the supplementary material.

2.3.2. LATENT TIMESTAMPS

Similarly, we hypothesize "If i were the author of document d, when would it be sent?" and define the "interaction-pattern-specific timing rate"

$$\xi_{idc} = \boldsymbol{\eta}_c^\top \boldsymbol{y}_{idc}, \qquad (10)$$

where $\boldsymbol{\eta}_c$ is Q–dimensional vector of coefficients with a Normal prior $\boldsymbol{\eta}_c \sim N(\boldsymbol{\mu}_\eta, \Sigma_\eta)$, and $\boldsymbol{y}_{idc}$ is a set of time-related covariates, which can be any feature that could affect

timestamps of the document. Some examples of the time-related covariates are described in Section 4.3.

The "timing rate" for author i is also computed from the weighted average of $\{\xi_{idc}\}_{c=1}^C$

$$\mu_{id} = \sum_{c=1}^C \pi_{dc} g^{-1}(\xi_{idc}), \qquad (11)$$

where $g(\cdot)$ is the appropriate link function such as identity, log, or inverse.

In modeling "when", we do not directly model the timestamp t_d. Instead, we assume the time-increment or "time to next document" (i.e., $\tau_d = t_d - t_{d-1}$) is drawn from a specific distribution in the exponential family. We follow the generalized linear model framework:

$$\begin{aligned} E(\tau_{id}) &= \mu_{id}, \\ V(\tau_{id}) &= V(\mu_{id}), \end{aligned} \qquad (12)$$

where τ_{id} is a positive real number. Possible choices of distribution include Exponential, Weibull, Gamma, and log-normal[1] distributions, which are commonly used in time-to-event modeling. Based on the choice of distribution, we may introduce any additional parameter (e.g., σ_τ^2) to account for the variance.

2.3.3. ACTUAL DATA

Finally, we choose the actual author, recipients, and timestamp—which will be observed—by selecting the author–recipient-set pair with the smallest time-increment (Snijders, 1996; 2017):

$$\begin{aligned} a_d &= \text{argmin}_i(\tau_{id}), \\ \boldsymbol{r}_d &= \boldsymbol{u}_{a_d d}, \\ t_d &= t_{d-1} + \tau_{a_d d}. \end{aligned} \qquad (13)$$

Therefore, it is an author-driven process in that the author of a document determines its recipients and its timestamp, based on the author's urgency to send the document to chosen recipients.

3. Posterior Inference

Given that we only observe the authors, recipients, timestamps, and tokens $\{(a_d, \boldsymbol{r}_d, t_d, \boldsymbol{w}_d)\}_{d=1}^D$ in real-world, our inference goal is to invert the generative process to obtain the posterior distribution over the unknown parameters, conditioned on the observed data and hyperparamters $\alpha, \beta, \boldsymbol{m}, \boldsymbol{\mu}_b, \Sigma_b, \boldsymbol{\mu}_\eta, \Sigma_\eta, \mu_\delta, \sigma_\delta^2$. After integrating out Φ

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[1]lognormal distribution is not exponential family but we can take the log-transformation and apply $\mu = E(\log(\tau_{id})) = \mu_{id}$ and $\sigma_\tau^2 = V(\log(\tau_{id})) = V(\mu_{id})$ using identity link function $g = I$.

and Θ using Dirichlet-multinomial conjugacy (Griffiths & Steyvers, 2004), we draw the samples using Markov chain Monte Carlo (MCMC) methods, repeatedly resampling the value of each parameter from its conditional posterior given the observed data, hyperparamters, and the current values of the other parameters. We express each parameters conditional posterior in a closed form using the data augmentation schemes in $\boldsymbol{u}$ (Tanner & Wong, 1987). In this section, we outline a Metropolis-within-Gibbs sampling algorithm and each latent variable's conditional posterior.

Since u_{idj} is a binary random variable, new values may be sampled directly using

$$
\begin{aligned}
P(u_{idj} = 1|\boldsymbol{u}_{id\setminus j}, \boldsymbol{z}, \boldsymbol{l}, \boldsymbol{b}, \delta, \boldsymbol{x}) &\propto \exp\{\delta + \lambda_{idj}\} \\
P(u_{idj} = 0|\boldsymbol{u}_{id\setminus j}, \boldsymbol{z}, \boldsymbol{l}, \boldsymbol{b}, \delta, \boldsymbol{x}) &\propto \mathrm{I}(\|\boldsymbol{u}_{id\setminus j}\|_1 > 0),
\end{aligned} \quad (14)
$$

which naturally prevent from the instances where the sender has no recipients to send the document.

The topic-interaction pattern assignment l_k is a discrete random variable and can be also directly sampled using

$$
\begin{aligned}
&P(l_k = c|\boldsymbol{l}_{\setminus k}, \boldsymbol{z}, \boldsymbol{b}, \boldsymbol{\eta}, \delta, \boldsymbol{u}, \boldsymbol{a}, \boldsymbol{r}, \boldsymbol{t}, \boldsymbol{x}, \boldsymbol{y}) \\
&\propto \prod_{d=1}^{D} \Big(\prod_{i=1}^{A} \frac{\exp\left\{\log\big(\mathrm{I}(\|\boldsymbol{u}_{id}\|_1 > 0)\big) + \sum_{j\neq i}(\delta + \lambda_{idj})u_{idj}\right\}}{Z(\delta, \boldsymbol{\lambda}_{id})} \\
&\quad \times \varphi_\tau(\tau_{a_d d}; \mu_{a_d d}, \sigma_\tau^2) \times \prod_{i\neq a_d} \big(1 - \Phi_\tau(\tau_{a_d d}; \mu_{id}, \sigma_\tau^2)\big) \Big),
\end{aligned} \quad (15)
$$

where φ_τ and Φ_τ are the probability density function (pdf) and cumulative distribution function (cdf) of the specified distribution of time-increments, respecitvely.

The conditional posterior for topic assignment z_{dn} is derived by multiplying the two sampling equations of LDA:

$$
\begin{aligned}
&p(z_{dn} = k|\boldsymbol{z}_{\setminus dn}, \boldsymbol{l}, \boldsymbol{b}, \boldsymbol{\eta}, \delta, \boldsymbol{u}, \boldsymbol{w}, \boldsymbol{a}, \boldsymbol{r}, \boldsymbol{t}, \alpha, \beta, \boldsymbol{m}) \\
&\propto (N_{dk,\setminus dn} + \alpha m_k) \times \frac{N_{w_{dn}k,\setminus dn} + \frac{\beta}{V}}{N_{k,\setminus dn} + \beta} \\
&\times \prod_{i=1}^{A} \frac{\exp\left\{\log\big(\mathrm{I}(\|\boldsymbol{u}_{id}\|_1 > 0)\big) + \sum_{j\neq i}(\delta + \lambda_{idj})u_{idj}\right\}}{Z(\delta, \boldsymbol{\lambda}_{id})} \\
&\times \varphi_\tau(\tau_{a_d d}; \mu_{a_d d}, \sigma_\tau^2) \times \prod_{i\neq a_d} \big(1 - \Phi_\tau(\tau_{a_d d}; \mu_{id}, \sigma_\tau^2)\big),
\end{aligned} \quad (16)
$$

where the subscript $\setminus dn$ denote the exclusion of document d and n^{th} element in document d, and $N_{w_{dn}k,\setminus dn}$ is the number of tokens assigned to topic k whose type is the same as that of w_{dn}, excluding w_{dn} itself.

New values for continuous random variables δ, $\boldsymbol{b}$, and $\boldsymbol{\eta}$ and σ_τ^2 (if applicable) cannot be sampled directly from their conditional posteriors, but may instead be obtained using

the Metropolis–Hastings algorithm. With uninformative priors (i.e., $N(0,\infty)$), the conditional posterior over δ and $\boldsymbol{b}$ is

$$
\prod_{d=1}^{D} \prod_{i=1}^{A} \frac{\exp\left\{\log\big(\mathrm{I}(\|\boldsymbol{u}_{id}\|_1 > 0)\big) + \sum_{j\neq i}(\delta + \lambda_{idj})u_{idj}\right\}}{Z(\delta, \boldsymbol{\lambda}_{id})}, \quad (17)
$$

where the two variables share the conditional posterior and thus can be jointly sampled. Likewise, assuming uninformative priors on $\boldsymbol{\eta}$ (i.e., $N(0,\infty)$) and σ_τ^2 (i.e., half-Cauchy(∞)), the conditional posterior is

$$
\prod_{d=1}^{D} \Big(\varphi_\tau(\tau_{a_d d}; \mu_{a_d d}, \sigma_\tau^2) \times \prod_{i\neq a_d} \big(1 - \Phi_\tau(\tau_{a_d d}; \mu_{id}, \sigma_\tau^2)\big) \Big). \quad (18)
$$

Although the IPTM is a highly complex model with a lot of latent variables, it yields an efficient inference algorithm by taking advantage of the two main parts of the likelihood repeatedly appear in the sampling equations—one from the latent recipients (Section 2.3.1) and another from the latent timestamps (Section 2.3.2). In addition, for better performance and interpretability of the topics we infer, we adopt the hyperparameter optimization technique for α and $\boldsymbol{m}$ called "new fixed-point iterations using the Digamma recurrence relation" in (Wallach, 2008), for every outer iteration.

4. Applications to Email Networks

The IPTM is intended for any network with timestamped, text-valued ties, however, in our application of the model we focus on the analysis of email network—a canonical example of dynamic textual communication. Via this application, we demonstrate that the IPTM is effective at predicting and explaining continuous-time textual communications.

4.1. Data

Our data come from the North Carolina county government email dataset collected by (ben Aaron et al., 2017) that includes internal email corpora covering the inboxes and outboxes of managerial-level employees of North Carolina county governments. Out of over twenty counties, we chose Dare County to 1) see whether and how communication networks surrounding a notable national emergency—Hurricane Sandy—differed from those surrounding other governmental functions, and 2) limit the scope of this initial application. The Dare County email network contains 2,247 emails, sent and received by 27 department managers over a period of 3 months (September–November) in 2012.

To verify that our model is applicable beyond the Dare County email network, we also performed two validation experiments using the Enron data set (Klimt & Yang, 2004).

We took a subset of the original data such that we only include emails between actors who sent over 300 emails, and actors who received over 300 emails from the chosen senders. Emails that were not sent to at least one other active actor were discarded, which resulted in a total of 6,613 emails involving 30 actors.

4.2. Dynamic Network Features

A set of network features $\boldsymbol{x}$ in recipient generating process can be flexibly tailored for various dataset (Section 2.3.1). For Dare County email network and the Enron data set, we employed a suite of eight different effects (Perry & Wolfe, 2013)—outdegree, indegree, send, receive, 2-send, 2-receive, sibling, and cosibling—to capture common network properties such as popularity, centrality, reciprocity, and transitivity. 2 illustrates the definition of each dynamic network statistic.

We assumed that each network feature has potentially different effects within a number of time intervals (i.e., recency effect), and partitioned the interval $[-\infty, t_d)$ into 4 sub-intervals with equal length in the log-scale. Because the Dare County email network only spans 12 weeks in length, we disregarded the time interval before 16 days and focused on three time intervals prior to t_d: 96–384 hours, 24–96 hours, and 0–24 hours. For the Enron data set, we simply changed the time unit from hour to day. We then computed each of the network feature within each time interval such that each $\boldsymbol{x}_{idjc}$ consists of 24 different network statistics. We provide detailed formulations and interpretations of the network statistics in the supplementary material.

4.3. Timestamp Specifications

To effectively model the timestamp of documents, we used another set of covariates $\boldsymbol{y}_{idc}$. We first included author-specific intercepts to account for individual differences in document-sending behavior. Because both of our data sets consist of organizational emails, we added two temporal features which possibly affect "when to send": an indicator of weekends/weekdays and an indicator of AM/PM when the document $d-1$ was sent.

Our preliminary analysis revealed that the Dare County
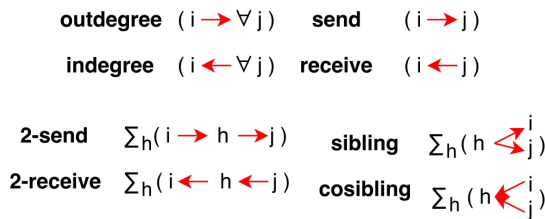

Figure 2. Eight dynamic network statistics used for the Dare County email network and the Enron dataset.

email networks and the Enron data set showed the best fitting when we assume lognormal distribution on the observed time-increments—i.e., $\log(\tau_{a_d d}) \sim N(\mu_{a_d d}, \sigma_\tau^2)$—compared to Gamma or Weibull distributions. We also observed significant lack-of-fit for single parameter distribution (e.g., Exponential distribution) since it failed to capture the variance in time-increments. Therefore, we chose lognormal distribution.

5. Predictive Analysis

We conducted a set of posterior predictive experiments—1) out-of-sample tie predictions, 2) topic coherence, and 3) posterior predictive checks—to gauge the IPTM's predictive performance as compared to alternative modeling approaches.

5.1. Out-of-Sample Tie Predictions

We evaluated the IPTM's ability to predict ties in textual communications from either the Dare County email network or the Enron data set, conditioned on the text of those emails and "training" part of the data. We separately formed a test split of each three components—author, recipients, and timestamps—by randomly selecting emails with probability 0.2 such that an email may observe all, some, or none of tie information. Any missing variables were imputed by drawing samples from their joint posterior distribution using an appropriately-modified version of the Metropolis-within-Gibbs algorithm in Section 3, and we ran inference to update the latent variables given the imputed data. We iterated imputation and inference multiple times to obtain enough samples of predicted values, and this procedure is provided in Algorithm 1.

Algorithm 1 Out-of-Sample Tie Predictions

Input: data $\{(a_d, \boldsymbol{r}_d, t_d, \boldsymbol{w}_d)\}_{d=1}^D$, sample size N
Separate test splits for $(a_d, \boldsymbol{r}_d, t_d)$ and hide the "test"
Initialize the parameters $(\boldsymbol{l}, \boldsymbol{z}, \boldsymbol{b}, \boldsymbol{\eta}, \delta, \boldsymbol{u})$.
for $n = 1$ **to** N **do**
    Impute "test" authors, recipients, and timestamps conditioned on the parameter values
    Note (to be deleted):
    $a_d \sim$ Multinom$(p_1, \ldots, p_A)$ where p_i is Eq. 18 for
    $a_d = i$ without product over d,
    $\boldsymbol{r}_d$ from Eq. 14, conditioned on a_d,
    t_d from Eq. 12—i.e., $\log(\tau_d) \sim N(\mu_{a_d d}, \sigma_\tau^2)$
    Update the parameters given the imputed data
end for
Summarize the predicted values after burn-in

We compared the IPTM's performance with that of baseline—the IPTM with $C = 1$—which unlinks the text and networks. Add "Why there is no comparable baseline
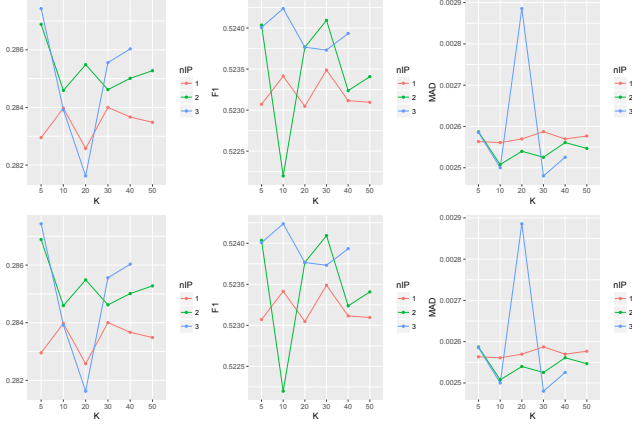
Figure 3. Average F1 score, AUC, MAE of out-of-sample tie predictions. *Top*: Dare County email network. *Bottom*: the Enron dataset.

outside this model?" We also varied the number of interaction patterns C from 1 to 3 and the number of topics K from 1 to 50 (Dare) or 100 (Enron) as a grid-search based hyperparameter selection process. For each combinations of C and K, predicted values of tie data were then compared to the true values to yield: F_1 scores for author predictions, multiclass version of the area under the ROC curve (AUC) measure (Hand & Till, 2001) for recipient predictions, and median absolute error (MAE) on timestamp predictions. We show the tie prediction results, averaged over test documents of each component, in Figure 3 (Plots to be updated). Although our model is intended for exploratory analysis, it achieves better link prediction performance than the baseline, validating our assumption that the IPTM acheives better predictive performance when topic-based contents are accounted to infer the parameters that govern the generation of tie data—authors, recipients, and timestamps.

5.2. Topic Coherence

Topic coherence metrics (Mimno et al., 2011) are often used to evaluate the semantic coherence in topic models. In order to demonstrate that the IPTM's incorporation of network features improves the ability of modeling text, we compared the coherence of topics inferred using our model with the coherence of topics inferred using LDA. Instead of re-fitting the data using standard LDA algorithms, we used the topic assignments from the IPTM with $C = 1$, which reduces the IPTM to LDA in terms of topic assignments. We varied the number of interaction patterns and the number of topics as in Section 5.1, and drew five samples from the joint posterior distribution over the latent variables. We evaluated the topics resulting from each sample and averaged over the five samples, where the results are shown in Figure 4. Combined with the findings in Section 5.1, this result demonstrates that the IPTM can achieve good predictive

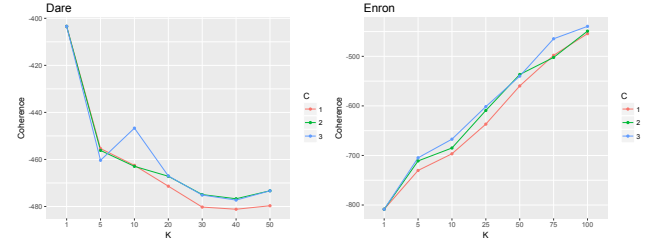performance while producing coherent topics.

Figure 4. Average topic coherence scores: (*left*) Dare County email network. (*right*) the Enron data set.

5.3. Posterior Predictive Checks

Finally, we performed posterior predictive checks (Rubin et al., 1984) to evaluate the appropriateness of the model specification for the Dare County email network. We formally generated entirely new data, by simulating ties and contents $\{(a_d, \boldsymbol{r}_d, t_d, \boldsymbol{w}_d)\}_{d=1}^D$ from the genenerative process in Section 2, conditional upon a set of inferred parameter values from the inference in Section 3. We specified the number of interaction patterns as $C = ?$ and the number of topics as $K = ?$, which yielded the best performance in Section 5.1. For the test of goodness-of-fit in terms of network dynamics, we defined multiple network statistics that summarize meaningful aspects of the Dare County email network: indegree distribution for author activities, outdegree distribution for recipient activities, recipient size distribution, document time-increments distribution, the edgewise shared partner distribution, and the geodesic distance distribution. We then generated 100 synthetic networks and texts from the posterior predictive distribution implied by the IPTM and Dare County email network. We applied each discrepancy function to each synthetic network to yield the distributions over the values of the six network statistics

As shown in Figure 5 (Plots to be updated), the IPTM shows "good fit" for the Dare County email network in that the observed data is not an outlier with respect to the distributions of new data drawn from the posterior predictive distribution. The IPTM generated synthetic networks with indegree distribution, outdegree distribution, recipient size, document time-increments, and edgewise shared partners that are very similar to those of the Dare County email network, showing that the model captures some important work features of the data including spreadness and transitivity.

6. Exploratory Analysis

Our model is primarily intended as an exploratory analysis tool for time-stamped textual communication. Our main goal in this exploratory analysis was to test three hypotheses: 1) personal or social topics (if any) would exhibit strong
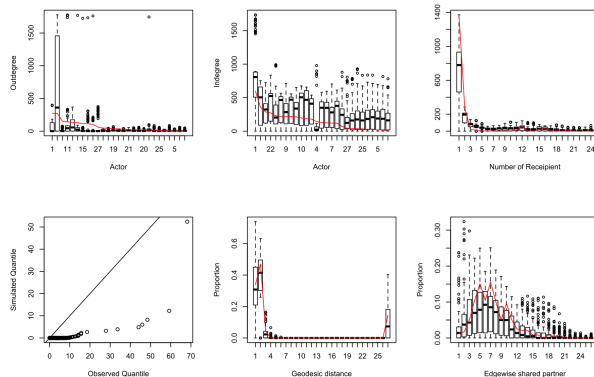
Figure 5. Posterior predictive checks for the Dare County email network: (a) outdegree, (b) indegree, (c) recipient size, (d) QQplot of time-increments, (e) geodesic distance, and (f) edgewise shared partners.

reciprocity and transitivity in tie formation, 2) topics about dissemination of information would be characterized by a lack of reciprocity, and 3) topics about Hurricane Sandy would exhibit a very different interaction pattern from the normal day-to-day conversations.

6.1. Topic Assignments

6.2. Interaction Pattern Coefficients

7. Summary

The IPTM is, to our knowledge, the first model to be capable of jointly modeling the author, recipients, timestamps and contents in time stamped text-valued networks. The IPTM incorporates innovative components, including the modeling of multicast tie formation and the conditioning of ERGM style network generative features on topic-based content. The application to North Carolina county government email data demonstrates, among other capabilities, the effectiveness at the IPTM in separating out both the content and relational structure underlying the normal day-to-day function of an organization and the management of a highly time-sensitive event—Hurricane Sandy. Finally, although we presented the IPTM in the context of email networks, the IPTM is applicable to a variety of networks in which ties are attributed with textual documents. These include, for example, economic sanctions sent between countries and legislation attributed with sponsors and co-sponsors.

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