

# Homework 9

● Graded

## Student

NATHAN LEUNG

## Total Points

20 / 20 pts

### Question 1

1

5 / 5 pts

- 0 pts Correct

- 1 pt Minor issue(s)

- 3 pts Major issue(s)

- 5 pts Completely incorrect

### Question 2

2

5 / 5 pts

- 0 pts Correct

- 1 pt Minor issue(s)

- 3 pts Major issue(s)

- 5 pts Completely incorrect

### Question 3

3

5 / 5 pts

- 0 pts Correct

- 1 pt Minor issue(s)

- 3 pts Major issue(s)

- 5 pts Completely incorrect

### Question 4

4

5 / 5 pts

- 0 pts Correct

- 1 pt Minor issue(s)

- 3 pts Major issue(s)

- 5 pts Completely incorrect



Question assigned to the following page: [1](#)

NATHAN LEUNG  
MATH 170A  
8 MARCH 2023

# HOMEWORK 9

$X_1, X_2 \sim \text{Uniform}(0, 1)$  and independent

$\alpha \in (0, 1)$  fixed

$$Y(\omega) = \begin{cases} X_1(\omega) & \text{if } X_1(\omega) > \alpha \\ X_2(\omega) & \text{otherwise} \end{cases}$$

a CDF of  $Y$

$$F_Y(y) = P(Y \leq y)$$

~~scratches~~

$$= P(Y \leq y \cap X_1 > \alpha) + P(Y \leq y \cap X_1 \leq \alpha)$$

$$= P(X_1 \leq y \cap X_1 > \alpha) + P(X_2 \leq y \cap X_1 \leq \alpha)$$

~~def of  $Y$~~

~~scratches~~

independence

$$= P(\alpha < X_1 \leq y) + P(X_2 \leq y) P(X_1 \leq \alpha)$$

~~scratches~~

$$= \begin{cases} 0 & \text{if } y \leq 0 \\ \alpha y & \text{if } 0 < y < \alpha \\ (\alpha - \alpha^2) + \alpha y & \text{if } \alpha \leq y < 1 \\ 1 & \text{if } y \geq 1 \end{cases}$$

$\alpha \in (0, 1)$

by def. of uniform distribution on  $[0, 1]$

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# Homework 9

1a We can differentiate to get the PDF  
cont.

$$F_Y(y) = \begin{cases} 0 & \text{if } y \leq 0 \\ \alpha y & \text{if } 0 < y \leq \alpha \\ (\alpha - \alpha)y + \alpha & \text{if } \alpha \leq y < 1 \\ 1 & \text{if } y \geq 1 \end{cases}$$

$$f_Y(y) = \frac{d}{dy} F_Y(y) = \begin{cases} 0 & \text{if } y \leq 0 \\ \alpha & \text{if } 0 < y \leq \alpha \\ 1-\alpha & \text{if } \alpha \leq y < 1 \\ 0 & \text{if } y \geq 1 \end{cases}$$

$$1b E[Y] = \int_{-\infty}^{\infty} x p_Y(x) dx$$

$$= \int_0^1 x p_Y(x) dx \quad P_Y(\text{FO}) \text{ when } y \notin (0, 1)$$

$$= \int_0^\alpha x \alpha dx + \int_\alpha^1 x (1-\alpha) dx$$

$$= \alpha \int_0^\alpha x dx + (1-\alpha) \int_\alpha^1 x dx$$

$$= \alpha \left[ \frac{x^2}{2} \right]_{x=0}^{x=\alpha} + (1-\alpha) \left[ \frac{x^2}{2} \right]_{x=\alpha}^{x=1}$$

$$= \alpha \left[ \frac{\alpha^2}{2} \right] + (1-\alpha) \left[ \frac{1}{2} - \frac{\alpha^2}{2} \right]$$

$$= \frac{\alpha^3}{2} + (1-\alpha) \left[ \frac{1 - \alpha^2}{2} \right]$$

$$= \frac{\alpha^3}{2} + \frac{(1+\alpha)(1-\alpha^2)}{2}$$

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# Homework

1b  
cont.

$$\begin{aligned} E[Y] &= \frac{\alpha^3}{2} + \frac{-\alpha^2 + \alpha - \alpha^3}{2} \\ &= \boxed{\frac{1 + \alpha - \alpha^2}{2}} \end{aligned}$$

1c  $E[Y] = \frac{1 + \alpha - \alpha^2}{2}$

To maximize, we take the derivative.

$$f(\alpha) = \frac{1 + \alpha - \alpha^2}{2}$$

$$\begin{aligned} f'(\alpha) &= \frac{d}{d\alpha} \left( \frac{1 + \alpha - \alpha^2}{2} \right) \\ &= \frac{1}{2} \frac{d}{d\alpha} (1 + \alpha - \alpha^2) \\ &= \frac{1}{2} (1 - 2\alpha) \end{aligned}$$

Then we solve for points where  $\frac{d}{d\alpha}$  is 0:

$$\frac{1}{2} (1 - 2\alpha) = 0$$

$$1 - 2\alpha = 0$$

$$1 = 2\alpha$$

$$\alpha = \frac{1}{2} \quad \text{a critical point}$$

$$f''(\alpha) = \frac{d}{d\alpha} \frac{1}{2} (1 - 2\alpha)$$

$$= \frac{1}{2} (-2) = -1$$

So by the 2nd derivative test,  $\alpha = \frac{1}{2}$

is a local max.

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# HOMENDEK 9

LC Since  $\alpha \in (0, 1)$  and

cont

$$f(0) = \frac{1}{2} \quad f(1) = \frac{1}{2}$$

$$f\left(\frac{1}{2}\right) = \frac{1 + \frac{1}{2} - \frac{1}{4}}{2} = \frac{\frac{5}{4}}{2} = \frac{5}{8}$$

$$\frac{5}{8} > \frac{1}{2}$$

The point  $\boxed{\alpha = \frac{1}{2}}$  maximizes  $E[Y]$ .

$$\text{Id } Z = \max(X_1, X_2)$$

$$F_Z(z) = P(Z \leq z)$$

$$= P(X_1 \leq z \cap X_2 \leq z) \quad \text{def of max}$$

$$= P(X_1 \leq z) P(X_2 \leq z) \quad \text{independent}$$

$$= \begin{cases} 0 & \text{if } x \leq 0 \\ z^2 & \text{if } 0 < x < 1 \\ 1 & \text{if } x \geq 1 \end{cases} \quad \text{def of uniform}$$

$$f_Z(z) = \begin{cases} 0 & \text{if } x \leq 0 \\ 2z & \text{if } 0 < x < 1 \\ 0 & \text{if } x \geq 1 \end{cases} \quad \text{take derivative}$$

$$E[Z] = \int_{-\infty}^{\infty} z f_Z(z) dz = \int_0^1 z \cdot 2z dz = \int_0^1 2z^2 dz$$

$$= 2 \left[ \frac{z^3}{3} \right]_0^1 = 2 \left( \frac{1}{3} \right) = \frac{2}{3}$$

L - J 2 - 0

C

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# HOMEWORK

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$$2 \quad X = \sqrt{-2 \ln(U)} \cos(2\pi V) \quad U, V \text{ uniform } (0,1)$$

$$Y = \sqrt{-2 \ln(U)} \sin(2\pi V) \quad \text{independent}$$

Define ~~the region~~ the region

$$A(x, y) = \{(s, t) \in \mathbb{R}^2 : s \leq x \text{ and } t \leq y\}.$$

Then we can write

$$F_{X,Y}(x, y) = P(X \leq x \text{ and } Y \leq y)$$

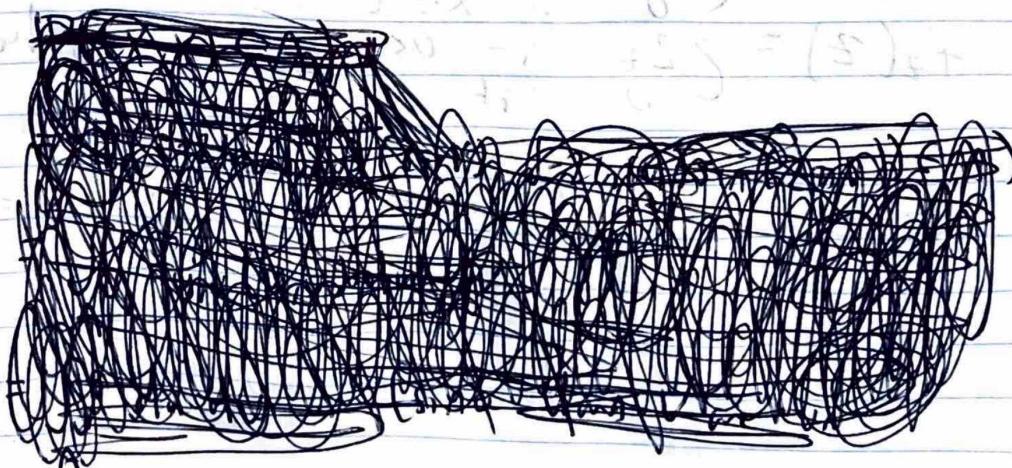
$$= P((X, Y) \in A(x, y))$$

Also, define the function ~~the function~~

$$g_A(x, y) = \begin{cases} 1 & \text{if } (x, y) \in A \\ 0 & \text{otherwise} \end{cases}$$

any region A

Since  $X = \sqrt{-2 \ln(U)} \cos(2\pi V)$  and  $Y = \sqrt{-2 \ln(U)} \sin(2\pi V)$ , we can write (for some region A)



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# HOMEWORK 9

2.  $P((X, Y) \in A)$

cont.

$$= \iint_{\mathbb{R}^2} g_A(\sqrt{-2\ln(u)} \cos(2\pi v), \sqrt{-2\ln(u)} \sin(2\pi v)) f_{u,v}(u,v) du dv$$

$\mathbb{R}^2$

$$= \iint_{[0,1]^2} g_A(\sqrt{-2\ln(u)} \cos(2\pi v), \sqrt{-2\ln(u)} \sin(2\pi v)) du dv$$

$[0,1]^2$

since  
 $U, V \sim$   
Uniform(0,1)

We proceed by changing variables. Let

$$r = \sqrt{-2\ln(u)} \quad \theta = 2\pi v$$

We have the change-of-variables transform

$$T(u, v) = (\sqrt{-2\ln(u)}, 2\pi v)$$

and Jacobian

$$J_T(r, \theta) = \begin{vmatrix} \frac{\partial u}{\partial r} & \frac{\partial u}{\partial \theta} \\ \frac{\partial v}{\partial r} & \frac{\partial v}{\partial \theta} \end{vmatrix}$$

Solving for the inverse transform, we have

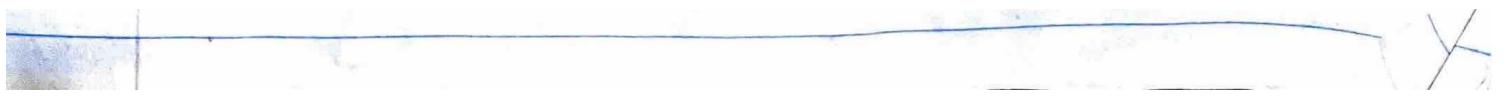
$$r = \sqrt{-2\ln(u)}$$

$$r^2 = -2\ln(u)$$

$r \geq 0$  always since we take principal  $\sqrt{\phantom{x}}$

$$-\frac{1}{2}r^2 = \ln(u)$$

$$u = e^{-\frac{1}{2}r^2}$$



X

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# Homework

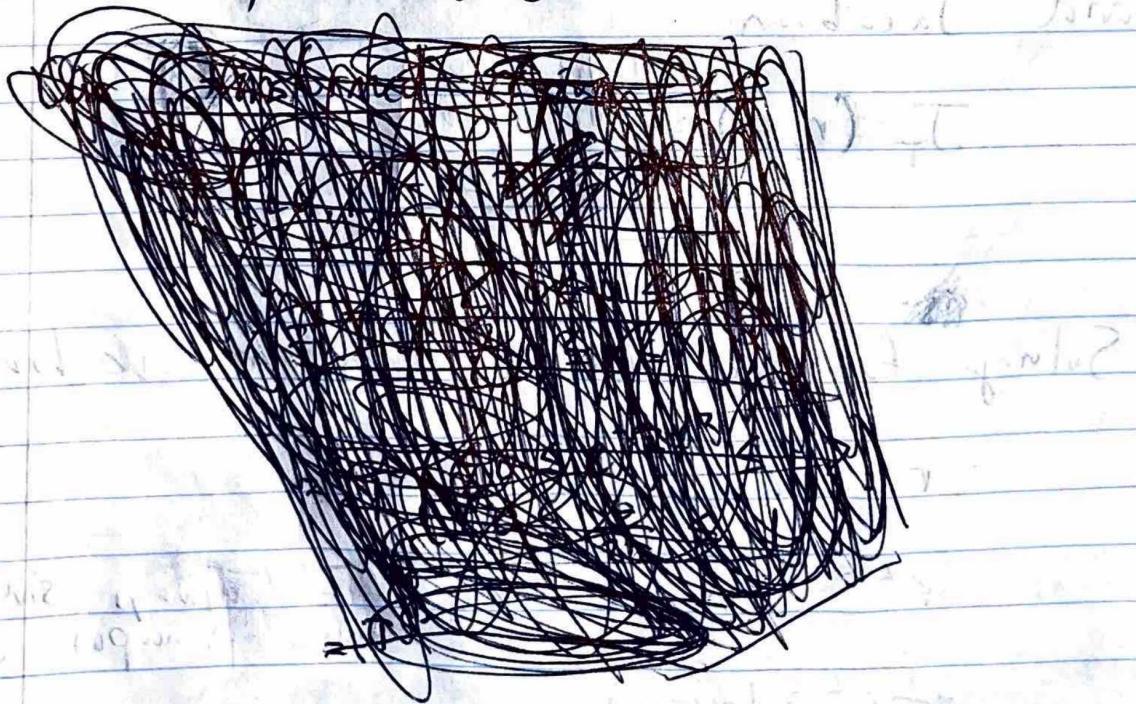
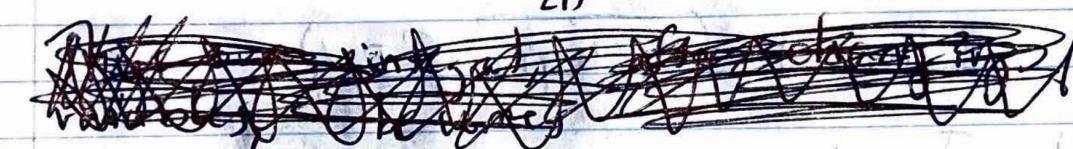
2.  $\theta = 2\pi V$ , so  $V = \frac{\theta}{2\pi}$ .

Thus,  $\frac{dV}{dr} = -r e^{-\frac{1}{2}r^2} \frac{du}{d\theta} = 0$

$$\frac{dV}{dr} = 0 \quad \frac{dV}{d\theta} = \frac{1}{2\pi}$$

and  $J(r, \theta) = \begin{vmatrix} -re^{-\frac{1}{2}r^2} & 0 \\ 0 & \frac{1}{2\pi} \end{vmatrix}$

$$= -\frac{1}{2\pi} r e^{-\frac{1}{2}r^2}$$



The transformed region is



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# The NETWORK

2.  $\{(u, v) : 0 \leq u \leq 1 \text{ and } 0 \leq v \leq 1\}$

$$= \{(r, \theta) : 0 \leq e^{-\frac{1}{2}r^2} \leq 1 \text{ and } 0 \leq \frac{\theta}{2\pi} \leq 1\}$$

$$= \{(r, \theta) : -\infty \leq -\frac{1}{2}r^2 \leq 0 \text{ and } 0 \leq \theta \leq 2\pi\}$$

In is  
order  
preserving

$$= \{(r, \theta) : -\infty \leq r^2 \leq 0 \text{ and } 0 \leq \theta \leq 2\pi\}$$

$$= \{(r, \theta) : -\infty \leq r \leq \infty \text{ and } 0 \leq \theta \leq 2\pi\}$$

$$= \mathbb{R} \times [0, 2\pi]$$

squares  
are  
always  
positive,  
so inequality  
always  
holds

Thus, our integral after changing variables  
from  $(u, v)$  to  $(r, \theta)$  becomes

$$\iint_{\mathbb{R} \times [0, 2\pi]} g_A(r \cos \theta, r \sin \theta) \left| -\frac{1}{2\pi} r e^{-\frac{1}{2}r^2} \right| dr d\theta$$

$$= \iint_{\mathbb{R} \times [0, 2\pi]} g_A(r \cos \theta, r \sin \theta) \cdot \frac{1}{2\pi} e^{-\frac{1}{2}r^2} r dr d\theta$$

We change variable once more. Let

$$X = r \cos \theta \quad \text{and} \quad Y = r \sin \theta$$

$$\text{so } r = \sqrt{x^2 + y^2} \quad \text{and} \quad \theta = \text{atan} 2(Y, X)$$

$\tan^{-1}(Y/X)$  (two-argument arctangent)

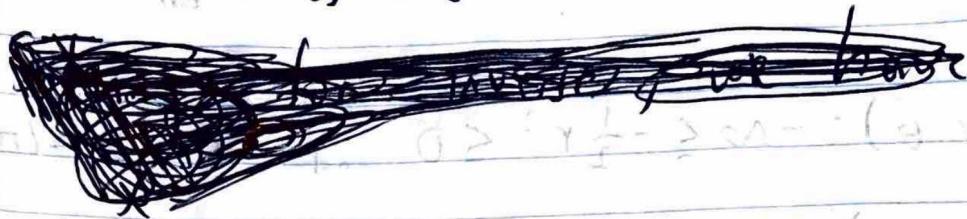
$$2 \begin{cases} \tan^{-1}\left(\frac{y}{x}\right) & \text{if } x \geq 0 \\ \tan^{-1}\left(\frac{y}{x}\right) + \pi & \text{if } x < 0 \end{cases}$$

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# Homework

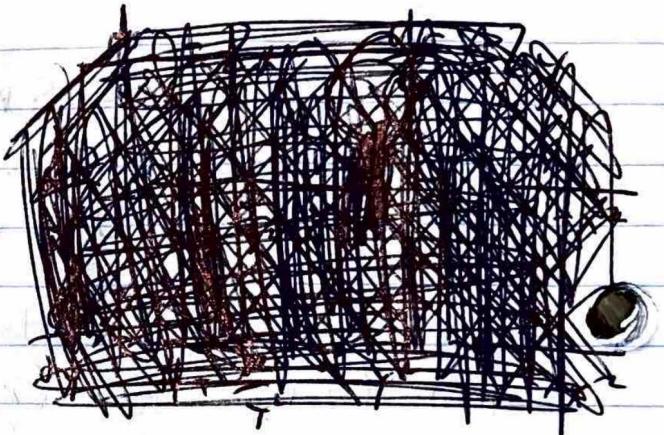
2 cont. This gives us a change of variables transform

$$S((r, \theta)) = (r\cos\theta, r\sin\theta)$$



with Jacobian

$$J_S = \begin{vmatrix} \frac{\partial r}{\partial x} & \frac{\partial r}{\partial y} \\ \frac{\partial \theta}{\partial x} & \frac{\partial \theta}{\partial y} \end{vmatrix}$$



$$\frac{\partial r}{\partial x} = \frac{1}{2} (x^2 + y^2)^{-\frac{1}{2}} \cdot 2x = \frac{x}{\sqrt{x^2 + y^2}}$$

$$\frac{\partial r}{\partial y} = \frac{1}{2} (x^2 + y^2)^{-\frac{1}{2}} \cdot 2y = \frac{y}{\sqrt{x^2 + y^2}}$$

$$\frac{\partial \theta}{\partial x} = \frac{1}{1 + \frac{x^2}{y^2}} \cdot \frac{1}{y} = \frac{1}{y^2 + \frac{x^2}{y^2}} = \frac{y}{y^2 + x^2}$$

$$\frac{\partial \theta}{\partial y} = \frac{1}{1 + \frac{x^2}{y^2}} \cdot \frac{-x}{y^2} = \frac{-x}{y^2 + x^2}$$

(Aug 2009) - 08:11 (x) 1st C

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# HOMEWORK

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2  
Ques.

$$\begin{aligned}
 J_s &= \begin{vmatrix} \frac{x}{\sqrt{x^2+y^2}} & \frac{y}{\sqrt{x^2+y^2}} \\ \frac{y}{x^2+y^2} & \frac{-x}{x^2+y^2} \end{vmatrix} \\
 &= \frac{-x^2}{(x^2+y^2)\sqrt{x^2+y^2}} - \frac{y^2}{(x^2+y^2)\sqrt{x^2+y^2}} = \frac{-(x^2+y^2)}{(x^2+y^2)\sqrt{x^2+y^2}} \\
 &= \frac{-1}{\sqrt{x^2+y^2}}
 \end{aligned}$$

~~Region~~ The transformed region is

$$\{(r, \theta) : r \in \mathbb{R}, 0 \leq \theta \leq 2\pi\}$$

$$=\{(x, y) : \sqrt{x^2+y^2} \in \mathbb{R}, 0 \leq \operatorname{atan}^2(y/x) \leq 2\pi\}$$

$\supseteq \mathbb{R}^2$  since both conditions are always true  
and  $\operatorname{atan}^2$  is defined on all  $\mathbb{R}^2$

Thus our integral becomes

$$\iint_{\mathbb{R}^2} g_A(x, y) \cdot \frac{1}{2\pi} e^{-\frac{1}{2}(x^2+y^2)} \sqrt{x^2+y^2} \left| \frac{-1}{\sqrt{x^2+y^2}} \right| dx dy$$

$$= \iint_{\mathbb{R}^2} g_A(x, y) \cdot \frac{1}{2\pi} e^{-\frac{1}{2}(x^2+y^2)} dx dy$$

~~100% 100% 100% 100% 100% 100% 100% 100% 100% 100%~~

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# Homework 9

2. That is, we have the equality  
cont.

$$P(X, Y) \in A$$

$$= \iint_{\mathbb{R}^2} g_A(x, y) \cdot \frac{1}{2\pi} e^{-\frac{1}{2}(x^2+y^2)} dx dy$$

Note that  $F_{X,Y}(x, y)$

$$= P(X \leq x \text{ and } Y \leq y)$$

$$= P((X, Y) \in A(x, y)) \quad \text{as defined earlier}$$

$$= \iint_{\mathbb{R}^2} g_{A(x,y)}(s, t) \cdot \frac{1}{2\pi} e^{-\frac{1}{2}(s^2+t^2)} ds dt$$

$$= \int_{-\infty}^x \int_{-\infty}^y \frac{1}{2\pi} e^{-\frac{1}{2}(s^2+t^2)} ds dt \quad \begin{matrix} \text{by definition} \\ \text{of } A(x, y) \end{matrix}$$

We want to show that  $X$  and  $Y$  are both  $N(0, 1)$  distributed and independent.

We have

$$F_x(x) = P(X \leq x) = \iint_{-\infty}^{x, \infty} \frac{1}{2\pi} e^{-\frac{1}{2}(s^2+t^2)} ds dt \quad \begin{matrix} \text{compute marginal} \\ \text{for } X \end{matrix}$$

$$= \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}s^2} \left( \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}t^2} dt \right) ds$$

$\infty$   $\infty$   
(S and so we can are symmetric about  
 $-\infty$  switch)

Question assigned to the following page: [2](#)

HöMELDungen 9

$$2. F_x(x) = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}s^2} (1) ds$$

$$F \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}s^2} ds$$

$$\text{So } f_x(x) = \frac{\partial}{\partial x} F_x(x) = \frac{\partial}{\partial x} \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}s^2} ds \\ = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2}$$

which is exactly the PDF for a  $N(0, 1)$  distributed random variable. Hence  $X \sim N(0, 1)$ .  
A symmetric argument shows  $Y \sim N(0, 1)$ .

Finally, to show independence, note that

$$F_{X,Y}(x,y) = \int_{-\infty}^x \int_{-\infty}^y \frac{1}{2\pi} e^{-\frac{1}{2}(s^2 + t^2)} ds dt$$

So their joint density function is

$$\begin{aligned} \frac{\partial}{\partial x \partial y} F_{X,Y}(x,y) &= \frac{1}{2\pi} e^{-\frac{1}{2}(s^2 + t^2)} \\ &= \left( \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2} \right) \left( \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}y^2} \right) \end{aligned}$$

Since the joint density function can be expressed as a product of a function of  $x$  and a function of  $y$ ,  $X$  and  $Y$  must

be independent.

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# Homework P9

3  $X \sim X_v^2$ ,  $(Y \sim X_k^2)$  independent

WTS  $X+Y \sim X_{k+v}^2$

We let  $Z = X+Y$  and proceed using the convolution formula.

$$f_Z(z) = \int_{-\infty}^{\infty} f_X(x) f_Y(z-x) dx$$

$$= \int_0^{\infty} f_X(x) f_Y(z-x) dx$$

since pdf of  $X_k^2$   
random variable is 0  
for negative values

Note it

$$z \leq 0, \rightarrow = \int_0^z f_X(x) f_Y(z-x) dx$$

thus  $f_Z(z) = 0$

Assume  $z > 0$

when  $x > z$ ,  
 $z-x < 0$  so  
 $f_Y(z-x) = 0$

$$= \int_0^z \left( \frac{1}{2^{v/2} \Gamma(v/2)} x^{\frac{v}{2}-1} e^{-\frac{x}{2}} \right) \left( \frac{1}{2^{k/2} \Gamma(k/2)} (z-x)^{\frac{k}{2}-1} e^{-\frac{(z-x)}{2}} \right) dx$$

$$= \frac{1}{2^{(k+v)/2} \Gamma(\frac{v}{2}) \Gamma(\frac{k}{2})} \int_0^z x^{\frac{v}{2}-1} (z-x)^{\frac{k}{2}-1} e^{-\frac{z}{2}} dx$$

(let  $A = \frac{1}{2^{(k+v)/2} \Gamma(\frac{v}{2}) \Gamma(\frac{k}{2})}$ ). Then we can

write

$$= Ae^{-\frac{z}{2}} \int_0^z x^{\frac{v}{2}-1} (z-x)^{\frac{k}{2}-1} dx$$

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# HOMEWORK

3 We make (the) substitution  
cont.

$$u = \frac{x}{z} \quad x = uz$$

$$du = \frac{1}{z} dx \quad \text{Note } z \neq 0 \text{ means this}$$

$$zdu = dx \quad \text{substitution is ok}$$

and we get

$$f_z(z) = Ae^{-z^{1/2}} \int_0^{\frac{v}{2}-1} (uz) (z - uz) \cdot z du$$

$$= Ae^{-z^{1/2}} \int_0^{\frac{v}{2}-1} z^{\frac{v}{2}-1} u^{\frac{v}{2}-1} z^{\frac{k}{2}-1} (1-u)^{\frac{k}{2}-1} z du$$

$$= Ae^{-z^{1/2}} z^{\frac{v+k}{2}-1} \int_0^{\frac{v}{2}-1} u^{\frac{v}{2}-1} (1-u)^{\frac{k}{2}-1} du$$

$$\text{let } C = \int_0^{\frac{v}{2}-1} u^{\frac{v}{2}-1} (1-u)^{\frac{k}{2}-1} du$$

$$\text{Then we have } f_z(z) = Ae^{-z^{1/2}} z^{\frac{v+k}{2}-1} \cdot C$$

Let  $W \sim \text{Gamma}\left(\frac{v+k}{2}, \frac{1}{2}\right) = \chi_{\frac{v+k}{2}}^2$ . By the definition of a Gamma/chi-squared distribution,

$$f_w(w) = \begin{cases} \frac{1}{\Gamma(\frac{v+k}{2}) 2^{\frac{v+k}{2}/2}} w^{\frac{v+k}{2}-1} e^{-\frac{1}{2}w} & \text{if } w > 0 \\ 0 & \text{otherwise} \end{cases}$$

We can write  $f_z$  in terms of  $f_w$  like so:

~~for  $\left(\frac{1}{2}, \frac{1}{2}\right)$~~

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# Homework

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276 from before

3

$$\text{cont } f_2(z) = AC f_w(z) \cdot \Gamma\left(\frac{v+k}{2}\right) 2^{\frac{(v+k)}{2}}$$

Since  $f_2(z)$  is a PDF,

$$\int_{-\infty}^{\infty} f_2(z) dz = 1$$

That is,

$$\int_{-\infty}^{\infty} AC f_w(z) \Gamma\left(\frac{v+k}{2}\right) \cdot 2^{\frac{(v+k)}{2}} dz = 1$$

$$AC \cdot \Gamma\left(\frac{v+k}{2}\right) \cdot 2^{\frac{(v+k)}{2}} \underbrace{\int_{-\infty}^{\infty} f_w(z) dz}_{\text{also 1}} = 1$$

$$AC \cdot \Gamma\left(\frac{v+k}{2}\right) \cdot 2^{\frac{(v+k)}{2}} = 1$$

$$\frac{\Gamma\left(\frac{v+k}{2}\right) \cdot 2^{\frac{(v+k)}{2}}}{\sum_{n=0}^{k-1} \Gamma\left(\frac{v}{2}\right) \Gamma\left(\frac{k}{2}\right)} C = 1$$

$$C = \frac{\Gamma\left(\frac{v}{2}\right) \Gamma\left(\frac{k}{2}\right)}{\Gamma\left(\frac{v+k}{2}\right)}$$

$$\int_0^1 u^{\frac{v}{2}-1} (1-u)^{\frac{k}{2}-1} du = \frac{\Gamma(\frac{v}{2}) \Gamma(\frac{k}{2})}{\Gamma(\frac{v+k}{2})}$$

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# HOMEWORK 9

3 And  
cont.

$$f_z(z) = A C f_w(z) \Gamma\left(\frac{v+k}{2}\right) z^{\frac{(v+k)}{2}}$$

$$= \frac{\Gamma\left(\frac{v+k}{2}\right) z^{\frac{(v+k)}{2}}}{2^{\frac{(k+v)}{2}} \Gamma\left(\frac{v}{2}\right) \Gamma\left(\frac{k}{2}\right)} \cdot \frac{\Gamma\left(\frac{v}{2}\right) \Gamma\left(\frac{k}{2}\right)}{\Gamma\left(\frac{v+k}{2}\right)} f_w(z)$$

$$= f_w(z)$$

i.e.  $z$  and  $w$  have the same PDF  
and hence are distributed identically.

Thus,  $\boxed{z \sim \chi_{v+k}^2}$ .

WILLIAM M. TAYLOR

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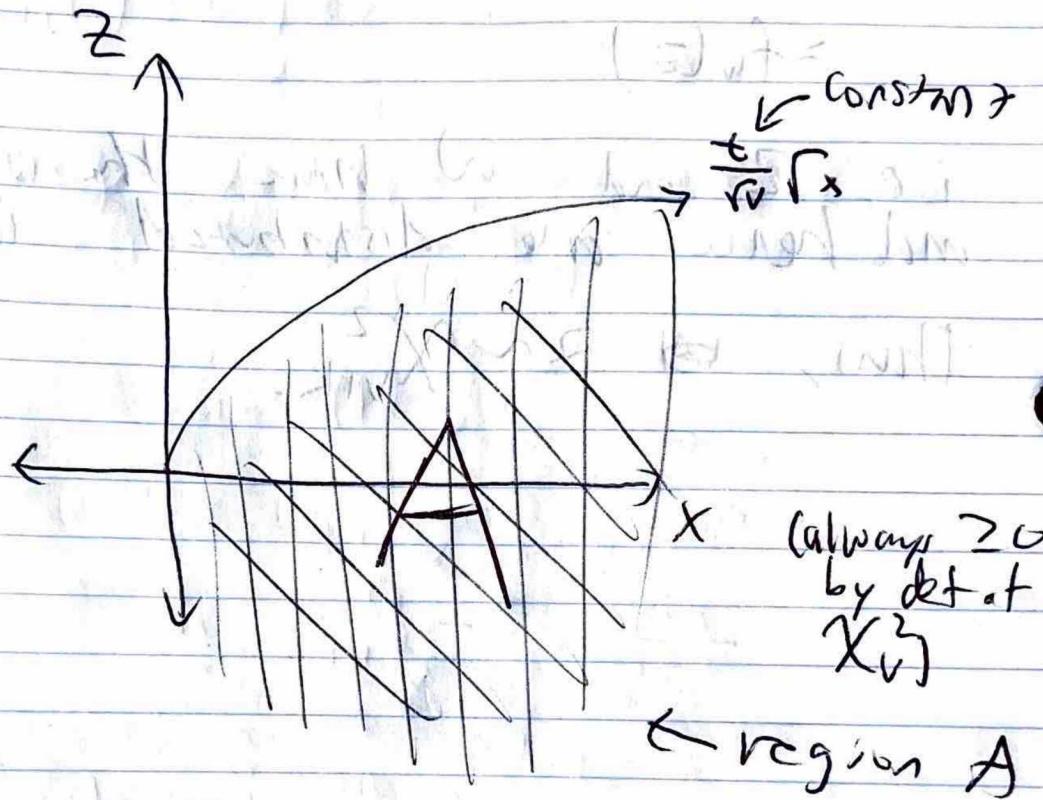
# HOMEWORK 9

4  $X \sim \chi^2_1$      $Z \sim N(0, 1)$ , independent

$$T = Z \sqrt{\frac{X}{\chi}}$$

$$P(T \leq t) = P(Z \sqrt{\frac{X}{\chi}} \leq t)$$

$$= P(Z \leq t \sqrt{\frac{X}{\chi}}) = P(Z \leq \frac{t}{\sqrt{\chi}} \sqrt{X})$$



$$P(T \leq t) \Leftrightarrow P(Z \leq \frac{t}{\sqrt{\chi}} \sqrt{X})$$

$$\Leftrightarrow P((Z, X) \in A)$$

$$P((Z, X) \in A) = \iint_{\substack{Z \leq \frac{t}{\sqrt{\chi}} \sqrt{X} \\ A}} f_Z(z) f_X(x) dz dx \quad \text{since independent}$$

$$= \int_0^\infty \int f_z(z) f_x(x) dz dx$$

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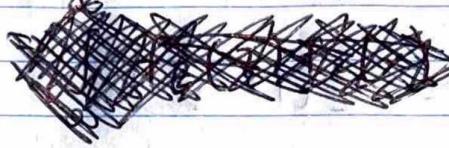
# HOMEWORK 9

4 We can differentiate to find the PDF  
 (cont)

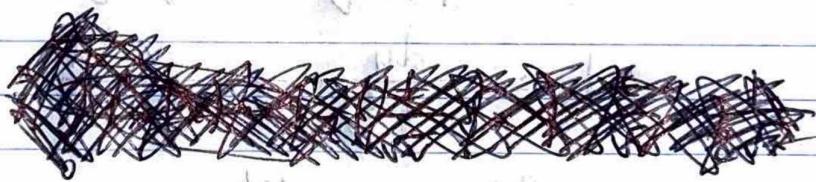
$$f_T(t) = \frac{d}{dt} P(T \leq t)$$

$$= \frac{d}{dt} \int_0^{\infty} \int_{-\infty}^{\frac{t}{\sqrt{x}}} f_Z(z) f_X(x) dz dx$$

$$= \int_0^{\infty} \frac{d}{dt} \int_{-\infty}^{\frac{t}{\sqrt{x}}} f_Z(z) f_X(x) dz dx$$
Leibniz integral rule



$$= \int_0^{\infty} f_X(x) \left( \frac{d}{dt} \int_{-\infty}^{\frac{t}{\sqrt{x}}} f_Z(z) dz \right) dx$$



$$= \int_0^{\infty} f_X(x) \cdot f_Z\left(\frac{t}{\sqrt{x}}\right) \cdot \frac{1}{\sqrt{x}} dx$$

Lichniz  
integral rule/  
Chain rule

$$= \int_0^{\infty} \left( \frac{1}{2^{v/2} \Gamma(v/2)} x^{-\frac{v}{2}-1} e^{-\lambda x} \right) \left( \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(\frac{t}{\sqrt{x}})^2} \right) \left( \frac{\sqrt{x}}{\sqrt{v}} \right) dx$$

$$= \frac{1}{2^{(v+1)/2} \sqrt{\pi v} \Gamma(v/2)} \int_0^{\infty} x^{\frac{v+1}{2}-1} e^{-\frac{x}{2} - \frac{1}{2}(\frac{t}{\sqrt{x}})^2} dx$$



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# HOMEWORK 9

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cont.  $f_T(t) = \frac{1}{2^{(v+1)/2} \Gamma(v/2) \sqrt{\pi v}} \int_0^{\infty} x^{\frac{v+1}{2}-1} e^{-\frac{x}{2}(1+\frac{t^2}{v})} dx$

Let  $C = \frac{1}{2^{(v+1)/2} \Gamma(\frac{v}{2}) \sqrt{\pi v}}$

Then  $f_T(t) = C \int_0^{\infty} x^{\frac{v+1}{2}-1} e^{-\frac{x}{2}(1+\frac{t^2}{v})} dx$

We change variables:

$$\frac{u}{2} = -\frac{x}{2}(1 + \frac{t^2}{v})$$

$$u = x(1 + \frac{t^2}{v})$$

$$x^2 = \frac{u}{1 + \frac{t^2}{v}} = \frac{uv}{v+t^2}$$

$$dx = \frac{du}{1 + \frac{t^2}{v}} = \frac{v du}{v+t^2}$$

$$f_T(t) = C \int_0^{\infty} \left( \frac{uv}{v+t^2} \right)^{\frac{v+1}{2}-1} e^{-\frac{u}{2}} \frac{v du}{v+t^2}$$

$$= C \int_0^{\infty} u^{\frac{v+1}{2}-1} \left( \frac{v}{v+t^2} \right)^{\frac{v+1}{2}-1} e^{-\frac{u}{2}} \left( \frac{v}{v+t^2} \right) du$$

$$= C \int_0^{\infty} u^{\frac{v+1}{2}-1} \left( \frac{v}{v+t^2} \right)^{\frac{v+1}{2}-1} e^{-\frac{u}{2}} du.$$

$$\begin{aligned} & \left( \frac{x}{\sqrt{v+t^2}} \right)^{\frac{1}{2}} - \frac{1}{2} - \frac{1}{2} \int_0^{\infty} u^{\frac{v+1}{2}} e^{-\frac{u}{2}} du \\ &= C \left( \frac{x}{\sqrt{v+t^2}} \right)^{\frac{v+1}{2}} \int_0^{\infty} u^{\frac{v+1}{2}} e^{-\frac{u}{2}} du \end{aligned}$$

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# Homework 9

4. Let  $W \sim \chi_{v+1}^2$ . Then

$$f_W(w) = \begin{cases} \frac{w^{(v+1)/2}}{2^{(v+1)/2} \Gamma(\frac{v+1}{2})} e^{-w/2} & w \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

And so on  $[0, \infty)$ ,  $f_W(w) \cdot 2^{(v+1)/2} \Gamma(\frac{v+1}{2}) = w^{(v+1)/2-1} e^{-w/2}$ . Substituting this into the integral, we get

$$f_T(t) = C \left( \frac{v}{v+t^2} \right)^{\frac{v+1}{2}} \int_0^\infty f_W(u) \cdot 2^{(v+1)/2} \Gamma(\frac{v+1}{2}) du$$

$$= C \left( \frac{v}{v+t^2} \right)^{\frac{v+1}{2}} 2^{(v+1)/2} \Gamma(\frac{v+1}{2}) \int_0^\infty f_W(u) du$$

$= 1$  since pdf

$$= \frac{2^{(v+1)/2} \Gamma(\frac{v+1}{2})}{2^{(v+1)/2} \Gamma(\frac{v}{2}) \sqrt{\pi v}} \left( \frac{v}{v+t^2} \right)^{\frac{v+1}{2}}$$

$$= \frac{\Gamma(\frac{v+1}{2})}{\Gamma(\frac{v}{2}) \sqrt{\pi v}} \left( \frac{v}{v+t^2} \right)^{\frac{v+1}{2}}$$

$$= \frac{\Gamma(\frac{v+1}{2})}{\Gamma(\frac{v}{2}) \sqrt{\pi v}} \left( \frac{v}{v+t^2} \right)^{-\left(\frac{v+1}{2}\right)}$$

$$= \frac{\Gamma(\frac{v+1}{2})}{\Gamma(\frac{v}{2}) \sqrt{\pi v}} \left( 1 + \frac{t^2}{v} \right)^{-\left(\frac{v+1}{2}\right)}$$

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# Homework 9

4. Since  $f_T(t)$  is a PDF, we have

$$1 = \int_{-\infty}^{\infty} f_T(t) dt = \int_{-\infty}^{\infty} \frac{\Gamma(\frac{v+1}{2})}{\Gamma(\frac{v}{2})\sqrt{\pi v}} (1 + \frac{t^2}{v})^{-\frac{v+1}{2}} dt$$

$$\approx \frac{\Gamma(\frac{v+1}{2})}{\Gamma(\frac{v}{2})\sqrt{\pi v}} \int_{-\infty}^{\infty} (1 + \frac{t^2}{v})^{-\frac{v+1}{2}} dt$$

And so

$$\int_{-\infty}^{\infty} (1 + \frac{t^2}{v})^{-\frac{v+1}{2}} dt = \frac{\Gamma(\frac{v}{2})\sqrt{\pi v}}{\Gamma(\frac{v+1}{2})}$$

$$4 \cdot \left(\frac{1}{\sqrt{e}} + 1\right) = 4 \cdot \frac{1}{\sqrt{e}} + 4 =$$

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