

# Linear Regression (part 2)

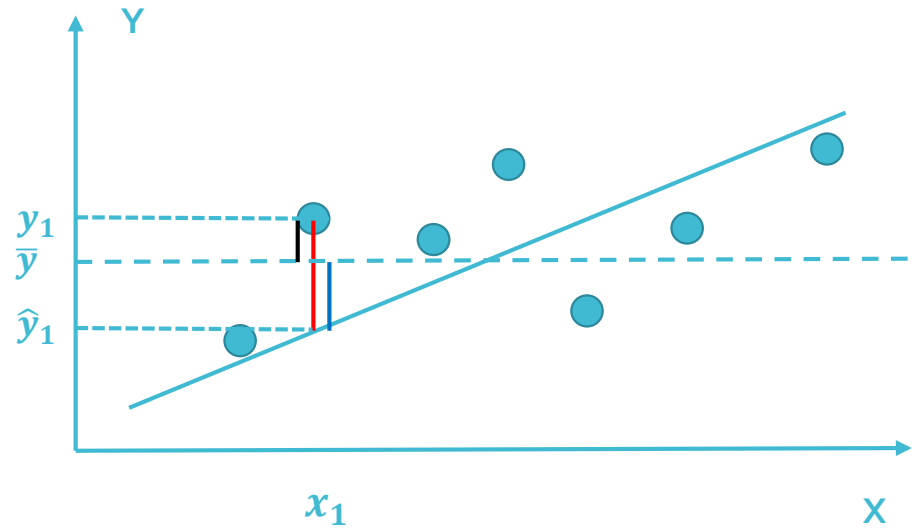
DSA 6000: Data Science and Analytics, Fall 2019

Wayne State University

# Assessing the Accuracy of the Model

- Residual Sum of Squares (RSS):  $\sum_{i=1}^n (y_i - \hat{f}(X_i))^2$ , measures the amount of variability in  $Y$  that is left unexplained after performing the regression.
- Residual Standard Error (RSE) is an estimate of the standard deviation of  $\epsilon$ .
  - $RSE = \sqrt{\left(\frac{1}{n-p-1}\right) RSS}$
  - Note that RSE depends on  $p$ , so adding a useless predictor to the model increases  $\left(\frac{1}{n-p-1}\right)$ , overall RSE might also increase
- RSE represents the average amount that the response will deviate from the true regression line. It is a measure of the *lack of fit* of the model to the data, in the units of  $Y$ .
- $R^2$  statistic: the proportion of variance in  $Y$  that is explained by the model.
  - $R^2 = \frac{TSS - RSS}{TSS}$ , where  $TSS = \sum (y_i - \bar{y})^2$  is the total sum of squares.
  - Adding an extra predictor will always **increase**  $R^2$
  - Adjusted  $R^2$  accounts for the model complexity

# Decompose the TSS



- Suppose the linear model is fit by the OLS method, then
- Total variation in Y is decomposed into two parts:
  - Variation explained by the model
  - Variation left unexplained by the model
- *Total SS = Explained SS + Residual SS*
- $\sum_{i=1}^n (y_i - \bar{y}_i)^2 = \sum_{i=1}^n (\bar{y}_i - \hat{y}_i)^2 + \sum_{i=1}^n (y_i - \hat{y}_i)^2$

# About the F-statistic

- F-statistic is used for testing whether at least one of the predictors has a significant effect on the response variable.
- $H_0: \beta_1 = \beta_2 = \dots = \beta_p = 0$
- $H_1$ : at least one  $\beta_j$  is non-zero
- A large F-statistic value will lead to rejection of  $H_0$ . The rejection threshold depends on both  $n$  and  $p$ .
- Why use F test since we already have the t test?
- For a model with many predictors (i.e., large  $p$ ), it can happen that the p-value for some individual predictor(s) is small (e.g.,  $< 0.05$ ), but the model as a whole fails the F test (i.e., fail to reject  $H_0$ ).
  - For instance, if there are 100 variables, all unrelated to  $Y$ , the p-values for about 5% of the variables will be below 0.05 **by chance**. We would expect to see about 5 small p-values even in the absence of any true association between the predictors and the response.
  - F-statistic is immune to this type of fallacy.

# Interpreting LR outputs

```
Call:
lm(formula = sales ~ TV + radio, data = ad, subset = trainIndex)

Residuals:
    Min       1Q   Median       3Q      Max
-8.8720 -0.8629  0.2989  1.1603  2.9526

Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)  2.839160   0.373493   7.602 4.17e-12 ***
TV           0.045219   0.001721  26.275 < 2e-16 ***
radio        0.191949   0.010121  18.965 < 2e-16 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.767 on 137 degrees of freedom
Multiple R-squared:  0.8881,    Adjusted R-squared:  0.8865
F-statistic: 543.8 on 2 and 137 DF,  p-value: < 2.2e-16
```

Everything else held equal, one more unit of ad expense on TV will on average increase sales by 0.045 unit.

# Model with Interaction Effects

```
> summary(lm(sales~ TV + radio, data= ad))
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )	
(Intercept)	2.92110	0.29449	9.919	<2e-16	***
TV	0.04575	0.00139	32.909	<2e-16	***
radio	0.18799	0.00804	23.382	<2e-16	***

- **Model:**  $\text{Sales} = 2.921 + 0.046 \cdot \text{TV} + 0.188 \cdot \text{radio} + \epsilon$
- **Interpretation:** e.g., regardless of the spending on TV advertisement, holding it at the same level, a unit change in radio advertisement will cause 0.188 unit of change in sales in the same direction

```
> summary(lm(sales~ TV + radio + TV:radio, data= ad))
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )	
(Intercept)	6.750e+00	2.479e-01	27.233	<2e-16	***
TV	1.910e-02	1.504e-03	12.699	<2e-16	***
radio	2.886e-02	8.905e-03	3.241	0.0014	**
TV:radio	1.086e-03	5.242e-05	20.727	<2e-16	***

- **Model:**  $\text{Sales} = 6.75 + 0.019 \cdot \text{TV} + 0.029 \cdot \text{radio} + 0.001 \cdot \text{TV} \cdot \text{radio} + \epsilon$
- **Interpretation:** The effect of radio advertising on sales now depends on the amount of the TV advertising.
- Positive interaction effect is called **synergy**, negative called **friction**.

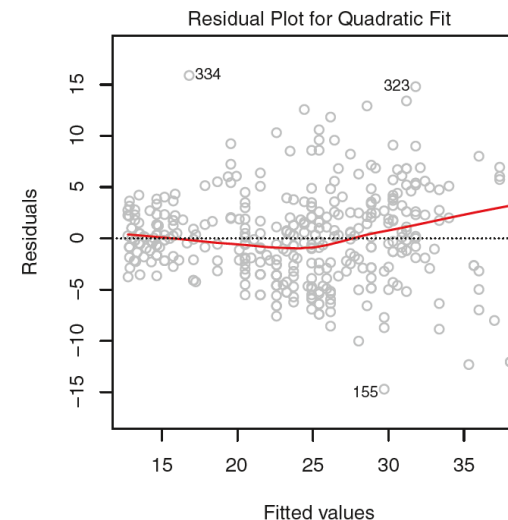
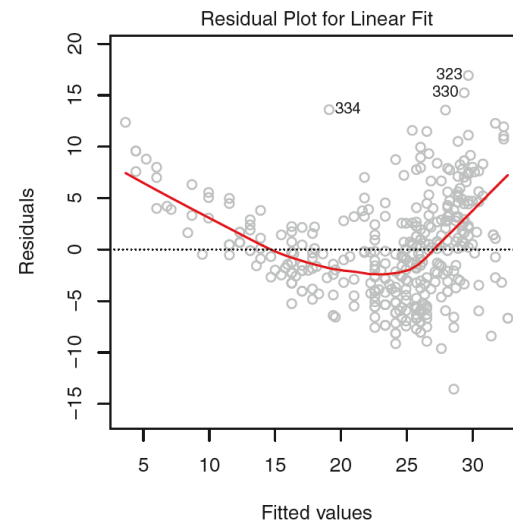
# Nonlinear relationships

```
> summary(lm(mpg ~ poly(horsepower, 2), data = Auto))
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )	
(Intercept)	23.4459	0.2209	106.13	<2e-16	***
poly(horsepower, 2)1	-120.1377	4.3739	-27.47	<2e-16	***
poly(horsepower, 2)2	44.0895	4.3739	10.08	<2e-16	***

- **Model:**  $\text{mpg} = 23.5 - 120 \cdot \text{horsepower} + 44 \cdot \text{horsepower}^2 + \epsilon$
- It models the nonlinear relationship between mpg and horsepower
- Coefficients are estimated by OLS, so it is still a linear fit.
- What makes you try including a higher order term of horsepower in the first place? The diagnostic plots.



# Deciding on Important Variables

- Variable Selection is studied extensively in Chapter 6.
- Three classical approaches:
  - **Forward selection:** start with null model, add a single variable at a time whose addition results in the lowest RSS among all possible single-variable additions.
  - **Backward selection:** start with full model, remove the variable with the largest p-value, refit and repeat, until all p-values are small enough.
  - **Mixed selection:** start with null model, add a single variable at a time as in Forward selection, if some variable in the updated model has a large p-value, remove it; repeat until all variables in the model have a small p-value and all remaining variables would have a high p-value if included in the model.



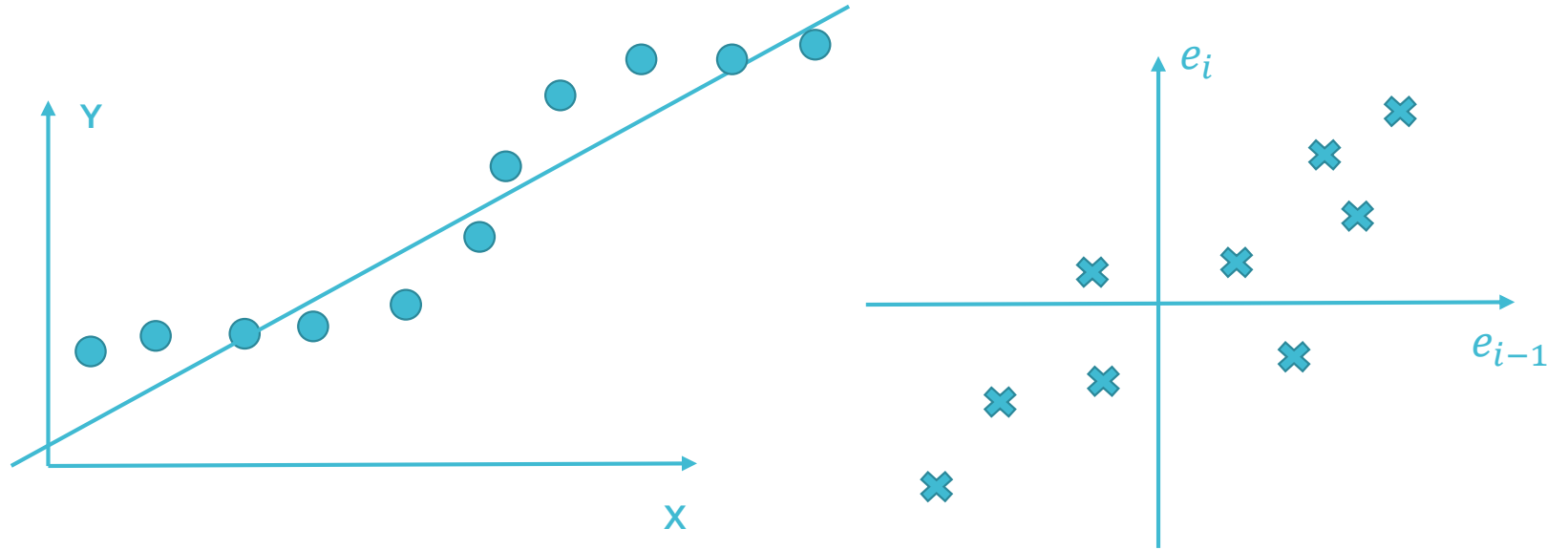
# Potential Problems in a Linear Model

- Linear regression makes strong assumptions about data
- Real data rarely conform to all those assumptions
- Potential problems:
  - Nonlinearity of the response-predictor relationship
  - Correlation of error terms
  - Non-constant variance of the error terms
  - Outliers
  - High-leverage points
  - Collinearity

# Nonlinearity

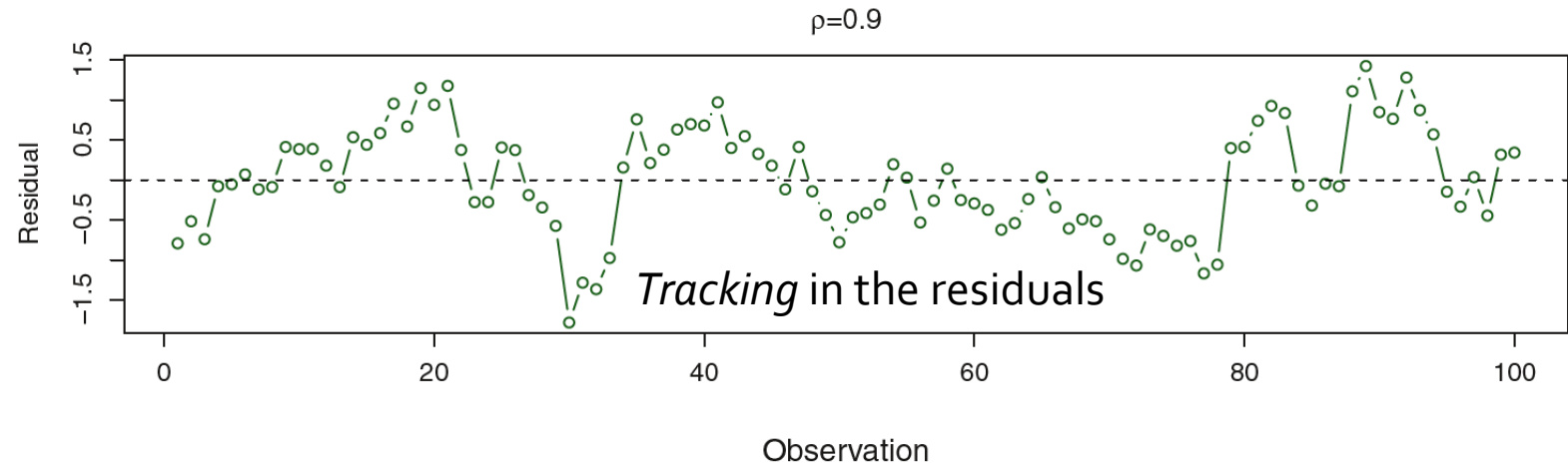
- The relationship between Y and X is nonlinear, but we fit a linear model
- Can be revealed by diagnostic plots (e.g., the residual plot)
- Solutions:
  - Apply nonlinear transformation on predictors, or include higher-order terms of predictors, and fit a linear model using the transformed variables (how the coefficients are interpreted will change)
  - Use nonlinear models (Chapter 7)

# Correlation of error terms



- Autocorrelation: statistical dependence of errors on preceding errors
- Can be tested using Durbin-Watson test
- Positive autocorrelation can make  $SE(\hat{\beta}_j)$  an underestimate of  $\sigma_{\beta_j}$ , and can over-estimate  $R^2$

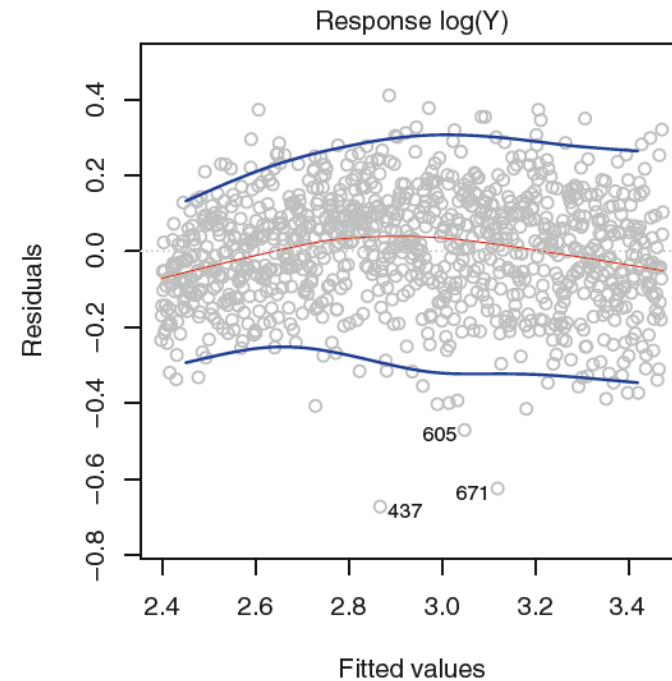
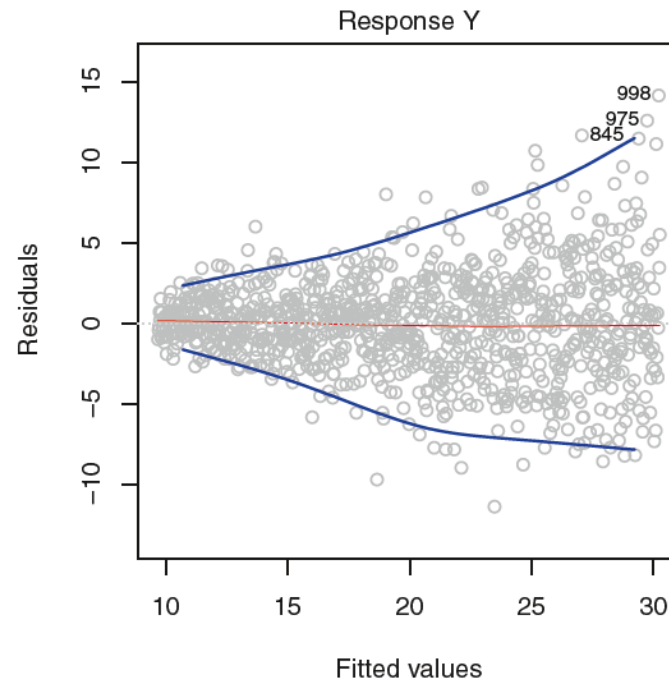
# Correlation of error terms



- Occurs most often in data taken from observations over time
- Can occur outside the time series data.
- May be due to undiscovered nonlinearity or to missing variables
- Good experimental design can mitigate the risk of such correlations

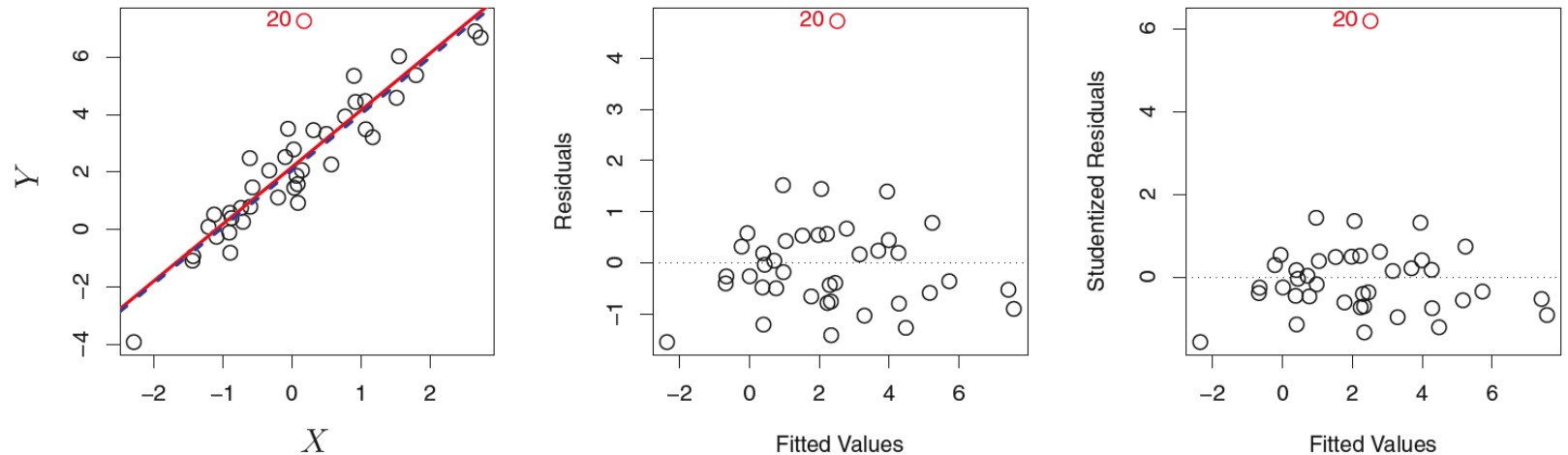
# Non-constant variance of the error terms (Heteroscedasticity)

- Violates the assumption:  $Var(\epsilon_i) = \sigma^2$
- Funnel shape in the residual plot
- If each  $X_i$  is an average of  $n_i$  raw data points, can use Weighted Least Squares (using  $n_i$  as the weight for obs  $i$ ) to mitigate
- Nonlinear transform, e.g.,  $\log(Y)$ ,  $\sqrt{Y}$  can also mitigate, but introduce nonlinearity



# Outliers

- An outlier is a point for which  $y_i$  is far from the value predicted by the model.
- An outlier may not drastically affect the coefficient estimate, but it can hurt the model's explanatory power, e.g., the R-squares.



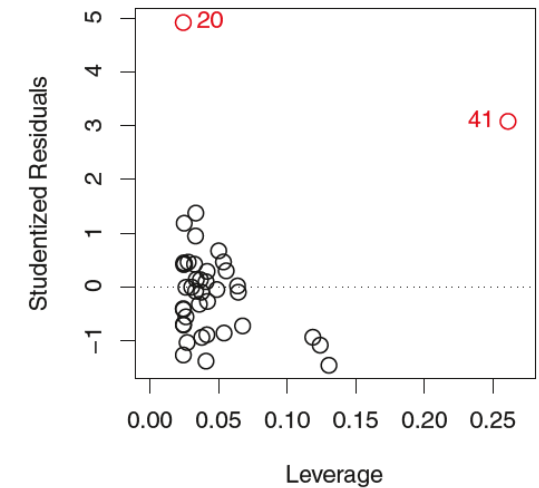
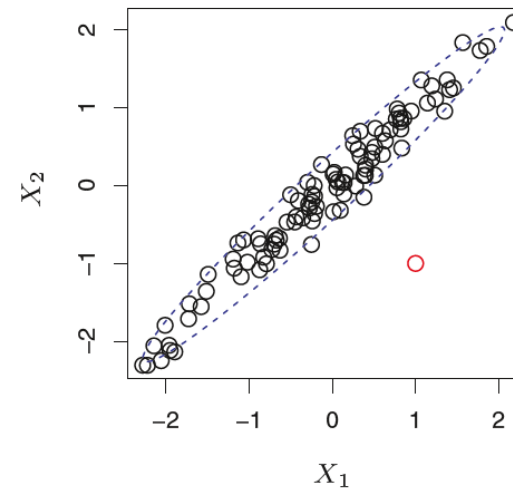
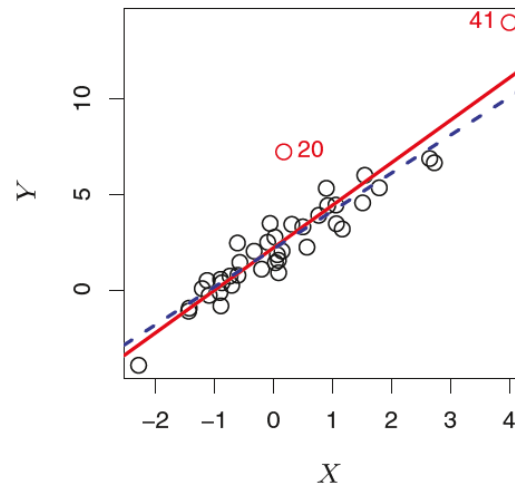
**Left:** Exclusion of the outlier obs 20 caused little change in the fit (red solid line v.s. black dashed line)

**Middle:** Residual plot can reveal an outlier

**Right:** We would expect all studentized residuals to fall between -3 to 3. Any residual outside the normal range can be regarded as an outlier

# High-leverage points

- A high-leverage point has an unusual independent variable value
- Easy to identify in simple linear regression, hard to identify graphically in multiple linear regression
- We can use the *leverage statistic* to quantify the leverage of each point
- The average leverage for all the observations is always equal to  $(p + 1)/n$
- So if an observation's leverage greatly exceeds the average, it is a high-leverage point



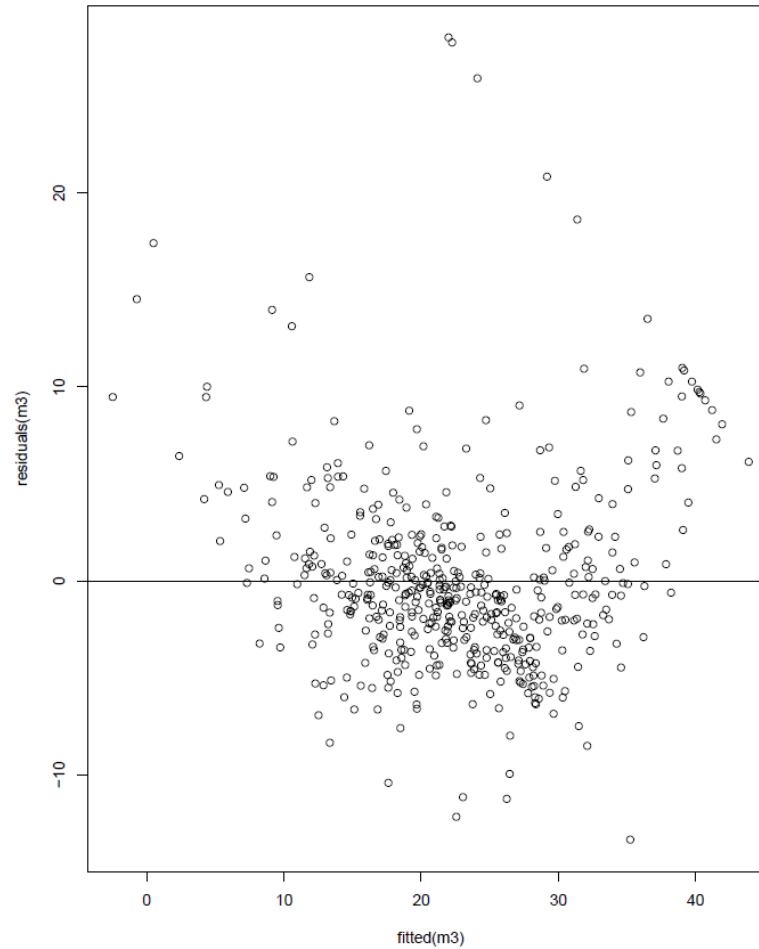
# Collinearity

- **Collinearity** refers to the situation in which two or more independent variables are closely related to one another.
- It reduces the accuracy of the coefficient estimate
  - A variable may have a significant effect on the response, but we may failed to recognize it (i.e., fail to reject  $H_0: \beta_j = 0$ ) due to its large standard error caused by collinearity
- Pairwise collinearity can be detected by pairwise scatter plots (pairs() function in R), but scatter plots cannot detect **multicollinearity**.
- Multicollinearity can be assessed by variance inflation factor (VIF).
  - VIF of a variable  $j$  is  $1/(1 - R_j^2)$ , where  $R_j^2$  is the R-square statistic obtained by fitting a linear model using  $X_j$  as the response and using all other predictors as independent variables.
- $VIF > 5$  indicates problematic multicollinearity due to presence of the variable. The variable should be removed.

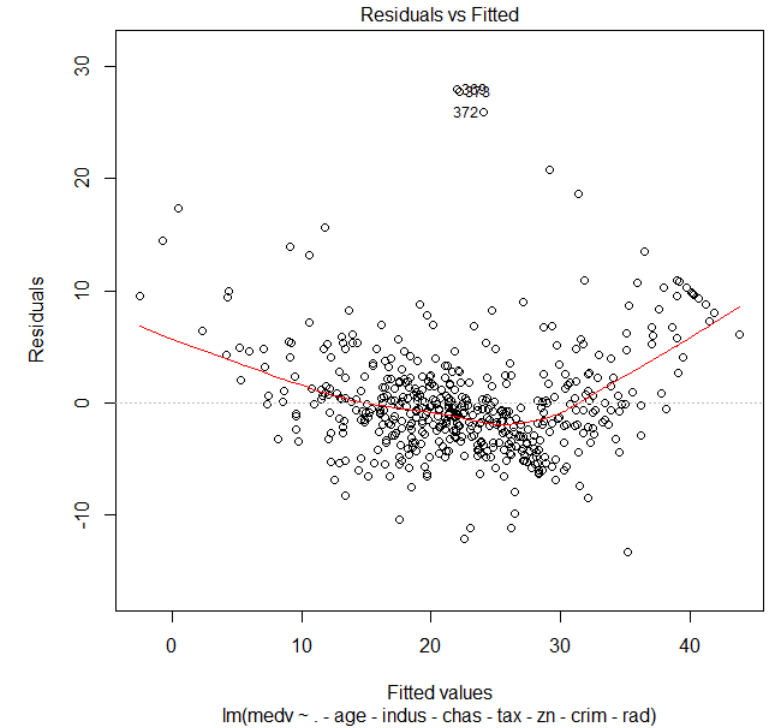


# Residual plot

```
> plot(fitted(m3), residuals(m3))  
> abline(h=0)  
> |
```



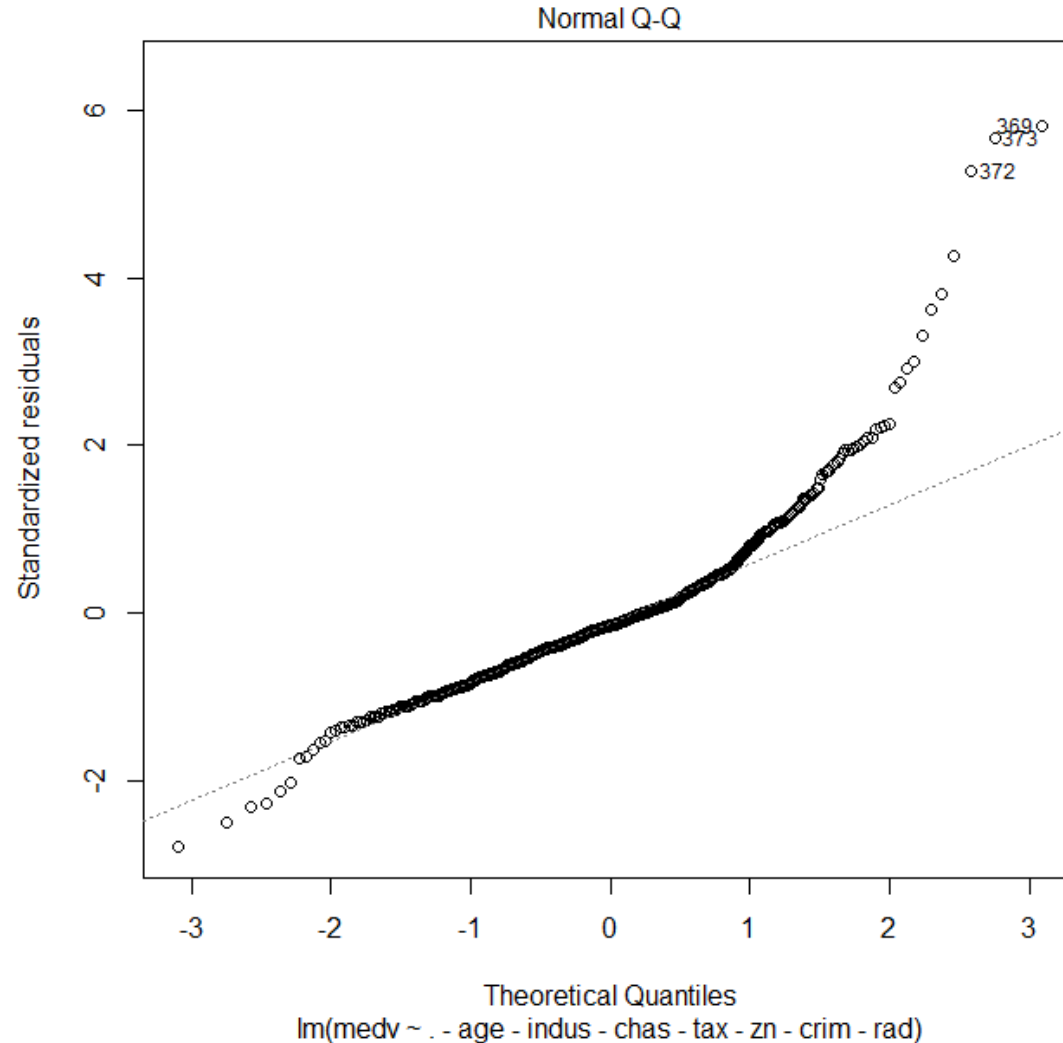
```
> plot(m3, which = 1)  
>
```



- The red curve is the LOWESS (LOcally WEighted Scatter-plot Smoother) curve.
- Can reveal nonlinearity, heteroscedasticity and outliers

# Quantile-quantile plot

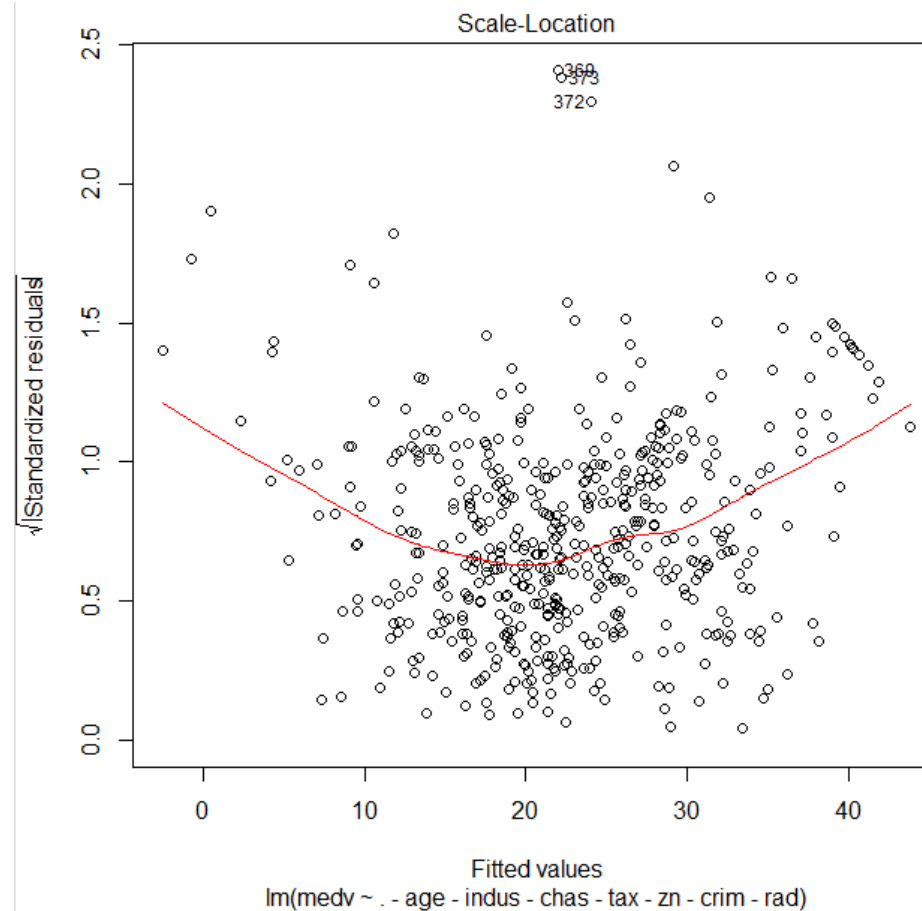
```
> plot(sort(rnorm(length(residuals(m3))), 0, 1)), sort(scale(residuals(m3))))  
> |
```



- Checks if the residuals are normally distributed.
- If yes, the points will lie along a straight line.
- Scattered off-the-line points are outliers
- Nonlinear shape indicates nonlinear relationship between response and predictors.

# Scale-location plot

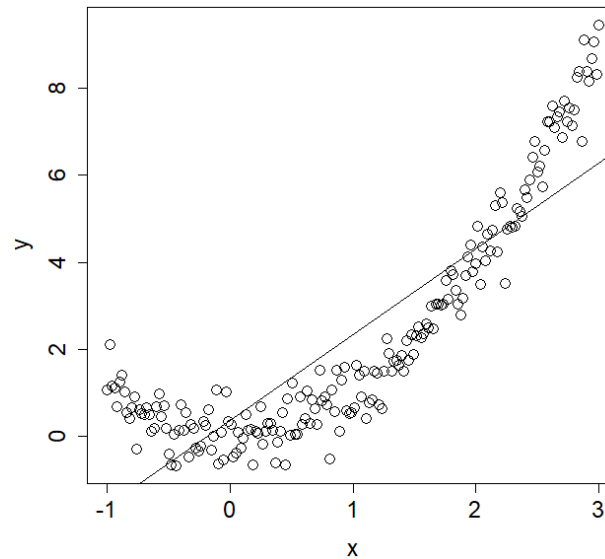
```
> plot(m3, which = 3)  
> |
```



- It plots the square root of the absolute value of the standardized residuals against the fitted values.
- The red line is the LOWESS curve fitted to these points.
- We want to see a level straight LOWESS curve.
  - Curved line indicates nonlinearity
  - Slanted line indicates heteroscedasticity

# Nonlinear relationship

```
> abline(lm(y~x))
> x = seq(-1,3,length.out = 200)
> y = x^2 + rnorm(200, 0, 0.5)
> plot(x, y, cex=1.5, cex.lab=1.5, cex.axis=1.5)
> abline(lm(y~x))
> |
```



```
> plot(lm(y~x), cex=1.5, cex.lab=1.5, cex.axis=1.5)
Hit <Return> to see next plot:
Hit <Return> to see next plot:
Hit <Return> to see next plot:
Hit <Return> to see next plot:
> |
```

