

Maximum Likelihood

Consider IID random samples X_1, X_2, \dots, X_n where X_i is a sample from the density function $f(X_i|\theta)$. We are going to introduce a new way of choosing parameters called Maximum Likelihood Estimation (MLE). We want to select that parameters (θ) that make the observed data the most likely. *Note that we are now using notation that shows that the density of X depends on its parameters, θ .*

First we define the likelihood of our data give parameters θ :

$$L(\theta) = \prod_{i=1}^n f(X_i|\theta)$$

This is the probability of all of our data. It evaluates to a product because all X_i are independent. Now we chose the value of θ that maximizes the likelihood function. Formally $\hat{\theta} = \underset{\theta}{\operatorname{argmax}} L(\theta)$.

A cool property of argmax is that since \log is a monotone function, the argmax of a function is the same as the argmax of the \log of the function! That's nice because logs make the math simpler. Instead of using likelihood, you should instead use log likelihood: $LL(\theta)$.

$$LL(\theta) = \log \prod_{i=1}^n f(X_i|\theta) = \sum_{i=1}^n \log f(X_i|\theta)$$

To use a maximum likelihood estimator, first write the log likelihood of the data given your parameters. Then chose the value of parameters that maximize the log likelihood function. Argmax can be computed in many ways. Most require computing the first derivative of the function.

Bernoulli MLE Estimation

Consider IID random variables X_1, X_2, \dots, X_n where $X_i \sim \operatorname{Ber}(p)$. First we are going to write the PMF of a Bernoulli in a crazy way: The probability mass function $f(X_i|p) = p^{X_i}(1-p)^{1-X_i}$. Wow! Whats up with that? First convince yourself that when $X_i = 0$ and $X_i = 1$ this returns the right probabilities. We write the PMF this way because its derivable.

Now let's do some MLE estimation:

$$\begin{aligned} L(\theta) &= \prod_{i=1}^n p^{X_i}(1-p)^{1-X_i} \\ LL(\theta) &= \sum_{i=1}^n \log p^{X_i}(1-p)^{1-X_i} \\ &= \sum_{i=1}^n X_i(\log p) + (1-X_i)\log(1-p) \\ &= Y \log p + (n-Y)\log(1-p) \end{aligned} \quad \text{where } Y = \sum_{i=1}^n X_i$$

Great Scott! Now we simply need to chose the value of p that maximizes our log-likelihood. One way to do that is to find the first derivative and set it equal to 0.

$$\begin{aligned} \frac{\delta LL(p)}{\delta p} &= Y \frac{1}{p} + (n-Y) \frac{-1}{1-p} = 0 \\ \hat{p} &= \frac{Y}{n} = \frac{\sum_{i=1}^n X_i}{n} \end{aligned}$$

All that work and we get the same thing as method of moments and sample mean...

Normal MLE Estimation

Consider IID random variables X_1, X_2, \dots, X_n where $X_i \sim N(\mu, \sigma^2)$.

$$\begin{aligned} L(\theta) &= \prod_{i=1}^n f(X_i | \mu, \sigma^2) \\ &= \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(X_i - \mu)^2}{2\sigma^2}} \\ LL(\theta) &= \sum_{i=1}^n \log \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(X_i - \mu)^2}{2\sigma^2}} \\ &= \sum_{i=1}^n \left[-\log(\sqrt{2\pi\sigma}) - \frac{1}{2\sigma^2} (X_i - \mu)^2 \right] \end{aligned}$$

If we chose the values of $\hat{\mu}$ and $\hat{\sigma}^2$ that maximize likelihood, we get: $\hat{\mu} = \frac{1}{n} \sum_{i=1}^n X_i$ and $\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \hat{\mu})^2$.

Linear Transform Plus Noise

Assume that $Y = \theta X + Z$ where $Z \sim N(0, \sigma^2)$ and X is an unknown distribution. The equations imply that $Y|X \sim N(\theta X, \sigma^2)$. Chose a value of θ that maximizes the probability of the data: $(X_1, Y_1), (X_2, Y_2), \dots, (X_n, Y_n)$.

We approach this problem by finding a function for the log likelihood of the data given θ . Then we find the value of θ that maximizes the log likelihood function. To start, use the PDF of a Normal to express the probability of $Y|X, \theta$:

$$f(Y_i | X_i, \theta) = \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(Y_i - \theta X_i)^2}{2\sigma^2}}$$

Now we are ready to write the likelihood function, then take its log to get the log likelihood function:

$$\begin{aligned} L(\theta) &= \prod_{i=1}^n f(Y_i, X_i, \theta) && \text{Let's break up this joint} \\ &= \prod_{i=1}^n f(Y_i | X_i, \theta) f(X_i) && f(X_i) \text{ is independent of } \theta \\ &= \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(Y_i - \theta X_i)^2}{2\sigma^2}} f(X_i) && \text{Substitute in the definition of } f(Y_i | X_i) \end{aligned}$$

$$\begin{aligned} LL(\theta) &= \log L(\theta) \\ &= \log \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(Y_i - \theta X_i)^2}{2\sigma^2}} f(X_i) && \text{Substitute in } L(\theta) \\ &= \sum_{i=1}^n \log \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(Y_i - \theta X_i)^2}{2\sigma^2}} + \sum_{i=1}^n \log f(X_i) && \text{Log of a product is the sum of logs} \\ &= m \log \frac{1}{\sqrt{2\pi}} - \frac{1}{2\sigma^2} \sum_{i=1}^m (Y_i - \theta X_i)^2 + \sum_{i=1}^n \log f(X_i) \end{aligned}$$

Remove constant multipliers and terms that don't include θ . We are left with trying to find a value of θ that maximizes:

$$\begin{aligned} \hat{\theta} &= \underset{\theta}{\operatorname{argmax}} - \sum_{i=1}^m (Y_i - \theta X_i)^2 \\ &= \underset{\theta}{\operatorname{argmin}} \sum_{i=1}^m (Y_i - \theta X_i)^2 \end{aligned}$$

This result says that the value of θ that makes the data most likely is one that minimizes the squared error of predictions of Y . We will see in a few days that this is the basis for linear regression.