

# Assignment 2 (ML for TS) - MVA 2023/2024

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## 1 Introduction

**Objective.** The goal is to better understand the properties of AR and MA processes, and do signal denoising with sparse coding.

**Warning and advice.**

- Use code from the tutorials as well as from other sources. Do not code yourself well-known procedures (e.g. cross validation or k-means), use an existing implementation.
- The associated notebook contains some hints and several helper functions.
- Be concise. Answers are not expected to be longer than a few sentences (omitting calculations).

**Instructions.**

- Fill in your names and emails at the top of the document.
- Hand in your report (one per pair of students) by Tuesday 5<sup>th</sup> December 11:59 PM.
- Rename your report and notebook as follows:  
FirstnameLastname1\_FirstnameLastname1.pdf and  
FirstnameLastname2\_FirstnameLastname2.ipynb.  
For instance, LaurentOudre\_CharlesTruong.pdf.
- Upload your report (PDF file) and notebook (IPYNB file) using this link:  
[docs.google.com/forms/d/e/1FAIpQLSfCqMXSDU9jZJbYUMmeLCXbVeckZYNiDpPI4hRUwcJ2cBHQM](https://docs.google.com/forms/d/e/1FAIpQLSfCqMXSDU9jZJbYUMmeLCXbVeckZYNiDpPI4hRUwcJ2cBHQM)

## 2 General questions

A time series  $\{y_t\}_t$  is a single realisation of a random process  $\{Y_t\}_t$  defined on the probability space  $(\Omega, \mathcal{F}, P)$ , i.e.  $y_t = Y_t(w)$  for a given  $w \in \Omega$ . In classical statistics, several independent realisations are often needed to obtain a “good” estimate (meaning consistent) of the parameters of the process. However, thanks to a stationarity hypothesis and a “short-memory” hypothesis, it is still possible to make “good” estimates. The following question illustrates this fact.

### Question 1

An estimator  $\hat{\theta}_n$  is consistent if it converges in probability when the number  $n$  of samples grows to  $\infty$  to the true value  $\theta \in \mathbb{R}$  of a parameter, i.e.  $\hat{\theta}_n \xrightarrow{\mathcal{D}} \theta$ .

- Recall the rate of convergence of the sample mean for i.i.d. random variables with finite variance.
- Let  $\{Y_t\}_{t \geq 1}$  a wide-sense stationary process such that  $\sum_k |\gamma(k)| < +\infty$ . Show that the sample mean  $\bar{Y}_n = (Y_1 + \dots + Y_n)/n$  is consistent and enjoys the same rate of convergence as the i.i.d. case. (Hint: bound  $\mathbb{E}[(\bar{Y}_n - \mu)^2]$  with the  $\gamma(k)$  and recall that convergence in  $L_2$  implies convergence in probability.)

### Answer 1

### 3 AR and MA processes

#### Question 2 Infinite order moving average MA( $\infty$ )

Let  $\{Y_t\}_{t \geq 0}$  be a random process defined by

$$Y_t = \varepsilon_t + \psi_1 \varepsilon_{t-1} + \psi_2 \varepsilon_{t-2} + \cdots = \sum_{k=0}^{\infty} \psi_k \varepsilon_{t-k}$$

where  $(\psi_k)_{k \geq 0} \subset \mathbb{R}$  ( $\psi = 1$ ) are square summable, i.e.  $\sum_k \psi_k^2 < \infty$  and  $\{\varepsilon_t\}_t$  is a zero mean white noise of variance  $\sigma_\varepsilon^2$ . (Here, the infinite sum of random variables is the limit in  $L_2$  of the partial sums.)

- Derive  $\mathbb{E}(Y_t)$  and  $\mathbb{E}(Y_t Y_{t-k})$ . Is this process weakly stationary?
- Show that the power spectrum of  $\{Y_t\}_t$  is  $S(f) = \sigma_\varepsilon^2 |\phi(e^{-2\pi i f})|^2$  where  $\phi(z) = \sum_j \psi_j z^j$ . (Assume a sampling frequency of 1 Hz.)

The process  $\{Y_t\}_t$  is a moving average of infinite order. Wold's theorem states that any weakly stationary process can be written as the sum of the deterministic process and a stochastic process which has the form (2).

#### Answer 2

Deriving  $\mathbb{E}(Y_t)$  and  $\mathbb{E}(Y_t Y_{t-k})$ :

$$\begin{aligned} \mathbb{E}(Y_t) &= \sum_{k=0}^{\infty} \mathbb{E}[\psi_k \varepsilon_{t-k}] \\ &= \sum_{k=0}^{\infty} \psi_k \mathbb{E}[\varepsilon_{t-k}] \\ &= 0 \quad (\varepsilon_k \text{ is a zero mean white noise}) \end{aligned}$$

$$\begin{aligned} \mathbb{E}(Y_t Y_{t-k}) &= \mathbb{E} \left[ \sum_{i=0}^{\infty} \psi_i \varepsilon_{t-i} \sum_{j=0}^{\infty} \psi_j \varepsilon_{t-k-j} \right] \\ &= \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \psi_i \psi_j \underbrace{\mathbb{E}[\varepsilon_{t-i} \varepsilon_{t-k-j}]}_{=0 \text{ if } t-i \neq t-k-j} \\ &= \sigma_\varepsilon^2 \sum_{i=k}^{\infty} \psi_i \psi_{i-k} \end{aligned}$$

The mean function,  $\mathbb{E}$  is constant and does not depend on time  $t$ . The covariance between  $Y_t$  and  $Y_{t+k}$  depends only on the lag  $k$  and not on time  $t$ . We conclude that the process is weakly stationary.

**Show that the power spectrum of  $\{Y_t\}_t$  is  $S(f) = \sigma_\varepsilon^2 |\phi(e^{-2\pi i f})|^2$  where  $\phi(z) = \sum_j \psi_j z^j$**

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$$\begin{aligned}
\mathbb{E}(Y_t Y_{t-\tau}) &= \mathbb{E} \left[ \sum_{i=0}^{\infty} \psi_i \varepsilon_{t-i} \sum_{j=0}^{\infty} \psi_j \varepsilon_{t-\tau-j} \right] \\
&= \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \psi_i \psi_j \mathbb{E}[\varepsilon_{t-i} \varepsilon_{t-\tau-j}] \\
&= \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \psi_i \psi_j \gamma_\varepsilon(i - j - \tau)
\end{aligned}$$

Then, we can write the power spectrum as follows:

$$\begin{aligned}
S(f) &= \sum_{\tau=-\infty}^{\infty} e^{-2i\pi f \tau} \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \psi_i \psi_j \gamma_\varepsilon(i - j - \tau) \\
&= \sigma_\varepsilon^2 \sum_{\tau=-\infty}^{\infty} e^{-2i\pi f \tau} \sum_{i=0}^{\infty} \psi_i \psi_{i-\tau} \\
&= \sigma_\varepsilon^2 \sum_{i=0}^{\infty} \psi_i \sum_{\tau=-\infty}^{\infty} \psi_{i-\tau} e^{-2i\pi f \tau}
\end{aligned}$$

Now, let's make a change of variables. Let  $j = i - \tau$ . Then, the equation becomes:

$$\begin{aligned}
S(f) &= \sigma_\varepsilon^2 \sum_{j=-\infty}^{\infty} \psi_j \sum_{i=j}^{\infty} \psi_i e^{-2i\pi f(i-j)} \\
&= \sigma_\varepsilon^2 \sum_{j=-\infty}^{\infty} \psi_j \sum_{i=j}^{\infty} \psi_i e^{-2i\pi f i} e^{2i\pi f j} \\
&= \sigma_\varepsilon^2 \sum_{j=-\infty}^{\infty} \psi_j e^{2i\pi f j} \sum_{i=j}^{\infty} \psi_i e^{-2i\pi f i} \\
&= \sigma_\varepsilon^2 \left( \sum_{j=-\infty}^{\infty} \psi_j e^{2i\pi f j} \right) \left( \sum_{i=-\infty}^{\infty} \psi_i e^{-2i\pi f i} \right) \\
&= \sigma_\varepsilon^2 |\phi(e^{-2\pi i f})|^2
\end{aligned}$$

### Question 3 AR(2) process

Let  $\{Y_t\}_{t \geq 1}$  be an AR(2) process, i.e.

$$Y_t = \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \varepsilon_t$$

with  $\phi_1, \phi_2 \in \mathbb{R}$ . The associated characteristic polynomial is  $\phi(z) := 1 - \phi_1 z - \phi_2 z^2$ . Assume that  $\phi$  has two distinct roots (possibly complex)  $r_1$  and  $r_2$  such that  $|r_i| > 1$ . Properties on the roots of this polynomial drive the behaviour of this process.

- Express the autocovariance coefficients  $\gamma(\tau)$  using the roots  $r_1$  and  $r_2$ .
- Figure 1 shows the correlograms of two different AR(2) processes. Can you tell which one has complex roots and which one has real roots?
- Express the power spectrum  $S(f)$  (assume the sampling frequency is 1 Hz) using  $\phi(\cdot)$ .
- Choose  $\phi_1$  and  $\phi_2$  such that the characteristic polynomial has two complex conjugate roots of norm  $r = 1.05$  and phase  $\theta = 2\pi/6$ . Simulate the process  $\{Y_t\}_t$  (with  $n = 2000$ ) and display the signal and the periodogram (use a smooth estimator) on Figure 2. What do you observe?

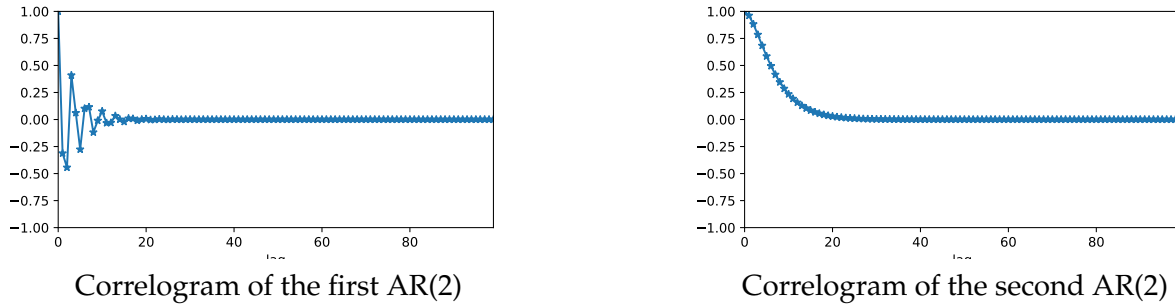


Figure 1: Two AR(2) processes

### Answer 3

Express the autocovariance coefficients  $\gamma(\tau)$  using the roots  $r_1$  and  $r_2$

$$\gamma(\tau) = E[Y_t Y_{t-\tau}]$$

$$Y_{t-\tau} = \phi_1 Y_{t-\tau+1} + \phi_2 Y_{t-\tau+2} + \varepsilon_{t-\tau}$$

$$\begin{aligned} \gamma(\tau) &= E[Y_t (\phi_1 Y_{t-\tau+1} + \phi_2 Y_{t-\tau+2} + \varepsilon_{t-\tau})] \\ &= \phi_1 E[Y_t Y_{t-\tau+1}] + \phi_2 E[Y_t Y_{t-\tau+2}] + E[Y_t \varepsilon_{t-\tau}] \\ \gamma(\tau) &= \begin{cases} \phi_1 \gamma(\tau-1) + \phi_2 \gamma(\tau-2), & \text{if } \tau \neq 0 \\ \phi_1 \gamma(\tau-1) + \phi_2 \gamma(\tau-2) + \sigma_\varepsilon^2, & \text{if } \tau = 0 \end{cases} \end{aligned}$$

Here,  $\sigma^2$  is the variance of the error term  $\varepsilon_t$ .

The characteristic polynomial of this equation is:  $\phi(z) = 1 - \phi_1 z - \phi_2 z^2$ . This characteristic polynomial is the reciprocal of the polynomial  $z^2 - \phi_1 z - \phi_2$ . If the roots of this polynomial are  $r_1$  and  $r_2$ , then roots of our characteristic polynomial are  $1/r_1$  and  $1/r_2$

The final solution for the autocovariance function is thus:

$$\gamma(\tau) = c_1 \left( \frac{1}{r_1} \right)^\tau + c_2 \left( \frac{1}{r_2} \right)^\tau + \sigma_\varepsilon^2 \delta_0(\tau)$$

A more detailed solution can be found [here, at the section 1.3](#)

When  $r_1, r_2 \in \mathbb{C}$ , ie.  $r_1 = re^{i\theta}$  and  $r_2 = re^{-i\theta}$ , where  $r$  is the magnitude of the roots and  $\theta$  is the phase.

$$\gamma(\tau) = E[Y_t Y_{t-\tau}] = \frac{1}{r^\tau} (c_1 \cos(\tau\theta) + c_2 \sin(\tau\theta))$$

This equation represents a damped oscillatory behavior, which is characteristic of AR processes with complex roots.

**The correlogram on the left corresponds to an AR(2) process with complex roots, while the correlogram on the right corresponds to an AR(2) process with real roots.**

**Express the power spectrum  $S(f)$  (assume the sampling frequency is 1 Hz) using  $\phi(\cdot)$ .**

$$Y_t = \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \varepsilon_t$$

We know that:

$$\phi(z) = 1 - \phi_1 z - \phi_2 z^2$$

Applying the lag operator  $L$  to the AR(2) process gives:

$$\phi(L)Y_t = (1 - \phi_1 L - \phi_2 L^2)Y_t = \varepsilon_t$$

This equation can be rewritten as:

$$Y_t = \frac{\varepsilon_t}{\phi(L)} = \frac{\varepsilon_t}{1 - \frac{1}{r_1}L - \frac{1}{r_2}L^2} = \frac{\varepsilon_t}{(1 - \frac{1}{r_1}L)(1 - \frac{1}{r_2}L)}$$

Each term in the product  $(1 - \frac{1}{r_1}z)^{-1}(1 - \frac{1}{r_2}z)^{-1}$  can be expanded as a power series in  $z$ :

$$\begin{aligned} (1 - \frac{1}{r_1}z)^{-1} &= 1 + \frac{1}{r_1}z + \frac{1}{r_1^2}z^2 + \frac{1}{r_1^3}z^3 + \dots = \sum_{i=0}^{\infty} \frac{1}{r_1^i} z^i, \\ (1 - \frac{1}{r_2}z)^{-1} &= 1 + \frac{1}{r_2}z + \frac{1}{r_2^2}z^2 + \frac{1}{r_2^3}z^3 + \dots = \sum_{j=0}^{\infty} \frac{1}{r_2^j} z^j. \end{aligned}$$

The product of these two power series gives the power series expansion of  $\psi^{-1}(z)$

$$\begin{aligned} \phi^{-1}(z) &= \left( \sum_{i=0}^{\infty} \frac{1}{r_1^i} z^i \right) \left( \sum_{j=0}^{\infty} \frac{1}{r_2^j} z^j \right) \\ &= \sum_{k=0}^{\infty} \left( \sum_{i=0}^k \frac{1}{r_1^i r_2^{k-i}} \right) z^k. \end{aligned}$$

This shows that the AR(2) process can be represented as an MA( $\infty$ ) process with MA coefficients given by the terms in the power series expansion of  $\psi^{-1}(z)$

$$Y_t = \sum_{k=0}^{\infty} \left( \sum_{i=0}^k \frac{1}{r_1^i r_2^{k-i}} \right) \varepsilon_{t-k}.$$

$$S(f)_{\text{AR}(2)} = \sigma_{\varepsilon}^2 |\phi(e^{-2\pi i f})|^{-2} = \frac{\sigma_{\varepsilon}^2}{|\phi(e^{-2\pi i f})|^2}.$$

This equation gives the power spectrum of the AR(2) process, which is inversely proportional to the square of the magnitude of the Fourier transform of the MA coefficients.

**Choose  $\phi_1$  and  $\phi_2$  such that the characteristic polynomial has two complex conjugate of norm  $r = 1.05$  and phase  $\theta = 2\pi/6$**

$$1 - \phi_1 z - \phi_2 z^2 = \left(1 - \frac{1}{r_1} z\right) \left(1 - \frac{1}{r_2} z\right)$$

Expanding the right-hand side, we get:

$$1 - \phi_1 z - \phi_2 z^2 = 1 - \left(\frac{1}{r_1} + \frac{1}{r_2}\right) z + \frac{1}{r_1 r_2} z^2$$

Comparing coefficients, we can see that  $\phi_1 = \left(\frac{1}{r_1} + \frac{1}{r_2}\right)$  and  $\phi_2 = \frac{-1}{r_1 r_2}$ .

Given that  $r_1 = r \exp(i\theta)$  and  $r_2 = r \exp(-i\theta)$ , we have:

$$\phi_1 = \frac{2\cos(\theta)}{r}$$

$$\phi_2 = -\frac{1}{r^2}$$

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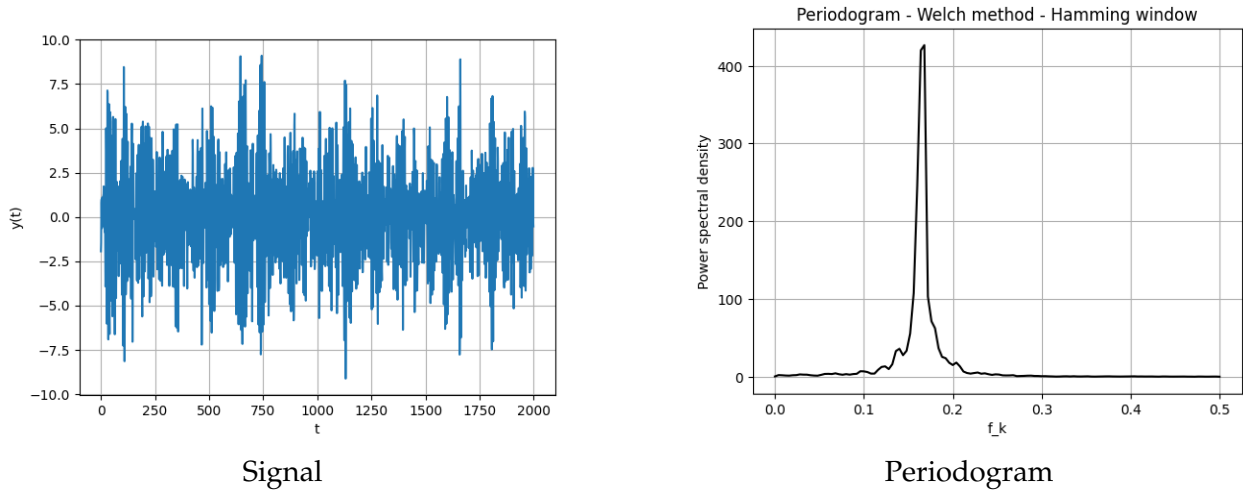


Figure 2: AR(2) process

A peak is observed at a frequency of  $f = 0.17$ . This peak signifies a strong periodic component in the time series at this frequency. The frequency corresponds to the phase of the roots of the AR(2) process, which was set to  $2\pi/6$ . This frequency is the reciprocal of the period, hence the peak at  $f = 0.17$ .



## 4 Sparse coding

The modulated discrete cosine transform (MDCT) is a signal transformation often used in sound processing applications (for instance to encode a MP3 file). A MDCT atom  $\phi_{L,k}$  is defined for a length  $2L$  and a frequency localisation  $k$  ( $k = 0, \dots, L - 1$ ) by

$$\forall u = 0, \dots, 2L - 1, \quad \phi_{L,k}[u] = w_L[u] \sqrt{\frac{2}{L}} \cos\left[\frac{\pi}{L} \left(u + \frac{L+1}{2}\right) \left(k + \frac{1}{2}\right)\right]$$

where  $w_L$  is a modulating window given by

$$w_L[u] = \sin\left[\frac{\pi}{2L} \left(u + \frac{1}{2}\right)\right].$$

### Question 4 *Sparse coding with OMP*

For the signal provided in the notebook, learn a sparse representation with MDCT atoms. The dictionary is defined as the concatenation of all shifted MDCT atoms for scales  $L$  in  $[32, 64, 128, 256, 512, 1024]$ .

- For the sparse coding, implement the Orthogonal Matching Pursuit (OMP). (Use convolutions to compute the correlations coefficients.)
- Display the norm of the successive residuals and the reconstructed signal with 10 atoms.

### Answer 4

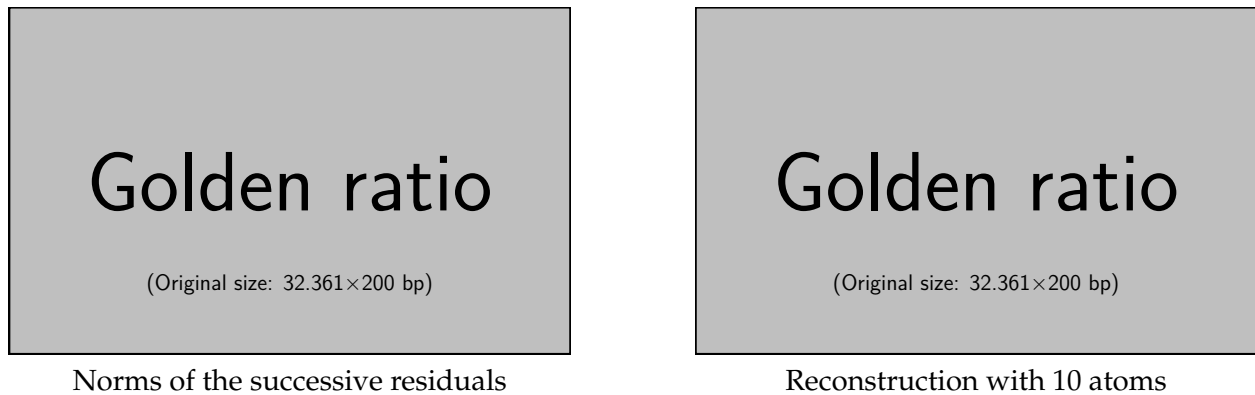


Figure 3: Question 4