Tree-based Multiple Hypothesis Testing with General FWER Control

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1. Mathematical Formulation

Here I consider a testing procedure without data splitting, and only about a case on the 'multiplicative trees'.

In the tree-structured multiple hypotheses setting, we say X_t is a data set to construct each tth p-value. Let $p_t a$ be the p-value of the parent node and p_{t+1} be the p-value of the subsequent child of that parent node. When we fix some $\alpha \in (0,1]$, our basic observation is

$$P(p_{t+1} > \alpha | p_t < \alpha_1) < P(p_{t+1} > \alpha | p_t < \alpha_2),$$

whenever $\alpha_1 \leq \alpha_2$. This is because, we can understand our procedure is actually using

$$p'_{t+1} = \max(p_1, ...p_t, p_{t+1}),$$

as a p-value implemented in the t+1th node, where p_i is all other ancestors of the p_{t+1} . This can be another (maybe equivalent) understanding of the **monotonicity condition** in Goeman and Solari (2010).

From this we can recall the **positive regression dependency on each one from a subset** I_0 (or **PRDS on** I_0) from Benjamini and Yekutieli (2001), which states when D is a non-decreasing set(so whenever $x \in D$ and $x \le y$, then $y \in D$),

$$P(X \in D|X_i = X)$$

is non-decreasing in x.

We also provide some notation for the tree-structured hypotheses. Let us have a tree with depth of d where for some fixed integer k, k is a number of children for every parent node (so we consider *multiplicative trees*). Then we say N is the number of total nodes and n as a number of total clusters, where

$$N = \sum_{r=1}^{d} k^{r-1}$$
$$n = 1 + \sum_{r=1}^{d-1} k^{r-1}.$$

2. Weak FWER Control

Here we consider about the weak FWER control, so when we assume all of the nulls are true. We first impose an assumption, called 'conditional super-uniformity', following Robertson et al. (2023). For the p-values p_i , we define $R_i = \mathbb{1}(p_i \leq \alpha_i)$, where $\mathbb{F}^t = \sigma(R_1, ..., R_t)$ and we let α_t as a \mathbb{F}^{t-1} -measurable function of $(R_1, ..., R_{t-1})$.

Definition 1. The null *p*-values are said to be **conditionally super-uniform** if $P(p_t \le \alpha_t | \mathbb{F}^{t-1}) \le \alpha_t$ for any \mathbb{F}^{t-1} -measurable α_t .

We also give additional notation before proposing our theorem on the weak FWER control. We say C_i is an ordered set of parents of the *i*th node on the tree, and d_i is the depth of the *i*th node. Then we see $|C_i| = d_i + 1$. Denote $C_{i,j}$ as the *j*th element of C_i . For example, consider a binary tree and the case i = 8. Then $C_8 = \{1, 2, 4\}$, and $C_{8,2} = 4$. We are also able to check $|C_8| = 3$, and $d_8 = 2$.

Now we state our theorem.

Theorem 1. If the null p-values are conditionally super-uniform, then the procedure with the critical value functions for every tth node

$$\alpha_t = \begin{cases} \frac{\alpha}{1+k\alpha}, & \text{if all of the } i\text{th tests where } i \in C_t \text{ are rejected,} \\ 0, & \text{otherwise,} \end{cases}$$

implies that the weak FWER is controlled at level α .

Proof. Assume that the all of the nulls are true. First, note that for any ith node and $\alpha \in (0,1)$, since we assumed conditional super-uniformity on the null p-values,

$$\begin{split} P(p_i \leq \alpha) &= P(p_i \leq \alpha \text{ and } \forall j \in C_i, \ p_j \leq \alpha) + P(p_i \leq \alpha \text{ and } \exists j \in C_i \text{ such that } P_j > \alpha) \\ &= P(p_i \leq \alpha \text{ and } \forall j \in C_i, \ p_j \leq \alpha) \\ &= P(p_1 \leq \alpha) P(p_{C_{i,2}} \leq \alpha | p_1 \leq \alpha) ... P(p_i \leq \alpha | p_1 \leq \alpha, ..., p_{C_{i,d_i}} \leq \alpha) \\ &< \alpha^{d_i + 1} \end{split}$$

holds. Therefore,

$$\begin{split} P(V \geq 1) &= P\Big((\bigcap_{i=1}^{N} \{p_i > \alpha_i\})^c\Big) \\ &= P\Big(\bigcup_{i=1}^{N} \{p_i \leq \alpha_i\}\Big) \\ &\leq \sum_{i=1}^{N} P(p_i \leq \alpha_i) \\ &= \underbrace{P(p_1 \leq \alpha_1)}_{p\text{-value on the root node}} + \underbrace{P(p_2 \leq \alpha_2) + \ldots + P(p_{1+k} \leq \alpha_{1+k})}_{k \text{ p-values on the second level}} + \underbrace{P(p_{2+k} \leq \alpha_{2+k}) + P(p_{1+k+k^2} \leq \alpha_{1+k+k^2})}_{k^2 \text{ p-values on the third level}} + \underbrace{P(p_{N-k^{d-1}+1} \leq \alpha_{N-k^{d-1}+1}) + \ldots + P(p_N \leq \alpha_N)}_{k^{d-1} \text{ p-values on the final level}} \\ &\leq \sum_{i=1}^{d} k^{j-1} \alpha_j^i \\ &\leq \sum_{i=1}^{\infty} k^{j-1} \Big(\frac{\alpha}{1+k\alpha}\Big)^j = \alpha, \end{split}$$

which gives us a desired conclusion.

3. General FWER/FDR Control

Now we assume that, whenever j is a child node of i in the tree structure, $X_j \subseteq X_i$. Note that this is a much more realistic than the conditional super-uniformity assumption.

4. Miscellanea

My idea is that,

- Since $p_1, ..., p_n$ (so the p-values after local adjustments) are still has a PRDS property, we can try to directly apply Benjamini and Yekutieli (2001)'s approach to get general FDR or FWER control.
- We can also consider which kind of local adjustments give valid and more powerful tests.

I will further check about,

- Idea of data splitting used in the recent simulation.
- Prove explicitly the validness of locally adjusted p-values from different methods.

REFERENCES

- Yoav Benjamini and Daniel Yekutieli. The control of the false discovery rate in multiple testing under dependency. *The Annals of Statistics*, 38:6:3782–3810, 2001.
- Jell J. Goeman and Aldo Solari. The sequential rejection principle of familywise error control. *The Annals of Statistics*, 38:6:3782–3810, 2010.
- David S. Robertson, James M. S. Wason, and Aaditya Ramdas. Online multiple hypothesis testing. *Statistical Science*, 38:4:557–575, 2023.