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11.105. X 满足二项式分布, 故 $f(X|\theta) = \binom{n}{x} \cdot \theta^x (1-\theta)^{n-x}$, $\theta \in (0, 1)$ 先验分布 $\pi(\theta) = 4\theta^3$ 故 (X, θ) 的联合密度函数 $h(X, \theta) = f(X|\theta) \cdot \pi(\theta) = 4 \binom{n}{x} \cdot \theta^{x+3} (1-\theta)^{n-x}$ 因此, 后验分布的密度函数 $\pi(\theta|X) = \frac{h(X, \theta)}{\int h(X, \theta) d\theta} = \frac{4 \binom{n}{x} \theta^{x+3} (1-\theta)^{n-x}}{4 \binom{n}{x} \int_0^1 \theta^{x+3} (1-\theta)^{n-x} d\theta} = \frac{\theta^{x+3} (1-\theta)^{n-x}}{B(x+4, n-x+1)}$
($\theta \in (0, 1)$)11.106 随机变量服从二项式分布, 即 $f(X|\theta) = \binom{10}{x} \theta^x (1-\theta)^{10-x}$ 先验分布 $\pi(\theta) = 1$ 联合密度函数 $h(X, \theta) = f(X|\theta) \cdot \pi(\theta) = \binom{10}{x} \theta^x (1-\theta)^{10-x}$ 由此知后验分布的密度函数 $\pi(\theta|X) = \frac{h(X, \theta)}{\int h(X, \theta) d\theta} = \frac{\binom{10}{x} \theta^x (1-\theta)^{10-x}}{\binom{10}{x} \int_0^1 \theta^x (1-\theta)^{10-x} d\theta} = \frac{\theta^x (1-\theta)^{10-x}}{B(x+1, 11-x)}$, $\theta \in (0, 1)$ 令 $x=2$, 有 $\pi(\theta|X) = \frac{\theta^2 (1-\theta)^8}{B(3, 9)}$ 11.109 设 X_1, X_2, \dots, X_n 为 n 个样本, 由于随机变量服从泊松分布, 故:

$$f(x|\lambda) = \frac{e^{-\lambda} \lambda^{x_1}}{x_1!} \cdot \frac{e^{-\lambda} \lambda^{x_2}}{x_2!} \cdots \frac{e^{-\lambda} \lambda^{x_n}}{x_n!} = e^{-n\lambda} \cdot \frac{\lambda^{n\bar{x}}}{x_1! x_2! \cdots x_n!}$$

由于先验分布满足 Gamma 分布, 故 $\pi(\lambda) = \frac{\lambda^{\alpha-1} e^{-\lambda/\beta}}{\beta^\alpha \Gamma(\alpha)}$, $\lambda > 0$

$$\text{因此, } h(X, \lambda) = f(X|\lambda) \pi(\lambda) = \frac{e^{-\lambda(n\bar{x}+1)/\beta} \cdot \lambda^{n\bar{x}+\alpha-1}}{x_1! x_2! \cdots x_n! \beta^\alpha \Gamma(\alpha)}$$

$$\text{故 } \pi(\lambda|X) = \frac{h(X, \lambda)}{\int h(X, \lambda) d\lambda} = \frac{e^{-\lambda(n\bar{x}+1)/\beta} \cdot \lambda^{n\bar{x}+\alpha-1}}{\int_0^\infty e^{-\lambda(n\bar{x}+1)/\beta} \cdot \lambda^{n\bar{x}+\alpha-1} d\lambda} = \frac{e^{-\lambda \cdot \frac{n\bar{x}+1}{\beta}} \cdot \lambda^{n\bar{x}+\alpha-1}}{(\frac{\beta}{n\bar{x}+1})^{n\bar{x}+\alpha} \Gamma(n\bar{x}+\alpha)}, \lambda > 0$$

$$\text{令 } \alpha=2, \beta=5, \bar{x}=14, \text{ 有 } \pi(\lambda|X) = \frac{e^{-\lambda \cdot \frac{5}{6}} \lambda^{15}}{(\frac{5}{6})^{16} \Gamma(16)} \sim I(16, \frac{5}{6})$$

$$\text{故平均值为 } 16 \times \frac{5}{6} = \frac{40}{3}, \text{ 方差 } 16 \times (\frac{5}{6})^2 = \frac{100}{9}$$

11.112 样本服从正态分布 $N(\bar{X}, \sigma^2)$, 故:

$$f(X|\theta) = \frac{1}{(2\pi)^{n/2} \sigma^n} \cdot e^{-\frac{\sum (X_i - \theta)^2}{2\sigma^2}}$$

由于 $\sum_{k=1}^n (X_i - \theta)^2 = \sum_{k=1}^n (X_i - \bar{X})^2 + n(\theta - \bar{X})^2$, 又关注带 θ 的项, 有:

$$f(X|\theta) \propto e^{-\frac{n(\theta - \bar{X})^2}{2\sigma^2}}$$

$$\text{先验分布 } \sim N(\mu, \nu^2), \text{ 有 } \pi(\theta) = \frac{1}{\sqrt{2\pi\nu^2}} \cdot e^{-\frac{(\theta - \mu)^2}{2\nu^2}} \propto e^{-\frac{(\theta - \mu)^2}{2\nu^2}} \propto e^{-\frac{[\theta - \frac{n\nu^2\bar{X} + \mu\nu^2}{n\nu^2 + \sigma^2}]^2}{2 \cdot \frac{\sigma^2\nu^2}{n\nu^2 + \sigma^2}}}$$

$$\text{又 } \pi(\theta|X) \propto f(X|\theta) \cdot \pi(\theta) \propto e^{-\frac{n(\theta - \bar{X})^2}{2\sigma^2} - \frac{(\theta - \mu)^2}{2\nu^2}} = e^{-\frac{[\theta - \frac{n\nu^2\bar{X} + \mu\nu^2}{n\nu^2 + \sigma^2}]^2}{2 \cdot \frac{\sigma^2\nu^2}{n\nu^2 + \sigma^2}}}, \theta \in (-\infty, \infty)$$

$$\text{令 } \bar{x}=2, n=20, \sigma=0.3, \mu=1.5, \nu^2=0.1$$

$$\text{有 } \mu_{\text{post}} = \frac{n\nu^2\bar{x} + \mu\nu^2}{n\nu^2 + \sigma^2} = 1.98, \nu_{\text{post}}^2 = \frac{\sigma^2\nu^2}{n\nu^2 + \sigma^2} = 0.00431$$

即 后验分布 $\sim N(1.98, 0.00431)$