

Lean on your statistics: the simplification of the balance intercept problem

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Introduction

As the field of epidemiology evolves, there are growing interest to employ more computation approaches to solve analytic problems. Among them, simulation is one of the most accessible concept. Previous literature argues the importance of simulation in epidemiology education and research. While seeing the power of computational analytic tools, we caution the excess reliance on the computation in analytic problem solving and neglect of some fundamental statistics theory. In the article, we demonstrate how a simple and basic statistics knowledge can simply analytic problems by visiting a particular simulation problem, the balance intercept.

The balance intercept problem was first introduced by Rudolph et al. (2021) to address the problem of controlling the marginal probability of binary outcomes when constructing a simulation study. The authors proposed to numerically calculate the “balance intercept” to replace the standard intercept in simulation procedures. This same problem was later revisited by Robertson, Steingrimsen, and Dahabreh (2021) who discovered that deriving the balance intercept analytically using Rudolph et al. (2021) can produce inaccurate estimation and hence unable to control the marginal probability at the desired level. Instead, Robertson, Steingrimsen, and Dahabreh (2021) proposed a numeric solution to solve for the balance intercept for binomial simulation with a logistic link function. Later, Zivich and Ross (2022) did xxx for multi-level categorical variable.

In the rest of this article, we look into the balance intercept problem with different statistical lenses. Specifically, we summarize what statistical problem the balance intercept manifest/represent, how the original authors set up the problem in a reference coding system, and how to translate the problem into the effect coding scope which provides a much simple solution without algorithmic integration.

In order to develop the statistical perspective about the balance intercept problem, we need to first introduce the coding scheme concept. Coding scheme describe the way a categorical variable is enumerated in the regression system. Most commonly used coding schemes in analysis include reference coding (also known as dummy coding) and the effect coding. The reference coding creates $p - 1$ binary columns (i.e. the value of each cell is either 0 or 1) for a categorical variable with p levels. The reference coding provides very easy effect comparison within levels of the categorical variables. In comparison, the effect coding also enumerates a p -level categorical variable into a matrix with $p - 1$ columns but the value of the cells normally include 0, 1, and -1. The effect coding provides interpretation that how the mean of each level deviates from the mean of means, and is preferred coding scheme in experimental design. For more technical details, we defer to [TODO: add textbook, probably INTRODUCING ANOVA and ANCOVA by andrew rutherford]. To note, the selection of coding scheme doesn't limit to the . In other words, both coding scheme can be applied ubiquitously without limiting to certain the model of choice, and link function of choice.

Here, we adapt the toy example in xxx to illustrate the difference and interpretation of the two coding scheme.

The balance intercept is an adjustment to the conditional probability of the reference level of a categorical covariate. Nevertheless, both methods require extensive calculations, and there could be a more straightforward solution.

[TODO: add what the problem is, what the authors problem. degree of freedom.]

based on the reference coding scheme

I proposed an alternative simulation strategy to control the marginal probability of a binary outcome. Instead of constructing the simulation based on the reference coding scheme, we encourage using effect coding, specifically deviation coding, to construct the design matrix of categorical covariates. The theoretical basis of this proposal is that the intercept term of the effect coding model (regardless of the link function or parametric assumptions) is the mean of the group means, which coincides with the marginal probability of a binary outcome when the groups are balanced. Hence, to simulate data from balanced designs, no additional calculation is needed compared to the previous proposals. In the case of unbalanced design, it is very intuitive to adjust the simulation equations and requires minimum arithmetic calculation. (See examples below) Besides its straightforwardness, the proposed solution should not incur any knowledge burden as the effect coding is commonly introduced in introductory statistics classes.

Effect Coding Primer

1. What is the reference coding, and why people prefer it.
2. What is the effect coding, and why people prefer that.
3. How to translate the effect coding and reference coding.

Balance Intercept of effect Coding

1. Treatment of continuous variable
2. Treatment of categorical variables
3. Generalizability of the approach

Balanced design

It is very simple to control the marginal probability for balanced design when the simulation is based on the effect coding. No calculations is required as the intercept in an effect coding model is the grand mean, in the binary outcome case, the marginal probability. Here we illustrate the process following the simple additive probability example in Rudolph et al. (2021), where the target marginal probability is 0.3, the effect size as in risk difference is 0.2, and a balanced covariate with 2 levels.

```
set.seed(123)

n <- 10000

# Marginal probabilities of each variable
p.y <- 0.3
p.x <- 0.5
rd <- 0.2

# Example 2: Generate L, X, and Y -----
X <- rbinom(n, 1, p.x)

# Generate X with marginal prob 0.5
dev_coding <- contr.sum(2) # Deviation Coding with 2 levels
```

```

X_design_dev <- cbind( 1,    # Adding intercept column
                      dev_coding[X+1,]) # Construct the design matrix

# The design matrix with effect coding can be more easily construct with model.matrix function

beta_vec <- c(p.x,    # Intercept term, the marginal probability for balanced design
              -rd/2) # Set up conditional prob for reference level
Y <- rbinom(n, 1, X_design_dev %*% beta_vec)

```

To validate the simulation, we can see the marginal probability 0.5041, and the conditional probability of the two levels are 0.3985 and 0.6122 respectively. Hence the simulation matches with the desired design.

```
mean(Y);
```

```
## [1] 0.5041
```

```
mean(Y[X==0]);
```

```
## [1] 0.3984576
```

```
mean(Y[X==1]);
```

```
## [1] 0.6121788
```

One can use the following code to examine data quality from the modeling perspective.

```

summary(glm(Y~X, family = binomial(link="identity"))) # Reference coding model
summary(glm(Y~X_design_dev-1, family = binomial(link="identity"))) # Effect coding model

```

Unbalanced design

When the groups are not balanced, the simulation with effect coding is less straightforward compared to the balanced case, mainly because the equality between intercept and grand mean doesn't hold. The intercept needs to be adjusted based on the conditional probability of one of the levels (default to the reference level in the reference coding scheme). This adjustment of intercept requires some arithmetic calculation; nevertheless, in the author's biased view, the complexity is still manageable and requires less calculation than the previous proposals. We demonstrate the simulation procedure for unbalanced design with a toy example. The simulation settings are similar to the example above except that we change the group ratio to 8:2 and the effect size to 0.4.

To calculate the new intercept, we need first to establish the conditional probability for one of the levels (by default the $X = 0$ level in this example). As we know that the marginal probability can be expressed

$$Pr(Y = 1) = \frac{n_1 Pr(Y = 1|X = 0) + n_2 Pr(Y = 1|X = 1)}{n_1 + n_2} = \frac{n_1 Pr(Y = 1|X = 0) + n_2 (Pr(Y = 1|X = 0) + RD)}{n_1 + n_2}$$

where RD is the effect size in risk difference, n_1 and n_2 are the group sample size for $X = 0$ and $X = 1$ respectively. Given $Pr(Y = 1)$, n_1 , n_2 and RD, we can easily derive the conditional probability of $X=0$,

$$Pr(Y = 1|X = 0) = \frac{(n_1 + n_2)Pr(Y = 1) - n_2RD}{n_1 + n_2}.$$

The intercept, a_0 , as the mean of the group means can be calculated with

$$a_0 = \frac{2Pr(Y = 1|X = 0) + RD}{2}.$$

The simulation procedure translates to the toy example as

```
set.seed(123)

n <- 10000

# Marginal probabilities of each variable
p.y <- 0.3
p.x <- 0.8      # Imbalanced design
rd <- 0.2

cond.p <- (n*p.y - n*(p.x)*rd)/n
a.0 <- cond.p + rd/2

# Example 2: Generate L, X, and Y -----
X <- rbinom(n, 1, p.x)

# Generate X with marginal prob 0.5
dev_coding <- contr.sum(2) # Deviation Coding with 2 levels
X_design_dev <- cbind( 1, # Adding intercept column
                      dev_coding[X+1,]) # Construct the design matrix

# The design matrix with effect coding can be more easily construct with model.matrix function

beta_vec <- c(a.0, # Intercept term, the calculated mean of group means
              -rd/2) # Set up conditional prob for reference level
eta <- X_design_dev %*% beta_vec
eta[eta<0] <- 0
eta[eta>1] <- 1
Y <- rbinom(n, 1, eta)
```

A quick examination shows the simulated data matches with the expectation.

```
mean(Y);

## [1] 0.2976

mean(Y[X==0]);

## [1] 0.1380195
```

```
mean(Y[X==1]);
```

```
## [1] 0.3362315
```

```
# summary(glm(Y~X, family = binomial(link="identity"))) # Reference coding model  
# summary(glm(Y~X_design_dev-1, family = binomial(link="identity"))) # Effect coding model
```

To note, if we ignore the group ratio change and don't adjust the intercept in the simulation procedure (for example, use the balanced-design simulation procedure without any modification), the observed marginal probability deviates from the target marginal probability and the deviation is more obvious with more extreme value of effect size and unbalanced group ratio.

Conclusion

In this report, we propose to use the effect coding scheme in simulations to address the problem of controlling marginal probability of binary outcomes. We provide preliminary evidence that the proposal works for both balanced and unbalanced designs via toy examples. Compared to the previous solutions that based on the reference coding scheme, our proposed solution requires less calculation than the approaches to derive analytic and numeric approximation of the balance point. Particularly, it requires modest calculation when the study design is balanced.

The problem of controlling marginal probability in essence is to find the conditional probability for the reference group, as all the other simulation parameters and the model degrees of freedom are considered fixed.

In this report, we only consider easy simulation scenarios, i.e. binary covariates, one covariate, identify link function, to demonstrate the feasibility of this simulation strategy. We anticipate with the levels of a covariate and the number of covariates growing, the calculation complexity would grow but still be manageable. We will provide a more delicate equation, particularly for the unbalanced design, to generalize for those situation. We will also conduct larger scale of simulation studies to evaluate the efficacy of the proposed solution.

The reason why this is necessary * significantly reduces the complexity of the problem, leveraging a statistics concept covered in the introductory level statistics class * Require less numeric programming, which can easily produce errors for people without numeric programming background. The simulation strategy translates easily to other programming e.g. *SAS*, *STATA*, whose implementations are not shown in this article.

The author doesn't argue if analytic approach and computational approach is superior. Instead, we advocate for a balanced emphasis on both computational skills and analytic thinking.

References

- Robertson, Sarah E, Jon A Steingrimsdottir, and Issa J Dahabreh. 2021. "Using Numerical Methods to Design Simulations: Revisiting the Balancing Intercept." *American Journal of Epidemiology*, November, kwab264. <https://doi.org/10.1093/aje/kwab264>.
- Rudolph, Jacqueline E, Jessie K Edwards, Ashley I Naimi, and Daniel J Westreich. 2021. "SIMULATION IN PRACTICE: THE BALANCING INTERCEPT." *American Journal of Epidemiology* 190 (8): 1696–98. <https://doi.org/10.1093/aje/kwab039>.
- Zivich, Paul N, and Rachael K Ross. 2022. "RE: "Using Numerical Methods to Design Simulations: Revisiting the Balancing Intercept"." *American Journal of Epidemiology*, May. <https://doi.org/10.1093/aje/kwac083>.