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# Background

"It is extremely unlikely that the true (effect) function f(X) (on the outcome) is actually linear in X."

— Hastie, Tibshirani, and Friedman (2009) PP. 139

#### Question

How to model nonlinear effects for survival outcome in **high-dimensional** setting?

Following all necessary assumptions, a Cox proportional hazards model with event time  $t_i$  and predictors  $x_{ij}$ , j = 1, ..., p, is expressed as

$$h(t_i) = h_0(t_i) \exp(\sum_{j=1}^{p} B_j(x_{ij})), \quad i = 1, \ldots, n.$$

Spline functions

Background

$$B(x) = \sum_{k=1}^{K} \beta_k b_k(x) \equiv \boldsymbol{X}^T \boldsymbol{\beta}$$

 $b_k(x)$  are the *basis functions*, possibly truncated power basis and b-spline basis. (Wood 2017)

# Function Smoothing

- ► Smoothing penalty pen<sub> $\lambda$ </sub>( $B_i(X_i)$ ) =  $\lambda \int B''(X)^2 dx = \lambda \beta^T S\beta$ 
  - ▶ The smoothing penalty matrix **S** is known given **X**
  - ▶ **S** is symmetric and positive semi-definite
- Penalized Partial Likelihood Function

$$pl(\beta) = \sum_{i=1}^{n} d_i \log \frac{\exp(\beta^T \mathbf{x}_i)}{\sum_{i' \in R(t_i)} \exp(\beta^T b s x_{i'})} - \sum_{j=1}^{p} \operatorname{pen}_{\lambda}(B_j(X_j)),$$

lacktriangle The smoothing parameter  $\lambda$  is a tuning parameter, selected via cross-validation

# High-dimensional Additive Cox Model

#### Primary Challenges:

- Jointly model signal sparsity versus function smoothness
  - Smooth penalty only overfits the model
  - Sparsity penalty only overshinks the coefficients
  - Damage predictive performance
- Adaptive shrinkage
  - Global shrinkage assumes similar function smoothness
- Bi-level selection that simultaneously answers
  - if a variable is predictive to the outcome,  $B_j(X_j) = 0$
  - ightharpoonup if a variable has a nonlinear relationship with the outcome,  $B_j(X_j)=\beta_jX_j$

# Bayesian Hierarchical Additive Model

- Two-part spike-and-slab LASSO prior for spline functions
  - Variable selection via inclusion indicator
  - ▶ Bi-level selection accounting for effect hierarchy
  - Adaptive shrinkage via Bayesian regularization
- EM-Coordinate Descent algorithm
  - Expedited computation
  - Seamless variable selection via sparse solution

# Two-part Spike-and-slab LASSO (SSL) Prior

Follow Marra and Wood (2011), a spline function  $B(X) = \mathbf{X}^T \boldsymbol{\beta}$  can be decomposed to linear and nonlinear components with respect to the smoothing penalty matrix S

$$\boldsymbol{X}^{T}\boldsymbol{\beta} = X^{0}\boldsymbol{\beta} + \boldsymbol{X}^{*}\boldsymbol{\beta}^{*}$$

Two-part spike-and-slab LASSO prior

$$egin{aligned} eta_j | \gamma_j, s_0, s_1 &\sim \mathsf{DE}(0, (1 - \gamma_j) s_0 + \gamma_j s_1) \ eta_{jk}^* | \gamma_j^*, s_0, s_1 &\stackrel{\mathsf{iid}}{\sim} \mathsf{DE}(0, (1 - \gamma_j^*) s_0 + \gamma_j^* s_1), k = 1, \dots, K - 1 \end{aligned}$$

- $ightharpoonup \gamma_i$  controls the inclusion of linear component
- $\triangleright \gamma_i^*$  controls the inclusion of nonlinear component

# Effect hierarchy assumes lower-order effects are more likely to be active than higher-order effects

lacktriangle Structured prior on latent indicators  $\gamma_j$  and  $\gamma_j^*$ 

$$\gamma_j | heta_j \sim extit{Bin}(\gamma_j | 1, heta_j), \quad \gamma_j^* | \gamma_j, heta_j \sim extit{Bin}(1, \gamma_j heta_j),$$

Simplification via analytic integration

$$\gamma_j^* | \theta_j \sim \textit{Bin}(1, \theta_j^2),$$

Adaptive shrinkage

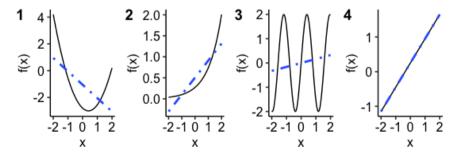
$$\theta_i \sim \text{Beta}(a, b)$$

### EM-Cooridante Descent Algrithm

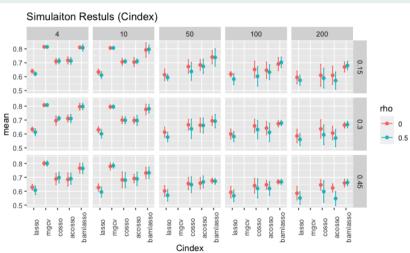
We are interested in estimating  $\Theta = \{\beta, \theta\}$  using optimization based algorithm for scalability purpose

- ightharpoonup Treat  $\gamma$ s as the "missing data" in the EM procedure
  - Construct the conditional expectation of the posterior density function
- Decompose the posterior density function to two pieces
  - $ightharpoonup L_1$  penalized partial likelihood function of eta
  - **Posterior density function of**  $\theta$
- Maximize the two pieces independently
  - ightharpoonup Optimize  $\beta$  with coordinate descent algorithm
  - ightharpoonup Optimize  $\theta$  with beta-binomial conjugate relationship

- $ightharpoonup n_{train} = 500, \ n_{test} = 1000$
- p = 4, 10, 50, 100, 200
- Survival and censoring time follow Weibull distribution
  - ► Censoring rate is controlled at {0.15, 0.3, 0.45}



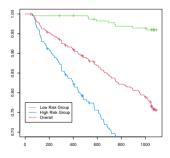
#### Prediction Performance



A Scalable and Flexible Cox Proportional Hazards Model for High-Dimensional Survival Prediction and Eunctional Selection

# Emory Cardiovascular Biobank

- All-cause mortality among patents undergoing cardiac catheterization
- Sample size N=454 and number of features p=200
- 5-knot cubic spline



#### Conclusion

- ▶ A scalable and flexible Cox Model for high-dimensional survival data analysis
  - Two-part spike-and-slab LASSO prior for spline functions
    - ▶ Jointly model signal sparsity and function smoothness with adaptive regularization
    - ▶ Bi-level selection that accounts the effect hierarchy principle
  - EM-Coordinate Descent algorithm
    - Computation advantage and sparse solution
- R package: BHAM
  - Ancillary functions for high-dimensional formulation
  - ► Model summary and variable selection
  - Website via boyiguo1.github.io/BHAM

#### References

#### References I

- Hastie, Trevor, Robert Tibshirani, and Jerome Friedman. 2009. *The Elements of Statistical Learning: Data Mining, Inference, and Prediction*. Springer Science & Business Media.
- Marra, Giampiero, and Simon N Wood. 2011. "Practical Variable Selection for Generalized Additive Models." *Computational Statistics & Data Analysis* 55 (7): 2372–87.
- Wood, Simon N. 2017. Generalized additive models: An introduction with R, second edition. https://doi.org/10.1201/9781315370279.