# Spike-and-Slab LASSO Generalized Additive Models and Fast Algorithms for High-Dimensional Data Analysis

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#### Outline

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  - EM-Coordinate Descent algorithm
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#### Background

# Non-linear Effect Modeling

"It is extremely unlikely that the true (effect) function f(X) (on the outcome) is actually linear in X."

- Hastie, Tibshirani, and Friedman (2009) PP. 139
- Traditional modeling approaches
  - ▶ Categorization of continuous variable, polynomial regression
  - Simple but may be statistically flawed
- Machine learning methods
  - ▶ Black-box algorithms: Random forests, neural network
  - Predict accurate but too complicated for interpretation

Firstly formalized by Hastie and Tibshirani (1987)

$$y_i \stackrel{\text{iid}}{\sim} EF(\mu_i, \phi), \quad i = 1, \dots, n$$

$$\mu_i = g^{-1}(\beta_0 + \sum_{j=1}^p B_j(x_j))$$

where  $B_j(x_j)$  is a smoothing function,  $g(\cdot)$  is a link function,  $\phi$  is the dispersion parameter \* Objective: to estimate smoothing functions  $B_j(x_j)$  \* Applications in biomedical research: \* Dose-response curve \* Time-varying effect

## High-dimensional GAM

- Grouped penalty models
  - Grouped lasso penalty (Ravikumar et al. 2009; Huang, Horowitz, and Wei 2010). grouped SCAD penalty (Wang, Chen, and Li 2007; Xue 2009)
  - Sparse penalty induces excess shrinkage, causing inaccurate interpolation of non-linear effect
- Bavesian Hierarchical Models
  - Grouped spike-and-slab priors (Scheipl, Fahrmeir, and Kneib 2012; Yang and Narisetty 2020), grouped spike-and-slab lasso prior(Bai et al. 2020; Bai 2021)
  - Mostly Markov chain Monte Carlo methods for model fitting
  - Computational inefficiency causes scaling problems in high-dimensional data analysis

# Other challenges

- Bi-level selection
  - ► To detect a smoothing function is linear and nonlinear effects
  - ▶ All-in-all-out selection reduces the ability of result interpretation
- Uncertainty inferences
  - Penalized models doesn't provide uncertainty measures
  - ► Challenging to estimate the effective degree of freedom for each smoothing functions

## Objectives

- ➤ To develop statistical models that improve curve interpolation and outcome predicting
  - Local adaption of sparse penalty and smooth penalty
  - ▶ Bi-level selection for linear and nonlinear effect
- To develop a fast and scalable algorithm
- To implement a user-friendly statistical software

## Bayesian Hierarchical Additive Model (BHAM)

Given the data  $\{X_i, y_i\}_{i=1}^n$  where  $X_i \in \mathbb{R}^p$ ,  $y_i \in \mathbb{R}$  and p >> n, we have the generalized additive model

$$y_i \stackrel{\text{i.i.d.}}{\sim} EF(\mu_i, \phi),$$
 $g(\mu_i) = g^{-1}(\beta_0 + \sum_{j=1}^p B_j(x_j)), \quad i = 1, \dots, n.$ 

The smoothing function can be written in a matrix form  $B_j(x_j) = \beta_j^T \mathbf{X}_j$ , where  $\beta_j$  are the coefficients of the smoothing function and  $\mathbf{X}_j$  is the basis matrix of dimension  $K_j$ .

#### Smoothing Function Reparameterization

Smoothing penalty from Smoothing spline regression (Wood2017?)

$$\lambda_j \int B_j''(x) dx = \lambda_j \beta_j^T \mathbf{S}_j \beta_j,$$

where  $S_i$  is a known smoothing penalty matrix.

 $\triangleright$  Isolate the linear and nonlinear componentsvia eigendecomposing  $S_i$ 

$$\mathbf{X}\boldsymbol{\beta} = X^0\boldsymbol{\beta} + \mathbf{X}^*\boldsymbol{\beta}^*$$

- Renefits
  - Motivate bi-level selection
  - Implicit modeling of function smoothness
  - Reduce computation load with conditionally independent prior of basis coefficients

## Two-part Spike-and-slab LASSO (SSL) Prior

SSL prior for linear coefficients and group SSL priors for nonlinear coefficients

$$egin{aligned} eta_j | \gamma_j, s_0, s_1 &\sim \mathsf{DE}(0, (1-\gamma_j) s_0 + \gamma_j s_1) \ eta_{jk}^* | \gamma_j^*, s_0, s_1 &\stackrel{\mathsf{iid}}{\sim} \mathsf{DE}(0, (1-\gamma_j^*) s_0 + \gamma_j^* s_1), k = 1, \dots, K_j \end{aligned}$$

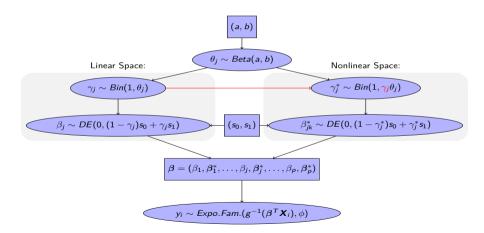
Effect hierarchy enforced latent inclusion indicators  $\gamma_j$  and  $\gamma_i^*$  for bi-level selection

$$\gamma_j | heta_j \sim \textit{Bin}(\gamma_j | 1, heta_j), \quad \gamma_j^* | \gamma_j, heta_j \sim \textit{Bin}(1, \gamma_j heta_j),$$

Local adaptivity of signal sparsity and function smoothness

$$\theta_j \sim \mathsf{Beta}(a,b)$$

#### Visual Representation



# EM-Cooridante Descent Algrithm for Scalable Model Fitting

We are interested in estimating  $\Theta = \{\beta, \theta, \phi\}$  using optimization based algorithm for scalability purpose

- Basic Ideas
  - ightharpoonup Treat  $\gamma$ s as the "missing data" in the EM procedure
  - P Quantify the expectation of log posterior density function of  $\Theta$  with respect to  $\gamma$  conditioning on  $\Theta^{(t-1)}$
  - Maximize the independent parts of the objective function using Coordinate Descent algorithm and closed-form equations
- Previous applications in high-dimensional data analysis
  - ► EMVS (Ročková and George 2014), Spike-and-slab lasso (Ročková and George 2018)
  - ► BhGLM (Yi et al. 2019)

#### Decomposition of Objective Function

We aim to maximize the log posterior density of  $\Theta$  by averaging over all possible values of  $\gamma$ 

$$\log f(\Theta, \gamma | \mathbf{y}, \mathbf{X}) = Q_1(\beta, \phi) + Q_2(\gamma, \theta),$$

 $L_1$ -penalized likelihood function of  $\beta$ ,  $\phi$ 

$$Q_1 \equiv Q_1(oldsymbol{eta},\phi) = \log f(\mathbf{y}|oldsymbol{eta},\phi) + \sum_{j=1}^p \left[ \log f(eta_j|\gamma_j) + \sum_{k=1}^{\mathcal{K}_j} \log f(eta_{jk}^*|\gamma_{jk}^*) 
ight]$$

Posterior density of  $\theta$  given data points  $\gamma$ s

$$Q_2 \equiv Q_2(\gamma, oldsymbol{ heta}) = \sum_{j=1}^p \left[ (\gamma_j + \gamma_j^*) \log heta_j + (2 - \gamma_j - \gamma_j^*) \log (1 - heta_j) 
ight] + \sum_{j=1}^p \log f( heta_j).$$

 $Q_1$  and  $Q_2$  are independent conditional on  $\gamma$ s

## Summary of EM-Coordinate Descent Algorithm

- ► E-step
  - Formulate  $E_{\gamma|\Theta^{(t)}}\left[Q(\Theta,\gamma)\right]=E(Q_1)+E(Q_2)$ 
    - $ightharpoonup E(Q_1)$  is a penalized likelihood function of  $\beta, \phi$
    - $E(Q_2)$  is a posterior density of  $\theta$  given  $E(\gamma)$
    - $ightharpoonup E(Q_1)$  and  $E(Q_2)$  are conditionally independent
  - lacktriangle Calculate  $E(\gamma_j)$  and  $E(\gamma_j^*)$ , and penalties by Bayes' theorem
- M-step:
  - Use Coordinate Descent to fit the penalized model in  $E(Q_1)$  to update  $\beta, \phi$
  - ▶ Closed form calculation via  $E(Q_2)$  to update  $\theta$

## Tuning Parameter Selection

- $\triangleright$   $s_0$  and  $s_1$  are tuning parameters
- Empirically,  $s_1$  has extremely small effect on changing the estimates
- Focus on tuning s<sub>0</sub>
- ► Consider a sequence of L ordered values  $\{s_0^I\}$ :  $0 < s_0^1 < s_0^2 < \dots < s_0^L < s_1$
- $\triangleright$  Cross-validation to choose optimal value for  $s_0$

# Simulation Study

# Simulation Study

- ▶ Follow the data generating process introduced in Bai et al. (2020).
- $n_{train} = 500, n_{test} = 1000$
- p = 4, 10, 50, 200

$$\mu = 5\sin(2\pi x_1) - 4\cos(2\pi x_2 - 0.5) + 6(x_3 - 0.5) - 5(x_4^2 - 0.3),$$

- $f_j(x_j) = 0$  for j = 5, ..., p.
- lacksquare 2 types of outcome: Gaussian ( $\phi=1$ ), Binomial
- Splines are constructed using 10 knots
- ▶ 50 Iterations

## Comparison & Metircs

- Methods of comparison
  - Proposed model BHAM
  - Linear LASSO model as the benchmark
  - ► mgcv (Wood 2004)
  - COSSO (Zhang and Lin 2006) and adaptive COSSO(Storlie et al. 2011)
  - Sparse Bayesian GAM (Bai 2021)
  - spikeSlabGAM [TODO: add ]
- Metrics
  - ightharpoonup Prediction:  $R^2$  for continuous outcomes, out-of-sample AUC for binary outcomes
  - Variable Selection Performance:

#### Simulation Results

- ▶ The proposed method works better in low, medium, high settings than other state-of-art methods
- ► SB-GAM works better in ultra-high setting

4       0.94 (0.01)       0.89 (0.04)       0.90 (0.02)       0.90 (0.02)       0.94 (0.01)       0.93 (0.02)         10       0.93 (0.01)       0.87 (0.03)       0.87 (0.03)       0.85 (0.03)       0.92 (0.04)       0.92 (0.02)         50       0.92 (0.01)       0.87 (0.02)       0.83 (0.02)       0.83 (0.02)       0.76 (0.04)       0.92 (0.02)         200       0.88 (0.01)       0.86 (0.02)       0.81 (0.06)       0.81 (0.08)       -       0.92 (0.02)	р	EM-IWLS	EM-CD	COSSO	ACOSSO	mgcv	SB-GAM
50 <b>0.92 (0.01)</b> 0.87 (0.02) 0.83 (0.02) 0.83 (0.02) 0.76 (0.04) 0.92 (0.02)	4	0.94 (0.01)	0.89 (0.04)	0.90 (0.02)	0.90 (0.02)	0.94 (0.01)	0.93 (0.01)
	10	0.93 (0.01)	0.87 (0.03)	0.87 (0.03)	0.85 (0.03)	0.92 (0.04)	0.92 (0.01)
200 0.99 (0.01) 0.96 (0.02) 0.91 (0.06) 0.91 (0.09)	50	0.92 (0.01)	0.87 (0.02)	0.83 (0.02)	0.83 (0.02)	0.76 (0.04)	0.92 (0.01)
200 0.88 (0.01) 0.80 (0.02) 0.81 (0.00) 0.81 (0.08) - 0.92 (0.0	200	0.88 (0.01)	0.86 (0.02)	0.81 (0.06)	0.81 (0.08)	-	0.92 (0.01)

#### Conclusion

#### Conclusion

- Propose a scalable Bayesian Hierarchical Additive Model (BHAM) for high-dimensional data analysis
  - Organic balance between sparse penalty and smooth penalty
  - ▶ Bi-level selection for linear- and non-linear effects
  - Uncertainty measures provided
- R package: BHAM
  - Ancillary functions for high-dimensional formulation
  - Model summary and variable selection
  - Covariate adjustment without penalty
  - Website via boyiguo1.github.io/BHAM

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