## Ch. 16 – Statistical Inference

Def'n: <u>Estimation</u> is the assignment of value(s) to a population parameter based on a value of the corresponding sample statistic.

An estimator is a rule used to calculate an estimate.

An <u>estimate</u> is a specific value of an estimator.

Note: in this chapter, always assuming an SRS.

## - Notation:

- Let  $\theta$  be a generic parameter.
- Let  $\hat{\theta}$  be an estimator a statistic calculated from a random sample
- Consequently,  $\hat{\theta}$  is an r.v. with mean  $E(\hat{\theta}) = \mu_{\hat{\theta}}$  and std. dev.  $\sigma_{\hat{\theta}}$

Def'n: A <u>point estimate</u> is a *single number* that is our "best guess" for the parameter.

 $\rightarrow$  like a *statistic*, but more precise towards parameter estimation.

An <u>interval estimate</u> is an *interval of numbers* within which the parameter value is believed to fall.

Generic large sample confidence intervals:

Def'n: A <u>confidence interval (CI)</u> for a parameter  $\theta$  is an interval estimate of plausible values for  $\theta$ . With a chosen degree of confidence, the CI's construction is such that the value of  $\theta$  is captured between the statistics L and U, the lower and upper endpoints of the interval, respectively.

The <u>confidence level</u> of a CI estimate is the success rate of the *method* used to construct the interval (as opposed to confidence in any particular interval). The generic notation is  $100(1-\alpha)\%$ . Typical values are 90%, 95%, and 99%.

Ex16.1) Using 95% and the upcoming method to construct a CI, the method is "successful" 95% of the time. That is, if this method was used to generate an interval estimate over and over again with different samples, in the long run, 95% of the resulting intervals would capture the true value of  $\theta$ .

Many large-sample CIs have the form:

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point estimate \pm (critical value) \times (standard error)
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where "point estimate" is a statistic  $\hat{\theta}$  used to estimate parameter  $\theta$ ,

"standard error" is a statistic  $\hat{\sigma}_{\hat{\theta}}$  used to estimate std. dev. of estimator  $\hat{\theta}$ ,

"critical value" is a fixed number z defined so that if Z has std. norm. dist'n, then  $P(-z \le Z \le z) = 1 - \alpha = \text{confidence level}$ 

The product of the "standard error" and "critical value" is the *margin of error*.

Note: critical value z often denoted by  $z_{\alpha/2}$ , where the notation reflects  $P(Z > z) = \alpha/2$ .

Ex16.2) If the confidence level is 95%, what is the critical value?

Table 16X0 – Critical values for usual confidence levels, using three decimal places

$100(1-\alpha)\%$	α	α/2	$z_{\alpha/2}$
90%	0.10	0.050	1.645
95%	0.05	0.025	1.960
99%	0.01	0.005	2.576

The estimator  $\hat{\theta}$  and its standard error  $\hat{\sigma}_{\hat{\theta}}$  are defined so that, when the sample size n is sufficiently large, the sampling distribution of

$$\frac{\hat{\theta} - \theta}{\hat{\sigma}_{\hat{n}}} \stackrel{\sim}{\sim} N(0, 1)$$

Thus,

$$P\left(-z \le \frac{\hat{\theta} - \theta}{\hat{\sigma}_{\hat{\theta}}} \le z\right) \approx 1 - \alpha$$

Algebraic manipulation yields

$$P(\hat{\theta} - z\hat{\sigma}_{\hat{\theta}} \le \theta \le \hat{\theta} + z\hat{\sigma}_{\hat{\theta}}) \approx 1 - \alpha$$

Large Sample CI for Population Proportion

Recall the 3 rules regarding the general properties of the sampling distribution of  $\hat{p}$ .

Then, when n is large, a  $(1 - \alpha)100\%$  CI for p is

$$\hat{p} \pm z_{\alpha/2} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

Note that *n* being large also allows for the standard error to use  $\hat{p}$  since *p* is unknown.

## Assumptions:

- 1.  $n\hat{p} \ge 15$  and  $n(1-\hat{p}) \ge 15$ ,
- 2. the sample can be regarded as a random sample from the population of interest.

Ex16.3) A survey of 1356 random adults asked them to pick out the funniest city name in a list. 923 chose "Keokuk", 74 chose "Walla Walla", and 359 chose "Seattle". Let *p* be the proportion of all adults who would have answered "Seattle" had they been polled. Construct and interpret a 95% confidence interval for *p*.

Direct interpretation:

Never write  $P(\hat{p}_L \le p \le \hat{p}_U) = 0.95$ . Wrong conceptual interpretation.

Correct conceptual interpretation: If many samples were obtained and corresponding intervals calculated, about 95% of the intervals would cover *p*.

Note that the interval is not appropriate for small samples. Such an interval is obtainable, but not in this course.

The margin of error for a CI:

- 1. Increases as the confidence level increases.
- 2. Decreases as the sample size increases.

Ex16.4) Using the data from Ex16.3),

- a) If the confidence level is 99%, what is the new confidence interval?
- b) If n = 2712, what is the new confidence interval (assuming  $\hat{p}$  stays the same)?

Choosing the sample size:

Consider the CI as 
$$\hat{p} \pm ME$$
, where  $ME = z_{\alpha/2} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$ 

Recall that *ME* is the <u>margin of error</u>. The width of the CI is 2*ME*. Now, we still want to see how large a sample size is required; hence, we rearrange to

$$n \approx \hat{p}(1-\hat{p})\left(\frac{z_{\alpha/2}}{ME}\right)^2 = p*(1-p*)\left(\frac{z_{\alpha/2}}{ME}\right)^2$$

Round up n to next integer. Replace  $\hat{p}$  by a prior estimate. If we don't have such information, then how to make n as large as possible? By choosing  $\hat{p} = 0.5$ , we maximize  $\hat{p}(1-\hat{p})$  and get a conservative choice for n. This choice is most common. If, however, we expect  $\hat{p}$  to be close to 0 or 1, say  $\hat{p} \le 0.1$ , then we could set  $\hat{p} = 0.1$  to obtain a smaller n. In this situation, though, we would usually want a smaller ME.

Ex16.5) a) If you wish to conduct a poll so that the margin of error is at most 3 percentage points with 99% confidence, what is the minimum sample size required?

b) How would *n* change if you knew  $\hat{p} \le 0.1$ ?