## Mediation analysis with a binary outcome

Denote by Y an oucome, M a mediator, E an exposure, and C possible confounders (between Y and E, or Y and M, or E and M).

- Y(e, m) = Y(e, m, c) will be the counterfactual outcome, i.e. the outcome value had the exposure be set to e, the mediator to m, and the confounder set to c.
- M(e) = M(e, c) will be the counterfactual mediator, i.e. the mediator value had the exposure be set to e and the confounder set to c.

We define (all at a fixed confounder value c):

- the total effect as  $Y(e, M(e)) Y(e^*, M(e^*))$
- the natural direct effect as  $Y(e, M(e)) Y(e^*, M(e))$
- the indirect direct effect as  $Y(e^*, M(e)) Y(e^*, M(e^*))$

## 1 Theory

We consider the following models:

logit (
$$\mathbb{P}[Y = 1|E, M, C]$$
) =  $\beta_0 + \beta_1 E + \beta_2 M + \beta_3 C$   
 $\mathbb{E}[M|E, C] = \alpha_0 + \alpha_1 E + \alpha_3 C$ 

Therefore the counterfactual probability can be expressed as:

$$Y(e, M(e^*), c) = \frac{1}{1 + \exp^{-\beta_0 - \beta_1 e - \beta_2 M(e^*) + \beta_3 c}}$$

$$= \frac{1}{1 + \exp^{-(\beta_0 + \beta_2 \alpha_0) - (\beta_1 e + \beta_2 \alpha_1 e^*) + (\beta_3 + \beta_2 \alpha_3) c}}$$

$$\log it \left( \mathbb{P} \left[ Y(e, M(e^*), c) \right] \right) = (\beta_0 + \beta_2 \alpha_0) + (\beta_1 e + \beta_2 \alpha_1 e^*) + (\beta_3 + \beta_2 \alpha_3) c$$

So the odd for the total effect is:

$$OR^{TE} = \frac{\mathbb{P}[Y(e, M(e))] / \mathbb{P}[Y(e, M(e))]}{\mathbb{P}[Y(e^*, M(e^*))] / \mathbb{P}[Y(e^*, M(e^*))]} = \exp((\beta_1 + \beta_2 \alpha_1)(e - e^*))$$

for the natural direct effect:

$$OR^{NDE} = \frac{\mathbb{P}[Y(e, M(e))] / \mathbb{P}[Y(e, M(e))]}{\mathbb{P}[Y(e^*, M(e))] / \mathbb{P}[Y(e^*, M(e))]} = \exp(\beta_1(e - e^*))$$

for the natural indirect effect:

$$OR^{NIE} = \frac{\mathbb{P}[Y(e^*, M(e))] / \mathbb{P}[Y(e^*, M(e))]}{\mathbb{P}[Y(e^*, M(e^*))] / \mathbb{P}[Y(e^*, M(e^*))]} = \exp(\beta_2 \alpha_1 (e - e^*))$$

This matches formula in VanderWeele and Vansteelandt (2010)

Note that we can also express the total effect, natural direct effect, and natural indirect effect on the probability scale, e.g.:

$$\begin{split} & \mathbb{P}\left[Y(e, M(e), c) = 1\right] - \mathbb{P}\left[Y(e^*, M(e), c) = 1\right] \\ & = \frac{1}{1 + \exp^{-(\beta_0 + \beta_2 \alpha_0) - (\beta_1 e + \beta_2 \alpha_1 e) + (\beta_3 + \beta_2 \alpha_3)c}} - \frac{1}{1 + \exp^{-(\beta_0 + \beta_2 \alpha_0) - (\beta_1 e^* + \beta_2 \alpha_1 e) + (\beta_3 + \beta_2 \alpha_3)c}} \end{split}$$

which will be a function of the confounder value. To get a single estimate we could average it out over the confounder distribution in our population.

## 2 In •

```
library(medflex)
data(UPBdata)
head(UPBdata)
sum(is.na(UPBdata))
```

```
att attbin attcat
                              negaff initiator gender educ age UPB
1 1.0005617
                 1
                        M 0.8404610
                                        myself
                                                                 1
2 -0.7085889
                 0
                        L -1.2574650
                                                         M 42
                                           both
                                                    Μ
                                                                 0
                 0 L -1.2022564 both
1 M -0.3741277 ex-partner
3 -0.7085889
                                                    F
                                                         H 43
4 0.6061423
                1
                                                    M
                                                         H 52
                                                                 1
5 0.2117230
                 1
                        M 1.9446325 ex-partner
                                                         M 32
                                                                 1
             1
                        H -0.8157964 ex-partner
6 2.0523467
                                                    М
                                                         H 47
                                                                 0
[1] 0
```

Manually

```
e.lm <- glm(negaff ~ factor(attbin) + gender + educ + age, data = UPBdata)
e.logit <- glm(UPB ~ attbin + negaff + gender + educ + age, family =
   binomial("logit"), data = UPBdata)</pre>
```

```
exp(coef(e.logit)["attbin"])
exp(coef(e.logit)["negaff"]*coef(e.lm)["factor(attbin)1"])
```

```
attbin
1.550122
negaff
1.41916
   Using Medflex
expData <- neWeight(negaff \sim factor(attbin) + gender + educ + age, family
   = gaussian, data = UPBdata)
neMod1 <- neModel(UPB \sim attbin0 + attbin1 + gender + educ + age, family =
   binomial("logit"), expData = expData)
exp(cbind(estimate = coef(neMod1),confint(neMod1))[c("attbin01", "attbin11
   "),])
         estimate 95% LCL 95% UCL
attbin01 1.485757 0.946311 2.294548
attbin11 1.421865 1.188970 1.678737
e.lm <- glm(negaff \sim factor(attbin), data = UPBdata)
e.logit <- glm(UPB \sim attbin + negaff, family = binomial("logit"), data =
   UPBdata)
exp(coef(e.logit)["attbin"])
exp(coef(e.logit)["negaff"]*coef(e.lm)["factor(attbin)1"])
 attbin
1.541862
 negaff
1.430096
expData <- neWeight(negaff \sim factor(attbin), family = gaussian, data =
neMod1 \leftarrow neModel(UPB \sim attbin0 + attbin1, family = binomial("logit"),
   expData = expData)
exp(cbind(estimate = coef(neMod1),confint(neMod1))[c("attbin01", "attbin11
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attbin01 1.467847 0.9419685 2.253654
attbin11 1.437643 1.2003462 1.711461
expData <- neImpute(UPB ~ factor(attbin) + negaff + gender + educ + age,
```

```
expData <- neImpute(UPB ~ factor(attbin) + negaff + gender + educ + age,
    family = binomial("logit"), data = UPBdata)
neMod1 <- neModel(UPB ~ attbin0 + attbin1 + gender + educ + age, family =
    binomial("logit"), expData = expData, se = "robust")
exp(cbind(estimate = coef(neMod1),confint(neMod1))[c("attbin01", "attbin11"), ])</pre>
```

```
estimate 95% LCL 95% UCL attbin01 1.494111 0.9833185 2.270238 attbin11 1.405922 1.2006238 1.646326
```

## 3 References

VanderWeele, T. J. and Vansteelandt, S. (2010). Odds ratios for mediation analysis for a dichotomous outcome. *American journal of epidemiology*, 172(12):1339–1348.