

## 18. The t-interval

Chapter 14 introduced the binomial, Poisson, uniform, and normal distributions. In this chapter, we introduce another distribution, the t-distribution. Unlike the distributions of Chapter 14, the t-distribution arises from the need to make accurate statistical inferences, not from any particular kind of data (e.g., successes or failures in a binomial distribution, or events happening over time in a Poisson distribution). In Chapter 17, we calculated confidence intervals (CIs) using the normal distribution and z-scores. In doing so, we made the assumption that the sample standard deviation ( $s$ ) was the same as the population standard deviation ( $\sigma$ ),  $s = \sigma$ . In other words, we assumed that we knew what  $\sigma$  was, which is almost never true. For large enough sample sizes (i.e., high  $N$ ), this is not generally a problem, but for lower sample we need to be careful.

If there is a difference between  $s$  and  $\sigma$ , then our CIs will also be wrong. More specifically, the uncertainty between our sample estimate ( $s$ ) and the true standard deviation ( $\sigma$ ) is expected to increase the deviation of our sample mean ( $\bar{x}$ ) from the true mean ( $\mu$ ). This means that if we are using the *sample* standard deviation instead of the *population* standard deviation (which is pretty much always), then the shape of the standard normal distribution from Chapter 17 (Figure 17.2) will be wrong. The correct shape will be “wider and flatter” (Sokal and Rohlf, 1995), with more probability density at the extremes and less in the middle of the distribution (Box et al., 1978). What this means is that if we use z-scores when calculating CIs using  $s$ , our CIs will not be wide enough, and we will think that we have more confidence in the mean than we really do. Instead of using the standard normal distribution, we need to use a t-distribution<sup>1</sup>.

The difference between the standard normal distribution and t-distribution depends on our sample size,  $N$ . As  $N$  increases, we become more confident that the sample variance will be close to the true population variance (i.e., the deviation of  $s^2$  from  $\sigma^2$  decreases). At low  $N$ , our t-distribution is much wider and flatter than the standard normal distribution. As  $N$  becomes large<sup>2</sup>, the t-distribution becomes basically indistinguishable

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<sup>1</sup>This is also called the “Student’s t-distribution”. It was originally discovered by the head brewer of Guinness in Dublin in the early 20th century (Box et al., 1978). The brewer, W. S. Gosset, published under the pseudonym “A. Student” because Guinness had a policy of not allowing employees to publish (Miller and Miller, 2004).

<sup>2</sup>How large  $N$  needs to be for the t-distribution to be considered close enough to the normal distribution is subjective. The two distributions get closer and closer as  $N \rightarrow \infty$ , but Sokal and Rohlf (1995) suggest that they are indistinguishable for all intents and purposes once  $N > 30$ . It is always safe to use the t-distribution when calculating confidence intervals, which is what all statistical programs such as Jamovi or R will do by default, so there is no need to worry about these kinds of arbitrary cutoffs in this case.

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from the standard normal distribution. For calculating CIs from a sample, especially for small sample sizes, it is therefore best to use *t*-scores instead of *z*-scores. The idea is the same; we are just multiplying the standard errors by a different constant to get our CIs. For example, in [Chapter 17](#), we multiplied the standard error of 20 cat masses by  $z = 1.96$  because 95% of the probability density lies between  $z = -1.96$  and  $z = 1.96$  in the standard normal distribution. In truth, we should have multiplied by -2.093 because we only had a sample size of  $N = 20$ . Figure 18.1 shows the difference between the standard normal distribution and the more appropriate *t*-distribution<sup>3</sup>.

Note that in Figure 18.1, a *t*-distribution with 19 degrees of freedom (*df*) is shown. The *t*-distribution is parameterised using *df*, and we lose a degree of freedom when calculating  $s^2$  from a sample size of  $N = 20$ , so  $df = 20 - 1 = 19$  is the correct value (see [Chapter 12](#) for a brief explanation). For calculating CIs, *df* will always be  $N - 1$ , and this will be taken care of automatically in statistical programs such as Jamovi and R<sup>4</sup>.

Recall from [Chapter 17](#) that our body mass measurements of 20 cats had a sample mean of  $\bar{x} = 4.1$  kg and sample standard deviation of  $s = 0.6$  kg. We calculated the lower 95% CI to be  $LCI_{95\%} = 4.041$  and the upper 95% CI to be  $UCI_{95\%} = 4.159$ . We can now repeat the calculation using the *t*-score 2.093 instead of the *z*-score 1.96. Our corrected lower 95% CI is,

$$LCI_{95\%} = 4.1 - \left( 2.093 \times \frac{0.6}{20} \right) = 4.037$$

Our upper 95% confidence interval is,

$$UCI_{95\%} = 4.1 + \left( 2.093 \times \frac{0.6}{20} \right) = 4.163.$$

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<sup>3</sup>We can define the *t*-distribution mathematically ([Miller and Miller, 2004](#)), but it is an absolute beast,

$$f(t) = \frac{\Gamma\left(\frac{v+1}{2}\right)}{\sqrt{\pi v} \Gamma\left(\frac{v}{2}\right)} \left(1 + \frac{t^2}{v}\right)^{-\frac{v+1}{2}}.$$

In this equation,  $v$  is the degrees of freedom. The  $\Gamma()$  is called a “gamma function”, which is basically the equivalent of a factorial function, but for any number  $z$  (not just integers), such that  $\Gamma(z + 1) = \int_0^\infty x^z e^{-x} dx$  (where  $z > -1$ , or, even more technically, the real part of  $z > -1$ ). If  $z$  is an integer  $n$ , then  $\Gamma(n + 1) = n!$  ([Borowski and Borwein, 2005](#)). What about the rest of the *t* probability density function? Why is it all so much? The reason is that it is the result of 2 different probability distributions affecting *t* independently, a standard normal distribution and a Chi-square distribution ([Miller and Miller, 2004](#)). We will look at the Chi-square in [Week 9](#). Suffice to say that underlying mathematics of the *t*-distribution is not important for our purposes in applying statistical techniques.

<sup>4</sup>Another interesting caveat, which Jamovi and R will take care of automatically (so we do not actually have to worry about it), is that when we calculate  $s^2$  to map *t*-scores to probability densities in the *t*-distribution, we multiply the sum of squares by  $1/N$  instead of  $1/(N - 1)$  ([Sokal and Rohlf, 1995](#)). In other words, we no longer need to correct the sample variance  $s^2$  to account for bias in estimating  $\sigma^2$  because the *t*-distribution takes care of this for us.

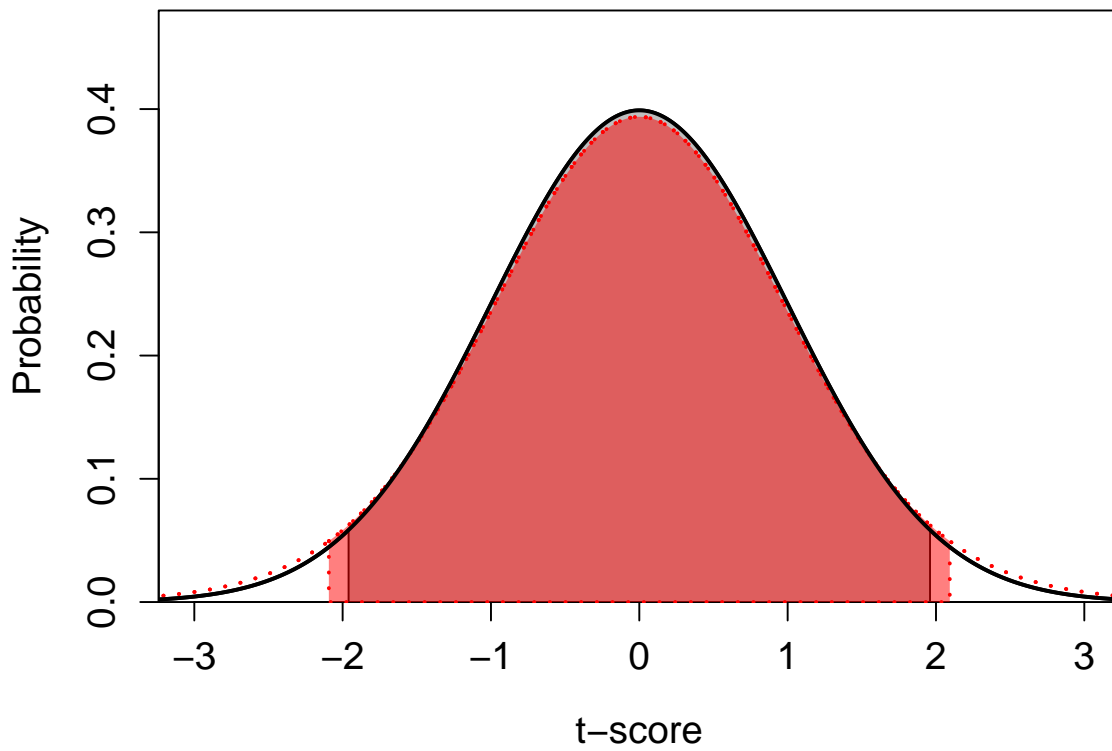


Figure 18.1.: A standard normal probability distribution showing 95 per cent of probability density surrounding the mean (grey). On top of the standard normal distribution in grey, red dotted lines show a t-distribution with 19 degrees of freedom. Red shading shows 95 per cent of the probability density of the t-distribution.

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The confidence intervals have not changed too much. By using the *t*-distribution, the LCI changed from 4.041 to 4.037, and the UCI changed from 4.159 to 4.163. In other words, we only needed our CIs to be a bit wider ( $4.163 - 4.037 = 0.126$  for the using *t*-scores versus  $4.159 - 4.041 = 0.118$  using *z*-scores). This is because a sample size of 20 is already large enough for the *t*-distribution and standard normal distribution to be very similar (Figure 18.1). But for lower sample sizes ( $N$ ) and therefore fewer degrees of freedom ( $df = N - 1$ ), the difference between the shapes of these distributions gets more obvious (Figure 18.2).

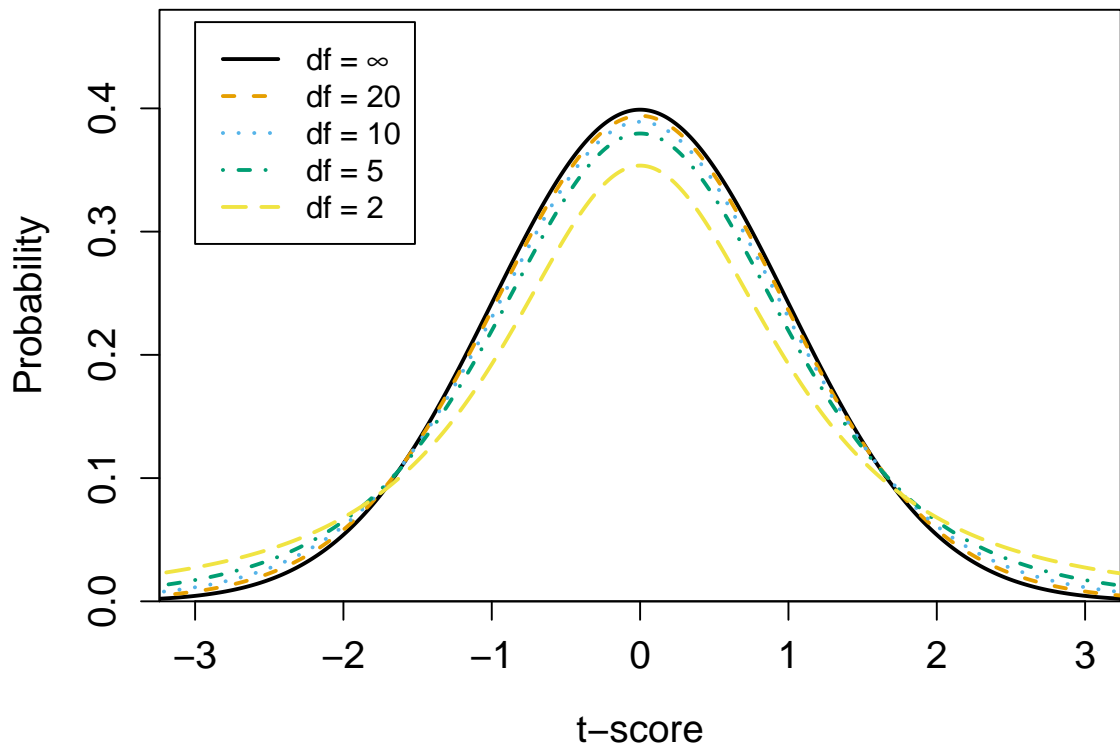


Figure 18.2.: A *t*-distribution with infinite degrees of freedom (*df*) is shown in black; this distribution is identical to the standard normal distribution. Other *t*-distributions with the same mean and standard deviation, but different degrees of freedom, are indicated by curves of different colours and line types.

The main point of Figure 18.2 is that as degrees of freedom decreases, the *t*-distribution becomes wider, with more probability density in the tails. Figure 18.2 is quite busy, so we have made an [interactive application](#) to make visualising the *t*-distribution easier.

[Click here](#) for an interactive application to visualise *t*-scores

Note that *t*-scores do not need to be used when making binomial confidence intervals. Using *z*-scores is fine.

The t-distribution is important throughout most of the rest of this module. It is not just used for calculating confidence intervals. The t-distribution also plays a critical role in hypothesis-testing, which is the subject of [Week 6](#) and applied throughout the rest of the book. The t-distribution is therefore very important for understanding most of the statistical techniques presented in this book.