University of Oslo

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Project 1

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https://github.com/bragewiseth/MachineLearningProjects

Contents

Abstract

In this project, we explored the use of different methods for solving linear regres- sion on topological data. We emplyed OLS, ridge and lasso methods for linear regression on simulated height data using *Franke's Function*, Where the goal was to fit polynomi als to minimize the mean square error. As well as real topological data with the same goal.

Keywords: Linear Regression, Fitting Polynomials, Scaling, Bias-Variance

1. Introduction

To get a smooth start on this project we begin to construct some fake topological data. This can be constructed with *Franke's Function*. This gives us a nice dataset to start linear regression. A natrual start would be to use ordinary least squares to try to fit a polynomial to some degree. This is done by constructing a design matrix X with the polynomial terms. Then we can use the closed form solution to find the coefficients β. This is done by minimizing the mean square error (MSE) between the data and the model.

$$\beta = (X^T X)^{-1} X^T \mathbf{y} \tag{1}$$

1. Introduction

Briefly introduce the problem or task you are addressing. Mention the importance of linear regression models and scaling techniques in machine learning. State the purpose of the report, i.e., to explore OLS, Ridge, and Lasso regression models, and their impact on bias-variance trade-off through scaling.

2. Data

Describe the dataset you are using for your analysis. Include information about the features, target variable, and any preprocessing steps applied (e.g., data cleaning).

4. Linear Regression Models 4.1 Ordinary Least Squares (OLS)

Explain the concept of OLS regression. Discuss its advantages and limitations. Present the results of applying OLS to your dataset.

4.2 Ridge Regression

Explain the concept of Ridge regression. Discuss how it addresses overfitting and its use of L2 regularization. Present the results of applying Ridge regression to your dataset.

4.3 Lasso Regression

Explain the concept of Lasso regression. Discuss how it addresses overfitting and its use of L1 regularization. Present the results of applying Lasso regression to your dataset.

5. Scaling Techniques

Discuss the importance of feature scaling in machine learning. Explain common scaling techniques such as Standardization (Z-score scaling) and Min-Max scaling. Apply these scaling techniques to your dataset.

6. Bias-Variance Trade-Off

Define bias and variance in the context of machine learning models. Discuss how regularization (Ridge and Lasso) impacts the bias-variance trade-off. Present the results of your analysis regarding bias and variance for the different regression models.

7. Results and Discussion

Summarize the performance metrics of OLS, Ridge, and Lasso models. Discuss any observed trends or differences in model performance. Analyze the impact of scaling on model performance.

8. Conclusion

Summarize the key findings of your analysis. Discuss the implications of your results for the problem/task you introduced in the introduction.

2. Data

3. Linear Regression Models

$$MSE(\mathbf{y}, \tilde{\mathbf{y}}) = \frac{1}{n} \sum_{i=0}^{n-1} (y_i - \tilde{y}_i)^2,$$

$$R^2(\mathbf{y}, \tilde{\mathbf{y}}) = 1 - \frac{\sum_{i=0}^{n-1} (y_i - \tilde{y}_i)^2}{\sum_{i=0}^{n-1} (y_i - \bar{y})^2},$$

$$\bar{y} = \frac{1}{n} \sum_{i=0}^{n-1} y_i.$$

3.1 Ordinary Least Squares (OLS)

3.2 Ridge Regression

3.3 Lasso Regression

4. Results and Discussion

As we increase the order of the polynomial fit we see a corresponding decrease in MSE and increase in R2.

We see that OLS remains constant since it is not a function of . For the MSE from the training set we se that the MSE wil increase for larger values of for rigde regression. This is expected since the OLS wil give the best approximation for the data it has been trained on (compared to Rigde), while rigde punishes the weights from becomming too large and therefore gives a not so tight fit. This can explain why rigde does better than OLS for some values of . As stated before the OLS wil try to fit the model to the best of its ability to the training data, this includes the noise which might not generalize to the test set, hence giving rigde a greater score. The term for this is called ***overfitting***. And balancing between overfitting and underfitting is called the ***bias variance tradeoff***.

4.1 Confidence Interval

The assumption we have made is that there exists a continuous function f (x) and a normal distributed error N (0, 2) which describes our data y = f (x) + We then approximate this function f (x) with our model ˜y from the solution of the linear regression equations (ordinary least squares OLS), that is our function f is approximated by ˜y where we minimized (y − ˜y)2, with ˜y = X . The matrix X is the so-called design or feature matrix. yi N (Xi,∗ , 2), that is y follows a normal distribution with mean value X and variance 2. We can use this when we define a so-called confidence interval for the parameters . A given parameter j is given by the diagonal matrix element of the above matrix. 6

5. The Topic of Scaling

Table 1: Unscaled sample design matrix fitting one-dimensional polynomial of degree 5

1.	0.	0.	0.	0.	0.
1.	0.25	0.0625	0.01562	0.00391	0.00098
1.	0.5	0.25	0.125	0.0625	0.03125
1.	0.75	0.5625	0.42188	0.31641	0.2373
1.	1.	1.	1.	1.	1.

Table 2: Scaled sample design matrix fitting one-dimensional polynomial of degree 5

0.	-1.41421	-1.0171	-0.83189	-0.728	-0.66226
0.	-0.70711	-0.84758	-0.7903	-0.71772	-0.65971
0.	0.	-0.33903	-0.49913	-0.56348	-0.58075
0.	0.70711	0.50855	0.29116	0.10488	-0.0433
0.	1.41421	1.69516	1.83016	1.90431	1.94603

the code for generating this output [1]

MSE for OLS on unscaled data:	0.010349396022903145
MSE for OLS on scaled data:	0.010349396024145656
MSE for Ridge on unscaled data:	0.02106077418650843
MSE for Ridge on scaled data:	0.01782525371566323

the code for generating this output

First x_{unscaled} and x_{scaled} is not that different, the original data was close to zero-centered and not that spread out, which means that initially by just looking at the data scaling is not that necessary. When we add the polynomial terms we can now see that some of the entries of X_{unscaled} get really small as an example $0.1^5 = 0.00001$ this makes the columns of X_{unscaled} live in their own order of magnitude and scaling should be considered to bring them back to the same order of magnitude. The act of not scaling results in β spanning from -50 to 48 while scaling gives a smaller span from -10 to 13. Now this alone may not justify why we should scale this data set, as scaled and unscaled OLS yields the same MSE. However when doing ridge regression the cost function is directly dependent on the magnitude of β_i. Now with each β_i varying alot, some are getting more penalized than others. As we can see the MSE for unscaled data is much higher than for scaled data in the ridge case. This leads us to conclude that we should scale the data, making it easier to tweak λ and giving us nicer numbers to work with

6. Bias & Variance

$$\text{Var}(\tilde{\mathbf{y}}) = \mathbb{E}\left[(\tilde{\mathbf{y}} - \mathbb{E}[\tilde{\mathbf{y}}])^2\right] = \mathbb{E}[\tilde{\mathbf{y}}^2] - \mathbb{E}[\tilde{\mathbf{y}}]^2$$
$$\mathbb{E}[\tilde{\mathbf{y}}^2] = \mathbb{E}[\tilde{\mathbf{y}}]^2 + \text{Var}(\tilde{\mathbf{y}})$$

$$\begin{aligned}\mathbb{E}[\mathbf{y}^2] &= \mathbb{E}[\mathbf{f}+\epsilon]^2 = \mathbb{E}[\mathbf{f}^2 + 2\mathbf{f}\epsilon + \epsilon^2] \\ &= \mathbb{E}[\mathbf{f}^2] + 2\mathbb{E}[\mathbf{f}\epsilon] + \mathbb{E}[\epsilon^2] \\ &= \mathbb{E}[\mathbf{f}^2] + 2\mathbb{E}[\mathbf{f}]\mathbb{E}[\epsilon] + \mathbb{E}[\epsilon^2] \\ &= \mathbb{E}[\mathbf{f}^2] + \sigma^2\end{aligned}$$

$$\begin{aligned}\mathbb{E}[\mathbf{y}\tilde{\mathbf{y}}] &= \mathbb{E}[\mathbf{f}\tilde{\mathbf{y}} + \epsilon\tilde{\mathbf{y}}] \\ &= \mathbb{E}[\mathbf{f}\tilde{\mathbf{y}}] + \mathbb{E}[\epsilon\tilde{\mathbf{y}}] \\ &= \mathbb{E}[\mathbf{f}\tilde{\mathbf{y}}] + \mathbb{E}[\epsilon]\mathbb{E}[\tilde{\mathbf{y}}] \\ &= \mathbf{f}\mathbb{E}[\tilde{\mathbf{y}}]\end{aligned}$$

$$\begin{aligned}\mathbb{E}\left[(\mathbf{y}-\tilde{\mathbf{y}})^2\right] &= \mathbb{E}[\mathbf{y}^2] - 2\mathbb{E}[\mathbf{y}\tilde{\mathbf{y}}] + \mathbb{E}[\tilde{\mathbf{y}}^2] \\ &= \mathbf{f}^2 + \sigma^2 - 2\mathbf{f}\mathbb{E}[\tilde{\mathbf{y}}] + \mathbb{E}[\tilde{\mathbf{y}}^2] \\ &= \mathbf{f}^2 + \sigma^2 - 2\mathbf{f}\mathbb{E}[\tilde{\mathbf{y}}] + \mathbb{E}[\tilde{\mathbf{y}}]^2 + \mathrm{Var}(\tilde{\mathbf{y}}) \\ &= (\mathbf{f} - \mathbb{E}[\tilde{\mathbf{y}}])^2 + \mathrm{Var}(\tilde{\mathbf{y}}) + \sigma^2\end{aligned}$$

$$\begin{aligned}\mathrm{Bias}[\tilde{y}] = \mathbb{E}\left[(\mathbf{y} - \mathbb{E}\left[\tilde{\mathbf{y}}\right])^2\right] &= \mathbb{E}[\mathbf{y}^2] - 2\mathbb{E}[\mathbf{y}]\mathbb{E}[\tilde{\mathbf{y}}] + \mathbb{E}[\tilde{\mathbf{y}}^2] \\ &= \mathbf{f}^2 + \sigma^2 - 2\mathbf{f}\mathbb{E}[\tilde{\mathbf{y}}] + \mathbb{E}[\tilde{\mathbf{y}}]^2 \\ &= (\mathbf{f} - \mathbb{E}[\tilde{\mathbf{y}}])^2 + \sigma^2\end{aligned}$$

To find $\mathbb{E}[\tilde{\mathbf{y}}]$ we can use bootstrap to get different samples for $\tilde{\mathbf{y}}$ and then take the mean of that distribution. The term *variance* refers to how spread out our observations are. We can for our case think of the observations $\tilde{y}_i$ as either individual predictions, or the mean of predictions from one test set.
In a sense the bias resembles the mathematical defninition for variance, We can see that the bias term is a distance measure of $\mathbf{y}$ from the expected value of $\tilde{\mathbf{y}}$. If we take the individual observations approach this can be tough of as A high bias will then result in a large band around $\tilde{\mathbf{y}}$ in which y_i will exist. The variance term is a measure of how much the $\tilde{\mathbf{y}}$ varies around its own mean. In other terms a high variance model will experience large variance in the predicted values $\tilde{\mathbf{y}}$ from one test set to another. Or as a result of this a large variance in score if you will. Now these two terms are in a sense opposites, it is likely that a model with high bias will have low variance and vice versa. This is something we as designers of the model get to tune, if we value a lower bias more than we value a low variance or vice versa, we can choose a model accordingly

Appendix

SVD

svd

adding these two we get

$$\frac{\partial \left[\frac{1}{n}||\mathbf{y}-\mathbf{X}\beta||_2^2+\lambda||\beta||_2^2\right]}{\partial\beta} = 0 = \frac{2}{n}\left(\mathbf{X}^T\mathbf{X}+\lambda\mathbf{I}\right)\beta - \frac{2}{n}\mathbf{X}^T\mathbf{y} \implies \beta = \left(\mathbf{X}^T\mathbf{X}+\lambda\mathbf{I}\right)^{-1}\mathbf{X}^T\mathbf{y}$$

The minimization of ridge addition alone is

$$\min_{\beta\in\mathbb{R}^p}\frac{1}{n}\lambda||\beta||_2^2 = \frac{2}{n}\lambda\beta$$

We know that

$$\min_{\beta\in\mathbb{R}^p}\frac{1}{n}||\mathbf{y}-\mathbf{X}\beta||_2^2 = \frac{2}{n}\mathbf{X}^T\mathbf{X}\beta - \frac{2}{n}\mathbf{X}^T\mathbf{y}$$

$$\min_{\beta\in\mathbb{R}^p}\frac{1}{n}\sum_{i=0}^{n-1}(y_i-\tilde{y}_i)^2 = \frac{1}{n}||\mathbf{y}-\mathbf{X}\beta||_2^2,$$

$$\min_{\beta\in\mathbb{R}^p}\frac{1}{n}||\mathbf{y}-\mathbf{X}\beta||_2^2+\lambda||\beta||_2^2$$

$$\hat{\beta}_{\rm OLS} = \left(\mathbf{X}^T\mathbf{X}\right)^{-1}\mathbf{X}^T\mathbf{y},$$

$$\hat{\beta}_{\rm Ridge} = \left(\mathbf{X}^T\mathbf{X}+\lambda\mathbf{I}\right)^{-1}\mathbf{X}^T\mathbf{y},$$

Use the singular value decomposition of an $n\times p$ matrix $\mathbf{X}$ (our design matrix)

$$\mathbf{X} = \mathbf{U}\mathbf{\Sigma}\mathbf{V}^T,$$

where $\mathbf{U}$ and $\mathbf{V}$ are orthogonal matrices of dimensions $n\times n$ and $p\times p$, respectively, and $\mathbf{\Sigma}$ is an $n\times p$ matrix which contains the ingular values only. This material was discussed during the lectures of week 35.

Show that you can write the OLS solutions in terms of the eigenvectors (the columns) of the orthogonal matrix $\mathbf{U}$ as

$$\tilde{\mathbf{y}}_{\rm OLS} = \mathbf{X}\beta = \sum_{j=0}^{p-1}\mathbf{u}_j\mathbf{u}_j^T\mathbf{y}.$$

$$\mathbf{X}^T\mathbf{X} = \mathbf{V}\mathbf{\Sigma}^T\mathbf{U}^T\mathbf{U}\mathbf{\Sigma}\mathbf{V}^T = \mathbf{V}\mathbf{\Sigma}^T\mathbf{\Sigma}\mathbf{V}^T = \mathbf{V}\mathbf{\Sigma}^2\mathbf{V}^T$$

$$\tilde{\mathbf{y}}_{\rm OLS} = \mathbf{X}\beta = \mathbf{X}\left(\mathbf{X}^T\mathbf{X}\right)^{-1}\mathbf{X}^T\mathbf{y} = \mathbf{U}\mathbf{\Sigma}\mathbf{V}^T\left(\mathbf{V}\mathbf{\Sigma}^2\mathbf{V}^T\right)^{-1}\mathbf{V}\mathbf{\Sigma}^T\mathbf{U}^T\mathbf{y} = \mathbf{U}\mathbf{\Sigma}\mathbf{V}^T\mathbf{V}^{T^{-1}}\mathbf{\Sigma}^{2^{-1}}\mathbf{V}^{-1}\mathbf{V}\mathbf{\Sigma}^T\mathbf{U}^T\mathbf{y}$$

using the orthogonality of $\mathbf{V}$ and $\mathbf{U}$ we get. Multiplying with $\mathbf{\Sigma}$ removes columns from $\mathbf{U}$ with eigenvalues equal to zero.

$$\tilde{\mathbf{y}}_{\text{OLS}} = \mathbf{U}\mathbf{U}^T\mathbf{y} = \sum_{j=0}^{p-1} \mathbf{u}_j \mathbf{u}_j^T \mathbf{y}$$

$$\mathbf{y} = f(\mathbf{x}) + \varepsilon$$

$$\tilde{\mathbf{y}} = \mathbf{X}\beta.$$

$$\mathbb{E}(y_i) = \sum_j x_{ij}\beta_j = \mathbf{X}_{i,*}\,\beta$$

$$Var(y_i) = \sigma^2$$
$$Var(\hat{\beta}) = \sigma^2\,(\mathbf{X}^T\mathbf{X})^{-1}.$$

Math Behind the Confidence Interval section

We can assume that y follows some function f with some noise ϵ

$$\mathbf{y} = f(\mathbf{x}) + \epsilon$$
$$\mathbb{E}[\mathbf{y}] = \mathbb{E}[f(\mathbf{x}) + \epsilon] = \mathbb{E}[f(\mathbf{x})] + \mathbb{E}[\epsilon_i]$$

The expected value of ϵ_i is 0, $f(x)$ is a non-stochastic variable and is approimated by $\mathbf{X}\beta$

$$\mathbb{E}[y_i] = \mathbf{X_{i,*}}\beta$$

The variance is defined as

$$\begin{aligned}
Var(y_i) &= \mathbb{E}\big[(y_i - \mathbb{E}[y_i])^2\big] = \mathbb{E}\left[y_i^2 - 2y_i\mathbb{E}[y_i] + \mathbb{E}[y_i]^2\right] = \mathbb{E}[y_i^2] - \mathbb{E}[y_i]^2 \\
Var(y_i) &= \mathbb{E}\big[\left(\mathbf{X}_{i,*}\beta\right)^2 + 2\epsilon\mathbf{X}_{i,*}\beta + \epsilon^2\big] - (\mathbf{X}_{i,*}\beta)^2 \\
Var(y_i) &= (\mathbf{X}_{i,*}\beta)^2 + 2\mathbb{E}[\epsilon]\mathbf{X}_{i,*}\beta + \mathbb{E}[\epsilon^2] - (\mathbf{X}_{i,*}\beta)^2 \\
Var(y_i) &= \mathbb{E}[\epsilon^2] = \sigma^2
\end{aligned}$$

for the expected value of β we can insert the definition of β from earlier

$$\mathbb{E}[\hat{\beta}] = \mathbb{E}[(\mathbf{X}^\top\mathbf{X})^{-1}\mathbf{X}^T\mathbf{Y}] = (\mathbf{X}^T\mathbf{X})^{-1}\mathbf{X}^T\mathbb{E}[\mathbf{Y}] = (\mathbf{X}^T\mathbf{X})^{-1}\mathbf{X}^T\mathbf{X}\beta = \beta$$

$$\begin{aligned}
Var(\hat{\beta}) &= \mathbb{E}\{[\beta - \mathbb{E}(\beta)][\beta - \mathbb{E}(\beta)]^T\} \\
&= \mathbb{E}\{[(\mathbf{X}^T\mathbf{X})^{-1}\,\mathbf{X}^T\mathbf{y} - \beta]\,[(\mathbf{X}^T\mathbf{X})^{-1}\,\mathbf{X}^T\mathbf{y} - \beta]^T\} \\
&= (\mathbf{X}^T\mathbf{X})^{-1}\,\mathbf{X}^T\,\mathbb{E}\{\mathbf{y}\,\mathbf{y}^T\}\,\mathbf{X}\,(\mathbf{X}^T\mathbf{X})^{-1} - \beta\,\beta^T \\
&= (\mathbf{X}^T\mathbf{X})^{-1}\,\mathbf{X}^T\,\{\mathbf{X}\,\beta\,\beta^T\,\mathbf{X}^T + \sigma^2\}\,\mathbf{X}\,(\mathbf{X}^T\mathbf{X})^{-1} - \beta\,\beta^T \\
&= \beta\,\beta^T + \sigma^2\,(\mathbf{X}^T\mathbf{X})^{-1} - \beta\,\beta^T \;\; = \;\; \sigma^2\,(\mathbf{X}^T\mathbf{X})^{-1},
\end{aligned}$$

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