

9.3

a) $\theta = P\left(\sum_{i=1}^5 iX_i \geq 21.6\right), X_i \sim \text{exp}(1)$

so $E[\theta] = \frac{1}{n} \sum_{k=1}^n \mathbb{1}\left(\sum_{i=1}^5 i^{(k)} X_i^{(k)} \geq 21.6\right) \quad i^{(k)} = i \quad \forall k$

to estimate θ we do the following

- n times {
- 1) $X = []$
 - 2) $U_i \sim \text{unif}(0,1)$ generation for $i \in \{1,2,3,4,5\}$
 - 3) if ~~if~~ $\sum_{i=1}^5 -i \log(U_i) \geq 21.6$, $X.append(1)$
 - else $X.append(0)$
 - 4) return $\frac{\text{sum}(X)}{\text{len}(X)}$

b) to estimate θ using antithetic variables we do the following

- n times {
- 1) $X = []$
 - 2) generate $U_i \sim \text{unif}(0,1)$ for i
 - 3) set $X_i^{(1)} = -\log(U_i), X_i^{(2)} = -\log(1-U_i)$
 - 4) if $\sum_{i=1}^5 i X_i^{(1)} \geq 21.6$ or $\sum_{i=1}^5 i X_i^{(2)} \geq 21.6$, $X.append(1)$
 - else $X.append(0)$
 - 5) return $\frac{\text{sum}(X)}{\text{len}(X)}$

c) This is not efficient as variance is reduced by only a factor of ~ 1.006 after 10K runs of each estimator generating 1k samples.

this implements antithetic estimator

$$E[\theta] = \frac{1}{2} \sum_{k=1}^n \mathbb{1}\left(\sum_{i=1}^5 i^{(k)} X_i^{(1),k} \geq 21.6\right) + \mathbb{1}\left(\sum_{i=1}^5 i^{(k)} X_i^{(2),k} \geq 21.6\right)$$

c.

Raw E: 0.038569, Antithetic E: 0.038298

Raw Var: 3.7e-05

Antithetic Var: 1.8e-05, Reduction over Raw: 2.0634

The use of antithetic variables is effective.

9.10

a. First note $E[I]$, $P(X \leq a) = F(a)$, $F = P(X \leq x)$

$$E[XI] = \int_0^a x f(x) dx$$

$$\begin{aligned} \text{So } \text{cov}(X, I) &= E[XI] - E[X]E[I] \\ &= \int_0^a x f(x) dx - F(a) \int_0^{\infty} x f(x) dx \end{aligned}$$

$$\text{and } \text{var}(X) = E[X^2] - E[X]^2 = \int x^2 f(x) dx - \left(\int x f(x) dx \right)^2$$

$$\begin{aligned} \text{var}(I) &= E[I^2] - E[I]^2 = E[I] (1 - E[I]) \\ &= F(a)(1 - F(a)) \end{aligned}$$

Then b/c % reduction of var :

$$= \frac{\text{var}(I) + c^2 \text{var}(X) + \text{cov}(I, X) \cancel{\text{var}(I)}}{\text{var}(I)}$$

$$= \frac{\text{var}(I)(1 - \text{corr}^2(I, X)) \cancel{\text{var}(I)}}{\text{var}(I)} = 1 - \text{corr}^2(I, X) = \frac{\text{cov}^2(I, X)}{\text{var}(I)\text{var}(X)}$$

So if $X \sim \text{unif}(0, 1)$, $F(x) = x$, $f(x) = 1$,

$$\text{cov}(X, I) = \int_0^a x dx - a \int_0^1 x dx = \frac{a^2 - a}{2}$$

$$\text{var}(X) = \int_0^1 x^2 dx - \left(\int_0^1 x dx \right)^2 = \frac{1}{3} - \frac{1}{4} = \frac{1}{12}$$

$$\text{var}(I) = a(1-a)$$

$$\text{thus } \text{corr}^2(X, I) = 3a(1-a) \Rightarrow \% \text{ reduction} = 1 - 3a(1-a)$$

9.10. cont.

b. if $X \sim \text{exp}(1)$ then $F(x) = 1 - e^{-x}$, $f(x) = e^{-x}$

$$\begin{aligned} \text{then } \text{cov}(I, X) &= \int_0^a x e^{-x} dx - (1 - e^{-a}) \int_0^{\infty} x e^{-x} dx \\ &= -a e^{-a} - \int_0^{\infty} -e^{-x} dx - (1 - e^{-a}) \left(x e^{-x} - \int_0^{\infty} -e^{-x} dx \right) \\ &= -a e^{-a} - e^{-a} + 1 - 1 + e^{-a} \\ &= -a e^{-a} \end{aligned}$$

$$\text{var}(X) = \frac{1}{\lambda^2} = 1$$

$$\begin{aligned} \text{var}(I) &= F(a)(1 - F(a)) = 1 - e^{-a}(1 - (1 - e^{-a})) \\ &= e^{-a}(1 - e^{-a}) \end{aligned}$$

and the $\frac{1}{2}$ reduction in var \Rightarrow $1 - \text{corr}^2(I, X) = \frac{\text{var}(I) + \text{var}(X) - \text{var}(I+X)}{\text{var}(I) \text{var}(X)}$

$$= \frac{e^{-a}(1 - e^{-a}) + 1 - (1 - e^{-a})^2}{e^{-a}(1 - e^{-a})} = \frac{1 - e^{-a}}{e^{-a}(1 - e^{-a})} = 1 - \frac{(-a e^{-a})^2}{e^{-a}(1 - e^{-a})}$$

c. Since $1 - \text{corr}^2(I, X) = \frac{\text{cov}^2(I, X)}{\text{var}(I) \text{var}(X)}$ all we need to show is $\text{cov}(I, X) < 0$

$$\begin{aligned} \text{cov}(I, X) &= E[XI] - E[X]E[I] = 1 - F(a) \int_0^a x f(x) dx - F(a) \int_0^{\infty} x f(x) dx \\ &= -(1 - F(a)) \int_0^a F(x) dx + a F(a)(1 - F(a)) - F(a) \int_a^{\infty} x f(x) dx \\ &= -(1 - F(a)) \int_0^a F(x) dx + a F(a)(1 - F(a)) - (1 - F(a)) \bar{X} F(a), \quad \bar{X} = E[X|F] \\ &= -(1 - F(a)) \int_0^a F(x) dx + (\bar{X} - a) F(a) \end{aligned}$$

then b/c $F(a) < 1$, $F(x) \geq 0$, $\bar{X} \geq a$ we must conclude

$$\text{cov}(I, X) < 0 \Rightarrow \text{corr}(I, X) = \frac{\text{cov}(I, X)}{\sqrt{\text{var}(I) \text{var}(X)}} < 0$$

9.18

$$y \sim N(1, 1), X \sim N(y, 4), y = Y$$

$$\text{so } X_{y=y} \sim N(y, 4), \theta = P(X > 1)$$

a) raw estimator $E[\theta] = \frac{1}{n} \sum_{k=1}^n \mathbb{1}(X_k > 1)$ we do the following

1) $X = []$

2) Gen $y = Y \sim N(1, 1)$

3) Gen $X \sim N(y, 4)$

4) if $X < 1$, $X.append(1)$

else, $X.append(0)$

5) return $\frac{\text{sum}(X)}{\text{len}(X)}$

} n times

$$= P_{y=y}(X > 1 | y)$$

b) First note $P(X > 1 | y = y_i) = P\left(\frac{X - y_i}{2} > \frac{1 - y_i}{2} | y_i\right) = P\left(\frac{X - y_i}{2} > z_{y_i} | y_i\right)$

so the conditional estimator is

$$E[\theta_c] = \frac{1}{n} \sum_{k=1}^n \mathbb{1}\left(\frac{X_k - y_k}{2} > z_{y_k}\right) = \frac{1}{n} \sum_{k=1}^n A_{y_k}$$

and do the following to simulate it

1) $X = []$

2) Gen $y = Y \sim N(1, 1)$

3) find z_y (z-score of y) = $\frac{y-1}{1}$

4) find $A_{z_y} = P(X > y)$

5) $X.append(A_{z_y})$

6) average over X

9.18. cont.

c) using antithetic variables and conditioning we get an estimator of

$$E[\theta_{c,a}] = \frac{1}{2} \sum_{i=1}^n \mathbb{1}\left(\frac{X_i - Y_i^{(u)}}{2} > z_{y^{(u)}}^{(u)}\right) + \mathbb{1}\left(\frac{X_i - Y_i^{(u)}}{2} > z_{y^{(u)}}^{(u)}\right)$$

which we simulate by the following

- n times {
- 1) $X = []$
 - 2) Gen $y_1 \sim N(1, \frac{1}{2}), y_2 \sim N(0, \frac{1}{2})$
 - 3) set $A = y_1 + y_2, B = y_1 - y_2$
 - 4) find z_A, z_B (respective z-scores)
 - 5) find $A_{z_A} = P(X > z_A), B_{z_B} = P(X > z_B)$
 - 6) $X.append(A_{z_A} + B_{z_B})$
 - 7) average over X
 - 8) return $\frac{avg(X)}{2}$

Note that for

$$y_1 \sim N(1, \frac{1}{2}), y_2 \sim N(0, \frac{1}{2})$$

that $E[y_1 + y_2] = 1, \text{cov}(y_1 + y_2) < 0$

$$E[y_1 - y_2] = 1, \text{cov}(y_1 - y_2) < 0$$

d) we can use a control variate of Y so $E[Y] = 1$ and our new estimator is

$$E[\theta_{c,c}] = \sum_{i=1}^n A_{z_{y_n}} - \gamma(y_n - 1) \text{ where } \gamma = \text{cov}(z_{y_n}, Y)$$

we find γ empirically and simulate the entire process by

- n times {
- 1) $X = []$
 - 2) Gen $y_n \sim N(1, 1)$
 - 3) find z_y
 - 4) find $A_{z_{y_n}}$
 - 5) $X.append(A_{z_{y_n}} - \gamma(y_n - 1))$
 - 6) $X.append(A_{z_{y_n}} - \gamma(y_n - 1))$
 - 7) return $avg(X)$
- a) $y_{cov} = [], z_{cov} = []$
 b) $y_{cov}.append(y_n \sim N(1, 1))$
 c) $z_{cov}.append(A_{z_{y_n}})$
 d) $\gamma = \text{cov}(y_{cov}, z_{cov})$

e) f) g) h)

Raw E: 0.50137, Conditional E: 0.498847

Conditional + Antithetic Var E: 0.499851, Conditional + Control Var E: 0.500073

Raw Var: 0.250001

Conditional Var: 0.083237, Reduction over Raw: 3.0035

Conditional + Antithetic Var: 0.041684, Reduction over Conditional: 5.9975

Conditional + Control Var: 0.003683, Reduction over Conditional: 67.8849

i) Exact value of $\theta = 0.5$

9.19

a) $U \sim \text{unif}(0,1)$, $C \sim \text{poisson}(\frac{15}{.5+U})$

$I = \begin{cases} 1 & \text{if } C \geq 20 \\ 0 & \text{otherwise} \end{cases}$ $P = E[I] = \frac{1}{n} \sum_{i=1}^n \mathbb{1}(C_i \geq 20)$ is a raw estimator

So to simulate the estimator we do the following

- 1) $X = []$
- 2) Gen $U \sim \text{unif}(0,1)$
- 3) Gen $C \sim \text{poisson}(\frac{15}{.5+U})$
- 4) if $C \geq 20$, $X.append(1)$
else, $X.append(0)$
- 5) return $\text{avg}(X)$

b) A conditional estimator is as follows
 $E[I|U] = P(C \geq 20|U) = 1 - P(C \leq 19|U) = 1 - \sum_{k=0}^{19} \frac{e^{-\lambda} \lambda^k}{k!}$, $\lambda = \frac{15}{.5+U}$

then adding a control variable of $g(U) = \frac{15}{.5+U} = \frac{15}{E(U)+U}$

we get $= E[I|U] + \gamma(\frac{15}{.5+U} - E[\frac{15}{.5+U}])$

$E[I] = E[I|U] + \gamma(U-.5)$

$E[I] = 1 - \sum_{k=0}^{19} \frac{e^{-\lambda} \lambda^k}{k!} + \gamma(U-.5)$ when $\lambda = \frac{15}{.5+U}$, $\gamma = \frac{1}{\lambda}$

So to simulate we do the following

- $\frac{1}{n}$ times {
- 1) $X = []$
 - 2) Gen $U \sim \text{unif}(0,1)$
 - 3) set $\lambda = \frac{15}{.5+U}$
 - 4) $a = 1 - \sum_{k=0}^{19} \frac{e^{-\lambda} \lambda^k}{k!} + \gamma(U-.5) > 1$, $X.append(a)$
 ~~$a = 1 - \sum_{k=0}^{19} \frac{e^{-\lambda} \lambda^k}{k!} + \gamma(U-.5)$~~
 - 5) return $\text{avg}(X)$

9.19. cont.

c) note that $1 - \sum_{k=0}^{19} \frac{e^{-\lambda} \lambda^k}{k!}$, $\lambda = \frac{15}{.5+u}$ and

$$1 - \sum_{k=0}^{19} \frac{e^{-\delta} \delta^k}{k!}, \delta = \frac{15}{1.5-u}$$

have the same distribution and are negatively correlated.

thus a conditional and antithetic estimator is as follows

$$E[I] = \frac{1}{2} \left[\sum_{i=1}^2 \left(1 - \sum_{k=0}^{19} \frac{e^{-\lambda_{(i)}} \lambda_{(i)}^k}{k!} + 1 \right) \right], \lambda_{(1)} = \frac{15}{.5+u}, \lambda_{(2)} = \frac{15}{1.5-u}$$

which we simulate by the following

- n times {
- 1) $X = []$
 - 2) Gen $u \sim \text{unif}(0,1)$
 - 3) set $\lambda_{(1)} = \frac{15}{.5+u}, \lambda_{(2)} = \frac{15}{1.5-u}$
 - 4) set $\beta_{(1)} = 1 - \sum_{k=0}^{19} \frac{e^{-\lambda_{(1)}} \lambda_{(1)}^k}{k!}$
 $\beta_{(2)} = 1 - \sum_{k=0}^{19} \frac{e^{-\lambda_{(2)}} \lambda_{(2)}^k}{k!}$
 - 5) $X.append(\beta_{(1)} + \beta_{(2)})$
 - 6) return $\frac{\text{avg}(X)}{2}$

d)

Raw E: 0.29208, Conditional + Control Var E: 0.297733, Conditional + Antithetic Var E: 0.290481

Raw Var: 0.206771

Conditional + Control Var: 0.096605, Reduction over Raw: 2.1404

Conditional + Antithetic Var: 0.016109, Reduction over Raw: 12.8357

9.21

$$\begin{aligned}
 a) \quad \Theta &= P(S_n > c) = \sum_i P(S_n > c | m_n = X_i) P(m_n = X_i) \\
 &= \sum_i P(S_n > c | m_n = X_i) \frac{1}{n} \\
 &= \frac{1}{n} \sum_i \frac{P(S_n > c, m_n = X_n)}{P(m_n = X_n)} \\
 &= \frac{1}{n} \sum_i \frac{P(S_n > c, m_n = X_n)}{\frac{1}{n}} = \sum_i P(S_n > c, m_n = X_n) \\
 &= n P(S_n > c, m_n = X_n)
 \end{aligned}$$

$$\begin{aligned}
 b) \quad P(S_n > c, m_n = X_n | \{X_i\}_{i=1}^{n-1}) &= P(S_n > c | \{X_i\}_{i=1}^n) P(m_n = X_n | \{X_i\}_{i=1}^{n-1}) \\
 &= P(S_n > c) P(S_{n-1} \leq c) P(m_n = X_n) P(m_{n-1} \neq X_{n-1}) \\
 &= P(S_n > c) (1 - P(S_n > c)) P(m_n = X_n) P(m_{n-1} \neq X_{n-1}) \\
 &= \frac{(\Theta - \Theta^2)(n-1)}{n^2}
 \end{aligned}$$

9.24

a), b), c), d)

Raw E: 35.75642 Antithetic E: 34.727933

Control of S E: 35.81240339742783, Control of S and I E: 35.810366

Raw Var: 356.785666

Antithetic Var: 76.625111, Reduction over Raw: 4.6562

Control of S Var: 76.625111, Reduction over Raw: 3.5848

Control on S and I Var: 76.625111, Reduction over Raw: 4.3727

9.24

e) note that $E[\bar{T}_i | N_i] = (N_i + 1)\mu$, μ is mean service time

so all we must know is $f(\cdot)$ for N_i . Observe the following

given $N_1 = 0$ we can say

$P(N_1 = 0) = 1$ then for N_2 we can say

$P(N_2 = 0) = \frac{1}{3}$, $P(N_2 = 1) = \frac{2}{3}$ and extrapolate this out to

for $k \geq 3$

$$P(N_k = 0) = \sum_{i=0}^{k-3} P(N_{k-1} = i) \left(\frac{1}{3}\right)^{i+1}$$

$$P(N_k = j) = \left(\frac{2}{3}\right) \sum_{i=j-1}^{k-3} P(N_{k-1} = i) \left(\frac{1}{3}\right)^{i+1-j} \quad \forall 1 \leq j \leq k-1$$

5.1.4

a) A raw estimator is as follows

$$\begin{aligned} E[z] &= E[P(X \in A : \{(x, y) : x, y \geq a\})] \\ &= \frac{1}{n} \sum_{i=1}^n \mathbb{1}(X \in A) \\ &= \frac{1}{n} \sum_{i=1}^n \mathbb{1}(x \geq a \text{ and } y \geq a), \quad x, y \sim N(\bar{0}, C) \\ &\quad \text{where } C = \begin{pmatrix} 4 & -1 \\ -1 & 4 \end{pmatrix}, \bar{0} = (0, 0) \end{aligned}$$

We can simulate through the following

- 1) $X = []$
- 2) Gen $x, y \sim N(\bar{0}, C)$
- 3) if $x \geq a$ and $y \geq a$, $X.append(1)$
else, $X.append(0)$
- 4) return $\text{avg}(X)$

from X we can easily find $E[X]$, $\text{var}(X)$ and the associated 95% CI

b) pretty clearly $(a, a)^T$ maximizes $N(0, C)$ over $A = \{(x, y) : x, y \geq a\}$

$$\text{So we just find } E[z] = \frac{1}{n} \sum \mathbb{1}(x \geq a, y \geq a), \quad x, y \sim N((a, a)^T, C)$$

by the following

- 1) $X = []$
- 2) Gen $\bar{x}, \bar{y} \sim N((a, a)^T, C)$
- 3) if $\bar{x} \geq a$, $\bar{y} \geq a$, $X.append(1)$
else $X.append(0)$
- 4) return $\text{avg}(X)$

c) we replace C with $\bar{C} = \delta C$ for some $\delta \in \mathbb{R}$ in the above methods

Crude MC simulation, $a = 1$

E: 0.06641, Var: 0.062000331903319036

95% CI: (0.06486670000476241, 0.06795329999523758)

Importance Sampling MC simulation, $a = 1$

Var: 0.01991542623748882

95% CI: (0.06422883527214322, 0.06597819105987184)

Crude MC simulation, $a = 3$
E: 0.00139, Var: 0.0013880817808178082
95% CI: (0.0011590803584807677, 0.0016209196415192322)
Importance Sampling MC simulation, $a = 3$
Var: 2.0454874191127908e-05
95% CI: (0.0013465624310992428, 0.0014026261302517906)

Crude MC simulation, $a = 10$
E: 0, Var: 0
95% CI: (nan, nan)
Importance Sampling MC simulation, $a = 10$
Var: 1.0093322023063004e-32
95% CI: (1.1200873995965199e-17, 1.244624975505494e-17)

Importance Sampling MC simulation, $a = 1$, $\delta = 0.001$
Var: 3.5908585150546344e-294
95% CI: (6.386310202812891e-149, 8.735307354325856e-149)
Importance Sampling MC simulation, $a = 1$, $\delta = 2$
Var: 0.039487006633100515
95% CI: (0.09565626437097965, 0.09811952547760372)
Importance Sampling MC simulation, $a = 1$, $\delta = 10$
Var: 0.08920734914819273
95% CI: (0.14933945547357524, 0.15304185822852984)

Importance Sampling MC simulation, $a = 3$, $\delta = 0.001$
Var: 0.0
95% CI: (nan, nan)
Importance Sampling MC simulation, $a = 3$, $\delta = 2$
Var: 0.0007208506592266336
95% CI: (0.009701072039568772, 0.010033889416399955)
Importance Sampling MC simulation, $a = 3$, $\delta = 10$
Var: 0.02263870018307892
95% CI: (0.06889521052734648, 0.0707603404399142)

Importance Sampling MC simulation, $a = 10$, $\delta = 0.001$
Var: 0.0
95% CI: (nan, nan)
Importance Sampling MC simulation, $a = 10$, $\delta = 2$
Var: 6.302214284788805e-18
95% CI: (3.848106698983523e-10, 4.159299775524978e-10)
Importance Sampling MC simulation, $a = 10$, $\delta = 10$
Var: 9.539689167495977e-06
95% CI: (0.0008941180978331379, 0.0009324050200036513)

5.4.1

A raw estimator is $E[\theta] = \frac{1}{k} \sum_{i=1}^k \mathbb{1}\left(\sum_{n=1}^M X_n D_n > C = 45\right)$ which we simulate by

- k times {
- 1) $X = []$
 - 2) Gen $P_n \sim \text{Beta}(1, 19)$, $X_n \sim \mathcal{N}(3, 1)$ for $n \in [1, 100]$
 - 3) Find $D_n \sim \text{Bern}(P_n) \forall n = 1, \dots, 100$
 - 4) Find $L = \sum_{n=1}^M D_n X_n$
 - 5) if $L > 45$, $X.append(1)$
else $X.append(0)$
 - 6) return $\text{avg}(X)$

then b/c $E\left[\sum_{n=1}^M X_n D_n \mid \sum_{n=1}^M D_n\right]$

$$\begin{aligned}
 &= \sum_{n=1}^M (E[X_n D_n \mid \sum_{n=1}^M D_n]) = \sum_{n=1}^M (E[X_n \mid \sum_{n=1}^M D_n] E[D_n \mid \sum_{n=1}^M D_n]) \\
 &= \sum_{n=1}^M (E[X_n] E[D_n \mid \sum_{n=1}^M D_n]) \\
 &= \mu \sum_{n=1}^M (E[D_n \mid \sum_{n=1}^M D_n]) = \mu E\left[\sum_{n=1}^M D_n \mid \sum_{n=1}^M D_n\right] \\
 &= \mu \sum_{n=1}^M D_n
 \end{aligned}$$

so $P(L > 45 \mid X) = 1 - \Phi\left(\frac{X - \mu \sum_{n=1}^M D_n}{\sqrt{\sum_{n=1}^M D_n}}\right)$ where $\Phi(\cdot)$ is the CDF of $\mathcal{N}(0, 1)$ at (\cdot)

so we simulate the conditional estimator of $E[\theta] = \frac{1}{k} \sum_{i=1}^k A_{Z_{D_n}}$ where $A_{Z_D} = 1 - \Phi\left(\frac{45 - \mu \sum_{n=1}^M D_n}{\sqrt{\sum_{n=1}^M D_n}}\right)$

- by
- 1) $X = []$
 - 2) Gen $P_n \sim \text{Beta}(1, 19)$, $D_n \sim \text{Bern}(P_n) \forall n$
 - 3) find $X = \frac{\mu \sum_{n=1}^M D_n}{\sqrt{\sum_{n=1}^M D_n}} = \beta$
 - 4) find A_{Z_β}
 - 5) $X.append(A_{Z_\beta})$
 - 6) return $\text{avg}(X)$