

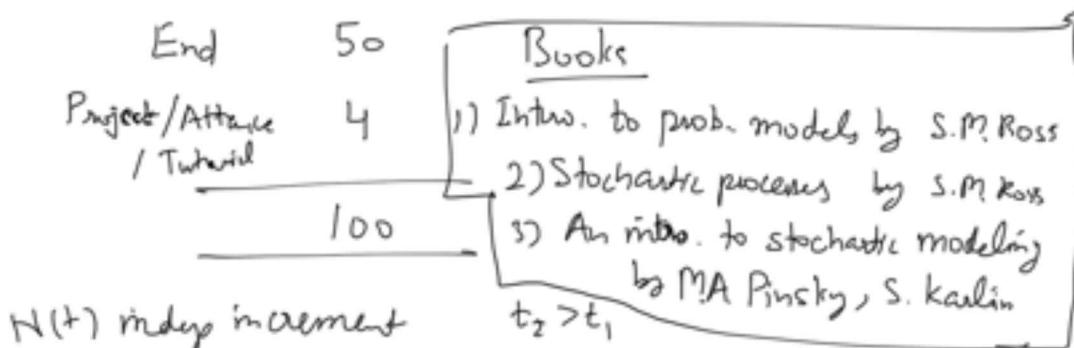
MA 41017/MA 60067

Stochastic Processes / Stochastic Process and Simulation  
Marks

CT-1 8 (Fr.) 3 Feb. 11:10 - 12:10

Mid 30

CT-2 8 (Fr.) 31 March 11:10 - 12:10



$$P(N(t_2) - N(t_1) = x_2 - x_1, N(t_1) = x_1)$$

$$\rightarrow = P(N(t_2) - N(t_1) = x_2 - x_1) P(N(t_1) = x_1)$$

$$= P(N(t_2 - t_1) = x_2 - x_1) P(N(t_1) = x_1)$$

$N(t)$  stationary  
fahren

$$E(X) = \overline{E(X|y)}$$

$$E(X|y=y) = \sum_x x p_{X|Y=y}(x) = \phi(y)$$

$$E(X|Y) = \phi(Y)$$

$$E(E(X|Y)) = E(\phi(Y)) = \sum_y \phi(y) p_Y(y)$$

$$= \sum_y \sum_x x \left( \underbrace{p_{X|Y=y}(x)}_{\frac{p(x,y)}{p_Y(y)}} \right) p_Y(y)$$

$$= \sum_y \sum_x x p(y, x)$$

$$= \sum_y \sqrt{\sum_x h(x,y)} - \sum_y \dots$$

$$-\sum_{x \in \mathbb{Z}} P_x(x) = \sum_{x \in \mathbb{Z}} P_x(x) = E(X)$$

$X_i \sim \text{Pois}(x_i)$ ,  $i=1,2$

$$\underbrace{X_1, X_2}_{\text{random}} \quad M_{X_i}(t) = e^{\lambda_i(e^t - 1)}, i=1,2$$

$$S_2 = X_1 + X_2 \sim \text{Pois}(\lambda_1 + \lambda_2)$$

$$\begin{aligned} M_{S_2}(t) &= M_{X_1}(t) M_{X_2}(t) \\ &= e^{(\lambda_1 + \lambda_2)(e^t - 1)} \end{aligned}$$

Syllabus : DTMC, Poiss. Process and related distributions, CTMC, queuing theory, renewal processes, martingales, Brownian Motion, simulation.

Stochastic Process (S.P.)  
is a family of random variables (or)  $\{X(t), t \in T\}$ , defined on a given probability space, indexed by the parameter  $t, t \in T$   
values assumed by  $X(t) \in S$  indexed  
called states State Space

T parameter space or time space

(1) discrete state, discrete parameter Sp

(2) " " , continuous + "

(3) continuous " , " , " , "

(4) " , " , discrete " , "

Example Consider a queuing system with jobs arriving at random points in time, queuing for service and departing from the system.

..... improving now the system after service completion.

a)  $X(t)$  # of jobs in the system at time  $t$

$$\{X(t), t \in T\}$$

$X(t) \in \{0, 1, 2, \dots\} = S$  discrete state,

$$T = [0, \infty)$$

continuous parameter SP.

b)  $w_k$  time that the  $k^{\text{th}}$  customer has to wait in the system before receiving service.

$$\{w_k, k \in T\}$$

$w_k \in [0, \infty) = S$  continuous state,

$$T = \{1, 2, \dots\}$$

discrete parameter SP

c)  $y(t)$ : cumulative service requirement (expeniture) of all jobs in the system at time  $t$ .

$$y(t) \in [0, \infty) = S \quad T = [0, \infty)$$

$\{y(t)\}$  is continuous state, continuous parameter

d)  $N_k$  # of jobs in the system at the time of departure of the  $k^{\text{th}}$  customer (after service completion).

$$N_k \in \{0, 1, 2, \dots\} = S$$

$$T = \{1, 2, \dots\}$$

$\{N_k, k \in T\}$  as discrete state, discrete parameter

— X —

SP.

Discrete time Markov Chain: (DTMC)

SP.  $\{X_n, n = 0, 1, 2, \dots\}$  that takes values on a finite or countable number of values

$[0, 1, 2, \dots] = S \Rightarrow$  discrete state space

discrete parameter S.P.  $\{X_n\}$

$$i, j, i_0, i_1, \dots \in S$$

$$P(X_{n+1} = j | X_n = i) = P(i \rightarrow j)$$

$$P(X_{n+1} = j | X_n = i, X_{n-1} = i_{n-1}, \dots, X_0 = i_0)$$

$$= P(X_{n+1} = j | X_n = i)$$

$$= P_{ij}(n, n+1)$$

$P_{ij}^{(1)}$  stationary transition probability  
 initial state      final state      or  
 Homogeneous M.C.

$$= P_{ij}$$

$$P^{(1)} = P = \begin{matrix} i \\ \downarrow \\ 0 \\ \text{transition probability matrix} \\ \text{tpm} \end{matrix} = 2 \begin{bmatrix} j \rightarrow & 0 & 1 & 2 & \dots \\ P_{00} & P_{01} & P_{02} & \dots \\ P_{10} & P_{11} & P_{12} & \dots \\ P_{20} & P_{21} & P_{22} & \dots \\ \hline \dots & \hline \dots & \hline \dots & \hline \dots \end{bmatrix}$$

$$0 \leq P_{ij} \leq 1, \forall i, \forall j$$

$$\sum_j P_{ij} = 1 \quad \text{for fixed } i$$

—x—

Example 1 Consider a game of ladder climbing.

There are 5 levels in the game, level 1 is lowest (bottom) and level 5 is the highest (top). A player starts at the bottom. Each time, a fair coin is tossed. If it turns up heads, the player moves up one rung. If tails, the player moves down to the very bottom. Once at the top level, the player moves to the very bottom if tail turns up and stays at the top if head turns up.

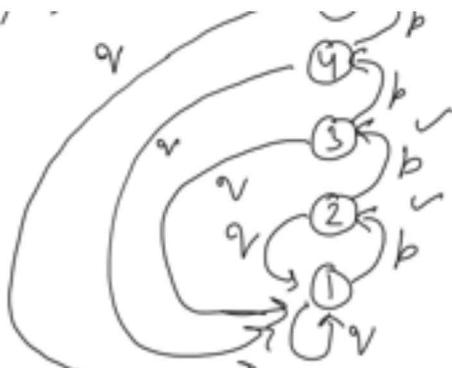
Let  $X_n$  be the level of the game in the  $n$ th step / transition. tpm of  $X_n$

$$X_n \in \{1, 2, 3, 4, 5\} = S \quad \xrightarrow{\text{TPM}} \xleftarrow{\text{S}}$$

$$P_{ij} = \frac{1}{2}$$

$X_n$  DTMC

$$P_{ij} = P(X_{n+1}=j | X_n=i) \\ = P(X_1=j | X_0=i)$$



tpm

$$P = \begin{bmatrix} q & p & 0 & 0 & 0 \\ q & 0 & p & 0 & 0 \\ q & 0 & 0 & p & 0 \\ q & 0 & 0 & 0 & p \\ q & 0 & 0 & 0 & p \end{bmatrix}$$

$p+q=1$   
 $p=q=\frac{1}{4}$

$$P(X_{n+1}=j | \underline{X_n=i}, X_{n-1}=i_{n-1}, \dots, X_0=i_0) \\ = P(X_{n+1}=j | \underline{X_n=i})$$

— X —

Example: Let  $(X_n)_{n=0,1,2,\dots}$  be a sequence of i.i.d.  
(independently & identically distributed)

discrete rv. with  $P(X_1=j) = \left(\frac{1}{2}\right)^{j+1}$ ,  $\forall j = 0, 1, 2, \dots$

Determine whether each of the following chain  
is Markovian or not. If so find its corresponding  
state space (S) and tpm

(i)  $\{S_n\}_{n=0,1,2,\dots}$  where  $S_n = \sum_{i=1}^n X_i$

(ii)  $\{M_n\}_{n=0,1,2,\dots}$  where  $M_n = \max\{X_1, X_2, \dots, X_n\}$

For (i)  $S_n \in \{0, 1, 2, \dots\}$   $S_{n+1} = S_n + X_{n+1}$

$$P_{ij} = P(S_{n+1}=j | S_n=i)$$

$$S_n=i \quad \xrightarrow{S_{n+1}=j} \quad 0 \quad ! \quad 2 \quad 3 \quad \dots \quad n$$

$$t \text{pm} \quad P = \begin{vmatrix} 0 & p_0 & p_1 & p_2 & p_3 & \dots \\ 1 & 0 & p_0 & p_1 & p_2 & \dots \\ 2 & 0 & 0 & p_0 & p_1 & \dots \\ \vdots & - & - & - & - & \dots \\ \vdots & - & - & - & - & \dots \end{vmatrix} = \begin{bmatrix} \frac{1}{2} & \frac{1}{4} & \frac{1}{8} & \frac{1}{16} & \dots \\ 0 & \frac{1}{2} & \frac{1}{4} & \frac{1}{8} & \dots \\ 0 & 0 & \frac{1}{2} & \frac{1}{4} & \dots \\ - & - & - & - & \dots \end{bmatrix}$$

Example ( Transformation of a process into M.C.)

Suppose that whether or not it rains today depends on previous weather conditions through the last two days. Suppose that if it has rained for the past two days, then it will rain tomorrow with prob. ( $w_r$ ) 0.7; if it has rained today but not yesterday, then it will rain tomorrow w.p 0.5; if it has rained yesterday but not today, then it will rain tomorrow w.p 0.4; if it has not rained in the past two days, then it will rain tomorrow w.p 0.2.

Set  $X_n$  state at any time  $n$  is determined by the weather condition during both that day and the previous day

State $X_n$	Rained yesterday	Rained today	
0	✓	✓	
1	✗	✓	
2	✓	✗	
3	✗	✗	

$$S = \{0, 1, 2, 3\}$$

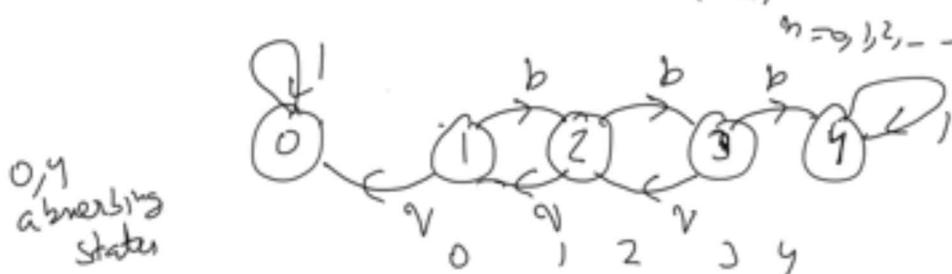
t pm yesterday	today		$X_{n+1} = ?$
	0	1	
0	✓✓	0.7 0	0.3 0
1	✗✓	0.5 0	0.5 0

$$\Gamma = \begin{matrix} & \leftarrow v \times \\ 3 \times 3 & \begin{bmatrix} 0 & 0.4 & 0 & 0.6 \\ 0 & 0.2 & 0 & 0.8 \end{bmatrix} \end{matrix}$$

Example 1 Particle performs a random walk in states,  $\{0, 1, 2, 3, 4\}$ . It remains in state 0 and 4 with probability 1. It moves from state  $n$  ( $0 < n < 4$ ) to  $n+1$  with prob  $b$ ; and from state  $n$  to  $n-1$  with prob  $a = 1-b$ .

Let  $X_n$  : position of particle at time/step  $n$ .

$$X_n \in \{0, 1, 2, 3, 4\} = S \quad (X_n) \text{ M.C.}$$



$$t_{pm} \quad P = \begin{bmatrix} 0 & 1 & 0 & 0 & 0 \\ 1 & a & 0 & b & 0 \\ 2 & 0 & a & 0 & b \\ 3 & 0 & 0 & a & 0 \\ 4 & 0 & 0 & 0 & 1 \end{bmatrix}$$

$n$ -step transition probability:

$$i, j \in S \quad (X_n) \text{ M.C.} \quad S = \{0, 1, 2, \dots\}$$

$$P_{ij}^{(n)} = P(X_{m+n} = j | X_m = i) = P(X_n = j | X_0 = i)$$

$$n\text{-step } t_{pm} \rightarrow P^{(n)} = \left( (P_{ij}^{(n)}) \right)_{0 \leq i, j \leq n} = \begin{bmatrix} P_{00}^{(n)} & P_{01}^{(n)} & P_{02}^{(n)} & \dots \\ P_{10}^{(n)} & P_{11}^{(n)} & P_{12}^{(n)} & \dots \\ \vdots & \vdots & \vdots & \ddots \end{bmatrix}$$

$$\sum_j P_{ij}^{(n)} = 1 \quad \text{for fixed } i$$

$$0 \leq P_{ij} \leq 1 \quad \forall i, \forall j$$

Chapman kolmogorov equations  $i, j, k \in S$

$$P_{ij}^{(m+n)} = \sum_k P_{ik}^{(m)} P_{kj}^{(n)} = \sum_k P_{ik}^{(n)} P_{kj}^{(m)}$$

$$\xrightarrow{(i,j) \text{ th cell}} P_{ij}^{(m+n)} = P(X_{m+n} = j \mid X_0 = i)$$

$$\xleftarrow{\text{JP}} = \sum_k P(X_{m+n} = j, X_n = k \mid X_0 = i)$$

$$\begin{aligned} P(A) &= P\left(\bigcup_i (A \cap E_i)\right) \\ &= \sum_i P(A \cap E_i) \end{aligned}$$

$$= \sum_k P(X_{m+n} = j \mid X_n = k, X_0 = i) P(X_n = k \mid X_0 = i)$$

$$P(AB|c) = P(A|Bc) P(B|c)$$

$$= \sum_k P(X_{m+n} = j \mid X_n = k) P(X_n = k \mid X_0 = i)$$

$$= \sum_k P_{kj}^{(n)} P_{ik}^{(m)}$$

$$\begin{pmatrix} - & \overline{-} & \overline{-} \\ P_{i0}^{(n)} & P_{i1}^{(n)} & - \\ - & - & - \end{pmatrix} \begin{pmatrix} - & P_{0j}^{(m)} & - \\ - & P_{1j}^{(m)} & - \\ - & - & - \end{pmatrix}$$

$$= P^{(m+n)}$$

$$P^{(m+n)} = P^{(n)} P^{(m)} = P^{(m)} P^{(n)}$$

$$P^{(1)} = P$$

$$P^{(2)} = P^{(1)} P^{(1)} = P \cdot P = P^2$$

$$P^{(n)} = P^n$$

—x—

pmt of step  $X_n$

$\{ \in S = \{0, 1, 2, \dots\}$

$$P(X_n=i) = p_i^{(n)}$$

pmt of  $X_n$

$$\underline{p}^{(n)} = (P(X_n=0), P(X_n=1), \dots) = (p_0^{(n)}, p_1^{(n)}, \dots)$$

pmt of  $X_0$

$$\underline{p}^{(0)} = (p_0^{(0)}, p_1^{(0)}, \dots)$$

$P \leftarrow tpm$

claim  $\underline{p}^{(n)} = \underline{p}^{(n-1)} P = \dots = \underline{p}^{(0)} P^n$

$$\begin{aligned} p_i^{(n)} &= p_0^{(n)} p_{0i} + p_1^{(n-1)} p_{1i} + \dots \\ &= \sum_k p_k^{(n-1)} p_{ki} \end{aligned}$$

so  $p_i^{(n)} = P(X_n=i)$

$$= \sum_k P(X_n=i, X_{n-1}=k)$$

$$= \sum_k P(X_n=i | X_{n-1}=k) P(X_{n-1}=k)$$

$$= \sum_k p_{ki} p_k^{(n-1)}$$

$$(p_0^{(n)}, p_1^{(n)}, \underline{\cancel{p_2^{(n)}}}) = (\underline{p_0^{(n-1)}}, \underline{p_1^{(n-1)}} \dots) \left( \begin{matrix} p_{00} & p_{01} \\ \vdots & \vdots \\ p_{0n} & p_{1n} \end{matrix} \right)$$

$$\underline{p}^{(n)} = \underline{p}^{(n-1)} P$$

—x—

Example:  $\{X_n\}$  M.C.  $S = \{1, 2, 3\}$

	1	2	3
1	0.1	0.5	0.4

$$P(X_0=1) = 0.6 \quad P(X_0=2) = 0.2 \quad P(X_0=3) = 0.2$$

$$P(X_0=1) = 0.7, \quad P(X_0=2) = 0.2, \quad P(X_0=3) = 0.1$$

$$\textcircled{a} \quad P(X_0=2, X_1=3, X_2=2, X_3=2)$$

$$= P(X_3=2, X_2=3, X_1=3, X_0=2)$$

$$= P(X_3=2 | X_2=3, X_1=3, X_0=2) P(X_2=3 | X_1=3, X_0=2)$$

$$\cdot P(X_1=3 | X_0=2) P(X_0=2)$$

$$= P(X_3=2 | X_2=3) P(X_2=3 | X_1=3) P(X_1=3 | X_0=2) \\ P(X_0=2)$$

$$= P_{32} P_{23} P_{13} P(X_0=2) = 0.4 \times 0.3 \times 0.2 \times 0.2 \\ = 0.0048$$

$$\textcircled{b} \quad P(X_2=3, X_1=3 | X_0=2)$$

$$= P(X_2=3 | X_1=3, X_0=2) P(X_1=3 | X_0=2)$$

$$= P(X_2=3 | X_1=3) P(X_1=3 | X_0=2)$$

$$= P_{33} P_{23} = 0.3 \times 0.2$$

$$\textcircled{c} \quad P(X_3=2, X_0=2, X_1=3)$$

$$= P(X_3=2, X_1=3, X_0=2)$$

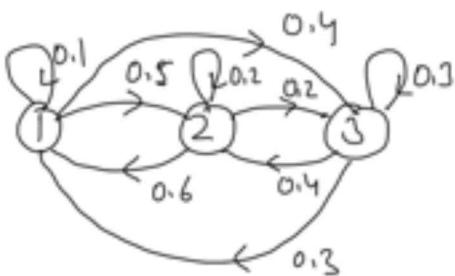
$$= P(X_3=2 | X_1=3, X_0=2) P(X_1=3 | X_0=2) P(X_0=2)$$

$$= P(X_3=2 | X_1=3) P(X_1=3 | X_0=2) P(X_0=2)$$

$$= \underline{P_{32}^{(2)}} P_{23} P(X_0=2) = 0.35 \times 0.2 \times 0.2 \\ = 0.014$$

$$\frac{P^{(2)}}{-^{(2)}} = P^2 = P \cdot P \quad k \in S = \{1, 2, 3\}$$

$$\begin{aligned}
 P_{22} &= \sum_k P_{3k} P_{k2} \\
 &= P_{31} P_{12} + P_{32} P_{22} + P_{33} P_{32} \\
 &= 0.3 \times 0.5 + 0.4 \times 0.2 + 0.3 \times 0.4 \\
 &= 0.35
 \end{aligned}$$



$$(d) P(X_2 = 3) = \underline{\underline{p}}_3^{(2)} \quad P \leftarrow \text{tpm}$$

$$\underline{\underline{p}}^{(2)} = \underline{\underline{p}}^{(1)} P$$

$$\underline{\underline{p}}^{(1)} = \underline{\underline{p}}^{(0)} P$$

$$\underline{\underline{p}}^{(0)} = (0.7, 0.2, 0.1)$$

$$\begin{aligned}
 \underline{\underline{p}}^{(1)} &= (0.7, 0.2, 0.1) \begin{pmatrix} 0.1 & 0.5 & 0.4 \\ 0.6 & 0.2 & 0.2 \\ 0.3 & 0.4 & 0.3 \end{pmatrix} \\
 &= (0.22, 0.43, 0.35)
 \end{aligned}$$

$$\underline{\underline{p}}^{(2)} = \underline{\underline{p}}^{(1)} P$$

$$\begin{aligned}
 \underline{\underline{p}}_3^{(2)} &= 0.22 \times 0.4 + 0.43 \times 0.2 + 0.35 \times 0.3 \\
 &= 0.279
 \end{aligned}$$

Example Consider a two-state M.C.  $(X_n)$  having state

space  $S = \{0, 1\}$  with tpm

$$Q = \begin{pmatrix} 0 & 1 \\ \frac{1}{2} & \frac{1}{2} \\ 1 & \frac{2}{3} \end{pmatrix}$$

Whether  $Z_n = (X_{n-1}, X_n)$  is a M.C.? If so

determine state space and tpm of  $\{Z_n\}$ .

Sol.  $\{Z_n\}$  is a M.C.

S = {(0,0), (0,1), (1,0), (1,1)}					
		(0,0)	(0,1)	(1,0)	(1,1)
		1/2	1/2	0	0
(0,0)		0	0	1/3	2/3
(0,1)		1/2	1/2	0	0
(1,0)		0	0	1/3	2/3
(1,1)					

$$P_{(i,j),(k,l)} = P(Z_{n+1} = (k,l) | Z_n = (i,j))$$

$$\begin{aligned}
 & i, j, k, l \in \{0, 1\} \\
 P(Z_{n+1} = (1,1) | Z_n = (0,1)) &= \frac{P(A|BC)}{P(ABC)} = \frac{P(ABC)}{P(BC)P(B)} = \frac{P(ABC)}{P(B)P(C)} = P(A|BC) \\
 & = P(X_n = 1, X_{n+1} = 1 | X_{n-1} = 0, X_n = 1) \\
 & = P(X_{n+1} = 1 | X_n = 1, X_{n-1} = 0) \\
 & = P(X_{n+1} = 1 | X_n = 1) = \theta_{11} = \frac{2}{3}
 \end{aligned}$$

Classification of states:

$$\{X_n\} \text{ M.C. } S = \{0, 1, 2, 3, \dots\}$$

Def<sup>n</sup>  $i \rightarrow j$ , state  $j$  is accessible from state  $i$  if  $P_{ij}^{(n)} > 0$  for some  $n$ .

Def<sup>n</sup>  $i \leftrightarrow j$ , states  $i$  and  $j$  communicate with each other  
if  $i \rightarrow j$  and  $j \rightarrow i$

Result  $i \leftrightarrow j, j \leftrightarrow k \Rightarrow i \leftrightarrow k$

Set  $\exists n, m \text{ st. } P_{ij}^{(n)} > 0, P_{jk}^{(m)} > 0$

$$P_{ik}^{(m+n)} = \sum_l P_{il}^{(n)} P_{lk}^{(m)} \quad (\text{C k equilibrium})$$

$$\geq \tilde{P}_{ij}^{(n)} P_{jk}^{(m)}$$

$$> 0$$

$i \rightarrow k$ . Similarly  $k \rightarrow i \therefore i \leftrightarrow k$

Def<sup>b</sup> M.C  $(X_n)$  is irreducible or (connected) if every state communicate with every other state otherwise reducible.

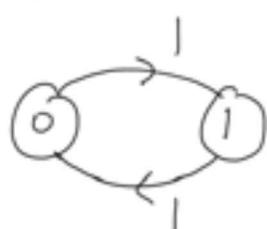
Def<sup>b</sup> period of state  $i$

$$d(i) = \text{gcd } \{ n \mid P_{ii}^{(n)} > 0 \}$$

(If  $P_{ii}^{(n)} = 0 \forall n \geq 1$ , define  $d(i) = 0$ )

Example: ①  $(X_n) \quad S = \{0, 1\}$

$$P = \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix}$$



Irreducible M.C.

$0 \leftrightarrow 1$   
Class  $\{0, 1\}$

$$d(0) = \text{gcd } \{ 2, 4, 6, \dots \} \quad P_{00}^{(n)} > 0$$

$$= 2 = d(1)$$

② M.C  $S = \{0, 1, 2\}$  having tpm

$$P = \begin{pmatrix} 1/2 & 1/2 & 0 \\ 1/2 & 1/4 & 1/4 \\ 0 & 1/3 & 2/3 \end{pmatrix}$$



$0 \leftrightarrow 1 \leftrightarrow 2$       Class = {0, 1, 2}  
Irreducible  $M \subset$

$$d(0) = \inf \{ d(1, 2, 3, \dots) \} = 1 \quad P_{00}^{(n)} > 0 \\ = d(1) = d(2)$$

$\overbrace{\hspace{10em}}$        $(X_n, n \in S = \{0, 1, 2, \dots\})$

For state  $i \in S$

$f_{ii}^{(n)}$  =  $P(X_n=i, X_k \neq i, k=1, \dots, n-1 | X_0=i)$   
 recurrence probability of first visit to state  $i$  in  
 time  $n$  transitions / steps, starting from state  $i$

$$f_{ii}^{(0)} = 1$$

$$f_{ii} = f_{ii}^{(1)} + f_{ii}^{(2)} + f_{ii}^{(3)} + \dots$$

$\hookrightarrow$  probability of ever visiting state  $i$ , starting from state  $i$

$\rightarrow f_{ii} = 1$ , i.e., return to state  $i$  is certain,  
 starting from state  $i$   
 i recurrent state

$\rightarrow f_{ii} < 1$ , i.e., return to state  $i$  is uncertain  
 i transient state

$$P_{ij}^{(n)} = P(X_{m+n}=j | X_m=i)$$

$$\text{Let } I_n = \begin{cases} 1 & \text{if } X_n = i \\ 0 & \text{if } X_n \neq i \end{cases}$$

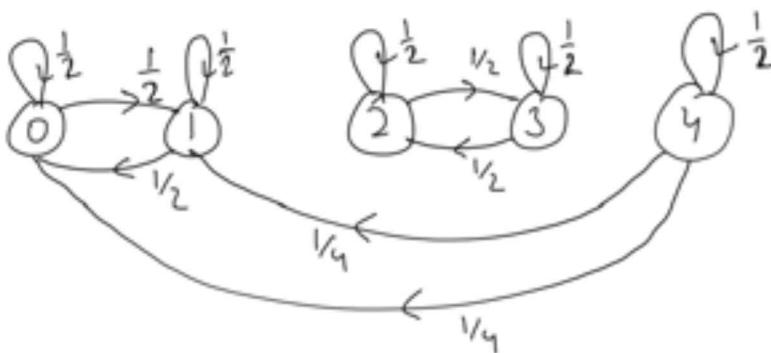
$\sum_{n=1}^{\infty} I_n$  : # of time periods, the process is in  
 state  $i$

$$\dots \xrightarrow{\infty} 1 \dots \xrightarrow{\infty} \dots \dots \dots$$

$$\begin{aligned}
 E\left(\sum_{n=1}^{\infty} I_n \mid X_0 = i\right) &= \sum_{n=1}^{\infty} E(I_n \mid X_0 = i) \\
 &= \sum_{n=1}^{\infty} [1 \cdot P(X_n = i \mid X_0 = i) + 0 \cdot P(X_n \neq i \mid X_0 = i)] \\
 &= \sum_{n=1}^{\infty} P_{ii}^{(n)} \\
 i \text{ recurrent} \Leftrightarrow f_{ii} &= 1 \Leftrightarrow \sum_{n=1}^{\infty} P_{ii}^{(n)} = \infty \\
 i \text{ transient} \Leftrightarrow f_{ii} &< 1 \Leftrightarrow \sum_{n=1}^{\infty} P_{ii}^{(n)} < \infty
 \end{aligned}$$

Example Consider a M.C. having states 0, 1, 2, 3, 4  
and tpm

$$D = \begin{bmatrix} 0 & 1 & 2 & 3 & 4 \\ \frac{1}{2} & \frac{1}{2} & 0 & 0 & 0 \\ 1 & \frac{1}{2} & \frac{1}{2} & 0 & 0 \\ 2 & 0 & 0 & \frac{1}{2} & \frac{1}{2} \\ 3 & 0 & 0 & \frac{1}{2} & \frac{1}{2} \\ 4 & \frac{1}{4} & \frac{1}{4} & 0 & 0 & \frac{1}{2} \end{bmatrix}$$



$0 \leftrightarrow 1$ ,  $2 \leftrightarrow 3, 4$       Reducible M.C.  
 (clan  $\{0, 1\}$ ,  $\{2, 3\}$ ,  $\{4\}$ )  
 ↓      ↗ transient  
 recurrent

$$\begin{aligned}
 f_{00} &= f_{00}^{(1)} + f_{00}^{(2)} + f_{00}^{(3)} + \dots \\
 &= \frac{1}{2} + \frac{1}{4} + \frac{1}{8} + \dots = \frac{1}{2} \left[ 1 + \frac{1}{2} + \frac{1}{4} + \dots \right] \\
 &= \frac{1}{2} \times \frac{1}{1 - \frac{1}{2}} = 1 \quad \text{0 recurrent.}
 \end{aligned}$$

$$f_{44} = f_{44}^{(1)} + f_{44}^{(2)} + \dots = \frac{1}{2} + 0 + 0 + \dots < 1$$

4 transient.

— x —

P1,  $i \leftrightarrow j$ ,  $i$  recurrent  $\Rightarrow j$  recurrent

Given  $\begin{cases} i \leftrightarrow j \rightarrow \exists n, m \text{ st. } P_{ij}^{(n)} > 0, P_{ji}^{(m)} > 0 \\ i \text{ recurrent} \Leftrightarrow \sum_v P_{ii}^{(v)} = \infty \end{cases}$

$$P_{jj}^{(m+n+v)} \geq P_{ji}^{(m)} P_{ii}^{(v)} P_{ij}^{(n)} \quad [\text{Using } C_k = 1]$$

$$\begin{aligned}
 \sum_v P_{jj}^{(m+n+v)} &\geq P_{ji}^{(m)} P_{ij}^{(n)} \underbrace{\sum_v P_{ii}^{(v)}}_{=\infty} \\
 &= \infty
 \end{aligned}$$

$\Rightarrow j$  recurrent.

P2  $i \leftrightarrow j$ ,  $i$  transient  $\Rightarrow j$  transient

P3 In a finite state M.C all states can not be transient.

P4 In a finite state, irreducible M.C. all states are recurrent.

self/why P1 and P3.

Def Let  $i$  recurrent

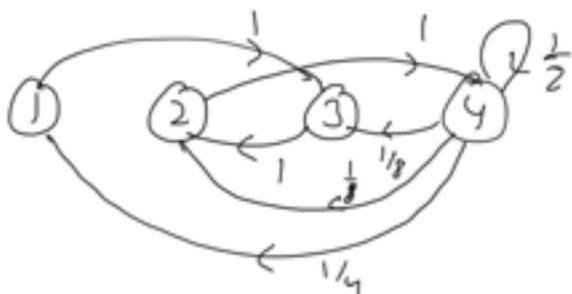
$$m_{ii} = \sum_{n=1}^{\infty} n f_{ii}^{(n)} \quad \text{mean recurrence time}$$

If  $m_{ii} = \infty$ ,  $i$  null recurrent

If  $m_{ii} < \infty$ ,  $i$  non-null recurrent/positive recurrent

Example:  $(X_t)$  M.C.  $S = \{1, 2, 3, 4\}$

$$P^S = \begin{matrix} & \begin{matrix} 1 & 2 & 3 & 4 \end{matrix} \\ \begin{matrix} 1 \\ 2 \\ 3 \\ 4 \end{matrix} & \begin{pmatrix} 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 1 \\ 0 & 1 & 0 & 0 \\ \frac{1}{4} & \frac{1}{8} & \frac{1}{8} & \frac{1}{2} \end{pmatrix} \end{matrix}$$



$1 \leftrightarrow 2 \leftrightarrow 3 \leftrightarrow 4$  Irreducible M.C.

Class  $\{1, 2, 3, 4\}$  finite state

all states are positive recurrent.

$$f_{44} = f_{44}^{(1)} + f_{44}^{(2)} + f_{44}^{(3)} + f_{44}^{(4)} + f_{44}^{(5)} + \dots$$

$$= \frac{1}{2} + \frac{1}{8} + \frac{1}{8} + \frac{1}{4} + 0 + 0 + \dots$$

$$\leq 1 \quad 4 \text{ recurrent}$$

$$m_{44} = \sum_{n=1}^{\infty} n f_{44}^{(n)} = 1 \times \frac{1}{2} + 2 \times \frac{1}{8} + 3 \times \frac{1}{8} + 4 \times \frac{1}{4} + 0 + \dots$$
$$= \frac{17}{8} < \infty$$

→ Finite state, irreducible M.C. all states are +ve recurrent

→ Irreducible M.C., all states are either +ve recurrent or null recurrent or transient.

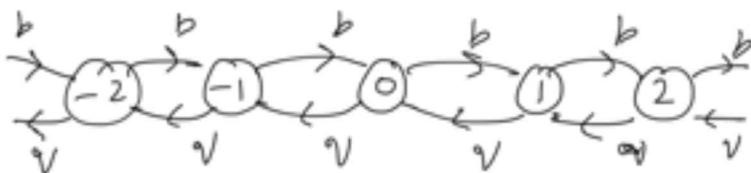
Example One-dimensional ...

— — — unbounded random walk

$$S = \{-\dots, -2, -1, 0, 1, 2, \dots\}$$

$X_n$  position of particle at nth step

$$P_{i,i+1} = p ; P_{i,i-1} = q = 1-p, P_{i,j} = 0, j \neq i-1, i+1)$$



$$P_{ii}^{(n)} = \begin{cases} \binom{2m}{m} p^m (1-p)^m & n=2m \\ 0 & n \neq 2m \end{cases}, m=1, 2, 3, \dots$$

$$= \begin{cases} a_m & , n=2m \\ 0 & , n \neq 2m \end{cases}$$

Ratio-test

$$\lim_{m \rightarrow \infty} \frac{a_{m+1}}{a_m} = \begin{cases} < 1, \sum a_m \text{ converges} \\ > 1, \sum a_m \text{ diverges} \end{cases}$$

$$\frac{a_{m+1}}{a_m} = \frac{\binom{2m+2}{m+1} p^{m+1} (1-p)^{m+1}}{\binom{2m}{m} p^m (1-p)^m}$$

$$= \frac{(2m+2)(2m+1)}{(m+1)(m+1)} p (1-p)$$

$$\lim_{m \rightarrow \infty} \frac{a_{m+1}}{a_m} = 4 p (1-p)$$

$$= \begin{cases} 1 & ; p = \frac{1}{2} \\ < 1 & ; p \neq \frac{1}{2} \end{cases}$$

$$p \neq \frac{1}{2} \Rightarrow \sum_n P_{ii}^{(n)} < \infty, i.e., i \text{ transient}$$

$b \neq \frac{1}{2}$  irreducible M.C all states are transient

Show  $b = \frac{1}{2}$  is accurate  $\dots - - -$  recurrent

Food  $P_2 = \frac{1-(3)^2}{1-3^5} = \frac{1-9}{1-243} = \frac{8}{242} = \frac{4}{121} \approx \frac{1}{30}$

$\frac{N}{b} = 3$   $i=2, N=5$

Gambler's ruin problem: Stuck  $1 - P_2$

initial capital Rs  $i$   $\xrightarrow{\text{aim}}$  Rs  $N$   
 $i = 0, 1, 2, \dots, N$

$$P(Z_i = +1) = b; P(Z_i = -1) = q = 1-b$$

$Z_i$ :  $i^{th}$  bet / step / transition.

$X_n$ : player's fortune after  $n^{th}$  bet/steps  
 $\in \{0, 1, 2, \dots, N\}$  M.C



$P_{00} = 1, P_{NN} = 1, N$  recurrent/transient  
 $0, N$  recurrent/transient  
 $1, 2, \dots, N-1$  transient

typm  $P = \begin{bmatrix} 1 & 0 & 0 & \cdots & 0 \\ q & 0 & b & 0 & \cdots & 0 \\ 0 & q & 0 & b & \cdots & 0 \\ \vdots & \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \cdots & 1 \end{bmatrix}$

$T_0$ : time he broke

$$T_0 = \inf \{n : X_n = 0\}$$

$T_N$ : time he has Rs. N

$$T_N = \inf \{n : X_n = N\}$$

$P_i = P(T_N < T_0)$  : Prob. that starting with  $i$ ,  
the gambler's fortune will reach  
 $N$  before reaching 0.

$$\begin{aligned} &= P(T_N < T_0 | Z_1 = -1) \frac{P(Z_1 = -1)}{} \\ &\quad + P(T_N < T_0 | Z_1 = 1) \frac{P(Z_1 = 1)}{} \\ &= q P_{i-1} + p P_{i+1} \end{aligned}$$

$$\Rightarrow q P_i + p P_i = q P_{i-1} + p P_{i+1}$$

$$\Rightarrow P_{i+1} - P_i = \frac{q}{p} (P_i - P_{i-1})$$

$$i=1 \quad P_2' - P_1 = \frac{q}{p} P_1 \quad P_0 = 0, P_N = 1$$

$$i=2 \quad P_3' - P_2' = \frac{q}{p} (P_2 - P_1) = \left(\frac{q}{p}\right)^2 P_1$$

$$\frac{P_i - P_{i-1}'}{\overline{P_i - P_1}} = \left(\frac{q}{p}\right)^{i-1} P_1$$

$$\overline{P_i - P_1} = P_1 \left( \frac{q}{p} + \left(\frac{q}{p}\right)^2 + \dots + \left(\frac{q}{p}\right)^{i-1} \right)$$

$$\Rightarrow P_i = P_1 \left( 1 + \frac{q}{p} + \left(\frac{q}{p}\right)^2 + \dots + \left(\frac{q}{p}\right)^{i-1} \right)$$

$$P_i = \begin{cases} \frac{1 - \left(\frac{q}{p}\right)^i}{1 - \frac{q}{p}} P_1 & \text{if } \frac{q}{p} \neq 1 \\ i P_1 & \text{if } \frac{q}{p} = 1 \end{cases}$$

$$\because P_N = 1 \Rightarrow P_1 = \begin{cases} \frac{1 - \frac{q}{p}}{1 - \left(\frac{q}{p}\right)^N} & \text{if } \frac{q}{p} \neq 1 \\ \frac{1}{N} & \text{if } \frac{q}{p} = 1 \end{cases}$$

$$P_1 = \begin{cases} \frac{1 - \left(\frac{q}{p}\right)^i}{1 - \left(\frac{q}{p}\right)^N} & \text{if } \frac{q}{p} \neq 1 \Leftrightarrow p \neq \frac{1}{2} \\ \frac{i}{N} & \text{if } \frac{q}{p} = 1 \Leftrightarrow p = \frac{1}{2} \end{cases}$$

$N \rightarrow \infty$

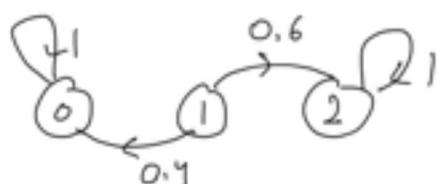
$$P_i = \begin{cases} 0, & p = \frac{1}{2} \\ 0, & p < \frac{1}{2} \Leftrightarrow \frac{v}{p} > 1 \\ 1 - \left(\frac{v}{p}\right)^i, & p > \frac{1}{2} \Leftrightarrow \frac{v}{p} < 1 \end{cases}$$

—x—

Example 1 ① tpm

$$P = \begin{matrix} & 0 & 1 & 2 \\ 0 & 1 & 0 & 0 \\ 1 & 0.4 & 0 & 0.6 \\ 2 & 0 & 0 & 1 \end{matrix}$$

Starting in 1, determine the prob. that M.C ends in state 0.



$$p = 0.6, v = 0.4 \quad p \neq \frac{1}{2}$$

$$i=1, N=2 \quad \frac{v}{p} = \frac{2}{3}$$

$$\rightarrow 1 - P_1 = 1 - \frac{1 - \left(\frac{2}{3}\right)}{1 - \left(\frac{2}{3}\right)^2} = 0.4$$

② The probability of the thrower winning in the dice game called "Craps" is  $p = 0.49$ . Suppose Player A is the thrower and begins the game with \$5, and Player B, the opponent, begins with \$10. What is the probability that player A goes bankrupt before player B? Assume that the bet is \$1 per round.

$$i = 5, N = 15$$

$$p = 0.49, v = 1-p = 0.51$$

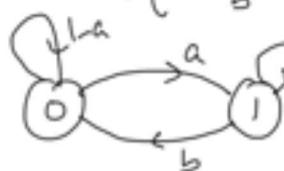
$$1 - P_E = 1 - \underline{1 - \left(\frac{0.51}{0.49}\right)^5}$$

$$1 - \left(\frac{0.51}{0.49}\right)^{15}$$

—x—

Limiting prob.

$$\text{try } P = \begin{bmatrix} 0 & 1 \\ 1-a & a \\ b & 1-b \end{bmatrix}, \quad 0 \leq a, b \leq 1$$



$$a = b = 0 \text{ or } a = b = 1$$

we take separately

$$P^n = \begin{bmatrix} \frac{b + a(1-a-b)^n}{a+b} & \frac{a - a(1-a-b)^n}{a+b} \\ \frac{b - b(1-a-b)^n}{a+b} & \frac{a + b(1-a-b)^n}{a+b} \end{bmatrix}$$

$$\text{Set } P_{00} = 1-a, \quad P_{01} = a, \quad P_{10} = b, \quad P_{11} = 1-b$$

$$\begin{aligned} P_{00}^{(n)} &= P_{00}^{(n-1)} P_{00} + P_{01}^{(n-1)} P_{10} \\ &= (1-a) P_{00}^{(n-1)} + b P_{01}^{(n-1)} \quad \because P_{00}^{(n-1)} + P_{01}^{(n-1)} = 1 \\ &= (1-a) P_{00}^{(n-1)} + b (1 - P_{00}^{(n-1)}) \\ &= b + (1-a-b) \underbrace{P_{00}^{(n-1)}}_{n \geq 1} \end{aligned}$$

$$\begin{aligned} &= b + b(1-a-b) + b(1-a-b)^2 + \dots + b(1-a-b)^{n-2} \\ &\quad + \underbrace{P_{00}^{(n-1)}(1-a-b)^{n-1}}_{(1-a)} \end{aligned}$$

$$= b \sum_{k=0}^{n-2} (1-a-b)^k + \underbrace{(1-a)(1-a-b)^{n-1}}_{\checkmark}$$

$$\frac{1 - (1-a-b)^{n-1}}{1 - (1-a-b)} = \frac{\cancel{1 - (1-a-b)^{n-1}}}{\cancel{a+b}}$$

$$= \frac{b}{a+b} + \frac{a(1-a-b)}{a+b}$$

$$\begin{aligned} & \frac{(1-a-b)^{n-1}}{a+b} \left[ \frac{1-a-b}{a+b} \right] \\ & \frac{a+b-a^2-ab}{a+b} \\ & = \frac{a(1-a-b)}{a+b} \end{aligned}$$

$$\lim_{n \rightarrow \infty} P^n = \begin{pmatrix} \frac{b}{a+b} & \frac{a}{a+b} \\ \frac{b}{a+b} & \frac{a}{a+b} \end{pmatrix} = \begin{pmatrix} \pi_0 & \pi_1 \\ \pi_0 & \pi_1 \end{pmatrix}$$

$$\begin{aligned} \lim_{n \rightarrow \infty} p_i^{(n)} &= \lim_{n \rightarrow \infty} P(X_n=i) \\ p_i &= \tilde{p}^{(0)} P^n \xrightarrow{n \rightarrow \infty} \begin{pmatrix} & \\ & \end{pmatrix} = \left( \frac{b}{a+b}, \frac{a}{a+b} \right) \\ &= (\pi_0, \pi_1) \end{aligned}$$



$$P = \begin{pmatrix} 0 & 1 \\ 0 & 1 \\ 1 & 0 \end{pmatrix} \quad P^2 = \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix} \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix} = \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}$$

$a=b=1$   
periodic with period 2  
 $d(P)=2$

$$P^3 = P ; P^4 = P^2$$

$$\begin{aligned} &(\alpha, 1-\alpha) \quad \tilde{p}^{(1)} = \tilde{p}^{(0)} P = (1-\alpha, \alpha) \\ &\tilde{p}^{(1)} = \tilde{p}^{(0)} P = (\alpha, 1-\alpha) \\ &\text{limiting prob DNE.} \end{aligned}$$



$$a=b=0$$

$$P = \begin{pmatrix} 0 & 1 \\ 1 & 0 \\ 0 & 1 \end{pmatrix}$$

$$\begin{matrix} 0 \\ 0 \\ 1 \end{matrix} \quad \begin{matrix} 0 \\ 1 \\ 0 \end{matrix}$$

$$P^n = I$$

→ Finite state, irreducible, aperiodic (period 1),  
limiting prob exist

$$\pi_j = \lim_{n \rightarrow \infty} p_j^{(n)}$$

$$p_j^{(n)} = P(X_n=j)$$

$$\left( \underline{b}^{(n)} = \underline{b}^{(n-1)} P \right) \rightarrow \underline{P_j^{(n)}} = \sum_i b_i^{(n-1)} P_{ij}$$

Take limit

$$\left. \begin{array}{l} \underline{\pi}_j = \sum_i \pi_i P_{ij} \\ \sum_i \pi_i = 1 \end{array} \right\} \rightarrow \underline{\pi} = \underline{\pi} P$$

$$\sum_i \pi_i = 1$$

$$\underline{\pi} = (\pi_0, \pi_1, \dots)$$

Regular tpm:

Given tpm  $P$  is regular if  $P^k$  has all element  $> 0$  for some  $k$ .

$$\text{regular } N=3 \quad P = \begin{pmatrix} + & + & 0 \\ + & 0 & + \\ + & 0 & + \end{pmatrix} \quad P^2 = \begin{pmatrix} + & + & 0 \\ + & 0 & + \\ + & 0 & + \end{pmatrix} \begin{pmatrix} + & + & 0 \\ + & 0 & + \\ + & 0 & + \end{pmatrix}$$

state  
 $N$

$$P \quad \underline{P^N}$$

$$= \begin{pmatrix} + & + & + \\ + & + & + \\ + & + & + \end{pmatrix}$$

$$N=2 \quad P^4$$

$$P = \begin{pmatrix} 0 & + \\ + & 0 \end{pmatrix}; P^2 = \begin{pmatrix} + & 0 \\ 0 & + \end{pmatrix}$$

not regular

$$P^4, P^8, P^2$$

Show that finite state aperiodic irreducible  
 $M \subseteq S$  is regular and recurrent

Thm  $\circledast$  Let  $P$  regular tpm  $S = \{0, 1, \dots, N\}$ . Then

limiting prob  $\underline{\pi} = (\pi_0, \pi_1, \dots, \pi_N)$  is unique

Set " to equations

$$\underline{\pi} = \underline{\pi} P \quad \left\{ \begin{array}{l} \pi_j = \sum_{k=0}^N \pi_k P_{kj}, \quad j=0, 1, \dots, N \\ \sum_{k=0}^N \pi_k = 1 \end{array} \right.$$

Since  $M \subseteq S$  regular we have  $\lim_{n \rightarrow \infty} P^n$

$$\lim_{n \rightarrow \infty} P_{ij}^{(n)} = \pi_j \quad ; \quad \sum_{j=0}^N \pi_j = 1$$

$$P_{ij}^{(n)} = \sum_{k=0}^N P_{ik}^{(n-1)} P_{kj}$$

$$\text{Take Limit as } n \rightarrow \infty \quad P_{ij}^{(n)} \rightarrow \pi_j \text{ & } P_{ik}^{(n-1)} \rightarrow \pi_k$$

$$\pi_j = \sum_{k=0}^N \pi_k P_{kj}$$

$$\sum_{j=0}^N \pi_j = 1$$

T.S. sol<sup>n</sup> is unique

$$\exists x_0, \dots, x_N \text{ st.} \quad \xrightarrow{\quad (1) \quad}$$

$$x_j = \sum_{k=0}^N x_k P_{kj}, \quad j = 0, 1, \dots, N$$

$$P_{jl} \downarrow$$

$$\sum_{j=0}^N x_j P_{jl} = \sum_{j=0}^N \sum_{k=0}^N x_k P_{kj} P_{jl}$$

$\swarrow \text{using (1)}$

$$x_l = \sum_{k=0}^N x_k \left( \sum_{j=0}^N P_{kj} P_{jl} \right) \quad \searrow P_{kl}^{(2)}$$

$$\Rightarrow x_l = \sum_{k=0}^N x_k P_{kl}^{(2)}$$

$$x_l = \sum_{k=0}^N x_k P_{kl}^{(n)}, \quad l = 0, 1, \dots, N$$

$$\text{as } n \rightarrow \infty \quad P_{kl}^{(n)} \rightarrow \pi_l$$

$$x_l = \sum_{k=0}^N x_k \pi_l = \pi_l \left( \sum_{k=0}^N x_k \right) = \pi_l$$

$\xrightarrow{\quad 1 \quad}$

Example An NCD system has disjoint classes

$E_0$  (no discount),  $E_1$  (20% discount) and  $E_2$  (40% discount)

Movement in the system is determined by the rule

when in class  $i$  -  $\rightarrow$  class  $j$  -  $\rightarrow$  class  $k$  -  $\dots$

... steps from one discount level (or stays in  $E_0$ ) with one claim in a year, and return to a level if no discount if more than one claim is made. A claim free year results in a step up to a higher discount level (or one remains in class  $E_2$  if already there).

NCD class	$E_0$	$E_1$	$E_2$
% discount	6	20	40
Annual premium	100	80	60 ✓

- If we suppose that for someone in this scheme the prob. of one claim in a year is 0.2 while the prob. of two or more claims is 0.1. Find then
- (i) In long run what proportion of time is the person in each of the discount classes
  - (ii) Find the annual premium paid.

$$E_i \equiv i$$

$$i=0,1,2$$

$$P = \begin{pmatrix} 0 & 1 & 2 \\ 0 & 0.3 & 0.7 \\ 1 & 0.3 & 0 \\ 2 & 0.1 & 0.2 \end{pmatrix}$$



Class = {0, 1, 2} irreducible, aperiodic finite state

limiting prob. exist and same as stationary state prob.

$$\sum_{i=0}^2 \pi_i = 1$$

$$\underline{\pi} = (\pi_0, \pi_1, \pi_2)$$

$$\rightarrow \sum_{i=0}^2 \pi_i = 1$$

$$\Rightarrow \sum \pi_i = 0.3\pi_0 + 0.3\pi_1 + 0.1\pi_2$$

$$\begin{cases} \pi_1 = 0.7\pi_0 + 0.2\pi_2 \\ \pi_0 + \pi_1 + \pi_2 = 1 \end{cases}$$

$$\Rightarrow \pi_0 = 0.1860, \pi_1 = 0.2442, \pi_2 = 0.5698$$

on annual premium paid

$$= 0.1860 \times 100 + 60 \times 0.2442 + 60 \times 0.5698$$

$$= 72.324$$

Doubly stochastic matrix

tpm P

$$\sum_k p_{ik} = \sum_i p_{ik} = 1$$

$$\left( \begin{array}{ccc|c} & & & p_1 \\ & & & p_2 \\ & & & p_3 \\ \hline 1 & & & \\ 2 & & & \\ 3 & & & \\ \hline & 1 & 1 & 1 \end{array} \right)$$

$$\text{eg} \quad P = \begin{pmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} \end{pmatrix}$$

let P doubly stochastic, regular

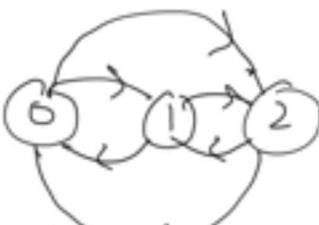
$$\underset{\text{limiting prob}}{\lim} \pi^s \left( \frac{1}{N}, \frac{1}{N}, \dots, \frac{1}{N} \right) \xrightarrow{ss [0, 1, \dots, N-1]}$$

$$\begin{cases} \pi_j^s = \sum_k \pi_k p_{kj} \\ \sum_k \pi_k = 1 \end{cases}$$

$$\frac{1}{N} = \sum_k \frac{1}{N} p_{kj} = \frac{1}{N} \left( \sum_k p_{kj} \right)$$

$\therefore \pi^s = \left( \frac{1}{N}, \dots, \frac{1}{N} \right)$  is unique since  
it's uniformly doubly stochastic tpm

$$\begin{matrix} \text{tpm} & \xrightarrow{x} & 0 & 1 & 2 \\ \begin{matrix} ss [0, 1, 2] \\ N=3 \end{matrix} & | & \begin{pmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & 0 & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} & 0 \end{pmatrix} \end{matrix}$$



$$\pi = \left( \frac{1}{3}, \frac{1}{3}, \frac{1}{3} \right)$$

irreducible, aperiodic,  
initial state  $\pi$ , s.  
doubly stochastic

e.g. doubly stochastic tpm may or may not be symmetric

not symmetric

$$P = \begin{pmatrix} 7/12 & 0 & 5/12 \\ 2/12 & 6/12 & 4/12 \\ 3/12 & 6/12 & 3/12 \end{pmatrix} \rightarrow P = \begin{pmatrix} 0 & 1 & 2 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \end{pmatrix}$$

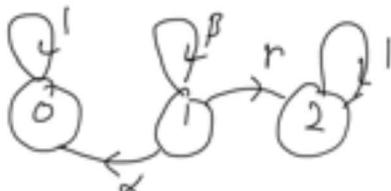
regular

Simple first step analysis

$$P = \begin{pmatrix} 0 & 1 & 2 \\ 1 & 0 & 0 \\ \alpha & \beta & r \\ 0 & 0 & 1 \end{pmatrix}$$

$$\alpha + \beta + r = 1$$

$$0 < \alpha, \beta, r < 1$$



$$T = \min \{n : X_n = 0 \text{ or } X_n = 2\}$$

↳ time of absorption of the

mean time to absorption  $\omega = E(T | X_0 = 1)$  process  
starting from state 1

$$u_1 = P(X_T = 0 | X_0 = 1)$$

0,2 absorbing  
1 transient  
 $S = \{0, 1, 2\}$   
initial

↳ Prob. of ultimate absorption in state 0, starting from state 1.

$$u_1 = P(X_T = 0 | X_0 = 1)$$

$$u_1 = \sum_{k \in S} P(X_T = 0, X_1 = k | X_0 = 1)$$

$$u_1 = \sum_{k \in S} P(X_T = 0 | X_1 = k, X_0 = 1) P(X_1 = k | X_0 = 1)$$

$$u_1 = \sum_{k \in S} P(X_T=0 | X_1=k) p_{1k}$$

$$u_1 = \underbrace{P(X_T=0 | X_1=0)p_{10}}_1 + \underbrace{P(X_T=0 | X_1=1)p_{11}}_{u_1} + \underbrace{P(X_T=0 | X_1=2)p_{12}}_{\beta}$$

$$u_1 = \alpha + \beta u_1 \Rightarrow (1-\beta)u_1 = \alpha \Rightarrow u_1 = \frac{\alpha}{1-\beta}$$

$$\nu = E(T | X_0=1)$$

$$= 1 + \underbrace{\alpha \times 0 + \beta \times 0}_{X_1=0 \text{ or } X_1=2} + \underbrace{\beta \times \nu}_{X_1=1}$$

$$\Rightarrow \nu = 1 + \beta \nu \Rightarrow (1-\beta)\nu = 1 \Rightarrow \nu = \frac{1}{1-\beta}$$

eg

$$u_1 = P(X_T=0 | X_0=1) \rightarrow 1 \quad \begin{pmatrix} 0 & 1 & 2 & 3 \\ 1 & 0 & 0 & 0 \\ p_{00} & p_{11} & p_{12} & p_{13} \\ p_{20} & p_{21} & p_{22} & p_{23} \\ 0 & 0 & 0 & 1 \end{pmatrix} \quad \begin{array}{l} 0, 3 \text{ absorbing} \\ 1, 2 \text{ transient} \\ 0 < p_{ij} < 1 \end{array}$$

$$\begin{cases} u_1 = p_{00} + p_{11}u_1 + p_{12}u_2 \\ u_2 = p_{20} + p_{21}u_1 + p_{22}u_2 \end{cases}$$

$$\nu_i = E(X_T | X_0=i), i=1,2$$

$$\nu_1 = 1 + p_{11}\nu_1 + p_{12}\nu_2$$

$$\nu_2 = 1 + p_{21}\nu_1 + p_{22}\nu_2$$

Finite state m.c.  $S = \{0, 1, \dots, N\}$

$0, 1, \dots, N-1 \rightarrow \text{transient}$

$N, \dots, N \rightarrow \text{absorbing}$

$$\text{trpm } P = \begin{pmatrix} Q & R \end{pmatrix} \quad \begin{array}{c|cc} & \multicolumn{2}{c}{N+1} \\ \hline & \alpha & N-\alpha \end{array}$$

$$\begin{pmatrix} 0 & I \end{pmatrix} \begin{matrix} N+1 \\ N-N+1 \end{matrix} \begin{matrix} O_{N \times N} & R_{N \times N-N+1} \\ O_{N-N+1 \times N} & I_{N-N+1 \times N-N+1} \end{matrix}$$

$$k \in \{0, \dots, N\}, \quad i \in \{0, 1, \dots, N-1\}$$

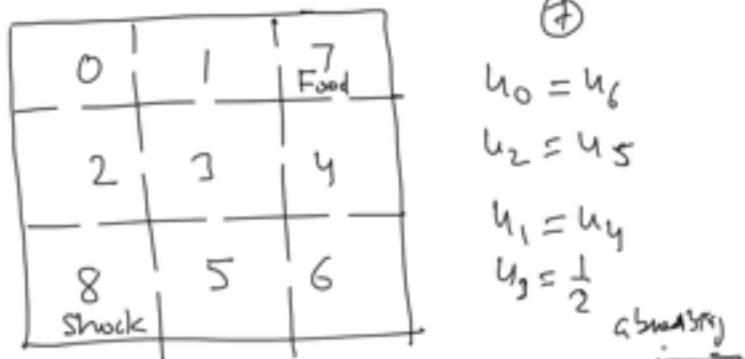
$$u_{ik} = u_i = P(\text{abnormal } k \mid X_0=i)$$

$$= P_{ik} x_1 + \sum_{\substack{j=0 \\ j \neq k}}^N P_{ij} x_0 + \sum_{j=0}^{N-1} P_{ij} u_j$$

$$= P_{ik} + \sum_{j=0}^{N-1} P_{ij} u_j, \quad i \in \{0, 1, \dots, N-1\}$$

$$v_i = 1 + \sum_{j=0}^{N-1} P_{ij} v_j \quad \text{for } i \in \{0, 1, \dots, N-1\}$$

Example



tpm	0	1	2	3	4	5	6	7	8
0		$\frac{1}{2}$	$\frac{1}{2}$						
1		$\frac{1}{3}$			$\frac{1}{3}$			$\frac{1}{3}$	
2		$\frac{1}{3}$			$\frac{1}{3}$			$\frac{1}{3}$	
3			$\frac{1}{4}$	$\frac{1}{4}$		$\frac{1}{4}$	$\frac{1}{4}$		
4					$\frac{1}{3}$			$\frac{1}{3}$	$\frac{1}{3}$
5					$\frac{1}{3}$			$\frac{1}{3}$	$\frac{1}{3}$
6						$\frac{1}{2}$	$\frac{1}{2}$		
7								1	
8								1	

abnormality

$$u_{i=7} = u_i$$

$$u_0 = \frac{1}{2} u_1 + \frac{1}{2} u_2 \quad 7 \quad u_0 = \frac{1}{2}$$

$$u_1 = \frac{1}{3}u_0 + \frac{1}{3}u_3 + \frac{1}{3} \quad \left| \begin{array}{l} u_1 = \frac{2}{3} \\ u_2 = \frac{1}{3} \\ u_3 = \frac{1}{2} \end{array} \right.$$

Mean time spent in transient state!

## Finite state M.C.

$T = \{1, 2, \dots, t\}$  set of transient states

$$P_T = \begin{bmatrix} P_{11} & \cdots & P_{1t} \\ \vdots & & \vdots \\ P_{t1} & \cdots & P_{tt} \end{bmatrix}$$

$$i, j \in T$$

$s_{ij}$  = expected # of times period the M.C. is in state  $j$  given that it starts in state  $i$

$$I_{n,j} = \sum_{i=0}^n a_i x_{n-i} \quad S_{ij} = \begin{cases} 1 & i=j \\ 0 & \text{o.w.} \end{cases}$$

$$s_{ij} = s_{ij} + E\left(\sum_{n=1}^{\infty} I_{n,j} | X_0=i\right)$$

$$= \delta_{ij} + \sum_{n=1}^{\infty} \left( E(I_{h,j} | X_0=i) \right)$$

$$1 \times P(X_1 = j | X_0 = i) + 0 \times P(X_1 \neq j | X_0 = i)$$

$$P_{ij}^{(n)}$$

$$= s_{ij} + \sum_{n=1}^{\infty} P_{ij}^{(n)} \quad \text{_____} \quad (\ast \ast)$$

$$= \delta_{ij} + \sum_{n=1}^{\infty} \sum_k p_{ik} \underbrace{p_{kj}^{(n-1)}}_{\substack{i,j \in T \\ k \in S}},$$

$$= \delta_{ij} + \sum_k P_{ik} \left( \sum_{n=1}^{\infty} P_{kj}^{(n-1)} \right)$$

$$\delta_{kj} + \sum_{n=2}^{\infty} P_{kj}^{(n-1)}$$

$$\underbrace{\delta_{kj} + \sum_{n=1}^{\infty} p_{kj}^{(n)}}_{\downarrow} \quad \text{def} \quad \delta_{kj}$$

$$= \delta_{ij} + \sum_k p_{ik} \delta_{kj}$$

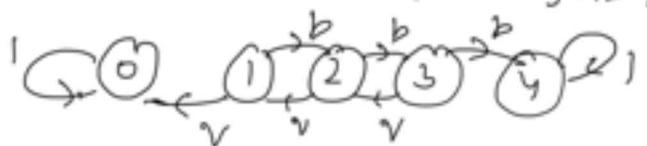
$$\delta_{ij} = \delta_{ij} + \sum_{k=1}^t p_{ik} \delta_{kj}, \quad \leftarrow$$

since it is impossible to go from a recurrent state to a transient state  $\Rightarrow \delta_{kj} = 0$ , when  $k$  is a recurrent state  
 $S = (I - P_T)^{-1}$

$$S = I + P_T S$$

$$\Rightarrow (I - P_T) S = I \Rightarrow S = (I - P_T)^{-1}$$

Example Gambler ruin problem  $p=0.4, N=4$



transient 1, 2, 3

$$P = \begin{bmatrix} 0 & 1 & 2 & 3 & 4 \\ 1 & 0 & 0 & 0 & 0 \\ 2 & 0 & 0 & p & 0 \\ 3 & 0 & 0 & 0 & p \\ 4 & 0 & 0 & 0 & 1 \end{bmatrix} \xrightarrow{\text{absorbing}} P_T$$

$$I = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix}; P_T = \begin{bmatrix} 0 & 0.4 & 0 \\ 0.6 & 0 & 0.4 \\ 0 & 0.6 & 0 \end{bmatrix}$$

$$(I - P_T) = \begin{bmatrix} 1 & -0.4 & 0 \\ -0.6 & 1 & 0 \\ 0 & -0.6 & 1 \end{bmatrix}$$

$$(\delta_{ij}) = S = (I - P_T)^{-1} = \begin{bmatrix} 1 & 1.46 & 2 \\ 2 & 1.15 & 1.92 \\ 1 & 0.76 & 0.76 \end{bmatrix}^{-1}$$

$$3 \left[ 0.69 \quad 1.15 \quad 1.46 \right]$$

$$\delta_{2,3} = 0.76; \quad \underline{\delta_{2,1} = 1.15}$$

$f_{ij}$ : prob. that M.C ever makes a transition into state  $j$  given that it starts in state  $i$

$$\begin{aligned} f_{2,1} &= f_{2,1}^{(1)} + f_{2,1}^{(2)} + f_{2,1}^{(3)} + \dots \\ &= \nu + p\nu^2 + p^2\nu^3 + p^3\nu^4 + \dots \\ &= \nu + p\nu [\nu + p\nu^2 + p^2\nu^3 + \dots] \\ &= \nu + p\nu f_{2,1} \end{aligned}$$

$$\Rightarrow f_{2,1} = \frac{\nu}{1-p\nu} = \frac{0.6}{1-0.4 \times 0.6} = 0.78$$



$$f_{2,1} = 1 - \frac{1 - \left(\frac{\nu}{p}\right)^4}{1 - \left(\frac{\nu}{p}\right)^3} = 1 - \frac{1 - \left(\frac{0.6}{0.4}\right)^4}{1 - \left(\frac{0.6}{0.4}\right)^3} = 0.78$$

$$\delta_{ij} = E(\text{time}_{ij} | \text{start } i)$$

$$\begin{aligned} &= \underbrace{E(\text{time}_{ij} | \text{start } i, \text{ever transit to } j)}_{\delta_{ij}} \cdot f_{ij} \\ &\quad + \underbrace{E(\text{time}_{ij} | \text{start } i, \text{never transit to } j)}_{\delta_{ij}'} \cdot (1-f_{ij}) \\ &\quad \downarrow \delta_{ij}' \\ &= (\delta_{ij} + \lambda_{jj}) \cdot f_{ij} + \delta_{ij} (1-f_{ij}) \end{aligned}$$

$$= (\delta_{ij} + \lambda_{jj}) \cdot f_{ij} + \delta_{ij} (1-f_{ij})$$

$$= \delta_{ij} + \lambda_{jj} f_{ii}$$

$$\Rightarrow f_{ij} = \frac{\delta_{ij} - \delta_{jj}}{\lambda_{jj}} \quad \delta_{2,1} = 1.15, \quad \delta_{2,2} = 0$$

$$f_{2,1} = \frac{1.15 - 0}{1.46} = 0.78 \quad \lambda_{1,1} = 1.46$$

Particular Case:

$$P_S = \begin{pmatrix} Q & R \\ 0 & I \end{pmatrix}$$

$0, 1, \dots, n-1$  terminated

$n \rightarrow N$  absorption

$$S = I + QS \Rightarrow S = (I - Q)^{-1}$$

$$\lambda_{ij} = \delta_{ij} + \sum_{k=0}^{n-1} P_{ik} \delta_{kj} \quad ; i, j = 0, 1, \dots, n-1$$

$T \rightarrow$  time of absorption

$$T = \min \{ n : 0 \leq X_n \leq N \}$$

$$\delta_{ij} = E \left( \sum_{n=0}^{T-1} 1(X_n=j) \mid X_0=i \right)$$

$$\nu_i = E(T \mid X_0=i)$$

$$\sum_{j=0}^{n-1} \sum_{n=0}^{T-1} 1(X_n=j) = \sum_{n=0}^{T-1} \underbrace{\sum_{j=0}^{n-1} 1(X_n=j)}_{1} \leq T$$

$$\sum_{j=0}^{n-1} \delta_{ij} = E \left( \sum_{j=0}^{n-1} \sum_{n=0}^{T-1} 1(X_n=j) \mid X_0=i \right)$$

$$= E(T \mid X_0=i)$$

$$= \nu_i$$

$$\boxed{\sum_{j=0}^{n-1} \delta_{ij} = \nu_i}$$

reducible M.M.

$$(1) \quad t_{pm}$$

$$\pi = (\pi_0, \pi_1)$$

$$\pi = \pi P$$

$$\pi_0 + \pi_1 = 1$$

$$\pi_0 = \frac{1}{2} \pi_0 + \frac{1}{4} \pi_1$$

$$\pi_1 = \frac{1}{4} \pi_0 + \frac{3}{4} \pi_1$$

$$\pi_0 = \frac{1}{3}, \pi_1 = \frac{2}{3}$$

$$P = \begin{bmatrix} 0 & 1 & 2 & 3 \\ \frac{1}{2} & \frac{1}{2} & 0 & 0 \\ \frac{1}{4} & \frac{3}{4} & 0 & 0 \\ 0 & 0 & \frac{1}{3} & \frac{2}{3} \\ 0 & 0 & \frac{2}{3} & \frac{1}{3} \end{bmatrix}$$

Class  $[0, 1, 2, 3]$   
 doubly stochastic  
 regular  
 recurrent aperiodic  
 recurrent aperiodic

$$\pi_0 = \frac{1}{2}, \pi_1 = \frac{1}{2}$$

$$P = \begin{bmatrix} P_1 & 0 \\ 0 & P_2 \end{bmatrix}, P^2 = \begin{bmatrix} P_1 & 0 \\ 0 & P_2 \end{bmatrix} \begin{bmatrix} P_1 & 0 \\ 0 & P_2 \end{bmatrix} = \begin{bmatrix} P_1^2 & 0 \\ 0 & P_2^2 \end{bmatrix}$$

$$P^n = \begin{bmatrix} P_1^n & 0 \\ 0 & P_2^n \end{bmatrix}$$

$$\lim_{n \rightarrow \infty} P^n \leq \begin{pmatrix} \lim_{n \rightarrow \infty} P_1^n & 0 \\ 0 & \lim_{n \rightarrow \infty} P_2^n \end{pmatrix} \leq \begin{pmatrix} \pi_0^{(1)} \pi_1^{(1)} & 0 \\ 0 & \pi_0^{(2)} \pi_1^{(1)} \end{pmatrix}$$

$$\lim_{n \rightarrow \infty} P^n = \begin{bmatrix} 0 & \frac{1}{3} & \frac{2}{3} & 0 & 0 \\ 1 & \frac{1}{3} & \frac{2}{3} & 0 & 0 \\ 2 & 0 & 0 & \frac{1}{2} & \frac{1}{2} \\ 3 & 0 & 0 & \frac{1}{2} & \frac{1}{2} \end{bmatrix}$$

$$\pi_0 = \frac{1}{2} \pi_0 + \frac{1}{4} \pi_1$$

$$\pi_0 + \pi_1 = 1$$

$$(2)$$

$$\pi_0^{(1)} = \frac{1}{3}$$

$$\pi_1^{(1)} = \frac{2}{3}$$

$$\rightarrow \frac{1}{3} \begin{bmatrix} 0 & 1 & 2 & 3 \\ \frac{1}{2} & \frac{1}{2} & 0 & 0 \\ \frac{1}{4} & \frac{3}{4} & 0 & 0 \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{4} & \frac{1}{4} \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

$$C_1 = u_2 = \frac{1}{4} + \frac{1}{4} + \frac{1}{4} u_2$$

$$\Rightarrow u_2 = \frac{2}{3}$$



Class  $[0, 1, 2, 3]$   
 recurrent aperiodic  
 transient absorbing

$$\lim_{n \rightarrow \infty} P^n = 0 \begin{bmatrix} 0 & 1 & 2 & 3 \\ \frac{1}{3} & \frac{2}{3} & 0 & 0 \\ \frac{1}{3} & \frac{2}{3} & 0 & 0 \\ \frac{2}{3} \times \frac{1}{3} & \frac{2}{3} \times \frac{2}{3} & 0 & \frac{1}{3} \\ 0 & 0 & 0 & 1 \end{bmatrix}$$

(3)

$$P^n = 0 \begin{bmatrix} 0 & 1 & 2 & 3 & 4 & 5 \\ \frac{1}{2} & \frac{1}{2} & 0 & 0 & 0 & 0 \\ \frac{1}{3} & \frac{2}{3} & 0 & 0 & 0 & 0 \\ \frac{1}{3} & 0 & 0 & \frac{1}{3} & \frac{1}{6} & \frac{1}{6} \\ \frac{1}{6} & \frac{1}{6} & \frac{1}{6} & 0 & \frac{1}{3} & \frac{1}{6} \\ 0 & 0 & 0 & 0 & 0 & 1 \\ 0 & 0 & 0 & 0 & 1 & 0 \end{bmatrix}$$

$$u_2 = \frac{1}{3} + \frac{1}{3} u_3$$

$$u_3 = \frac{2}{6} + \frac{1}{6} u_2$$

$$u_2 \quad u_3$$

$$\frac{8}{17} \quad \frac{7}{17}$$

$$\pi_0, \pi_1, \pi_2, \pi_3, \pi_4, \pi_5$$

$$C_1 = \{0, 1\}, C_2 = \{2, 3\}, C_3 = \{4, 5\}$$

recurrent, aperiodic      transient      recurrent period=2

$$\lim_{n \rightarrow \infty} P^n = 0 \begin{bmatrix} 0 & 1 & 2 & 3 & 4 & 5 \\ \frac{2}{5} & \frac{3}{5} & 0 & 0 & 0 & 0 \\ \frac{2}{5} & \frac{3}{5} & 0 & 0 & 0 & 0 \\ \frac{8}{17} \times \frac{2}{5} & \frac{8}{17} \times \frac{3}{5} & 0 & 0 & \times & \times \\ \frac{2}{17} \times \frac{2}{5} & \frac{2}{17} \times \frac{3}{5} & 0 & 0 & \times & \times \\ 0 & 0 & 0 & 0 & \times & \times \\ 0 & 0 & 0 & 0 & \times & \times \end{bmatrix}$$

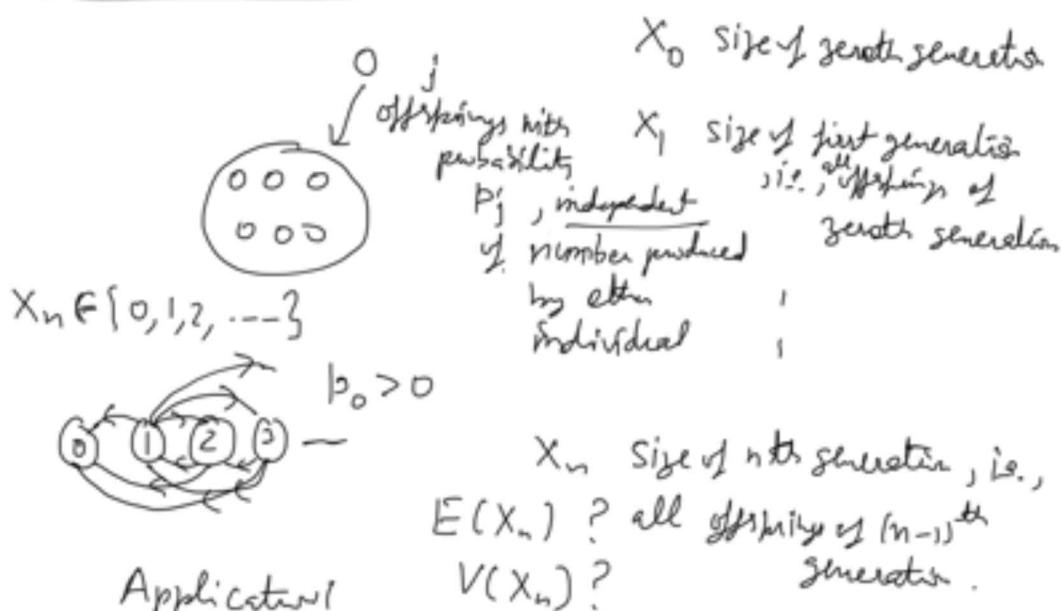
For time av X DNE

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{m=0}^{n-1} P^m = 0 \begin{bmatrix} 2/5 & 3/5 & 0 & 0 & 0 & 0 \\ 2/5 & 3/5 & 0 & 0 & 0 & 0 \\ \frac{8}{17} \times \frac{2}{5} & \frac{8}{17} \times \frac{3}{5} & 0 & 0 & \frac{9}{17} \times \frac{1}{2} & \frac{9}{17} \times \frac{1}{2} \\ \frac{2}{17} \times \frac{2}{5} & \frac{2}{17} \times \frac{3}{5} & 0 & 0 & \frac{10}{17} \times \frac{1}{2} & \frac{10}{17} \times \frac{1}{2} \end{bmatrix}$$

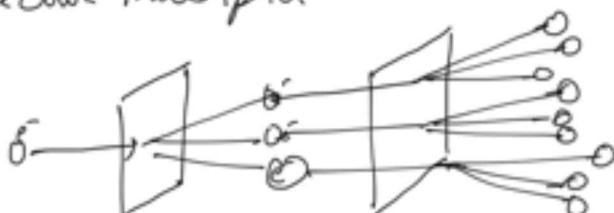
$$4 \begin{pmatrix} & & & & 17^0 & 17^2 \\ 0 & 0 & 0 & 0 & \frac{1}{2} & \frac{1}{2} \\ 5 & 0 & 0 & 0 & 0 & \frac{1}{2} \end{pmatrix}$$

Endemic → inderogable, specific, +ve recurrent

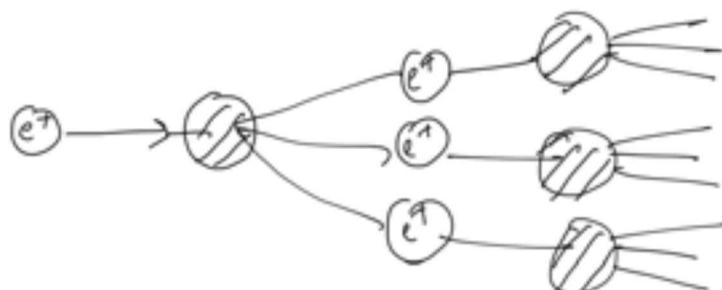
## Branching process:



## Electron multiplier



## Nuclear chain reaction



Suppose  $X_0 = 1$  Survival of family name?

mean # of offspring of a single individual:  $\mu = \sum_{j=0}^{\infty} j p_j$

$$\text{Var. } \sigma^2 = \sum_{j=0}^{\infty} (j-\mu)^2 p_j$$

$(n-1)^{\text{st}}$  general term      1    2    ...     $x_{n-1}$

$$\text{size of } n^{\text{th}} \text{ generation } X_n = \sum_{i=1}^n Z_i$$

$Z_i$  # of offspring of  $i^{\text{th}}$  individual of  $(n-1)^{\text{th}}$  generation  $E(Z_i) = \mu, V(Z_i) = \sigma^2$

$$\begin{aligned} E(X_n) &= E\left(E(X_n | X_{n-1})\right) \\ &= E\left(E\left(\sum_{i=1}^{X_{n-1}} Z_i | X_{n-1}\right)\right) \end{aligned}$$

$$\begin{aligned} &= E(X_{n-1} \mu) \\ &= \mu E(X_{n-1}) \\ &= \mu^2 E(X_{n-2}) \\ &= \mu^n E(X_0) \\ &= \mu^n \end{aligned}$$

$\overbrace{E\left(\sum_{i=1}^{X_{n-1}} Z_i | X_{n-1}=x\right)}$

$= E\left(\sum_{i=1}^x Z_i\right) = \sum_{i=1}^x \underbrace{E(Z_i)}_{\mu}$

$= x\mu \quad \checkmark$

$$V(X_n) = \underbrace{E(V(X_n | X_{n-1}))}_{\mu_{n-1} \sigma^2} + \underbrace{V(E(X_n | X_{n-1}))}_{\mu_{n-1} \mu}$$

$$\begin{aligned} &= E(X_{n-1} \sigma^2) + V(X_{n-1} \mu) \\ &= \sigma^2 E(X_{n-1}) + \mu^2 V(X_{n-1}) \\ &= \sigma^2 \mu^{n-1} + \mu^2 V(X_{n-1}) \end{aligned}$$

$\overbrace{V\left(\sum_{i=1}^{X_{n-1}} Z_i | X_{n-1}=x\right)}$

$= V\left(\sum_{i=1}^x Z_i\right)$

$= \sum_{i=1}^x \underbrace{V(Z_i)}_{\sigma^2}$

$= x \sigma^2$

$$= \sigma^2 \mu^{n-1} + \mu^2 V(X_{n-1}) \quad \text{---} \otimes$$

$$= \sigma^2 \mu^{n-1} + \mu^2 [\sigma^2 \mu^{n-2} + \mu^2 V(X_{n-2})]$$

$$= \sigma^2 [\mu^{n-1} + \mu^n] + \mu^4 V(X_{n-1})$$

$$\begin{aligned}
 &= \sigma^2 [\mu^{n-1} + \mu^n] + \mu^n [\sigma^2 \mu^{n-3} + \mu^2 V(X_{n-3})] \\
 &= \sigma^2 [\mu^{n-1} + \mu^n + \mu^{n+1}] + \mu^n V(X_{n-3}) \\
 &\quad \vdots \\
 &= \sigma^2 [\mu^{n-1} + \mu^n + \dots + \mu^{2n-2}] + \mu^{2n} V(X_0) \rightarrow 0 \\
 &= \sigma^2 \mu^{n-1} [1 + \mu + \dots + \mu^{n-1}]
 \end{aligned}$$

$$\begin{aligned}
 &= \begin{cases} \sigma^2 \mu^{n-1} \left( \frac{1-\mu^n}{1-\mu} \right) & \text{if } \mu \neq 1 \\ n \sigma^2 & \text{if } \mu = 1 \end{cases} \\
 u_{n+1} &= P(X_{n+1} = 0) = \sum_{j=0}^{\infty} P(X_{n+1} = 0 | X_1 = j) p_j \\
 &\quad \downarrow \\
 &\quad [P(X_n = 0)]^j \\
 \boxed{u_{n+1} = \sum_{j=0}^{\infty} u_n^j p_j} \\
 &\quad \longrightarrow
 \end{aligned}$$

$\Pi_0$  prob. of ultimate extinction, i.e., prob. that the population will eventually die out (under the assumption that  $X_0 = 1$ )

$$\Pi_0 = \lim_{n \rightarrow \infty} P(X_n = 0 | X_0 = 1)$$

$$\rightarrow \Pi_0 = 1 \quad \underline{\text{if } \mu < 1}$$

$$\begin{aligned}
 \mu^n &= E(X_n) = \sum_{j=1}^{\infty} j P(X_n = j) \\
 &\geq \sum_{j=1}^{\infty} 1 \cdot P(X_n = j)
 \end{aligned}$$

$$= P(X_n \geq 1)$$

$$\lim_{n \rightarrow \infty} P(X_n \geq 1) = 0 \Rightarrow \lim_{n \rightarrow \infty} P(X_n = 0) = 1$$

$$\Rightarrow \pi_0 = 1$$

$$\rightarrow \pi_0 = 1 \text{ if } \mu = 1$$

$\rightarrow$  When  $\mu > 1$

$$\pi_0 = P(\text{prob die out})$$

$$= \sum_{j=0}^{\infty} P(\text{prob die out} | X_j = j) p_j$$

$$\Rightarrow \boxed{\pi_0 = \sum_{j=0}^{\infty} \pi_0^j p_j} \rightarrow \pi_0 \text{ is the smallest positive number which is the solution of this equation}$$

Example:  $\boxed{1} \quad \pi_0 = \frac{1}{2}, p_1 = \frac{1}{4}, p_2 = \frac{1}{4}$

$$\pi_0 = 1 \quad \mu = 0 \times \frac{1}{2} + 1 \times \frac{1}{4} + 2 \times \frac{1}{4} = \frac{3}{4} < 1$$

$$\boxed{x_0 = n} \quad \pi_0^n = 1 = 1$$

$$\boxed{2} \quad \underline{x_0 = 1} \quad p_0 = \frac{1}{4}, p_1 = \frac{1}{4}, p_2 = \frac{1}{2} \quad \pi_0$$

$$\mu = \frac{1}{4} + \frac{1}{4} = \frac{5}{4} > 1$$

$$\pi_0 = \sum_j \pi_0^j p_j$$

$$\Rightarrow \pi_0 = \frac{1}{4} + \frac{1}{4} \pi_0 + \frac{1}{2} \pi_0^2$$

$$\Rightarrow \pi_0 = 1, \frac{1}{2} \Rightarrow \boxed{\pi_0 = \frac{1}{2}}$$

$$\boxed{x_0 = n} \quad -1 \dots n$$

$$\pi_0 = \left(\frac{1}{2}\right)$$

—x—

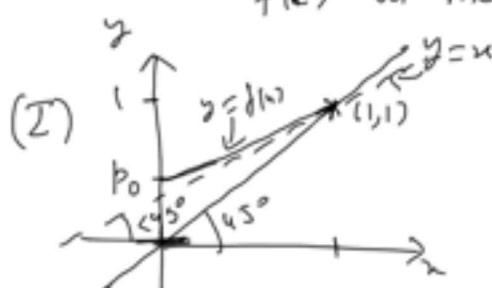
$$x = f(x) \rightarrow \textcircled{1}^*, \text{ where } f(x) = \sum_{j=0}^{\infty} x^j p_j$$

$$f(0) = p_0 > 0 \quad 0 \leq x \leq 1,$$

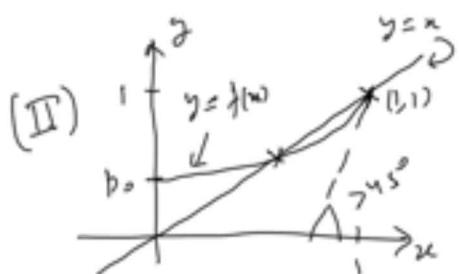
$$f(1) = \sum_{j=0}^{\infty} p_j = 1 \quad f(x) = p_0 + x p_1 + x^2 p_2 + x^3 p_3 + \dots$$

$$f'(x) > 0 \quad ; \quad f''(x) > 0 \quad f''(x) = 2p_2 + 6xp_3 + \dots$$

$f(x)$  is increasing and convex.  $f'(1) = \mu$



$$f(x) > x \quad \forall x \in (0, 1) \Leftrightarrow f'(1) \leq 1 \Leftrightarrow \mu \leq 1$$



$$f(x) = x \text{ for some } x \in (0, 1) \Leftrightarrow \mu > 1$$

Let  $p_0 > 0$ ,  $p_1 + p_2 < 1$ , let  $\pi$  satisfy  $\textcircled{1}^*$

We will show by induction  $\pi \geq P(X_n=0), \forall n$

$$\pi = \sum_{j=0}^{\infty} \pi^j p_j \geq p_0 = P(X_1=0) \quad \text{XXXX}$$

Assume  $\text{XXXX}$  holds for  $n$ , i.e.,  $\pi \geq P(X_n=0)$

$$P(X_{n+1}=0) = \sum_{j=0}^{\infty} (P(X_{n+1}=0 | X_1=j)) p_j$$

$$[P(X_n=0)]^j$$

$$= \sum_j (P(X_n=0))^j p_j$$

$$\leq \sum_j \pi^j p_j$$

$$= \pi$$

Using mathematical induction  $\Pi \geq P(X_n=0), \forall n$

$$\Pi \geq \lim_{n \rightarrow \infty} P(X_n=0) = \Pi_0$$

Branching process & generating function:

v.v.  $\xi_i \geq 0$ , integer valued s.t.  $P(\xi_i=k) = p_k, k=0,1,2,\dots$

$$\text{generating function } \phi(s) = E(s^{\xi_i}) = \sum_{k=0}^{\infty} s^k p_k$$

$$= p_0 + p_1 s + p_2 s^2 + \dots, \quad 0 \leq s \leq 1$$

$$\frac{d\phi(s)}{ds} \Big|_{s=0} = p_1 \quad s \frac{1}{2!} \frac{d^2\phi(s)}{ds^2} \Big|_{s=0} = p_2 \quad ;$$

$$\frac{1}{k!} \frac{d^k \phi(s)}{ds^k} \Big|_{s=0} = p_k$$

Indep. r.v.  $\xi_i$  having generating function  $\phi_i(s)$

$X = \sum_{i=1}^n \xi_i$  generating function

$$\phi_X(s) = E(s^{\sum_{i=1}^n \xi_i})$$

$$= E(s^{\xi_1}) \cdots E(s^{\xi_n})$$

$$= \phi_1(s) \phi_2(s) \cdots \phi_n(s)$$

$$\frac{d\phi(s)}{ds} \Big|_{s=1} = p_1 + 2p_2 + 3p_3 + \dots = E(\xi)$$

$$\frac{d^2\phi(s)}{ds^2} \Big|_{s=1} = E(\xi(\xi-1)) = E(\xi^2) - E(\xi)$$

$$V(\xi) = E(\xi^2) - (E(\xi))^2$$

$$= \frac{d^2\phi(s)}{ds^2} \Big|_{s=1} + \frac{d\phi(s)}{ds} \Big|_{s=1} - \left( \frac{d\phi(s)}{ds} \Big|_{s=1} \right)^2$$

Example :  $\xi_i \sim \text{Pois}(l)$

$$l = m = \gamma = \lambda \cdot k$$

$$P_k = P(\xi = k) = \frac{e^{-\lambda} \lambda^k}{k!}, \quad k = 0, 1, 2, \dots$$

$$\begin{aligned}\phi(s) &= E(s^\xi) = \sum_{k=0}^{\infty} s^k \frac{e^{-\lambda} \lambda^k}{k!} \\ &= e^{-\lambda} \sum_{k=0}^{\infty} \frac{(\lambda s)^k}{k!} \\ &= e^{-\lambda} e^{\lambda s} \\ &= e^{-\lambda(1-s)} \quad \text{for } |s| < 1\end{aligned}$$

$$E(\xi) = \lambda$$

$$V(\xi) = \lambda$$

— X —

Branching process  $X_n \leftarrow$  at stage  $n$ , population.

$$\text{Let offspring dist } p_k = P(\xi = k); \quad \phi(s) = E(s^\xi) = \sum_{k=0}^{\infty} s^k p_k$$

$$u_n = P(X_n = 0)$$

$$= \sum_{k=0}^{\infty} u_{n-1}^k p_k$$

$$= \phi(u_{n-1})$$

$$u_0 = 0, \quad u_1 = \phi(u_0), \quad u_2 = \phi(u_1), \dots$$

$$\pi_0 \quad \text{smallest soln of } u = \phi(u)$$

$$\pi_\infty = 1 - \pi_0$$

Ex parent has no offspring w/p  $\frac{1}{4}$

$$2 \quad , \quad \text{w/p } \frac{3}{4}$$

$$\begin{aligned}\phi(s) &= E(s^\xi) = \sum_k s^k p_k \\ &= p_0 + s^2 p_2 = \frac{1}{4} + \frac{3}{4} s^2\end{aligned}$$

$$u_n = \phi(u_{n-1}) = \frac{1}{4} + \frac{3}{4} u_{n-1}^2$$

$$u_0 = 0$$

$$u_1 = \frac{1}{4}, \quad u_2 = \frac{1}{4} + \frac{3}{4} (u_1)^2 = \frac{1}{4} + \frac{3}{4} \times \frac{1}{16}$$

