

Transformations and priors that we use

January 25, 2014

In this document, we lay out three aspects of model implementation:

1. Order that parameters are passed in to the estimation algorithm;
2. Transformations that we use to get from unconstrained to constrained parameters; and
3. Prior distributions of the *unconstrained* parameters, including any Jacobians from the transformations.

Parameters are passed in the following order:

1. θ_{12} , columnwise, with no transformation.
2. \bar{c} , the average c_j , across J brands (no transformation);
3. $\log \text{sd } c$, the log of the standard deviation of c_j ;
4. c_{off} , the standardized offset of c_j from the mean. Thus, $c_j = \text{sd}(c) (c_{\text{off}} - \bar{c})$.
5. \bar{u} , the average u_j , across J brands (no transformation);
6. $\log \text{sd } u$, the log of the standard deviation of u_j ;
7. u_{off} , the standardized offset of u_j from the mean. Thus, $u_j = \text{sd}(u) (u_{\text{off}} - \bar{u})$.
8. ϕ , which is included only if the H matrix is included. Columnwise, no transformation;
9. $\text{logit } \delta$, a scalar parameter;
10. \log diagonal elements for V_1 (see below for details);
11. factors for V_1 , columnwise, some elements transformed (see details);
12. \log diagonal elements for V_2 (see below for details);
13. factors for V_2 , columnwise, some elements transformed (see details);
14. \log scale parameter for W_1 (see below);
15. transformed Cholesky factors for W_1 (see below);
16. \log diagonal elements for W_1 (see below for details);
17. factors for W_1 , columnwise, some elements transformed (see details);

0.1 θ_{12}

The parameters for θ_{12} are passed in columnwise, with no transformations.

We apply a matrix normal prior, passing in a mean matrix, and the lower Cholesky decompositions for the two covariance matrices. The matrix normal is a standard parameterization.

0.2 c and u

The average effects \bar{c} and \bar{u} are unconstrained, with normal priors. The priors on $\text{sd } c$ and $\text{sd } u$ are half-T. Hyperparameters are σ_c and ν_c (similar for u). The Jacobians of the transformations are $\text{sd } c$ and $\text{sd } u$. Therefore,

$$\pi(\log \text{sd}(c)) = \frac{2\Gamma(\frac{\nu_c+1}{2})}{\Gamma(\frac{\nu_c}{2})\sigma_c\sqrt{\nu_c\pi}} \left[1 + \frac{1}{\nu_c} \left(\frac{\text{sd}(c)}{\sigma_c} \right)^2 \right]^{-\frac{\nu_c+1}{2}} \text{sd}(c) \quad (1)$$

$$\pi(\log \text{sd}(u)) = \frac{2\Gamma(\frac{\nu_u+1}{2})}{\Gamma(\frac{\nu_u}{2})\sigma_u\sqrt{\nu_u\pi}} \left[1 + \frac{1}{\nu_u} \left(\frac{\text{sd}(u)}{\sigma_u} \right)^2 \right]^{-\frac{\nu_u+1}{2}} \text{sd}(u) \quad (2)$$

The prior for each c_j is normal, with mean \bar{c} and standard deviation $\text{sd}(c)$. To operationalize this, we immediately transform using the offsets

$$c_j = \text{sd}_c (c_{\text{off}} + \bar{c}) \quad (3)$$

$$u_j = \text{sd}_u (u_{\text{off}} + \bar{u}) \quad (4)$$

$$(5)$$

This transformation lets us give each c_{off} and u_{off} a standard normal prior (mean=0, sd=1).

0.3 ϕ

The coefficient matrix ϕ is passed in columnwise, with no transformation. It is included only if we are including the H matrix in the model.

We apply a matrix normal prior, passing in a mean matrix, and the lower Cholesky decompositions for the two covariance matrices. The matrix normal is a standard parameterization.

0.4 δ

We pass in $\text{logit } \delta$, and transform so

$$\delta = \frac{\exp(\text{logit } \delta)}{1 + \exp(\text{logit } \delta)} \quad (6)$$

The prior on δ is a beta distribution with parameters a_δ and b_δ . The Jacobian of the transformation is $d\delta = \delta(1 - \delta)$. Therefore,

$$\pi(\text{logit } \delta) = \frac{\Gamma(a_\delta + b_\delta)}{\Gamma(a_\delta)\Gamma(b_\delta)} \delta^{a_\delta} (1 - \delta)^{b_\delta} \quad (7)$$

Note that when taking the log of this prior, $\log \delta = \text{logit } \delta - \log(1 + \exp(\text{logit } \delta))$, and $\log(1 - \delta) = -\log(1 + \exp(\text{logit } \delta))$. This can be useful for avoiding numerical issues that would result from computing unnecessary logarithms.

0.5 V_1 and V_2

The covariance matrices V_1 and V_2 are structured similarly, so we will consider a general matrix V . We will let V take a factor-analytic structure, where x is a matrix of factors and Σ is a diagonal matrix with all positive elements. We construct the matrix as

$$V = xx' + \Sigma \quad (8)$$

Let's start with Σ , and let the i^{th} element of the diagonal be Σ_{ii} . The prior on each Σ_{ii} is half-T. Since we are passing in $\log \Sigma_{ii}$ instead, we need to multiply each half-T density by Σ_{ii} (the Jacobian of the transformation).

We arrange x to have each column be a factor. For identification (need a reference for this), the upper triangle of x is zero, and the elements of the diagonal are all positive. Therefore, if x has k rows and n columns, there are only $kn - \frac{1}{2}n(n + 1)$ unconstrained parameters and $\frac{1}{2}n(n - 1)$ positive parameters. The unconstrained parameters all have T priors, and the positive ones have half-T priors. Only the half-T densities need to be multiplied by the Jacobian of the transformation.

When passing in the parameters, $\log \text{diag } \Sigma$ comes first, followed by the elements of x , column-wise.

If there are no factors, then $V = \Sigma$.

0.6 W

Now, it gets fun. We partition W so the upper left corner is αW_1 and the lower right corner is W_2 . The upper right and lower left corners are all zero. W_2 has the same factor-analytic structure as V_1 and V_2 .

For identification, all elements of the diagonal of W_1 must be the same. Therefore, we treat αW_1 as a scaled correlation matrix, where $\alpha > 0$ is the scale parameter. W_1 is a symmetric, positive definite matrix with all ones on the diagonal. The prior on $\log \alpha$ is a half-T prior, multiplied by α (the Jacobian).

We place an LKJ prior on W_1 , with parameter η (see LKJ paper and Stan manual). W_1 has $J + 1$ rows/columns.

$$\pi(W_1) = C|W_1|^{\eta-1} \quad (9)$$

where

$$C = 2^{\sum_{i=1}^J (2\eta - 2 + J + 1 - i)(J + 1 - i)} \prod_{i=1}^J \left[\mathbb{B} \left(\eta + \frac{J-i}{2}, \eta + \frac{J-i}{2} \right) \right]^{J+1-i} \quad (10)$$

Let y be the vector of the $d = \binom{J+1}{2}$ unconstrained parameters, arranged columnwise in a lower triangular matrix. Let $z_{ij} = \tanh(y_{ij})$ (this is a Fisher transformation). Following the procedure in the Stan manual, construct another lower triangular matrix x as follows:

1. $x_{11} = 1$
2. For $i = 2 \dots J+1$, $x_{i1} = z_{i1}$ (copy first column)
3. For $i = 2 \dots J+1$, $x_{ii} = \prod_{k=1}^{i-1} \sqrt{1 - z_{ik}^2}$ (diagonal elements)
4. For $j = 2 \dots J+1$ and $i = j+1 \dots J+1$, $x_{ij} = z_{ij} \prod_{k=1}^{j-1} \sqrt{1 - z_{ik}^2}$ (remaining off-diagonal elements)

This transformation will ensure that $W_1 = xx'$ is a valid correlation matrix.

Through the magic of Mathematica, we get the following Jacobian of the transformation

$$\mathcal{J}_{W_1} = \prod_{j=1}^{d-1} \prod_{i=j+1}^d [\text{sech}(y_{ij})]^{d-j+1} \quad (11)$$

$$= \prod_{j=1}^{d-1} \prod_{i=j+1}^d \left[1 - z_{ij}^2 \right]^{\frac{d-j+1}{2}} \quad (12)$$

The operator sech is the hyperbolic secant. The second line comes from the identity $\text{sech}^2(x) = 1 - \tanh^2(x)$.

Mathematica also tells us that

$$|W_1| = \prod_{j=1}^{d-1} \prod_{i=j+1}^d \left[1 - z_{ij}^2 \right] \quad (13)$$

If $\eta = 1$, the distribution is uniform over all correlation matrices. For $\eta > 1$, there is a mode at the identity matrix, and for $\eta < 1$ there is a trough.