

M342W

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$$g(x) = \hat{y} = \underbrace{\bar{y}_{\text{red}}}_{b_0} + \underbrace{(\bar{y}_{\text{green}} - \bar{y}_{\text{red}})}_{b_1} x, \text{ let } n_g = \sum x_i, p_g = \bar{x} = \frac{n_g}{n}$$

$$n_r = n - n_g$$

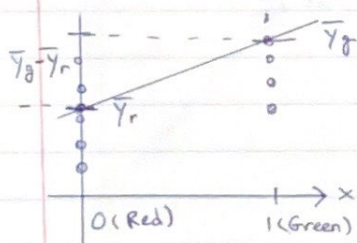
$$\bar{y} = \frac{1}{n} (\sum y_i) = \frac{1}{n} (\sum_{i: \text{green}} y_i + \sum_{i: \text{red}} y_i) = \frac{\sum y_i}{n} \cdot \frac{n_g}{n_g} + \frac{\sum y_i}{n} \cdot \frac{n_r}{n_r}$$

$$= p_g \frac{\sum y_i}{n_g} + (1 - p_g) \frac{\sum y_i}{n_r} = p_g \bar{y}_g + (1 - p_g) \bar{y}_r$$

$$b_1 = \frac{\sum x_i y_i - \frac{(\sum x_i)(\sum y_i)}{n}}{\sum x_i^2 - \frac{(\sum x_i)^2}{n}} = \frac{n_g \bar{y}_g - n p_g \bar{y}}{n_g - n p_g^2} \cdot \frac{1}{n} = \frac{p_g \bar{y}_g - p_g \bar{y}}{p_g - p_g^2} = \frac{\bar{y}_g - \bar{y}}{1 - p_g}$$

$$\bar{y} = \frac{\bar{y}_g - (p_g \bar{y}_g + (1 - p_g) \bar{y}_r)}{1 - p_g} = \frac{(1 - p_g) \bar{y}_g - (1 - p_g) \bar{y}_r}{1 - p_g} = \bar{y}_g - \bar{y}_r$$

$$b_0 = \bar{y} - b_1 \bar{x} = p_g \bar{y}_g + (1 - p_g) \bar{y}_r - (\bar{y}_g - \bar{y}_r) p_g = \bar{y}_r$$



What if $x \in \{\text{red, green, blue}\}$? This is then $p=2$ and we need an OLS solution for $p>1$. But intuitively...

$$g(x) = \begin{cases} \bar{y}_r & \text{if } x = \text{red} \\ \bar{y}_g & \text{if } x = \text{green} \\ \bar{y}_b & \text{if } x = \text{blue} \end{cases} = \underbrace{\bar{y}_r}_{b_0} + \underbrace{(\bar{y}_g - \bar{y}_r)}_{b_1} \underbrace{1_{x=\text{green}}}_{x_1} + \underbrace{(\bar{y}_b - \bar{y}_r)}_{b_2} \underbrace{1_{x=\text{blue}}}_{x_2}$$

How well does g predict? We need a "model performance metric". In the SVM this was accuracy or misclassification error. Here, it ~~will~~ can ~~be~~ also be what we use internally in the algorithm.

$$SSE := \sum_{i=1}^n e_i^2 = \sum (y_i - g(x_i))^2. \text{ Is SSE interpretable?}$$

Is SSE interpretable? No, let's take the mean at least, call that mean squared error (MSE):

* $MSE = \frac{1}{n-2} SSE$, but this is still in the squared unit of the phenomenon so it's still uninterpretable. We can take the square root of MSE called root mean squared error (RMSE):

$$RMSE = \sqrt{\frac{1}{n-2} \sum e_i^2} = \sqrt{MSE}$$

RMSE is in the same unit as y (just like the standard deviation is in the same unit as the random variable). Also, from the CLT,

$$[g(x) \pm 1.96 \cdot RMSE]$$

is approx a 95% confidence interval for the true y at that x

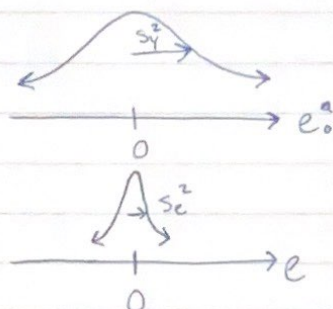
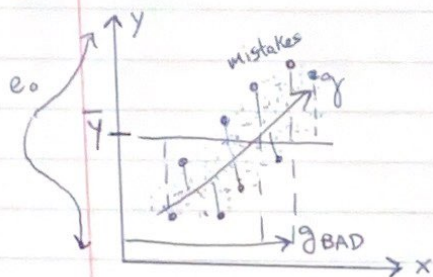
"standard deviation of the residuals" s_e

RMSE is a very important metric in regression models.

Another important error/performance metric is "R-squared" which is the "proportion of variance explained". We will now explain this definition.

Consider the null model $g_0 = \bar{y}$. What is the SSE of this model? Let's call it SSE_0 .

$$SSE_0 = \sum_{i=1}^n e_{0,i}^2 = \sum_{i=1}^n (y_i - \bar{y})^2 = \underbrace{SST}_{\text{Sum of Squares total}} = (n-1)s_y^2$$

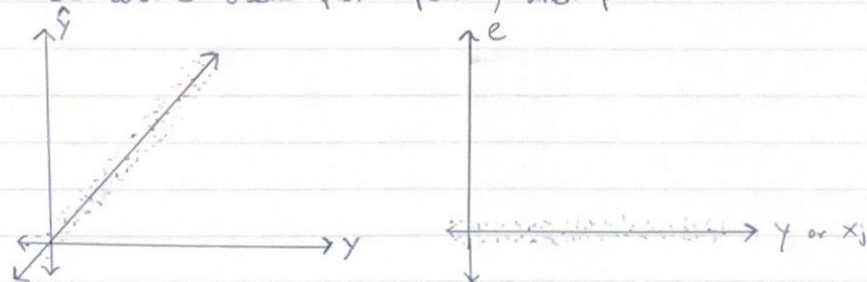


$$\frac{SSE}{SST} = \frac{(n-1)S_e^2}{(n-1)S_y^2} = \frac{S_e^2}{S_y^2}$$

$$\frac{SST - SSE}{SST} = \frac{(n-1)S_y^2 - (n-1)S_e^2}{(n-1)S_y^2} = \frac{\overbrace{S_y^2 - S_e^2}^{\Delta S^2}}{S_y^2} = \frac{\Delta S^2}{S_y^2} = R^2$$

R^2 can never be more than 100%. But R^2 can be negative. This occurs when $S_e^2 > S_y^2$ meaning the model is predicting worse than $y_0 = \bar{y}$.

Here's another useful plot especially when $p > 1$:



$R^2 = 1$ implies $RMSE = 0$

$R^2 \uparrow$ implies $RMSE \downarrow$

If $R^2 = 99\%$, does this mean the model is for sure "good"? No.

Because if the initial variance was so very large, even a 99% reduction wouldn't result in a small residual variance i.e. $RMSE$ still could be high after 99% variance reduction.

We now would like to generalize the least squares estimation algorithm to cases where $p > 1$. Let's begin with $p = 2$.

$$\mathcal{H} = \{ \omega_0 + \omega_1 x_1 + \omega_2 x_2 : \underbrace{\omega_0, \omega_1, \omega_2}_{\vec{\omega} \in \mathbb{R}^3} \in \mathbb{R} \}$$

$$SSE = \sum_{i=1}^n e_i^2 = \sum (y_i - \hat{y}_i)^2 = \sum (y_i - \omega_0 - \omega_1 x_{1,i} - \omega_2 x_{2,i})^2$$

$$b_0 = \underset{\omega_0 \in \mathbb{R}}{\operatorname{argmin}} \{SSE\}, \quad b_1 = \underset{\omega_1 \in \mathbb{R}}{\operatorname{argmin}} \{SSE\}, \quad b_2 = \underset{\omega_2 \in \mathbb{R}}{\operatorname{argmin}} \{SSE\}$$

This problem can be solved more simply with matrix algebra and a matrix equation:

$$\mathcal{D} = \langle X, \vec{y} \rangle, \quad \text{let } X = [\vec{1}_n \ \vec{X}_1 \ \vec{X}_2] = \begin{bmatrix} 1 & x_{11} & x_{12} \\ 1 & x_{21} & x_{22} \\ \vdots & \vdots & \vdots \\ 1 & x_{n1} & x_{n2} \end{bmatrix}$$

$$\text{e.g. } \hat{y}_i = \vec{x}_i \vec{\omega}$$

$$\vec{\hat{y}} = X \vec{\omega} = \begin{bmatrix} \omega_0 + \omega_1 x_{11} + \omega_2 x_{12} \\ \omega_0 + \omega_1 x_{21} + \omega_2 x_{22} \\ \vdots \\ \omega_0 + \omega_1 x_{n1} + \omega_2 x_{n2} \end{bmatrix}$$

$$\text{define } \vec{e} := \vec{y} - \vec{\hat{y}}$$

$$SSE = \sum_{i=1}^n e_i^2 = \underbrace{\vec{e}^T \vec{e}}_{\vec{\hat{y}}^T \vec{y}} = (\vec{y} - \vec{\hat{y}})^T (\vec{y} - \vec{\hat{y}}) = (\vec{y}^T - \vec{\hat{y}}^T) (\vec{y} - \vec{\hat{y}})$$

$$= \vec{y}^T \vec{y} - \vec{\hat{y}}^T \vec{y} - \underbrace{\vec{y}^T \vec{\hat{y}}}_{\vec{\hat{y}}^T \vec{y}} + \vec{\hat{y}}^T \vec{\hat{y}} = \vec{y}^T \vec{y} - 2 \vec{\hat{y}}^T \vec{y} + \vec{\hat{y}}^T \vec{\hat{y}}$$

$$= \vec{y}^T \vec{y} - 2 (X \vec{\omega})^T \vec{y} + (X \vec{\omega})^T X \vec{\omega} = \vec{y}^T \vec{y} - 2 \vec{\omega}^T X^T \vec{y} + \vec{\omega}^T X^T X \vec{\omega}$$