

Exact Importance Sampling for Affine Processes

Brennan Hall

Supervised by: Alex Shkolnik

bhall@ucsb.edu

June 4, 2019

Department of Statistics & Applied Probability

University of California, Santa Barbara

Setting

For a Borel set $\mathbb{D} \subseteq \mathbb{R}^d$, a homogeneous \mathbb{D} -valued Markov process Y on a probability space $(\Omega, \mathcal{F}, \mathbf{P})$ is called *affine* if, letting $\mathcal{U} = \{u \in \mathbb{C}^d \mid \sup_{x \in \mathbb{D}} \Re(u \cdot x) < \infty\}$, for every pair $(t, u) \in [0, \infty) \times \mathcal{U}$ there are $\phi(t, u) \in \mathbb{C}$ and $\varphi(t, u) \in \mathbb{C}^d$ s.t.

$$(1) \quad \mathbf{E}_y(e^{u \cdot Y_t}) = e^{\phi(t, u) + \varphi(t, u) \cdot y}$$

where $y \in \mathbb{D}$ is the starting point of Y .

e.g. Brownian motion, Lévy, OU, Bessel, CIR.

Affine structure

An affine jump process Y on state space $\mathbb{D} \subset \mathbb{R}^d$ can be expressed as the solution to an SDE of the form

$$(2) \quad dX_t = a(X_t)dt + b(X_t)dW_t + dJ_t$$

where W is a standard Brownian Motion, J is a marked point process with jump size measure K where a, b and K are affine functions of the form

$$(3) \quad \begin{aligned} a(x) &= a + x_1\alpha^1 + \cdots + x_d\alpha^d \\ b(x) &= b + x_1\beta^1 + \cdots + x_d\beta^d \\ K(x, d\xi) &= m(d\xi) + x_1\mu^1(d\xi) + \cdots + x_d\mu^d(d\xi) \end{aligned}$$

and $a(x)$ is p.s.d. in $\mathbb{R}^{d \times d}$, $b(x) \in \mathbb{R}^d$ and K is a (nondegenerate) Radon measure on \mathbb{R}^d s.t.. $\int_{\mathbb{R}^d} (|\xi|^2 \wedge 1) K(x, d\xi) < \infty$.

Simple example

Consider the strong solution $X \in \mathbb{R}_+$ to the SDE

$$(4) \quad dX_t = \kappa(\mu - X_t)dt + \sigma\sqrt{X_t}dW_t + \delta dN_t$$

for $\kappa, \mu, \sigma, \delta, x \in \mathbb{R}_+$ and W , a standard Brownian motion. Arrivals of counting process N occur with an intensity X .

For $n \in \mathbb{N}$ and $T \in [0, \infty)$, compute the tail probability

$$(5) \quad \mathbf{P}(\mathcal{E}_n) = \mathbf{P}(N_T \geq n).$$

Monte Carlo: sample $\mathbf{1}_{\mathcal{E}_n}$ and average.

Let $\{\mathcal{E}_n\}_{n \in \mathbb{N}}$ denote a rare-event sequence:

n – rare-event parameter (*problem specific*),

$\mathbf{P}(\mathcal{E}_n) \rightarrow 0$ as $n \uparrow \infty$.

The Monte Carlo (MC) estimator $\mathbf{1}_{\mathcal{E}_n}$ is highly inefficient.

$$(6) \quad \text{Relative Error: } \frac{\sqrt{\mathbf{Var}(\mathbf{1}_{\mathcal{E}_n})}}{\mathbf{P}(\mathcal{E}_n)} \sim 1/\sqrt{\mathbf{P}(\mathcal{E}_n)}$$

Classic Approach

Entails a change of measure from \mathbf{P} to \mathbf{Q} .

Let $Z \geq 0$ on (Ω, \mathcal{F}) satisfy $\mathbf{E}(Z) = 1$. Then,

$$(7) \quad \mathbf{Q}(\mathcal{A}) = \mathbf{E}(\mathbf{1}_{\mathcal{A}} Z) \quad \mathcal{A} \in \mathcal{F}.$$

Given an equivalent of measure, i.e., $Z > 0$,

Sample the estimator $\mathbf{1}_{\mathcal{E}_n}/Z$ under \mathbf{Q} , i.e.,

$$(8) \quad \mathbf{E}_{\mathbf{Q}}(\mathbf{1}_{\mathcal{E}_n}/Z) = \mathbf{P}(\mathcal{E}_n).$$

Exponential change of measure

For $\theta \in \mathbb{R}$, take (Siegmund, 1976)

$$(9) \quad m(\theta) = \mathbf{E}(e^{\theta N_T})$$

and consider $\mathbf{Q} = Z\mathbf{P}$ for $Z = e^{\theta N_T} / m(\theta)$.

The probability \mathbf{Q} is the (unique) asymptotically optimal importance sampling measure for estimator $\mathbf{1}_{\mathcal{E}_n} / Z$ of $\mathbf{P}(\mathcal{E}_n)$,
i.e., per fixed level of precision, the number of importance sampling trials of $\mathbf{1}_{\mathcal{E}_n} / Z$ is subexponential in n (efficient).

$$\liminf_{n \uparrow \infty} \frac{\log \mathbf{Var}_{\mathbf{Q}}(\mathbf{1}_{\mathcal{E}_n} / Z)}{\log \mathbf{P}(\mathcal{E}_n)} = 2.$$

The process $Y = (N, X)$ is affine ($\mathbf{E}(e^{u \cdot Y_t}) = e^{\phi(t,u) + \varphi(t,u) \cdot y}$).

We can compute $m(\theta)$. More generally,

$$(10) \quad \mathbf{E}(e^{\theta(N_T - N_t)} \mid \mathcal{F}_t) = e^{p(t) + q(t)X_t}.$$

where for $p(T) = q(T) = 0$ the (p, q) solve the ODEs

$$(11) \quad \begin{aligned} \dot{p} &= -\kappa \mu q \\ \dot{q} &= \kappa q - \frac{1}{2} \sigma^2 q^2 - (e^{\theta + \delta q} - 1) \end{aligned}$$

How do we sample (N, X) under \mathbf{Q} ?

ECM dynamics (time-dependence)

The dynamics of $Y = (N, X)$ under \mathbf{Q} are

- N admits an intensity $e^{\theta + \delta q} X$.
- X starts at x and, for B a standard Brownian motion, solves

$$(12) \quad dX_t = \kappa_q(t) (\mu_q(t) - X_t) dt + \sigma \sqrt{X_t} dB_t + \delta dN_t$$

where for the ODE solutions (p, q) we have

$$(13) \quad \kappa_q = \kappa - \sigma^2 q, \quad \mu_q = \frac{\kappa \mu}{\kappa_q}.$$

The drift is now time dependent! Introduces sampling bias that can materially effect small probability estimates.

Two sources of bias in sampling X :

1. Standard method to sample the jump times through a time change of a standard Poisson process. Compensator of N

$$(14) \quad A_T = \int_0^t X_s ds$$

is computed by approximation.

2. Time-dependent drift coefficient makes transition law of X unknown under \mathbf{Q} .

New Approach

Approximate ECM allows for exact sampling of the process $Y = (N, X)$ under \mathbf{Q} .

- Define a time grid $0 = t_0 < t_1 < \dots < t_{K_n} = T$
- Approximate X with step functions X^n along time grid
- Approximate time-dependent variables with step functions $p_n(t)$ and $q_n(t)$ along time grid
- Simulate exactly over each time grid.
- Bias can be corrected

Let $q_n(t) = q(t_k)$, $\kappa_k = \kappa_{q_n}(t)$, $\mu_k = \mu_{q_n}(t)$ and $X_t^n = X_{t_k}$ on each $[t_k, t_{k+1})$.

The dynamics of $Y = (N, X)$ under \mathbf{Q} are given by

- N admits an intensity $e^{\theta + \delta q_n} X^n$.
- X starts at x and, for B a standard Brownian motion, solves

$$dX_t = \kappa_k(\mu_k - X_t)dt + \sigma\sqrt{X_t}dB_t + \delta dN_t \quad t_k \leq t < t_{k+1}$$

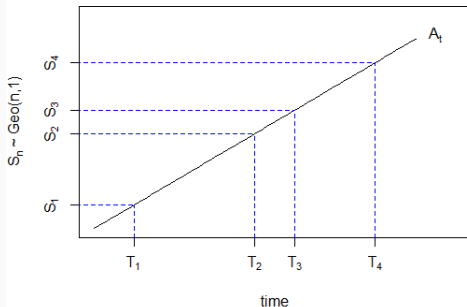
The (N, X) may be sampled exactly under \mathbf{Q} on each $[t_{k-1}, t_k)$.

Sampling jump times

Over each grid interval $[t_{k-1}, t_k)$ N has \mathbf{Q} -compensator

$$(15) \quad A_T = \int_0^T X_s^n ds = A_{t_{k-1}} + \int_{t_{k-1}}^T X_s^n ds$$

$$T_j = \inf \{t > T_{j-1} : A_t \geq S_j\}.$$



Bias correction term included in importance sampling measure

$$\mathbf{Q} = Z_n \mathbf{P}.$$

Z_n may be expressed as $Z_n = Q_n Z$ for $Z = e^{\theta N_T} / m(\theta)$ and

$$Q_n = D_n^{ecm} \times D_n^{rate}$$

$$D_n^{ecm} = \exp \left\{ \sum_{k=1}^{K_n} \dot{q}_n(t_{k-1}) \int_{t_{k-1}}^{t_k} X_s ds - X_{t_{k-1}} (q_n(t_k) - q_n(t_{k-1})) \right\}$$

$$D_n^{rate} = \exp \left\{ \int_0^T e^{\theta + q_n(s)} (X_s - X_s^n) ds \right\} \prod_{j=1}^{N_T} \left(\frac{X_{T_j^-}}{X_{T_j^-}^n} \right)$$

$Q_n \Rightarrow 1$, ie. approximate ECM converges to ECM.

What we have so far

- ECM is optimal importance sampling measure - for $Z = e^{\theta N_T} / m(\theta)$

$$(16) \qquad \mathbf{Q} = Z\mathbf{P}$$

- Two sources of bias in sampling
 - Approximation required to sample jump times via compensator time change
 - ECM introduces sampling bias for affine jump processes due to time dependent drift
- Approximate ECM allows for exact sampling and preserves (asymptotic) optimality.

Theorem. The importance measure $\mathbf{Q} = Z_n \mathbf{P}$ is asymptotically optimal for estimating $\mathbf{P}(\mathcal{E}_n)$ and $\mathbf{1}_{\mathcal{E}_n}/Z_n$ may be sampled exactly.

Conditional Monte Carlo method uses

$$(17) \quad \mathbf{E}_{\mathbf{Q}}(\mathbf{1}_{\mathcal{E}_n} L_T) = \mathbf{E}_{\mathbf{Q}}(\mathbf{1}_{\mathcal{E}_n} \mathbf{E}_{\mathbf{Q}}(L_T | \mathcal{H}))$$

$$(18) \quad = \mathbf{E}_{\mathbf{Q}}(\mathbf{1}_{\mathcal{E}_n} L_T^{\mathcal{H}})$$

with $\mathcal{H} = \sigma(\{N_{\tau_l}, X_{\tau_l}\})$ and $L_T^{\mathcal{H}} = e^{-\theta N_T} m(\theta) G_T \mathbf{E}_{\mathbf{Q}}(H_T | \mathcal{H})$

Define $\{\tau_l\}$ as the sequence of sorted times $\{t_k\} \cup \{T_j\}$.

Sample estimator $L_T^{\mathcal{H}} \mathbf{1}_{\mathcal{G}_n}$:

$$G_T = \exp \left\{ \sum_{k=1}^{K_n} X_{t_{k-1}} \left(e^{\theta + q_n(t_{k-1})} - (q_n(t_k) - q_n(t_{k-1})) \right) \right\} \prod_{j=1}^{N_T} \left(\frac{X_{T_j^-}}{X_{T_j^-}^n} \right)$$

$$H_T = \prod_{l=1}^{K_n + N_T} \mathcal{L}_{\tau_{l-1}, X_{\tau_{l-1}}} \left(\tau_l, X_{\tau_l} \mid \dot{q}_n(\tau_{l-1}) + e^{\theta + q_n(\tau_{l-1})} \right)$$

where $\mathcal{L}_{(s,x)}(t, y \mid z) = \mathbf{E} \left(\exp \left\{ z \int_s^t \tilde{X}_z dz \right\} \mid \tilde{X}_s, \tilde{X}_t = y \right)$

is the bridge transform.

Affine models have numerous applications.

The classical exponential change of measure approach leads to biased estimators (leads to a time-dependent drift coefficient).

We develop an approximate ECM scheme which inherits asymptotic optimality properties but allows for exact sampling.

Numerical results currently in progress.