HARVARD EXTENSION SCHOOL

EXT CSCI E-106 Model Data Class Group Project Template

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In this project, our aim is to classify the probability of a passenger surviving the Titanic crash of 1912. We used a variety of linear and non-linear models to deduce the most accurate model and provide long-term stability in our predictions.

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Classify whether a passenger on board the maiden voyage of the RMS Titanic in 1912 survived given their age, sex and class. Sample-Data-Titanic-Survival.csv to be used in the Final Project

Variable	Description
pclass	Passenger Class $(1 = 1st; 2 = 2nd; 3 = 3rd)$
survived	Survival $(0 = No; 1 = Yes)$
name	Name
sex	Sex
age	f Age
sibsp	# of siblings / spouses aboard the Titanic
parch	# of parents / children aboard the Titanic
ticket	Ticket number
fare	Passenger fare
cabin	Cabin number
embarked	Port of Embarkation ($C = Cherbourg; Q =$
	Queenstown; $S = Southampton$)
boat	Lifeboat ID, if passenger survived
body	Body number (if passenger did not survive and
	body was recovered
home.dest	The intended home destination of the passenger

1 Instructions:

0. Join a team with your fellow students with appropriate size (Up to Nine Students total) If you have not group by the end of the week of April 11 you may present the project by yourself or I will randomly assign other stranded student to your group. I will let know the final groups in April 11.

- 1. Load and Review the dataset named "Titanic_Survival_Data.csv" 2. Create the train data set which contains 70% of the data and use set.seed (15). The remaining 30% will be your test data set.
- 3. Investigate the data and combine the level of categorical variables if needed and drop variables as needed. For example, you can drop id, Latitude, Longitude, etc.
- 4. Build appropriate model to predict the probability of survival.
- 5. Create scatter plots and a correlation matrix for the train data set. Interpret the possible relationship between the response.
- 6. Build the best models by using the appropriate selection method. Compare the performance of the best logistic linear models.
- 7. Make sure that model assumption(s) are checked for the final model. Apply remedy measures (transformation, etc.) that helps satisfy the assumptions.
- 8. Investigate unequal variances and multicollinearity.
- 9. Build an alternative to your model based on one of the following approaches as applicable to predict the probability of survival: logistic regression, classification Tree, NN, or SVM. Check the applicable model assumptions. Explore using a negative binomial regression and a Poisson regression.
- 10. Use the test data set to assess the model performances from above.
- 11. Based on the performances on both train and test data sets, determine your primary (champion) model and the other model which would be your benchmark model.
- 12. Create a model development document that describes the model following this template, input the name of the authors, Harvard IDs, the name of the Group, all of your code and calculations, etc..

Due Date: May 12 2025 1159 pm hours EST Notes No typographical errors, grammar mistakes, or misspelled words, use English language All tables need to be numbered and describe their content in the body of the document All figures/graphs need to be numbered and describe their content All results must be accurate and clearly explained for a casual reviewer to fully understand their purpose and impact Submit both the RMD markdown file and PDF with the sections with appropriate explanations. A more formal.

2 Executive Summary

We have tested five different models of which four were logistics regression based and one was a poisson model. The best model with regards to predictive accuracy was a "filtered" logistic regression (filtered because we removed insignificant predictors) with an accuracy of 84.18%.

The final model shows that the most significant predictors of survival are: age, sex, and deck_F. Below we show the final calculated function:

survived ~ 5.74 - 0.05 age - 0.025 sibsp + 0.80 pclass_1 + 1.82 pclass_2 - 3.27 sex_M + 0.83 embarked

3 Introduction

The Titanic was a British-registered ship that set sail on its maiden voyage on April 10th, 1912 with 2,240 passengers and crew on board. On April 15th, 1912, the ship struck an iceberg, split in half, and sank to the bottom of the ocean (National Oceanic and Atmospheric Administration (NOAA), 2023). In this report, we are going to analyze the data in the Titanic.csv file and use it to determine the best model for predicting whether someone on board would live or die. By creating this model, we hope to understand what factors a passenger could have taken into account in order to reduce their risk of death during the trip. We cleaned the data and split into into a train/test split in order to properly train our models. We created simple linear models, multivariate linear models, logistic models (both binomial and poisson), a regression tree, and a neural network model. The train sample size was 916 data points (70.03%) and the test sample size was 392 data points (29.97%). We built the models after examining the data and determining which predictor variables we thought would be most relevant for survival rate. Once we had our variables and training data, we created the models and examined the performance of the models on both training and testing data to determine if they were robust. We also examined if the model assumptions appeared to hold for each model.

4 Description of the data and quality

Based on the data cleaning we were able to only remove 2 rows from the data set. We used median imputation as well as KNN for various columns. We also dummified several categorical columns. We found that leaving sibsp and parch as continuous as opposed to categorical increased their contributions to the model performance (See Appendix A). Further, we also extracted the deck number and found that removing deck_G from the model increased its performance.

4.1 Loading the data

```
odata <- read.csv("../data/Titanic_Survival_Data.csv")
cat("Size of entire data set:", nrow(odata), "\n")</pre>
```

Size of entire data set: 1310

4.2 Removing un-needed columns

Name: Removing because names have no inference on surivival (inference)

ticket: Ticket No. will also likely not have an influence in survival

boat: This is highly correlated to the survival dependant variable since people who made it on a boat likely survived

body: This is highly correlated to the survival dependant variable since people who's body was recovered did not survive.

home.dest: The destination likely has nothing to do with the survival

```
data.clean = odata[, !(names(odata) %in% c("name", "ticket", "boat", "body", "home.dest"))]
```

4.3 Data Augmentation

We extracted the deck letter from the cabin since it could potentially correlate to the survival.

```
#Extract deck letter from cabin
data.clean$deck <- substr(data.clean$cabin, 1,1)
# Remove cabin col:
data.clean$cabin <- NULL</pre>
```

4.4 Initial Check for Missing values

We see that age and deck have the most amount of missing data, therefore we proceed to impute them.

```
print(plot_missing_barchart(data.clean))
```

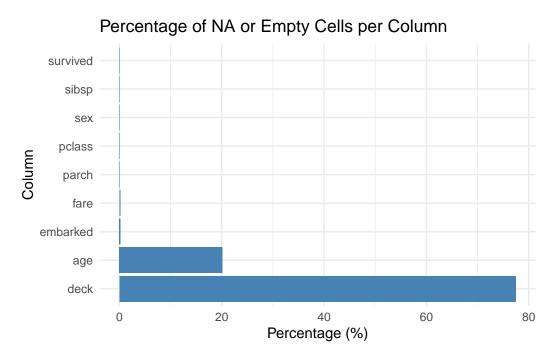


Figure 1: Percentage of Missing Values

4.5 Imputing data

Below we impute Age using the median value in that column.

For deck we use KNN to impute the missing deck values.

After imputing these two columns we can see that the largest amount of missing data is $\sim 0.2\%$ which is quite small and can be removed.

```
# ---- Age----
#Replace NAs in age column with Median value
set.seed (1023)
median_age <- median(data.clean$age, na.rm = TRUE)
data.clean <- data.clean %>%
  mutate(age = ifelse(is.na(age), median_age, age))
```

```
# ---- deck----
# For deck, since its a category, we decided to use KNN to impute the column:
# Install if not already installed
# install.packages("VIM")
library(VIM)
Warning: package 'VIM' was built under R version 4.4.3
Loading required package: colorspace
Loading required package: grid
VIM is ready to use.
Suggestions and bug-reports can be submitted at: https://github.com/statistikat/VIM/issues
Attaching package: 'VIM'
The following object is masked from 'package:datasets':
    sleep
# Replace "" with NA in the 'deck' column
data.clean$deck[data.clean$deck == ""] <- NA
# Convert 'cabin' to factor
data.clean$deck <- as.factor(data.clean$deck)</pre>
# Apply kNN imputation just to Cabin column
set.seed (1023)
data.clean <- kNN(data.clean, variable = "deck", k = 5)
# Check that NAs were imputed
# sum(is.na(data.clean$deck))
                                    # Original
# sum(is.na(data.clean.imputed$deck)) # After
# Remove indicator col:
data.clean$deck_imp <- NULL
```

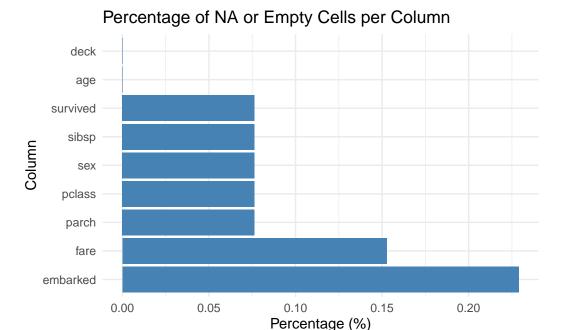


Figure 2: Percentage of Missing Values after Imputation

4.6 Dummifying Columns:

We dummify pclass, sex, embarked and deck. We leave sibsp and parch as continuous variables as we observed that dummifying these columns leads to smaller significance (See Appendix A), whilst leaving them as continuous maximizes their contributions to the models explanatory power.

```
# Dummifying pclass:
data.clean$pclass_1 = ifelse(data.clean$pclass == 1, 1, 0)
data.clean$pclass_2 = ifelse(data.clean$pclass == 2, 1, 0)
# Dummifying sex:
data.clean$sex_M = ifelse(data.clean$sex == 'male', 1, 0)
# Dummifying embarked:
data.clean$embarked_C = ifelse(data.clean$embarked == 'C', 1, 0)
data.clean$embarked_Q = ifelse(data.clean$embarked == 'Q', 1, 0)
# Dummifying deck:
data.clean$deck_A = ifelse(data.clean$deck == 'A', 1, 0)
data.clean$deck B = ifelse(data.clean$deck == 'B', 1, 0)
data.clean$deck_C = ifelse(data.clean$deck == 'C', 1, 0)
data.clean$deck_D = ifelse(data.clean$deck == 'D', 1, 0)
data.clean$deck_E = ifelse(data.clean$deck == 'E', 1, 0)
data.clean$deck_F = ifelse(data.clean$deck == 'F', 1, 0)
data.clean$deck_G = ifelse(data.clean$deck == 'G', 1, 0)
# Removing Dummified cols:
data.clean = subset(data.clean, select = -c(pclass, sex, embarked,deck))
```

4.7 Remove NA rows and deck_G

Below we remove NA rows, which turned out to be only 2 after proper cleaning and imputation. We also removed deck_G as we observed that it has a large skew in the data distribution with only 13 people allocated in this deck. It was observed that this variable lead to erroneous predictions in the model.

```
# Plot histogram of the 'values' column
hist(data.clean$deck_G,
    main = "Histogram of Values",
    xlab = "Values",
    col = "skyblue",
    border = "white")
```

Histogram of Values

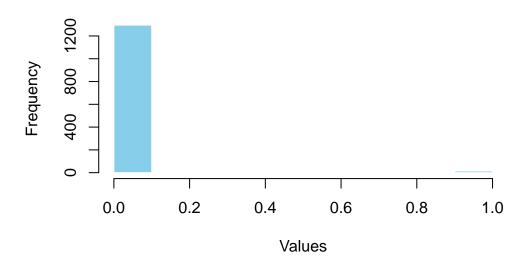


Figure 3: Histogram of Deck_G

```
# Removing deck_G col:
data.clean = subset(data.clean, select = -c(deck_G))

data.clean = na.omit(data.clean)
cat(nrow(odata) - nrow(data.clean), 'rows were removed from original dataset')
```

2 rows were removed from original dataset

4.8 Divide into Test / Train

Finally we divide into 70% training data and 30% test data.

```
set.seed (1023)
train_indices = sample(1 : nrow(data.clean), size = 0.7005*nrow(data.clean), replace = FALSE)
train = data.clean[train_indices,]
test = data.clean[-train_indices,]
cat("We are using:", nrow(train)/nrow(data.clean) * 100, '% of the data for training')
```

4.9 EDA

Using the training data set we use a variety of method to draw some initial conclusions:

- Histogram: Showing that more people in their late teens up to late thirties survived.
- Bar chart showing that more people died than survived
- Bar chart showing that a higher number of people survived when they had less siblings on board.
- Correlation matrix shows that sex and Deck_F are highly negatively correlated to survival. There is a soft positive correlation to pclass_1.
- There is a high correlation between pclass_1 and fare, this justifies that one of these predictors can potentially be removed.
- The scatter plots did not give us much more information on the relation between the predictors and the dependent variable.

```
# Histogram showing that more people in their late teens up to late thirties survived.
ggplot(train, aes(age)) +
geom_histogram(bins=30)
```

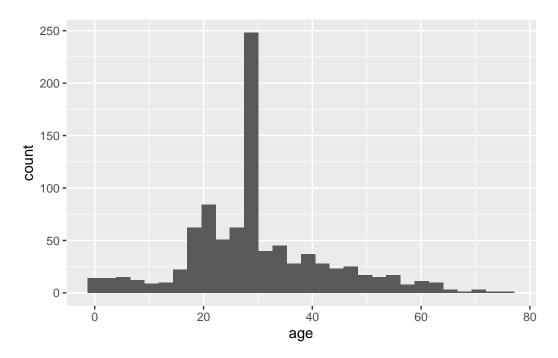


Figure 4: Histogram of survival vs age

```
# Bar chart showing that more people died than survived
ggplot(train, aes(survived)) +
  geom_bar()
```

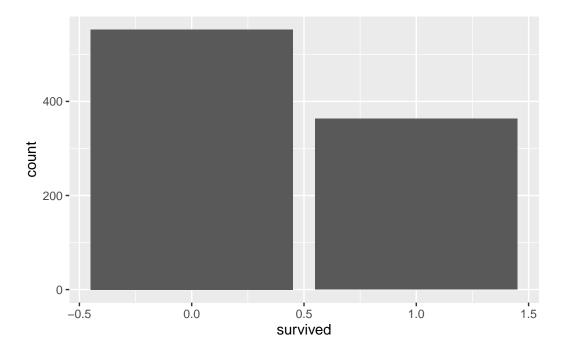


Figure 5: Barchart of survival

```
# Bar chart showing that a higher number of people survived when they had less
# siblings on board.
ggplot(train, aes(sibsp, survived)) +
  geom_bar(stat='identity')
```

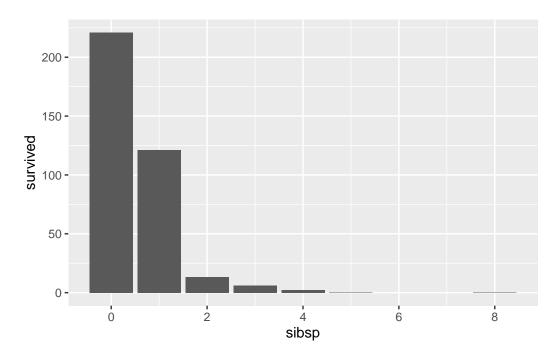


Figure 6: Barchart of survival vs Num. of siblings

cor(train)

survived age sibsp parch fare

```
survived
          1.00000000 -0.04599216 -0.02754365 0.095720424 0.24962042
         -0.04599216
                   1.00000000 -0.15495239 -0.122885529 0.16447722
age
         -0.02754365 -0.15495239 1.00000000 0.355216328 0.16529618
sibsp
parch
         0.24962042  0.16447722  0.16529618  0.209878267  1.00000000
fare
pclass_1 0.28432063 0.34847627 -0.02230662 -0.016451014 0.59116304
          pclass_2
         -0.53377358 0.05385130 -0.12706675 -0.243501842 -0.20177640
sex M
embarked_C 0.14757622 0.05701305 -0.06121524 -0.001165488 0.27208779
embarked_Q -0.02542950 -0.03373092 -0.06022107 -0.093253422 -0.12699238
deck_A
          deck_B
          deck_C
        0.17861407 0.15235605 0.02822826 -0.049048850 0.30882422
deck_D
         \mathtt{deck}_{\mathtt{E}}
deck_F
         -0.47745405 -0.29217392 0.06499841 0.015200439 -0.42777519
           pclass_1
                     pclass_2
                                  sex M
                                         embarked_C embarked_Q
         survived
         0.34847627 0.01898825 0.05385130 0.057013047 -0.03373092
age
         -0.02230662 -0.06908924 -0.12706675 -0.061215242 -0.06022107
sibsp
         -0.01645101 -0.01769672 -0.24350184 -0.001165488 -0.09325342
parch
         0.59116304 -0.12563735 -0.20177640 0.272087792 -0.12699238
fare
         1.00000000 -0.30255006 -0.11706927 0.285611052 -0.16373859
pclass 1
pclass_2
         -0.30255006 1.00000000 -0.01057862 -0.143425262 -0.14105035
         -0.11706927 -0.01057862 1.00000000 -0.040151739 -0.09800681
sex_M
embarked_C 0.28561105 -0.14342526 -0.04015174 1.000000000 -0.16362864
embarked_Q -0.16373859 -0.14105035 -0.09800681 -0.163628636 1.00000000
deck_A
         0.28790374 -0.08080438 0.05888172 0.167298858 -0.05745333
deck_B
          0.42601897 -0.11327680 -0.09133967 0.182056642 -0.08108169
        0.48876162 -0.07908768 -0.09801461 0.169699800 -0.08023736
\mathtt{deck}_{\mathtt{C}}
deck_D
        0.15200710 0.04093807 -0.03323992 0.248144405 -0.07224766
deck E
         0.05822613 -0.02663532 -0.08900107 -0.078527161 -0.08770239
         deck F
             deck A
                       deck B
                                 deck C
                                            deck D
                                                      deck E
         survived
          0.11573201 0.12638063 0.15235605 0.10150497 0.11691545
age
sibsp
         -0.06315491 -0.01583908 0.02822826 -0.02726397 -0.07405912
parch
         -0.05405176 0.07625540 -0.04904885 -0.02902970 -0.01633207
fare
         0.06359762  0.45645334  0.30882422  0.01777467  -0.01633414
pclass_1
         0.28790374  0.42601897  0.48876162  0.15200710  0.05822613
         -0.08080438 -0.11327680 -0.07908768 0.04093807 -0.02663532
pclass_2
          0.05888172 -0.09133967 -0.09801461 -0.03323992 -0.08900107
sex_M
embarked_C 0.16729886 0.18205664 0.16969980 0.24814440 -0.07852716
embarked_Q -0.05745333 -0.08108169 -0.08023736 -0.07224766 -0.08770239
\mathtt{deck}_{\mathtt{A}}
         1.00000000 -0.04614047 -0.06005247 -0.04955727 -0.06365303
deck_B
         -0.04614047 1.00000000 -0.08474977 -0.06993828 -0.08983109
deck C
         -0.06005247 -0.08474977 1.00000000 -0.09102568 -0.11691645
deck D
         -0.04955727 -0.06993828 -0.09102568 1.00000000 -0.09648328
         -0.06365303 -0.08983109 -0.11691645 -0.09648328 1.00000000
deck E
deck_F
         -0.22835159 -0.32226388 -0.41943104 -0.34612824 -0.44457878
             deck F
survived
         -0.47745405
age
         -0.29217392
```

```
sibsp
            0.06499841
            0.01520044
parch
           -0.42777519
fare
pclass_1
           -0.72064119
            0.14341225
pclass_2
sex_M
            0.15235799
embarked_C -0.31968879
embarked Q 0.21273851
deck_A
           -0.22835159
deck_B
           -0.32226388
deck_C
           -0.41943104
deck_D
           -0.34612824
           -0.44457878
deck_E
            1.00000000
deck_F
```

```
pairs(train[c(1:4,7,13)])
```

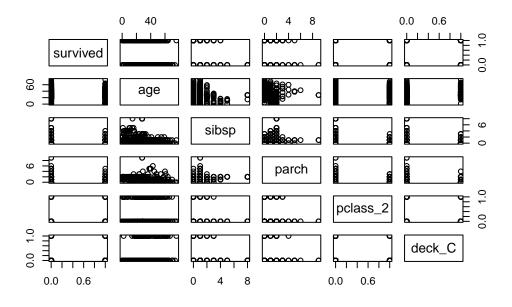


Figure 7: Scatter plots of all variables in train data

Since this data is mainly categorical, the scatterplot and correlation matrix are not very useful.

(Statology, 2025) is used to develop the correlation values between our categorical columns. This describes the use of pysch and reompanion.

```
#install.packages("psych")
library(psych) # [@statology2025] to understand how this works
```

Warning: package 'psych' was built under R version 4.4.3

Attaching package: 'psych'

```
The following objects are masked from 'package:ggplot2':
    %+%, alpha
tetrachoric(train[, c("survived", "sex_M")])
Call: tetrachoric(x = train[, c("survived", "sex_M")])
tetrachoric correlation
         srvvd sex_M
survived 1.00
        -0.75 1.00
sex_M
with tau of
survived sex M
   0.26 -0.35
tetrachoric(train[, c("survived", "pclass_1")])
Call: tetrachoric(x = train[, c("survived", "pclass_1")])
tetrachoric correlation
         srvvd pcl_1
survived 1.00
pclass_1 0.46 1.00
with tau of
survived pclass_1
    0.26
           0.69
tetrachoric(train[, c("survived", "pclass_2")])
Call: tetrachoric(x = train[, c("survived", "pclass_2")])
tetrachoric correlation
         srvvd pcl_2
survived 1.00
pclass_2 0.11 1.00
 with tau of
survived pclass_2
    0.26
           0.77
#install.packages("rcompanion")
library(rcompanion) # Reference 4 to understand how this works.
Warning: package 'rcompanion' was built under R version 4.4.3
Attaching package: 'rcompanion'
```

```
The following object is masked from 'package:psych': phi
```

```
cramerV(train$survived, train$sex)
```

Cramer V 0.5338

```
library(corrplot)
```

corrplot 0.95 loaded

```
cor_matrix <- cor(train)#[,1]
corrplot(cor_matrix, method = "circle")</pre>
```

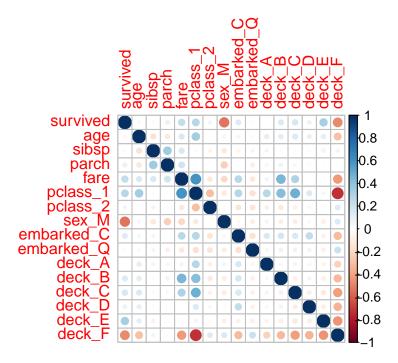


Figure 8: Correlation Matrix

5 Model Development Process

The data was properly cleaned and divided into train/test in the prior section.

Here we train a binary model. The Q-Q plot shows that the residuals are indeed normally distributed so a transformation is potentially not necessary.

The statistical comparison between test and train data shows that the model is very stable with an accuracy of $\sim 84\%$ for both.

We also analyzed the VIF and we see that there is high degree of correlation between the decks, this provides justification to remove some of the decks as predictors.

```
library(car)
```

```
Warning: package 'car' was built under R version 4.4.3
Loading required package: carData
Warning: package 'carData' was built under R version 4.4.3
Attaching package: 'car'
The following object is masked from 'package:psych':
    logit
The following object is masked from 'package:dplyr':
    recode
The following object is masked from 'package:purrr':
    some
# Log model on train data:
set.seed (1023)
lmod <- glm(survived ~ ., family = binomial, data = train)</pre>
# summary(lmod)
vif(lmod)
                sibsp
                            parch
                                        fare
                                               pclass_1
                                                           pclass_2
                                                                          sex_M
       age
  1.487819
             1.241500
                         1.283338
                                    1.720620
                                                4.661296
                                                           1.574259
                                                                       1.619104
embarked_C embarked_Q
                           \mathtt{deck}_{\mathtt{A}}
                                      deck_B
                                                deck_C
                                                            deck_D
                                                                         deck_E
  1.433942
             1.368562
                       5.122014
                                    6.161202
                                              9.446197
                                                          7.824253
                                                                       7.280808
    deck_F
 18.174852
y_hat_log_train <- predict(lmod, data = train, type="response")</pre>
predictions_log_train <- ifelse(y_hat_log_train > 0.5, 1, 0)
y_hat_log_test<-predict(lmod, newdata = test, type="response")</pre>
predictions_log_test <- ifelse(y_hat_log_test > 0.5, 1, 0)
confusion_matrix_log_train <- confusionMatrix(as.factor(predictions_log_train), as.factor(train$surv
base.model.accuracy = confusion_matrix_log_train$overall['Accuracy']
base.model.f1 = confusion_matrix_log_train$byClass['F1']
base.model.train.summary = data.frame(
 Accuracy = base.model.accuracy,
 F1 = base.model.f1
)
```

```
row.names(base.model.train.summary) <- 'base.model.train'

confusion_matrix_log_test <- confusionMatrix(as.factor(predictions_log_test), as.factor(test$survived)
base.model.accuracy = confusion_matrix_log_test$overall['Accuracy']
base.model.f1 = confusion_matrix_log_test$byClass['F1']
base.model.test.summary = data.frame(
    Accuracy = base.model.accuracy,
    F1 = base.model.f1
)
row.names(base.model.test.summary) <- 'base.model.test'</pre>
```

data.frame(rbind(base.model.train.summary, base.model.test.summary))

Table 2: Summary Stats. of base log. model

	Accuracy	F1
base.model.train base.model.test	0.000 -000	$\begin{array}{c} 0.8115543 \\ 0.7758007 \end{array}$

summary(lmod)

```
Call:
glm(formula = survived ~ ., family = binomial, data = train)
Coefficients:
             Estimate Std. Error z value Pr(>|z|)
                                   6.042 1.52e-09 ***
(Intercept) 5.386e+00 8.913e-01
age
           -5.396e-02 9.051e-03 -5.962 2.49e-09 ***
           -2.538e-01 1.205e-01 -2.106 0.03520 *
sibsp
           1.371e-02 1.185e-01 0.116 0.90790
parch
fare
           -6.976e-05 2.289e-03 -0.030 0.97568
pclass_1
           7.953e-01 4.670e-01
                                   1.703 0.08856
pclass_2
           1.829e+00 3.011e-01
                                   6.073 1.25e-09 ***
sex_M
           -3.274e+00 2.566e-01 -12.756 < 2e-16 ***
embarked_C 8.299e-01 2.877e-01 2.884 0.00393 **
embarked_Q 7.944e-01 3.602e-01 2.205 0.02744 *
           -2.166e+00 1.011e+00 -2.143 0.03212 *
deck_A
           -1.510e+00 1.006e+00 -1.501 0.13338
\mathtt{deck}_{\mathtt{B}}
           -1.646e+00 9.553e-01 -1.723 0.08484 .
\mathtt{deck}_{\mathtt{C}}
deck_D
           -2.616e+00 9.310e-01 -2.810 0.00495 **
           4.274e-01 8.975e-01
deck_E
                                  0.476 0.63393
           -4.342e+00 8.572e-01 -5.065 4.08e-07 ***
deck_F
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

(Dispersion parameter for binomial family taken to be 1)

Null deviance: 1230.15 on 915 degrees of freedom Residual deviance: 634.78 on 900 degrees of freedom

AIC: 666.78

Number of Fisher Scoring iterations: 6

```
par(mfrow=c(2,2))
plot(lmod)
```

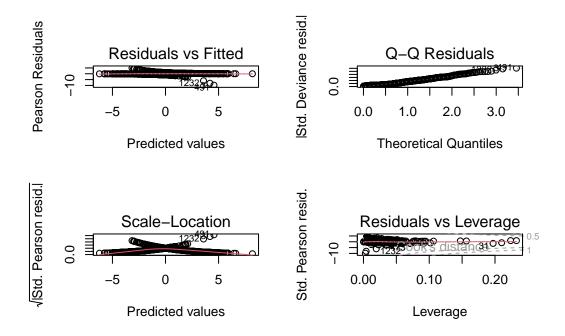


Figure 9: 4x4 standard plots for log. model

Plots show that a linear model is not appropriate for this data.

6 Model Performance Testing

We compare four different models:

- 1) Base log. model (lmod)
- 2) Model with insig. pred. removed (filtered.model)
- 3) Stepwise model (step.model)
- 4) Model with high vif pred. removed (vif.model)

When comparing the accuracy and F1 score on all models the filtered model was the highest performer and we decided to use that as the champion model until now. We suspect that multicollinearity among the decks has an impact on the model performance which is why removing these collinear predictors and filtering for insignificant variables improved the model slightly against the based model (+0.3%).

```
Function to remove Insig. Predictors one by one
backward eliminate = function(model, alpha = 0.05) {
 repeat {
   d1 = drop1(model, test = "F")
   # Get p-values excluding intercept row
   pvals = d1^{r(>F)}[-1]
   # Stop if all predictors are significant or only intercept left
   if( all(is.na(pvals)) || max(pvals, na.rm = TRUE) <= alpha ){</pre>
    print("all variable are signifcant")
    break
   }
   # Remove the term with max p-value
   term_to_remove = rownames(d1)[-1][which.max(pvals)]
   cat("Removing:", term_to_remove, "with p-value", max(pvals, na.rm = TRUE), "\n")
   model = update(model, paste(". ~ . -", term_to_remove))
 return(model)
```

```
Function to remove Cooks Outliers
# model_formula: A formula object, e.g. PSA.level ~ .
# data: A data frame containing the variables in the model.
# threshold: A numeric value indicating the Cook's D threshold (default 0.5).
# print: If TRUE (default) will print the rows beign removed.
# returns: A list with the final model and the filtered dataset.
# Example usage:
# result = remove_cooks_outliers(PSA.level ~ ., mydata)
# summary(result$model)
# str(result$filtered_data)
remove_cooks_outliers = function(model_formula, data, threshold = 0.5,
                            print = TRUE)
 {
 all_high_cd_rows = data.frame() # to store all removed rows
 repeat {
   model = glm(model_formula, family = binomial, data = data)
   cooksD = cooks.distance(model)
   high_cd_indices = which(cooksD > threshold)
   if (length(high_cd_indices) == 0) { # If there are no more outliers
     break
   }
```

```
if (print == TRUE){
    cat("Removing rows with Cook's D >", threshold, ":\n", high_cd_indices, "\n")
   # Save these outliers before removing them
   high_cd_rows = data[high_cd_indices, ]
   all_high_cd_rows = rbind(all_high_cd_rows, high_cd_rows)
   # Update data by removing high CD rows for next iteration.
   data = data[-high_cd_indices, ]
 }
 final model = glm(model formula, family = binomial, data = data)
 return(list(model = final model, filtered data = data, high cd data = all high cd rows))
Model with insig. variables removed
binary.model.filtered = backward_eliminate(lmod)
Warning in drop1.glm(model, test = "F"): F test assumes 'quasibinomial' family
Removing: fare with p-value 0.9710719
Warning in drop1.glm(model, test = "F"): F test assumes 'quasibinomial' family
Removing: parch with p-value 0.894389
Warning in drop1.glm(model, test = "F"): F test assumes 'quasibinomial' family
Removing: deck_E with p-value 0.5871368
Warning in drop1.glm(model, test = "F"): F test assumes 'quasibinomial' family
[1] "all variable are significant"
step wise model with insig. variables removed
library(olsrr)
Warning: package 'olsrr' was built under R version 4.4.3
```

Attaching package: 'olsrr'

The following object is masked from 'package:MASS':

cement

The following object is masked from 'package:datasets':

rivers

ols_step_both_p(lmod,p_enter=0.1,p_remove=0.05,details=FALSE)

Stepwise Summary

Step	Variable	AIC	SBC	SBIC	R2	Adj. R2
0	Base Model	1293.366	1303.006	-28281799.310	0.00000	0.00000
1	sex_M (+)	988.183	1002.643	-55309468.385	0.28491	0.28413
2	deck_F (+)	757.110	776.390	-92005909.772	0.44556	0.44435
3	age (+)	723.096	747.196	-99521155.302	0.46694	0.46518
4	pclass_2 (+)	697.456	726.376	-105700242.392	0.48278	0.48051
5	deck_E (+)	674.236	707.976	-111671761.113	0.49683	0.49406
6	sibsp (+)	668.302	706.862	-113609066.331	0.50117	0.49788
7	deck_D (+)	664.319	707.699	-115088607.776	0.50442	0.50059
8	<pre>embarked_Q (+)</pre>	663.468	711.668	-115792422.465	0.50596	0.50160
9	embarked_Q (-)	664.319	707.699	-115088607.776	0.50442	0.50059

Final Model Output

Model Summary

R	0.710	RMSE	0.344
R-Squared	0.504	MSE	0.119
Adj. R-Squared	0.501	Coef. Var	87.272
Pred R-Squared	0.495	AIC	664.319
MAE	0.262	SBC	707.699

RMSE: Root Mean Square Error

MSE: Mean Square Error MAE: Mean Absolute Error

AIC: Akaike Information Criteria SBC: Schwarz Bayesian Criteria

ANOVA

	Sum of Squares	DF	Mean Square	F	Sig.
Regression	110.541	7	15.792	132.026	0.0000
Residual	108.606	908	0.120		
Total	219.147	915			

Parameter Estimates

model	Beta	Std. Error	Std. Beta	t	Sig	lower	upper
(Intercept)	1.131	0.043		26.050	0.000	1.046	1.216
sex_M	-0.464	0.024	-0.457	-19.025	0.000	-0.512	-0.416
deck_F	-0.447	0.031	-0.445	-14.444	0.000	-0.508	-0.387
age	-0.006	0.001	-0.171	-6.871	0.000	-0.008	-0.005
pclass_2	0.147	0.028	0.124	5.193	0.000	0.091	0.202
deck_E	0.169	0.043	0.108	3.942	0.000	0.085	0.253
sibsp	-0.031	0.011	-0.068	-2.848	0.005	-0.052	-0.010
deck_D	-0.123	0.050	-0.064	-2.439	0.015	-0.222	-0.024

```
Call:
glm(formula = survived ~ sex_M + deck_F + age + deck_E + pclass_2 +
    pclass_1 + sibsp + fare, family = binomial, data = train)
```

Coefficients:

```
Estimate Std. Error z value Pr(>|z|)
(Intercept) 4.039203 0.452351 8.929 < 2e-16 ***
        -3.223591 0.242196 -13.310 < 2e-16 ***
sex_M
        -2.635131 0.348620 -7.559 4.07e-14 ***
deck F
        age
        1.898805 0.387471 4.901 9.56e-07 ***
deck_E
        1.407566 0.270017 5.213 1.86e-07 ***
pclass_2
        0.453560 0.388459 1.168 0.24297
pclass_1
        sibsp
         0.001986
                 0.002090 0.951 0.34181
fare
```

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for binomial family taken to be 1)

Null deviance: 1230.15 on 915 degrees of freedom Residual deviance: 654.59 on 907 degrees of freedom

AIC: 672.59

Number of Fisher Scoring iterations: 6

```
# Removing fare from stepwise model since its insig.:
binary.model.stepwise = update(binary.model.stepwise, paste(". ~ . -", 'fare'))
summary(binary.model.stepwise)
```

```
Call:
glm(formula = survived ~ sex_M + deck_F + age + deck_E + pclass_2 +
   pclass_1 + sibsp, family = binomial, data = train)
Coefficients:
          Estimate Std. Error z value Pr(>|z|)
(Intercept) 4.089766 0.450105 9.086 < 2e-16 ***
         -3.246187
                  0.241244 -13.456 < 2e-16 ***
sex_M
        -2.659167  0.348576  -7.629  2.37e-14 ***
\mathtt{deck}_{\mathtt{F}}
         age
         1.874294   0.386905   4.844   1.27e-06 ***
\mathtt{deck}_{\mathtt{E}}
         1.432280    0.269447    5.316    1.06e-07 ***
pclass_2
         pclass_1
         sibsp
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
   Null deviance: 1230.15 on 915 degrees of freedom
Residual deviance: 655.53 on 908 degrees of freedom
AIC: 671.53
Number of Fisher Scoring iterations: 6
Model with influential points removed
# Applying function to remove influential points via Cooks Distance:
filtering.result = remove_cooks_outliers(survived ~ ., data = train)
cat(nrow(filtering.result$high_cd_data), 'rows were identified as outliers')
O rows were identified as outliers
# plot(logmod)
# Since there were now rows identified as outliers, then this model will be the
# same as the initial binary model and need no be considered in the final
# model compare.
Model with high VIF preds removed
library(car)
vif(lmod)
                                    pclass_1
                                            pclass_2
                                                       sex M
     age
            sibsp
                     parch
                              fare
 1.487819
        1.241500
                  1.283338
                           1.720620
                                    4.661296
                                            1.574259
                                                     1.619104
embarked_C embarked_Q
                    deck A
                             deck B
                                    deck C
                                             deck D
                                                       deck E
 1.433942
          1.368562
                  5.122014
                           6.161202
                                    9.446197
                                            7.824253
                                                     7.280808
```

```
deck_F
18.174852
```

```
vif.model = glm(survived ~ . -deck_F, family = binomial, data = train)
vif(vif.model)
                 sibsp
                            parch
                                         fare
                                                 pclass_1
                                                            pclass_2
                                                                            sex_M
       age
  1.477874
              1.230404
                          1.274009
                                     1.720786
                                                 4.741756
                                                             1.481905
                                                                         1.468964
embarked_C embarked_Q
                            deck_A
                                       \mathtt{deck}_{\mathtt{B}}
                                                   deck_C
                                                               deck_D
                                                                           deck_E
  1.420600
              1.325908
                          1.986762
                                     2.357338
                                                 3.004607
                                                             1.736194
                                                                         1.557128
# Removing deck F from the model eliminates the multicollinearity completely.
```

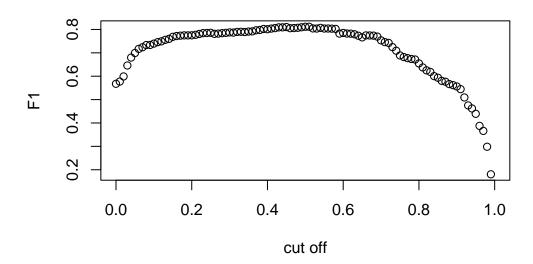
```
Function to calc. cutoff
cutoff.prg<-function(pred,act){</pre>
# pred<-predicted_probabilities</pre>
# act<-true_labels</pre>
p < -seq(0,1,0.01)
n<-length(p)
out <- matrix (0, nrow=n, ncol=12)
for(i in 1:n){
predictions <- ifelse(pred >p[i], 1, 0)
confusion_matrix <- confusionMatrix(as.factor(predictions),as.factor(act),mode="prec_recall", positive to the confusion of th
out[i,]<-cbind(p=p[i],t(confusion_matrix[[4]]))</pre>
dimnames(out)[[2]] <-c("p", "Sensitivity", "Specificity", "Pos Pred Value", "Neg Pred Value", "Precision",
out
}
# Finding the optimal cutoff
observations = train$survived
prob <- predict(lmod, train, type="response")</pre>
```

```
Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), :
Levels are not in the same order for reference and data. Refactoring data to
match.
Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), :
```

test_cutoff<-cutoff.prg(prob,observations)</pre>

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

```
plot(test_cutoff[,1],test_cutoff[,8],xlab="cut off",ylab="F1")
```



```
optimal.cutoff = test_cutoff[which.max(test_cutoff[,8]),]
optimal.cutoff[1]
```

р 0.5

```
observations.train = train$survived
observations.test = test$survived
# Confusion matrix on base log model on train data
y_hat_prob = predict(lmod, train, type="response")
predictions.binary.model.train <- ifelse(y_hat_prob > optimal.cutoff[1], 1, 0)
confusion.matrix.binary.model.train <- confusionMatrix(as.factor(predictions.binary.model.train),as.</pre>
base.model.accuracy = confusion.matrix.binary.model.train$overall['Accuracy']
base.model.f1 = confusion.matrix.binary.model.train$byClass['F1']
base.model.train.summary = data.frame(
  Accuracy = base.model.accuracy,
 F1 = base.model.f1
row.names(base.model.train.summary) <- 'base.model.train'</pre>
# Confusion matrix on base log model on test data
y_hat_prob = predict(lmod, test, type="response")
predictions.binary.model.test <- ifelse(y_hat_prob > optimal.cutoff[1], 1, 0)
confusion.matrix.binary.model.test <- confusionMatrix(as.factor(predictions.binary.model.test),as.fac
base.model.accuracy.test = confusion.matrix.binary.model.test$overall['Accuracy']
base.model.f1.test = confusion.matrix.binary.model.test$byClass['F1']
base.model.test.summary = data.frame(
```

```
Accuracy = base.model.accuracy.test,
  F1 = base.model.f1.test
row.names(base.model.test.summary) <- 'base.model.test'</pre>
# Confusion matrix on stepwise log model on train data
y_hat_prob = predict(binary.model.stepwise, train, type="response")
predictions.binary.model.step.train <- ifelse(y_hat_prob > optimal.cutoff[1], 1, 0)
confusion.matrix.binary.model.step.train <- confusionMatrix(as.factor(predictions.binary.model.step.
stepwise.model.accuracy = confusion.matrix.binary.model.step.train$overall['Accuracy']
stepwise.model.f1 = confusion.matrix.binary.model.step.train$byClass['F1']
stepwise.model.train.summary = data.frame(
  Accuracy = stepwise.model.accuracy,
  F1 = stepwise.model.f1
row.names(stepwise.model.train.summary) <- 'stepwise.model.train'
# Confusion matrix on stepwise log model on test data
y_hat_prob = predict(binary.model.stepwise, test, type="response")
predictions.binary.model.step.test <- ifelse(y_hat_prob > optimal.cutoff[1], 1, 0)
confusion.matrix.binary.model.step.test <- confusionMatrix(as.factor(predictions.binary.model.step.te
stepwise.model.accuracy = confusion.matrix.binary.model.step.test$overall['Accuracy']
stepwise.model.f1 = confusion.matrix.binary.model.step.test$byClass['F1']
stepwise.model.test.summary = data.frame(
  Accuracy = stepwise.model.accuracy,
  F1 = stepwise.model.f1
row.names(stepwise.model.test.summary) <- 'stepwise.model.test'</pre>
# Confusion matrix on log model w/ insig. pred. removed
y_hat_prob = predict(binary.model.filtered, train, type="response")
predictions.binary.model.filtered.train <- ifelse(y_hat_prob > optimal.cutoff[1], 1, 0)
confusion.matrix.binary.model.filtered.train <- confusionMatrix(as.factor(predictions.binary.model.fi
filtered.model.accuracy = confusion.matrix.binary.model.filtered.train$overall['Accuracy']
filtered.model.f1 = confusion.matrix.binary.model.filtered.train$byClass['F1']
filtered.model.train.summary = data.frame(
  Accuracy = filtered.model.accuracy,
  F1 = filtered.model.f1
row.names(filtered.model.train.summary) <- 'filtered.model.train'</pre>
# Confusion matrix on log model w/ insig. pred. removed
y_hat_prob = predict(binary.model.filtered, test, type="response")
predictions.binary.model.filtered.test <- ifelse(y_hat_prob > optimal.cutoff[1], 1, 0)
confusion.matrix.binary.model.filtered.test <- confusionMatrix(as.factor(predictions.binary.model.fil
filtered.model.accuracy = confusion.matrix.binary.model.filtered.test$overall['Accuracy']
filtered.model.f1 = confusion.matrix.binary.model.filtered.test$byClass['F1']
filtered.model.test.summary = data.frame(
```

```
Accuracy = filtered.model.accuracy,
 F1 = filtered.model.f1
row.names(filtered.model.test.summary) <- 'filtered.model.test'
# Confusion matrix on log model w/ high vif pred. removed
y_hat_prob = predict(vif.model, train, type="response")
predictions.binary.model.vif.train <- ifelse(y_hat_prob > optimal.cutoff[1], 1, 0)
confusion.matrix.binary.model.vif.train <- confusionMatrix(as.factor(predictions.binary.model.vif.train
vif.model.accuracy = confusion.matrix.binary.model.vif.train$overall['Accuracy']
vif.model.f1 = confusion.matrix.binary.model.vif.train$byClass['F1']
vif.model.train.summary = data.frame(
  Accuracy = vif.model.accuracy,
  F1 = vif.model.f1
row.names(vif.model.train.summary) <- 'vif.model.train'</pre>
# Confusion matrix on log model w/ high vif pred. removed
y_hat_prob = predict(vif.model, test, type="response")
predictions.binary.model.vif.test <- ifelse(y_hat_prob > optimal.cutoff[1], 1, 0)
confusion.matrix.binary.model.vif.test <- confusionMatrix(as.factor(predictions.binary.model.vif.test
vif.model.accuracy = confusion.matrix.binary.model.vif.test$overall['Accuracy']
vif.model.f1 = confusion.matrix.binary.model.vif.test$byClass['F1']
vif.model.test.summary = data.frame(
  Accuracy = vif.model.accuracy,
 F1 = vif.model.f1
row.names(vif.model.test.summary) <- 'vif.model.test'</pre>
data.frame(rbind(base.model.train.summary, base.model.test.summary,
                 stepwise.model.train.summary, stepwise.model.test.summary,
                 filtered.model.train.summary, filtered.model.test.summary,
                 vif.model.train.summary,vif.model.test.summary))
```

Table 3: Comparison of all log models performance

	Accuracy	F1
base.model.train	0.8504367	0.8115543
base.model.test	0.8392857	0.7758007
stepwise.model.train	0.8231441	0.7692308
stepwise.model.test	0.8367347	0.7611940
filtered.model.train	0.8504367	0.8115543
filtered.model.test	0.8418367	0.7801418
vif.model.train	0.8351528	0.7911480
vif.model.test	0.8392857	0.7789474

7 Challenger Models

We decided to build a Poisson model for our challenger model. After building our original poisson model, we first check to see if the variables are important.

Ho: Variables are not important

Ha: Variable are important

library (MASS)

Since we have a p-value of 0, we reject the null hypothesis. Variables are important.

Our poisson model has three significant variables, pclass 2, sex M, and deck F with an alpha of 0.05.

According to the poisson regression:

Train base Poisson model:

The odds of survival increases by 37.6% when the passenger is second class. ((exp(0.31918) - 1) * 100)

The odds of survival decreases by 68.14% when the passenger is male. ((exp(-1.144) - 1) * 100)

The odds of survival decreases by 64.76% when the passenger is from deck G. ((exp(-1.043) - 1) * 100)

Our Poisson Regression has an accuracy of 74.74% with optimal cutoff(.17) based on max F1 score (0.71954).

```
poissonReg_full <- glm(survived ~ .,family=poisson, train)</pre>
# compare the fitted model to the null model and calculate if the variables are important
summary(poissonReg_full)
Call:
glm(formula = survived ~ ., family = poisson, data = train)
Coefficients:
              Estimate Std. Error z value Pr(>|z|)
(Intercept) 7.787e-01 3.449e-01 2.258 0.023946 *
            -1.609e-02 4.355e-03 -3.695 0.000220 ***
age
sibsp
            -6.954e-02 6.979e-02 -0.996 0.319019
parch
            -5.055e-04 6.756e-02 -0.007 0.994030
            9.289e-05 1.025e-03 0.091 0.927778
fare
           3.590e-01 2.047e-01 1.753 0.079536 .
pclass_1
pclass_2
           5.102e-01 1.532e-01 3.329 0.000871 ***
           -1.095e+00 1.201e-01 -9.112 < 2e-16 ***
{\tt sex\_M}
embarked_C 1.072e-01 1.349e-01 0.795 0.426660
embarked_Q
           4.537e-01 2.162e-01 2.098 0.035872 *
            -6.157e-01 4.732e-01 -1.301 0.193282
\mathtt{deck}_{\mathtt{A}}
            -6.494e-01 4.130e-01 -1.572 0.115868
deck_B
            -6.891e-01 3.943e-01 -1.748 0.080527 .
\mathtt{deck}_{\mathtt{C}}
deck_D
            -7.195e-01 3.962e-01 -1.816 0.069347 .
            -2.300e-01 3.556e-01 -0.647 0.517786
\mathtt{deck}_{\mathtt{E}}
            -1.527e+00 3.419e-01 -4.467 7.94e-06 ***
deck_F
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for poisson family taken to be 1)
    Null deviance: 672.00 on 915 degrees of freedom
Residual deviance: 404.15 on 900 degrees of freedom
```

AIC: 1162.2

```
1-pchisq(672.00-409.95, length(coef(poissonReg_full)) - 1)
```

[1] 0

```
# We then reduce the model and see if the removed variables are significant
poissonReg_reduced <- glm(survived ~ pclass_2+sex_M+deck_F,family=poisson, train)
# Verify that all variables are significant and no more can be dropped
anova(poissonReg_reduced, poissonReg_full, test="Chi")</pre>
```

Resid. Df	Resid. Dev	Df	Deviance	Pr(>Chi)
912	430.4971	NA	NA	NA
900	404.1538	12	26.34334	0.0095947

drop1(poissonReg_reduced,test="Chi")

	Df	Deviance	AIC	LRT	Pr(>Chi)
	NA	430.4971	1164.497	NA	NA
$pclass_2$	1	436.9472	1168.947	6.450073	0.0110948
sex_M	1	540.1126	1272.113	109.615481	0.0000000
${\rm deck}_{\rm F}$	1	520.5287	1252.529	90.031585	0.0000000

```
### Train Data

# Testing against train data
predicted_probs <- predict(poissonReg_reduced, newdata = train, type = "response")

# Get the results of different cutoff values
trainResult<-cutoff.prg(predicted_probs,train$survived)</pre>
```

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to

match.

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

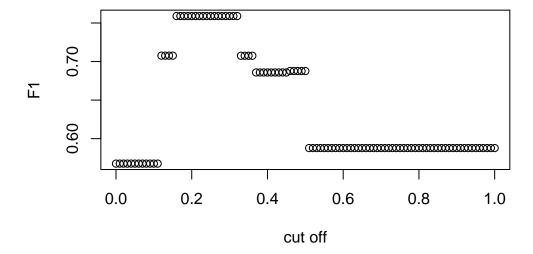
Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

We get get the optimal p cut off value based on the maximum F1 score
plot(trainResult[,1],trainResult[,8],xlab="cut off",ylab="F1")



```
poisson.model.cutoff = trainResult[which.max(trainResult[,8]),]

# Get the F1 and Accuracy for train data using the optimal p value
observations.train <- train$survived
predictions.poisson.model.train <- ifelse(predicted_probs > poisson.model.cutoff[1], 1, 0)

confusion.matrix.binary.model.poisson.train <- confusionMatrix(as.factor(predictions.poisson.model.train))</pre>
```

```
poisson.model.accuracy.train = confusion.matrix.binary.model.poisson.train$overall['Accuracy']
poisson.model.f1.train = confusion.matrix.binary.model.poisson.train$byClass['F1']
poisson.model.train.summary = data.frame(
 Accuracy = poisson.model.accuracy.train,
 F1 = poisson.model.f1.train
row.names(poisson.model.train.summary) <- 'poisson.model.train'
### Test Data
# Testing against test data
predicted_probs <- predict(poissonReg_reduced, newdata = test, type = "response")</pre>
# Get the F1 and Accuracy for test data using the optimal p value
observations.test <- test$survived
predictions.poisson.model.test <- ifelse(predicted_probs > poisson.model.cutoff[1], 1, 0)
confusion.matrix.binary.model.poisson.test <- confusionMatrix(as.factor(predictions.poisson.model.tes
poisson.model.accuracy.test = confusion.matrix.binary.model.poisson.test$overall['Accuracy']
poisson.model.f1.test = confusion.matrix.binary.model.poisson.test$byClass['F1']
poisson.model.test.summary = data.frame(
  Accuracy = poisson.model.accuracy.test,
 F1 = poisson.model.f1.test
row.names(poisson.model.test.summary) <- 'poisson.model.test'</pre>
data.frame(rbind(poisson.model.train.summary,poisson.model.test.summary ))
```

	Accuracy	F1
poisson.model.train poisson.model.test	0.7663755 0.7474490	$0.7590090 \\ 0.7195467$

8 Model Limitation and Assumptions

The champion model selected is the filtered logistic regression model due to its superior test accuracy (84.1%) and F1 score (0.78014). The filtered logistic regression model was developed through backward elimination procedure, applied to the full base model. This process iteratively removed predictors with high p-values until only statistically significant predictors remained. The base model is used as the benchmark due to its similar performance and broader feature set.

summary(binary.model.filtered)

Call: glm(formula = survived ~ age + sibsp + pclass_1 + pclass_2 + sex_M + embarked_C + embarked_Q + deck_A + deck_B + deck_C + deck_D + deck_F, family = binomial, data = train) Coefficients: Estimate Std. Error z value Pr(>|z|)(Intercept) 5.745434 0.507728 11.316 < 2e-16 *** age -0.053296 0.008919 -5.976 2.29e-09 *** -0.254771 0.116064 -2.1950.02816 * sibsp pclass_1 0.799195 0.449070 1.780 0.07513 . 1.827594 0.297584 6.141 8.18e-10 *** pclass_2 0.251332 -13.045 < 2e-16 *** sex_M -3.278525embarked_C 0.830036 0.282806 2.935 0.00334 ** 2.214 0.02684 * embarked_Q 0.785344 0.354744 -4.135 3.55e-05 *** $deck_A$ -2.5495790.616634 deck B -1.8968730.585234 -3.241 0.00119 ** deck C -2.032446 0.516558 -3.935 8.33e-05 *** deck_D -2.9959510.494319 -6.061 1.35e-09 *** -4.706551 0.404816 -11.626 < 2e-16 *** deck_F 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 Signif. codes: (Dispersion parameter for binomial family taken to be 1) Null deviance: 1230.1 on 915 degrees of freedom

Residual deviance: 635.0 on 903 degrees of freedom

AIC: 661

Number of Fisher Scoring iterations: 6

To further evaluate and benchmark the logistic regression models, we computed RMSE, R^2 , and MAE using the predicted probabilities against the actual binary outcomes. While traditional R^2 is not appropriate for logistic models, these metrics provide insight into how well the predicted probabilities align with observed outcomes. The Filtered model appears to have higher R^2 (0.4425>0.4342), lower RMSE(0.3945 < 0.3977) and MAE (0.1556 <0.1582). Additionally, model validation was done through train/test split with consistent performance across sets. The performance on test data and train data show minimal difference, indicating the model is robust, no over-itting problem.

To evaluate the stability and fit of the base and filtered logistic regression models, we conducted residual-based tests on their linear analogs. While logistic regression does not formally require normally distributed or homoscedastic residuals, these tests offer insights into model specification quality and residual behavior.

Breusch-Pagan Test: Both models exhibit statistically significant heteroskedasticity (p < 0.01)

Shapiro-Wilk Test: Both models show highly significant deviations from normality (p < 0.01)

Both models show similar residual behavior, with slight heteroskedasticity and non-normality. These findings do not violate logistic regression assumptions but suggest the model could be further improved through additional transformations or interaction terms.

Multicollinearity was significantly lower in the filtered model (max VIF: 3.69 vs. 16.63), supporting better generalization and interpretability.

```
library(car)

# Normality Test

# Breusch Pagan Test on Logistic Model

BPtest_basemodel <- ols_test_breusch_pagan(lmod)

BPtest_filtered.model <- ols_test_breusch_pagan(binary.model.filtered)

# Shapiro-Wilk test

SWtest_basemodel <- shapiro.test(residuals(lmod))

SWtest_filteredmodel <- shapiro.test(residuals(binary.model.filtered))

# Combine into unified rows

diagnostics_summary <- data.frame(
    Model = c("Base.model", "Filtered.model"),
    BP_Statistic = c(BPtest_basemodel$bp, BPtest_filtered.model$bp),
    BP_pvalue = c(BPtest_basemodel$p, BPtest_filtered.model$p),
    Shapiro_W = c(SWtest_basemodel$statistic, SWtest_filteredmodel$statistic),
    Shapiro_pvalue = c(SWtest_basemodel$p.value, SWtest_filteredmodel$p.value)
)</pre>
```

diagnostics_summary

Table 7: Summary of Diagnostic tests on Filtered vs Base models

Model	BP_Statistic	BP_pvalue	Shapiro_W	Shapiro_pvalue
Base.model	8.234753	0.0041096	0.9727993	0
Filtered.model	8.213280	0.0041585	0.9726405	0

```
# Run VIF on the logistic model
ols_vif_tol(lmod)
```

Table 8: VIF Results for Base Logistic Regression Model

Variables	Tolerance	VIF
age	0.8045607	1.242914
sibsp	0.8152516	1.226615
parch	0.7785122	1.284501
fare	0.5132230	1.948471
pclass_1	0.2345039	4.264322
pclass_2	0.7848041	1.274203
sex_M	0.8534444	1.171723
$embarked_C$	0.7524221	1.329041
embarked_Q	0.8549107	1.169713
deck_A	0.2520772	3.967038
$deck_B$	0.1505798	6.640995
deck_C	0.1062098	9.415324
deck_D	0.1509330	6.625458
$deck_E$	0.1122820	8.906149

Variables	Tolerance	VIF
deck_F	0.0522602	19.135037

ols_vif_tol(binary.model.filtered)

Table 9: VIF Results for Filtered Logistic Regression Model

Variables	Tolerance	VIF
age	0.8258163	1.210923
sibsp	0.9322261	1.072701
pclass_1	0.2639511	3.788581
$pclass_2$	0.7953422	1.257321
sex_M	0.9103965	1.098423
$embarked_C$	0.7737213	1.292455
$embarked_Q$	0.8651816	1.155827
$\operatorname{deck}_{-}A$	0.6538958	1.529295
$deck_B$	0.5054336	1.978499
$deck_C$	0.4299243	2.325991
$deck_D$	0.6025945	1.659491
deck_F	0.3615519	2.765855

```
# Base Model - Train
observations.train <- train$survived
y_hat_base_train <- predict(lmod, train, type = "response")</pre>
pred_base_train <- ifelse(y_hat_base_train > optimal.cutoff[1], 1, 0)
ModelTrain base <- data.frame(obs = observations.train, pred = pred_base_train)
log.train.base <- defaultSummary(ModelTrain_base)</pre>
# Base Model - Test
observations.test <- test$survived
y_hat_base_test <- predict(lmod, test, type = "response")</pre>
pred_base_test <- ifelse(y_hat_base_test > optimal.cutoff[1], 1, 0)
ModelTest_base <- data.frame(obs = observations.test, pred = pred_base_test)</pre>
log.test.base <- defaultSummary(ModelTest_base)</pre>
# Base Model - Train
observations.train <- train$survived
y_hat_base_train <- predict(binary.model.filtered, train, type = "response")</pre>
pred_base_train <- ifelse(y_hat_base_train > optimal.cutoff[1], 1, 0)
ModelTrain base <- data.frame(obs = observations.train, pred = pred_base_train)
filtered.log.train.base <- defaultSummary(ModelTrain_base)</pre>
# Base Model - Test
observations.test <- test$survived
y_hat_base_test <- predict(binary.model.filtered, test, type = "response")</pre>
pred_base_test <- ifelse(y_hat_base_test > optimal.cutoff[1], 1, 0)
ModelTest_base <- data.frame(obs = observations.test, pred = pred_base_test)</pre>
filtered.test.base <- defaultSummary(ModelTest_base)</pre>
```

	RMSE	Rsquared	MAE
log.train.base	0.3867342	0.4727588	0.1495633

0.4008919

0.3867342

0.3976975

0.4240145

0.4727588

0.4322221

0.1607143

0.1495633

0.1581633

Table 10: Comparison of stats between base and filtered log models:

9 Ongoing Model Monitoring Plan

log.test.base

filtered.log.train.base

filtered.test.base

In order to maintain the effectiveness of the model, we would need to continue to test it on new data. Since the Titanic was a rare event, we do not have a lot of new data to test on the model, but we can still be prepared in case new data were to become available. The first step in monitoring the model is to determine specific thresholds that we expect the model to stay above. We would want the model to maintain certain Accuracy and F1 values in order to determine that the model is working correctly. The first level we are monitoring is Accuracy or F1 values falling below 70%. Once one of those values falls below 70% we would examine the test data to make sure the data is valid. Assuming the data is valid, we will keep a closer eye on the model performance moving forward. Once both Accuracy and F1 values fall below 70%, or one of the values falls below 60%, we will retrain the model on more recent data.

One of the biggest concerns with our model is data drift. Since the Titanic sank over 100 years ago, the data that we are using from the model may not align with data relevant to ship travel today. Our expectation is that age is one data point that will look drastically different for ship travel today than it did for the Titanic simply due to the massive increase in life expectancy (verywellhealth, 2024). Due to this change, we will make sure to closely monitor age in new data and will adjust the model if we see the median age change by 20%.

Another concern of ours is that there will be concept drift. It is possible that over time, the relationships between variables in the model will change, resulting in the model losing effectiveness. In order to account for this, we will continue to monitor correlations between the model variables and will adjust the model if we notice any large (greater than 20%) changes in the correlations.

Overall, we will monitor our model to make sure it remains robust. Our assumptions are that a robust model will maintain both Accuracy and F1 values of over 70%, while the correlations in the model remain fairly constant and new data doesn't appear to vary drastically from our training data. We will look for the model to keep Accuracy and F1 values above 70%, while also making sure that median age of new data doesn't change by more than 20% and that correlations between variables in the model don't change by more than 20%. If any of these events were to happen, we would look to retrain our model on more recent data in order to maintain model robustness.

10 Conclusion

As we can see from the comparing all three models. The "filtered model" performed the best against the test data set. This is due to the removal insignificant variables until our model remained with only sigificant predictors. Multicollinarity seemed to be a big factor that impacted our models accuracy and F1. Once we removed indications multicollinarity, we could see that our models performance increased over the base model.

However, with the "poisson.model", we see a decrease in accuracy compared the base model. This decrease is expected due to the nature of the poisson regression. The regression is most useful when the response variable is

a count variable rather than a binary response. Thus a decrease in accuracy and F1 is expected against the base model.

As a conclusion, the model reduced via p-values filtering is the best model to predict if a passenger were to survive, with an accuracy of 84.28% and improved over the base model by 0.25%.

Table 11: Comparison of our base model, best performing developed model, and challenger model

	Accuracy	F1
base.model.test	0.8392857	0.7758007
filtered.model.test	0.8418367	0.7801418
poisson.model.test	0.7474490	0.7195467

Appendix A: Check if 'sibsp' and 'parch' should be continuous or categorical

We don't see significant improvement between modeling these predictors as continuous or categorical, therefore we decided to leave them as continuous.

```
library(car)
data.clean.ap1 = odata[, !(names(odata) %in% c("name", "ticket", "boat", "body", "home.dest"))]
Data Augmentation
#Extract deck letter from cabin
data.clean.ap1$deck <- substr(data.clean.ap1$cabin, 1,1)</pre>
# Remove cabin col:
data.clean.ap1$cabin <- NULL
Imputing data
# ---- Age----
#Replace NAs in age column with Median value
median_age <- median(data.clean.ap1$age, na.rm = TRUE)</pre>
data.clean.ap1 <- data.clean.ap1 %>%
 mutate(age = ifelse(is.na(age), median_age, age))
# ---- deck----
# For deck, since its a category, we decided to use KNN to impute the column:
# Install if not already installed
# install.packages("VIM")
library(VIM)
```

```
# Replace "" with NA in the 'deck' column
data.clean.ap1$deck[data.clean.ap1$deck == ""] <- NA
# Convert 'cabin' to factor
data.clean.ap1$deck <- as.factor(data.clean.ap1$deck)
# Apply kNN imputation just to Cabin column
data.clean.ap1 <- kNN(data.clean.ap1, variable = "deck", k = 5)
# Check that NAs were imputed
# sum(is.na(data.clean$deck))
                                  # Original
# sum(is.na(data.clean.imputed$deck)) # After
# Remove indicator col:
data.clean.ap1$deck_imp <- NULL
Dummify Cat. cols
# Dummifying pclass:
data.clean.ap1$pclass_1 = ifelse(data.clean.ap1$pclass == 1, 1, 0)
data.clean.ap1$pclass_2 = ifelse(data.clean.ap1$pclass == 2, 1, 0)
# Dummifying sex:
data.clean.ap1$sex_M = ifelse(data.clean.ap1$sex == 'male', 1, 0)
# Dummifying embarked:
data.clean.ap1$embarked_C = ifelse(data.clean.ap1$embarked == 'C', 1, 0)
data.clean.ap1$embarked_Q = ifelse(data.clean.ap1$embarked == 'Q', 1, 0)
# Dummifying deck:
data.clean.ap1$deck_A = ifelse(data.clean.ap1$deck == 'A', 1, 0)
data.clean.ap1$deck_B = ifelse(data.clean.ap1$deck == 'B', 1, 0)
data.clean.ap1$deck_C = ifelse(data.clean.ap1$deck == 'C', 1, 0)
data.clean.ap1$deck_D = ifelse(data.clean.ap1$deck == 'D', 1, 0)
data.clean.ap1$deck_E = ifelse(data.clean.ap1$deck == 'E', 1, 0)
data.clean.ap1$deck F = ifelse(data.clean.ap1$deck == 'F', 1, 0)
#data.clean.ap1$deck_G = ifelse(data.clean.ap1$deck == 'G', 1, 0) # removed due to causing issues
# Dummifying sibsp:
data.clean.ap1$sibsp_1 = ifelse(data.clean.ap1$sibsp == 1, 1, 0)
data.clean.ap1$sibsp_2 = ifelse(data.clean.ap1$sibsp == 2, 1, 0)
data.clean.ap1$sibsp_3 = ifelse(data.clean.ap1$sibsp == 3, 1, 0)
data.clean.ap1$sibsp_4 = ifelse(data.clean.ap1$sibsp == 4, 1, _0)
data.clean.ap1$sibsp_5 = ifelse(data.clean.ap1$sibsp == 5, 1, 0)
#data.clean.ap1$sibsp_8 = ifelse(data.clean.ap1$sibsp == 8, 1, 0) # removed due to causing issues
# Dummifying parch:
data.clean.ap1$parch_1 = ifelse(data.clean.ap1$parch == 1, 1, 0)
data.clean.ap1$parch_2 = ifelse(data.clean.ap1$parch == 2, 1, 0)
data.clean.ap1$parch_3 = ifelse(data.clean.ap1$parch == 3, 1, 0)
```

```
data.clean.ap1$parch_4 = ifelse(data.clean.ap1$parch == 4, 1, 0)
data.clean.ap1$parch_5 = ifelse(data.clean.ap1$parch == 5, 1, 0)
data.clean.ap1$parch_6 = ifelse(data.clean.ap1$parch == 6, 1, 0)
#data.clean.ap1$parch_9 = ifelse(data.clean.ap1$parch == 9, 1, 0) # removed due to causing issues
# Removing Dummified cols:
data.clean.ap1 = subset(data.clean.ap1, select = -c(pclass, sex, embarked, deck))#, sibsp, parch))
data.clean.ap1 = na.omit(data.clean.ap1)
cat(nrow(odata) - nrow(data.clean.ap1), 'rows were removed from original dataset')
2 rows were removed from original dataset
set.seed(567)
train_indices_ap1 = sample(1 : nrow(data.clean.ap1), size = 0.7005*nrow(data.clean.ap1), replace = Fi
train.ap1 = data.clean.ap1[train_indices_ap1,]
test.ap1 = data.clean.ap1[-train_indices_ap1,]
cat("We are using:", nrow(train.ap1)/nrow(data.clean.ap1) * 100, '% of the data for training')
We are using: 70.03058 % of the data for training
mulvar_model.ap1 <- lm(survived ~ ., data = train.ap1)</pre>
summary(mulvar_model.ap1)
Call:
lm(formula = survived ~ ., data = train.ap1)
Residuals:
                   Median
    Min
              1Q
                                3Q
                                        Max
-1.28431 -0.18993 -0.02983 0.19714 0.97447
Coefficients:
             Estimate Std. Error t value Pr(>|t|)
(Intercept) 1.1771844 0.1208120 9.744 < 2e-16 ***
           age
           -0.0387862 0.0222879 -1.740 0.08216 .
sibsp
           -0.0300588 0.0271365 -1.108 0.26830
parch
fare
           0.0003968 0.0002994 1.325 0.18550
pclass_1
           0.0811930 0.0556170 1.460 0.14468
           0.1305011 0.0315138 4.141 3.79e-05 ***
pclass_2
           -0.4220045 0.0253002 -16.680 < 2e-16 ***
sex_M
embarked_C 0.0477656 0.0322484
                                   1.481 0.13891
embarked_Q 0.1103169 0.0427122 2.583 0.00996 **
\mathtt{deck}_{\mathtt{A}}
           -0.3194255 0.1468546 -2.175 0.02988 *
deck_B
           -0.2768199 0.1381986 -2.003 0.04547 *
           -0.2682689 0.1346614 -1.992 0.04666 *
deck_C
           -0.3270119  0.1314397  -2.488  0.01303 *
deck_D
```

0.0136466 0.1251546 0.109 0.91320

deck_E

```
deck_F
            -0.5835976  0.1189681  -4.905  1.11e-06 ***
sibsp_1
            0.0752979 0.0356515
                                   2.112 0.03496 *
sibsp_2
           0.1073260 0.0801711 1.339 0.18101
sibsp_3
            -0.1027125 0.1199623 -0.856 0.39211
            -0.2242068 0.1301364 -1.723 0.08526 .
sibsp_4
sibsp_5
           -0.2237646 0.2026863 -1.104 0.26989
            0.1456157 0.0456425
parch_1
                                    3.190 0.00147 **
            0.1363899 0.0718664 1.898 0.05804 .
parch_2
parch_3
            0.2094703 0.1744398
                                    1.201 0.23014
parch_4
          -0.0890979 0.2045944 -0.435 0.66332
parch_5
             0.0600587 0.2394630
                                    0.251 0.80202
             0.0037453 0.2906192
                                  0.013 0.98972
parch_6
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.3413 on 889 degrees of freedom
Multiple R-squared: 0.5237,
                                Adjusted R-squared: 0.5098
F-statistic: 37.59 on 26 and 889 DF, p-value: < 2.2e-16
vif(mulvar_model.ap1)
                                             pclass 1
                                                        pclass_2
       age
                sibsp
                          parch
                                      fare
                                                                      sex_M
  1.461878
             3.469370
                       4.492844
                                   1.920452
                                             4.690456
                                                        1.313421
                                                                    1.156946
embarked_C embarked_Q
                          \mathtt{deck}_{-}\mathtt{A}
                                    deck_B
                                               \mathtt{deck}_{\mathtt{C}}
                                                          \mathtt{deck}_{\mathtt{D}}
                                                                     deck_E
             1.169222
                       5.198873
                                   8.907075 13.377776
  1.420534
                                                        9.336210 11.767550
    deck_F
             sibsp_1 sibsp_2 sibsp_3
                                            {\tt sibsp\_4}
                                                        sibsp_5
                                                                    parch_1
 26.496638
            1.868491 1.445820 1.463025
                                                        1.404482
                                            1.863122
                                                                    1.917338
   parch_2
             parch_3
                     parch_4 parch_5
                                            parch_6
  3.089996
             1.298950
                       1.431051
                                   1.471912
                                             1.446896
lmod.ap1 <- glm(as.factor(survived) ~ ., family = binomial, data = train.ap1)</pre>
summary(lmod.ap1)
Call:
glm(formula = as.factor(survived) ~ ., family = binomial, data = train.ap1)
Coefficients:
              Estimate Std. Error z value Pr(>|z|)
(Intercept) 5.266e+00 1.399e+00 3.765 0.000167 ***
age
            -5.266e-02 9.841e-03 -5.351 8.76e-08 ***
            -1.942e+00 1.285e+02 -0.015 0.987938
sibsp
            -1.671e+00 1.615e+02 -0.010 0.991745
parch
            7.583e-04 2.406e-03 0.315 0.752647
fare
           8.262e-01 4.862e-01 1.699 0.089253 .
pclass_1
            1.344e+00 3.064e-01
                                   4.385 1.16e-05 ***
pclass_2
sex_M
            -3.044e+00 2.476e-01 -12.296 < 2e-16 ***
embarked_C 6.262e-01 2.821e-01 2.220 0.026451 *
```

embarked_Q 1.147e+00 3.682e-01 3.116 0.001834 **

-2.578e+00 1.522e+00 -1.694 0.090192 .

-2.209e+00 1.496e+00 -1.477 0.139625

 $deck_A$

 $\mathtt{deck}_{\mathtt{B}}$

```
\mathtt{deck}_{\mathtt{C}}\mathtt{C}
            -2.266e+00 1.468e+00 -1.544 0.122706
            -2.772e+00 1.439e+00 -1.926 0.054095 .
deck D
            3.489e-01 1.422e+00 0.245 0.806183
deck_E
            -4.675e+00 1.374e+00 -3.403 0.000665 ***
deck_F
            2.202e+00 1.285e+02 0.017 0.986323
sibsp_1
sibsp_2
           4.142e+00 2.570e+02 0.016 0.987141
            4.166e+00 3.854e+02 0.011 0.991377
sibsp_3
            4.649e+00 5.139e+02 0.009 0.992783
sibsp_4
           -6.681e+00 1.233e+03 -0.005 0.995677
sibsp_5
            2.640e+00 1.615e+02 0.016 0.986959
parch_1
parch_2
             4.102e+00 3.230e+02 0.013 0.989869
             5.957e+00 4.845e+02 0.012 0.990190
parch_3
             6.123e+00 6.460e+02 0.009 0.992438
parch_4
             8.503e+00 8.076e+02 0.011 0.991599
parch_5
            -4.236e+00 1.760e+03 -0.002 0.998080
parch_6
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
    Null deviance: 1223.12 on 915 degrees of freedom
Residual deviance: 609.92 on 889 degrees of freedom
AIC: 663.92
Number of Fisher Scoring iterations: 15
vif(lmod.ap1)
         age
                    sibsp
                                  parch
                                                fare
                                                         pclass_1
                                                                       pclass_2
1.644869e+00 7.753684e+05 1.618468e+06 1.763642e+00 5.096422e+00 1.529833e+00
       {\tt sex\_M}
               embarked_C
                             embarked_Q
                                              \mathtt{deck}_{\mathtt{A}}
                                                           deck_B
                                                                         deck_C
1.439671e+00 1.505300e+00 1.370293e+00 1.035046e+01 1.551825e+01 2.526478e+01
      deck D
                   deck E
                                 deck F
                                             sibsp_1
                                                          sibsp_2
                                                                        sibsp 3
1.949011e+01 1.511758e+01 4.497328e+01 3.293441e+05 1.876474e+05 2.068774e+05
     sibsp_4
                  sibsp_5
                               parch_1
                                             parch_2
                                                          parch_3
                                                                        parch 4
1.995922e+05 1.372545e+00 3.184507e+05 6.909815e+05 2.223700e+05 2.249046e+05
     parch_5
                  parch_6
4.152894e+05 1.434758e+00
y_hat_mulvar_train.ap1<-predict(mulvar_model.ap1, data = train.ap1)</pre>
predictions_train.ap1 <- ifelse(y_hat_mulvar_train.ap1 > 0.5, 1, 0)
ModelTrain_mulvar.ap1<-data.frame(obs = train.ap1$survived, pred=predictions_train.ap1)</pre>
linear.train.ap1 <- defaultSummary(ModelTrain_mulvar.ap1)</pre>
y_hat_mulvar_test.ap1<-predict(mulvar_model.ap1, newdata = test.ap1)</pre>
predictions_test.ap1 <- ifelse(y_hat_mulvar_test.ap1 > 0.5, 1, 0)
ModelTest_mulvar.ap1<-data.frame(obs = test.ap1$survived, pred=predictions_test.ap1)
linear.test.ap1 <- defaultSummary(ModelTest_mulvar.ap1)</pre>
y_hat_log_train.ap1<-predict(lmod.ap1, data = train.ap1)</pre>
predictions_log_train.ap1 <- ifelse(y_hat_log_train.ap1 > 0.5, 1, 0)
```

```
ModelTrain_lmod.ap1<-data.frame(obs = train.ap1$survived, pred=predictions_log_train.ap1)
log.train.ap1 <- defaultSummary(ModelTrain_lmod.ap1)

y_hat_log_test.ap1<-predict(lmod.ap1, newdata = test.ap1)
predictions_log_test.ap1 <- ifelse(y_hat_log_test.ap1 > 0.5, 1, 0)
ModelTest_lmod.ap1<-data.frame(obs = test.ap1$survived, pred=predictions_log_test.ap1)
log.test.ap1 <- defaultSummary(ModelTest_lmod.ap1)

data.frame(rbind(linear.train.ap1,linear.test.ap1,log.train.ap1,log.test.ap1))</pre>
```

	RMSE	Rsquared	MAE
linear.train.ap1	0.3881430	0.4633288	0.1506550
linear.test.ap1	0.4040610	0.4221825	0.1632653
log.train.ap1	0.3951121	0.4460886	0.1561135
log.test.ap1	0.3912304	0.4439052	0.1530612

confusion_matrix_mulvar_train.ap1 <- confusionMatrix(as.factor(predictions_train.ap1), as.factor(train.ap1)
confusion_matrix_mulvar_train.ap1</pre>

Confusion Matrix and Statistics

Reference
Prediction 0 1
0 499 76
1 62 279

Accuracy : 0.8493

95% CI: (0.8245, 0.8719)

No Information Rate : 0.6124 P-Value [Acc > NIR] : <2e-16

Kappa : 0.6803

Mcnemar's Test P-Value: 0.2685

Precision: 0.8182 Recall: 0.7859 F1: 0.8017

Prevalence: 0.3876

Detection Rate : 0.3046 Detection Prevalence : 0.3723 Balanced Accuracy : 0.8377

'Positive' Class : 1

Confusion Matrix and Statistics

Reference

Prediction 0 1

0 215 32

1 32 113

Accuracy : 0.8367

95% CI: (0.7964, 0.8719)

No Information Rate : 0.6301 P-Value [Acc > NIR] : <2e-16

Kappa: 0.6498

Mcnemar's Test P-Value : 1

Precision: 0.7793 Recall: 0.7793

F1 : 0.7793

Prevalence: 0.3699

Detection Rate: 0.2883
Detection Prevalence: 0.3699

Balanced Accuracy: 0.8249

'Positive' Class : 1

confusion_matrix_log_train.ap1 <- confusionMatrix(as.factor(predictions_log_train.ap1), as.factor(trainconfusion_matrix_log_train.ap1)</pre>

Confusion Matrix and Statistics

Reference

Prediction 0 1

0 523 105

1 38 250

Accuracy : 0.8439

95% CI: (0.8187, 0.8668)

No Information Rate : 0.6124 P-Value [Acc > NIR] : < 2.2e-16

Kappa: 0.6593

Mcnemar's Test P-Value : 3.406e-08

Precision: 0.8681

Recall : 0.7042

F1: 0.7776

Prevalence: 0.3876

Detection Rate: 0.2729
Detection Prevalence: 0.3144

Balanced Accuracy: 0.8182

'Positive' Class: 1

confusion_matrix_log_test.ap1 <- confusionMatrix(as.factor(predictions_log_test.ap1), as.factor(test
confusion_matrix_log_test.ap1</pre>

Confusion Matrix and Statistics

Reference Prediction 0 1 0 227 40

1 20 105

Accuracy : 0.8469

95% CI: (0.8074, 0.8811)

No Information Rate : 0.6301 P-Value [Acc > NIR] : < 2e-16

Kappa: 0.662

Mcnemar's Test P-Value : 0.01417

Precision : 0.8400 Recall : 0.7241

F1: 0.7778
Prevalence: 0.3699

Detection Rate : 0.2679
Detection Prevalence : 0.3189
Balanced Accuracy : 0.8216

'Positive' Class : 1

```
#Replace NAs in age column with Median value
median_age <- median(data.clean.ap2$age, na.rm = TRUE)</pre>
data.clean.ap2 <- data.clean.ap2 %>%
 mutate(age = ifelse(is.na(age), median age, age))
# ---- deck----
# For deck, since its a category, we decided to use KNN to impute the column:
# Install if not already installed
# install.packages("VIM")
library(VIM)
# Replace "" with NA in the 'deck' column
data.clean.ap2$deck[data.clean.ap2$deck == ""] <- NA
# Convert 'cabin' to factor
data.clean.ap2$deck <- as.factor(data.clean.ap2$deck)
# Apply kNN imputation just to Cabin column
data.clean.ap2 <- kNN(data.clean.ap2, variable = "deck", k = 5)
# Check that NAs were imputed
# sum(is.na(data.clean$deck))
                                 # Original
# sum(is.na(data.clean.imputed$deck)) # After
# Remove indicator col:
data.clean.ap2$deck_imp <- NULL
Dummify Cat. cols
# Dummifying pclass:
data.clean.ap2$pclass_1 = ifelse(data.clean.ap2$pclass == 1, 1, 0)
data.clean.ap2$pclass_2 = ifelse(data.clean.ap2$pclass == 2, 1, 0)
# Dummifying sex:
data.clean.ap2$sex_M = ifelse(data.clean.ap2$sex == 'male', 1, 0)
# Dummifying embarked:
data.clean.ap2$embarked_C = ifelse(data.clean.ap2$embarked == 'C', 1, 0)
data.clean.ap2$embarked_Q = ifelse(data.clean.ap2$embarked == 'Q', 1, 0)
# Dummifying deck:
data.clean.ap2$deck_A = ifelse(data.clean.ap2$deck == 'A', 1, 0)
data.clean.ap2$deck_B = ifelse(data.clean.ap2$deck == 'B', 1, 0)
data.clean.ap2$deck_C = ifelse(data.clean.ap2$deck == 'C', 1, 0)
data.clean.ap2$deck_D = ifelse(data.clean.ap2$deck == 'D', 1, 0)
data.clean.ap2$deck_E = ifelse(data.clean.ap2$deck == 'E', 1, 0)
data.clean.ap2$deck_F = ifelse(data.clean.ap2$deck == 'F', 1, 0)
#data.clean.ap2$deck_G = ifelse(data.clean.ap2$deck == 'G', 1, 0) # removed due to causing issues
```

```
# Dummifying sibsp to 2 categories:
data.clean.ap2$sibsp_y = ifelse(data.clean.ap2$sibsp > 0, 1, 0)
# Dummifying parch to 2 categories:
data.clean.ap2$parch_y = ifelse(data.clean.ap2$parch > 0, 1, 0)
# Removing Dummified cols:
data.clean.ap2 = subset(data.clean.ap2, select = -c(pclass, sex, embarked, deck))#, sibsp, parch))
data.clean.ap2 = na.omit(data.clean.ap2)
cat(nrow(odata) - nrow(data.clean.ap2), 'rows were removed from original dataset')
2 rows were removed from original dataset
set.seed(567)
train_indices_ap2 = sample(1 : nrow(data.clean.ap2), size = 0.7005*nrow(data.clean.ap2), replace = FA
train.ap2 = data.clean.ap2[train_indices_ap2,]
test.ap2 = data.clean.ap2[-train_indices_ap2,]
cat("We are using:", nrow(train.ap2)/nrow(data.clean.ap2) * 100, '% of the data for training')
We are using: 70.03058 % of the data for training
mulvar_model.ap2 <- lm(survived ~ ., data = train.ap2)</pre>
summary(mulvar_model.ap2)
Call:
lm(formula = survived ~ ., data = train.ap2)
Residuals:
             1Q
                 Median
                             3Q
    Min
                                    Max
-1.27243 -0.19451 -0.02769 0.19202 0.96436
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 1.1720948 0.1212992 9.663 < 2e-16 ***
          age
          sibsp
          parch
          0.0003848 0.0003005 1.281 0.200641
fare
           0.1055453 0.0557446
pclass_1
                                1.893 0.058629 .
          0.1376903 0.0314471 4.378 1.34e-05 ***
pclass_2
sex_M
          -0.4270550 0.0253448 -16.850 < 2e-16 ***
embarked_C 0.0585880 0.0319987 1.831 0.067439 .
embarked_Q 0.1061446 0.0425937 2.492 0.012881 *
          -0.3315651 0.1464353 -2.264 0.023797 *
\mathtt{deck}_{\mathtt{A}}
          deck_B
          -0.2951467   0.1350558   -2.185   0.029120 *
deck C
deck_D
```

```
deck_E
            0.0078459 0.1256544 0.062 0.950226
           deck F
          sibsp_y
parch_y
           0.1573960 0.0479445 3.283 0.001067 **
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.3436 on 898 degrees of freedom
Multiple R-squared: 0.5124,
                              Adjusted R-squared: 0.5031
F-statistic: 55.5 on 17 and 898 DF, p-value: < 2.2e-16
vif(mulvar_model.ap2)
                                     fare pclass_1 pclass_2
      age
               sibsp
                          parch
                                                                   \mathtt{sex}_{\mathtt{M}}
                                  1.908310 4.649205
            2.359204
                       2.776007
                                                       1.290432
  1.380228
                                                                  1.145551
embarked_C embarked_Q deck_A
                                  deck_B
                                            \mathtt{deck}_{\mathtt{C}}\mathtt{C} \mathtt{deck}_{\mathtt{D}}\mathtt{D}
                                                                  deck E
  1.379990 \quad 1.147247 \quad 5.439805 \quad 8.740948 \quad 13.154666 \quad 9.148653 \quad 11.598396
   deck_F sibsp_y parch_y
 26.237760
            2.297316
                       3.213776
lmod.ap2 <- glm(as.factor(survived) ~ ., family = binomial, data = train.ap2)</pre>
summary(lmod.ap2)
Call:
glm(formula = as.factor(survived) ~ ., family = binomial, data = train.ap2)
Coefficients:
             Estimate Std. Error z value Pr(>|z|)
(Intercept) 5.1275081 1.3016899 3.939 8.18e-05 ***
           -0.0558992  0.0094958  -5.887  3.94e-09 ***
age
          -1.0266480 0.2640080 -3.889 0.000101 ***
sibsp
          -0.2791962 0.1939136 -1.440 0.149925
parch
           0.0007343 0.0024042 0.305 0.760042
fare
pclass_1 1.1421684 0.4862413 2.349 0.018825 *
           1.4015232  0.2992294  4.684  2.82e-06 ***
pclass_2
           -3.0275901 0.2424805 -12.486 < 2e-16 ***
sex_M
embarked_C 0.6732550 0.2814737 2.392 0.016762 *
embarked_Q 1.2284326 0.3679272 3.339 0.000841 ***
\mathtt{deck}_{\mathtt{A}}
          -2.5767916 1.4219365 -1.812 0.069960 .
          -2.2194317 1.4042525 -1.581 0.113991
\mathtt{deck}_{\mathtt{B}}
          -2.3552547 1.3682224 -1.721 0.085179 .
deck_C
deck D
          -2.9570245 1.3451465 -2.198 0.027928 *
           0.4470949 1.3202938 0.339 0.734886
deck E
deck F
          -4.4528239 1.2754924 -3.491 0.000481 ***
          1.3882952 0.3947966 3.516 0.000437 ***
sibsp_y
parch_y
           1.2304958 0.4199141 2.930 0.003386 **
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
```

Null deviance: 1223.12 on 915 degrees of freedom Residual deviance: 620.68 on 898 degrees of freedom

parch

fare

pclass_1

pclass_2

 ${\tt sex_M}$

AIC: 656.68

Number of Fisher Scoring iterations: 6

sibsp

```
vif(lmod.ap2)
```

age

```
1.619429
              3.641267
                          2.989704
                                      1.753639
                                                  5.122869
                                                              1.534060
                                                                          1.410059
embarked_C embarked_Q
                            deck_A
                                        \mathtt{deck}_{\mathtt{B}}
                                                    \mathtt{deck}_{\mathtt{C}}\mathtt{C}
                                                                deck_D
                                                                            deck_E
  1.487757
            1.363260
                          9.852461 13.562280 21.612617 16.823106 13.659175
    \mathtt{deck}_{\mathtt{F}}
               sibsp_y
                         parch_y
              3.482784
 39.454519
                          3.287406
y_hat_mulvar_train.ap2<-predict(mulvar_model.ap2, data = train.ap2)</pre>
predictions_train.ap2 <- ifelse(y_hat_mulvar_train.ap2 > 0.5, 1, 0)
ModelTrain_mulvar.ap2<-data.frame(obs = train.ap2$survived, pred=predictions_train.ap2)
linear.train.ap2 <- defaultSummary(ModelTrain_mulvar.ap2)</pre>
y_hat_mulvar_test.ap2<-predict(mulvar_model.ap2, newdata = test.ap2)</pre>
predictions_test.ap2 <- ifelse(y_hat_mulvar_test.ap2 > 0.5, 1, 0)
ModelTest_mulvar.ap2<-data.frame(obs = test.ap2$survived, pred=predictions_test.ap2)
linear.test.ap2 <- defaultSummary(ModelTest_mulvar.ap2)</pre>
y_hat_log_train.ap2<-predict(lmod.ap2, data = train.ap2)</pre>
predictions_log_train.ap2 <- ifelse(y_hat_log_train.ap2 > 0.5, 1, 0)
ModelTrain_lmod.ap2<-data.frame(obs = train.ap2$survived, pred=y_hat_log_train.ap2)
log.train.ap2 <- defaultSummary(ModelTrain_mulvar.ap2)</pre>
y_hat_log_test.ap2<-predict(lmod.ap2, newdata = test.ap2)</pre>
predictions_log_test.ap2 <- ifelse(y_hat_log_test.ap2 > 0.5, 1, 0)
ModelTest_lmod.ap2<-data.frame(obs = test.ap2$survived, pred=predictions_log_test.ap2)
log.test.ap2 <- defaultSummary(ModelTest_lmod.ap2)</pre>
data.frame(rbind(linear.train.ap2,linear.test.ap2,log.train.ap2,log.test.ap2))
```

	RMSE	Rsquared	MAE
linear.train.ap2	0.3909456	0.4570397	0.1528384
linear.test.ap2	0.4040610	0.4221825	0.1632653
log.train.ap2	0.3909456	0.4570397	0.1528384
log.test.ap2	0.4008919	0.4212961	0.1607143

confusion_matrix_mulvar_train.ap2 <- confusionMatrix(as.factor(predictions_train.ap2), as.factor(trainconfusion_matrix_mulvar_train.ap2)</pre>

Confusion Matrix and Statistics

Reference

Prediction 0 1

0 498 77

1 63 278

Accuracy : 0.8472

95% CI: (0.8222, 0.8699)

No Information Rate : 0.6124 P-Value [Acc > NIR] : <2e-16

Kappa: 0.6757

Mcnemar's Test P-Value: 0.2719

Precision: 0.8152

Recall : 0.7831

F1 : 0.7989

Prevalence: 0.3876

Detection Rate: 0.3035 Detection Prevalence: 0.3723

Balanced Accuracy: 0.8354

'Positive' Class : 1

confusion_matrix_mulvar_test.ap2 <- confusionMatrix(as.factor(predictions_test.ap2), as.factor(test.ap2)
confusion_matrix_mulvar_test.ap2</pre>

Confusion Matrix and Statistics

Reference

Prediction 0 1

0 215 32

1 32 113

Accuracy : 0.8367

95% CI: (0.7964, 0.8719)

No Information Rate : 0.6301 P-Value [Acc > NIR] : <2e-16

Kappa: 0.6498

Mcnemar's Test P-Value : 1

Precision: 0.7793

Recall: 0.7793

F1 : 0.7793

Prevalence: 0.3699

Detection Rate: 0.2883

Detection Prevalence: 0.3699

Balanced Accuracy: 0.8249

'Positive' Class : 1

confusion_matrix_log_train.ap2 <- confusionMatrix(as.factor(predictions_log_train.ap2), as.factor(trainconfusion_matrix_log_train.ap2</pre>

Confusion Matrix and Statistics

Reference

1 38 244

Prediction 0 1 0 523 111

Accuracy : 0.8373

95% CI: (0.8118, 0.8607)

No Information Rate : 0.6124 P-Value [Acc > NIR] : < 2.2e-16

Kappa : 0.6439

Mcnemar's Test P-Value: 3.669e-09

Precision: 0.8652 Recall: 0.6873 F1: 0.7661

Prevalence : 0.3876

Detection Rate : 0.2664

Detection Prevalence : 0.3079

Balanced Accuracy : 0.8098

'Positive' Class : 1

confusion_matrix_log_test.ap2 <- confusionMatrix(as.factor(predictions_log_test.ap2), as.factor(test
confusion_matrix_log_test.ap2</pre>

Confusion Matrix and Statistics

Reference

Prediction 0 1 0 227 43 1 20 102

Accuracy : 0.8393

95% CI : (0.7991, 0.8742)

No Information Rate : 0.6301 P-Value [Acc > NIR] : < 2.2e-16

Kappa : 0.6436

Mcnemar's Test P-Value: 0.005576

Precision : 0.8361 Recall : 0.7034

F1 : 0.7640

Prevalence: 0.3699
Detection Rate: 0.2602

Detection Prevalence : 0.3112
Balanced Accuracy : 0.8112

'Positive' Class : 1

References

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