HARVARD EXTENSION SCHOOL

EXT CSCI E-106 Model Data Class Group Project Template

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In this project, our aim is to classify the probability of a passenger surviving the Titanic crash of 1912. We used a variety of linear and non-linear models to deduce the most accurate model and provide long-term stability in our predictions.

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Classify whether a passenger on board the maiden voyage of the RMS Titanic in 1912 survived given their age, sex and class. Sample-Data-Titanic-Survival.csv to be used in the Final Project

Variable	Description
pclass	Passenger Class $(1 = 1st; 2 = 2nd; 3 = 3rd)$
survived	Survival $(0 = No; 1 = Yes)$
name	Name
sex	\mathbf{Sex}
age	${f Age}$
sibsp	# of siblings $/$ spouses aboard the Titanic
parch	# of parents / children aboard the Titanic
ticket	Ticket number
fare	Passenger fare
cabin	Cabin number
embarked	Port of Embarkation ($C = Cherbourg; Q =$
	Queenstown; $S = Southampton$)
boat	Lifeboat ID, if passenger survived
body	Body number (if passenger did not survive and
	body was recovered
home.dest	The intended home destination of the passenger

1 Instructions:

- 0. Join a team with your fellow students with appropriate size (Up to Nine Students total) If you have not group by the end of the week of April 11 you may present the project by yourself or I will randomly assign other stranded student to your group. I will let know the final groups in April 11.
- 1. Load and Review the dataset named "Titanic_Survival_Data.csv" 2. Create the train data set which contains 70% of the data and use set.seed (15). The remaining 30% will be your test data set.

- 3. Investigate the data and combine the level of categorical variables if needed and drop variables as needed. For example, you can drop id, Latitude, Longitude, etc.
- 4. Build appropriate model to predict the probability of survival.
- 5. Create scatter plots and a correlation matrix for the train data set. Interpret the possible relationship between the response.
- 6. Build the best models by using the appropriate selection method. Compare the performance of the best logistic linear models.
- 7. Make sure that model assumption(s) are checked for the final model. Apply remedy measures (transformation, etc.) that helps satisfy the assumptions.
- 8. Investigate unequal variances and multicollinearity.
- 9. Build an alternative to your model based on one of the following approaches as applicable to predict the probability of survival: logistic regression, classification Tree, NN, or SVM. Check the applicable model assumptions. Explore using a negative binomial regression and a Poisson regression.
- 10. Use the test data set to assess the model performances from above.
- 11. Based on the performances on both train and test data sets, determine your primary (champion) model and the other model which would be your benchmark model.
- 12. Create a model development document that describes the model following this template, input the name of the authors, Harvard IDs, the name of the Group, all of your code and calculations, etc..

Due Date: May 12 2025 1159 pm hours EST Notes No typographical errors, grammar mistakes, or misspelled words, use English language All tables need to be numbered and describe their content in the body of the document All figures/graphs need to be numbered and describe their content All results must be accurate and clearly explained for a casual reviewer to fully understand their purpose and impact Submit both the RMD markdown file and PDF with the sections with appropriate explanations. A more formal.

2 Executive Summary

This section will describe the model usage, your conclusions and any regulatory and internal requirements. In a real world scenario, this section is for senior management who do not need to know the details. They need to know high level (the purpose of the model, limitations of the model and any issues).

3 Introduction

This section needs to introduce the reader to the problem to be resolved, the purpose, and the scope of the statistical testing applied. What you are doing with your prediction? What is the purpose of the model? What methods were trained on the data, how large is the test sample, and how did you build the model?

The Titanic was a British-registered ship that set sail on its maiden voyage on April 10th, 1912 with 2,240 passengers and crew on board. On April 15th, 1912, the ship struck an iceberg, split in half, and sank to the bottom of the ocean (National Oceanic and Atmospheric Administration (NOAA), 2023). In this report, we are going to analyze the data in the Titanic.csv file and use it to determine the best model for predicting whether someone on board would live or die. By creating this model, we hope to understand what factors a passenger could have taken into account in order to reduce their risk of death during the trip. We cleaned the data and split into into a train/test split in order to properly train our models. We created simple linear models, multivariate linear models, logistic models (both binomial and poisson), a regression tree, and a neural network model. The train sample size was 916 data points (70.03%) and the test sample size was 392 data points (29.97%). We built the

models after examining the data and determining which predictor variables we thought would be most relevant for survival rate. Once we had our variables and training data, we created the models and examined the performance of the models on both training and testing data to determine if they were robust. We also examined if the model assumptions appeared to hold for each model.

4 Description of the data and quality

Based on the data cleaning we were able to only remove 2 rows from the data set. We used median imputation as well as KNN for various columns. We also dummified several categorical columns. We found that leaving sibsp and parch as continuous as opposed to categorical increased their contributions to the model performance Appendix A. Further, we also extracted the deck number and found that removing deck_G from the model increased its performance.

4.1 Loading the data

```
odata <- read.csv("../data/Titanic_Survival_Data.csv")
cat("Size of entire data set:", nrow(odata), "\n")</pre>
```

Size of entire data set: 1310

4.2 Removing un-needed columns

Name: Removing because names have no inference on survivival (inference)

ticket: Ticket No. will also likely not have an influence in survival

boat: This is highly correlated to the survival dependant variable since people who made it on a boat likely survived

body: This is highly correlated to the survival dependant variable since people who's body was recovered did not survive.

home.dest: The destination likely has nothing to do with the survival

```
data.clean = odata[, !(names(odata) %in% c("name", "ticket", "boat", "body", "home.dest"))]
```

4.3 Data Augmentation

We extracted the deck letter from the cabin since it could potentially correlate to the survival.

```
#Extract deck letter from cabin
data.clean$deck <- substr(data.clean$cabin, 1,1)
# Remove cabin col:
data.clean$cabin <- NULL</pre>
```

4.4 Initial Check for Missing values

We see that age and deck have the most amount of missing data, therefore we proceed to impute them.

print(plot_missing_barchart(data.clean))

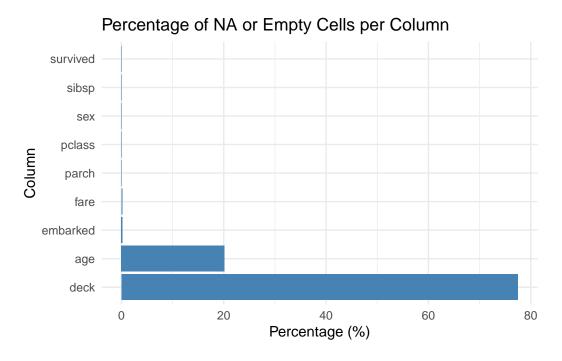


Figure 1: Percentage of Missing Values

4.5 Imputing data

Below we impute Age using the median value in that column.

For deck we use KNN to impute the missing deck values.

After imputing these two columns we can see that the largest amount of missing data is $\sim 0.2\%$ which is quite small and can be removed.

```
# ---- Age----
#Replace NAs in age column with Median value
median_age <- median(data.clean$age, na.rm = TRUE)
data.clean <- data.clean %>%
    mutate(age = ifelse(is.na(age), median_age, age))
# ---- deck----
# For deck, since its a category, we decided to use KNN to impute the column:
# Install if not already installed
# install.packages("VIM")
library(VIM)
```

Warning: package 'VIM' was built under R version 4.4.3

```
Loading required package: colorspace

Loading required package: grid

VIM is ready to use.

Suggestions and bug-reports can be submitted at: https://github.com/statistikat/VIM/issues

Attaching package: 'VIM'

The following object is masked from 'package:datasets':
    sleep

# Replace "" with NA in the 'deck' column
data.clean$deck[data.clean$deck == ""] <- NA

# Convert 'cabin' to factor
```

```
data.clean$deck[data.clean$deck == ""] <- NA

# Convert 'cabin' to factor
data.clean$deck <- as.factor(data.clean$deck)

# Apply kNN imputation just to Cabin column
data.clean <- kNN(data.clean, variable = "deck", k = 5)

# Check that NAs were imputed
# sum(is.na(data.clean$deck)) # Original
# sum(is.na(data.clean.imputed$deck)) # After

# Remove indicator col:
data.clean$deck_imp <- NULL</pre>
```

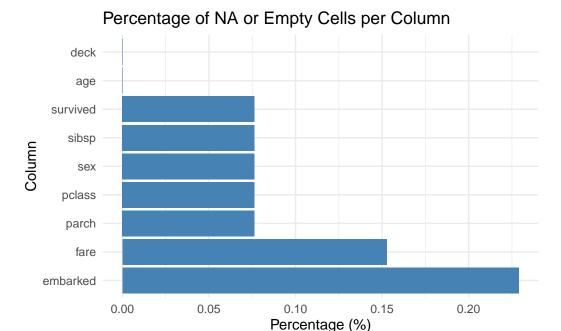


Figure 2: Percentage of Missing Values after Imputation

4.6 Dummifying Columns:

We dummify pclass, sex, embarked and deck. We leave sibsp and parch as continuous variables as we observed that dummifying these columns leads to smaller significance Appendix A, whilst leaving them as continuous maximizes their contributions to the models explanatory power.

```
# Dummifying pclass:
data.clean$pclass_1 = ifelse(data.clean$pclass == 1, 1, 0)
data.clean$pclass_2 = ifelse(data.clean$pclass == 2, 1, 0)
# Dummifying sex:
data.clean$sex_M = ifelse(data.clean$sex == 'male', 1, 0)
# Dummifying embarked:
data.clean$embarked_C = ifelse(data.clean$embarked == 'C', 1, 0)
data.clean$embarked_Q = ifelse(data.clean$embarked == 'Q', 1, 0)
# Dummifying deck:
data.clean$deck_A = ifelse(data.clean$deck == 'A', 1, 0)
data.clean$deck B = ifelse(data.clean$deck == 'B', 1, 0)
data.clean$deck_C = ifelse(data.clean$deck == 'C', 1, 0)
data.clean$deck_D = ifelse(data.clean$deck == 'D', 1, 0)
data.clean$deck_E = ifelse(data.clean$deck == 'E', 1, 0)
data.clean$deck_F = ifelse(data.clean$deck == 'F', 1, 0)
data.clean$deck_G = ifelse(data.clean$deck == 'G', 1, 0)
# Removing Dummified cols:
data.clean = subset(data.clean, select = -c(pclass, sex, embarked,deck))
```

4.7 Remove NA rows and deck_G

Below we remove NA rows, which turned out to be only 2 after proper cleaning and imputation. We also removed deck_G as we observed that it has a large skew in the data distribution with only 13 people allocated in this deck. It was observed that this variable lead to erroneous predictions in the model.

```
# Plot histogram of the 'values' column
hist(data.clean$deck_G,
    main = "Histogram of Values",
    xlab = "Values",
    col = "skyblue",
    border = "white")
```

Histogram of Values

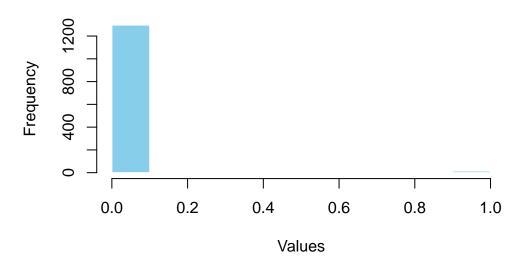


Figure 3: Histogram of Deck_G

```
# Removing deck_G col:
data.clean = subset(data.clean, select = -c(deck_G))

data.clean = na.omit(data.clean)
cat(nrow(odata) - nrow(data.clean), 'rows were removed from original dataset')
```

2 rows were removed from original dataset

4.8 Divide into Test / Train

Finally we divide into 70% training data and 30% test data.

```
set.seed (1023)
train_indices = sample(1 : nrow(data.clean), size = 0.7005*nrow(data.clean), replace = FALSE)
train = data.clean[train_indices,]
test = data.clean[-train_indices,]
cat("We are using:", nrow(train)/nrow(data.clean) * 100, '% of the data for training')
```

4.9 EDA

Using the training data set we use a variety of method to draw some initial conclusions:

- Histogram: Showing that more people in their late teens up to late thirties survived.
- Bar chart showing that more people died than survived
- Bar chart showing that a higher number of people survived when they had less siblings on board.
- Correlation matrix shows that sex and Deck_F are highly negatively correlated to survival. There is a soft positive correlation to pclass_1.
- There is a high correlation between pclass_1 and fare, this justifies that one of these predictors can potentially be removed.
- The scatter plots did not give us much more information on the relation between the predictors and the dependent variable.

```
# Histogram showing that more people in their late teens up to late thirties survived.
ggplot(train, aes(age)) +
geom_histogram(bins=30)
```

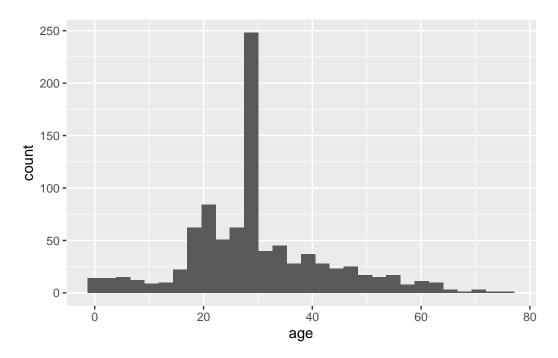


Figure 4: Histogram of survival vs age

```
# Bar chart showing that more people died than survived
ggplot(train, aes(survived)) +
  geom_bar()
```

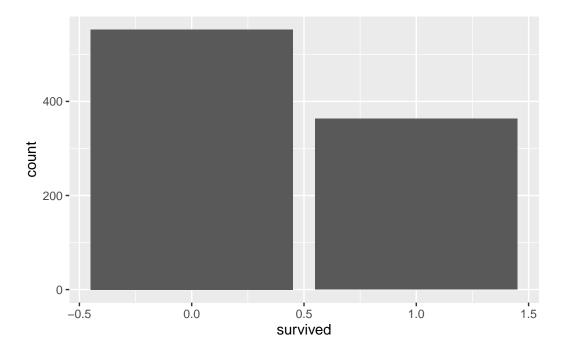


Figure 5: Barchart of survival

```
# Bar chart showing that a higher number of people survived when they had less
# siblings on board.
ggplot(train, aes(sibsp, survived)) +
  geom_bar(stat='identity')
```

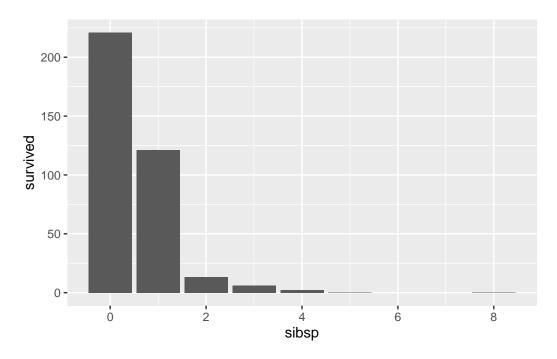


Figure 6: Barchart of survival vs Num. of siblings

cor(train)

survived age sibsp parch fare

```
survived
           1.00000000 - 0.04599216 - 0.027543645 0.095720424 0.249620417
          -0.04599216 1.00000000 -0.154952394 -0.122885529 0.164477215
age
          -0.02754365 -0.15495239 1.000000000 0.355216328 0.165296185
sibsp
parch
          0.09572042 -0.12288553  0.355216328  1.000000000  0.209878267
         0.24962042 0.16447722 0.165296185 0.209878267 1.000000000
fare
pclass_1 0.28432063 0.34847627 -0.022306615 -0.016451014 0.591163043
           pclass_2
          -0.53377358 0.05385130 -0.127066747 -0.243501842 -0.201776401
sex M
embarked_C 0.14757622 0.05701305 -0.061215242 -0.001165488 0.272087792
embarked_Q -0.02542950 -0.03373092 -0.060221066 -0.093253422 -0.126992379
deck_A
           0.02818626 \quad 0.12375210 \quad -0.063208812 \quad -0.046447535 \quad 0.060156186
deck_B
           0.17102793 0.11384754 -0.018056185 0.073894445 0.458426488
deck_C
         0.15622066 0.15548653 0.032030626 -0.054359774 0.314379308
deck_D
         0.11169173  0.07390772  -0.003562241  0.002702881  0.007239167
         0.32472414 0.11407881 -0.082487921 -0.030305432 -0.017794021
\mathtt{deck}_{\mathtt{E}}
deck_F
          -0.47931430 -0.28313314 0.052474332 0.010856728 -0.420674788
             pclass_1
                        pclass_2
                                      sex_M embarked_C embarked_Q
          survived
          0.34847627 0.01898825 0.05385130 0.057013047 -0.03373092
age
          -0.02230662 -0.06908924 -0.12706675 -0.061215242 -0.06022107
sibsp
          -0.01645101 -0.01769672 -0.24350184 -0.001165488 -0.09325342
parch
          0.59116304 -0.12563735 -0.20177640 0.272087792 -0.12699238
fare
          1.00000000 -0.30255006 -0.11706927 0.285611052 -0.16373859
pclass 1
pclass_2
          -0.30255006 1.00000000 -0.01057862 -0.143425262 -0.14105035
          -0.11706927 -0.01057862 1.00000000 -0.040151739 -0.09800681
sex_M
embarked_C 0.28561105 -0.14342526 -0.04015174 1.000000000 -0.16362864
embarked_Q -0.16373859 -0.14105035 -0.09800681 -0.163628636 1.00000000
deck_A
          0.27817486 -0.07214490 0.05726877 0.179563214 -0.06045425
deck_B
          0.42157158 -0.12290262 -0.10513911 0.197253957 -0.08030780
\mathtt{deck}_{\mathtt{C}}
         0.49783147 -0.09377853 -0.08157273 0.166128834 -0.07869333
deck_D
         deck E
          0.06047826 -0.03335428 -0.07748047 -0.085668330 -0.08697346
          deck F
               deck A
                          deck B
                                     deck C
                                                 deck D
                                                            deck E
          survived
          0.12375210 0.11384754 0.15548653 0.073907716 0.11407881
age
sibsp
          -0.06320881 -0.01805619 0.03203063 -0.003562241 -0.08248792
parch
          -0.04644753 0.07389444 -0.05435977 0.002702881 -0.03030543
fare
          0.06015619 0.45842649 0.31437931 0.007239167 -0.01779402
pclass_1
          0.27817486 0.42157158 0.49783147 0.116048985 0.06047826
          -0.07214490 \ -0.12290262 \ -0.09377853 \ \ 0.041176226 \ -0.03335428
pclass_2
           0.05726877 - 0.10513911 - 0.08157273 - 0.065993951 - 0.07748047
sex_M
embarked_C 0.17956321 0.19725396 0.16612883 0.288111512 -0.08566833
embarked_Q -0.06045425 -0.08030780 -0.07869333 -0.092805348 -0.08697346
\mathtt{deck}_{\mathtt{A}}
          1.00000000 -0.04808710 -0.06241531 -0.055570446 -0.06660453
deck_B
          -0.04808710 1.00000000 -0.08291288 -0.073820128 -0.08847788
deck C
          -0.06241531 -0.08291288 1.00000000 -0.095815832 -0.11484105
deck D
          -0.05557045 -0.07382013 -0.09581583 1.000000000 -0.10224685
          -0.06660453 -0.08847788 -0.11484105 -0.102246852 1.00000000
deck E
deck_F
           \hbox{-0.23590750 } \hbox{-0.31338100 } \hbox{-0.40675710 } \hbox{-0.362149521 } \hbox{-0.43405805} 
               deck F
survived
          -0.47931430
age
          -0.28313314
```

```
sibsp
            0.05247433
            0.01085673
parch
           -0.42067479
fare
pclass_1
           -0.70753043
            0.15770708
pclass_2
sex_M
            0.15681273
embarked_C -0.35310325
embarked Q 0.22529873
deck_A
           -0.23590750
deck_B
           -0.31338100
deck_C
           -0.40675710
deck_D
           -0.36214952
           -0.43405805
deck_E
            1.00000000
deck_F
```

```
pairs(train[c(1:4,7,13)])
```

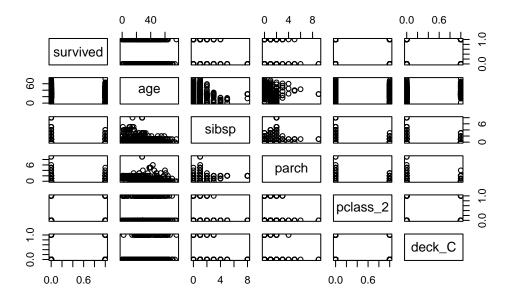


Figure 7: Scatter plots of all variables in train data

Since this data is mainly categorical, the scatterplot and correlation matrix are not very useful.

(Statology, 2025) is used to develop the correlation values between our categorical columns. This describes the use of pysch and recompanion.

```
#install.packages("psych")
library(psych) # [@statology2025] to understand how this works
```

Warning: package 'psych' was built under R version 4.4.3

Attaching package: 'psych'

```
The following objects are masked from 'package:ggplot2':
    %+%, alpha
tetrachoric(train[, c("survived", "sex_M")])
Call: tetrachoric(x = train[, c("survived", "sex_M")])
tetrachoric correlation
         srvvd sex_M
survived 1.00
        -0.75 1.00
sex_M
with tau of
survived sex M
   0.26 -0.35
tetrachoric(train[, c("survived", "pclass_1")])
Call: tetrachoric(x = train[, c("survived", "pclass_1")])
tetrachoric correlation
         srvvd pcl_1
survived 1.00
pclass_1 0.46 1.00
with tau of
survived pclass_1
    0.26
           0.69
tetrachoric(train[, c("survived", "pclass_2")])
Call: tetrachoric(x = train[, c("survived", "pclass_2")])
tetrachoric correlation
         srvvd pcl_2
survived 1.00
pclass_2 0.11 1.00
 with tau of
survived pclass_2
    0.26
           0.77
#install.packages("rcompanion")
library(rcompanion) # Reference 4 to understand how this works.
Warning: package 'rcompanion' was built under R version 4.4.3
Attaching package: 'rcompanion'
```

```
The following object is masked from 'package:psych':

phi
```

```
cramerV(train$survived, train$sex)
```

```
Cramer V 0.5338
```

```
library(corrplot)
```

corrplot 0.95 loaded

```
cor_matrix <- cor(train)#[,1]
corrplot(cor_matrix, method = "circle")</pre>
```

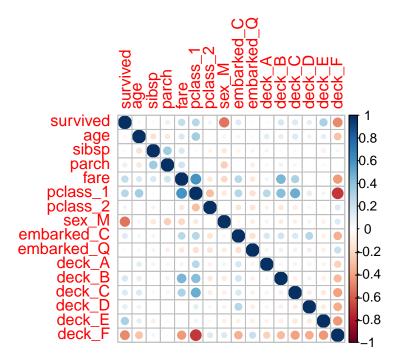


Figure 8: Correlation Matrix

5 Model Development Process

The data was properly cleaned and divided into train/test in the prior section.

Here we train a binary model. The Q-Q plot shows that the residuals are indeed normally distributed so a transformation is potentially not necessary.

The statistical comparison between test and train data shows that the model is very stable with an accuracy of $\sim 84\%$ for both.

We also analyzed the VIF and we see that there is high degree of correlation between the decks, this provides justification to remove some of the decks as predictors.

```
library(car)
```

```
Warning: package 'car' was built under R version 4.4.3
Loading required package: carData
Warning: package 'carData' was built under R version 4.4.3
Attaching package: 'car'
The following object is masked from 'package:psych':
    logit
The following object is masked from 'package:dplyr':
    recode
The following object is masked from 'package:purrr':
    some
# Log model on train data:
lmod <- glm(survived ~ ., family = binomial, data = train)</pre>
# summary(lmod)
vif(lmod)
                                                pclass_1
                                                            pclass_2
                 sibsp
                            parch
                                         fare
                                                                           sex_M
       age
  1.493399
             1.258340
                         1.281859
                                     1.727034
                                                 4.267734
                                                           1.560100
                                                                        1.604117
                                                 deck_C
embarked_C embarked_Q
                           deck_A
                                       \mathtt{deck}_{\mathtt{B}}
                                                             \mathtt{deck}_{\mathtt{D}}
                                                                          deck_E
  1.513511
             1.397623
                         4.532575
                                     4.848833
                                                 7.604802
                                                            6.662169
                                                                        6.091048
    deck_F
 14.919819
y_hat_log_train <- predict(lmod, data = train, type="response")</pre>
predictions_log_train <- ifelse(y_hat_log_train > 0.5, 1, 0)
y_hat_log_test<-predict(lmod, newdata = test, type="response")</pre>
predictions_log_test <- ifelse(y_hat_log_test > 0.5, 1, 0)
confusion_matrix_log_train <- confusionMatrix(as.factor(predictions_log_train), as.factor(train$surv
base.model.accuracy = confusion_matrix_log_train$overall['Accuracy']
base.model.f1 = confusion_matrix_log_train$byClass['F1']
base.model.train.summary = data.frame(
  Accuracy = base.model.accuracy,
  F1 = base.model.f1
row.names(base.model.train.summary) <- 'base.model.train'</pre>
```

```
confusion_matrix_log_test <- confusionMatrix(as.factor(predictions_log_test), as.factor(test$survived)
base.model.accuracy = confusion_matrix_log_test$overall['Accuracy']
base.model.f1 = confusion_matrix_log_test$byClass['F1']
base.model.test.summary = data.frame(
    Accuracy = base.model.accuracy,
    F1 = base.model.f1
)
row.names(base.model.test.summary) <- 'base.model.test'</pre>
```

data.frame(rbind(base.model.train.summary, base.model.test.summary))

Table 2: Summary Stats. of base log. model

	Accuracy	F1
base.model.train base.model.test	0.0 0 - 0	$0.8071625 \\ 0.7730496$

summary(lmod)

```
Call:
glm(formula = survived ~ ., family = binomial, data = train)
Coefficients:
              Estimate Std. Error z value Pr(>|z|)
(Intercept) 5.3284111 0.8287318 6.430 1.28e-10 ***
age
           -0.0557904 0.0090406 -6.171 6.78e-10 ***
           -0.2761926 0.1235912 -2.235 0.02544 *
sibsp
           0.0250325 0.1186751 0.211 0.83294
parch
           0.0001907 0.0023082 0.083 0.93415
fare
           0.8206018 0.4454290 1.842 0.06544 .
pclass_1
pclass_2
           1.8580626 0.2991637 6.211 5.27e-10 ***
           -3.2522858 0.2555063 -12.729 < 2e-16 ***
sex_M
embarked_C 0.6313152 0.2939442 2.148 0.03173 *
embarked_Q 0.8927717 0.3642358 2.451 0.01424 *
\mathtt{deck}_{\mathtt{A}}
           -1.9122914 0.9044640 -2.114 0.03449 *
deck_B
           -1.2596191 0.9199589 -1.369 0.17093
\mathtt{deck}_{\mathtt{C}}
           -1.6453843 0.8501358 -1.935 0.05294 .
           -2.3295283 0.8315684 -2.801 0.00509 **
deck_D
\mathtt{deck}_{\mathtt{E}}
           0.4168948 0.8021132
                                    0.520 0.60324
deck_F
           -4.3018644 0.7733641 -5.563 2.66e-08 ***
___
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
    Null deviance: 1230.15 on 915 degrees of freedom
```

Residual deviance: 635.57 on 900 degrees of freedom

AIC: 667.57

Number of Fisher Scoring iterations: 6

```
par(mfrow=c(2,2))
plot(lmod)
```

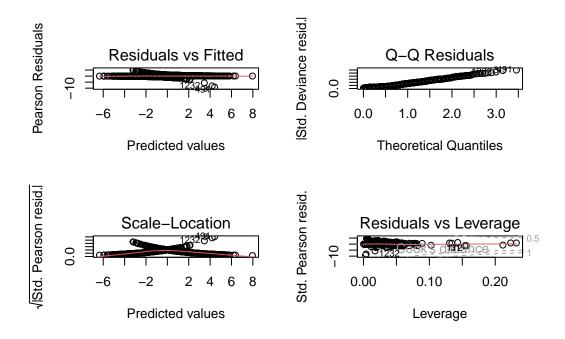


Figure 9: 4x4 standard plots for log. model

Plots show that a linear model is not appropriate for this data.

6 Model Performance Testing

We compare four different models: 1) Base log. model 2) Model with insig. pred. removed 3) Stepwise model 4) Model with high vif pred. removed

When comparing the accuracy and F1 score on all models the base model was still the highest performer and we decided to use that as the champion model until now.

```
# Stop if all predictors are significant or only intercept left
if( all(is.na(pvals)) || max(pvals, na.rm = TRUE) <= alpha ){
    print("all variable are significant")
    break
}

# Remove the term with max p-value
    term_to_remove = rownames(d1)[-1][which.max(pvals)]
    cat("Removing:", term_to_remove, "with p-value", max(pvals, na.rm = TRUE), "\n")
    model = update(model, paste(". ~ . -", term_to_remove))
}
return(model)
}</pre>
```

```
Function to remove Cooks Outliers
# model formula: A formula object, e.g. PSA.level ~ .
# data: A data frame containing the variables in the model.
# threshold: A numeric value indicating the Cook's D threshold (default 0.5).
# print: If TRUE (default) will print the rows beign removed.
# returns: A list with the final model and the filtered dataset.
# Example usage:
# result = remove_cooks_outliers(PSA.level ~ ., mydata)
# summary(result$model)
# str(result$filtered data)
remove_cooks_outliers = function(model_formula, data, threshold = 0.5,
                             print = TRUE)
 {
 all_high_cd_rows = data.frame() # to store all removed rows
 repeat {
   model = glm(model_formula, family = binomial, data = data)
   cooksD = cooks.distance(model)
   high_cd_indices = which(cooksD > threshold)
   if (length(high_cd_indices) == 0) { # If there are no more outliers
     break
   }
   if (print == TRUE){
     cat("Removing rows with Cook's D >", threshold, ":\n", high_cd_indices, "\n")
   }
   # Save these outliers before removing them
   high_cd_rows = data[high_cd_indices, ]
   all_high_cd_rows = rbind(all_high_cd_rows, high_cd_rows)
   # Update data by removing high CD rows for next iteration.
   data = data[-high_cd_indices, ]
 }
```

```
final_model = glm(model_formula, family = binomial, data = data)
 return(list(model = final_model, filtered_data = data, high_cd_data = all_high_cd_rows))
Model with insig. variables removed
binary.model.filtered = backward_eliminate(lmod)
Warning in drop1.glm(model, test = "F"): F test assumes 'quasibinomial' family
Removing: fare with p-value 0.9215538
Warning in drop1.glm(model, test = "F"): F test assumes 'quasibinomial' family
Removing: parch with p-value 0.7803802
Warning in drop1.glm(model, test = "F"): F test assumes 'quasibinomial' family
Removing: deck_E with p-value 0.5506678
Warning in drop1.glm(model, test = "F"): F test assumes 'quasibinomial' family
[1] "all variable are significant"
step wise model with insig. variables removed
library(olsrr)
Warning: package 'olsrr' was built under R version 4.4.3
Attaching package: 'olsrr'
The following object is masked from 'package:MASS':
   cement
The following object is masked from 'package:datasets':
   rivers
ols_step_both_p(lmod,p_enter=0.1,p_remove=0.05,details=FALSE)
```

Stepwise Summary

Step	Variable	AIC	SBC	SBIC	R2	Adj. R2
0	Base Model	1293.366	1303.006	-28281799.310	0.00000	0.00000
1	sex_M (+)	988.183	1002.643	-55309468.385	0.28491	0.28413
2	deck_F (+)	757.429	776.709	-91941771.510	0.44537	0.44415
3	age (+)	725.480	749.580	-99004330.024	0.46555	0.46379
4	pclass_2 (+)	696.998	725.918	-105806102.781	0.48304	0.48077
5	deck_E (+)	673.931	707.671	-111746149.031	0.49700	0.49423
6	sibsp (+)	667.234	705.794	-113874510.309	0.50175	0.49846
7	deck_D (+)	663.423	706.804	-115313992.417	0.50490	0.50108
8	embarked_Q (+)	661.553	709.753	-116277909.273	0.50699	0.50264
9	embarked_Q (-)	663.423	706.804	-115313992.417	0.50490	0.50108

Final Model Output

Model Summary

R	0.711	RMSE	0.344
R-Squared	0.505	MSE	0.118
Adj. R-Squared	0.501	Coef. Var	87.229
Pred R-Squared	0.495	AIC	663.423
MAE	0.261	SBC	706.804

RMSE: Root Mean Square Error

MSE: Mean Square Error MAE: Mean Absolute Error

AIC: Akaike Information Criteria SBC: Schwarz Bayesian Criteria

ANOVA

	Sum of Squares	DF	Mean Square	F	Sig.
Regression Residual Total	110.647 108.500 219.147	7 908 915	15.807 0.119	132.282	0.0000

Parameter Estimates

model	Beta	Std. Error	Std. Beta	t	Sig	lower	upper
(Intercept)	1.128	0.044		25.918	0.000	1.043	1.213
sex_M	-0.465	0.024	-0.458	-19.077	0.000	-0.513	-0.418
deck_F	-0.450	0.031	-0.450	-14.496	0.000	-0.511	-0.389
age	-0.006	0.001	-0.170	-6.842	0.000	-0.008	-0.005
pclass_2	0.156	0.028	0.132	5.491	0.000	0.100	0.211

```
0.105 3.842 0.000 0.081
-0.071 -2.957 0.003 -0.054
     sibsp -0.032
                        0.011
                                               -2.957 0.003 -0.054
                                                                        -0.011
                                    -0.064
                                             -2.404 0.016 -0.211
    deck_D -0.116
                         0.048
                                                                         -0.021
# Fit model with stepwise parms only:
binary.model.stepwise = glm(survived ~ sex_M + deck_F + age + deck_E +
                          pclass_2 + pclass_1 + sibsp + fare,
                         family = binomial, data = train)
summary(binary.model.stepwise)
Call:
glm(formula = survived ~ sex_M + deck_F + age + deck_E + pclass_2 +
   pclass_1 + sibsp + fare, family = binomial, data = train)
Coefficients:
           Estimate Std. Error z value Pr(>|z|)
(Intercept) 4.027276 0.443292 9.085 < 2e-16 ***
sex_M
         -3.228025 0.242958 -13.286 < 2e-16 ***
deck_F
         -2.649341 0.335751 -7.891 3.00e-15 ***
         age
deck_E
         1.823622   0.376762   4.840   1.30e-06 ***
pclass_2
pclass_1
          0.531123 0.376285 1.411 0.15810
       sibsp
          0.001937 0.002089 0.927 0.35379
fare
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
   Null deviance: 1230.15 on 915 degrees of freedom
Residual deviance: 652.91 on 907 degrees of freedom
AIC: 670.91
Number of Fisher Scoring iterations: 6
# Removing fare from stepwise model since its insig.:
binary.model.stepwise = update(binary.model.stepwise, paste(". ~ . -", 'fare'))
summary(binary.model.stepwise)
Call:
glm(formula = survived ~ sex_M + deck_F + age + deck_E + pclass_2 +
   pclass_1 + sibsp, family = binomial, data = train)
Coefficients:
           Estimate Std. Error z value Pr(>|z|)
(Intercept) 4.075133 0.441121 9.238 < 2e-16 ***
         -3.250356 0.241996 -13.431 < 2e-16 ***
sex_M
```

0.249

deck_E 0.165

0.043

```
deck_F
         -2.671528
                   0.335679 -7.959 1.74e-15 ***
                   0.008763 -6.266 3.70e-10 ***
         -0.054914
age
         1.802546  0.376305  4.790  1.67e-06 ***
deck_E
pclass_2
          1.497212  0.269700  5.551  2.83e-08 ***
         0.658065 0.351293 1.873 0.06103 .
pclass_1
         -0.305711
                   0.113440 -2.695 0.00704 **
sibsp
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
   Null deviance: 1230.15 on 915 degrees of freedom
Residual deviance: 653.81 on 908 degrees of freedom
AIC: 669.81
Number of Fisher Scoring iterations: 6
Model with influential points removed
# Applying function to remove influential points via Cooks Distance:
filtering.result = remove_cooks_outliers(survived ~ ., data = train)
cat(nrow(filtering.result$high_cd_data),'rows were identified as outliers')
O rows were identified as outliers
# plot(logmod)
# Since there were now rows identified as outliers, then this model will be the
# same as the initial binary model and need no be considered in the final
# model compare.
Model with high VIF preds removed
library(car)
vif(lmod)
             sibsp
                      parch
                                fare
                                     pclass_1 pclass_2
                                                          sex_M
     age
                                     4.267734 1.560100
 1.493399
          1.258340 1.281859
                            1.727034
                                                        1.604117
embarked_C embarked_Q
                              deck B
                                      \mathtt{deck}_{\mathsf{C}}
                                                deck_D
                                                        deck_E
                     deck_A
 1.513511
          1.397623 4.532575 4.848833 7.604802 6.662169 6.091048
   deck F
14.919819
vif.model = glm(survived ~ . -deck_F, family = binomial, data = train)
vif(vif.model)
     age
             sibsp
                      parch
                                fare
                                     pclass_1 pclass_2
                                                          sex_M
 1.481259
          1.236486
                   1.275733
                            1.729266
                                     4.369376
                                              1.472982
                                                        1.452735
embarked C embarked Q
                              deck B
                                      deck C
                                               deck D
                                                         deck E
                     deck A
 1.511473
          1.348123
                  1.988884
                            2.269651
                                    2.863440
                                              1.737448
                                                        1.558559
```

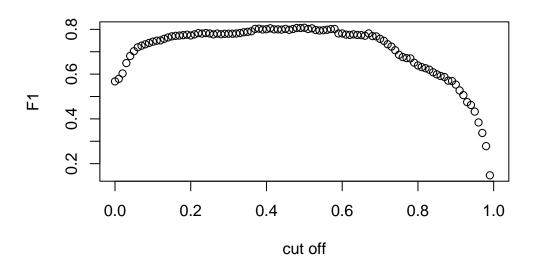
Removing deck_F from the model eliminates the multicollinearity completely.

```
#
                                                                                        Function to calc. cutoff
cutoff.prg<-function(pred,act){</pre>
# pred<-predicted_probabilities</pre>
# act<-true_labels
p < -seq(0,1,0.01)
n<-length(p)
out<-matrix(0,nrow=n,ncol=12)</pre>
for(i in 1:n){
predictions <- ifelse(pred >p[i], 1, 0)
confusion_matrix <- confusionMatrix(as.factor(predictions),as.factor(act),mode="prec_recall", positive transfer of the confusion of the confus
out[i,]<-cbind(p=p[i],t(confusion_matrix[[4]]))</pre>
}
dimnames(out)[[2]] <-c("p", "Sensitivity", "Specificity", "Pos Pred Value", "Neg Pred Value", "Precision",
out
}
# Finding the optimal cutoff
observations = train$survived
prob <- predict(lmod, train, type="response")</pre>
test_cutoff<-cutoff.prg(prob,observations)</pre>
```

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

```
plot(test_cutoff[,1],test_cutoff[,8],xlab="cut off",ylab="F1")
```



```
optimal.cutoff = test_cutoff[which.max(test_cutoff[,8]),]
optimal.cutoff[1]
```

p 0.5

```
observations.train = train$survived
observations.test = test$survived
# Confusion matrix on base log model on train data
y_hat_prob = predict(lmod, train, type="response")
predictions.binary.model.train <- ifelse(y_hat_prob > optimal.cutoff[1], 1, 0)
confusion.matrix.binary.model.train <- confusionMatrix(as.factor(predictions.binary.model.train),as.</pre>
base.model.accuracy = confusion.matrix.binary.model.train$overall['Accuracy']
base.model.f1 = confusion.matrix.binary.model.train$byClass['F1']
base.model.train.summary = data.frame(
  Accuracy = base.model.accuracy,
 F1 = base.model.f1
row.names(base.model.train.summary) <- 'base.model.train'</pre>
# Confusion matrix on base log model on test data
y_hat_prob = predict(lmod, test, type="response")
predictions.binary.model.test <- ifelse(y_hat_prob > optimal.cutoff[1], 1, 0)
confusion.matrix.binary.model.test <- confusionMatrix(as.factor(predictions.binary.model.test),as.fac
base.model.accuracy.test = confusion.matrix.binary.model.test$overall['Accuracy']
base.model.f1.test = confusion.matrix.binary.model.test$byClass['F1']
base.model.test.summary = data.frame(
```

```
Accuracy = base.model.accuracy.test,
  F1 = base.model.f1.test
row.names(base.model.test.summary) <- 'base.model.test'</pre>
# Confusion matrix on stepwise log model on train data
y_hat_prob = predict(binary.model.stepwise, train, type="response")
predictions.binary.model.step.train <- ifelse(y_hat_prob > optimal.cutoff[1], 1, 0)
confusion.matrix.binary.model.step.train <- confusionMatrix(as.factor(predictions.binary.model.step.
stepwise.model.accuracy = confusion.matrix.binary.model.step.train$overall['Accuracy']
stepwise.model.f1 = confusion.matrix.binary.model.step.train$byClass['F1']
stepwise.model.train.summary = data.frame(
  Accuracy = stepwise.model.accuracy,
  F1 = stepwise.model.f1
row.names(stepwise.model.train.summary) <- 'stepwise.model.train'
# Confusion matrix on stepwise log model on test data
y_hat_prob = predict(binary.model.stepwise, test, type="response")
predictions.binary.model.step.test <- ifelse(y_hat_prob > optimal.cutoff[1], 1, 0)
confusion.matrix.binary.model.step.test <- confusionMatrix(as.factor(predictions.binary.model.step.te
stepwise.model.accuracy = confusion.matrix.binary.model.step.test$overall['Accuracy']
stepwise.model.f1 = confusion.matrix.binary.model.step.test$byClass['F1']
stepwise.model.test.summary = data.frame(
  Accuracy = stepwise.model.accuracy,
  F1 = stepwise.model.f1
row.names(stepwise.model.test.summary) <- 'stepwise.model.test'</pre>
# Confusion matrix on log model w/ insig. pred. removed
y_hat_prob = predict(binary.model.filtered, train, type="response")
predictions.binary.model.filtered.train <- ifelse(y_hat_prob > optimal.cutoff[1], 1, 0)
confusion.matrix.binary.model.filtered.train <- confusionMatrix(as.factor(predictions.binary.model.fi
filtered.model.accuracy = confusion.matrix.binary.model.filtered.train$overall['Accuracy']
filtered.model.f1 = confusion.matrix.binary.model.filtered.train$byClass['F1']
filtered.model.train.summary = data.frame(
  Accuracy = filtered.model.accuracy,
  F1 = filtered.model.f1
row.names(filtered.model.train.summary) <- 'filtered.model.train'</pre>
# Confusion matrix on log model w/ insig. pred. removed
y_hat_prob = predict(binary.model.filtered, test, type="response")
predictions.binary.model.filtered.test <- ifelse(y_hat_prob > optimal.cutoff[1], 1, 0)
confusion.matrix.binary.model.filtered.test <- confusionMatrix(as.factor(predictions.binary.model.fil
filtered.model.accuracy = confusion.matrix.binary.model.filtered.test$overall['Accuracy']
filtered.model.f1 = confusion.matrix.binary.model.filtered.test$byClass['F1']
filtered.model.test.summary = data.frame(
```

```
Accuracy = filtered.model.accuracy,
 F1 = filtered.model.f1
row.names(filtered.model.test.summary) <- 'filtered.model.test'
# Confusion matrix on log model w/ high vif pred. removed
y_hat_prob = predict(vif.model, train, type="response")
predictions.binary.model.vif.train <- ifelse(y_hat_prob > optimal.cutoff[1], 1, 0)
confusion.matrix.binary.model.vif.train <- confusionMatrix(as.factor(predictions.binary.model.vif.train
vif.model.accuracy = confusion.matrix.binary.model.vif.train$overall['Accuracy']
vif.model.f1 = confusion.matrix.binary.model.vif.train$byClass['F1']
vif.model.train.summary = data.frame(
  Accuracy = vif.model.accuracy,
  F1 = vif.model.f1
row.names(vif.model.train.summary) <- 'vif.model.train'</pre>
# Confusion matrix on log model w/ high vif pred. removed
y_hat_prob = predict(vif.model, test, type="response")
predictions.binary.model.vif.test <- ifelse(y_hat_prob > optimal.cutoff[1], 1, 0)
confusion.matrix.binary.model.vif.test <- confusionMatrix(as.factor(predictions.binary.model.vif.test
vif.model.accuracy = confusion.matrix.binary.model.vif.test$overall['Accuracy']
vif.model.f1 = confusion.matrix.binary.model.vif.test$byClass['F1']
vif.model.test.summary = data.frame(
  Accuracy = vif.model.accuracy,
 F1 = vif.model.f1
row.names(vif.model.test.summary) <- 'vif.model.test'</pre>
data.frame(rbind(base.model.train.summary, base.model.test.summary,
                 stepwise.model.train.summary, stepwise.model.test.summary,
                 filtered.model.train.summary, filtered.model.test.summary,
                 vif.model.train.summary,vif.model.test.summary))
```

Table 3: Comparison of all log models performance

	Accuracy	F1
base.model.train	0.8471616	0.8071625
base.model.test	0.8367347	0.7730496
stepwise.model.train	0.8220524	0.7687943
stepwise.model.test	0.8418367	0.7669173
filtered.model.train	0.8471616	0.8066298
filtered.model.test	0.8367347	0.7730496
vif.model.train	0.8329694	0.7877947
vif.model.test	0.8367347	0.7746479

7 Challenger Models

Build an alternative model based on one of the following approaches to predict survival as applicable:logistic regression, decision tree, NN, or SVM, Poisson regression or negative binomial. Check the applicable model

assumptions. Apply in-sample and out-of-sample testing, back testing and review the comparative goodness of fit of the candidate models. Describe step by step your procedure to get to the best model and why you believe it is fit for purpose.

```
library(MASS)
# Train base Poisson model:
poissonReg <- glm(survived ~ .,family=poisson, train)</pre>
summary(poissonReg)
Call:
glm(formula = survived ~ ., family = poisson, data = train)
Coefficients:
             Estimate Std. Error z value Pr(>|z|)
(Intercept) 7.519e-01 3.472e-01 2.166 0.030341 *
           -1.668e-02 4.337e-03 -3.847 0.000119 ***
age
sibsp
           -7.014e-02 6.994e-02 -1.003 0.315889
           1.132e-02 6.704e-02 0.169 0.865869
parch
           9.688e-05 1.025e-03 0.095 0.924696
fare
            3.485e-01 1.987e-01 1.754 0.079421 .
pclass_1
           5.468e-01 1.515e-01 3.609 0.000307 ***
pclass_2
sex_M
           -1.088e+00 1.199e-01 -9.069 < 2e-16 ***
embarked_C 7.474e-02 1.362e-01 0.549 0.583046
           5.098e-01 2.183e-01 2.335 0.019548 *
embarked_Q
deck_A
           -5.113e-01 4.521e-01 -1.131 0.258009
           -5.677e-01 4.064e-01 -1.397 0.162402
deck_B
deck_C
           -6.456e-01 3.912e-01 -1.650 0.098877 .
           -6.580e-01 3.820e-01 -1.722 0.084985 .
deck D
deck_E
           -1.870e-01 3.526e-01 -0.530 0.595781
deck_F
           -1.540e+00 3.413e-01 -4.513 6.38e-06 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for poisson family taken to be 1)
    Null deviance: 672.00 on 915 degrees of freedom
Residual deviance: 400.84 on 900 degrees of freedom
AIC: 1158.8
Number of Fisher Scoring iterations: 5
# compare the fitted model to the null model
1-pchisq(672.34-390.72, length(coef(poissonReg)))
[1] 0
```

	Df	Deviance	Resid. Df	Resid. Dev	Pr(>Chi)
NULL	NA	NA	915	671.9954	NA
age	1	1.1812909	914	670.8141	0.2770933
sibsp	1	0.7166530	913	670.0975	0.3972440
parch	1	5.6694845	912	664.4280	0.0172625
fare	1	25.6933009	911	638.7347	0.0000004
$pclass_1$	1	27.9481171	910	610.7866	0.0000001
$pclass_2$	1	17.1639017	909	593.6227	0.0000343
sex_M	1	118.6444922	908	474.9782	0.0000000
$embarked_C$	1	1.6665466	907	473.3116	0.1967217
$embarked_Q$	1	0.0858234	906	473.2258	0.7695556
$\operatorname{deck}_{-}A$	1	0.0631207	905	473.1627	0.8016298
$deck_B$	1	0.0077426	904	473.1549	0.9298829
$deck_C$	1	0.0326392	903	473.1223	0.8566319
$deck_D$	1	0.5268888	902	472.5954	0.4679178
$deck_E$	1	57.6272332	901	414.9682	0.0000000
deck_F	1	14.1318730	900	400.8363	0.0001704

drop1(poissonReg,test="Chi")

	Df	Deviance	AIC	LRT	Pr(>Chi)
	NA	400.8363	1158.836	NA	NA
age	1	416.0444	1172.044	15.2080992	0.0000963
sibsp	1	401.8958	1157.896	1.0594618	0.3033380
parch	1	400.8646	1156.865	0.0282941	0.8664193
fare	1	400.8452	1156.845	0.0088926	0.9248705
pclass_1	1	403.9035	1159.903	3.0671630	0.0798888
$pclass_2$	1	413.5888	1169.589	12.7524939	0.0003555
sex_M	1	489.9790	1245.979	89.1427013	0.0000000
$embarked_C$	1	401.1362	1157.136	0.2998976	0.5839466
$embarked_Q$	1	405.8500	1161.850	5.0137077	0.0251474
$\operatorname{deck}_{-}A$	1	402.0796	1158.080	1.2432486	0.2648459
$deck_B$	1	402.6560	1158.656	1.8197039	0.1773485
$\operatorname{deck}_{-}C$	1	403.3130	1159.313	2.4766883	0.1155453
$deck_D$	1	403.5115	1159.511	2.6751598	0.1019251
$deck_E$	1	401.1067	1157.107	0.2703827	0.6030752
$\operatorname{deck}_{-}F$	1	414.9682	1170.968	14.1318730	0.0001704

```
# Testing against test data
predicted_probs <- predict(poissonReg, newdata = test, type = "response")
# Get the results of different cutoff values
testResult<-cutoff.prg(predicted_probs,test$survived)</pre>
```

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to

match.

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

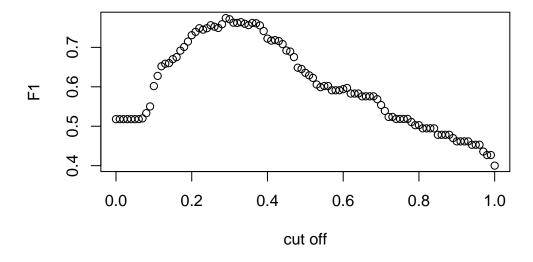
Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

Warning in confusionMatrix.default(as.factor(predictions), as.factor(act), : Levels are not in the same order for reference and data. Refactoring data to match.

We get get the optimal p cut off value based on the maximum F1 score
plot(testResult[,1],testResult[,8],xlab="cut off",ylab="F1")



testResult[which.max(testResult[,8]),]

p	Sensitivity	Specificity
0.2900000	0.8759124	0.7921569
Pos Pred Value	Neg Pred Value	Precision
0.6936416	0.9223744	0.6936416
Recall	F1	Prevalence
0.8759124	0.7741935	0.3494898
Detection Rate	Detection Prevalence	Balanced Accuracy
0.3061224	0.4413265	0.8340346

8 Model Limitation and Assumptions

The champion model selected is the filtered logistic regression model due to its superior train accuracy (85.04%) and F1 score (0.8099). The filtered logistic regression model was developed through backward elimination procedure, applied to the full base model. This process iteratively removed predictors with high p-values until only statistically significant predictors remained, based on threshold (p <0.05). The base model is used as the benchmark due to its similar performance and broader feature set.

```
\label{eq:Survived} \begin{subarray}{l} \beg
```

To further evaluate and benchmark the logistic regression models, we computed RMSE, R^2 , and MAE using the predicted probabilities against the actual binary outcomes. While traditional R^2 is not appropriate for logistic models, these metrics provide insight into how well the predicted probabilities align with observed outcomes. The Filtered model appears to have higher R^2 (0.4425>0.4342), lower RMSE(0.3945 < 0.3977) and MAE (0.1556 <0.1582). Additionally, model validation was done through train/test split with consistent performance across sets. The performance on test data and train data show minimal difference, indicating the model is robust, no overfitting problem.

To evaluate the stability and fit of the base and filtered logistic regression models, we conducted residual-based tests on their linear analogs. While logistic regression does not formally require normally distributed or homoscedastic residuals, these tests offer insights into model specification quality and residual behavior.

Breusch-Pagan Test: Both models exhibit statistically significant heteroskedasticity (p < 0.01)

Shapiro-Wilk Test: Both models show highly significant deviations from normality (p < 0.01)

Both models show similar residual behavior, with slight heteroskedasticity and non-normality. These findings do not violate logistic regression assumptions but suggest the model could be further improved through additional transformations or interaction terms.

Multicollinearity was significantly lower in the filtered model (max VIF: 3.69 vs. 16.63), supporting better generalization and interpretability. The VIF is less than 5, confirmed no severe multicollinearity.

```
library(car)
# Normality Test
# Breusch Pagan Test on Logistic Model
BPtest_basemodel <- ols_test_breusch_pagan(lmod)</pre>
BPtest_filtered.model <- ols_test_breusch_pagan(binary.model.filtered)
# Shapiro-Wilk test
SWtest_basemodel <- shapiro.test(residuals(lmod))</pre>
SWtest_filteredmodel <- shapiro.test(residuals(binary.model.filtered))</pre>
# Combine into unified rows
diagnostics_summary <- data.frame(</pre>
 Model = c("Base.model", "Filtered.model"),
 BP_Statistic = c(BPtest_basemodel$bp, BPtest_filtered.model$bp),
  BP pvalue = c(BPtest basemodel$p, BPtest filtered.model$p),
  Shapiro_W = c(SWtest_basemodel$statistic, SWtest_filteredmodel$statistic),
  Shapiro_pvalue = c(SWtest_basemodel$p.value, SWtest_filteredmodel$p.value)
)
```

Table 6: Summary of Diagnostic tests on Filtered vs Base models

Model	BP_Statistic	BP_pvalue	Shapiro_W	Shapiro_pvalue
Base.model	10.11660	0.0014694	0.9752266	0
Filtered.model	10.31519	0.0013194	0.9751365	0

Run VIF on the logistic model
ols_vif_tol(lmod)

Table 7: VIF Results for Base Logistic Regression Model

Variables	Tolerance	VIF
age	0.8112422	1.232677
sibsp	0.8133135	1.229538
parch	0.7793306	1.283152
fare	0.5101132	1.960349
pclass_1	0.2468157	4.051607
$pclass_2$	0.7943336	1.258917
sex_M	0.8512441	1.174751
$embarked_C$	0.7176662	1.393405
$embarked_Q$	0.8520166	1.173686
$\operatorname{deck}_{-}A$	0.2572099	3.887875
$deck_B$	0.1656314	6.037501
$deck_C$	0.1183078	8.452526
$deck_D$	0.1511669	6.615204
$deck_E$	0.1232885	8.111060
$\operatorname{deck}_{-}\!F$	0.0562819	17.767709

ols_vif_tol(binary.model.filtered)

Table 8: VIF Results for Filtered Logistic Regression Model

Variables	Tolerance	VIF
age	0.8251918	1.211839
sibsp	0.9323496	1.072559
pclass_1	0.2772087	3.607391
$pclass_2$	0.8044472	1.243090
sex_M	0.9085314	1.100677
$embarked_C$	0.7371420	1.356591
$embarked_Q$	0.8625868	1.159304
$\operatorname{deck}_{-}A$	0.6525446	1.532462
$deck_B$	0.5110353	1.956812
$\operatorname{deck}_{-}C$	0.4334379	2.307135
$deck_D$	0.5807436	1.721930
$deck_F$	0.3568101	2.802611

```
# Base Model - Train
observations.train <- train$survived
y_hat_base_train <- predict(lmod, train, type = "response")</pre>
pred_base_train <- ifelse(y_hat_base_train > optimal.cutoff[1], 1, 0)
ModelTrain_base <- data.frame(obs = observations.train, pred = pred_base_train)
log.train.base <- defaultSummary(ModelTrain_base)</pre>
# Base Model - Test
observations.test <- test$survived
y_hat_base_test <- predict(lmod, test, type = "response")</pre>
pred_base test <- ifelse(y_hat_base_test > optimal.cutoff[1], 1, 0)
ModelTest_base <- data.frame(obs = observations.test, pred = pred_base_test)
log.test.base <- defaultSummary(ModelTest_base)</pre>
# Base Model - Train
observations.train <- train$survived
y hat base train <- predict(binary.model.filtered, train, type = "response")</pre>
pred_base_train <- ifelse(y_hat_base_train > optimal.cutoff[1], 1, 0)
ModelTrain_base <- data.frame(obs = observations.train, pred = pred_base_train)</pre>
filtered.log.train.base <- defaultSummary(ModelTrain_base)</pre>
# Base Model - Test
observations.test <- test$survived
y_hat_base_test <- predict(binary.model.filtered, test, type = "response")</pre>
pred_base_test <- ifelse(y_hat_base_test > optimal.cutoff[1], 1, 0)
ModelTest base <- data.frame(obs = observations.test, pred = pred base_test)
filtered.test.base <- defaultSummary(ModelTest_base)</pre>
```

```
data.frame(rbind(log.train.base,log.test.base,filtered.log.train.base,filtered.test.base))
```

Table 9: Comparison of stats between base and filtered log models:

	RMSE	Rsquared	MAE
log.train.base	0.3909456	0.4631895	0.1528384
log.test.base	0.4040610	0.4177733	0.1632653
filtered.log.train.base	0.3909456	0.4627873	0.1528384
filtered.test.base	0.4040610	0.4177733	0.1632653

9 Ongoing Model Monitoring Plan

In order to maintain the effectiveness of the model, we would need to continue to test it on new data. Since the Titanic was a rare event, we do not have a lot of new data to test on the model, but we can still be prepared in case new data were to become available. The first step in monitoring the model is to determine specific thresholds that we expect the model to stay above. We would want the model to maintain certain R², RMSE, and MAE values in order to determine that the model is working correctly. One of the biggest concerns with our model is data drift. Since the Titanic sank over 100 years ago, the data that we are using from the model may not align with today relevant to ship travel today.

10 Conclusion

Summarize your results here. What is the best model for the data and why?

Appendix A: Check if 'sibsp' and 'parch' should be continuous or categorical

We don't see significant improvement between modeling these predictors as continuous or categorical, therefore we decided to leave them as continuous.

```
library(car)
data.clean.ap1 = odata[, !(names(odata) %in% c("name", "ticket", "boat", "body", "home.dest"))]
Data Augmentation
#Extract deck letter from cabin
data.clean.ap1$deck <- substr(data.clean.ap1$cabin, 1,1)</pre>
# Remove cabin col:
data.clean.ap1$cabin <- NULL
Imputing data
# ---- Age----
#Replace NAs in age column with Median value
median_age <- median(data.clean.ap1$age, na.rm = TRUE)</pre>
data.clean.ap1 <- data.clean.ap1 %>%
 mutate(age = ifelse(is.na(age), median_age, age))
# ---- deck----
# For deck, since its a category, we decided to use KNN to impute the column:
# Install if not already installed
# install.packages("VIM")
library(VIM)
# Replace "" with NA in the 'deck' column
data.clean.ap1$deck[data.clean.ap1$deck == ""] <- NA
# Convert 'cabin' to factor
data.clean.ap1$deck <- as.factor(data.clean.ap1$deck)
# Apply kNN imputation just to Cabin column
data.clean.ap1 <- kNN(data.clean.ap1, variable = "deck", k = 5)
# Check that NAs were imputed
# sum(is.na(data.clean$deck))
                            # Original
# sum(is.na(data.clean.imputed$deck)) # After
```

```
# Remove indicator col:
data.clean.ap1$deck_imp <- NULL
Dummify Cat. cols
# Dummifying pclass:
data.clean.ap1$pclass_1 = ifelse(data.clean.ap1$pclass == 1, 1, 0)
data.clean.ap1$pclass_2 = ifelse(data.clean.ap1$pclass == 2, 1, 0)
# Dummifying sex:
data.clean.ap1$sex_M = ifelse(data.clean.ap1$sex == 'male', 1, 0)
# Dummifying embarked:
data.clean.ap1$embarked_C = ifelse(data.clean.ap1$embarked == 'C', 1, 0)
data.clean.ap1$embarked Q = ifelse(data.clean.ap1$embarked == 'Q', 1, 0)
# Dummifying deck:
data.clean.ap1$deck_A = ifelse(data.clean.ap1$deck == 'A', 1, 0)
data.clean.ap1$deck_B = ifelse(data.clean.ap1$deck == 'B', 1, 0)
data.clean.ap1$deck_C = ifelse(data.clean.ap1$deck == 'C', 1, 0)
data.clean.ap1$deck_D = ifelse(data.clean.ap1$deck == 'D', 1, 0)
data.clean.ap1$deck_E = ifelse(data.clean.ap1$deck == 'E', 1, 0)
data.clean.ap1$deck_F = ifelse(data.clean.ap1$deck == 'F', 1, 0)
#data.clean.ap1$deck G = ifelse(data.clean.ap1$deck == 'G', 1, 0) # removed due to causing issues
# Dummifying sibsp:
data.clean.ap1$sibsp_1 = ifelse(data.clean.ap1$sibsp == 1, 1, 0)
data.clean.ap1$sibsp_2 = ifelse(data.clean.ap1$sibsp == 2, 1, 0)
data.clean.ap1$sibsp_3 = ifelse(data.clean.ap1$sibsp == 3, 1, 0)
data.clean.ap1$sibsp_4 = ifelse(data.clean.ap1$sibsp == 4, 1, 0)
data.clean.ap1$sibsp_5 = ifelse(data.clean.ap1$sibsp == 5, 1, 0)
#data.clean.ap1$sibsp_8 = ifelse(data.clean.ap1$sibsp == 8, 1, 0) # removed due to causing issues
# Dummifying parch:
data.clean.ap1$parch_1 = ifelse(data.clean.ap1$parch == 1, 1, 0)
data.clean.ap1$parch_2 = ifelse(data.clean.ap1$parch == 2, 1, 0)
data.clean.ap1$parch_3 = ifelse(data.clean.ap1$parch == 3, 1, 0)
data.clean.ap1$parch_4 = ifelse(data.clean.ap1$parch == 4, 1, 0)
data.clean.ap1$parch_5 = ifelse(data.clean.ap1$parch == 5, 1, 0)
data.clean.ap1$parch_6 = ifelse(data.clean.ap1$parch == 6, 1, 0)
#data.clean.ap1$parch 9 = ifelse(data.clean.ap1$parch == 9, 1, 0) # removed due to causing issues
# Removing Dummified cols:
data.clean.ap1 = subset(data.clean.ap1, select = -c(pclass, sex, embarked, deck))#, sibsp, parch))
data.clean.ap1 = na.omit(data.clean.ap1)
cat(nrow(odata) - nrow(data.clean.ap1), rows were removed from original dataset)
```

```
set.seed(567)
train_indices_ap1 = sample(1 : nrow(data.clean.ap1), size = 0.7005*nrow(data.clean.ap1), replace = FA
train.ap1 = data.clean.ap1[train_indices_ap1,]
test.ap1 = data.clean.ap1[-train indices ap1,]
cat("We are using:", nrow(train.ap1)/nrow(data.clean.ap1) * 100, '% of the data for training')
We are using: 70.03058 % of the data for training
mulvar_model.ap1 <- lm(survived ~ ., data = train.ap1)</pre>
summary(mulvar_model.ap1)
Call:
lm(formula = survived ~ ., data = train.ap1)
Residuals:
   Min
           1Q Median
                          3Q
                                Max
-1.2866 -0.1941 -0.0224 0.1911 0.9739
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 1.1796068 0.1203338
                                9.803 < 2e-16 ***
          age
          -0.0401162 0.0221954 -1.807 0.07104 .
sibsp
          -0.0274112 0.0270379 -1.014 0.31095
parch
fare
           0.0004356 0.0002985 1.459 0.14481
pclass_1
          0.1097829 0.0542228 2.025 0.04320 *
          0.1381701 0.0312063 4.428 1.07e-05 ***
pclass_2
sex M
          -0.4129619  0.0252816  -16.335  < 2e-16 ***
embarked_C 0.0461410 0.0323127
                                1.428 0.15366
          0.1210184 0.0426024 2.841 0.00461 **
embarked Q
deck_A
          -0.3123626  0.1372942  -2.275  0.02314 *
\mathtt{deck}_{\mathtt{B}}
deck_C
          -0.3040370 0.1331009 -2.284 0.02259 *
          \mathtt{deck}_{\mathtt{D}}
          0.0041655 0.1243264
                              0.034 0.97328
\mathtt{deck}_{\mathtt{E}}
          deck_F
           0.0567316 0.0355896
                               1.594 0.11128
sibsp_1
sibsp_2
          0.1212082 0.0797596
                               1.520 0.12895
sibsp_3
          -0.1501858 0.1195550 -1.256 0.20937
sibsp_4
          -0.2298081 0.1296372 -1.773 0.07662 .
sibsp_5
          -0.2293432 0.2019073 -1.136 0.25631
          parch_1
           0.1456666  0.0715147  2.037  0.04196 *
parch_2
parch_3
           0.2704118 0.1739380
                                1.555 0.12039
parch_4
          -0.0799175 0.2038416 -0.392 0.69511
parch_5
           0.0679073 0.2385574
                                0.285 0.77597
parch_6
           0.0143690 0.2895211
                                0.050 0.96043
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
Residual standard error: 0.34 on 889 degrees of freedom
Multiple R-squared: 0.5273, Adjusted R-squared: 0.5135
F-statistic: 38.14 on 26 and 889 DF, p-value: < 2.2e-16
```

vif(mulvar_model.ap1)

```
age
                sibsp
                          parch
                                       fare
                                              pclass_1
                                                         pclass_2
                                                                       sex_M
 1.460949
            3.467000
                        4.494487
                                   1.922834
                                              4.492433
                                                         1.297793
                                                                    1.164100
embarked_C embarked_Q
                          deck_A
                                     deck_B
                                                \mathtt{deck}_{\mathtt{C}}
                                                          deck_D
                                                                      deck_E
 1.437150
            1.172137
                       5.135825
                                   8.715727 13.411354
                                                        9.672668 12.015361
   deck_F
             {	t sibsp\_1}
                      sibsp_2
                                 sibsp_3
                                            {	t sibsp\_4}
                                                        sibsp_5
                                                                     parch_1
                                            1.863037
                                                         1.404398
26.716246
            1.876291 1.441990 1.464254
                                                                    1.911901
  parch_2
            parch_3 parch_4
                                 parch_5
                                            parch_6
 3.083293
            1.301393
                       1.431435
                                   1.472004
                                              1.446997
```

lmod.ap1 <- glm(as.factor(survived) ~ ., family = binomial, data = train.ap1)
summary(lmod.ap1)</pre>

Call:

glm(formula = as.factor(survived) ~ ., family = binomial, data = train.ap1)

Estimate Std. Error z value Pr(>|z|)

Coefficients:

```
(Intercept) 5.253e+00 1.384e+00 3.797 0.000147 ***
age
           -5.481e-02 9.965e-03 -5.500 3.79e-08 ***
sibsp
           -1.946e+00 1.291e+02 -0.015 0.987976
            -1.642e+00 1.623e+02 -0.010 0.991928
parch
fare
           1.057e-03 2.428e-03 0.435 0.663393
pclass_1
           1.156e+00 4.819e-01
                                    2.398 0.016480 *
pclass_2
           1.496e+00 3.095e-01 4.834 1.34e-06 ***
sex_M
            -2.982e+00 2.464e-01 -12.103 < 2e-16 ***
embarked_C 6.410e-01 2.868e-01
                                    2.235 0.025424 *
embarked_Q
            1.316e+00 3.735e-01 3.523 0.000427 ***
deck_A
            -2.866e+00 1.503e+00 -1.907 0.056492 .
\mathtt{deck}_{\mathtt{B}}
            -2.498e+00 1.479e+00 -1.689 0.091220 .
deck_C
            -2.624e+00 1.445e+00 -1.817 0.069262 .
\mathtt{deck}_{\mathtt{D}}
           -2.971e+00 1.415e+00 -2.100 0.035711 *
deck_E
            3.196e-01 1.404e+00
                                  0.228 0.819965
            -4.753e+00 1.358e+00 -3.499 0.000467 ***
\mathtt{deck}_{\mathtt{F}}
sibsp_1
            2.065e+00 1.291e+02
                                    0.016 0.987240
             4.356e+00 2.583e+02
                                    0.017 0.986546
sibsp_2
sibsp_3
             3.607e+00 3.874e+02
                                    0.009 0.992572
             4.557e+00 5.166e+02
sibsp_4
                                    0.009 0.992962
sibsp_5
            -6.723e+00 1.240e+03 -0.005 0.995674
parch_1
             2.597e+00 1.623e+02
                                    0.016 0.987237
parch_2
             4.192e+00 3.247e+02
                                    0.013 0.989698
parch_3
             6.236e+00 4.870e+02
                                    0.013 0.989785
parch_4
             6.227e+00 6.494e+02
                                    0.010 0.992349
parch_5
             8.544e+00 8.117e+02
                                    0.011 0.991602
parch_6
            -4.112e+00 1.770e+03 -0.002 0.998146
```

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 (Dispersion parameter for binomial family taken to be 1) Null deviance: 1223.12 on 915 degrees of freedom Residual deviance: 602.86 on 889 degrees of freedom AIC: 656.86 Number of Fisher Scoring iterations: 15 vif(lmod.ap1) age sibsp parch fare pclass_1 pclass_2 1.653195e+00 7.591027e+05 1.634339e+06 1.757536e+00 4.949336e+00 1.555093e+00 sex M embarked_C embarked Q deck_A deck_B deck C 1.404155e+00 1.549826e+00 1.415978e+00 1.012962e+01 1.475410e+01 2.517074e+01 deck D deck E sibsp_1 sibsp_2 sibsp_3 deck_F 1.980371e+01 1.464836e+01 4.346983e+01 3.294475e+05 1.808931e+05 1.937224e+05 sibsp_4 sibsp 5 parch 1 parch 2 parch 3 parch_4 1.974919e+05 1.372219e+00 3.139987e+05 6.986245e+05 2.469813e+05 2.280957e+05 parch 5 parch 6 4.056249e+05 1.434520e+00 y hat_mulvar_train.ap1<-predict(mulvar_model.ap1, data = train.ap1)</pre> predictions_train.ap1 <- ifelse(y_hat_mulvar_train.ap1 > 0.5, 1, 0) ModelTrain mulvar.ap1<-data.frame(obs = train.ap1\$survived, pred=predictions_train.ap1) linear.train.ap1 <- defaultSummary(ModelTrain_mulvar.ap1)</pre> y_hat_mulvar_test.ap1<-predict(mulvar_model.ap1, newdata = test.ap1)</pre> predictions_test.ap1 <- ifelse(y_hat_mulvar_test.ap1 > 0.5, 1, 0) ModelTest_mulvar.ap1<-data.frame(obs = test.ap1\$survived, pred=predictions_test.ap1) linear.test.ap1 <- defaultSummary(ModelTest_mulvar.ap1)</pre> y_hat_log_train.ap1<-predict(lmod.ap1, data = train.ap1)</pre> predictions_log_train.ap1 <- ifelse(y_hat_log_train.ap1 > 0.5, 1, 0) ModelTrain_lmod.ap1<-data.frame(obs = train.ap1\$survived, pred=predictions_log_train.ap1) log.train.ap1 <- defaultSummary(ModelTrain_lmod.ap1)</pre> y_hat_log_test.ap1<-predict(lmod.ap1, newdata = test.ap1)</pre> predictions_log_test.ap1 <- ifelse(y_hat_log_test.ap1 > 0.5, 1, 0) ModelTest_lmod.ap1<-data.frame(obs = test.ap1\$survived, pred=predictions_log_test.ap1) log.test.ap1 <- defaultSummary(ModelTest_lmod.ap1)</pre>

	RMSE	Rsquared	MAE
linear.train.ap1	0.3796114	0.4824536	0.1441048
linear.test.ap1	0.3976975	0.4365256	0.1581633

data.frame(rbind(linear.train.ap1,linear.test.ap1,log.train.ap1,log.test.ap1))

	RMSE	Rsquared	MAE
log.train.ap1 log.test.ap1	$\begin{array}{c} 0.3964912 \\ 0.4008919 \end{array}$	$\begin{array}{c} 0.4429840 \\ 0.4214242 \end{array}$	

confusion_matrix_mulvar_train.ap1 <- confusionMatrix(as.factor(predictions_train.ap1), as.factor(train.ap1)
confusion_matrix_mulvar_train.ap1</pre>

Confusion Matrix and Statistics

Reference

Prediction 0 1 0 502 73 1 59 282

Accuracy : 0.8559

95% CI : (0.8315, 0.878)

No Information Rate : 0.6124 P-Value [Acc > NIR] : <2e-16

Kappa : 0.6942

Mcnemar's Test P-Value: 0.2578

Precision: 0.8270 Recall: 0.7944

F1 : 0.8103 Prevalence : 0.3876 Detection Rate : 0.3079

Detection Prevalence : 0.3723 Balanced Accuracy : 0.8446

'Positive' Class : 1

confusion_matrix_mulvar_test.ap1 <- confusionMatrix(as.factor(predictions_test.ap1), as.factor(test.acconfusion_matrix_mulvar_test.ap1)</pre>

Confusion Matrix and Statistics

Reference

Prediction 0 1 0 216 31 1 31 114

Accuracy : 0.8418

95% CI: (0.8019, 0.8765)

No Information Rate : 0.6301 P-Value [Acc > NIR] : <2e-16

Kappa : 0.6607

Mcnemar's Test P-Value : 1

Precision : 0.7862 Recall : 0.7862

Prevalence: 0.3699

F1 : 0.7862

Detection Rate: 0.2908
Detection Prevalence: 0.3699
Balanced Accuracy: 0.8304

'Positive' Class : 1

confusion_matrix_log_train.ap1 <- confusionMatrix(as.factor(predictions_log_train.ap1), as.factor(trainconfusion_matrix_log_train.ap1)</pre>

Confusion Matrix and Statistics

Reference

Prediction 0 1 0 523 106 1 38 249

Accuracy : 0.8428

95% CI: (0.8176, 0.8658)

No Information Rate : 0.6124 P-Value [Acc > NIR] : < 2.2e-16

Kappa: 0.6568

Mcnemar's Test P-Value: 2.36e-08

Precision : 0.8676 Recall : 0.7014 F1 : 0.7757

Prevalence : 0.3876
Detection Rate : 0.2718
Detection Prevalence : 0.3133
Balanced Accuracy : 0.8168

'Positive' Class : 1

confusion_matrix_log_test.ap1 <- confusionMatrix(as.factor(predictions_log_test.ap1), as.factor(test
confusion_matrix_log_test.ap1</pre>

Confusion Matrix and Statistics

Reference

Prediction 0 1

```
0 226 42
1 21 103
```

Accuracy: 0.8393

95% CI: (0.7991, 0.8742)

No Information Rate : 0.6301 P-Value [Acc > NIR] : < 2e-16

Kappa : 0.6446

Mcnemar's Test P-Value: 0.01174

Precision : 0.8306 Recall : 0.7103

F1 : 0.7658 Prevalence : 0.3699 Detection Rate : 0.2628

Detection Prevalence : 0.3163 Balanced Accuracy : 0.8127

'Positive' Class: 1

```
library(car)
data.clean.ap2 = odata[, !(names(odata) %in% c("name", "ticket", "boat", "body", "home.dest"))]
Data Augmentation
#Extract deck letter from cabin
data.clean.ap2$deck <- substr(data.clean.ap2$cabin, 1,1)</pre>
# Remove cabin col:
data.clean.ap2$cabin <- NULL
Imputing data
# ---- Age----
#Replace NAs in age column with Median value
median_age <- median(data.clean.ap2$age, na.rm = TRUE)</pre>
data.clean.ap2 <- data.clean.ap2 %>%
 mutate(age = ifelse(is.na(age), median_age, age))
# ---- deck----
# For deck, since its a category, we decided to use KNN to impute the column:
# Install if not already installed
# install.packages("VIM")
library(VIM)
# Replace "" with NA in the 'deck' column
```

```
data.clean.ap2$deck[data.clean.ap2$deck == ""] <- NA
# Convert 'cabin' to factor
data.clean.ap2$deck <- as.factor(data.clean.ap2$deck)
# Apply kNN imputation just to Cabin column
data.clean.ap2 <- kNN(data.clean.ap2, variable = "deck", k = 5)
# Check that NAs were imputed
# sum(is.na(data.clean$deck))
                                  # Original
# sum(is.na(data.clean.imputed$deck)) # After
# Remove indicator col:
data.clean.ap2$deck_imp <- NULL
Dummify Cat. cols
# Dummifying pclass:
data.clean.ap2$pclass_1 = ifelse(data.clean.ap2$pclass == 1, 1, 0)
data.clean.ap2$pclass 2 = ifelse(data.clean.ap2$pclass == 2, 1, 0)
# Dummifying sex:
data.clean.ap2$sex_M = ifelse(data.clean.ap2$sex == 'male', 1, 0)
# Dummifying embarked:
data.clean.ap2$embarked_C = ifelse(data.clean.ap2$embarked == 'C', 1, 0)
data.clean.ap2$embarked_Q = ifelse(data.clean.ap2$embarked == 'Q', 1, 0)
# Dummifying deck:
data.clean.ap2$deck_A = ifelse(data.clean.ap2$deck == 'A', 1, 0)
data.clean.ap2$deck_B = ifelse(data.clean.ap2$deck == 'B', 1, 0)
data.clean.ap2$deck_C = ifelse(data.clean.ap2$deck == 'C', 1, 0)
data.clean.ap2$deck_D = ifelse(data.clean.ap2$deck == 'D', 1, 0)
data.clean.ap2$deck_E = ifelse(data.clean.ap2$deck == 'E', 1, 0)
data.clean.ap2$deck F = ifelse(data.clean.ap2$deck == 'F', 1, 0)
#data.clean.ap2$deck_G = ifelse(data.clean.ap2$deck == 'G', 1, 0) # removed due to causing issues
# Dummifying sibsp to 2 categories:
data.clean.ap2$sibsp_y = ifelse(data.clean.ap2$sibsp > 0, 1, 0)
# Dummifying parch to 2 categories:
data.clean.ap2$parch_y = ifelse(data.clean.ap2$parch > 0, 1, 0)
# Removing Dummified cols:
data.clean.ap2 = subset(data.clean.ap2, select = -c(pclass, sex, embarked, deck))#, sibsp, parch))
data.clean.ap2 = na.omit(data.clean.ap2)
cat(nrow(odata) - nrow(data.clean.ap2), 'rows were removed from original dataset')
```

vif(mulvar_model.ap2)

```
set.seed(567)
train_indices_ap2 = sample(1 : nrow(data.clean.ap2), size = 0.7005*nrow(data.clean.ap2), replace = FA
train.ap2 = data.clean.ap2[train_indices_ap2,]
test.ap2 = data.clean.ap2[-train_indices_ap2,]
cat("We are using:", nrow(train.ap2)/nrow(data.clean.ap2) * 100, '% of the data for training')
We are using: 70.03058 % of the data for training
mulvar_model.ap2 <- lm(survived ~ ., data = train.ap2)</pre>
summary(mulvar_model.ap2)
Call:
lm(formula = survived ~ ., data = train.ap2)
Residuals:
    Min
              1Q
                   Median
                               3Q
                                       Max
-1.27243 -0.19451 -0.02769 0.19202 0.96436
Coefficients:
             Estimate Std. Error t value Pr(>|t|)
(Intercept) 1.1720948 0.1212992 9.663 < 2e-16 ***
           age
           -0.0867184   0.0185029   -4.687   3.21e-06 ***
sibsp
           parch
           0.0003848 0.0003005 1.281 0.200641
fare
pclass_1
           0.1055453 0.0557446 1.893 0.058629 .
            0.1376903 0.0314471
                                 4.378 1.34e-05 ***
pclass_2
          -0.4270550 0.0253448 -16.850 < 2e-16 ***
sex_M
embarked_C 0.0585880 0.0319987 1.831 0.067439 .
embarked_Q 0.1061446 0.0425937 2.492 0.012881 *
\mathtt{deck}_{\mathtt{A}}
          -0.3315651 0.1464353 -2.264 0.023797 *
           -0.2852592  0.1389479  -2.053  0.040362 *
\mathtt{deck}_{\mathtt{B}}
\mathtt{deck}_{\mathtt{C}}
          -0.2951467  0.1350558  -2.185  0.029120 *
           \mathtt{deck}_{\mathtt{D}}
\mathtt{deck}_{\mathtt{E}}
           0.0078459 0.1256544 0.062 0.950226
deck_F
           -0.5745493  0.1193029  -4.816  1.72e-06 ***
           sibsp_y
parch_y
            0.1573960 0.0479445 3.283 0.001067 **
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.3436 on 898 degrees of freedom
                              Adjusted R-squared: 0.5031
Multiple R-squared: 0.5124,
F-statistic: 55.5 on 17 and 898 DF, p-value: < 2.2e-16
```

```
fare pclass_1 pclass_2
                                                                   sex_M
      age
               sibsp
                         parch
            2.359204
  1.380228
                      2.776007
                                 1.908310 4.649205 1.290432
                                                                1.145551
embarked_C embarked_Q deck_A
                                 deck_B
                                            deck_C deck_D deck_E
  1.379990 1.147247 5.439805 8.740948 13.154666 9.148653 11.598396
   deck_F
           sibsp_y parch_y
 26.237760
            2.297316
                      3.213776
lmod.ap2 <- glm(as.factor(survived) ~ ., family = binomial, data = train.ap2)</pre>
summary(lmod.ap2)
Call:
glm(formula = as.factor(survived) ~ ., family = binomial, data = train.ap2)
Coefficients:
             Estimate Std. Error z value Pr(>|z|)
(Intercept) 5.1275081 1.3016899 3.939 8.18e-05 ***
           age
          -1.0266480 0.2640080 -3.889 0.000101 ***
sibsp
           -0.2791962 0.1939136 -1.440 0.149925
parch
fare
           0.0007343 0.0024042 0.305 0.760042
          1.1421684 0.4862413 2.349 0.018825 *
pclass_1
           1.4015232  0.2992294  4.684  2.82e-06 ***
pclass_2
           -3.0275901 0.2424805 -12.486 < 2e-16 ***
sex_M
embarked_C 0.6732550 0.2814737 2.392 0.016762 *
embarked_Q 1.2284326 0.3679272 3.339 0.000841 ***
           -2.5767916 1.4219365 -1.812 0.069960 .
\mathtt{deck}_{\mathtt{A}}
          -2.2194317 1.4042525 -1.581 0.113991
deck_B
           -2.3552547 1.3682224 -1.721 0.085179 .
deck_C
deck_D
           -2.9570245 1.3451465 -2.198 0.027928 *
deck E
           0.4470949 1.3202938 0.339 0.734886
          -4.4528239 1.2754924 -3.491 0.000481 ***
deck_F
sibsp_y
          1.3882952 0.3947966 3.516 0.000437 ***
            1.2304958 0.4199141 2.930 0.003386 **
parch_y
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
   Null deviance: 1223.12 on 915 degrees of freedom
Residual deviance: 620.68 on 898 degrees of freedom
AIC: 656.68
Number of Fisher Scoring iterations: 6
vif(lmod.ap2)
      age
               sibsp
                         parch
                                     fare
                                           pclass_1
                                                      pclass_2
                                                                   sex_M
  1.619429
            3.641267
                      2.989704
                                 1.753639
                                           5.122869
                                                      1.534060
                                                                1.410059
embarked_C embarked_Q
                        deck_A
                                   deck_B
                                             \mathtt{deck}_{\mathtt{C}}
                                                        deck_D
                                                                  deck_E
  1.487757
            1.363260
                      9.852461 13.562280 21.612617 16.823106 13.659175
```

```
deck_F sibsp_y parch_y 39.454519 3.482784 3.287406
```

```
y_hat_mulvar_train.ap2<-predict(mulvar_model.ap2, data = train.ap2)</pre>
predictions_train.ap2 <- ifelse(y_hat_mulvar_train.ap2 > 0.5, 1, 0)
ModelTrain_mulvar.ap2<-data.frame(obs = train.ap2$survived, pred=predictions_train.ap2)</pre>
linear.train.ap2 <- defaultSummary(ModelTrain_mulvar.ap2)</pre>
y_hat_mulvar_test.ap2<-predict(mulvar_model.ap2, newdata = test.ap2)</pre>
predictions_test.ap2 <- ifelse(y_hat_mulvar_test.ap2 > 0.5, 1, 0)
ModelTest_mulvar.ap2<-data.frame(obs = test.ap2$survived, pred=predictions_test.ap2)</pre>
linear.test.ap2 <- defaultSummary(ModelTest_mulvar.ap2)</pre>
y_hat_log_train.ap2<-predict(lmod.ap2, data = train.ap2)</pre>
predictions_log_train.ap2 <- ifelse(y_hat_log_train.ap2 > 0.5, 1, 0)
ModelTrain_lmod.ap2<-data.frame(obs = train.ap2$survived, pred=y_hat_log_train.ap2)</pre>
log.train.ap2 <- defaultSummary(ModelTrain_mulvar.ap2)</pre>
y_hat_log_test.ap2<-predict(lmod.ap2, newdata = test.ap2)</pre>
predictions_log_test.ap2 <- ifelse(y_hat_log_test.ap2 > 0.5, 1, 0)
ModelTest_lmod.ap2<-data.frame(obs = test.ap2$survived, pred=predictions_log_test.ap2)
log.test.ap2 <- defaultSummary(ModelTest_lmod.ap2)</pre>
data.frame(rbind(linear.train.ap2,linear.test.ap2,log.train.ap2,log.test.ap2))
```

	RMSE	Rsquared	MAE
linear.train.ap2	0.3909456	0.4570397	0.1528384
linear.test.ap2	0.4040610	0.4221825	0.1632653
log.train.ap2	0.3909456	0.4570397	0.1528384
log.test.ap2	0.4008919	0.4212961	0.1607143

confusion_matrix_mulvar_train.ap2 <- confusionMatrix(as.factor(predictions_train.ap2), as.factor(train.ap2)
confusion_matrix_mulvar_train.ap2</pre>

Confusion Matrix and Statistics

Reference
Prediction 0 1
0 498 77
1 63 278

Accuracy : 0.8472

95% CI: (0.8222, 0.8699)

No Information Rate : 0.6124 P-Value [Acc > NIR] : <2e-16

Kappa : 0.6757

Mcnemar's Test P-Value : 0.2719

Precision : 0.8152 Recall : 0.7831

F1 : 0.7989 Prevalence : 0.3876

Detection Rate: 0.3035

Detection Prevalence : 0.3723 Balanced Accuracy : 0.8354

'Positive' Class : 1

confusion_matrix_mulvar_test.ap2 <- confusionMatrix(as.factor(predictions_test.ap2), as.factor(test.ap2)
confusion_matrix_mulvar_test.ap2</pre>

Confusion Matrix and Statistics

Reference

Prediction 0 1 0 215 32 1 32 113

Accuracy : 0.8367

95% CI : (0.7964, 0.8719)

No Information Rate : 0.6301 P-Value [Acc > NIR] : <2e-16

Kappa : 0.6498

Mcnemar's Test P-Value : 1

Precision : 0.7793 Recall : 0.7793

F1 : 0.7793

Prevalence: 0.3699
Detection Rate: 0.2883

Detection Prevalence: 0.3699
Balanced Accuracy: 0.8249

'Positive' Class : 1

confusion_matrix_log_train.ap2 <- confusionMatrix(as.factor(predictions_log_train.ap2), as.factor(train)
confusion_matrix_log_train.ap2</pre>

Confusion Matrix and Statistics

Reference

Prediction 0 1 0 523 111

Accuracy: 0.8373

95% CI: (0.8118, 0.8607)

No Information Rate : 0.6124 P-Value [Acc > NIR] : < 2.2e-16

Kappa: 0.6439

Mcnemar's Test P-Value : 3.669e-09

Precision : 0.8652 Recall : 0.6873 F1 : 0.7661

Prevalence: 0.3876
Detection Rate: 0.2664

Detection Prevalence : 0.3079 Balanced Accuracy : 0.8098

'Positive' Class : 1

confusion_matrix_log_test.ap2 <- confusionMatrix(as.factor(predictions_log_test.ap2), as.factor(test
confusion_matrix_log_test.ap2</pre>

Confusion Matrix and Statistics

Reference

Prediction 0 1

0 227 43

1 20 102

Accuracy : 0.8393

95% CI: (0.7991, 0.8742)

No Information Rate : 0.6301 P-Value [Acc > NIR] : < 2.2e-16

Kappa: 0.6436

Mcnemar's Test P-Value: 0.005576

Precision : 0.8361 Recall : 0.7034

F1 : 0.7640

Prevalence: 0.3699

Detection Rate: 0.2602

Detection Prevalence: 0.3112 Balanced Accuracy: 0.8112

'Positive' Class : 1

References

National Oceanic and Atmospheric Administration (NOAA). (2023). RMS titanic – history and significance. https://www.noaa.gov/office-of-general-counsel/gc-international-section/rms-titanic-history-and-significance Statology. (2025). How to measure correlation between categorical variables. https://www.statology.org/correlation-between-categorical-variables/