

Doubly robust estimation under a randomly censored covariate

Brian Richardson

Slides Available Online



Acknowledgements

Seong-Ho Lee, PhD Tanya Garcia, PhD

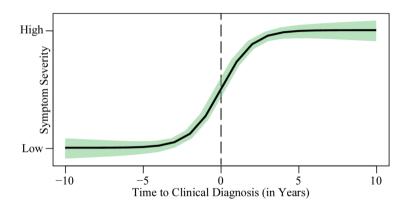


Yanyuan Ma, PhD

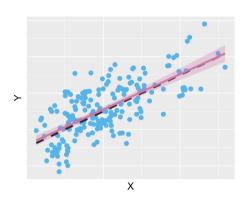


This research was supported by the National Institute of Environmental Health Sciences grant T32ES007018.

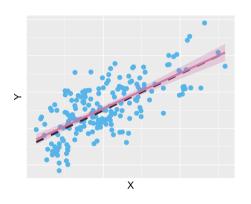
Huntington's Disease and Censored Covariates



Lotspeich et. al. (2024)

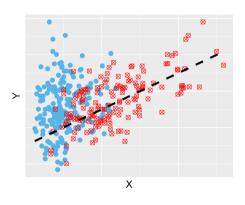


• Regression model: $E(Y) = \beta_0 + \beta_1 X$

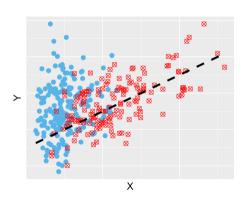


- Regression model: $E(Y) = \beta_0 + \beta_1 X$
- Estimate $\beta = (\beta_0, \beta_1)^T$ by solving:

$$\sum_{i=1}^{n} \underbrace{(Y_i - \beta_0 - \beta_1 X_i)(1, X_i)^T}_{\text{score function: } \mathbf{S}_{\beta}^F} = \mathbf{0}$$

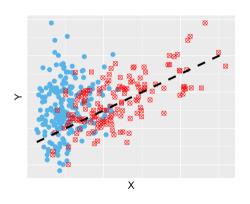


Problem: X is censored by a random censoring time C



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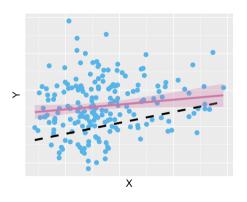
- $W = \min(X, C)$
- $\Delta = I(X \leq C)$



Problem: X is censored by a random censoring time C

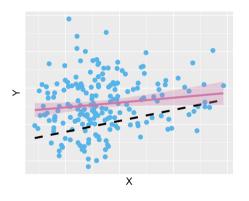
- $W = \min(X, C)$
- $\Delta = I(X \leq C)$
- assume: $C \perp \!\!\! \perp (X, Y)$

A Naive Approach



Naively treat W as X

A Naive Approach

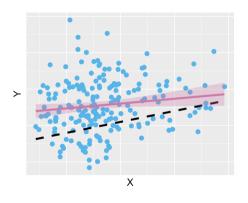


Naively treat W as X

• Estimate β by solving:

$$\sum_{i=1}^{n} \mathbf{S}_{\boldsymbol{\beta}}^{F}(Y_i, W_i, \boldsymbol{\beta}) = \mathbf{0}$$

A Naive Approach



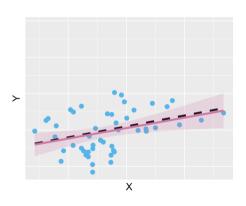
Naively treat W as X

• Estimate β by solving:

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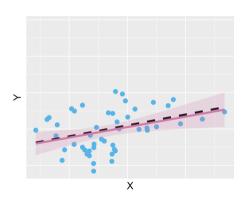
 \times Inconsistent estimator of β

Complete Case Analysis



Only use observations where *X* is *uncensored*

Complete Case Analysis

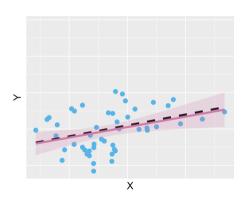


Only use observations where *X* is uncensored

• Estimate β by solving:

$$\sum_{i=1}^{n} \Delta_i \mathbf{S}_{oldsymbol{eta}}^F(Y_i, W_i, oldsymbol{eta}) = \mathbf{0}$$

Complete Case Analysis



Only use observations where *X* is uncensored

• Estimate β by solving:

$$\sum_{i=1}^n \Delta_i \mathbf{S}^F_{oldsymbol{eta}}(Y_i, W_i, oldsymbol{eta}) = \mathbf{0}$$

- ✓ Consistent
- Inefficient

$$f_{Y,W,\Delta}(y,w,\delta,\boldsymbol{\beta}, \boxed{\boldsymbol{\alpha}}) \propto \underbrace{\{f_{Y|X}(y,w,\boldsymbol{\beta})\}^{\delta}}_{\text{uncensored}} \underbrace{\left\{\int_{w}^{\infty} f_{Y|X}(y,x,\boldsymbol{\beta}) \boxed{f_{X}(x,\boldsymbol{\alpha})} dx\right\}^{1-\delta}}_{\text{censored}}$$

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$$\mathbf{S}_{\boldsymbol{\beta}}(y, w, \delta, \boldsymbol{\beta}) \equiv \partial \log f_{Y, W, \Delta}(y, w, \delta, \boldsymbol{\beta}, \alpha) / \partial \boldsymbol{\beta}$$

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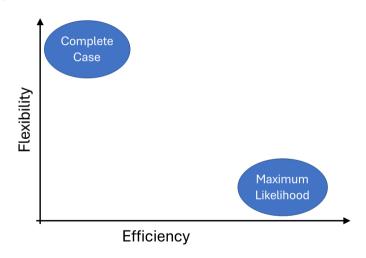
Pros

- \checkmark consistent estimator of β
- ✓ fully efficient

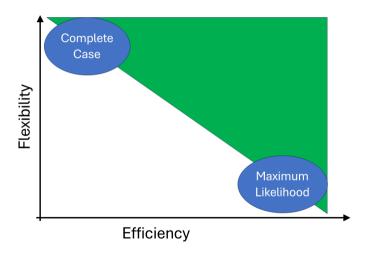
Cons

- X requires model for **nuisance** parameter f_X
- X inconsistent estimator when model for f_X is incorrect

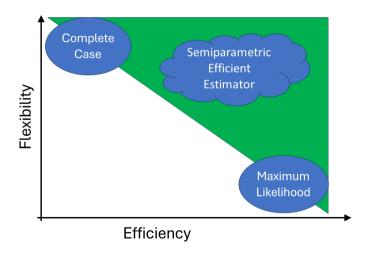
Existing Methods



Existing Opportunity



A New Approach

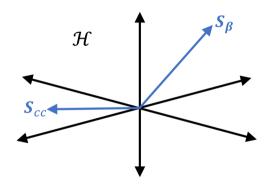


• goal: to find the estimating function resulting in a semiparametric efficient estimator

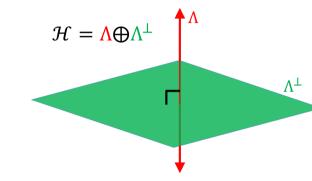
- goal: to find the estimating function resulting in a semiparametric efficient estimator
- semiparametric: infinite dimensional nuisance parameters f_X, f_C

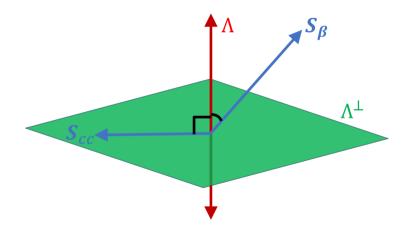
- goal: to find the estimating function resulting in a semiparametric efficient estimator
- semiparametric: infinite dimensional nuisance parameters f_X, f_C
- Geometric approach from Tsiatis (2006)

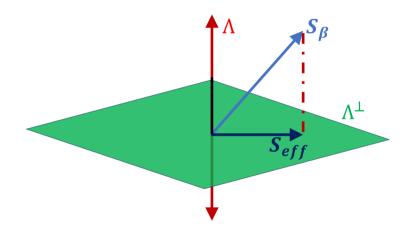
- Hilbert space of estimating functions
- covariance inner product $\langle \boldsymbol{h}, \boldsymbol{g} \rangle = \mathbb{E}(\boldsymbol{h}^T \boldsymbol{g})$



- construct Λ using
 nuisance scores
- ullet orthogonal complement Λ^{\perp}







The semiparametric efficient estimator $\widehat{\beta}_{\rm eff}$ is the solution to

$$\sum_{i=1}^n \mathbf{S}_{ ext{eff}}^F(Y_i,\,W_i,\Delta_i,oldsymbol{eta}) = \mathbf{0}$$

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- consistent and asymptotically normal
- ✓ **Doubly Robust**: $\widehat{\boldsymbol{\beta}}_{\text{eff}}$ is consistent if at least one of f_X , f_C is correctly specified
- ✓ Locally Efficiency: If f_X , f_C are both correctly specified, then $\widehat{\beta}_{\text{eff}}$ achieves the semiparametric efficiency bound

Simulation Setup

• $Y|X \sim N(\beta_0 + \beta_1 X, \sigma^2)$

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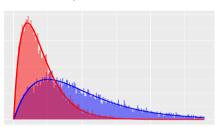
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- $X, C \sim$ gamma distributions

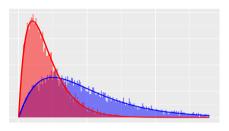
Simulation Setup

- $Y|X \sim N(\beta_0 + \beta_1 X, \sigma^2)$
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- $X, C \sim$ gamma distributions
- X, C possibly **misspecified** as exponential

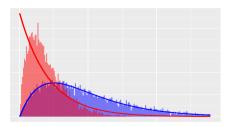
X, C correct



X, C correct

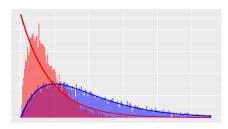


X correct, C incorrect

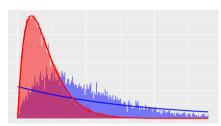


X, C correct

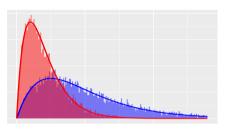
X correct, C incorrect



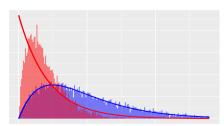
Xincorrect, Ccorrect



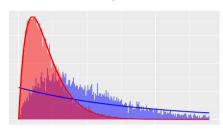
X, C correct



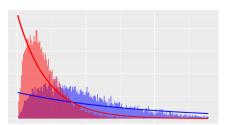
X correct, C incorrect



Xincorrect, Ccorrect

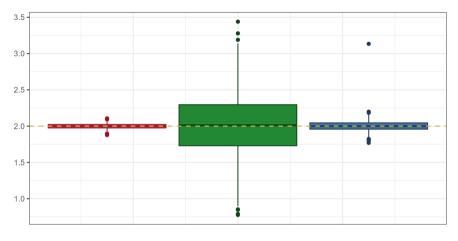


X, Cincorrect

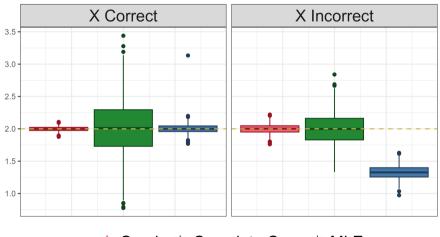




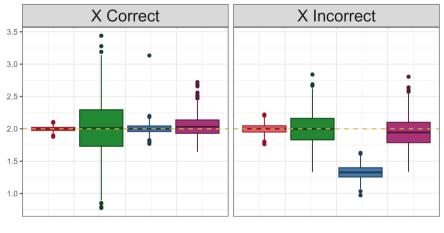
➡ Oracle ➡ Complete Case



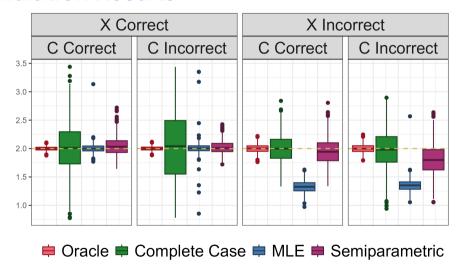
□ Oracle □ Complete Case □ MLE



□ Oracle □ Complete Case □ MLE



➡ Oracle ➡ Complete Case ➡ MLE ➡ Semiparametric



Generalizations

The methods presented here extend to:

• Nonlinear $E(Y|X) = m(X, \beta)$

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The methods presented here extend to:

- Nonlinear $E(Y|X) = m(X, \beta)$
- Additional uncensored covariates Z
 - $\mathrm{E}(Y|X,\mathbf{Z}) = m(X,\mathbf{Z},\boldsymbol{\beta})$
 - Nuisance distributions become $f_{X|\mathbf{Z}}, f_{C|\mathbf{Z}}, f_{\mathbf{z}}$

SPARCC: <u>Semiparametric Censored Covariate</u> Estimation





R package available at https://github.com/brian-d-richardson/sparcc

Thank you! Any questions?

Brian Richardson

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