Robust and efficient estimation in the presence of a randomly censored covariate

Brian Richardson

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Tanya Garcia, PhD



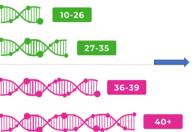
UNC Chapel Hill

This research was supported by the National Institutes of Neurological Disorders and Stroke (grant R01NS131225) and of Environmental Health Sciences (grant T32ES007018)

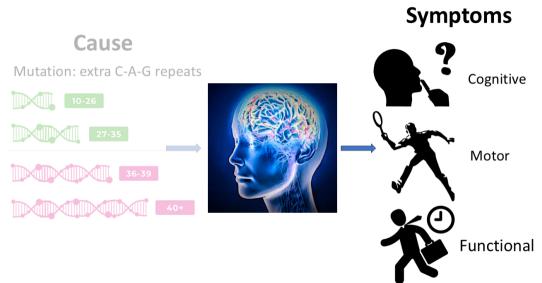


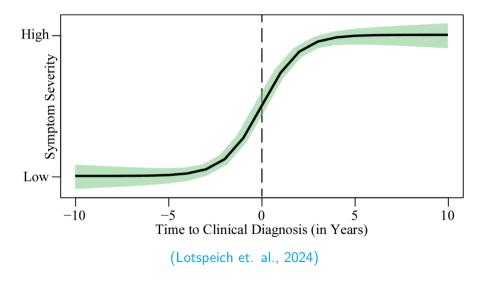
Cause

Mutation: extra C-A-G repeats









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• of interest: parameter β characterizing Y|X

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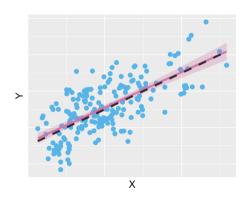
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"Humans are precious"

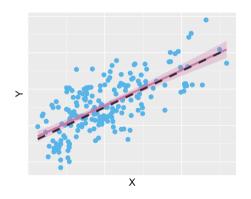
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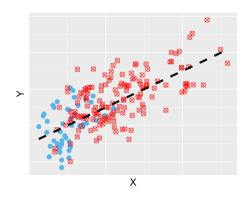
Want methods that are robust and efficient



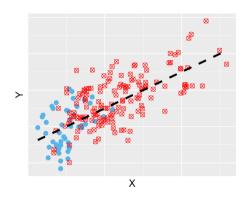
• Regression model: $E(Y|X) = \beta_0 + \beta_1 X$



- Regression model: $E(Y|X) = \beta_0 + \beta_1 X$
- Estimate $\beta = (\beta_0, \beta_1)^T$ with least squares/maximum likelihood
- Solve estimating equation $\sum_{i=1}^{n} (Y_i \beta_0 \beta_1 X_i) (1, X_i)^T = \mathbf{0}$

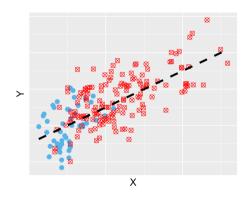


Problem: X is censored by a **random** censoring time C



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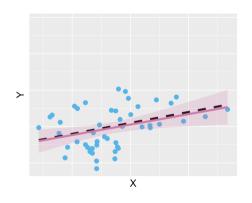
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Problem: X is censored by a **random** censoring time C

- $W = \min(X, C)$
- $\Delta = I(X \leq C)$
- assume: $C \perp \!\!\! \perp (X, Y)$

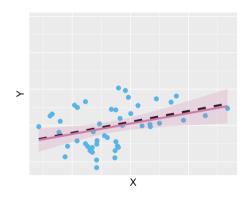
Complete Case Analysis



Only use uncensored observations

Solve estimating equation
$$\sum_{i=1}^{n} \Delta_i (Y_i - \beta_0 - \beta_1 W_i) (1, W_i)^T = \mathbf{0}$$

Complete Case Analysis

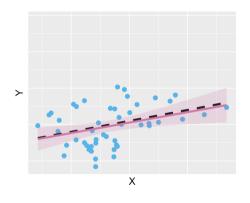


Only use uncensored observations

Solve estimating equation $\sum_{i=1}^{n} \Delta_{i} (Y_{i} - \beta_{0} - \beta_{1} W_{i}) (1, W_{i})^{T} = \mathbf{0}$

✓ Consistent

Complete Case Analysis



Only use uncensored observations

Solve estimating equation $\sum_{i=1}^{n} \Delta_i (Y_i - \beta_0 - \beta_1 W_i) (1, W_i)^T = \mathbf{0}$

- ✓ Consistent
- X Inefficient

$$f_{Y,W,\Delta}(y,w,\delta,\beta,\overline{\alpha}) \propto \underbrace{\{f_{Y|X}(y,w,\beta)\}^{\delta}}_{\text{uncensored}} \underbrace{\left\{\int_{w}^{\infty} f_{Y|X}(y,x,\beta) f_{X}(x,\alpha) dx\right\}^{1-\delta}}_{\text{censored}}$$

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$$\mathbf{S}_{\mathrm{ML}}(y,w,\delta,oldsymbol{eta}) \equiv rac{\partial}{\partialoldsymbol{eta}} \log f_{Y,W,\Delta}(y,w,\delta,oldsymbol{eta},oldsymbol{lpha}), \qquad \sum_{i=1}^{n} \mathbf{S}_{\mathrm{ML}}(Y_i,W_i,\Delta_i,oldsymbol{eta}) = \mathbf{0}$$

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- ✓ consistent
- ✓ fully efficient

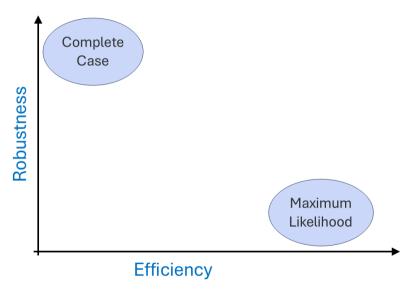
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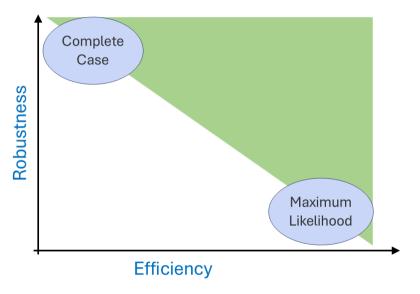
- ✓ consistent
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inconsistent when model for f_X is incorrect

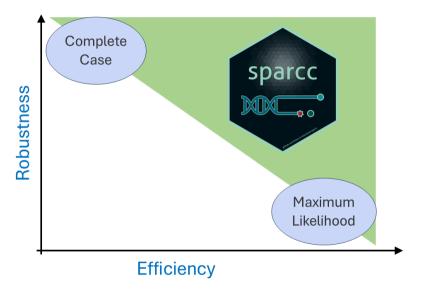
Existing Opportunity



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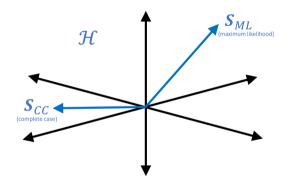
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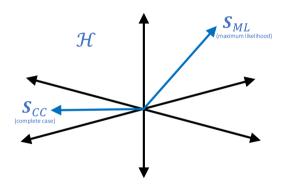
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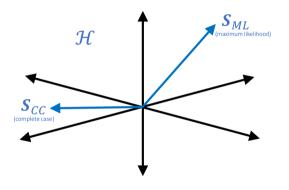
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- goal: derive the SPARCC estimator



• Hilbert space of estimating functions

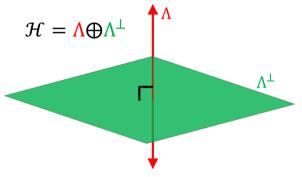


- Hilbert space of estimating functions
- covariance inner product $\langle \boldsymbol{h}, \boldsymbol{g} \rangle \equiv \mathrm{E}(\boldsymbol{h}^T \boldsymbol{g})$



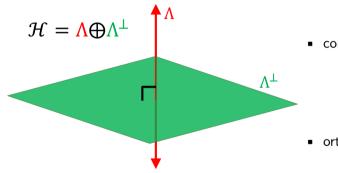
- Hilbert space of estimating functions
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- ullet orthogonal \Leftrightarrow uncorrelated

$$\mathbf{h} \perp \mathbf{g} \iff \langle \mathbf{h}, \mathbf{g} \rangle = 0$$



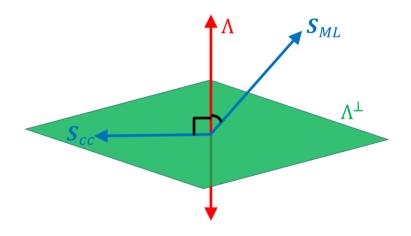
■ construct ∧ using **nuisance scores**

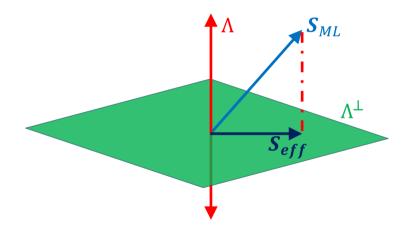
$$\partial \log f_{Y,W,\Delta}(y,w,\delta,\boldsymbol{\beta},\boldsymbol{\eta})/\partial \boldsymbol{\eta}$$



• construct Λ using nuisance scores $\partial \log f_{Y,W,\Delta}(y,w,\delta,\beta,\eta)/\partial \eta$

orthogonal complement Λ[⊥]





Implementing the SPARCC Estimator

The **SPARCC Estimator** $\widehat{\boldsymbol{\beta}}$ is the solution to

$$\sum_{i=1}^n \mathsf{S}_{ ext{eff}}(Y_i, W_i, \Delta_i, eta) = \mathbf{0}$$

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$$\mathbf{S}_{\mathrm{eff}}$$
 requires $\boldsymbol{\eta}=(f_X,f_C)$

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$$\mathbf{S}_{\mathrm{eff}}$$
 requires $\boldsymbol{\eta} = (f_X, f_C)$

Can be modeled either parametrically or nonparametrically

Properties of the SPARCC Estimator

With **parametric** f_X , f_C , $\widehat{\beta}$ is:

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- ✓ **semiparametric efficient** if f_X , f_C are both correctly specified

With **nonparametric** $f_X, f_C, \widehat{\beta}$ is:

- ✓ consistent
- ✓ semiparametric efficient

$$\underbrace{Y}_{\text{outcome}} | X, Z \sim N(\beta_0 + \beta_1 X + \beta_2 Z, \sigma^2)$$

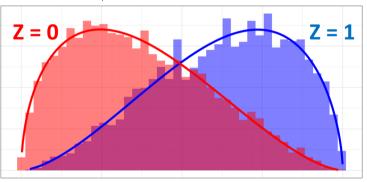
$$Y \mid X$$
, $Z \sim N(\beta_0 + \beta_1 X + \beta_2 Z, \sigma^2)$

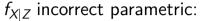
$$Y|X, \underbrace{Z}_{\text{uncensored}} \sim N(\beta_0 + \beta_1 X + \beta_2 Z, \sigma^2)$$

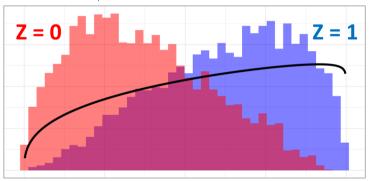
$$Y|X, Z \sim N(\underbrace{\beta_0 + \beta_1 X + \beta_2 Z, \sigma^2}_{\text{parameter of interest: } (\beta_0, \beta_1, \beta_2, \sigma^2)})$$

$$\boldsymbol{\eta} = (f_{X|Z}, f_{C|Z})$$

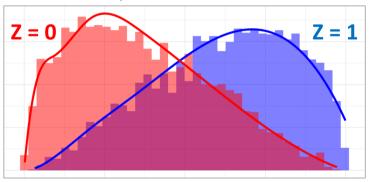
 $f_{X|Z}$ correct parametric:

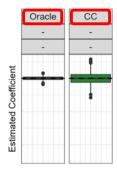


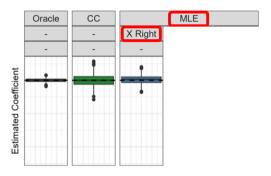


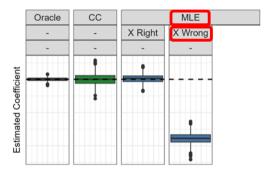


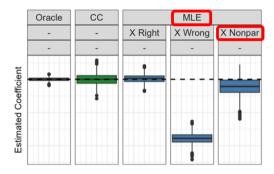
 $f_{X|Z}$ nonparametric:

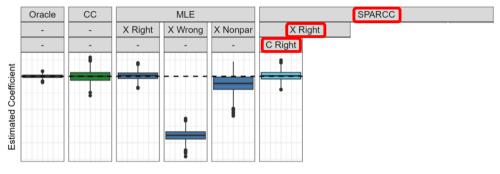


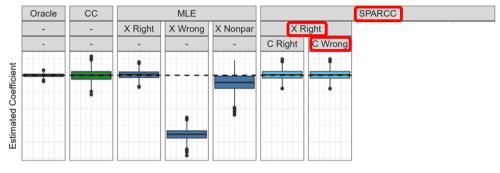


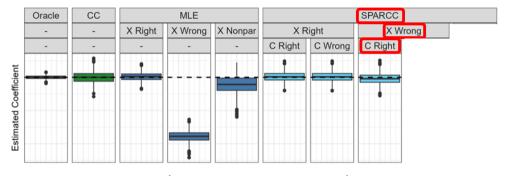


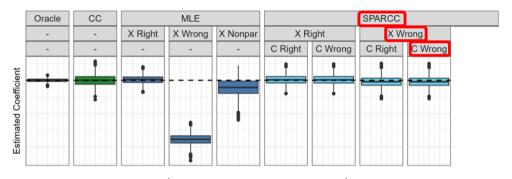


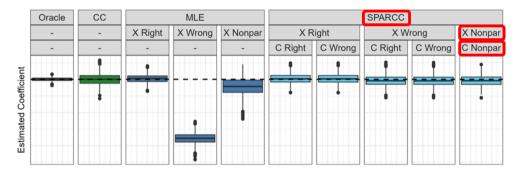




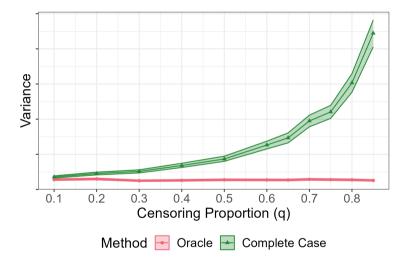




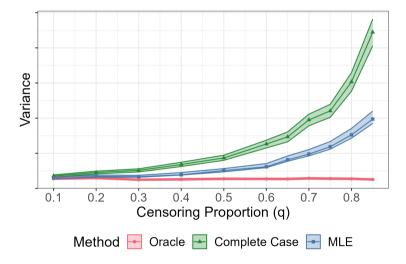




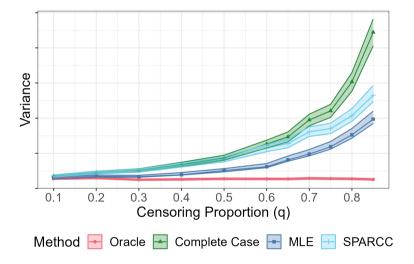
Simulation Results: Efficiency



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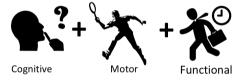


ENROLL-HD Study

• large, observational study of people with Huntington's disease or mutation

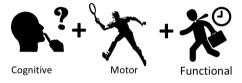
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- composite Unified Huntington Disease Rating Scale (cUHDRS) score



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CAP Score: product of age and mutation severity

Huntington's Disease Application

$$Y$$
_{CUHDRS} $|X, Z \sim N(\beta_0 + \beta_1 X + \beta_2 Z, \sigma^2)$

Huntington's Disease Application

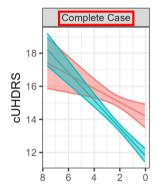
$$Y \mid X \atop \text{time to diagnosis}, Z \sim N(\beta_0 + \beta_1 X + \beta_2 Z, \sigma^2)$$

Huntington's Disease Application

$$Y|X, \underbrace{Z}_{\text{CAP group}} \sim N(\beta_0 + \beta_1 X + \beta_2 Z, \sigma^2)$$

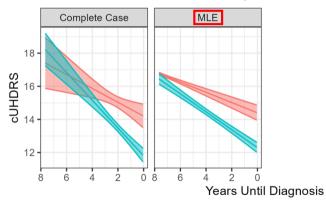
- **sample size**: *n* = 4530
- censoring rate: q = 81.9%

Estimated Mean cUHDRS (and 95% Confidence Intervals)

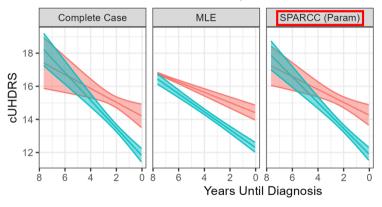


Years Until Diagnosis

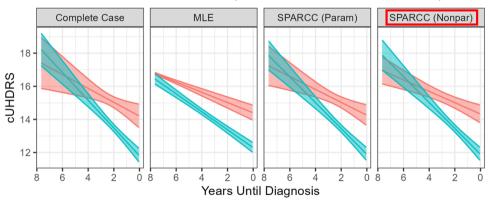
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Proposed **SPARCC** estimator has these properties

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Proposed **SPARCC** estimator has these properties

Future Work:

• implementation and computation for general models

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Proposed **SPARCC** estimator has these properties

Future Work:

- implementation and computation for general models
- longitudinal extension

SPARCC: \underline{S} emiparametric \underline{C} ensored \underline{C} ovariate Estimation



Paper on arXiv





 ${\sf GitHub}\ {\sf R}\ {\sf package}$