



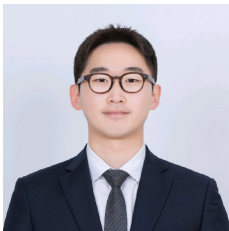
GILLINGS SCHOOL OF  
GLOBAL PUBLIC HEALTH

# Doubly robust estimation under a randomly censored covariate

Brian Richardson

# Acknowledgements

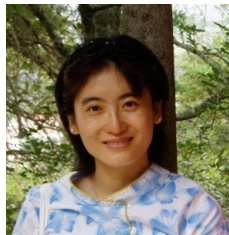
Seong-Ho Lee, PhD



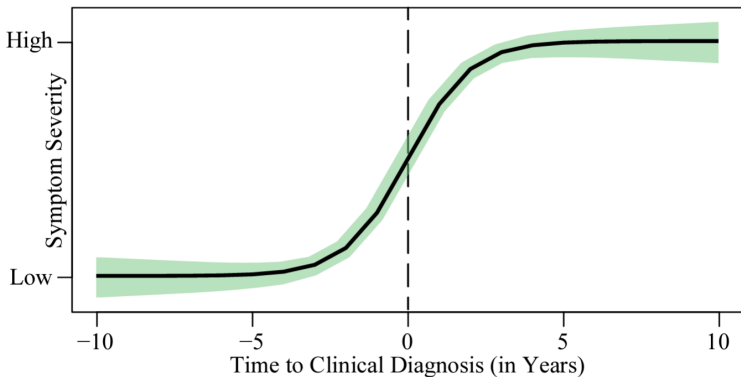
Tanya Garcia, PhD



Yanyuan Ma, PhD

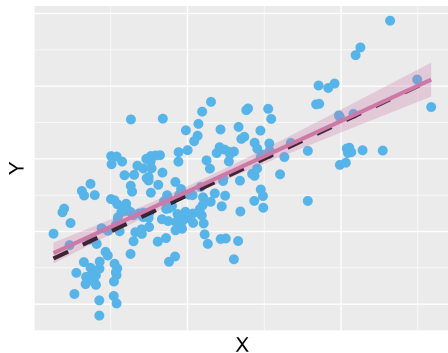


# Huntington's Disease and Censored Covariates



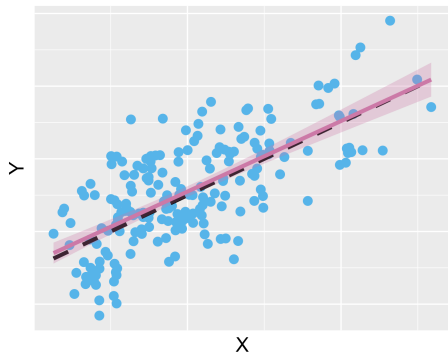
Lotspeich et. al. (2024)

# Censored Covariates: a Simple Example



- Regression model:  
 $E(Y) = \beta_0 + \beta_1 X$

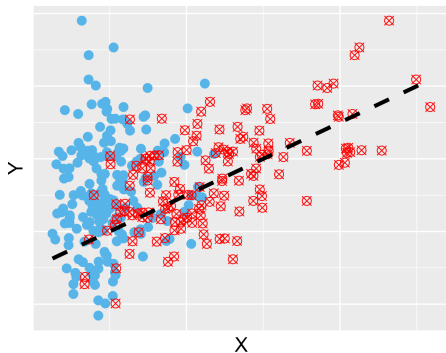
# Censored Covariates: a Simple Example



- Regression model:  
 $E(Y) = \beta_0 + \beta_1 X$
- Estimate  $\beta = (\beta_0, \beta_1)^T$  by solving:

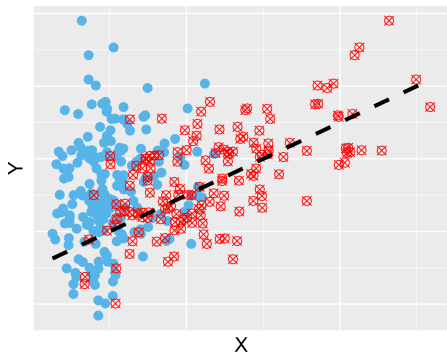
$$\sum_{i=1}^n \underbrace{(Y_i - \beta_0 - \beta_1 X_i)(1, X_i)^T}_{\text{score function: } \mathbf{S}_{\beta}^F} = \mathbf{0}$$

# Censored Covariates: a Simple Example



Problem:  $X$  is censored by a random censoring time  $C$

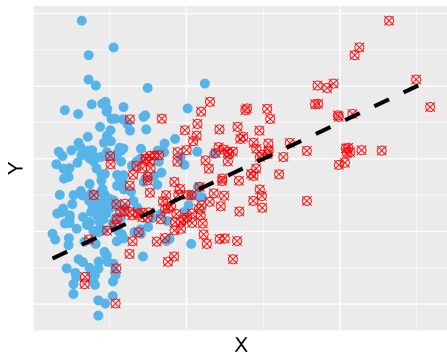
# Censored Covariates: a Simple Example



Problem:  $X$  is censored by a random censoring time  $C$

- $W = \min(X, C)$
- $\Delta = I(X \leq C)$

# Censored Covariates: a Simple Example

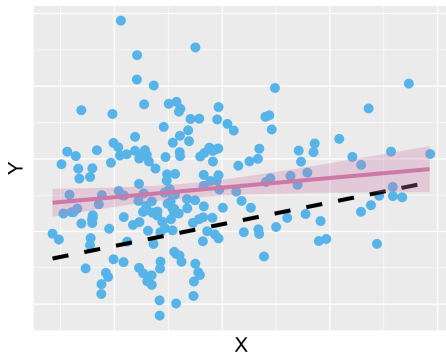


Problem:  $X$  is censored by a random censoring time  $C$

- $W = \min(X, C)$
- $\Delta = I(X \leq C)$
- assume:  $C \perp\!\!\!\perp (X, Y)$

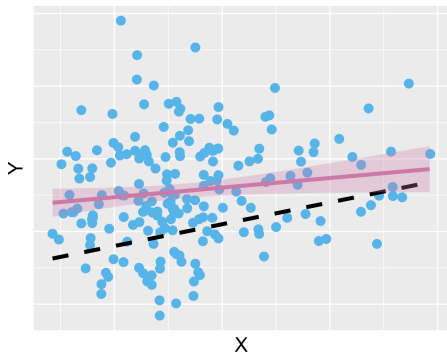


# A Naive Approach



Naively treat  $W$  as  $X$

# A Naive Approach

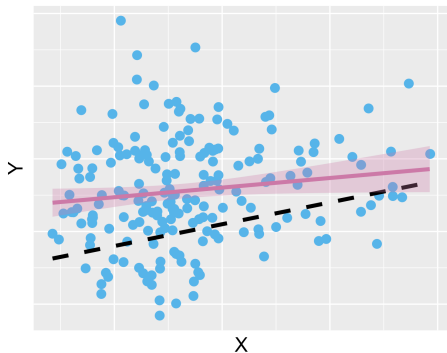


Naively treat  $W$  as  $X$

- Estimate  $\beta$  by solving:

$$\sum_{i=1}^n \mathbf{s}_{\beta}^F(Y_i, \boxed{W_i}, \beta) = \mathbf{0}$$

# A Naive Approach



Naively treat  $W$  as  $X$

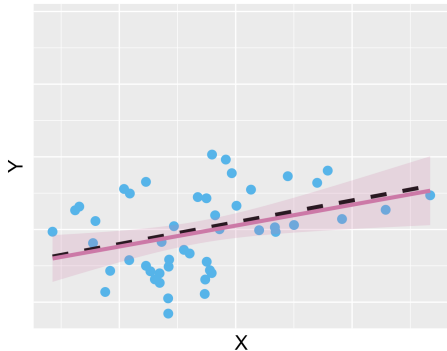
- Estimate  $\beta$  by solving:

$$\sum_{i=1}^n \mathbf{s}_{\beta}^F(Y_i, \boxed{W_i}, \beta) = \mathbf{0}$$

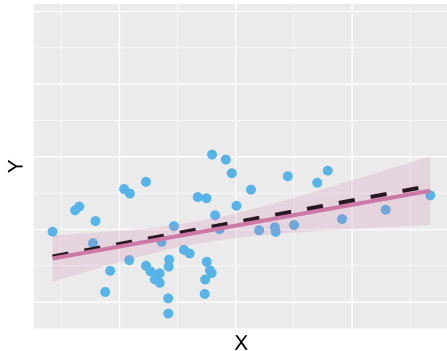
**✗** Inconsistent estimator of  $\beta$

# Complete Case Analysis

Only use observations where  $X$  is *uncensored*



# Complete Case Analysis

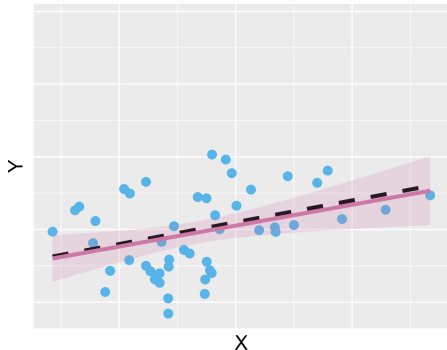


Only use observations where  $X$  is *uncensored*

- Estimate  $\beta$  by solving:

$$\sum_{i=1}^n \boxed{\Delta_i} \mathbf{s}_{\beta}^F(Y_i, W_i, \beta) = \mathbf{0}$$

# Complete Case Analysis



Only use observations where  $X$  is *uncensored*

- Estimate  $\beta$  by solving:

$$\sum_{i=1}^n \boxed{\Delta_i} \mathbf{s}_{\beta}^F(Y_i, W_i, \beta) = 0$$

✓ Consistent

✗ Inefficient

# Maximum Likelihood Estimation (MLE)

$$f_{Y,W,\Delta}(y, w, \delta, \boldsymbol{\beta}, \boldsymbol{\alpha}) \propto \underbrace{\{f_{Y|X}(y, w, \boldsymbol{\beta})\}^\delta}_{\text{uncensored}} \underbrace{\left\{ \int_w^\infty f_{Y|X}(y, x, \boldsymbol{\beta}) f_X(x, \boldsymbol{\alpha}) dx \right\}^{1-\delta}}_{\text{censored}}$$

# Maximum Likelihood Estimation (MLE)

$$f_{Y,W,\Delta}(y, w, \delta, \boldsymbol{\beta}, \boldsymbol{\alpha}) \propto \underbrace{\{f_{Y|X}(y, w, \boldsymbol{\beta})\}^\delta}_{\text{uncensored}} \underbrace{\left\{ \int_w^\infty f_{Y|X}(y, x, \boldsymbol{\beta}) f_X(x, \boldsymbol{\alpha}) dx \right\}^{1-\delta}}_{\text{censored}}$$

$$\mathbf{S}_{\boldsymbol{\beta}}(y, w, \delta, \boldsymbol{\beta}) \equiv \partial \log f_{Y,W,\Delta}(y, w, \delta, \boldsymbol{\beta}, \boldsymbol{\alpha}) / \partial \boldsymbol{\beta}$$



# Maximum Likelihood Estimation (MLE)

$$f_{Y,W,\Delta}(y, w, \delta, \boldsymbol{\beta}, \boldsymbol{\alpha}) \propto \underbrace{\{f_{Y|X}(y, w, \boldsymbol{\beta})\}^\delta}_{\text{uncensored}} \underbrace{\left\{ \int_w^\infty f_{Y|X}(y, x, \boldsymbol{\beta}) \boxed{f_X(x, \boldsymbol{\alpha})} dx \right\}^{1-\delta}}_{\text{censored}}$$

$$\mathbf{S}_\beta(y, w, \delta, \boldsymbol{\beta}) \equiv \partial \log f_{Y,W,\Delta}(y, w, \delta, \boldsymbol{\beta}, \boldsymbol{\alpha}) / \partial \boldsymbol{\beta}$$

$$\sum_{i=1}^n \mathbf{S}_\beta(Y_i, W_i, \Delta_i, \boldsymbol{\beta}) = \mathbf{0}$$

# Maximum Likelihood Estimation (MLE)

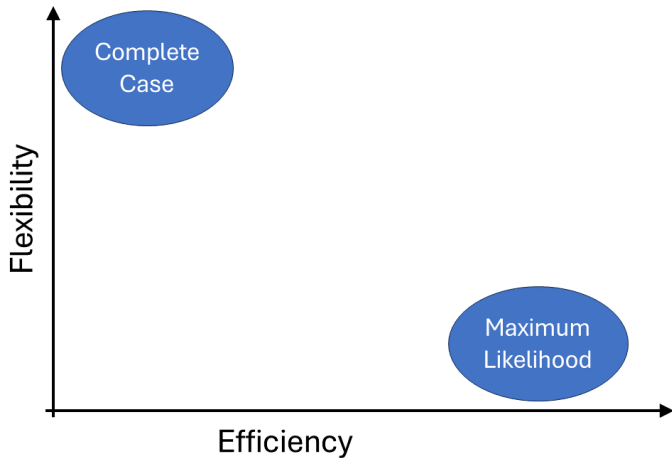
## Pros

- ✓ consistent estimator of  $\beta$
- ✓ fully efficient

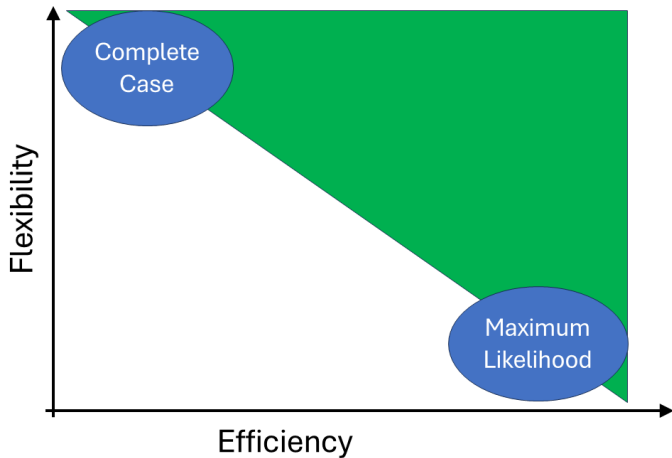
## Cons

- ✗ requires model for **nuisance parameter**  $f_X$
- ✗ inconsistent estimator when model for  $f_X$  is incorrect

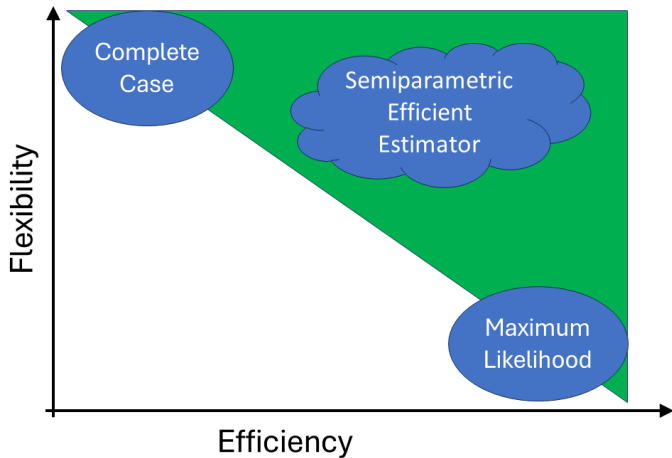
# Existing Methods



# Existing Opportunity



# A New Approach



# The Semiparametric Recipe

- **goal:** to find the estimating function resulting in a semiparametric efficient estimator

# The Semiparametric Recipe

- **goal:** to find the estimating function resulting in a semiparametric efficient estimator
- **semiparametric:** infinite dimensional nuisance parameters  $f_X, f_C$

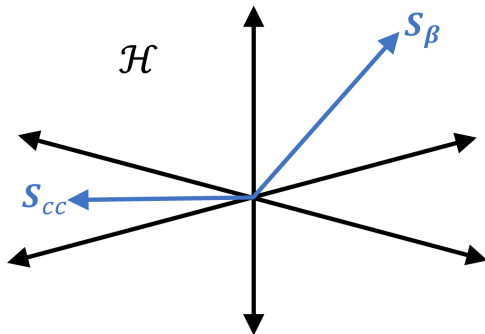
# The Semiparametric Recipe

- **goal:** to find the estimating function resulting in a semiparametric efficient estimator
- **semiparametric:** infinite dimensional nuisance parameters  $f_X, f_C$
- Geometric approach from Tsiatis (2006)



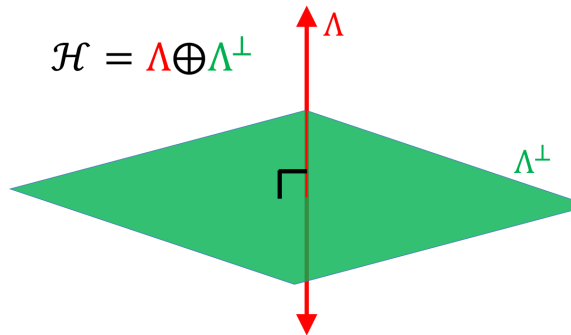
# The Semiparametric Recipe

- **Hilbert space** of estimating functions
- **covariance inner product**  
 $\langle h, g \rangle = E(h^T g)$

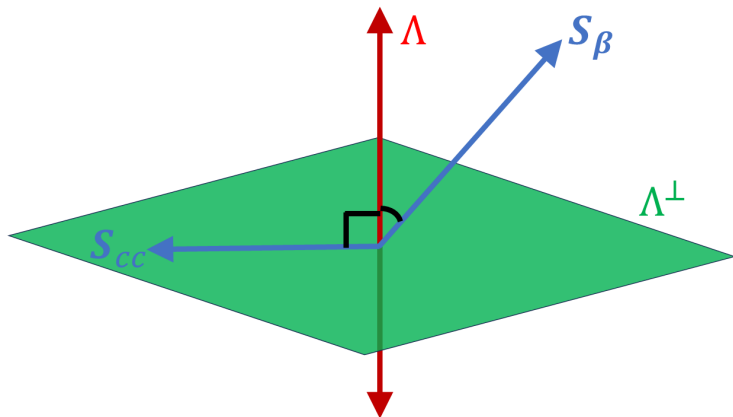


# The Semiparametric Recipe

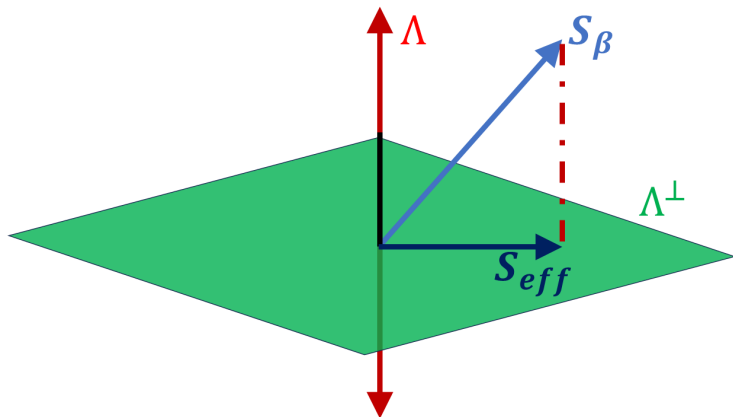
- construct  $\Lambda$  using **nuisance scores**
- orthogonal complement  $\Lambda^\perp$



# The Semiparametric Recipe



# The Semiparametric Recipe



# Properties of the Proposed Estimator

The **semiparametric efficient estimator**  $\hat{\beta}_{\text{eff}}$  is the solution to

$$\sum_{i=1}^n \mathbf{S}_{\text{eff}}^F(Y_i, W_i, \Delta_i, \beta) = \mathbf{0}$$

# Properties of the Proposed Estimator

The **semiparametric efficient estimator**  $\hat{\beta}_{\text{eff}}$  is the solution to

$$\sum_{i=1}^n \mathbf{S}_{\text{eff}}^F(Y_i, W_i, \Delta_i, \beta) = \mathbf{0}$$

- ✓ consistent and asymptotically normal

# Properties of the Proposed Estimator

The **semiparametric efficient estimator**  $\hat{\beta}_{\text{eff}}$  is the solution to

$$\sum_{i=1}^n \mathbf{s}_{\text{eff}}^F(Y_i, W_i, \Delta_i, \beta) = \mathbf{0}$$

- ✓ consistent and asymptotically normal
- ✓ **Doubly Robust:**  $\hat{\beta}_{\text{eff}}$  is consistent if at least one of  $f_X, f_C$  is correctly specified

# Properties of the Proposed Estimator

The **semiparametric efficient estimator**  $\hat{\beta}_{\text{eff}}$  is the solution to

$$\sum_{i=1}^n \mathbf{S}_{\text{eff}}^F(Y_i, W_i, \Delta_i, \beta) = \mathbf{0}$$

- ✓ consistent and asymptotically normal
- ✓ **Doubly Robust:**  $\hat{\beta}_{\text{eff}}$  is consistent if at least one of  $f_X, f_C$  is correctly specified
- ✓ **Locally Efficiency:** If  $f_X, f_C$  are *both* correctly specified, then  $\hat{\beta}_{\text{eff}}$  achieves the **semiparametric efficiency bound**



# Simulation Setup

- $Y|X \sim N(\beta_0 + \beta_1 X, \sigma^2)$

# Simulation Setup

- $Y|X \sim N(\beta_0 + \beta_1 X, \sigma^2)$
- large sample size  $n = 10,000$

# Simulation Setup

- $Y|X \sim N(\beta_0 + \beta_1 X, \sigma^2)$
- large sample size  $n = 10,000$
- high censoring rate  $q = P(X > C) = 0.8$

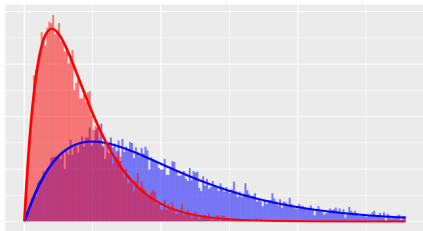
# Simulation Setup

- $Y|X \sim N(\beta_0 + \beta_1 X, \sigma^2)$
- large sample size  $n = 10,000$
- high censoring rate  $q = P(X > C) = 0.8$
- $X, C \sim$  gamma distributions

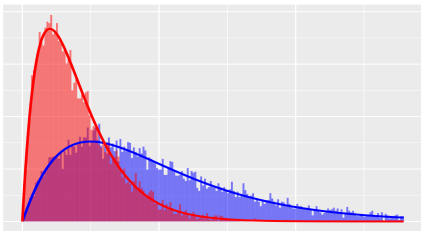
# Simulation Setup

- $Y|X \sim N(\beta_0 + \beta_1 X, \sigma^2)$
- large sample size  $n = 10,000$
- high censoring rate  $q = P(X > C) = 0.8$
- $X, C \sim$  gamma distributions
- $X, C$  possibly **misspecified** as exponential

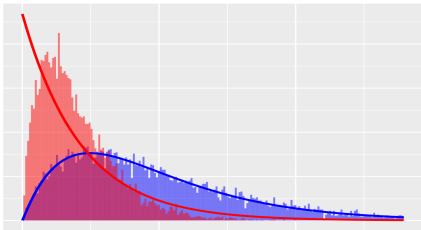
$X$ ,  $C$  correct



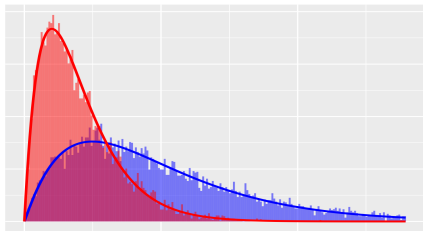
$X$ ,  $C$  correct



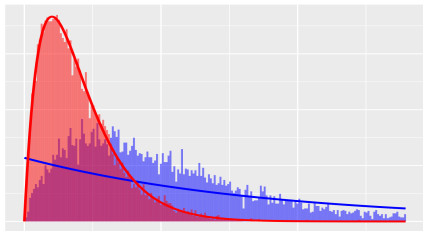
$X$  correct,  $C$  incorrect



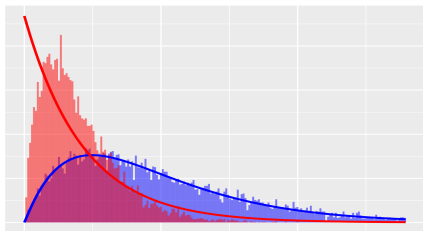
$X$ ,  $C$  correct



$X$  incorrect,  $C$  correct

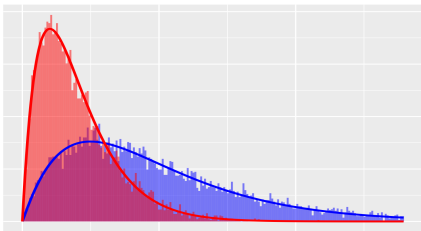


$X$  correct,  $C$  incorrect

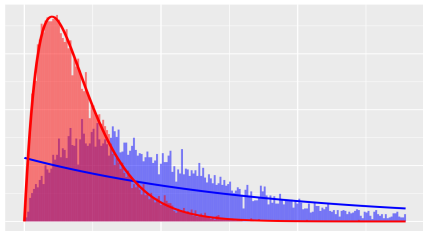




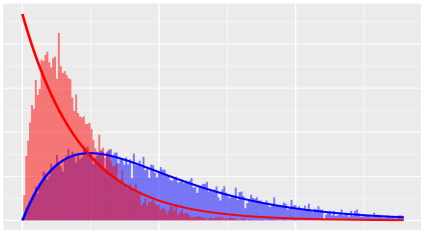
$X$ ,  $C$  correct



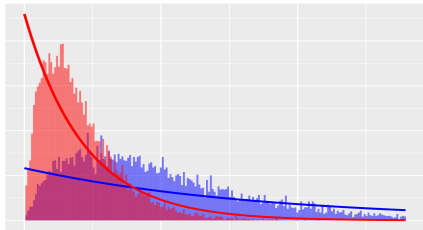
$X$  incorrect,  $C$  correct



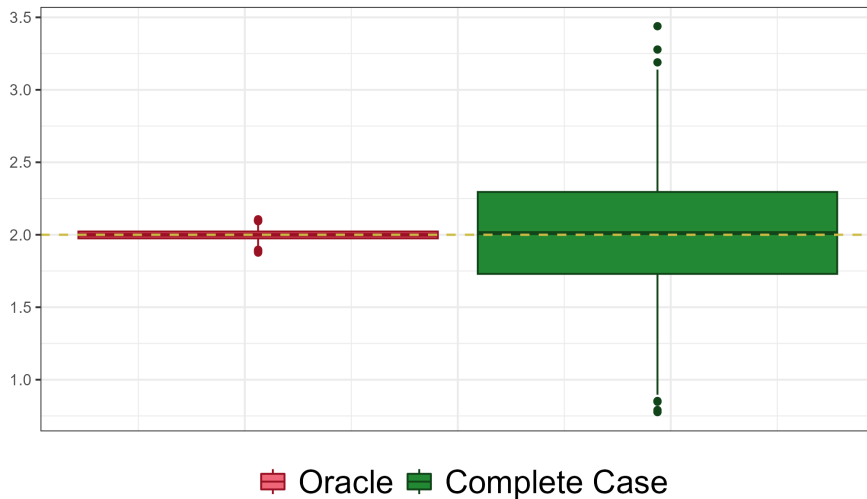
$X$  correct,  $C$  incorrect



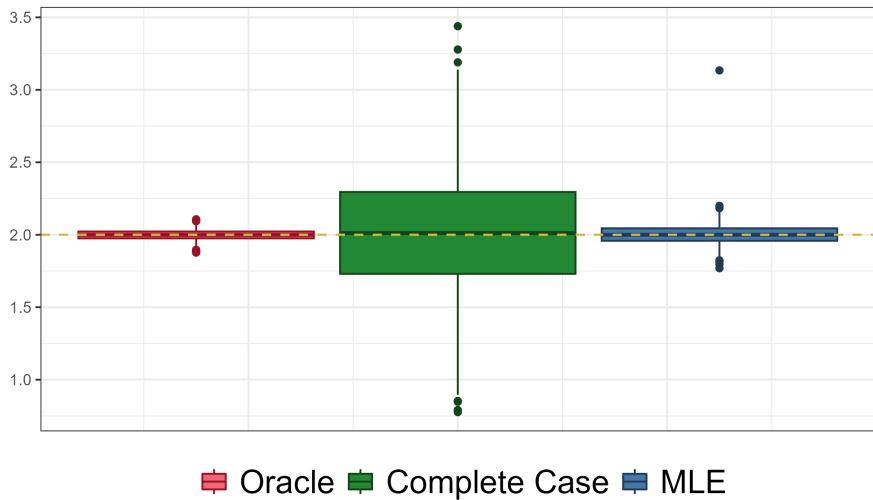
$X$ ,  $C$  incorrect



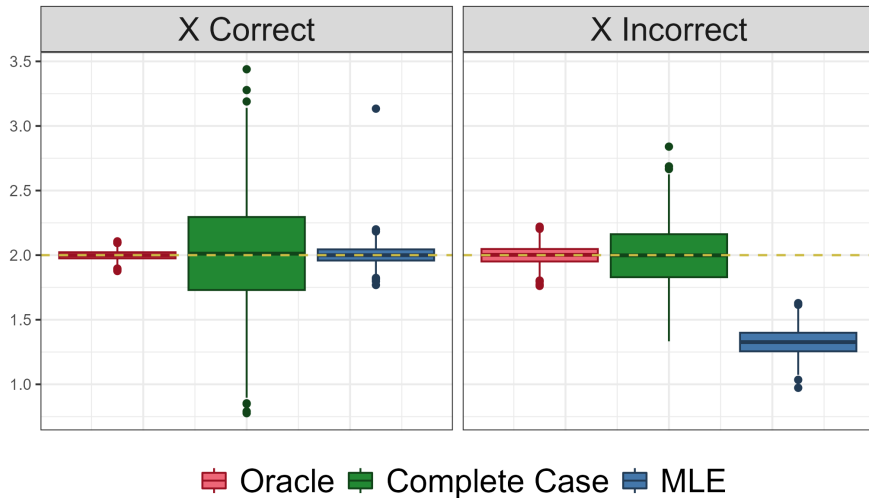
# Simulation Results



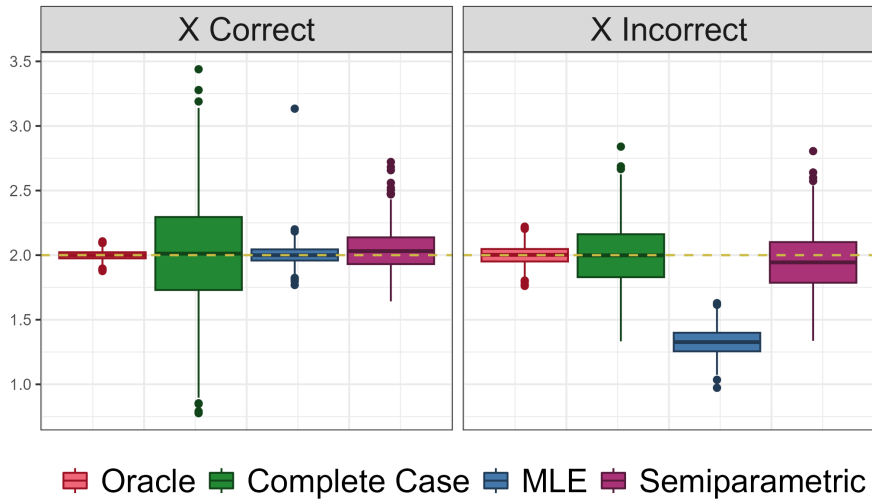
# Simulation Results



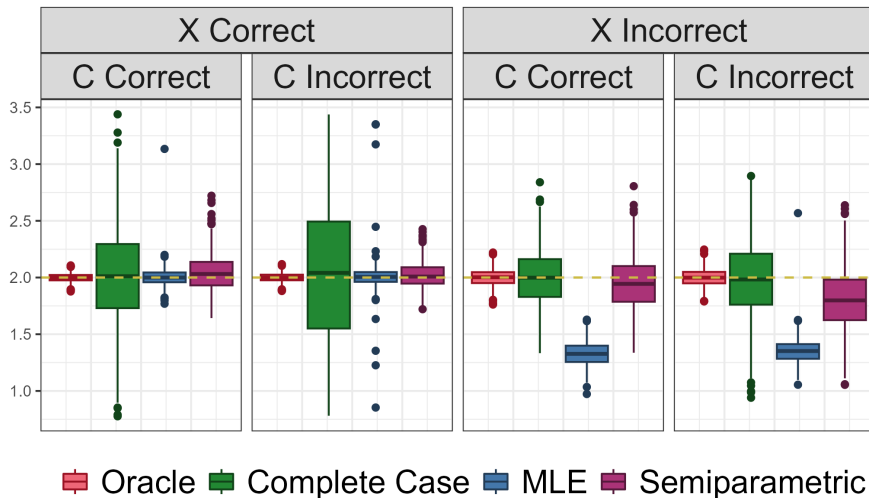
# Simulation Results



# Simulation Results



# Simulation Results



# Generalizations

The methods presented here extend to:

- Nonlinear  $E(Y|X) = m(X, \beta)$

# Generalizations

The methods presented here extend to:

- Nonlinear  $E(Y|X) = m(X, \beta)$
- Additional uncensored covariates  $\mathbf{Z}$ 
  - $E(Y|X, \mathbf{Z}) = m(X, \mathbf{Z}, \beta)$
  - Nuisance distributions become  $f_{X|\mathbf{Z}}, f_{C|\mathbf{Z}}, f_{\mathbf{Z}}$



# SPARCC: Semiparametric Censored Covariate Estimation



R package available at <https://github.com/brian-d-richardson/sparcc>

# Thank you! Any questions?

Brian Richardson

✉: [brichson@ad.unc.edu](mailto:brichson@ad.unc.edu)