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The utility of spatial model-based estimators of unobserved bycatch

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Table S1. Annual bycatch (tonnes) and bycatch rate (percent of hauls) for species selected from the US West Coast Groundfish Observer Program (WCGOP) dataset. All selected species are exclusively discarded. The summarized data are 35 440 post-IFQ hauls (4007 trips) observed in 2011–2015 in the area north of Cape Falcon, Oregon (45.77°N).

	2011		2012		2013		2014		2015	
Species	Catch (t)	% Hauls								
Big skate	25.2	10.2	33.9	10.8	24.1	9.1	68.2	17.9	34.0	18.5
Black skate	18.5	17.3	15.3	14.4	14.0	15.2	13.7	15.3	10.5	13.3
Brown catshark	19.3	45.6	21.5	43.5	24.3	45.4	25.4	45.4	22.9	45.8
California slickhead	9.3	12.3	6.3	8.1	6.4	9.0	5.3	9.3	4.7	6.7
Dungeness crab	120.1	27.6	137.8	32.7	98.2	25.3	105.0	31.9	86.8	30.7
Grenadier	116.8	34.0	121.9	29.8	108.1	29.8	64.0	26.0	42.0	22.5
Octopus	3.7	15.9	2.8	13.2	4.7	15.4	3.4	13.2	2.4	10.9
Pacific hake	147.6	55.1	165.8	58.2	148.0	54.2	122.7	56.2	143.8	60.7
Pacific halibut	61.0	29.3	62.3	30.3	63.7	27.1	53.8	33.9	65.9	36.2
Rosethorn rockfish	0.7	3.3	0.7	4.5	0.9	5.9	0.8	4.2	0.1	2.5
Sandpaper skate	25.9	44.9	33.0	48.4	35.0	51.8	33.9	53.9	34.3	55.4
Slender sole	18.7	20.7	35.2	23.6	46.7	26.9	31.7	31.3	28.2	31.2
Spiny dogfish shark	268.7	42.5	261.4	46.5	258.0	39.2	262.9	46.9	165.5	42.2
Spotted ratfish	50.7	37.5	58.7	42.3	69.0	41.9	57.3	44.4	59.4	48.8
Tanner crab	136.3	46.3	85.1	38.6	104.2	39.7	84.3	39.4	84.9	34.4

Effort = Log(Haul Duration)

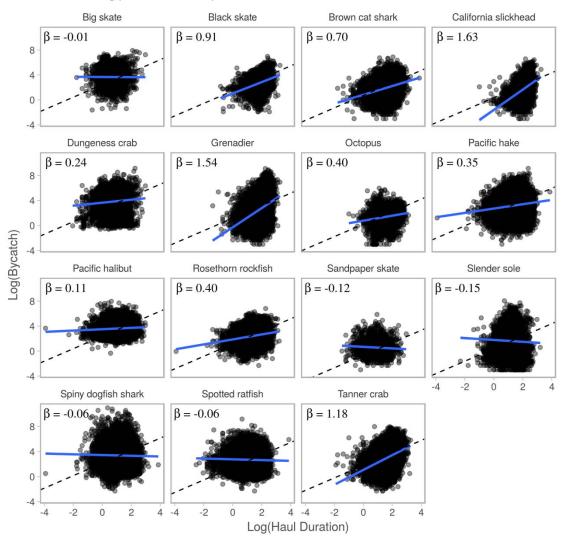


Figure S1. Estimated relationships between fishing effort (haul duration in hours) and bycatch (kg) for 15 species analyzed in the west coast groundfish trawl fishery. The slope terms β of loglog linear models are exponents of an assumed power law fit to each species, Bycatch = α Effort^{β}. Most β are much less than 1, indicating the relationship between bycatch and effort is either weak or not linear. Data ($n = 35\,440$) consist of observed hauls from the West Coast Groundfish Observer Program recorded in 2011–2015 in the area north of Cape Falcon, Oregon (45.77°N).

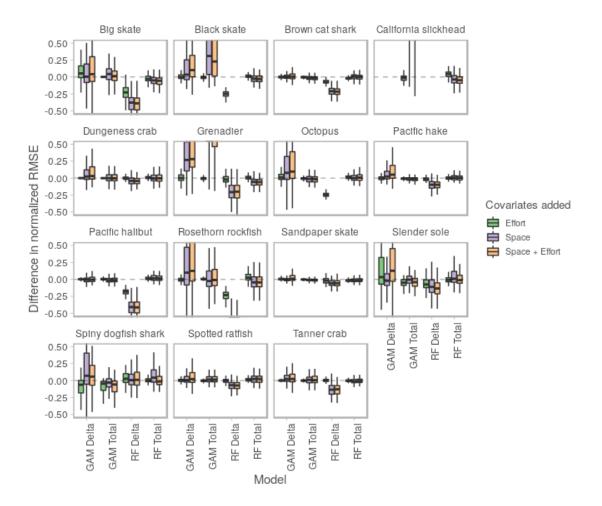


Figure S2. Change in predictive performance (normalized RMSE) when adding fishing effort and spatial location as covariates in each model. For many species, adding space to the GAM-Delta and GAM-Total models led to worse predictions (positive change in RMSE, above dashed line). On the other hand, adding space to the RF-Delta model consistently improved predictions (negative change in RMSE, below dashed line). For RF-Total, including space had either slightly improved predictions or had no effect. Adding effort had little effect for nearly all species and models and never had a larger effect than adding space.

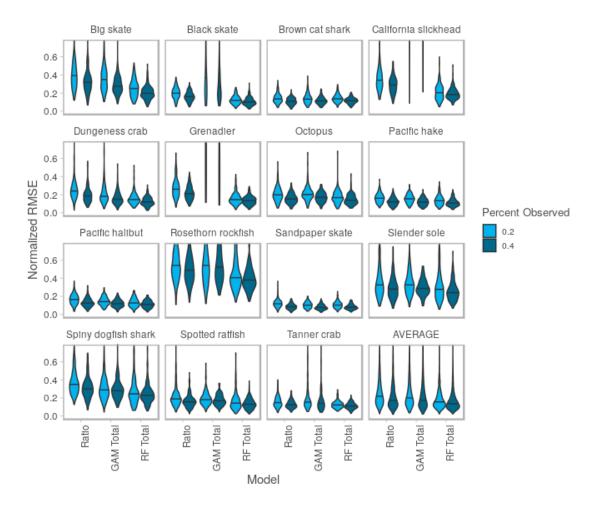


Figure S3. Predictive performance (normalized RMSE) for different levels of simulated observer coverage. Averaged across species, RF-Total had lower median RMSE than the ratio estimator, even at half the observer coverage (RF-Total at 20%: 0.155, Ratio at 40%: 0.180). GAM-Total failed to converge for 3/15 species.

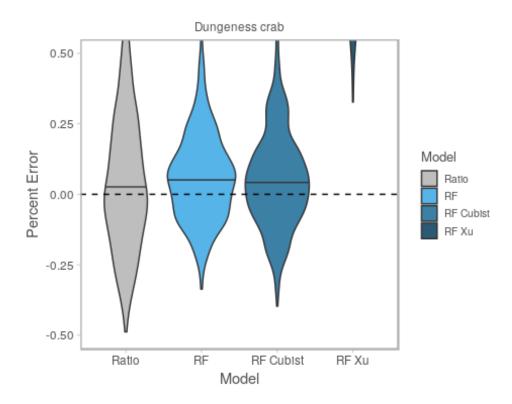


Figure S4. Performance of RF bias correction methods applied to Dungeness crab bycatch (percent error, PE, averaged across years 2011-2015). The ratio estimator is unbiased (median PE = 0.002). RF is positively biased (median PE = 0.055) and Cubist is less positively biased (median PE = 0.043). Cubist reduces bias by fitting a linear model in regression tree terminal nodes instead of using the data mean (Quinlan, 1992, 1993). The second method (Xu, 2013) fits a second RF model to the residuals of the original RF, but this method performed poorly (median PE = 1.107, off chart).

Random forest and GAM models

For each (of 15) species, we fit the models described below. **fac(x)** is shorthand for treating covariate *x* as a factor. All random forest models were fit with **ntree=500** and default **mtry**. All GAMs were fit with **family=Tweedie(p=tweedie_p[i])**, **method="REML"**. All data are at the haul-level.

- *Y*: response, bycatch/discards (in kg)
- *year*: year (5 levels: 2011–2015)
- *depth_interval*: depth of haul (3 levels: < 125, 126–250, > 250 fathoms)
- *season*: season (2 levels: summer, winter)
- bimonth: bimonthly period (6 levels: Jan–Feb, Mar–Apr, . . . , Nov–Dec)
- *logret*: log(retained target species catch in kg)
- *lat*: degrees latitude
- *long*: degrees longitude

For random forest models, we report the percent variance explained [pseudo-R2 = 1-mse Var(y)]. Random forest covariate effect plots were created using the **forestFloor R** package and custom code.

For GAM models, we report the output of **mgcv::summary.gam**, including percent deviance explained. GAM covariate effect plots were created using the **visreg R** package.

RF-Total

Designed to *mimic the stratified ratio estimator* by treating **year**, **season**, and **depth_interval** as factors. **bimonth** is included as linear and quadratic terms to avoid confounding with fac(season):

```
randomForest[Y \sim fac(year) + fac(season) + fac(depth\_interval) + bimonth + I(bimonth^2) + logret + lat + long + I(lat^2) + I(long^2)]
```

GAM-Total

Designed to *mimic the stratified ratio estimator* by treating **year**, **season**, and **depth_interval** as factors. **bimonth** is included as linear and quadratic terms to avoid confounding with **fac(season)**:

```
gam[Y \sim fac(year) + fac(season) + fac(depth interval) + bimonth + I(bimonth^2) + logret + s(lat, long, k=50)]
```

RF-Nonlinear

Allow random forest to fit covariates with *full non-linear flexibility* (not as factors):

```
randomForest(Y \sim year + julian_day + time + depth + gear + logret + lat + long)
```

New covariates introduced:

- *julian_day*: Julian day of year
- *time*: time of day in hours
- *depth*: depth of haul in fathoms

• *gear*: gear type (3 levels: groundfish trawl w/ footrope < 8 inches, groundfish trawl w/ footrope > 8 inches, pineapple trawl)

GAM-Nonlinear

Allow GAM to fit covariates with *full non-linear flexibility* (not as factors):

```
gam[Y \sim s(year, k=5) + s(julian\_day, k=5) + s(time, k=5) + s(depth, k=5) + gear + logret + s(lat, long, k=50)]
```