
Notes on the use of dae for design

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1 Introduction

The R ([R Core Team, 2019](#)) package `dae` ([Brien, 2019b](#)) provides functions useful in the design and anova of experiments. This document describes how to use some of them to produce layouts for experiments and to check some of their properties.

1.1 Functions to be used

The functions in `dae` fall into the following categories and those that will be covered in this document are listed and described:

1. Data

BIBDWheat.dat Data for a balanced incomplete block experiment.

Casuarina.dat Data for an experiment with rows and columns from [Williams et al. \(2002\)](#).

Cabinet1.des A design for one of the growth cabinets in an experiment with 50 lines and 4 harvests.

Exp249.mplot.des Systematic, main-plot design for an experiment to be run in a greenhouse.

Fac4Proc.dat Data for a 2^4 factorial experiment.

LatticeSquare.t49.des A Lattice square design for 49 treatments.

McIntyreTMV.dat The design and data from [McIntyre \(1955\)](#) two-phase experiment.

Oats.dat Data for an experiment to investigate nitrogen response of 3 oats varieties from [Yates \(1937\)](#).

Sensory3Phase.dat Data for the three-phase sensory evaluation experiment in [Brien and Payne \(1999\)](#).

Sensory3PhaseShort.dat Data for the three-phase sensory evaluation experiment in [Brien and Payne \(1999\)](#), but with short factor names.

SPLGrass.dat Data for an experiment to investigate the effects of grazing patterns on pasture composition.

2. Factor manipulation functions

fac.gen: Data for an experiment to investigate nitrogen response of 3 oats varieties.

fac.recode: Recodes the levels and values of a factor.

fac.combine: Combines several factors into one.

fac.divide: Divides a factor into several individual factors.

3. Design functions

designAnatomy: Given the layout for a design, obtain its anatomy via the canonical analysis of its projectors to show the confounding and aliasing inherent in the design.

designLatinSqrSys: Generate a systematic plan for a Latin Square design.

designBlocksGGPlot: Adds block boundaries to a plot produced by `designGGPlot`.

designGGPlot: A graphical representation of an experimental design based on labels stored in a `data.frame` using `ggplot2`.

designRandomize: Takes a systematic design and randomizes it according to the nesting (and crossing) relationships between the recipient(unit) factors for the randomization.

no.reps: Computes the number of replicates for an experiment.

summary.pcanon: Summarizes the anatomy of a design, being the decomposition of the sample space based on its canonical analysis, as produced by `designAnatomy`. The table produced includes the degrees of freedom and summary statistics of the canonical efficiency factors.

4. ANOVA functions

5. Matrix functions

6. Projector and canonical efficiency functions

efficiencies.pcanon: Produces a list containing the canonical efficiency factors for the joint decomposition of two or more sets of projectors (Brien and Bailey, 2009) obtained using `designAnatomy`.

7. Miscellaneous functions.

Documentation for these functions is available from the user manual via `vignette("dae-manual", package="dae")` and there are some notes that show how to use some of them in `vignette("DesignNotes", package="dae")`.

1.2 The paradigm

Fundamental to the approach in this document, and to using the functions described, is that a single allocation involves allocating a set of *allocated* factors to a set of *recipient* factors. In many designs, this allocation is achieved by randomization. However, sometimes there is systematic allocation or restricted allocation.

1.3 Notation used for mixed models

The general form for a mixed model is:

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\mathbf{u} + \mathbf{e},$$

where $\boldsymbol{\beta}$ is the vector of fixed parameters, \mathbf{u} is the vector of random effects, and \mathbf{e} is the vector of residuals corresponding to each observation. The matrices \mathbf{X} and \mathbf{Z} are the design matrices for the fixed and random effects, respectively. Generally, \mathbf{X} and $\boldsymbol{\beta}$ are conformably partitioned so that there is a separate submatrix and subvector for each fixed term. Similarly, \mathbf{Z} and \mathbf{u} are conformably partitioned according to the random terms.

A mixed model is expressed in symbolic form by list of the fixed terms, followed by a '|', and then a list of the random terms. Terms contributing to the residuals are underlined.

2 Single-allocation orthogonal design in R

This class of experiments covers the orthogonal standard or textbook experiments and these experiments must be single phase because they involve a single randomization, in the sense that the randomization can be achieved with a single permutation. Hence there will be two sets of factors, or tiers, one set being allocated to the other set. In `designRandomize`, these are referred to as the allocated and recipient sets of factors. They are also called the unit and treatment factors, respectively.

Firstly, initialize by loading the `dae` library. Also check the version that is loaded.

```
## Loading required package: ggplot2
## [1] '3.1.13'
```

```
options(width=100) @
```

2.1 Two potential designs for a 5×5 grid of plots

Suppose an experiment to investigate five treatments is to be conducted on 25 plots, the 25 plots being arranged in a 5×5 grid. Two possible designs are a randomized complete block design (RCBD) or a Latin square design (LSD). The factor-allocation diagram (Brien et al., 2011) for the RCBD is in Figure 1 and that for the LSD is in Figure 2.

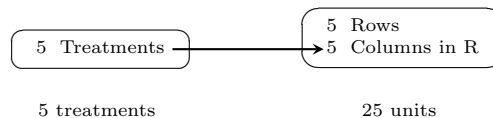


Figure 1: Factor-allocation diagram for an RCBD: treatments are allocated to units; the arrow indicates that the factor Treatments is randomized to Columns; Columns in R indicates that the Columns are considered to be nested within Rows for this randomization; R = Rows.

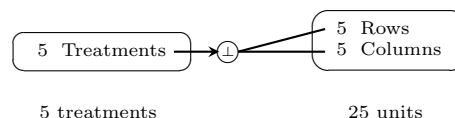


Figure 2: Factor-allocation diagram for an LSD: treatments are allocated to units; the arrow indicates that the allocation is randomized; the '⊥' at the end of the arrow indicates that an orthogonal design is used; the two lines from '⊥' indicates that the Treatments are allocated to the combinations of Rows and Columns using the design.

2.1.1 Produce the randomized layout for an RCBD

Use `designRandomize` to randomize the treatments according to an RCBD. The arguments to `designRandomize` that need to be set are (i) `allocated`, (ii) `nested.recipients`, (iii) `recipient`, and optionally, (iv) `seed`. The allocated factors are also referred to as treatment factors and the recipient factors as block or unit factors. A systematic arrangement of the allocated factors, corresponding to the values of the recipient factors, needs to be supplied and there are a number of ways of doing this.

In these notes, the general approach is to set up a systematic design in a `data.frame` with the values of both the recipient and the allocated factors. The `dae` function `fac.gen` will be used to generate the values of the recipient factors in standard order and often will also be used to generate the values of the allocated factors in an appropriate standard order for the design.

Then the allocated and recipient factors are supplied to `designRandomize` by subsetting the columns of the `data.frames` to just the appropriate factors for each argument. Note that the Treatments could also be supplied as a factor and the recipient factors can be specified directly to the `recipient` argument as a list, e.g. `list(Rows=b, Columns=t)`.

The randomized layout is obtained by permuting (i) Rows and (ii) Columns within Rows. Then the permuted Rows and Columns and the systematic Treatments are sorted so that Rows and Columns are in standard order.

In this example, the allocated factor is Treatments, with 5 levels, and the recipient factors are Rows and Columns, both with 5 levels. Suppose that Rows are to form the blocks.

Use the following R code to obtain and display the layout:

```
### Obtain the layout
b <- 5
t <- 5

### Set up a systematic design
RCBD.sys <- cbind(fac.gen(generate = list(Rows=b, Columns=t)),
                  fac.gen(generate = list(Treatments = LETTERS[1:t]),
                          times = b))

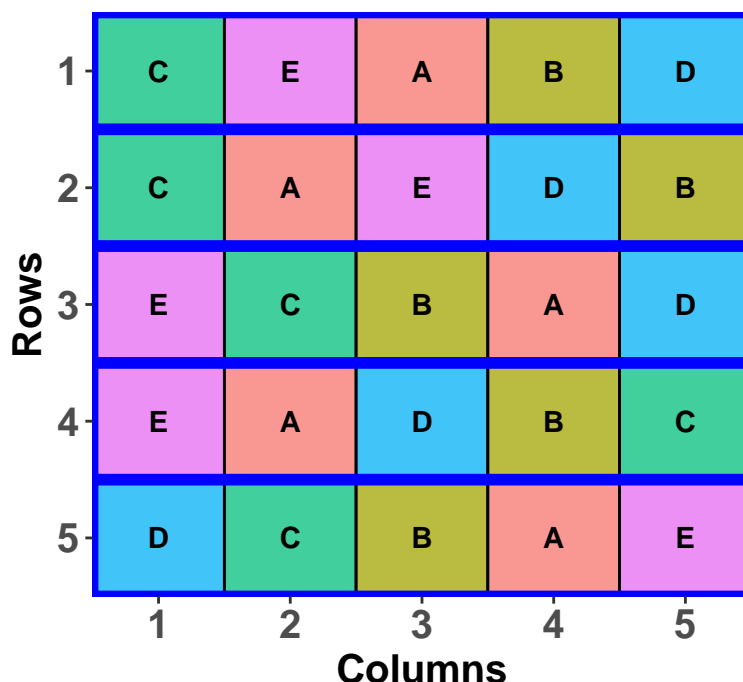
### Obtain the layout
RCBD.lay <- designRandomize(allocated = RCBD.sys["Treatments"],
                           recipient   = RCBD.sys[c("Rows", "Columns")],
                           nested.recipients = list(Columns = "Rows"),
                           seed        = 1134)

### Output the layout
RCBD.lay

##      Rows Columns Treatments
## 1      1      1          C
## 2      1      2          E
## 3      1      3          A
## 4      1      4          B
## 5      1      5          D
## 6      2      1          C
## 7      2      2          A
## 8      2      3          E
## 9      2      4          D
## 10     2      5          B
## 11     3      1          E
## 12     3      2          C
## 13     3      3          B
## 14     3      4          A
## 15     3      5          D
## 16     4      1          E
## 17     4      2          A
## 18     4      3          D
## 19     4      4          B
## 20     4      5          C
## 21     5      1          D
## 22     5      2          C
## 23     5      3          B
## 24     5      4          A
## 25     5      5          E

### Plot the layout
designGGPlot(RCBD.lay, labels = "Treatments", cellalpha = 0.75,
            blockdefinition = cbind(1,t))
```

Plot of Treatments



The function `fac.gen` is from the package `dae` and generates the factors in the `list` in standard order with the specified numbers of levels or the levels in supplied character or numeric vectors. The `seed` is specified to ensure that the same design is produced whenever `designRandomize` is run with these arguments.

2.1.2 Produce the randomized layout for an LSD

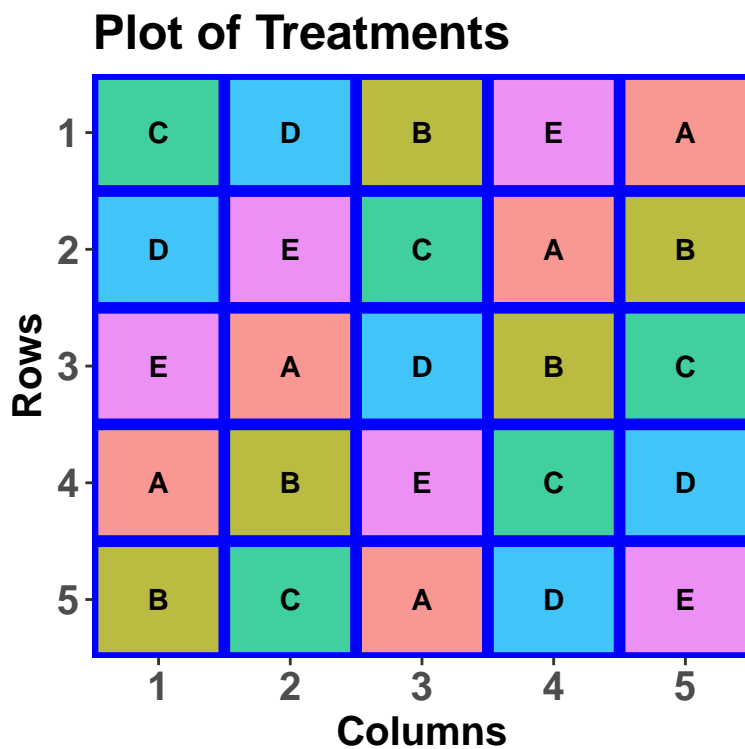
Use `designRandomize` to randomize the treatments according to an LSD, having obtained the systematic design using `fac.gen` and `designLatinSqrSys`. For this design, Rows and Columns are crossed; there are no nested factors. The layout can be obtained using the following R code:

```
b <- 5
t <- 5
### Set up a systematic design
LSD.sys <- cbind(fac.gen(list(Rows=b, Columns=t)),
                 Treatments = factor(designLatinSqrSys(t), labels = LETTERS[1:t]))
### Obtain the layout
LSD.lay <- designRandomize(allocated = LSD.sys["Treatments"],
                          recipient = LSD.sys[c("Rows", "Columns")],
                          seed      = 141)
### Output the layout
LSD.lay

##   Rows Columns Treatments
## 1     1      1          C
## 2     1      2          D
## 3     1      3          B
## 4     1      4          E
## 5     1      5          A
## 6     2      1          D
## 7     2      2          E
```

```
## 8      2      3      C
## 9      2      4      A
## 10     2      5      B
## 11     3      1      E
## 12     3      2      A
## 13     3      3      D
## 14     3      4      B
## 15     3      5      C
## 16     4      1      A
## 17     4      2      B
## 18     4      3      E
## 19     4      4      C
## 20     4      5      D
## 21     5      1      B
## 22     5      2      C
## 23     5      3      A
## 24     5      4      D
## 25     5      5      E

#'## Plot the layout
designGGPlot(LSD.lay, labels = "Treatments", cellalpha = 0.75,
            blockdefinition = cbind(1,1))
```



2.1.3 Check the properties of the designs

The properties of the designs can be investigated using `designAnatomy`.

Because these experiments involve a single randomization, they are two-tiered. That is, there are just two sets of factors involved in the randomization. As we have seen, the first set of factors is the set of allocated (treatment) factors and the second set is the set of recipient (unit) factors. Further there will be a set of projectors

associated with each tier and `designAnatomy` is used to do an eigenanalysis of the relationships between the two sets of projectors. The sets of projectors are specified to `designAnatomy` via model `formulae`, the formula for the recipient factors coming first in the `list` for `formulae`.

For both the RCBD and LSD the two sets of factors are (i) {Rows, Columns} and (ii) {Treatments}. What differs between the two designs is the nesting/crossing relationship between Rows and Columns and this will be expressed in the `formulae`.

Use the commands given below to produce the anatomies (skeleton anova tables) for the RCBD and LSD that have been obtained. Note that the 'Mean' source has been omitted from these tables, but can be included using `grandMean = TRUE` when calling `designAnatomy`.

```
#'## Anatomy for the RCBD
RCBD.canon <- designAnatomy(formulae = list(unit = ~ Rows/Columns,
                                           trt  = ~ Treatments),
                           data      = RCBD.lay)
summary(RCBD.canon)

##
##
## Summary table of the decomposition for unit & trt
##
## Source.unit   df1 Source.trt df2 aefficiency eefficiency order
## Rows          4
## Columns[Rows] 20 Treatments  4      1.0000      1.0000      1
##               Residual    16

#'## Anatomy for the LSD
LSD.canon <- designAnatomy(formulae = list(unit = ~ Rows*Columns,
                                           trt  = ~ Treatments),
                           data      = LSD.lay)
summary(LSD.canon)

##
##
## Summary table of the decomposition for unit & trt
##
## Source.unit   df1 Source.trt df2 aefficiency eefficiency order
## Rows          4
## Columns        4
## Rows#Columns  16 Treatments  4      1.0000      1.0000      1
##               Residual    12
```

Get the mixed-model terms for the analysis by rerunning the `summary` function with the `labels.swap` argument set to `TRUE`.

```
#'## Term-based anatomy for the RCBD
summary(RCBD.canon, labels.swap = TRUE)

##
##
## Summary table of the decomposition for unit & trt
##
## Term.unit     df1 Term.trt   df2 aefficiency eefficiency order
## Rows          4
## Rows:Columns  20 Treatments  4      1.0000      1.0000      1
##               Residual    16
```

```

#### Term-based anatomy for the LSD
summary(LSD.canon, labels.swap = TRUE)

##
##
## Summary table of the decomposition for unit & trt
##
## Term.unit      df1 Term.trt    df2 aefficiency eefficiency order
## Rows          4
## Columns        4
## Rows:Columns  16 Treatments    4      1.0000      1.0000      1
##              Residual      12

```

2.1.4 Questions

1. What is the advantage of specifying a **seed** in **designRandomize**?

It means that the design can be reproduced in subsequent executions of the R script.

2. With what unit term is Treatments confounded in these designs and what is the difference in the interpretation of this term between the two different designs, as demonstrated by the associated sources?

Treatments is confounded with the term Rows:Columns. For the RCBD, this term corresponds to the source Columns[Rows], or Rows within Columns. For the LSD, this term corresponds to the source Rows#Columns, or the interaction of Rows-and-Columns.

3. What would determine which of these two designs is used for a particular experiment?

In a discussion with the researcher, it needs to be determined whether overall Column differences can be ruled out. If they can, then the RCBD should be used; otherwise, the LSD would be used.

2.2 Split-plot from Yates (1937)

Yates (1937) describes a split-plot experiment that investigates the effects of three varieties of oats and four levels of Nitrogen fertilizer. The varieties are assigned to the main plots using a randomized complete block design with 6 blocks and the nitrogen levels are randomly assigned to the subplots in each main plot. The factor-allocation diagram for the experiment is in Figure 3.

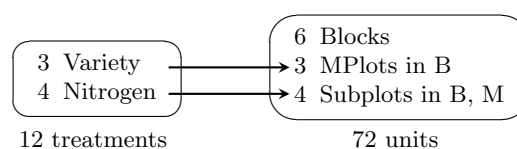


Figure 3: Factor-allocation diagram for a split-plot design: treatments are allocated to units; the arrows indicates that the factors Variety and Nitrogen are randomized to MPlots and Subplots, respectively; MPlots in B indicates that the MPlots are considered to be nested within Blocks for this randomization; Sunplots in B, M indicates that the Subplots are considered to be nested within Blocks and MPlots for this randomization; B = Blocks, M = MPlots

2.2.1 Produce the randomized experimental layout

Use **fac.gen** to obtain a systematic layout and then **designRandomize** to obtain a randomized layout for this experiment. Check the properties of the design, as illustrated in the following R code:

```

Oats.sys <- cbind(fac.gen(list(Blocks=6, MPlots=3, SubPlots=4)),
                 fac.gen(list(Variety=c("Victory", "Golden Rain", "Marvellous"),
                               Nitrogen=c(0,0.2,0.4,0.6)), times=6))
Oats.lay <- designRandomize(allocated = Oats.sys[c("Variety", "Nitrogen")],
                           recipient = Oats.sys[c("Blocks", "MPlots", "SubPlots")],
                           nested.recipients = list(MPlots = "Blocks",
                                                       SubPlots = c("MPlots", "Blocks")),
                           seed = 235805)

### Display design produced
Oats.lay

##      Blocks MPlots SubPlots      Variety Nitrogen
## 1         1      1        1  Marvellous      0.4
## 2         1      1        2  Marvellous      0
## 3         1      1        3  Marvellous      0.2
## 4         1      1        4  Marvellous      0.6
## 5         1      2        1    Victory      0
## 6         1      2        2    Victory      0.2
## 7         1      2        3    Victory      0.6
## 8         1      2        4    Victory      0.4
## 9         1      3        1 Golden Rain      0.2
## 10        1      3        2 Golden Rain      0.4
## 11        1      3        3 Golden Rain      0.6
## 12        1      3        4 Golden Rain      0
## 13        2      1        1  Marvellous      0.4
## 14        2      1        2  Marvellous      0.2
## 15        2      1        3  Marvellous      0
## 16        2      1        4  Marvellous      0.6
## 17        2      2        1    Victory      0.2
## 18        2      2        2    Victory      0
## 19        2      2        3    Victory      0.6
## 20        2      2        4    Victory      0.4
## 21        2      3        1 Golden Rain      0.6
## 22        2      3        2 Golden Rain      0.4
## 23        2      3        3 Golden Rain      0.2
## 24        2      3        4 Golden Rain      0
## 25        3      1        1 Golden Rain      0.2
## 26        3      1        2 Golden Rain      0.6
## 27        3      1        3 Golden Rain      0.4
## 28        3      1        4 Golden Rain      0
## 29        3      2        1  Marvellous      0.4
## 30        3      2        2  Marvellous      0.6
## 31        3      2        3  Marvellous      0
## 32        3      2        4  Marvellous      0.2
## 33        3      3        1    Victory      0.4
## 34        3      3        2    Victory      0.2
## 35        3      3        3    Victory      0
## 36        3      3        4    Victory      0.6
## 37        4      1        1  Marvellous      0
## 38        4      1        2  Marvellous      0.4
## 39        4      1        3  Marvellous      0.2
## 40        4      1        4  Marvellous      0.6
## 41        4      2        1 Golden Rain      0.2
## 42        4      2        2 Golden Rain      0.6

```

```
## 43      4      2      3 Golden Rain      0.4
## 44      4      2      4 Golden Rain      0
## 45      4      3      1      Victory      0.4
## 46      4      3      2      Victory      0
## 47      4      3      3      Victory      0.6
## 48      4      3      4      Victory      0.2
## 49      5      1      1 Golden Rain      0.2
## 50      5      1      2 Golden Rain      0
## 51      5      1      3 Golden Rain      0.6
## 52      5      1      4 Golden Rain      0.4
## 53      5      2      1  Marvellous      0
## 54      5      2      2  Marvellous      0.2
## 55      5      2      3  Marvellous      0.6
## 56      5      2      4  Marvellous      0.4
## 57      5      3      1      Victory      0.4
## 58      5      3      2      Victory      0.2
## 59      5      3      3      Victory      0.6
## 60      5      3      4      Victory      0
## 61      6      1      1  Marvellous      0
## 62      6      1      2  Marvellous      0.6
## 63      6      1      3  Marvellous      0.4
## 64      6      1      4  Marvellous      0.2
## 65      6      2      1      Victory      0.4
## 66      6      2      2      Victory      0.2
## 67      6      2      3      Victory      0
## 68      6      2      4      Victory      0.6
## 69      6      3      1 Golden Rain      0.6
## 70      6      3      2 Golden Rain      0.2
## 71      6      3      3 Golden Rain      0.4
## 72      6      3      4 Golden Rain      0

### Check its properties
Oats.canon <- designAnatomy(formulae = list(unit = ~ Blocks/MPlots/SubPlots,
                                           trt  = ~ Variety*Nitrogen),
                           data      = Oats.lay)
summary(Oats.canon, which.criteria = c("aeff", "order"))

##
##
## Summary table of the decomposition for unit & trt
##
## Source.unit      df1 Source.trt      df2 aefferency order
## Blocks           5
## MPlots[Blocks]   12 Variety          2      1.0000      1
##                  Residual          10
## SubPlots[Blocks:MPlots] 54 Nitrogen      3      1.0000      1
##                  Variety#Nitrogen  6      1.0000      1
##                  Residual          45
```

2.2.2 Analysis of variance (anova) for the Yields

After reading in the data, use the `aov` function to produce the anova as shown below. Note the use of the `Error` function to produce two Residual lines, one each for Wplots and Subplots (Note the change from MPlots to Wplots).

```

#### Read in data for actual experiment
data("Oats.dat")

#### Analyse by anova
oats.aov <- aov(Yield ~ Nitrogen*Variety +
               Error(Blocks/Wplots/Subplots), data=Oats.dat)
summary(oats.aov)

##
## Error: Blocks
##           Df Sum Sq Mean Sq F value Pr(>F)
## Residuals  5  15875     3175
##
## Error: Blocks:Wplots
##           Df Sum Sq Mean Sq F value Pr(>F)
## Variety    2   1786    893.2   1.485  0.272
## Residuals 10   6013    601.3
##
## Error: Blocks:Wplots:Subplots
##           Df Sum Sq Mean Sq F value    Pr(>F)
## Nitrogen    3  20021     6674  37.686 2.46e-12 ***
## Nitrogen:Variety  6    322      54   0.303   0.932
## Residuals   45   7969      177
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

The anova table shown here is the same as the anatomy, but in a different format.

2.2.3 Questions

1. In what sense does this design involve a single randomization?

In the sense that the randomization of both Nitrogen and Variety can be achieved with a single permutation of the units, the subplots.

2. What is the initial allocated mixed model for this design? Is it equivalent to a randomization model?

The initial allocation mixed model is $\text{Variety} + \text{Nitrogen} + \text{Variety}^{\wedge}\text{Nitrogen} \mid \text{Blocks} + \text{Blocks}^{\wedge}\text{MPlots} + \text{Blocks}^{\wedge}\text{MPlots}^{\wedge}\text{SubPlots}$. The initial allocation model is equivalent to a randomization model because the allocation was a randomization.

3. A factorial RCBD would involve randomizing the $3 \times 4 = 12$ treatments to the 12 subplots within each block. What has been achieved in using the split-plot design as compared to a factorial RCBD?

The precision of the Variety differences has been sacrificed to increase the precision of the Nitrogen differences. This is the case because the Residual mean square for $\text{MPlots}[\text{Blocks}]$ is substantially larger than that for $\text{Subplots}[\text{Blocks}^{\wedge}\text{MPlots}]$. If a factorial RCBD had been used, the Residual mean square for $\text{Plots}[\text{Blocks}]$ would be the weighted average of the two Residual mean squares from the split-plot experiment, the weight being the Residual degrees of freedom. That is, the value of the Residual mean square for the factorial RCBD would be between the values for the two Residual mean squares for the split-plot design.

2.3 A design for a petrol additives experiment

Box et al. (2005, Section 4.4) describes a car emission experiment that investigates 4 additives. It involves 4 cars being driven by 4 drivers. Here we investigate increasing the replication by repeating the experiment on two occasions. Suppose that the 4 cars differ between occasions.

In a `data.frame` called `LSRepeat.sys`, generate a systematic design using two 4×4 Latin squares to for allocating the 4 Additives to the 32 tests, being the combinations of the 2 Occasions x 4 Drivers x 4 Cars.

Now a comparison is made of two different ways of randomizing this design. Firstly, we retain the factors Occasions, Drivers and Cars from the systematic design. The factor-allocation diagram is in Figure 4.

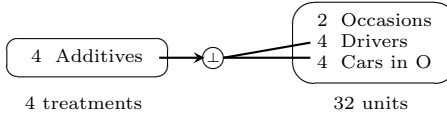


Figure 4: Factor-allocation diagram for repeated LSDs: treatments are allocated to units; the arrow indicates that the allocation is randomized; the ‘⊕’ at the end of the arrow indicates that an orthogonal design is used; the two lines from ‘⊕’ indicates that the Additives are allocated to the combinations of Drivers and Cars within Occasions using the design.

```

## Obtain a randomized layout with Cars nested within Occasions
LSRepeat2b.lay <- designRandomize(allocated = LSRepeat.sys["Additives"],
                                recipient   = LSRepeat.sys[c("Occasions", "Drivers",
                                                             "Cars")],
                                nested.recipients = list(Cars="Occasions"),
                                seed          = 194)

LSRepeat2b.lay

```

##	Occasions	Drivers	Cars	Additives
## 1	1	1	1	B
## 2	1	1	2	A
## 3	1	1	3	D
## 4	1	1	4	C
## 5	1	2	1	C
## 6	1	2	2	B
## 7	1	2	3	A
## 8	1	2	4	D
## 9	1	3	1	D
## 10	1	3	2	C
## 11	1	3	3	B
## 12	1	3	4	A
## 13	1	4	1	A
## 14	1	4	2	D
## 15	1	4	3	C
## 16	1	4	4	B
## 17	2	1	1	C
## 18	2	1	2	B
## 19	2	1	3	A
## 20	2	1	4	D
## 21	2	2	1	D
## 22	2	2	2	C
## 23	2	2	3	B
## 24	2	2	4	A
## 25	2	3	1	A
## 26	2	3	2	D
## 27	2	3	3	C
## 28	2	3	4	B
## 29	2	4	1	B
## 30	2	4	2	A
## 31	2	4	3	D

```
## 32      2      4      4      C

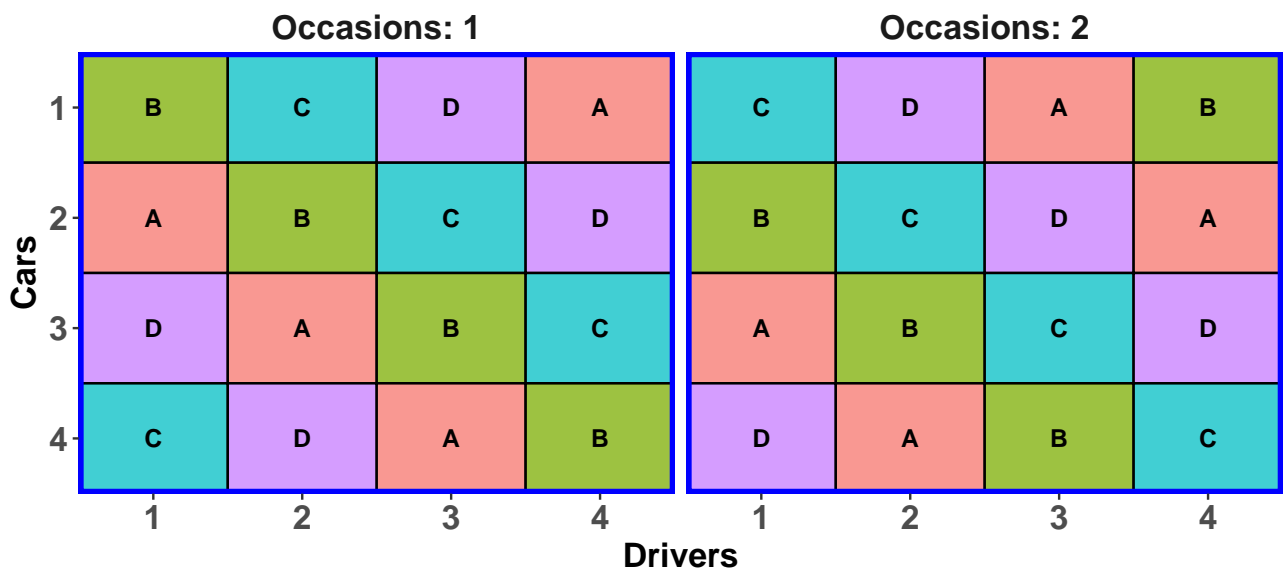
LSRepeat2b.canon <- designAnatomy(formulae = list(unit = ~ (Occasions/Cars)*Drivers,
                                                trt = ~ Additives),
                                data      = LSRepeat2b.lay)

summary(LSRepeat2b.canon)

##
##
## Summary table of the decomposition for unit & trt
##
## Source.unit      df1 Source.trt df2 aeffecticiency eeffecticiency order
## Occasions        1
## Cars[Occasions]  6
## Drivers           3
## Occasions#Drivers 3
## Cars#Drivers[Occasions] 18 Additives    3    1.0000    1.0000    1
##                      Residual    15

### Plot the layout
designGGPlot(LSRepeat2b.lay, row.factors = "Cars", column.factors = c("Occasions", "Drivers"),
            labels = "Additives", cellalpha = 0.75, blockdefinition = cbind(4,4))
```

Plot of Additives



Now we use only Drivers and Cars to do the randomization, but still attempt to include Occasions in the analysis. The new factor-allocation diagram is in Figure 5.

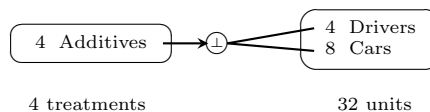


Figure 5: Factor-allocation diagram for repeated LSDs: treatments are allocated to units; the arrow indicates that the allocation is randomized; the '⊕' at the end of the arrow indicates that an orthogonal design is used; the two lines from '⊕' indicates that the Additives are allocated to the combinations of Drivers and Cars using the design.

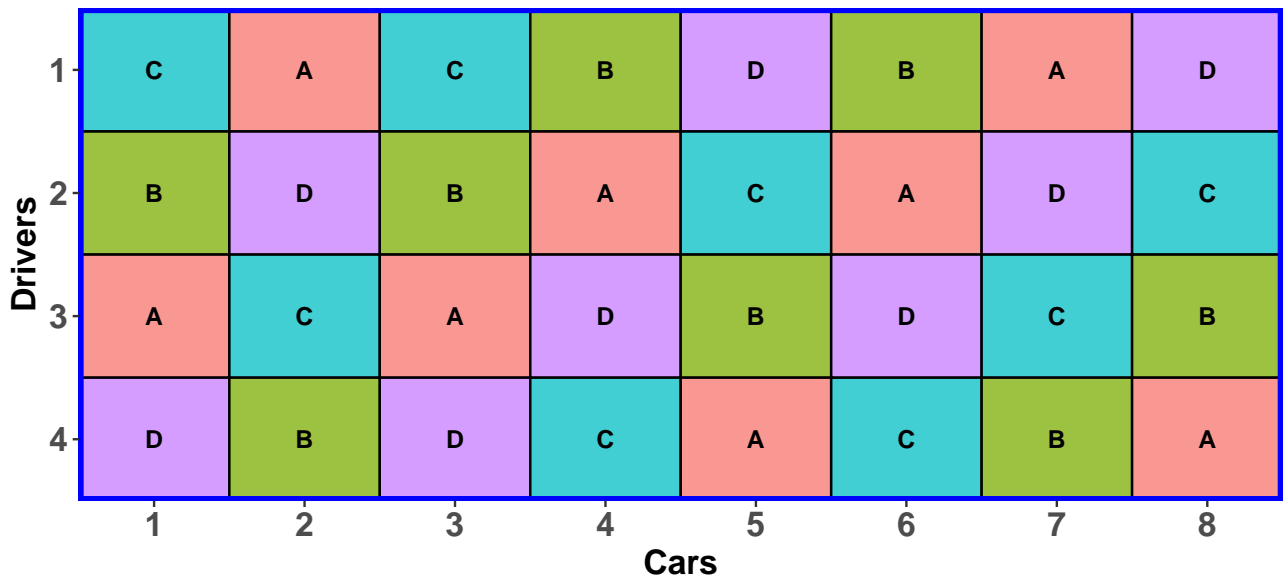
```

#### Obtain a randomized layout
LSRepeat.D8.sys <- LSRepeat.sys
LSRepeat.D8.sys$Cars <- with(LSRepeat.D8.sys, fac.combine(list(Occasions, Cars)))
LSRepeat.D8.sys <- with(LSRepeat.D8.sys, LSRepeat.D8.sys[order(Drivers,Cars),])
LSRepeat2b.D8.lay <- designRandomize(allocated = LSRepeat.D8.sys["Additives"],
                                     recipient = LSRepeat.D8.sys[c("Drivers", "Cars")],
                                     seed       = 149)

#### Plot the layout
designGGPlot(LSRepeat2b.D8.lay, row.factors = "Drivers", column.factors = "Cars",
            labels = "Additives", cellfillcolour.column = "Additives",
            cellalpha = 0.75, blockdefinition = cbind(4,8))

```

Plot of Additives



```

#### Get the Anatomy of the layout
LSRepeat2.D8.canon <- designAnatomy(formulae = list(unit = ~ Drivers*Cars,
                                                    trt   = ~ Additives),
                                   data      = LSRepeat2b.D8.lay)

summary(LSRepeat2.D8.canon)

##
##
## Summary table of the decomposition for unit & trt
##
## Source.unit df1 Source.trt df2 aefficiency eefficiency order
## Drivers      3
## Cars         7
## Drivers#Cars 21 Additives  3      1.0000      1.0000      1
##              Residual  18

#### Add Occasions to the analysis
LSRepeat2b.D8.lay$Occasions <- fac.recode(LSRepeat2b.D8.lay$Cars, rep(1:2, each=4))
LSRepeat2b.D8.lay

```



```
## Drivers Cars Additives Occasions
## 1 1 1 C 1
## 2 1 2 A 1
## 3 1 3 C 1
## 4 1 4 B 1
## 5 1 5 D 2
## 6 1 6 B 2
## 7 1 7 A 2
## 8 1 8 D 2
## 9 2 1 B 1
## 10 2 2 D 1
## 11 2 3 B 1
## 12 2 4 A 1
## 13 2 5 C 2
## 14 2 6 A 2
## 15 2 7 D 2
## 16 2 8 C 2
## 17 3 1 A 1
## 18 3 2 C 1
## 19 3 3 A 1
## 20 3 4 D 1
## 21 3 5 B 2
## 22 3 6 D 2
## 23 3 7 C 2
## 24 3 8 B 2
## 25 4 1 D 1
## 26 4 2 B 1
## 27 4 3 D 1
## 28 4 4 C 1
## 29 4 5 A 2
## 30 4 6 C 2
## 31 4 7 B 2
## 32 4 8 A 2

LSRepeat2b.D8.canon <- designAnatomy(formulae = list(unit = ~ (Occasions + Cars)*Drivers,
                                                    trt = ~ Additives),
                                     data = LSRepeat2b.D8.lay)

summary(LSRepeat2b.D8.canon)

##
##
## Summary table of the decomposition for unit & trt (based on adjusted quantities)
##
## Source.unit df1 Source.trt df2 aeffericiency eeffericiency order
## Occasions 1
## Cars[Occasions] 6
## Drivers 3
## Occasions#Drivers 3 Additives 3 0.1500 0.1250 2
## Cars#Drivers[Occasions] 18 Additives 3 0.8289 0.7500 2
## Residual 15
##
## The design is not orthogonal
```

2.3.1 Questions

1. What is the difference between the two randomizations?

For the first randomization, the Additives are randomized to the Cars within Occasions so that each Driver does all 4 Additives in the 4 Cars in an Occasion. The design is said to be resolved. This does not happen with the randomization based on only Drivers and Cars.

2. How do the two anatomies that include Occasions differ?

The first anatomy is orthogonal and does not have any information about Additives confounded with Cars#Drivers[Occasions]. On the other hand, the second anatomy, based on the layout where Occasions was not included in the randomization, is not orthogonal. Additives information is partially confounded with both Occasions#Drivers and Cars#Drivers[Occasions].

3. What effect does including Occasions#Drivers have on the anatomy?

Including Occasions#Drivers reduces the Residual DF by 3 (from 18 to 15).

2.4 An environmental experiment

Suppose an environmental scientist wants to investigate the effect on the biomass of burning areas of natural vegetation. There are available two areas separated by several kilometres for use in the investigation. It is only possible to either burn or not burn an entire area. The area to be burnt is randomly selected and the other area is to be left unburnt as a control. Further, 30 locations in each area are to be randomly sampled and the biomass measured at each location. The factor-allocation diagram for the experiment is in Figure 6.

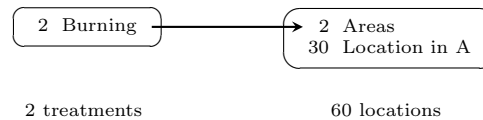


Figure 6: Factor-allocation diagram for the environmental experiment: treatments are allocated to locations; the arrow indicates that the factor Burning is randomized to Areas; Locations in A indicates that the Locations are considered to be nested within Areas; A = Areas.

Obtain the randomized layout for this experiment and check its properties.

```
### Obtain the layout
Burn.sys <- cbind(fac.gen(list(Areas=2, Locations=30)),
                  Burn = factor(rep(c("Burn", "NoBurn"), each=30)))
Burn.layout <- designRandomize(allocated = Burn.sys["Burn"],
                              recipient = Burn.sys[c("Areas", "Locations")],
                              nested.recipients = list(Locations = "Areas"),
                              seed = 872159)

### plot the design
designGGPlot(Burn.layout, labels = "Burn", row.factors = "Locations", column.factors = "Areas")
```

Plot of Burn

Locations	1	Burn	NoBurn
	2	Burn	NoBurn
	3	Burn	NoBurn
	4	Burn	NoBurn
	5	Burn	NoBurn
	6	Burn	NoBurn
	7	Burn	NoBurn
	8	Burn	NoBurn
	9	Burn	NoBurn
	10	Burn	NoBurn
	11	Burn	NoBurn
	12	Burn	NoBurn
	13	Burn	NoBurn
	14	Burn	NoBurn
	15	Burn	NoBurn
	16	Burn	NoBurn
	17	Burn	NoBurn
	18	Burn	NoBurn
	19	Burn	NoBurn
	20	Burn	NoBurn
	21	Burn	NoBurn
	22	Burn	NoBurn
	23	Burn	NoBurn
	24	Burn	NoBurn
	25	Burn	NoBurn
	26	Burn	NoBurn
	27	Burn	NoBurn
	28	Burn	NoBurn
	29	Burn	NoBurn
	30	Burn	NoBurn
		1	2
		Areas	

```

## Check its properties
Burn.canon <- designAnatomy(formulae = list(unit = ~ Areas/Locations,
                                             trt = ~ Burn),
                             data = Burn.lay)

summary(Burn.canon)

##
##
## Summary table of the decomposition for unit & trt
##
## Source.unit      df1 Source.trt df2 aefficiency eefficiency order
## Areas           1 Burn          1    1.0000    1.0000    1
## Locations[Areas] 58

```

2.4.1 Questions

1. How is the pseudo-replication involved in this experiment manifested in the anatomy?

Because (i) Areas and Burn are alongside each other in the anova table, (ii) they both have 1 degree of freedom, and (iii) the single canonical efficiency factor is one, then Areas and Burn are completely confounded. That is, the pseudoreplication has resulted in differences between Areas and between Burns being inextricably mixed up.

2. The randomization-based mixed model for the experiment is $\text{Burn} \mid \text{Areas} + \text{Areas}^{\wedge}\text{Locations}$. What difficulties do you anticipate in attempting to fit this model? How could the model be modified so that a fit can be obtained? [Brien and Demétrio \(2009\)](#) call models formed by removing terms to enable a fit to be achieved ‘models of convenience’. What dangers do you foresee in basing conclusions on the fitted model?

There will be a singularity in the model because Areas is confounded with Burn. A fit could be obtained by removing Areas from the random model. The problem is that a test of Burn would then be based on the ratio of variability in Burn differences to an estimate of the variance of Locations-within-Areas variability. This does not include Areas variability and so the denominator is likely to be underestimated; p-values based from this test are likely to be too small and significant differences are more likely to be declared where there are none as compared to when an estimate of Areas variability is included in the denominator of the F-statistic.

3 Single-allocation, nonorthogonal design in R

This class of experiments covers the nonorthogonal standard or textbook experiments.

3.1 Twenty treatments in an alpha design

The following table gives an alpha design for 20 treatments, taken from [Williams et al. \(2002, p.128\)](#). The design has 3 replicates, each of which contains 5 blocks of 4 plots. It is a resolved design in that each replicate contains a complete set of the treatments.

Table 1: Unrandomized alpha design for 20 treatments

Block	Replicate 1					Replicate 2					Replicate 3				
	1	2	3	4	5	1	2	3	4	5	1	2	3	4	5
	1	2	3	4	5	1	2	3	4	5	1	2	3	4	5
	6	7	8	9	10	7	8	9	10	6	8	9	10	6	7
	11	12	13	14	15	13	14	15	11	12	15	11	12	13	14
	16	17	18	19	20	19	20	16	17	18	17	18	19	20	16

The factor-allocation diagram for the experiment is in Figure 7.

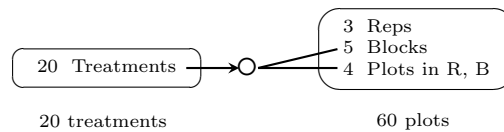


Figure 7: Factor-allocation diagram for the alpha design: treatments are allocated to units; the arrow indicates that the allocation is randomized; the ‘O’ at the end of the arrow indicates that a nonorthogonal design is used; the two lines from ‘O’ indicate that the Treatments are allocated to the combinations of Blocks and Plots using the design; Blocks in R indicates that the Blocks are considered to be nested within Reps for this randomization; Plots in R, B indicates that the Plots are considered to be nested within Reps and Blocks for this randomization; R = Reps; B = Blocks.

3.1.1 Produce the randomized layout for the alpha design and check its properties

Use `designRandomize` to obtain the randomized layout and `designAnatomy` to check its properties.

```
### Set up the systematic design
# Note that Treatments has been entered by rows within a replicate
alpha.sys <- cbind(fac.gen(list(Reps=3, Plots=4, Blocks=5)),
  Treats = factor(c(1:20,
    1:5, 7:10, 6, 13:15, 11, 12, 19, 20, 16:18,
    1:5, 8:10, 6, 7, 15, 11:14, 17:20, 16)))

### Obtain the layout
alpha.layout <- designRandomize(allocated = alpha.sys["Treats"],
  recipient = alpha.sys[c("Reps", "Plots", "Blocks")],
  nested.recipients = list(Blocks = "Reps",
    Plots = c("Reps", "Blocks")),
  seed = 918508)
alpha.layout <- with(alpha.layout, alpha.layout[order(Reps, Blocks, Plots), ])

### Check its properties
alpha.canon <- designAnatomy(formulae = list(units = ~ Reps/Blocks/Plots,
```

```

                                trts = ~ Treats),
                                which.criteria = "all",
                                data          = alpha.lay)
summary(alpha.canon, which.criteria = "all")

##
##
## Summary table of the decomposition for units & trts (based on adjusted quantities)
##
## Source.units      df1 Source.trts df2 aefficiency eefficiency mefficiency sefficiency xefficiency
## Reps              2
## Blocks[Reps]      12 Treats      12      0.2778      0.1667      0.3333      0.0152      0.4167
## Plots[Reps:Blocks] 45 Treats      19      0.7447      0.5833      0.7895      0.0365      1.0000
##                   Residual      26
## order dforthog
##
##      2      0
##      3      7
##
##
## The design is not orthogonal

```

The summary table shows us a number of summary statistics calculated from the canonical efficiency factors. They are:

aefficiency: the harmonic mean of the nonzero canonical efficiency factors.

mefficiency: the mean of the nonzero canonical efficiency factors.

eefficiency: the minimum of the nonzero canonical efficiency factors.

sefficiency: the variance of the nonzero canonical efficiency factors.

xefficiency: the maximum of the nonzero canonical efficiency factors.

order: the order of balance and is the number of unique nonzero canonical efficiency factors.

dforthog: the number of canonical efficiency factors that are equal to one.

For this example it can be seen that (i) an average 74.47%, as measured by the harmonic mean, or 78.95%, as measured by the arithmetic mean, of the information about Treats is confounded with the differences between plots within the reps-blocks combinations and (ii) there are 3 different efficiency factors associated with the 19 Treats degrees of freedom estimated from Plots[Reps:Blocks], the smallest of which is 0.5833 and 7 of which are one. In this case, where the treatments are equally replicated, it can be concluded that the mean variance of a normalized treatment contrast is inversely proportional to the harmonic mean of the canonical efficiency factors, that is, to 0.7447.

Get the mixed-model terms for the analysis by rerunning the summary function with the `labels.swap` argument set to `TRUE`.

```

##'## Obtain the terms for the design
summary(alpha.canon, which.criteria = "all", labels.swap = TRUE)

##
##
## Summary table of the decomposition for units & trts (based on adjusted quantities)
##
## Term.units      df1 Term.trts df2 aefficiency eefficiency mefficiency sefficiency xefficiency

```

```
## Reps                2
## Reps:Blocks         12 Treats      12      0.2778      0.1667      0.3333      0.0152      0.4167
## Reps:Blocks:Plots   45 Treats      19      0.7447      0.5833      0.7895      0.0365      1.0000
##                      Residual      26
## order dforthog
##
##      2      0
##      3      7
##
##
## The design is not orthogonal
```

3.1.2 Questions

1. What is the randomization-based mixed model for this experiment?

The trts term (Source.trts) provides the fixed term and the units terms (Source.units) provide the random terms. That is, Treats is assumed fixed and Repls, Blocks and Plots are assumed random. Hence, the symbolic mixed model is $Treats \mid Repls + Repls \wedge Blocks + \underline{Repls \wedge Blocks \wedge Plots}$.

2. In a mixed-model analysis, which unit terms might you fit as fixed terms? Why?

Repls is a definite candidate for the following reasons. Firstly, Repls has only two degrees of freedom and it will be difficult to estimate a variance component for it. Secondly, one does not want to estimate Treats from Repls (there is no Treats information between Repls).

3.2 Balanced incomplete block design from Joshi (1987)

Joshi (1987) gives an experiment to investigate six varieties of wheat that employs a balanced incomplete block design with 10 blocks, each consisting of three plots. The factor-allocation diagram for the experiment is in Figure 8.

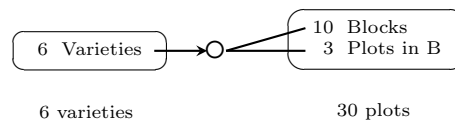


Figure 8: Factor-allocation diagram for the balanced incomplete block design: treatments are allocated to units; the arrow indicates that the allocation is randomized; the ‘O’ at the end of the arrow indicates that a nonorthogonal design is used; the two lines from ‘O’ indicates that the Varieties are allocated to the combinations of Blocks and Plots using the design; Plots in B indicates that the Plots are considered to be nested within Blocks for this randomization; B = Blocks.

3.2.1 Input the Yields and check properties of the design

Use the following R code to input the data for the experiment and check its properties.

```
#### Input the design and data
data("BIBDWheat.dat")

#### Check the properties of the design
bibdwheat.canon <- designAnatomy(formulae = list(units = ~ Blocks/Plots,
                                                trts = ~ Varieties),
                                data      = BIBDWheat.dat)
summary(bibdwheat.canon)
```

```
##
##
## Summary table of the decomposition for units & trts (based on adjusted quantities)
##
## Source.units  df1 Source.trts df2 aeffecticiency eeffecticiency order
## Blocks          9 Varieties    5      0.2000      0.2000      1
##                Residual      4
## Plots[Blocks]  20 Varieties    5      0.8000      0.8000      1
##                Residual     15
##
## The design is not orthogonal
```

From this it is clear that 80% of the information about Varieties is available from the Plots[Blocks] source; that is, 80% of the Varieties information is confounded with differences between plots within blocks. Of course, the remaining 20% is confounded with Blocks.

3.2.2 Anova for the Yields

```
## ## Perform an anova
summary(aov(Yield ~ Varieties + Error(Blocks/Plots), data = BIBDWheat.dat))

##
## Error: Blocks
##          Df Sum Sq Mean Sq F value Pr(>F)
## Varieties  5  196.6   39.32   0.582  0.718
## Residuals  4  270.4   67.59
##
## Error: Blocks:Plots
##          Df Sum Sq Mean Sq F value Pr(>F)
## Varieties  5 1156.4  231.29   4.021 0.0163 *
## Residuals 15  862.9   57.53
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

3.2.3 Questions

1. What is the value of xefficiency for Varieties when confounded with Plots[Blocks] for this design? Why?
It is 0.80 because there is only the one value for the canonical efficiency factor between these two sources
2. How many nonzero eigenvalues does $\mathbf{Q}_V \mathbf{Q}_{BP} \mathbf{Q}_V$ have?
It has 5 nonzero eigenvalues because there is 5 df of Varieties confounded with Plots[Blocks].

3.3 A design with rows and columns from Williams (2002)

Williams et al. (2002, p.144) provides an example of a resolved, Latinized, row-column design with four rectangles (blocks) each of six rows by ten columns. The experiment investigated differences between 60 provenances of a species of Casuarina tree, these provenances coming from 18 countries; the trees were inoculated prior to planting at two different times, time of inoculation being assigned to the four replicates so that each occurred in two replicates. At 30 months, diameter at breast height (Dbh) was measured.

The factor-allocation diagram for the experiment is in Figure 9.

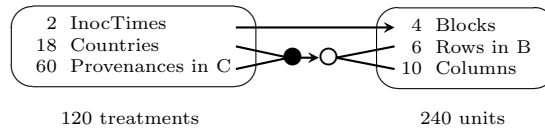


Figure 9: Factor-allocation diagram for the row-and-column design: treatments are allocated to units; the arrows indicates that the allocations are randomized; the two lines leading to the ‘●’ indicate that it is the combinations of Countries and Provenances that is allocated; the ‘○’ at the end of the lower arrow indicates that a nonorthogonal design is used; the two lines from ‘○’ indicates that the Countries and Provenances are allocated to the combinations of Rows and Columns using the design; Rows in B indicates that the Rows are considered to be nested within Blocks for this randomization; B = Blocks.

3.3.1 Input the design and check the properties of the design

Use the following R code to input the design and check its properties.

```
### Input the design
data("Casuarina.dat")
### Check the properties of the design
Casuarina.canon <- designAnatomy(formulae = list(units = ~ (Reps/Rows)*Columns,
                                                trts = ~ InocTime*(Countries+Provenances)),
                                data      = Casuarina.dat)

## Warning in projs.2canon(CombinedSets$Q[[ntiers]], struct[[ktier]]$Q): Provenances[Countries]
## and Countries are partially aliased in Rows[Reps]
## Warning in projs.2canon(CombinedSets$Q[[ntiers]], struct[[ktier]]$Q): Provenances[Countries]
## and Countries are partially aliased in Rows#Columns
## Warning in projs.2canon(CombinedSets$Q[[ntiers]], struct[[ktier]]$Q): Provenances[Countries]
## and Countries are partially aliased in Rows#Columns[Reps]
## Warning in projs.2canon(CombinedSets$Q[[ntiers]], struct[[ktier]]$Q): InocTime#Countries and
## Countries are partially aliased in Rows#Columns[Reps]
## Warning in projs.2canon(CombinedSets$Q[[ntiers]], struct[[ktier]]$Q): InocTime#Countries and
## Provenances[Countries] are partially aliased in Rows#Columns[Reps]
## Warning in projs.2canon(CombinedSets$Q[[ntiers]], struct[[ktier]]$Q): InocTime#Provenances[Countries]
## and Countries are partially aliased in Rows#Columns[Reps]
## Warning in projs.2canon(CombinedSets$Q[[ntiers]], struct[[ktier]]$Q): InocTime#Provenances[Countries]
## and Provenances[Countries] are partially aliased in Rows#Columns[Reps]
## Warning in projs.2canon(CombinedSets$Q[[ntiers]], struct[[ktier]]$Q): InocTime#Provenances[Countries]
## and InocTime#Countries are partially aliased in Rows#Columns[Reps]

summary(Casuarina.canon, which = c("aeff", "eeff", "order", "dforthog"))

##
##
## Summary table of the decomposition for units & trts (based on adjusted quantities)
##
## Source.units      df1 Source.trts              df2 aeffectivity eeffectivity order dforthog
## Reps              3 InocTime                  1      1.0000      1.0000      1          1
##                   Residual                    2
## Rows[Reps]        20 Countries                  17      0.0145      0.0018     17          0
##                   Provenances[Countries]        3      0.1622      0.1326      3          0
## Columns            9 Countries                  9      0.0137      0.0028      9          0
## Reps#Columns       27 Countries                  17      0.0134      0.0012     17          0
##                   Provenances[Countries]        10     0.2320      0.1596     10          0
## Rows#Columns[Reps] 180 Countries                 17      0.7611      0.5588     17          0
##                   Provenances[Countries]        42     0.6851      0.3429     42          0
```

```
##          InocTime#Countries          17          0.6808          0.4735          17          0
##          InocTime#Provenances[Countries] 42          0.5516          0.2009          42          0
##          Residual                      62
##
## Table of (partial) aliasing between sources derived from the same formula
##
## Source                                df Alias                                In                                aefficiency
## Provenances[Countries]                3 Countries                                Rows[Reps]                        0.1622
## Provenances[Countries]               10 Countries                                Reps#Columns                      0.2320
## Provenances[Countries]               42 Countries                                Rows#Columns[Reps]              0.6851
## InocTime#Countries                    17 Countries                                Rows#Columns[Reps]              0.6808
## InocTime#Countries                    17 Provenances[Countries] Rows#Columns[Reps]              0.6808
## InocTime#Provenances[Countries] 42 Countries                                Rows#Columns[Reps]              0.5516
## InocTime#Provenances[Countries] 42 Provenances[Countries] Rows#Columns[Reps]              0.5516
## InocTime#Provenances[Countries] 42 InocTime#Countries Rows#Columns[Reps]              0.5516
## eefficiency order dforthog
##      0.1326      3      0
##      0.1596     10      0
##      0.3429     42      0
##      0.4735     17      0
##      0.4735     17      0
##      0.2009     42      0
##      0.2009     42      0
##      0.2009     42      0
##
## The design is not orthogonal
```

Firstly, note that `designAnatomy` has automatically detected that Provenances is nested within Countries, even though Provenances has 60 unique levels: the sources for these two terms are Countries and Provenances[Countries] and these have 17 and 42 degrees of freedom when estimated from Rows # Columns[Reps], respectively. The total of these degrees of freedom is 59, one less than the number of Provenances, as expected.

Secondly, the partial aliasing evident in this design reflects a lack of (structure) balance between the treatment sources within each units source. This is an undesirable, but unavoidable, feature of the design for this experiment.

3.4 A resolved design for the wheat experiment that is near-A-optimal under a mixed model

Gilmour et al. (1995) provides an example of a wheat experiment for 25 Varieties in which a balanced lattice design was employed.

The factor-allocation diagram for the experiment is in Figure 10.

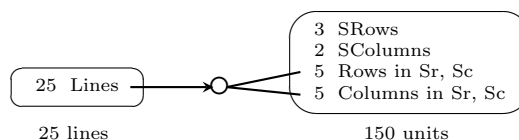


Figure 10: Factor-allocation diagram for the row-and-column design: treatments are allocated to units; the arrows indicates that the allocations are randomized; the ‘O’ at the end of the lower arrow indicates that a nonorthogonal design is used; the two lines from ‘O’ indicates that the Lines are allocated to the combinations of Rows and Columns using the design; Rows (Columns) in Sr, Sc indicates that the Rows (Columns) are considered to be nested within SRows and SColumns for this randomization; Sr = S(upper)Rows; Sc = S(upper)Columns.

3.4.1 Input the design and check the properties of the design

The design is available in the Wheat data set in the asremlPlus package (Brien, 2019a). Use the following R code to input the design, plot it and check its properties.

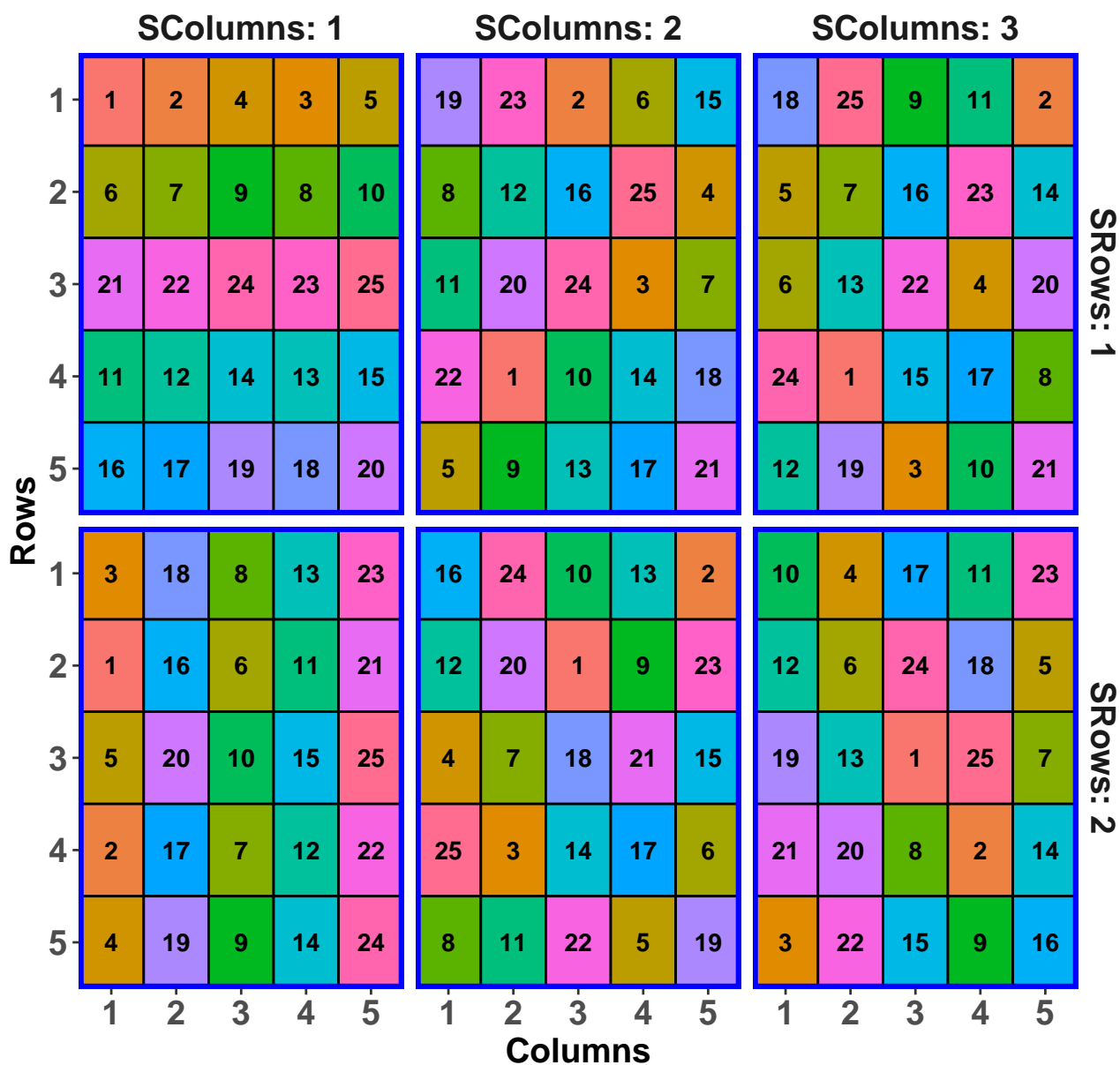
```
### Get the design
library(asremlPlus)

## ASReml-R needs to be loaded if the mixed-model functions are to be used.
##
## ASReml-R is available from VSNi. Please visit http://www.vsnr.co.uk/ for more information.

data(Wheat.dat)
latt.layout <- cbind(fac.gen(list(ARows = 10, AColumns = 15)),
                    fac.gen(list(SRows = 2, Rows = 5, SColumns = 3, Columns = 5)),
                    Wheat.dat["Variety"])

### Plot the design
#+ "LattDesign"
library(scales)
cell.colours <- hue_pal()(25)
designGGPlot(latt.layout, labels = "Variety",
            row.factors = c("SRows", "Rows"), column.factors = c("SColumns", "Columns"),
            blockdefinition = cbind(5,5))
```

Plot of Variety



```
## Check the properties of the design
latt.canon <- designAnatomy(formulae = list(units = ~ (SRows:SColumns)/(Rows*Columns),
                                           trts = ~ Variety),
                           data      = latt.lay)
summary(latt.canon, which.criteria = c("aeff", "order"))

##
##
## Summary table of the decomposition for units & trts (based on adjusted quantities)
##
## Source.units          df1 Source.trts df2 aefficiency order
## SRows:SColumns       5
## Rows[SRows:SColumns] 24 Variety      24      0.1667      1
```

```
## Columns[SRows:SColumns] 24 Variety 24 0.1667 1
## Rows#Columns[SRows:SColumns] 96 Variety 24 0.6667 1
## Residual 72
##
## The design is not orthogonal
```

4 Multiphase experiments in R

This class of experiments differs from those previously presented in that they often employ two or more randomizations or allocations, each to a different type of unit. As a result, there will be three or more sets of factors, or tiers, to deal with; further, when there are three sets of factors, three formula will need to be supplied to `designAnatomy`.

4.1 Athletic examples based on Brien et al. (2011)

Brien et al. (2011) give several designs for an athletic experiment that illustrate the basic principles to be employed in designing multiphase experiments. Here designs for two different multiphase scenarios are considered, both being based on a first-phase that is the testing phase and employs a split-unit design.

4.1.1 A standard single-phase athlete training experiment

First, a split-unit design is generated for an experiment in which the performance of an athlete when subject to nine different training conditions is tested. The nine training conditions are the combinations of three surfaces and three intensities of training. Also, assume that the prime interest is in surface differences, with intensities included to observe the surfaces over a range of intensities. The experiment is to involve 12 athletes, three per month for four consecutive months; each athlete undergoes three tests. The heart rate of the athlete is to be taken immediately upon completion of a test.

A split-plot design is to be employed for the experiment: the three intensities are randomized to the three athletes in each month and the three surfaces are randomized to the three tests that each athlete is to undergo. The factor-allocation diagram is shown in Figure 11. Generate a randomized layout for the experiment.

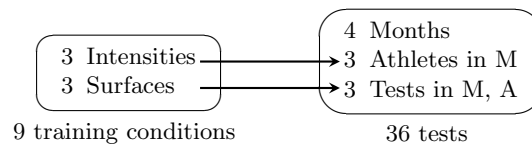


Figure 11: Factor-allocation diagram for the standard athlete training experiment: training conditions are randomized to tests; the two left-hand arrows indicate that the levels of Intensities and Surfaces are randomized to Athletes and Tests, respectively; M = Months; A = Athletes.

```
#'## Phase 1: Construct a systematic layout and generate a randomized layout for the first phase
split.sys <- cbind(fac.gen(list(Months = 4, Athletes = 3, Tests = 3)),
                  fac.gen(list(Intensities = LETTERS[1:3], Surfaces = 3),
                             times = 4))
split.lay <- designRandomize(allocated = split.sys[c("Intensities", "Surfaces")],
                             recipient = split.sys[c("Months", "Athletes", "Tests")],
                             nested.recipients = list(Athletes = "Months",
                                                         Tests = c("Months", "Athletes")),
                             seed = 2598)

split.lay
```

##	Months	Athletes	Tests	Intensities	Surfaces
## 1	1	1	1	B	3
## 2	1	1	2	B	2
## 3	1	1	3	B	1
## 4	1	2	1	C	2
## 5	1	2	2	C	1
## 6	1	2	3	C	3
## 7	1	3	1	A	1
## 8	1	3	2	A	2

```

## 9      1      3      3      A      3
## 10     2      1      1      B      1
## 11     2      1      2      B      2
## 12     2      1      3      B      3
## 13     2      2      1      A      3
## 14     2      2      2      A      2
## 15     2      2      3      A      1
## 16     2      3      1      C      1
## 17     2      3      2      C      3
## 18     2      3      3      C      2
## 19     3      1      1      B      1
## 20     3      1      2      B      3
## 21     3      1      3      B      2
## 22     3      2      1      C      2
## 23     3      2      2      C      3
## 24     3      2      3      C      1
## 25     3      3      1      A      2
## 26     3      3      2      A      3
## 27     3      3      3      A      1
## 28     4      1      1      A      3
## 29     4      1      2      A      2
## 30     4      1      3      A      1
## 31     4      2      1      B      1
## 32     4      2      2      B      2
## 33     4      2      3      B      3
## 34     4      3      1      C      1
## 35     4      3      2      C      3
## 36     4      3      3      C      2

### Get anatomy to check properties of the design
split.canon <- designAnatomy(formulae = list(tests = ~ Months/Athletes/Tests,
                                             cond  = ~ Intensities*Surfaces),
                             data      = split.lay)
summary(split.canon, which.criteria="none")

##
##
## Summary table of the decomposition for tests & cond
##
## Source.tests      df1 Source.cond      df2
## Months            3
## Athletes[Months]  8 Intensities        2
##                   Residual            6
## Tests[Months:Athletes] 24 Surfaces        2
##                   Intensities#Surfaces  4
##                   Residual            18

### Plot the design
#+ "SplitDes_v2"
split.lay <- within(split.lay,
                    Conditions <- fac.combine(list(Intensities, Surfaces),
                                                combine.levels = TRUE))
plt <- designGGPlot(split.lay, labels = "Conditions",
                    row.factors = "Tests", column.factors = c("Months", "Athletes"),

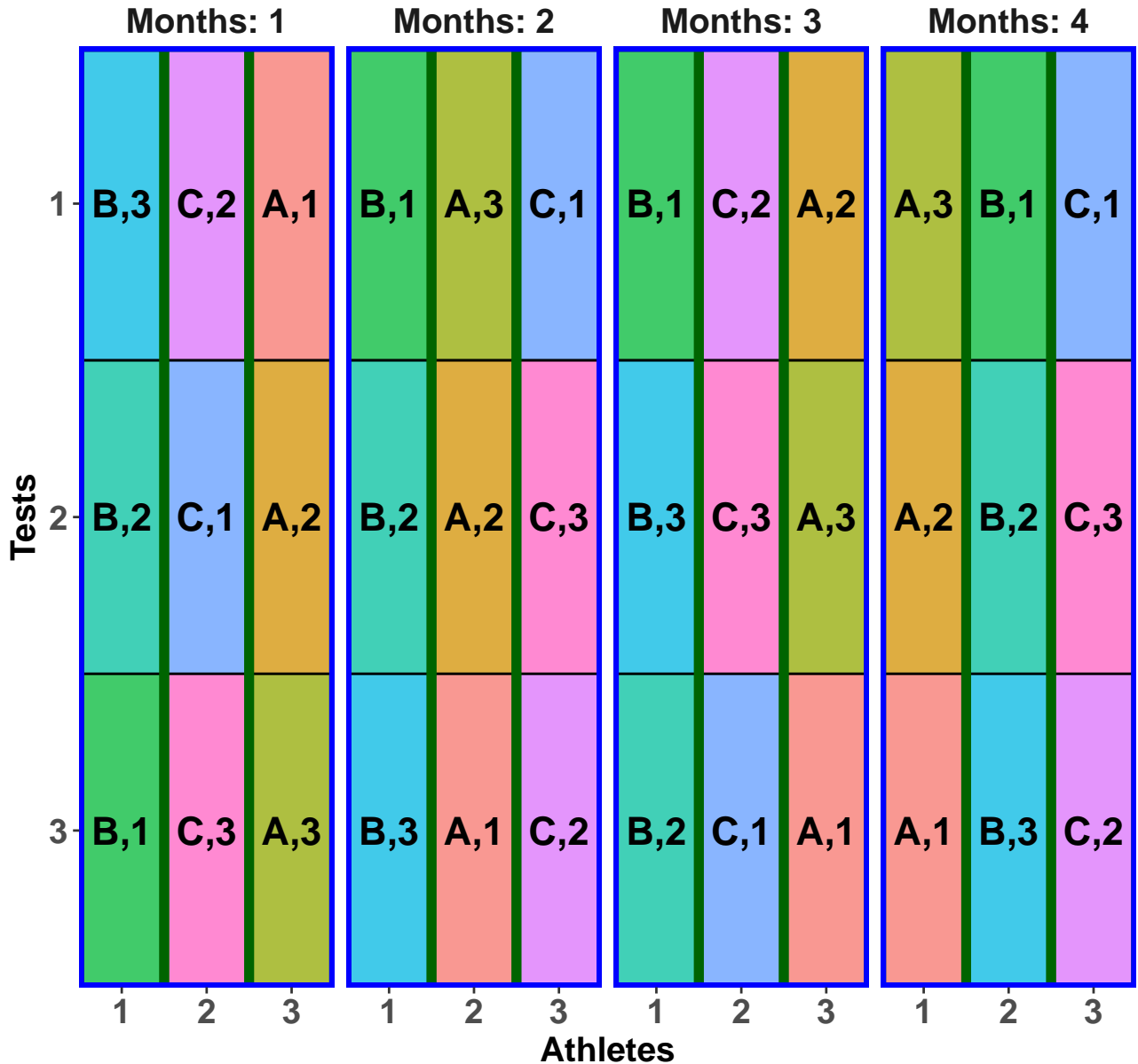
```

```

cellalpha = 0.75, size = 6,
blockdefinition = rbind(c(3,1)), blocklinecolour = "darkgreen",
printPlot = FALSE)
designBlocksGGPlot(plt, nrows = 3, ncolumns = 3, blockdefinition = rbind(c(3,3)))

```

Plot of Conditions



Question

1. Why was a split-plot design chosen for this experiment?

Because it is likely that variation between tests within an athlete will be smaller than variation between athletes within a month. Hence, because the prime interest is in Surfaces, they are assigned to tests within an athlete and will have better precision than Intensities, which have been assigned to the more variable athletes within a month.

4.1.2 A simple two-phase athlete training experiment

Suppose that, in addition to heart rate taken immediately upon completion of a test, the free haemoglobin is to be measured using blood specimens taken from the athletes after each test and transported to the laboratory for analysis. That is, a second laboratory phase is required to obtain the new response. In this phase, because the specimens become available monthly, the batch of specimens for one month are to be processed, in a random order, before those for the next month are available. The factor-allocation diagram for this experiment is in Figure 12, the dashed line indicating that Months are systematically allocated to Batches. The randomizations in this diagram are composed (Brien and Bailey, 2006) and is one of the two types of randomizations in a chain (Bailey and Brien, 2015). This means that the second-phase randomization only need to consider how the tests factors are to be assigned to locations; training conditions can be ignored in determining the second-phase design.

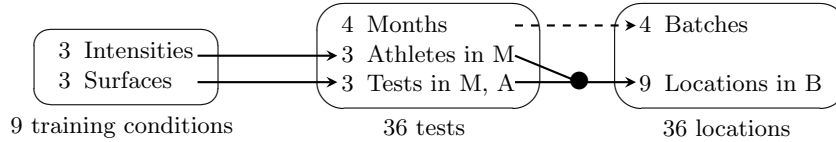


Figure 12: Factor-allocation diagram for the two-phase athlete training experiment: training conditions are randomized to tests and tests are allocated to locations; the two left-hand arrows indicate that the levels of Intensities and Surfaces are randomized to Athletes and Tests, respectively; the dashed arrow indicates that Months are systematically allocated to Batches; the '●' indicates that the combinations of the levels of Athletes and Tests are randomized to the Locations; M = Months; A = Athletes; B = Batches.

Using the following R code, obtain a layout for the second phase and check the properties of the layout. In doing this, the first-phase layout is randomized. However, because Months is not randomized to Batches, the argument `except` in `designRandomize` is used to effect the systematic allocation.

```

##' Generate a layout for a simple two-phase athlete training experiment
##'
##'## Phase 1 - the split-plot design that has already been generated.
##'## Phase 2 - randomize tests (and training conditions) to locations,
##'##           but Months assigned systematically to Batches
##'##           so except Batches from the randomization
eg1.lay <- designRandomize(allocated = split.lay,
                           recipient = list(Batches = 4, Locations = 9),
                           nested.recipients = list(Locations = "Batches"),
                           except = "Batches",
                           seed = 71230)

eg1.lay

##      Batches Locations Months Athletes Tests Intensities Surfaces Conditions
## 1          1          1      1         2     3           C         3      C,3
## 2          1          2      1         1     2           B         2      B,2
## 3          1          3      1         2     2           C         1      C,1
## 4          1          4      1         3     1           A         1      A,1
## 5          1          5      1         3     2           A         2      A,2
## 6          1          6      1         1     1           B         3      B,3
## 7          1          7      1         2     1           C         2      C,2
## 8          1          8      1         1     3           B         1      B,1
## 9          1          9      1         3     3           A         3      A,3
## 10         2          1      2         3     1           C         1      C,1
## 11         2          2      2         2     2           A         2      A,2
## 12         2          3      2         1     3           B         3      B,3
## 13         2          4      2         1     2           B         2      B,2
## 14         2          5      2         3     2           C         3      C,3
## 15         2          6      2         2     1           A         3      A,3

```

```
## 16      2      7      2      2      3      A      1      A,1
## 17      2      8      2      3      3      C      2      C,2
## 18      2      9      2      1      1      B      1      B,1
## 19      3      1      3      1      1      B      1      B,1
## 20      3      2      3      3      1      A      2      A,2
## 21      3      3      3      2      3      C      1      C,1
## 22      3      4      3      2      2      C      3      C,3
## 23      3      5      3      2      1      C      2      C,2
## 24      3      6      3      3      3      A      1      A,1
## 25      3      7      3      3      2      A      3      A,3
## 26      3      8      3      1      2      B      3      B,3
## 27      3      9      3      1      3      B      2      B,2
## 28      4      1      4      2      3      B      3      B,3
## 29      4      2      4      2      1      B      1      B,1
## 30      4      3      4      1      1      A      3      A,3
## 31      4      4      4      1      2      A      2      A,2
## 32      4      5      4      1      3      A      1      A,1
## 33      4      6      4      3      1      C      1      C,1
## 34      4      7      4      2      2      B      2      B,2
## 35      4      8      4      3      2      C      3      C,3
## 36      4      9      4      3      3      C      2      C,2

#'## Plot the layout
#* Athlete_eg1lay
eg1.lay$Conditions <- with(eg1.lay, fac.combine(list(Intensities, Surfaces),
                                                    combine=TRUE, sep=","))
designGGPlot(eg1.lay, labels = "Conditions",
             row.factors = "Locations", column.factors = "Batches",
             cellfillcolour.column = "Athletes", cellalpha = 0.75, size = 6,
             title = "Intensities-Surfaces combinations (coloured by Athletes)",
             blockdefinition = rbind(c(9,1)),
             ggplotFuncs = list(xlab("Batches (Months)")))
```

Intensities–Surfaces combinations (coloured by Athletes)

Locations	1	C,3	C,1	B,1	B,3
	2	B,2	A,2	A,2	B,1
	3	C,1	B,3	C,1	A,3
	4	A,1	B,2	C,3	A,2
	5	A,2	C,3	C,2	A,1
	6	B,3	A,3	A,1	C,1
	7	C,2	A,1	A,3	B,2
	8	B,1	C,2	B,3	C,3
	9	A,3	B,1	B,2	C,2
		1	2	3	4
		Batches (Months)			

Check the properties of the design.

```

#### Check properties of the design
eg1.canon <- designAnatomy(formulae = list(locs = ~ Batches/Locations,
                                          tests = ~ Months/Athletes/Tests,
                                          cond = ~ Intensities*Surfaces),
                          data      = eg1.lay)
summary(eg1.canon, which.criteria="none")

##
##
## Summary table of the decomposition for locs, tests & cond
##
## Source.locs      df1 Source.tests      df2 Source.cond      df3
## Batches          3 Months
## Locations[Batches] 32 Athletes[Months] 8 Intensities      2
##                                     Residual      6
##               Tests[Months:Athletes] 24 Surfaces      2
##                                     Intensities#Surfaces 4
##                                     Residual      18

```

Questions

1. What would be the allocation-based mixed model for this experiment, an allocation-based mixed model having the same terms as the randomization-based mixed model that would apply if all the allocations had been made by randomizing. Do you anticipate any problem in fitting it?

The allocation-based mixed model is formed by treating all training-conditions factors as fixed and the remaining factors as random. Hence, the symbolic mixed model is $\text{Intensities} + \text{Surfaces} + \text{Intensities}^{\wedge}\text{Surfaces} \mid \text{Months} + \text{Months}^{\wedge}\text{Athletes} + \text{Months}^{\wedge}\text{Athletes}^{\wedge}\text{Tests} + \text{Batches} + \text{Batches}^{\wedge}\text{Locations}$. The problem in fitting it would be that Months and Batches are confounded so that the variance model is singular.

2. Compare the units for the two phases in this experiment?

A unit in the first phase is a test conducted on an athlete in a particular month; in the second phase, a unit is a location of a test within a batch. That is, the unit in the first phase is an athlete's test and in the second phase is a blood specimen in a lab location.

3. What are the outcomes for the two phases for this experiment?

The outcome for the first phase is the heart rate for a test and a blood specimen from the test; the outcome for the second phase, is the free haemoglobin measured at a location.

4.1.3 Allowing for lab processing order in the athletic training example

Brien (2017) discusses a design, and its properties, that differs in the second phase from that described in Section 4.1.2: it assumes that lab processing order within a batch is important and so the second phase now requires a row-column design. However, one cannot consider a design for just Months, Athletes and Tests and ignore Intensities and Surfaces, as was done in the previous design. Indeed prime consideration needs to be given to Intensities and Surfaces. That is, a suitable cross-phase design is needed. However, the second-phase design has to be considered in that it must account for the split-unit nature of the first-phase design.

For the second-phase design, the Months are associated with Batches. Then each triple of consecutive locations in a batch are associated with a single athlete, one of those for the month associated with the batch. This leaves tests to be assigned to locations within triples. Thus, the cross-phase design will need to allocate efficiently an intensity to a location triple and surface to the locations within a triple.

The cross-phase design is a balanced factorial design (Hinkelmann and Kempthorne, 2005, Section 12.5) and can be constructed as follows:

1. a 3×4 extended Latin square, formed from a 3×3 Latin square by repeating one of its columns, will be used to allocate Intensities to the 3 Locations triples \times 4 Months.

2. A 3×4 extended Latin squares will be used to allocate Surfaces to the 3 Locations \times 4 Months within a triple; the same extended Latin square is used for the three triples.
3. To ensure no repeat Intensities-Surfaces combinations for a Location, the repeated columns for the two extended Latin squares will be associated with different Batches.

The factor-allocation diagram, for this design, is in Figure 13. In this diagram, the training conditions and tests panels are surrounded by a dashed rectangle and lines go from the training conditions sources to the lines from the test sources. This indicates that the result of the allocation in the first phase needs to be explicitly taken into account in the second-phase allocation. The randomizations involved have been called a randomized-inclusive randomizations (Brien and Bailey, 2006) and are one of the two types of randomizations in a chain (Bailey and Brien, 2015). Because Batches and Locations are crossed, the second phase randomization is achieved by independently permuting the Batches and Locations. A design with the same properties had been previously constructed by Rosemary Bailey (pers. comm.).

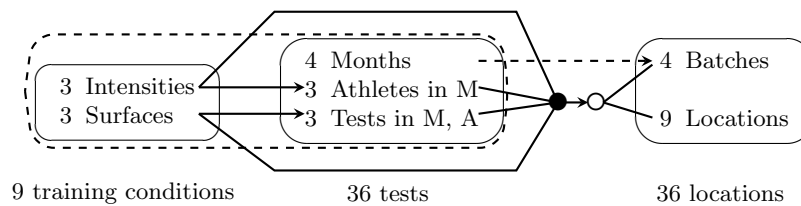


Figure 13: Factor-allocation diagram for the two-phase athlete training experiment with a row-column design for the second phase: training conditions are randomized to tests, then training conditions and tests are randomized to locations; the ‘●’ indicates that the observed combinations of the levels of Intensities, Surfaces, Athletes and Tests are randomized to locations; the ‘○’ indicates that a nonorthogonal design was used in this randomization to the combinations of the levels of Batches and Locations; the dashed arrow indicates that Months were systematically allocated to Batches; the dashed oval indicates that all factors from the first phase form a pseudotier and all are actively involved in determining the allocation to locations; M = Months and A = Athletes.

Use the following R code to obtain a layout for the new second phase design.

```
#'## Generate a systematic cross-phase design for Intensities and Surfaces
#'## It is based on an extended Latin square (ELS) for Intensities
#'## and four Latin squares for Surfaces, one for each Month such that the start rows are an ELS
#'+ Athlete_eg2sys_v3

eg2.phx.sys <- cbind(fac.gen(list(Batches = 4, Locations = 9)),
  data.frame(Intensities = factor(rep(c(designLatinSqrSys(3), c(3,2,1)),
    each = 3), labels = LETTERS[1:3]),
    Surfaces = factor(c(rep(1:3, times = 3),
      rep(1:3, times = 3),
      rep(c(2,3,1), times = 3),
      rep(c(3,1,2), times = 3))))))

eg2.phx.sys$Conditions <- with(eg2.phx.sys, fac.combine(list(Intensities, Surfaces),
  combine.levels = TRUE))

designGGPlot(eg2.phx.sys, labels = "Conditions",
  row.factors = "Locations", column.factors = "Batches",
  cellfillcolour.column = "Intensities", cellalpha = 0.75,
  celllinesize = 2, size = 6,
  title = "Intensities-Surfaces combinations in systematic design")
```

Intensities–Surfaces combinations in systematic design

1	A,1	B,1	C,2	C,3
2	A,2	B,2	C,3	C,1
3	A,3	B,3	C,1	C,2
4	B,1	C,1	A,2	B,3
5	B,2	C,2	A,3	B,1
6	B,3	C,3	A,1	B,2
7	C,1	A,1	B,2	A,3
8	C,2	A,2	B,3	A,1
9	C,3	A,3	B,1	A,2
	1	2	3	4
Locations	Batches			

```

### Second phase design
### Generate a systematic two-phase design by bringing in first-phase recipient factors
eg2.phx.sys$Months <- eg2.phx.sys$Batches
eg2.sys <- merge(split.lay, eg2.phx.sys) #merge on common factors Months, Intensities & Surfaces

### Allocate the second phase
eg2.lay <- designRandomize(allocated = eg2.sys[c("Months", "Athletes", "Tests",
                                                "Intensities", "Surfaces")],
                          recipient = eg2.sys[c("Batches", "Locations")],
                          except    = "Batches",
                          seed      = 243526)

head(eg2.lay)

##   Batches Locations Months Athletes Tests Intensities Surfaces
## 1      1         1      1         3     2           A         2
## 2      1         2      1         2     1           C         2
## 3      1         3      1         3     3           A         3
## 4      1         4      1         2     2           C         1
## 5      1         5      1         3     1           A         1
## 6      1         6      1         1     2           B         2

### Plot the layout
#+ Athlete_eg2lay_v3
eg2.lay$Conditions <- with(eg2.lay, fac.combine(list(Intensities, Surfaces),
                                                    combine=TRUE, sep=", "))
designGGPlot(eg2.lay, labels = "Conditions",
            row.factors = "Locations", column.factors = "Batches",
            cellfillcolour.column = "Athletes", cellalpha = 0.75, size = 6,
            title = "Intensities-Surfaces combinations (coloured by Athletes)",
            blockdefinition = rbind(c(9,1)),
            ggplotFuncs = list(xlab("Batches (Months)")))

```

Intensities–Surfaces combinations (coloured by Athletes)

1	A,2	B,2	C,3	C,1
2	C,2	A,2	B,3	A,1
3	A,3	B,3	C,1	C,2
4	C,1	A,1	B,2	A,3
5	A,1	B,1	C,2	C,3
6	B,2	C,2	A,3	B,1
7	C,3	A,3	B,1	A,2
8	B,1	C,1	A,2	B,3
9	B,3	C,3	A,1	B,2
	1	2	3	4

Batches (Months)

Check the properties of the design.


```

#### Check properties of the design
eg2.canon <- designAnatomy(formulae = list(locs = ~ Batches*Locations,
                                           tests = ~ Months/Athletes/Tests,
                                           cond = ~ Intensities*Surfaces),
                           data      = eg2.lay)
summary(eg2.canon, which.criteria =c("aefficiency", "order"))

##
##
## Summary table of the decomposition for locs, tests & cond (based on adjusted quantities)
##
## Source.locs      df1 Source.tests      df2 Source.cond      df3 aefficiency order
## Batches          3 Months              3              1.0000      1
## Locations        8 Athletes[Months]    2 Intensities      2 0.0625      1
##                  Tests[Months:Athletes] 6 Surfaces         2 0.0625      1
##                  Intensities#Surfaces    4 0.2500      1
## Batches#Locations 24 Athletes[Months]    6 Intensities      2 0.9375      1
##                  Residual                4 1.0000      1
##                  Tests[Months:Athletes] 18 Surfaces         2 0.9375      1
##                  Intensities#Surfaces    4 0.7500      1
##                  Residual                12 1.0000      1
##
## The design is not orthogonal

```

It is clear that Athletes[Months] and Tests[Months:Athletes] are not orthogonal to Locations and Batches#Locations, because the former sources are confounded with both of the latter sources. To examine the nature of the nonorthogonality, the skeleton anova for just the tests and locations tiers is obtained.

```

#### Examine the nonorthogonality between locations and tests
eg2.locstests.canon <- designAnatomy(formulae = list(locs = ~ Batches*Locations,
                                                    tests = ~ Months/Athletes/Tests),
                                     data      = eg2.lay)
summary(eg2.locstests.canon, which.criteria =c("aefficiency", "order"))

##
##
## Summary table of the decomposition for locs & tests
##
## Source.locs      df1 Source.tests      df2 aefficiency order
## Batches          3 Months              3 1.0000      1
## Locations        8 Athletes[Months]    2 1.0000      1
##                  Tests[Months:Athletes] 6 1.0000      1
## Batches#Locations 24 Athletes[Months]    6 1.0000      1
##                  Tests[Months:Athletes] 18 1.0000      1
##

```

Questions

1. What do you conclude about the confounding of Athletes[Months] and Tests[Months:Athletes] with Locations?

Since all efficiency factors are one, it is concluded that the 8 degrees of freedom for Athletes[Months] has been split into two orthogonal parts, one with 2 degrees of freedom which is confounded with Batches and the other with 6 degrees of freedom which is confounded with Batches:Locations. The source Tests[Months:Athletes] has been similarly partitioned.

2. Are the designs proposed for this experiment first-order balanced?

The design is first-order balanced, because the order of the efficiency factors is one for all confounded sources.

3. What has been the cost of allowing for order of processing in the lab? Is the cost acceptable? Why?

The cost has been that some information about Athletes[Months], along with Intensities, and some information about Tests[Months:Athletes], along with Surfaces and Intensities#Surfaces, has been confounded with Locations. The cost is acceptable, because the amount of information lost on the main effects is only 6.25% and on the interaction is 25%. The latter will be recovered in a REML-based mixed model analysis. However, the Residual degrees of freedom for Athletes[Months] has been reduced from 6 to 4 and for Tests[Months:Athletes] from 18 to 14. While the latter is unlikely to be seriously deleterious, the former is of concern.

4.2 McIntyre's (1955) two-phase experiment

McIntyre (1955) reports an investigation of the effect of four light intensities on the synthesis of tobacco mosaic virus in leaves of tobacco *Nicotiana tabacum* var. Hickory Pryor. It is a two-phase experiment: the first phase is a treatment phase, in which the four light treatments are randomized to the tobacco leaves, and the second phase is an assay phase, in which the tobacco leaves are randomized to the half-leaves of assay plants.

In the first phase, four successive leaves at defined positions on the stem were taken from each of eight plants of comparable age and vigour that had been inoculated with the virus. Arbitrarily grouping the plants into two sets of four, the four treatments were applied to the leaves, which had been separated from the plants and were sustained by flotation on distilled water, in a Latin square design for each set with tobacco plants as columns and leaf positions as rows; see Figure 15.

In the second phase, virus content of each tobacco leaf was assayed by expressing sap and inoculating half leaves of the assay plants, *Datura stramonium*, on which countable lesions would appear. Lots of eight sap samples were formed from pairs of tobacco plants, the pairs being comprised of a plant from each set in the treatment phase. The eight samples from a lot were assigned to four assay plants using one of four 4×4 Graeco-Latin square designs, with the leaves from a single tobacco plant assigned using one of the alphabets and the second tobacco plant using the other (see Figure 16). Actually, this design is a semi-Latin square (Bailey, 1992).

The factor-allocation diagram for the experiment is in Figure 14. Unfortunately, the randomization for this experiment was not described by McIntyre (1955). Because there are multiple squares in both phases, there are several possible randomizations depending on the effects anticipated as possible in the experiment. As shown by the nesting relations in the factor-allocation diagram, I have assumed that NicPlant is randomized within Sets and Posn randomized across Sets. Similarly, I have assumed that DatPlant was randomized with Lot and AssPosn across Lot. In the factor-allocation diagram, N_1 is a factor for the pairs of tobacco plants formed by taking a plant from each set in the first phase.

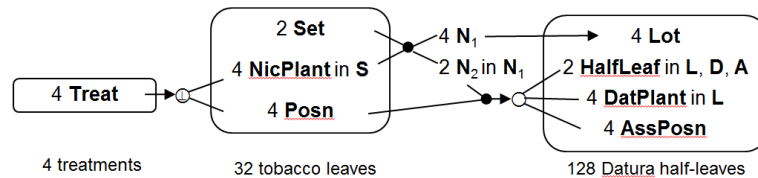


Figure 14: Factor-allocation diagram for McIntyre's (1955) two-phase experiment: treatments are randomized to tobacco leaves and tobacco leaves are randomized to Datura half-leaves; the arrow to the '⊗', the '⊗' and the two lines from the '⊗' indicate that Treat is randomized to the combinations of NicPlant and Posn using an orthogonal design; N_1 is a pseudofactor indexing the pairs of tobacco plants formed by taking a plant from each set in the first phase and N_2 is a pseudofactor indexing the tobacco plants within the pairs formed by taking a plant from each set in the first phase; N_1 is randomized to Lot in the second phase; the combinations of N_2 and Posn is randomized to the combinations of HalfLeaf, DatPlant and AssPosn using a nonorthogonal design, the latter indicated by the 'O'; S = Set; L = Lot; D = DatPlant; A = AssPosn.

Figure 15: Layout for the first phase of McIntyre's (1955) experiment[†]

Nicotiana Plants											
		1	2	3	4			1	2	3	4
Leaf	Position					Leaf	Position				
1		a	b	c	d			a	b	c	d
		1	5	9	13			17	21	25	29
2		b	a	d	c			c	d	a	b
		2	6	10	14			18	22	26	30
3		c	d	a	b			d	c	b	a
		3	7	11	15			19	23	27	31
4		d	c	b	a			b	a	d	c
		4	8	12	16			20	24	28	32

[†]The letter in each cell refers to the light intensity to be applied to the unit and the number to the unit.

Figure 16: Layout for the second phase of McIntyre's (1955) experiment[†]

		<i>Datura</i> Plants									
		1	2	3	4	5	6	7	8		
Assay Leaf	Position					Assay Leaf	Position				
1		1 17	2 20	3 18	4 19	5 23	6 22	7 24	8 21		
2		2 18	1 19	4 17	3 20	8 22	7 23	6 21	5 24		
3		3 19	4 18	1 20	2 17	7 21	8 24	5 22	6 23		
4		4 20	3 17	2 19	1 18	6 24	5 21	8 23	7 22		

		<i>Datura</i> Plants									
		9	10	11	12	13	14	15	16		
Assay Leaf	Position					Assay Leaf	Position				
1		9 28	10 25	11 27	12 26	13 30	14 31	15 29	16 32		
2		10 27	9 26	12 28	11 25	16 31	15 30	14 32	13 29		
3		11 26	12 27	9 25	10 28	15 32	16 29	13 31	14 30		
4		12 25	11 28	10 26	9 27	14 29	13 32	16 30	15 31		

[†]The numbers in the cell refer to the units from the first phase (tobacco leaves) to be assigned to the two half-leaves of the assay plant; they are in standard order for Set, then NicPlant followed by Position.

4.2.1 Check the properties of the randomized layout

Load the data and use `designAnatomy` to check the properties of the design.

```

#### Load data
data("McIntyreTMV.dat")
#### Check properties of the design
TMV.canon <- designAnatomy(formulae = list(assay = ~ ((Lot/DatPlant)*AssPosn)/HalfLeaf,
                                          test  = ~ (Set/NicPlant)*Posn,
                                          trt   = ~ Treat),
                          data      = McIntyreTMV.dat)
summary(TMV.canon, which.criteria=c("aeff", "ord"))

##
##
## Summary table of the decomposition for assay, test & trt (based on adjusted quantities)
##
## Source.assay          df1 Source.test          df2 Source.trt df3 aefficiency order
## Lot                  3 NicPlant[Set]          3          1.0000      1
## DatPlant[Lot]        12
## AssPosn              3
## Lot#AssPosn          9
## DatPlant#AssPosn[Lot] 36 Posn              3          0.5000      1
##                      Set#Posn              3          0.5000      1
##                      NicPlant#Posn[Set]    18 Treat          3          0.5000      1
##                      Residual             15          0.5000      1
##                      Residual             12
## HalfLeaf[Lot:DatPlant:AssPosn] 64 Set              1          1.0000      1
##                      NicPlant[Set]         3          1.0000      1
##                      Posn                  3          0.5000      1
##                      Set#Posn              3          0.5000      1
##                      NicPlant#Posn[Set]    18 Treat          3          0.5000      1
##                      Residual             15          0.5000      1
##                      Residual             36
##
## The design is not orthogonal

```

4.2.2 Questions

1. Is the variance matrix for this experiment based on two sets of terms that are orthogonal?

The variance matrix for this experiment is based on the factors in the tobacco leaves and Datura half-leaves tiers. The terms derived from the factors in these two tiers are not orthogonal. In particular, Set#Posn and NicPlant#Posn[Set] are partially confounded with both DatPlant#AssPosn[Lot] and HalfLeaf[Lot:DatPlant:AssPosn].

2. What are the advantages and disadvantages of a mixed -model analysis of the data from this experiment, as opposed to an anova?

The advantage of a mixed-model analysis is that combined estimates will be provided for Set#Posn, NicPlant#Posn[Set], and Treat. The disadvantages are (i) that not all random terms are well-estimated, some having small degrees of freedom, and cause problems in fitting the model, and (ii) the Wald F-statistics are only approximately distributed as F-distributions. On the other hand, an anova is not applicable because of the nonorthogonality between the sets of terms making up the variance matrix; at least some F-ratios will not be independently distributed.

4.3 A Plant Accelerator experiment with a split-unit design

This experiment involves the investigation of 75 wheat lines, of which 73 are Nested Association Mapping (NAM) wheat lines and the other two are two check lines, Scout and Gladius. It was conducted in 2014 in the Plant Accelerator, a facility in Adelaide with 4 Smarthouses. A Smarthouse is a large greenhouse with two areas within it: (i) a Table area at the southern end and (ii) a Conveyor area at the northern end — see Figure 17. The conveyor system has the capability of automatically moving and imaging around 500 pots per day. There are air conditioners placed down the western side of the Smarthouse, which creates a trend from west to east. Further, there is a north-south trend due to changes in light intensity (Brien et al., 2013).

The experiment involves two phases: the table and conveyor phases. The table phase is the establishment phase in which plants are germinated in pots on the tables where they undergo an early growth stage. In the conveyor phase, having placed the pots in carts on the conveyor system, the plants are automatically imaged and watered daily, being moved to a processing station by the conveyor system for this.

This experiment has a single plant per pot and these will be arranged in a 24×22 grid in both phases: 24 columns \times 22 locations in the table phase and 24 lanes \times 22 (2–23) positions, as shown in Figure 17. However, the 24 columns in the table phase run east-west and the 24 lanes in the conveyor phase run north-south. Because there are systematic trends in both phases to be accounted for in the analysis, the same layout will be used in both phases, but the table layout will be rotated clockwise through 90° . That is, Locations 1–22 will be in Positions 2–23, respectively, and the Column will be placed in the Lane with the same number.

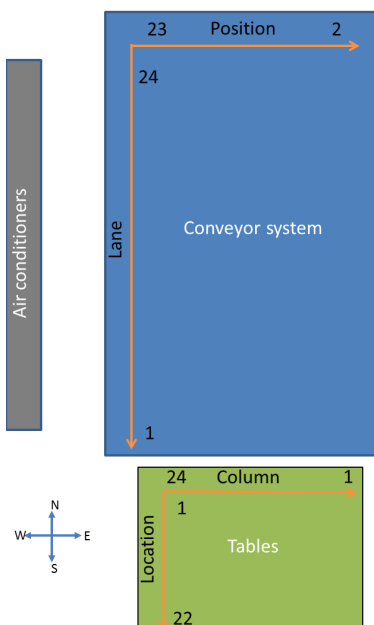


Figure 17: Schematic of Smarthouse for the Plant Accelerator experiment

The design employed for the experiment is a split-unit design in which two consecutive pots/carts form a main unit. The main-unit design uses a blocked design with rows and columns generated with DiGger (Coombes, 2009). It assigns Lines to main units, the Lines being unequally replicated; Scout and Gladius each occur on 12 main units (24 carts), 21 randomly-selected NAM lines each occur on 4 main units (8 carts) and the remaining 52 NAM lines each occur on 3 main units (6 carts). The subunit design merely randomizes Salt (0 mM NaCl, 100 mM NaCl) to the two carts in each main unit.

In the main-unit design, the blocks are, in the table phase, 6 Groups of 4 Columns and, in the conveyor phase, 6 Zones of 4 Rows (lanes). However, while the generated design is based on crossed rows and columns, it is known from past experience that, while there are differences between Zones, there are not differences between Rows within Zones (Brien et al., 2013) and none are anticipated between Columns within Groups on the tables. The columns of the main-unit design are indexed by 11 Pairs in the table phase and 11 MainPosns in the conveyor phase. The design generated with DiGger (Coombes, 2009) will be rerandomized so that the Lines are

randomized to 4 Columns within each Groups-Pairs combination and the 11 sets of Lines assigned by DiGger to the 11 Pairs will be rerandomized to Pairs. The factor-allocation diagram is shown in Figure 18.

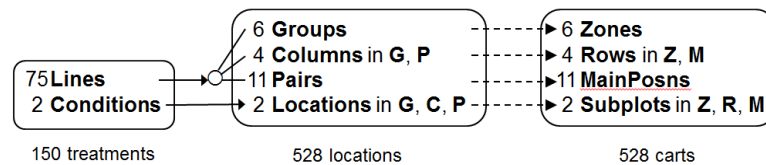


Figure 18: Factor-allocation diagram for the Plant Accelerator experiment: treatments are randomized to locations and locations are allocated to cars; the arrow to the ‘O’, the ‘O’ and the three lines from the ‘O’ indicate that Lines is randomized to the combinations of Groups, Columns and Pairs using a nonorthogonal design; the arrow from Conditions to Locations indicates that Conditions were randomized to Locations; the dashed arrows between the two panels on the right hand side indicate that the factors indexing locations were systematically assigned to those indexing the carts; G = Groups; P = Pairs; C = Columns; Z = Zones; M = MainPosns; R = Rows.

4.3.1 Produce the layout

Use the following instructions to load the main-unit design produced with DiGger and check its properties using `designAnatomy`.

```
### Load the main-unit design - it has Lines in row-column order
data("Exp249.munit.des")
Exp249.munit.des$Blocks <- factor(rep(1:6, each = 44))

### Check its properties
Exp249.munit.canon <- designAnatomy(formulae = list(cart = ~ (Blocks/Rows)*Cols,
                                                    treat = ~ Lines),
                                   data      = Exp249.munit.des)

summary(Exp249.munit.canon)

##
##
## Summary table of the decomposition for cart & treat (based on adjusted quantities)
##
## Source.cart      df1 Source.treat df2 aefficiency eefficiency order
## Blocks           5 Lines           5      0.1498      0.1422      5
## Rows[Blocks]     18 Lines          18      0.2685      0.1813     18
## Cols             10 Lines           10      0.2102      0.1769     10
## Blocks#Cols      50 Lines           50      0.1154      0.0142     50
## Rows#Cols[Blocks] 180 Lines          74      0.5816      0.2088     74
## Residual         106
##
## The design is not orthogonal
```

Expand main-unit design to produce the split-unit design, including a three-level factor Checks that compares Scout, Gladius and the mean of the NAM lines. Perhaps, produce a plot of the allocation of the Lines.

```
### Expand design to rerandomize lines and to assign salt treatments to locations
Exp249.alloc <- with(Exp249.munit.des,
                     data.frame(Lines = factor(rep(Lines, each=2), levels=1:75),
                               Checks = fac.recode(rep(Lines, each=2),
```

```

                                newlevels=c(rep(3, 73), 1 , 2),
                                labels = c("NAM","Scout","Gladius")),
                                Salt = factor(rep(1:2, times=264),
                                labels = c('0 NaCl','100 NaCl'))))
Exp249.recip <- fac.gen(list(Groups = 6, Cols = 4, Pairs = 11, Locations = 2))
Exp249.nest  <- list(Cols = c("Groups", "Pairs"),
                    Locations = c("Groups", "Cols", "Pairs"))
Exp249.lay <- designRandomize(allocated      = Exp249.alloc,
                             recipient      = Exp249.recip,
                             nested.recipients = Exp249.nest,
                             seed           = 51412)

### Add second-phase factors
### (to which the first-phase factors have been systematically allocated)
Exp249.lay <- cbind(fac.gen(list(Lanes = 24, Positions = 2:23)),
                   fac.gen(list(Zones = 6, Rows = 4, MainPosn = 11, Subunits = 2)),
                   Exp249.lay)

### Plot the assignment of Lines in the second-phase design - or see file that includes the output
Exp249.lay$Replication <- fac.recode(Exp249.lay$Lines, rep(1:3, c(21,52,2)))
designGGPlot(Exp249.lay, labels = "Lines", cellfillcolour.column = "Replication",
            colour.values = c("lightblue", "grey", "lightgreen"),
            row.factors = "Lanes", column.factors = "Positions",
            title = "Layout of Lines for optimized design",
            reverse.x = TRUE, reverse.y = FALSE, blockdefinition = cbind(4,22))

```

Layout of Lines for optimized design

24	11	11	7	7	17	17	3	3	50	50	39	39	71	71	18	18	74	74	10	10	75	75
23	75	75	69	69	24	24	67	67	28	28	4	4	74	74	45	45	13	13	51	51	31	31
22	26	26	20	20	35	35	12	12	44	44	52	52	68	68	38	38	61	61	41	41	55	55
21	2	2	22	22	9	9	30	30	36	36	58	58	27	27	72	72	16	16	57	57	8	8
20	41	41	64	64	43	43	75	75	48	48	30	30	13	13	70	70	19	19	47	47	12	12
19	34	34	74	74	40	40	6	6	31	31	2	2	62	62	21	21	53	53	59	59	29	29
18	20	20	5	5	50	50	68	68	14	14	45	45	8	8	55	55	11	11	75	75	15	15
17	44	44	26	26	72	72	60	60	73	73	16	16	1	1	74	74	33	33	46	46	51	51
16	54	54	60	60	21	21	61	61	63	63	48	48	34	34	75	75	75	75	50	50	33	33
15	37	37	53	53	18	18	65	65	59	59	74	74	49	49	23	23	8	8	58	58	32	32
14	45	45	6	6	42	42	4	4	15	15	57	57	7	7	69	69	2	2	72	72	17	17
13	74	74	66	66	3	3	24	24	19	19	14	14	73	73	11	11	39	39	35	35	13	13
12	10	10	44	44	5	5	8	8	1	1	43	43	39	39	60	60	28	28	29	29	26	26
11	65	65	42	42	70	70	27	27	18	18	75	75	54	54	9	9	6	6	67	67	74	74
10	19	19	40	40	33	33	46	46	24	24	37	37	14	14	16	16	23	23	64	64	25	25
9	71	71	75	75	15	15	66	66	12	12	56	56	38	38	58	58	22	22	74	74	57	57
8	56	56	55	55	29	29	13	13	47	47	5	5	20	20	68	68	43	43	32	32	49	49
7	25	25	1	1	75	75	18	18	9	9	27	27	41	41	31	31	17	17	65	65	54	54
6	21	21	75	75	7	7	62	62	2	2	6	6	36	36	52	52	10	10	23	23	63	63
5	61	61	4	4	74	74	74	74	42	42	3	3	64	64	73	73	40	40	71	71	11	11
4	4	4	37	37	25	25	47	47	10	10	62	62	15	15	49	49	20	20	16	16	14	14
3	69	69	48	48	56	56	9	9	74	74	51	51	5	5	7	7	67	67	53	53	46	46
2	12	12	52	52	30	30	59	59	38	38	19	19	75	75	66	66	21	21	36	36	22	22
1	70	70	32	32	34	34	17	17	75	75	63	63	35	35	28	28	1	1	74	74	3	3
	23	22	21	20	19	18	17	16	15	14	13	12	11	10	9	8	7	6	5	4	3	2

4.3.2 Check the properties of the design

The maximal allocation-based mixed model is $(\text{Checks} + \text{Lines}) * \text{Salt} \mid (\text{Zones} * \text{MainPosn}) / \text{Rows} / \text{Subunits} + (\text{Groups} * \text{Pairs}) / \text{Cols} / \text{Locations}$, with Checks nested within Lines. Use the `designAnatomy` to check the properties of the design for an analysis of data from an experiment based on this design.

```
## Check design properties
Exp249.canon <- designAnatomy(formulae = list(carts = ~ (Zones*MainPosn)/Rows/Subunits,
                                             tables = ~ (Groups*Pairs)/Cols/Locations,
                                             treats = ~ (Checks + Lines) * Salt),
                             data      = Exp249.lay)

## Warning in proj.2canon(CombinedSets$Q[[ntiers]], struct[[ktier]]$Q): Lines[Checks] and Checks
```


are partially aliased in MainPosn&Pairs

```
## Warning in proj.2canon(CombinedSets$Q[[ntiers]], struct[[ktier]]$Q): Lines[Checks] and Checks
are partially aliased in Zones#MainPosn&Groups#Pairs
## Warning in proj.2canon(CombinedSets$Q[[ntiers]], struct[[ktier]]$Q): Lines[Checks] and Checks
are partially aliased in Rows[Zones:MainPosn]&Cols[Groups:Pairs]
```

```
summary(Exp249.canon)
```

```
##
##
## Summary table of the decomposition for carts, tables & treats (based on adjusted quantities)
##
## Source.carts          df1 Source.tables          df2 Source.treats          df3
## Zones                5 Groups                5 Lines[Checks]          5
## MainPosn             10 Pairs                10 Checks              2
##                      Lines[Checks]          8
## Zones#MainPosn       50 Groups#Pairs          50 Checks              2
##                      Lines[Checks]          48
## Rows[Zones:MainPosn] 198 Cols[Groups:Pairs]    198 Checks              2
##                      Lines[Checks]          72
##                      Residual              124
## Subunits[Zones:MainPosn:Rows] 264 Locations[Groups:Pairs:Cols] 264 Salt              1
##                      Checks#Salt            2
##                      Lines#Salt[Checks]      72
##                      Residual              189
##
## aefficiency eefficiency order
##      0.1498      0.1422      5
##      0.0033      0.0031      2
##      0.2094      0.1809      8
##      0.2111      0.2049      2
##      0.1142      0.0145     48
##      0.7854      0.7792      2
##      0.6640      0.2632     66
##      1.0000      1.0000      1
##      1.0000      1.0000      1
##      1.0000      1.0000      1
##      1.0000      1.0000      1
##      1.0000      1.0000      1
##
## Table of (partial) aliasing between sources derived from the same formula
##
## Source          df Alias In          aefficiency eefficiency order
## Lines[Checks]   8 Checks MainPosn&Pairs          0.2094      0.1809      8
## Lines[Checks]  48 Checks Zones#MainPosn&Groups#Pairs          0.1142      0.0145     48
## Lines[Checks]  72 Checks Rows[Zones:MainPosn]&Cols[Groups:Pairs]          0.6640      0.2632     66
##
## The design is not orthogonal
```

Because, there is a one-to-one correspondence between the tables and carts sources, omit the tables formula and rerun — it will make the anova table more readable.

```

#### Check design properties, with tables omitted
Exp249.canon <- designAnatomy(formulae = list(carts = ~ (Zones*MainPosn)/Rows/Subunits,
                                             treats = ~ (Checks + Lines) * Salt),
                             data      = Exp249.lay)

## Warning in projs.2canon(CombinedSets$Q[[ntiers]], struct[[ktier]]$Q): Lines[Checks] and Checks
## are partially aliased in MainPosn
## Warning in projs.2canon(CombinedSets$Q[[ntiers]], struct[[ktier]]$Q): Lines[Checks] and Checks
## are partially aliased in Zones#MainPosn
## Warning in projs.2canon(CombinedSets$Q[[ntiers]], struct[[ktier]]$Q): Lines[Checks] and Checks
## are partially aliased in Rows[Zones:MainPosn]

summary(Exp249.canon)

##
##
## Summary table of the decomposition for carts & treats (based on adjusted quantities)
##
## Source.carts      df1 Source.treats      df2 aeffecticiency eeffecticiency order
## Zones             5 Lines[Checks]        5      0.1498      0.1422      5
## MainPosn          10 Checks                2      0.0033      0.0031      2
##                   Lines[Checks]          8      0.2094      0.1809      8
## Zones#MainPosn    50 Checks                2      0.2111      0.2049      2
##                   Lines[Checks]          48     0.1142      0.0145     48
## Rows[Zones:MainPosn] 198 Checks            2      0.7854      0.7792      2
##                   Lines[Checks]          72     0.6640      0.2632     66
##                   Residual              124
## Subunits[Zones:MainPosn:Rows] 264 Salt          1      1.0000      1.0000      1
##                   Checks#Salt            2      1.0000      1.0000      1
##                   Lines#Salt[Checks]     72      1.0000      1.0000      1
##                   Residual              189
##
## Table of (partial) aliasing between sources derived from the same formula
##
## Source      df Alias In      aeffecticiency eeffecticiency order
## Lines[Checks] 8 Checks MainPosn      0.2094      0.1809      8
## Lines[Checks] 48 Checks Zones#MainPosn 0.1142      0.0145     48
## Lines[Checks] 72 Checks Rows[Zones:MainPosn] 0.6640      0.2632     66
##
## The design is not orthogonal

```

4.3.3 Examine the properties of the design for an alternative analysis

However, rather than fit the allocation-based model, because it is known from past experience that once a linear trend for MainPosn has been fitted there are no deviations from this trend, the term `xMainPosn` is used to fit the trend; the term `xMainPosn` is a centred, linear covariate for MainPosn. Use the `designAnatomy` to check the properties of the design for an analysis based on a modified model, in which MainPosn in the random model has been replaced by `xMainPosn` in the fixed model, `Zones#MainPosn` has been omitted and `Rows[Zones:MainPosn]` has been replaced by `Mainunits[Zones]`.

```

#### Add factors and variates for new analysis
Exp249.lay <- within(Exp249.lay,
                    { xMainPosn <- as.numfac(MainPosn)
                      xMainPosn <- -(xMainPosn - mean(xMainPosn))
                    })

```

```

Mainunits <- fac.combine(list(Rows,MainPosn))
})
head(Exp249.lay)

##    Lanes Positions Zones Rows MainPosn Subunits Groups Cols Pairs Locations Lines Checks Salt
## 1      1          2     1     1        1          1      1     1     1          1     3 Gladius 100 NaCl
## 2      1          3     1     1        1          2      1     1     1          2     3 Gladius   0 NaCl
## 3      1          4     1     1        2          1      1     1     2          1    74      NAM 100 NaCl
## 4      1          5     1     1        2          2      1     1     2          2    74      NAM   0 NaCl
## 5      1          6     1     1        3          1      1     1     3          1     1 Gladius 100 NaCl
## 6      1          7     1     1        3          2      1     1     3          2     1 Gladius   0 NaCl
##    Replication Mainunits xMainPosn
## 1              1          1          5
## 2              1          1          5
## 3              3          2          4
## 4              3          2          4
## 5              1          3          3
## 6              1          3          3

#'## Check properties if only linear trend fitted
Exp249.canon <- designAnatomy(formulae = list(cart = ~ Zones/Mainunits/Subunits,
                                             treat = ~ xMainPosn +
                                                  (Checks + Lines) * Salt),
                             data      = Exp249.lay)
summary(Exp249.canon)

##
##
## Summary table of the decomposition for cart & treat (based on adjusted quantities)
##
## Source.cart      df1 Source.treat      df2 aefficiency eefficiency order
## Zones            5 Lines[Checks]      5      0.1500      0.1426      5
## Mainunits[Zones] 258 xMainPosn         1      1.0000      1.0000      1
##                  Checks               2      1.0000      1.0000      1
##                  Lines[Checks]        72      0.9879      0.8437      6
##                  Residual            183
## Subunits[Zones:Mainunits] 264 Salt      1      1.0000      1.0000      1
##                  Checks#Salt          2      1.0000      1.0000      1
##                  Lines#Salt[Checks]    72      1.0000      1.0000      1
##                  Residual            189
##
## Table of (partial) aliasing between sources derived from the same formula
##
## Source      df Alias      In      aefficiency eefficiency order
## Checks      2 xMainPosn treat    0.9995      0.9990      2
## Lines[Checks] 74 xMainPosn treat    0.9963      0.7851      2
## Checks#Salt   5 xMainPosn treat    0.9998      0.9990      2
## Lines#Salt[Checks] 149 xMainPosn treat    0.9982      0.7851      2
##
## The design is not orthogonal

```

The Table of (partial) aliasing shows that all `treat`- sources are partially aliased with `\verbxMainPosn=`, although they are not far from being orthogonal.

We have been able to check what information is available about Lines and Salt after adjustment for the linear

trend. In practice, a spline term might be needed to account for nonlinearity in the trend.

4.3.4 Questions

1. What advantages accrue from randomizing Lines within Groups[^]Pairs (Zones[^]MainPosn) as compared to the original DiGger design, in which they are randomized to Cols within Groups (Lanes within Zones) and to Pairs (MainPosn)?

The anatomy for the DiGger design shows us that all 74 degrees of freedom are estimable in Rows#Cols[Blocks] with average efficiency 0.582 and minimum efficiency 0.209. Compared to this, the anatomy for the rerandomized design shows that the NAM lines are estimable from Rows[Zones:MainPosn], the source equivalent to Rows#Cols[Blocks], with average efficiency 0.664 and minimum efficiency 0.263. Also, the Residual degrees of freedom for Rows#Cols[Blocks] have increased from 106 degrees of freedom in the original design to 124 degrees of freedom for Rows[Zones:MainPosn] in the rerandomized design. That is, one can expect the estimation of the Lines predictions and their standard errors to be more precise for the rerandomized design.

2. What effect does the use of a linear trend, as opposed to a set of effects, have on the analysis?

The efficiency for Lines has increased further so that the minimum is now 0.844 and the Residual degrees of freedom for Rows[Zones:MainPosn] now stands at 183. This allows one to consider ignoring information not estimable from Rows[Zones:MainPosn], while predictions will be adjusted for the trend across MainPosn.

4.4 Two-phase, wheat experiment with 49 lines

The first, or field, phase of a wheat trial for 49 lines is laid out as an RCBD with four blocks. The produce from the field trial is processed in the second, or laboratory, phase and the design employed is a balanced lattice square for 49 treatments that involves 4 replicates each consisting of a 7×7 square. In the laboratory phase there are four intervals each of which consists of 7 runs of a machine. In a run, samples are processed at seven consecutive times. This experiment is Example 2.2 from Bailey and Brien (2015), where its anova with expected mean squares is given. Its factor-allocation diagram is in Figure 19.

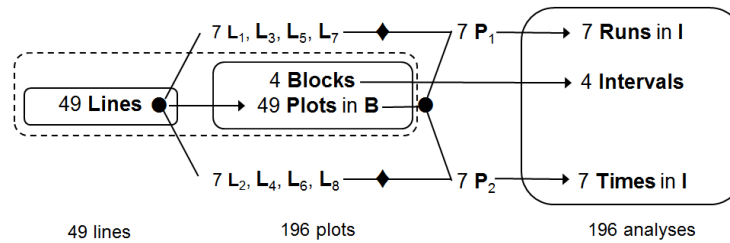


Figure 19: Factor-allocation diagram for the two-phase wheat variety experiment: lines are randomized to plots, then lines and plots are randomized to analyses; the arrow for Lines to Plots indicates that Lines are randomized to Plots; similarly, Blocks are randomized to intervals; L_1, L_3, L_5 and L_7 are pseudofactors that group the Lines for randomization to Runs and L_2, L_4, L_6 and L_8 are pseudofactors that group the Lines for randomization to Times; the two ‘♦’ symbols indicate that the pseudofactors for Lines determine the pseudofactors P_1 and P_2 for assigning Plots to Runs and Times, respectively; B = Blocks; I = Intervals.

4.4.1 Produce randomized layout for both phases and check its properties

```
### Generate a layout for the field phase
field.sys <- cbind(fac.gen(list(Blocks = 4, Plots = 49)),
                  fac.gen(list(Lines = 49), times=4))
field.lay <- designRandomize(allocated = field.sys["Lines"],
```

```

recipient      = field.sys[c("Blocks", "Plots")],
nested.recipients = list(Plots = "Blocks"),
seed           = 82522)

head(field.lay)

##   Blocks Plots Lines
## 1      1      1    48
## 2      1      2    10
## 3      1      3    23
## 4      1      4    31
## 5      1      5    36
## 6      1      6    11

##### Generate laboratory phase
##### Load a systematic balanced lattice square
data("LatticeSquare_t49.des")
##### Form systematic design
##### Add Intervals to field layout, merge the data frames and sort into lab phase order
field.lay$Intervals <- field.lay$Blocks
lab.alloc <- merge(LatticeSquare_t49.des, field.lay)
lab.alloc <- with(lab.alloc, lab.alloc[order(Intervals, Runs, Times),])
lab.alloc <- lab.alloc[c("Blocks", "Plots", "Lines")] #Reduce columns in lab.alloc
##### Randomize the design
lab.lay <- designRandomize(allocated      = lab.alloc,
                           recipient      = list(Intervals = 4, Runs = 7, Times = 7),
                           nested.recipients = list(Runs = "Intervals",
                                                       Times = "Intervals"),
                           seed           = 141797)

head(lab.lay)

##   Intervals Runs Times Blocks Plots Lines
## 1      1      1      1      4      41      1
## 2      1      1      2      4      43      49
## 3      1      1      3      4      36      25
## 4      1      1      4      4      13       9
## 5      1      1      5      4      10      33
## 6      1      1      6      4      44      41

##### Plot the design to show the allocation of Blocks, Plots and Lines in the lab phase
lab.lay$FieldFactors <- with(lab.lay, fac.combine(list(Blocks, Plots, Lines),
                                                    combine.levels = TRUE))

designGGPlot(lab.lay, labels = "FieldFactors",
             row.factors = c("Intervals", "Runs"), column.factors = "Times",
             title = "Allocation of Blocks, Plots and Lines in the lab phase",
             cellalpha = 0.75, blockdefinition = cbind(7, 7))

```

Allocation of Blocks, Plots and Lines in the lab phase

Runs	1	4,41,1	4,43,49	4,36,25	4,13,9	4,10,33	4,44,41	4,29,17	Intervals: 1	
	2	4,5,46	4,14,38	4,35,21	4,47,5	4,45,22	4,24,30	4,19,13		
	3	4,33,31	4,17,23	4,18,6	4,3,39	4,20,14	4,37,15	4,49,47		
	4	4,40,27	4,23,19	4,26,44	4,28,35	4,46,3	4,7,11	4,9,36		
	5	4,42,16	4,34,8	4,2,40	4,1,24	4,39,48	4,30,7	4,31,32		
	6	4,12,12	4,22,4	4,21,29	4,6,20	4,32,37	4,48,45	4,4,28		
	7	4,16,42	4,38,34	4,8,10	4,15,43	4,11,18	4,27,26	4,25,2		
	1	2,23,17	2,26,40	2,34,27	2,17,14	2,30,43	2,27,4	2,11,30	Intervals: 2	
	2	2,4,42	2,22,9	2,8,45	2,28,32	2,6,19	2,16,22	2,19,6		
	3	2,7,29	2,38,3	2,35,39	2,36,26	2,9,13	2,15,16	2,10,49		
	4	2,32,48	2,40,15	2,31,2	2,33,38	2,44,25	2,1,35	2,18,12		
	5	2,3,11	2,45,34	2,2,21	2,29,1	2,13,37	2,46,47	2,42,24		
	6	2,24,5	2,21,28	2,39,8	2,5,44	2,25,31	2,20,41	2,12,18		
	7	2,14,23	2,37,46	2,47,33	2,49,20	2,48,7	2,41,10	2,43,36		
	1	3,24,17	3,20,39	3,6,7	3,30,44	3,17,12	3,27,34	3,35,22	Intervals: 3	
	2	3,1,8	3,14,30	3,34,47	3,22,42	3,21,3	3,15,25	3,19,20		
	3	3,10,6	3,13,28	3,23,38	3,36,33	3,43,43	3,4,16	3,49,11		
	4	3,28,35	3,7,1	3,31,18	3,5,13	3,2,23	3,46,45	3,3,40		
	5	3,26,46	3,18,19	3,29,29	3,32,24	3,37,41	3,11,14	3,38,2		
	6	3,8,26	3,40,48	3,39,9	3,12,4	3,33,21	3,48,36	3,16,31		
	7	3,42,37	3,25,10	3,47,27	3,9,15	3,44,32	3,41,5	3,45,49		
	1	1,40,1	1,10,43	1,38,8	1,34,15	1,13,29	1,5,36	1,27,22	Intervals: 4	
	2	1,24,6	1,1,48	1,32,13	1,14,20	1,19,34	1,23,41	1,45,27		
	3	1,35,5	1,49,47	1,9,12	1,29,19	1,31,33	1,8,40	1,43,26		
	4	1,20,4	1,15,46	1,6,11	1,7,18	1,22,32	1,42,39	1,48,25		
	5	1,12,7	1,39,49	1,44,14	1,18,21	1,37,35	1,28,42	1,16,28		
	6	1,30,2	1,47,44	1,33,9	1,36,16	1,11,30	1,17,37	1,3,23		
	7	1,41,3	1,26,45	1,2,10	1,46,17	1,4,31	1,25,38	1,21,24		
		1	2	3	4	5	6	7	Times	

```

### Check properties of the design
wheat.canon <- designAnatomy(formulae = list(lab = ~ Intervals/(Runs*Times),
                                             field = ~ Blocks/Plots,
                                             treats = ~ Lines),
                             data = lab.lay)
summary(wheat.canon, which.criteria = c("aefficiency", "order"))

##
##
## Summary table of the decomposition for lab, field & treats (based on adjusted quantities)
##
## Source.lab          df1 Source.field  df2 Source.treats df3 aefficiency order
## Intervals          3 Blocks          3              1.0000  1

```

```
## Runs[Intervals]      24 Plots[Blocks]  24 Lines      24      0.2500      1
## Times[Intervals]     24 Plots[Blocks]  24 Lines      24      0.2500      1
## Runs#Times[Intervals] 144 Plots[Blocks] 144 Lines     48      0.7500      1
##                      Residual        96      1.0000      1
##
## The design is not orthogonal
```

Given, the nonorthogonality of Blocks:Plots evident in the anatomy, redo the table with just the lab and field tiers to investigate.

```
##### Check confounding of field sources with lab sources
wheat.labfield.canon <- designAnatomy(formulae = list(lab = ~ Intervals/(Runs*Times),
                                                    field = ~ Blocks/Plots),
                                     data = lab.lay)
summary(wheat.labfield.canon, which.criteria = c("aefficiency", "order"))

##
##
## Summary table of the decomposition for lab & field
##
## Source.lab      df1 Source.field  df2 aefficiency order
## Intervals       3 Blocks         3      1.0000      1
## Runs[Intervals] 24 Plots[Blocks] 24      1.0000      1
## Times[Intervals] 24 Plots[Blocks] 24      1.0000      1
## Runs#Times[Intervals] 144 Plots[Blocks] 144      1.0000      1
```

4.4.2 Question

1. Is the variance matrix for this experiment based on two sets of terms that are orthogonal?

Because all plots sources are confounded orthogonally with analyses sources, the variance matrix is indeed based on two sets of terms that are orthogonal.

4.5 Elaborate, two-phase, sensory experiment

Brien and Payne (1999) describe a two-phase sensory experiment, of which the first, or field, phase is a viticultural experiment and the second, or evaluation, phase involves the assessment of wine made from the produce of the first phase plots. In the field phase, two adjacent Youden squares are used to assign trellis treatments to the plots, a plot being a row-column combination within a square. Each plot is divided into two halfplots and two methods of pruning assigned at random to them. In the second phase, the halfplots from the field phase are randomized, using two Latin squares and an extended Youden design, to glasses in positions on a table for evaluation by judges. This experiment is Example 1.2 from Bailey and Brien (2015), where its anova, along with expected mean squares, is given. Its factor-allocation diagram is in Figure 20.

4.5.1 Check the properties of the randomized layout

Load the layout and use `designAnatomy` to check the properties of the design.

```
##### Load the layout
data("Sensory3PhaseShort.dat")

##### Examine the properties of the design
sensory.canon <- designAnatomy(formulae = list(eval = ~((Occ/Int/Sit)*Jud)/Posn,
                                              field = ~ (Row*(Sqr/Col))/Hplot,
```

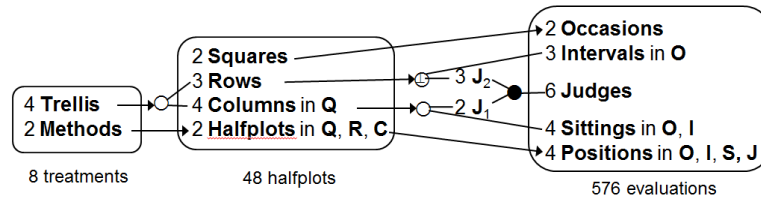


Figure 20: Factor-allocation diagram for the two-phase sensory experiment: treatments are randomized to halfplots, which are, in turn, randomized to evaluations; the arrow to the ‘O’, the ‘O’ and the two lines from the ‘O’ indicate that Trellis is randomized to the combinations of Rows and Columns using a nonorthogonal design; the single arrow between Methods and Halfplots indicates that Methods is randomized to Halfplots; the single arrows between the two right hand panels indicate that Squares are randomized to Occasions and Halfplots are randomized to Positions; J_1 and J_2 are two pseudofactors on Judges that split them into two sets of three; the Rows are randomized to the combinations of Intervals and J_2 using an orthogonal design, as indicated by the ‘ \oplus ’, and Columns are randomized to the combinations of J_1 and Sittings using a nonorthogonal design, as indicated by the ‘O’; Q= Squares; R = Rows; C = Columns; O = Occasions; I = Intervals; S = Sittings; J = Judges.

```

                                treats = ~Trel*Meth),
                                data     = Sensory3PhaseShort.dat)
summary(sensory.canon, which.criteria =c("aefficiency", "order"))

##
##
## Summary table of the decomposition for eval, field & treats (based on adjusted quantities)
##
## Source.eval      df1 Source.field      df2 Source.treats df3 aefficiency order
## Occ              1 Sqr                  1              1.0000      1
## Int[Occ]          4
## Sit[Occ:Int]      18 Col[Sqr]           6 Trel          3      0.0370      1
## Residual          12 Residual          3      0.3333      1
## Jud              5
## Occ#Jud           5
## Int#Jud[Occ]      20 Row                2              1.0000      1
## Row#Sqr           2              1.0000      1
## Residual          16
## Sit#Jud[Occ:Int]  90 Col[Sqr]           6 Trel          3      0.0741      1
## Residual          3      0.6667      1
## Row#Col[Sqr]      12 Trel          3      0.8889      1
## Residual          9      1.0000      1
## Posn[Occ:Int:Sit:Jud] 432 Hplot[Row:Sqr:Col] 24 Meth          1      1.0000      1
## Trel#Meth          3      1.0000      1
## Residual          20      1.0000      1
## Residual          408
##
## The design is not orthogonal

```

Note that 1/3 of Sqr:Col is partially confounded with Occ:Int:Sit and 2/3 with Occ:Int:Sit:Jud. Also, 1/9 of Trel is partially confounded with Sqr:Col and 8/9 with Row:Sqr:Col. The canonical efficiency factor for Trel in the two Sqr:Col sources is obtained by multiplying the canonical efficiency of 1/9 for Trel with that for the particular Sqr:Col source, yielding canonical efficiencies of 1/27 and 2/27.

4.5.2 Questions

1. Which is the nonorthogonal source amongst the field sources (`Source.field`) and what is its interblock and intrablock efficiency factors?

The only nonorthogonal field source is Sqr.Col. Its interblock efficiency factor is 1/3 and its intrablock efficiency factor is 2/3.

2. How would an intrablock analysis be achieved using, say, regression software?

To achieve an intrablock analysis requires careful specification of the order of fitting terms; a nonorthogonal source should not be estimated until after all nonorthogonal terms with which it is confounded, except the last, have been estimated. For this experiment, terms should be fitted in the following order:

$$\text{Occ*Jud} + \text{Row} + \text{Occ:Int}/(\text{Int} + \text{Sit}) + \text{Sqr.Col} + \text{Trel} + \text{Row:Sqr:Col} + \text{Occ:Int:Sit:Jud} + \text{Meth} + \text{Trel:Meth} + \text{Row:Sqr:Col:Hplot}.$$

This will leave a Residual that corresponds to Occ:Int:Sit:Jud:Posn.

5 Power and sample size for designed experiments

In power and sample size calculations, in addition to specifying delta, sigma, power and alpha, one has to supply a number of quantities that vary with the design of the experiment. The following table summarizes these for the common designs, giving the degrees of freedom of the denominator as a function of r , the pure replication of the treatments. Note that rm is the number of replicates in means being compared. For treatment means, this will be the product of r and a multiple, m , for the product of the number of levels of factors not involved in means being compared.

Design	m	rm	$df.num (\nu_1)$	$df.denom (\nu_2)$
CRD	1	r	$t - 1$	$t(r - 1)$
RCBD	1	b	$t - 1$	$(t - 1)(b - 1)$
LSD	1	$r(= t)$	$r - 1$	$(r - 1)(r - 2)$
Factorial				
A	b	br	$a - 1$	CRD $ab(r - 1)$
B	a	ar	$b - 1$	RCBD $(ab - 1)(r - 1)$
A:B	1	r	$(a - 1)(b - 1)$	LSD $(r - 1)(r - 2)$
Standard split-plot				
A	b	br	$a - 1$	$(a - 1)(r - 1)$
B	a	ar	$b - 1$	$(b - 1)(r - 1)$
A:B	1	r	$(a - 1)(b - 1)$	$a(b - 1)(r - 1)^\dagger$

[†]only approximate for effects not at the same level of A

5.1 Computing the power for given sample size

The function `power.exp` from the `dae` library is used for computing the power in detecting the difference between means for some, not necessarily proper, subset of the factors from a designed experiment.

The usage and arguments for this function are:

`power.exp(rm=5, df.num=1, df.denom=10, delta=1, sigma=1, alpha=0.05, print=FALSE)=`

rm: the number of observations used in computing a mean.

df.num: the degrees of freedom of the numerator of the F for testing the term involving the means;

df.denom: the degrees of freedom of the denominator of the F for testing the term involving the means;

delta: the true difference between a pair of means;

sigma: population standard deviation;

alpha: the significance level to be used;

print: T or F to have or not have a table of power calculation details printed out.

Note that the values given for the arguments in the above expression for `power.exp` are the default values assigned to the arguments if they are not set in a call to the function.

5.2 Example: Penicillin yield

Consider the penicillin example taken from [Box et al. \(2005\)](#). Suppose it was expected that the minimum difference between a pair of treatment means is 5 and that $\alpha = 0.05$. In the analysis of variance for this experiment, the Residual MSq was 18.83 so we will take $\sigma^2 \approx 20$. Also, $r = 5$ and $m = 1$. The `power.exp` call to compute the power, and the resulting output, is given below. Note that alpha is not set in this call and so the default value of 0.05 will be used. Also, the expressions `3 * (rm - 1)` and `sqrt(20)` will be evaluated prior to the call to the function. To get the correct value of `rm` used in evaluating the expression `3 * (rm - 1)`, `rm` needs to be set prior to calling `power.exp`.

```
rm <- 5
power.exp(rm=rm, df.num=3, df.denom=3*(rm-1), delta=5, sigma=sqrt(20), print=TRUE)

##   rm df.num df.denom alpha delta   sigma lambda   powr
##  1  5      3      12  0.05     5  4.472136  3.125  0.2159032
## [1] 0.2159032
```

That is, the power of the experiment is just over 0.2.

5.3 Computing the sample size to achieve specified power

The function `no.reps` from the `dae` library is used to compute the required number of pure replicates, r , of the treatments in a designed experiment to achieve a specified power in detecting a difference between the means for some, not necessarily proper, subset of the treatment factors. The usage and arguments for this function are as follows:

```
no.reps(multiple=1, df.num=1,
        df.denom=expression((df.num+1)*(r-1)),
        delta=1, sigma=1, alpha=0.05, power =0.8,
        tol = 0.025, print=FALSE)
```

multiple: the multiplier, m , which when multiplied by the number of pure replicates of a treatment, r , gives the number of observations (rm) used in computing means for the treatment factor subset; m is the replication arising from other treatment factors. However, for single treatment factor experiments the subset can only be the treatment factor and $m = 1$;

df.num: the degrees of freedom of the numerator of the F for testing the term involving the treatment factor subset;

df.denom: an expression for the degrees of freedom of the denominator of the F for testing the term involving the treatment factor subset it must involve r , the number of pure replicates, can involve other arguments to `no.reps` such as `multiple` and `df.num`, and must be enclosed in an expression function so that it is not evaluated when `no.reps` is called but will be evaluated as different values of r are tried during execution of `no.reps`;

delta: the true difference between a pair of means for some, not necessarily proper, subset of the treatment factors;

sigma: population standard deviation;

alpha: the significance level to be used;

power: the minimum power to be achieved;

tol: the maximum difference tolerated between the power required and the power computed in determining the number of replicates;

print: T or F to have or not have a table of power calculation details printed out.

5.4 Example II.1. Penicillin yield (continued)

We now determine the number of replicates required to achieve a power of 0.80 in detecting $\Delta = 5$ with $\alpha = 0.05$. We continue to take $\sigma^2 \approx 20$. The use of `no.reps` and the resulting output is as follows:

```
no.reps(multiple=1, df.num=3, df.denom=expression(df.num*(r-1)), delta=5,
        sigma=sqrt(20), power=0.8, print=FALSE)
```

```
## $nreps
## [1] 19
##
## $power
## [1] 0.8055926
```

That is, 19 reps will achieve a little more than the required power of 0.80.

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