

THE DESIGN AND MIXED-MODEL ANALYSIS OF EXPERIMENTS

PRACTICAL IX SOLUTIONS

IX.1 Let \mathbf{Y} be a n -vector of jointly-distributed random variables with

$$\psi = E[\mathbf{Y}] = \mathbf{X}_{AB}(\alpha\beta) \text{ and } \mathbf{V} = \sigma_U^2 \mathbf{I}_n,$$

where $(\alpha\beta)$ is the ab -vector of parameters specifying a different mean response for each A-B combinations,

\mathbf{X}_{AB} is the $n \times ab$ matrix equal to $\mathbf{I}_a \otimes \mathbf{I}_b \otimes \mathbf{1}_r$ that gives the combinations of A and B for each element of \mathbf{Y} ,

σ_U^2 is the variability arising from different units.

Prove that Then $\hat{\psi} = \mathbf{A}\mathbf{B}$ where $\mathbf{A}\mathbf{B}$ is the n -vector of means for the combinations of A and B, respectively.

The model $\psi = E[\mathbf{Y}] = \mathbf{X}_{AB}(\alpha\beta)$ is of full rank so that $\hat{\psi} = \mathbf{X}_{AB}(\mathbf{X}_{AB}'\mathbf{X}_{AB})^{-1}\mathbf{X}_{AB}'\mathbf{Y}$. Now, for the observations arranged in standard order of A then B then the replicates,

$$\begin{aligned} \mathbf{X}_{AB}(\mathbf{X}_{AB}'\mathbf{X}_{AB})^{-1}\mathbf{X}_{AB}' &= \mathbf{I}_a \otimes \mathbf{I}_b \otimes \mathbf{1}_r \left[(\mathbf{I}_a \otimes \mathbf{I}_b \otimes \mathbf{1}_r)' (\mathbf{I}_a \otimes \mathbf{I}_b \otimes \mathbf{1}_r) \right]^{-1} \mathbf{I}_a \otimes \mathbf{I}_b \otimes \mathbf{1}_r' \\ &= \mathbf{I}_a \otimes \mathbf{I}_b \otimes \mathbf{1}_r [\mathbf{I}_a \otimes \mathbf{I}_b \otimes \mathbf{1}_r' \mathbf{1}_r]^{-1} \mathbf{I}_a \otimes \mathbf{I}_b \otimes \mathbf{1}_r' \\ &= \frac{1}{r} \mathbf{I}_a \otimes \mathbf{I}_b \otimes \mathbf{1}_r (\mathbf{I}_a \otimes \mathbf{I}_b) \mathbf{I}_a \otimes \mathbf{I}_b \otimes \mathbf{1}_r' \\ &= \frac{1}{r} (\mathbf{I}_a \otimes \mathbf{I}_b \otimes \mathbf{1}_r) (\mathbf{I}_a \otimes \mathbf{I}_b \otimes \mathbf{1}_r') \\ &= \frac{1}{r} \mathbf{I}_a \otimes \mathbf{I}_b \otimes \mathbf{1}_r \mathbf{1}_r' \\ &= \frac{1}{r} \mathbf{I}_a \otimes \mathbf{I}_b \otimes \mathbf{J}_r \end{aligned}$$

and so $\hat{\psi}$ is clearly the n -vector of means for the combinations of A and B as claimed.

IX.2 Let \mathbf{P}_G , \mathbf{P}_A , \mathbf{P}_B , \mathbf{R}_G , \mathbf{R}_A and \mathbf{R}_B be as defined in lemma IX.2. Also, let $\mathbf{P}_{AB} = r^{-1} \mathbf{I}_a \otimes \mathbf{I}_b \otimes \mathbf{J}_r$ and $\mathbf{R}_{AB} = \mathbf{I} - \mathbf{P}_{AB}$.

- a) Prove that $\mathbf{R}_{AB} \mathbf{R}_B \mathbf{R}_A \mathbf{R}_G = \mathbf{R}_{AB}$, given that $\mathbf{P}_{AB} \mathbf{P}_A = \mathbf{P}_A \mathbf{P}_{AB} = \mathbf{P}_A$, $\mathbf{P}_{AB} \mathbf{P}_B = \mathbf{P}_B \mathbf{P}_{AB} = \mathbf{P}_B$ and $\mathbf{P}_{AB} \mathbf{P}_G = \mathbf{P}_G \mathbf{P}_{AB} = \mathbf{P}_G$

$$\begin{aligned} \mathbf{R}_{AB} \mathbf{R}_B \mathbf{R}_A \mathbf{R}_G &= (\mathbf{I} - \mathbf{P}_{AB})(\mathbf{I} - \mathbf{P}_B) \mathbf{R}_A \mathbf{R}_G \\ &= (\mathbf{I} - \mathbf{P}_{AB} - \mathbf{P}_B + \mathbf{P}_{AB} \mathbf{P}_B) \mathbf{R}_A \mathbf{R}_G \\ &= (\mathbf{I} - \mathbf{P}_{AB}) \mathbf{R}_A \mathbf{R}_G \\ &= (\mathbf{I} - \mathbf{P}_{AB}) \mathbf{R}_G \\ &= \mathbf{R}_{AB} \end{aligned}$$

- b) Prove that $\mathbf{P}_{AB} \mathbf{R}_B \mathbf{R}_A \mathbf{R}_G = \mathbf{P}_{AB} \mathbf{R}_A \mathbf{R}_B \mathbf{R}_G = \mathbf{P}_{AB} - \mathbf{P}_A - \mathbf{P}_B + \mathbf{P}_G$.

First note that since by lemma IX.1 $\mathbf{P}_A \mathbf{P}_B = \mathbf{P}_B \mathbf{P}_A$, $\mathbf{R}_A \mathbf{R}_B = (\mathbf{I} - \mathbf{P}_A)(\mathbf{I} - \mathbf{P}_B) = \mathbf{I} - \mathbf{P}_A - \mathbf{P}_B - \mathbf{P}_A \mathbf{P}_B = \mathbf{I} - \mathbf{P}_A - \mathbf{P}_B - \mathbf{P}_B \mathbf{P}_A = \mathbf{R}_B \mathbf{R}_A$. This proves that $\mathbf{P}_{AB} \mathbf{R}_B \mathbf{R}_A \mathbf{R}_G = \mathbf{P}_{AB} \mathbf{R}_A \mathbf{R}_B \mathbf{R}_G$.

Next using lemmas IX.1 and IX.2

$$\begin{aligned} \mathbf{P}_{AB} \mathbf{R}_A \mathbf{R}_B \mathbf{R}_G &= \mathbf{P}_{AB} (\mathbf{I} - \mathbf{P}_A)(\mathbf{I} - \mathbf{P}_B)(\mathbf{I} - \mathbf{P}_G) \\ &= (\mathbf{P}_{AB} - \mathbf{P}_A)(\mathbf{I} - \mathbf{P}_B - \mathbf{P}_G + \mathbf{P}_G) \\ &= (\mathbf{P}_{AB} - \mathbf{P}_A)(\mathbf{I} - \mathbf{P}_B) \\ &= \mathbf{P}_{AB} - \mathbf{P}_A - \mathbf{P}_{AB} \mathbf{P}_B + \mathbf{P}_A \mathbf{P}_B \\ &= \mathbf{P}_{AB} - \mathbf{P}_A - \mathbf{P}_B + \mathbf{P}_G \end{aligned}$$

IX.3 Let \mathbf{Y} be a n -vector of jointly-distributed random variables and $R(\alpha\beta | \alpha, \beta) = (\mathbf{A} \cdot \mathbf{B}_e)' (\mathbf{A} \cdot \mathbf{B}_e)$ where $\mathbf{A} \cdot \mathbf{B}_e = \mathbf{P}_{AB} \mathbf{R}_A \mathbf{R}_B \mathbf{R}_G \mathbf{Y} = (\mathbf{P}_{AB} - \mathbf{P}_A - \mathbf{P}_B + \mathbf{P}_G) \mathbf{Y}$.

Prove that the degrees of freedom of $R(\alpha\beta | \alpha, \beta)$ is $(a-1)(b-1)$ where a is the number of levels of A and b is the number of levels of B.

First use the direct product expressions of the \mathbf{P} s to show that $\text{trace}(\mathbf{P}_G) = 1$, $\text{trace}(\mathbf{P}_A) = a$, $\text{trace}(\mathbf{P}_B) = b$ and $\text{trace}(\mathbf{P}_{AB}) = ab$.

Now

$$\begin{aligned} \text{trace}(\mathbf{P}_{AB} - \mathbf{P}_A - \mathbf{P}_B + \mathbf{P}_G) &= \text{trace}(\mathbf{P}_{AB}) - \text{trace}(\mathbf{P}_A) - \text{trace}(\mathbf{P}_B) + \text{trace}(\mathbf{P}_G) \\ &= ab - a - b + 1 \\ &= (a-1)(b-1) \end{aligned}$$

IX.4 Suppose that an experiment is to be conducted to investigate the effect on the maize yield of two factors, Varieties (A, B and C) and Fertilizer treatments (Control, P, PM, PNK where P, M, N and K stand for Phosphorus, Manure, Nitrogen and Potash, respectively). It is planned to use a completely randomized design to assign the 12 Variety-Fertilizer combinations and it is estimated that the variance for this experiment will be about 4.

- a) It is desired to be able to detect a difference of at least 5 in the overall differences between a pair of Fertilizers. How many replicates should be observed if the power is to be 0.90 and the level of significance 0.05?

In this case you are concerned with the Fertilizer main effect. The columns of ANOVAPower.xls are filled out as follows:

Column	Column heading	Cell contents
A	sample size (r)	3
B	alpha	0.05
C	DF numerator	3
D	DF denominator	=12*(A5-1)
E	central F	=FINV(B5,C5,D5)
F	no. values in a mean (m)	=3*A5
G	delta	5
H	standard deviation	2
I	lambda	=F5*G5*G5/2/H5/H5
J	power	=1-NCF(E5,C5,D5,I5)

The completed worksheet has the following values:

sample size (r)	alpha	DF numerator	DF denominator	central F	no. values in a mean (m)	delta	standard deviation	lambda	power
3	0.05	3	24	3.0088	9	5	2.000000	28.125	0.9906

So 3 replicates needed to detect, with 90% power, a minimum change of 5 in the overall Fertilizer difference.

- b) What power would be achieved with the number of replicates that you have computed in detecting a change, between two varieties, of at least 5 in the difference between a pair of Fertilizers? How many replicates would be required to detect this latter difference with power 0.90 and significance level 0.05?

In this case, you are concerned with the power of an interaction effect. You need to change DF numerator to 6 and the no. values in a mean (m) to be equal to the sample size (r). The completed worksheet has the following values:

sample size (r)	alpha	DF numerator	DF denominator	central F	no. values in a mean (m)	delta	standard deviation	lambda	power
3	0.05	6	24	2.5082	3	5	2.000000	9.375	0.4887
7	0.05	6	72	2.2274	7	5	2.000000	21.875	0.9410

The number of replicates for the main effect was computed to be 3 and this achieves only a power of 0.4887 in detecting the interaction effect. To achieve a power of at least 0.90 requires 7 replicates.

- c) Use Genstat to obtain a randomized layout for this experiment with the number of replicates you finally computed in b). Use a seed of 552312 in generating the design and set up the treatment factors with the letter codes above by using them as labels.

The number of replicates computed was 7 and so we generate a completely randomized design with two factors Variety and Fertilizers with 3 and 4 levels and a Dummy factor with 7 levels. The steps given in the lecture notes were followed and the following Genstat output generated.

Genstat 5 Release 4.1 (PC/Windows NT) 04 April 2000 10:49:39
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Genstat 5 Fourth Edition - (for Windows)
Genstat 5 Procedure Library Release PL11

```

3  DESIGN
4  %WSPREAD Plots,Variety,Fertiliz
5  DELETE [redefine=yes] Variety,Fertiliz
6  FACTOR [modify=yes;nvalues=84;levels=3;labels=!t('A','B','C')] Variety
7  READ Variety; frepresentation=ordinal

Identifier      Values      Missing      Levels
Variety          84           0           3

11 FACTOR [modify=yes;nvalues=84;levels=4;labels=!t('Control','P','PM','PNK')\
12 ] Fertiliz
13 READ Fertiliz; frepresentation=ordinal

Identifier      Values      Missing      Levels
Fertiliz        84           0           4

```

```
18 PDESIGN [BLOCK=Plots; TREAT=Variety,Fertilizer]
```

```
*** Treatment combinations on each unit of the design ***
```

Plots

1	2	2
2	3	2
3	1	2
4	1	3
5	2	3
6	1	1
7	1	3
8	3	3
9	1	3
10	3	4
11	2	4
12	2	4
13	2	3
14	1	3
15	1	2
16	3	2
17	2	1
18	2	1
19	3	4
20	1	4
21	2	2
22	1	4
23	1	4
24	3	4
25	1	4
26	3	2
27	3	2
28	1	4
29	1	1
30	3	2
31	2	4
32	1	4
33	3	2
34	1	1
35	3	4
36	3	3
37	1	2
38	2	1
39	1	1
40	2	3
41	1	1
42	2	1
43	3	4
44	3	1
45	3	1
46	3	3
47	3	1
48	2	4
49	2	3
50	2	4
51	2	3
52	2	3
53	2	2
54	3	2
55	3	1
56	2	4
57	1	1
58	1	2
59	3	1
60	3	3
61	2	1
62	1	3
63	2	4
64	3	4
65	2	1

```

66  3 3
67  1 2
68  1 1
69  2 3
70  1 2
71  1 3
72  3 3
73  2 2
74  2 1
75  3 4
76  2 2
77  1 4
78  1 2
79  3 1
80  3 1
81  3 3
82  2 2
83  1 3
84  2 2

```

Treatment factors are listed in the order: Variety Fertiliz

Labels of Variety:

```

1      A
2      B
3      C

```

Labels of Fertiliz:

```

1      Control
2      P
3      PM
4      PNK

```

IX.5 A completely randomized experiment was conducted to investigate the effect of vitamin B₁₂ (0, 5 mg) and antibiotics (0, 40 mg) fed to swine. The response was the average daily gain in weight.

		Vitamin B ₁₂	
		0	5
Antibiotics	0	1.30	1.26
		1.19	1.21
		1.08	1.19
	40	1.05	1.52
		1.00	1.56
		1.05	1.55

What are the components of this experiment?

1. Observational unit – a swine
2. Response variable – Weight gain
3. Unrandomized factors – Swines
4. Randomized factors – Antibiotics, Vitamin
5. Type of study – Factorial CRD

What is the experimental structure for this experiment?

Structure	Formula
unrandomized	12 Swines
randomized	2 Antibiotic*2 Vitamin

What are the expected mean squares for the lines in the analysis of variance table based on the same dichotomization of the factors into random/fixed factors as for unrandomized/randomized factors?

Source	df	E[MSq]
Swines	11	
Antibiotic	1	$\sigma_S^2 + f_A(\psi)$
Vitamin	1	$\sigma_S^2 + f_V(\psi)$
Antibiotic.Vitamin	1	$\sigma_S^2 + f_{AV}(\psi)$
Residual	8	σ_S^2
Total	11	

Obtain the usual analysis for a two-factor factorial experiment using Genstat, including diagnostic checking and the examination of treatment differences.

Also perform an analysis of the data that involves a nested factorial structure by identifying an appropriate grouping of the treatments using your previous analysis.

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```

3  "Data taken from File: D:/ANALYSES/LM/MULTIFAC/FAC2SWIN.GSH"
4  DELETE [redefine=yes] Swine,Antibiot,Vitamin,WtGain
5  FACTOR [modify=yes;nvalues=12;levels=12] Swine
6  READ Swine; frepresentation=ordinal

```

Identifier	Values	Missing	Levels
Swine	12	0	12

```

8  FACTOR [modify=yes;nvalues=12;levels=!(0,5)] Antibiot
9  READ Antibiot; frepresentation=ordinal

```

Identifier	Values	Missing	Levels
Antibiot	12	0	2

```

11 FACTOR [modify=yes;nvalues=12;levels=!(0,40)] Vitamin
12 READ Vitamin; frepresentation=ordinal

```

```

Identifier   Values   Missing   Levels
Vitamin      12         0         2

14 VARIATE [nvalues=12] WtGain
15 READ WtGain

Identifier   Minimum   Mean   Maximum   Values   Missing
WtGain       1.000    1.247    1.560     12       0

17
18 PRINT Swine,Antibiot,Vitamin,WtGain

    Swine      Antibiot      Vitamin      WtGain
      1         0.000         0.00         1.300
      2         0.000        40.00         1.260
      3         0.000         0.00         1.190
      4         0.000        40.00         1.210
      5         0.000         0.00         1.080
      6         0.000        40.00         1.190
      7         5.000         0.00         1.050
      8         5.000        40.00         1.520
      9         5.000         0.00         1.000
     10         5.000        40.00         1.560
     11         5.000         0.00         1.050
     12         5.000        40.00         1.550

19 BLOCK Swine
20 TREAT Antibiot*Vitamin
21 ANOVA [FPROB=Y; PSE=LSD] WtGain

21.....

***** Analysis of variance *****

Variate: WtGain

Source of variation      d.f.        s.s.        m.s.        v.r.    F pr.

Swine stratum
Antibiot                  1    0.020833    0.020833     5.68    0.044
Vitamin                   1    0.218700    0.218700    59.65    <.001
Antibiot.Vitamin          1    0.172800    0.172800    47.13    <.001
Residual                   8    0.029333    0.003667
Total                     11    0.441667

* MESSAGE: the following units have large residuals.

Swine 1          0.110    s.e. 0.049
Swine 5         -0.110    s.e. 0.049

***** Tables of means *****

Variate: WtGain

Grand mean  1.247

Antibiot      0.00      5.00
              1.205    1.288

Vitamin       0.00     40.00
              1.112    1.382

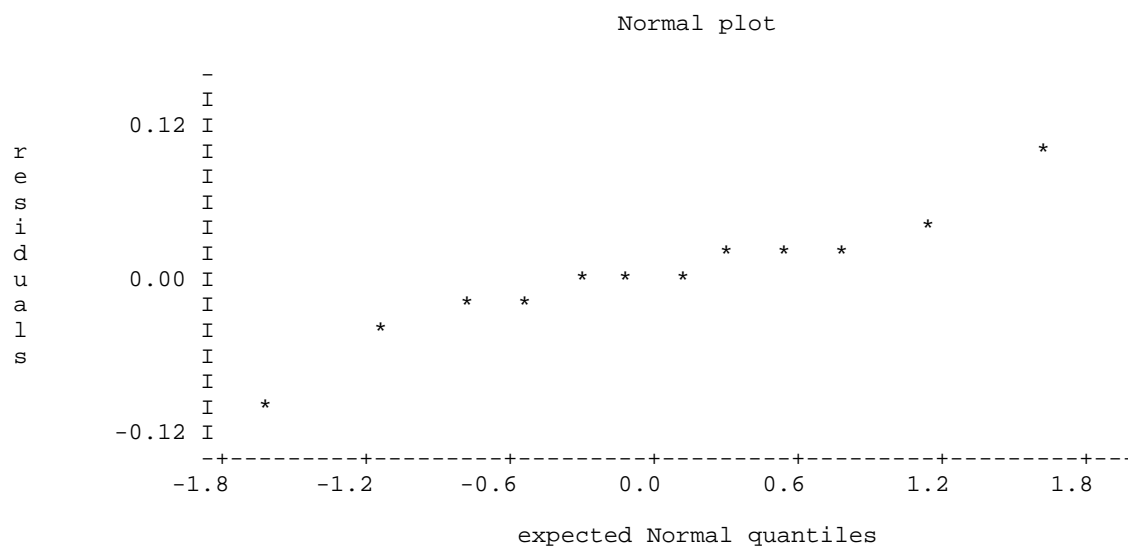
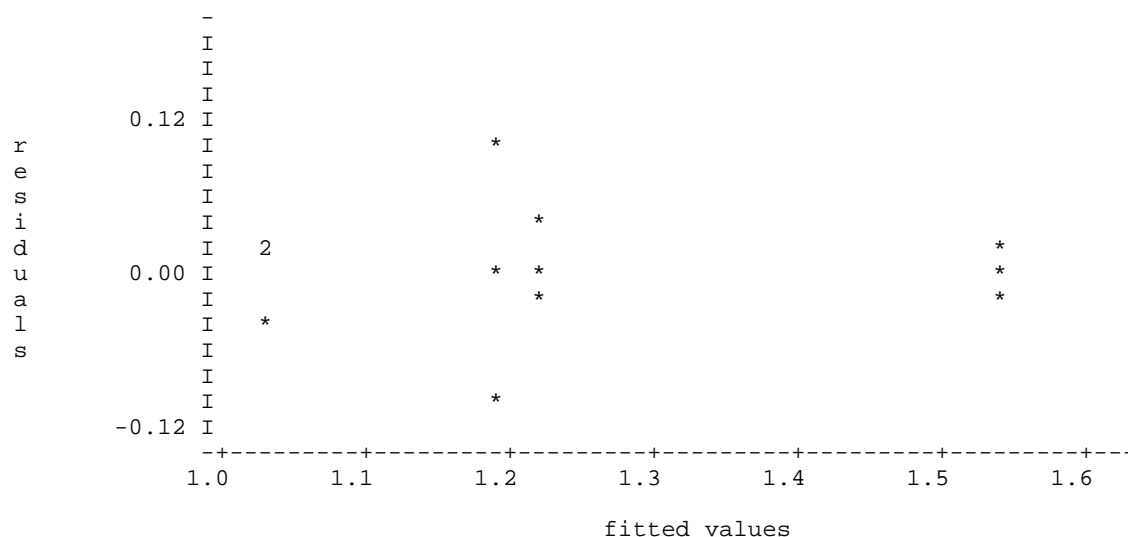
Antibiot Vitamin      0.00     40.00
  0.00              1.190    1.220
  5.00              1.033    1.543

```


*** Least significant differences of means (5% level) ***

Table	Antibiot	Vitamin	Antibiot Vitamin
rep.	6	6	3
d.f.	8	8	8
l.s.d.	0.0806	0.0806	0.1140

22 APLLOT METHOD=fit,normal



As this is a CRD we cannot perform Tukey's nonadditivity test on the Residual.

The residuals-versus-fitted-values plot displays an appropriate pattern except that one treatment (no Vitamin, no Antibiotic) appears to have a somewhat higher spread compared to the others. This is reflected in the Normal Probability plot where there are two extreme residuals. Otherwise the plot looks satisfactory. Without further evidence of the reason for these outliers, it is Advisable to leave them in the analysis.

The significance test for the example is as follows:

Step 1: Set up hypotheses

a) $H_0: \alpha_0 = \alpha_{40}$

H_1 : the population Antibiotic means are different

b) $H_0: \rho_0 = \rho_5$

H_1 : the population Vitamin means are different

c) H_0 : there is no interaction between Antibiotic and Vitamin

$$((\alpha\mu)_{ij} - (\alpha\mu)_{i.} - (\alpha\mu)_{.j} + (\alpha\mu)_{..}) = 0 \quad \text{for all } i, j$$

H_1 : there is an interaction between Antibiotic and Vitamin

$$((\alpha\mu)_{ij} - (\alpha\mu)_{i.} - (\alpha\mu)_{.j} + (\alpha\mu)_{..}) \neq 0 \quad \text{for some } i, j$$

Step 2: Calculate test statistics

The analysis of variance table for a two-factor CRD, with random factors being the unrandomized factors and fixed factors the randomized factors, is:

Source	df	SSq	MSq	E[MSq]	F	Prob
Swine	11	0.4417				
Antibiotic	1	0.0208	0.0208	$\sigma_S^2 + f_A(\psi)$	5.68	0.0443
Vitamin	1	0.2187	0.2187	$\sigma_S^2 + f_V(\psi)$	59.65	<0.001
Antibiotic.Vitamin	1	0.1728	0.1728	$\sigma_S^2 + f_{AV}(\psi)$	47.13	<0.001
Residual	8	0.0293	0.0037	σ_S^2		

Step 3: Decide between hypotheses

There is a significant interaction between Antibiotic and Vitamin level.

Because there is a significant interaction the means for each combination of Vitamin and Antibiotic are examined using a multiple comparison procedure.

		Vitamin B ₁₂		LSD(5%)
		0	5	
Antibiotics	0	1.19	1.22	0.1140
	40	1.03	1.54	

Examination of the means reveals that Vitamin has no effect with 0 Antibiotic, but there is a significant extra weight gain when Vitamin is used with 40mg of Antibiotic.

However, it appears that there might be a response when both Antibiotic is at 40 and Vitamin B₁₂ is at 5 and not a response with other combinations. To investigate this possibility we set up an analysis with a nested factorial structure.

```
33  FACPRODUCT FACTORS=!p(Antibiot,Vitamin); PRODUCT=Treats
34  FACTOR [LEV=2] Both
35  CALC Both=NEWLEVELS(Treats; !v(2,2,2,1))
36  BLOCK Swine
37  TREAT Both/Treats
38  ANOVA [FPROB=Y; PSE=LSD] WtGain
```

```
38.....
```

```
***** Analysis of variance *****
```

```
Variate: WtGain
```

Source of variation	d.f.	s.s.	m.s.	v.r.	F pr.
Swine stratum					
Both	1	0.352044	0.352044	96.01	<.001
Both.Treats	2	0.060289	0.030144	8.22	0.011
Residual	8	0.029333	0.003667		
Total	11	0.441667			

```
* MESSAGE: the following units have large residuals.
```

Swine 1	0.110	s.e. 0.049
Swine 5	-0.110	s.e. 0.049

```
***** Tables of means *****
```

```
Variate: WtGain
```

```
Grand mean 1.247
```

Both	1	2		
	1.543	1.148		
rep.	3	9		
Both	Treats	1	2	3
1				1.543
2		1.190	1.220	1.033

```
*** Least significant differences of means (5% level) ***
```

Table	Both	Both Treats
rep.	unequal	3
d.f.	8	8
l.s.d.	0.0931	0.1140

The analysis indicates that our proposed model does not fit the data as there are difference between the three other treatment combinations.

IX.6 An experiment was conducted on the survival of *Salmonella typhimurium* under three different levels of sorbic acid and six levels of water activity (a_w). A randomized complete block design was used with three blocks of eighteen plates to which the eighteen treatment combinations are applied. The data analyzed are the log(density/ml) measured seven days after the imposition of treatments.

Sorbic acid	a_w	Block		
		I	II	III
0	0.98	3604.72	4315.64	4146.42
	0.94	772.78	812.41	518.01
	0.90	354.25	395.44	464.05
	0.86	157.59	210.61	149.90
	0.82	127.74	74.44	91.84
	0.78	74.44	76.71	66.69
100	0.98	2079.74	2416.32	1978.31
	0.94	678.58	487.85	671.83
	0.90	149.90	196.37	323.76
	0.86	127.74	141.17	72.97
	0.82	72.97	83.93	65.37
	0.78	62.18	80.64	65.37
200	0.98	1261.43	1012.32	1326.10
	0.94	561.16	482.99	620.17
	0.90	181.27	164.02	228.15
	0.86	82.27	81.45	120.30
	0.82	70.81	71.52	79.04
	0.78	138.38	61.56	63.43

What are the components of this experiment?

1. Observational unit - a plate
2. Response variable - Density
3. Unrandomized factors - Blocks, Plates
4. Randomized factors - Activity, Sorbic
5. Type of study - Two-factor RCBD

What is the experimental structure for this experiment?

Structure	Formula
unrandomized	3 Blocks/18 Plates
randomized	6 Activity*3 Sorbic

What are the expected mean squares for the lines in the analysis of variance table based on the same division of the factors into random/fixed factors as for unrandomized/randomized factors?

Source	df	E[MSq]
Blocks	2	$\sigma_{BP}^2 + 18\sigma_B^2$
Blocks.Plates	51	
Activity	5	$\sigma_{BP}^2 + f_A(\psi)$
Sorbic	2	$\sigma_{BP}^2 + f_S(\psi)$
Activity.Sorbic	10	$\sigma_{BP}^2 + f_{AS}(\psi)$
Residual	34	σ_{BP}^2
Total	53	

The data is available in the file *Fac2Salm.gsh* in the directory *G:\Disciplinal\Genstat*. Analyze the data using Genstat, including diagnostic checking and the examination of treatment differences. If there is evidence of transformable non-additivity, use the YTRANSFORM procedure in Genstat to determine the best power transformation for this data. Reanalyze the data using the transformation that you identify.

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```
3 "Data taken from File: D:/ANALYSES/LM/MULTIFAC/FAC2SALM.GSH"
4 DELETE [redefine=yes] Block,Plates,Activity,Sorbic,Density,LnDensit
5 FACTOR [modify=yes;nvalues=54;levels=3] Block
6 READ Block; frepresentation=ordinal
```

Identifier	Values	Missing	Levels
Block	54	0	3

```
9 FACTOR [modify=yes;nvalues=54;levels=18] Plates
10 READ Plates; frepresentation=ordinal
```

Identifier	Values	Missing	Levels
Plates	54	0	18

```
13 FACTOR [modify=yes;nvalues=54;levels=!(0.98,0.94,0.9,0.86,0.82,0.78)\
14 ] Activity
15 READ Activity; frepresentation=ordinal
```

Identifier	Values	Missing	Levels
Activity	54	0	6

```
18 FACTOR [modify=yes;nvalues=54;levels=!(0,100,200)] Sorbic
19 READ Sorbic; frepresentation=ordinal
```

Identifier	Values	Missing	Levels
Sorbic	54	0	3

```
22 VARIATE [nvalues=54] Density
23 READ Density
```

Identifier	Minimum	Mean	Maximum	Values	Missing	Skew
Density	61.6	606.9	4315.6	54	0	

```

35  VARIATE [nvalues=54] LnDensit
36  READ LnDensit

```

Identifier	Minimum	Mean	Maximum	Values	Missing
LnDensit	4.120	5.519	8.370	54	0

```

41
42  PRINT Block,Plates,Sorbic,Activity,Density

```

Block	Plates	Sorbic	Activity	Density
1	1	0.0	0.9800	3604.7
2	1	0.0	0.9800	4315.6
3	1	0.0	0.9800	4146.4
1	2	0.0	0.9400	772.8
2	2	0.0	0.9400	812.4
3	2	0.0	0.9400	518.0
1	3	0.0	0.9000	354.2
2	3	0.0	0.9000	395.4
3	3	0.0	0.9000	464.1
1	4	0.0	0.8600	157.6
2	4	0.0	0.8600	210.6
3	4	0.0	0.8600	149.9
1	5	0.0	0.8200	127.7
2	5	0.0	0.8200	74.4
3	5	0.0	0.8200	91.8
1	6	0.0	0.7800	74.4
2	6	0.0	0.7800	76.7
3	6	0.0	0.7800	66.7
1	7	100.0	0.9800	2079.7
2	7	100.0	0.9800	2416.3
3	7	100.0	0.9800	1978.3
1	8	100.0	0.9400	678.6
2	8	100.0	0.9400	487.8
3	8	100.0	0.9400	671.8
1	9	100.0	0.9000	149.9
2	9	100.0	0.9000	196.4
3	9	100.0	0.9000	323.8
1	10	100.0	0.8600	127.7
2	10	100.0	0.8600	141.2
3	10	100.0	0.8600	73.0
1	11	100.0	0.8200	73.0
2	11	100.0	0.8200	83.9
3	11	100.0	0.8200	65.4
1	12	100.0	0.7800	62.2
2	12	100.0	0.7800	80.6
3	12	100.0	0.7800	65.4
1	13	200.0	0.9800	1261.4
2	13	200.0	0.9800	1012.3
3	13	200.0	0.9800	1326.1
1	14	200.0	0.9400	561.2
2	14	200.0	0.9400	483.0
3	14	200.0	0.9400	620.2
1	15	200.0	0.9000	181.3
2	15	200.0	0.9000	164.0
3	15	200.0	0.9000	228.1
1	16	200.0	0.8600	82.3
2	16	200.0	0.8600	81.5
3	16	200.0	0.8600	120.3
1	17	200.0	0.8200	70.8
2	17	200.0	0.8200	71.5
3	17	200.0	0.8200	79.0
1	18	200.0	0.7800	138.4
2	18	200.0	0.7800	61.6
3	18	200.0	0.7800	63.4

```

43 BLOCK Block/Plates
44 TREAT POL(Activity; 2)*POL(Sorbic; 2)
45 ANOVA [FPROB=Y; PSE=LSD] Density

```

45.....

***** Analysis of variance *****

Variate: Density

Source of variation	d.f.	s.s.	m.s.	v.r.	F pr.
Block stratum	2	11587.	5793.	0.36	
Block.Plates stratum					
Activity	5	38997864.	7799573.	487.17	<.001
Lin	1	24082914.	24082914.	1504.23	<.001
Quad	1	11539997.	11539997.	720.79	<.001
Deviations	3	3374953.	1124984.	70.27	<.001
Sorbic	2	2785828.	1392914.	87.00	<.001
Lin	1	2671748.	2671748.	166.88	<.001
Quad	1	114080.	114080.	7.13	0.012
Activity.Sorbic	10	9696336.	969634.	60.56	<.001
Lin.Lin	1	4623625.	4623625.	288.79	<.001
Quad.Lin	1	2879858.	2879858.	179.88	<.001
Lin.Quad	1	147971.	147971.	9.24	0.005
Dev.Lin	3	1883694.	627898.	39.22	<.001
Quad.Quad	1	87311.	87311.	5.45	0.026
Deviations	3	73876.	24625.	1.54	0.222
Residual	34	544343.	16010.		
Total	53	52035958.			

* MESSAGE: the following units have large residuals.

Block 1	Plates 1	-397.	s.e. 100.
Block 2	Plates 1	280.	s.e. 100.
Block 2	Plates 7	245.	s.e. 100.

***** Tables of means *****

Variate: Density

Grand mean 607.

Activity	0.98	0.94	0.90	0.86	0.82	0.78
	2460.	623.	273.	127.	82.	77.
Sorbic	0.00	100.00	200.00			
	912.	542.	367.			
Activity	Sorbic	0.00	100.00	200.00		
0.98		4022.	2158.	1200.		
0.94		701.	613.	555.		
0.90		405.	223.	191.		
0.86		173.	114.	95.		
0.82		98.	74.	74.		
0.78		73.	69.	88.		

*** Least significant differences of means (5% level) ***

Table	Activity	Sorbic	Activity Sorbic
rep.	9	18	3
d.f.	34	34	34
l.s.d.	121.2	85.7	210.0

[illegible]

Total 53 52035958.

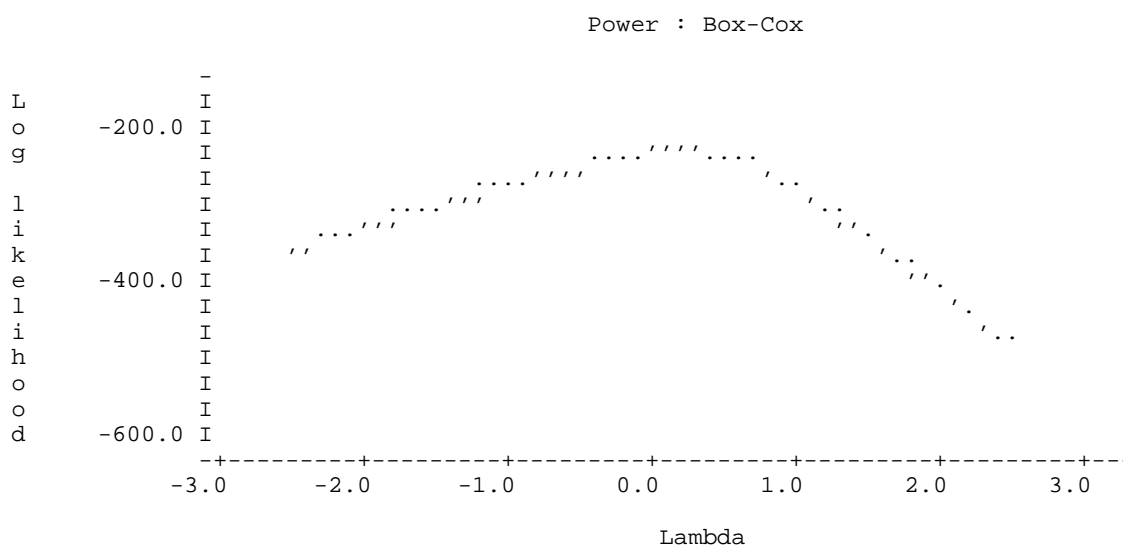
57 COVAR

We first perform the diagnostic checking. The residual-versus-fitted-values plot is displaying a funnel shape and the normal probability plot is not displaying a straight-line trend. The test for transformable non-additivity, summarized in the following table, is significant. So the assumptions are not met

Source	df	SSq	MSq	F	Prob
Block	2	11587.	5793.	0.59	
Block.Plates	51				
Activity	5	38997864.	7799573.	792.65	<.001
Sorbic	2	2785828.	1392914.	141.56	<.001
Activity.Sorbic	10	9696336.	969634.	98.54	<.001
Residual	34	544343.	16010.		
Nonadditivity	1	219628.	219628.	22.32	<.001
Deviations	33	324715.	9840		
Total	53	52035958.			

The following output is obtained using the YTRANSFORM procedure.

```
58 "Perform Box-Cox analysis"
59 YTRANSFORM [TERMS=Activity*Sorbic; LOWER=-2.5; UPPER=2.5; GRAPHICS=line] \
60 Density; SAVE=s
```



This plot indicates that the log-likelihood of λ is maximized at about 0.25. We apply this transformation to the data and reanalyze the transformed data. However, a log transformation ($\lambda = 0$) is also likely to be satisfactory.

```
61 "Analyze data raised to the power 0.25"
62 CALC TrDensity=Density**0.25
63 BLOCK Block/Plates
64 TREAT POL(Activity; 2)*POL(Sorbic; 2)
65 ANOVA [FPROB=Y; PSE=LSD] TrDensity
```

65.....

***** Analysis of variance *****

Variate: TrDensit

Source of variation	d.f.	s.s.	m.s.	v.r.	F pr.
Block stratum	2	0.00466	0.00233	0.05	
Block.Plates stratum					
Activity	5	104.81864	20.96373	441.51	<.001
Lin	1	89.25659	89.25659	1879.79	<.001
Quad	1	14.95629	14.95629	314.99	<.001
Deviations	3	0.60575	0.20192	4.25	0.012
Sorbic	2	3.65786	1.82893	38.52	<.001
Lin	1	3.49327	3.49327	73.57	<.001
Quad	1	0.16459	0.16459	3.47	0.071
Activity.Sorbic	10	4.53375	0.45337	9.55	<.001
Lin.Lin	1	2.78285	2.78285	58.61	<.001
Quad.Lin	1	0.33371	0.33371	7.03	0.012
Lin.Quad	1	0.00003	0.00003	0.00	0.982
Dev.Lin	3	1.35144	0.45048	9.49	<.001
Quad.Quad	1	0.01187	0.01187	0.25	0.620
Deviations	3	0.05385	0.01795	0.38	0.769
Residual	34	1.61439	0.04748		
Total	53	114.62930			

* MESSAGE: the following units have large residuals.

Block 1	Plates 18	0.407	s.e. 0.173
Block 3	Plates 9	0.406	s.e. 0.173

***** Tables of means *****

Variate: TrDensit

Grand mean 4.188

Activity	0.98	0.94	0.90	0.86	0.82	0.78
	6.883	4.980	4.006	3.323	2.998	2.940
Sorbic	0.00	100.00	200.00			
	4.539	4.110	3.916			
Activity	Sorbic	0.00	100.00	200.00		
0.98		7.959	6.811	5.878		
0.94		5.127	4.965	4.848		
0.90		4.480	3.828	3.711		
0.86		3.617	3.244	3.109		
0.82		3.132	2.931	2.930		
0.78		2.918	2.883	3.018		

*** Least significant differences of means (5% level) ***

Table	Activity	Sorbic	Activity Sorbic
rep.	9	18	3
d.f.	34	34	34
l.s.d.	0.2088	0.1476	0.3616

Total

53 114.62930

77 COVAR

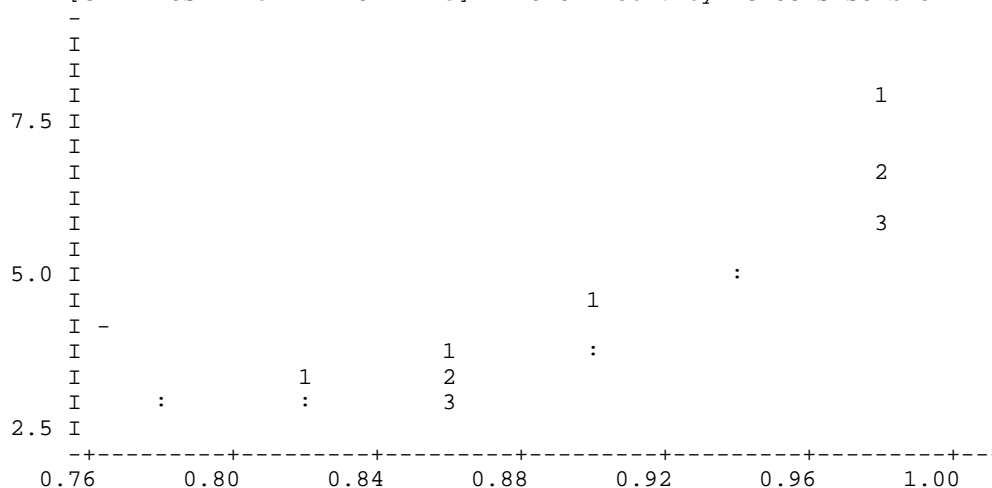
Source	df	SSq	MSq	F	Prob
Blocks	2	0.00466	0.00233	0.05	
Blocks.Plates	51				
Activity	5	104.81864	20.96373	441.51	<.001
Linear	1	89.25659	89.25659	1879.79	<.001
Quadratic	1	14.95629	14.95629	314.99	<.001
Deviations	3	0.60575	0.20192	4.25	0.012
Sorbic	2	3.65786	1.82893	38.52	<.001
Linear	1	3.49327	3.49327	73.57	<.001
Quadratic	1	0.16459	0.16459	3.47	0.071
Activity.Sorbic	10	4.53375	0.45337	9.55	<.001
A _{Linear} .B _{Linear}	1	2.78285	2.78285	58.61	<.001
A _{Quadratic} .B _{Linear}	1	0.33371	0.33371	7.03	0.012
A _{Linear} .B _{quadratic}	1	0.00003	0.00003	0.00	0.982
Deviations.B _{Linear}	1	1.35144	0.45048	9.49	<.001
A _{Quadratic} .B _{Quadratic}	1	0.01187	0.01187	0.25	0.620
Deviations	3	0.05385	0.01795	0.38	0.769
Residual	34	1.61439	0.04748		
Nonadditivity	1	0.01173	0.01173	0.24	0.626
Deviations	33	1.60266	0.04857		
Total	53	114.62930			

The diagnostic checking is now more satisfactory. The residual-versus-fitted-values and normal probability plots are acceptable, except for two outliers. The test for transformable non-additivity is now non-significant.

This analysis indicates that there are significant deviations from the $A_{\text{Quadratic}} \cdot B_{\text{Linear}}$ interaction term and so a response function of degree two does not fit the data. A plot of the Activity.Sorbic means is given in the following Genstat output.

```
78 "Examine trend in means"
```

```
79 AGRAPH [GRAPHICS=line; METHOD=line] XFACTOR=Activity; GROUPS=Sorbic
```



```

mean['0'] v. xlevel using symbol 1
mean['100'] v. xlevel using symbol 2
mean['200'] v. xlevel using symbol 3
sedbar v. ___ using symbol -

```

There appears to be a non-linear trend in the Activity means that differs between the levels of Sorbic acid. One could try to fit non-linear curves to these trends.

IX.7 The yields of an undesirable by-product of a process were measured from 12 runs in which 2 different catalysts and 2 different pressures were used; that is each combination was replicated 3 times. This experiment was repeated at two different laboratories. The data, given below, are the percentage of by-product produced and it is available in the file *Fac2ByPr.gsh* in the directory *G:\Disciplinal\Genstat*.

Catalyst Laboratory	I		II		
	A	B	A	B	
Pressure	High	53	27	40	45
		43	45	32	12
		45	57	29	69
	Low	42	32	61	54
		95	27	24	60
		60	98	11	26

The components of this experiment are:

1. Observational unit – a run
2. Response variable – % By-product
3. Unrandomized factors – Laboratories, Tests
4. Randomized factors – Catalyst, Pressure
5. Type of study – Two-factor RCBD

In this experiment interactions between laboratories and treatments are likely to be of interest so that the experimental structure for this experiment would be:

Structure	Formula
unrandomized	2 Labs/12 Tests
randomized	2 Catalyst*2 Pressure*Labs

What are the expected mean squares for the lines in the analysis of variance table based on the same dichotomization of the factors into random/fixed factors as for unrandomized/randomized factors?

Source	df	E[MSq]				
Labs	1	σ_{LT}^2	$+3\sigma_{PCL}^2$	$+6\sigma_{CL}^2$	$+6\sigma_{PL}^2$	$+12\sigma_L^2$
Labs.Tests	22					
Pressure	1	σ_{LT}^2	$+3\sigma_{PCL}^2$		$+6\sigma_{PL}^2$	$+f_P(\psi)$
Catalyst	1	σ_{LT}^2	$+3\sigma_{PCL}^2$	$+6\sigma_{CL}^2$		$+f_C(\psi)$
Pressure.Catalyst	1	σ_{LT}^2	$+3\sigma_{PCL}^2$			$+f_{PC}(\psi)$
Pressure.Labs	1	σ_{LT}^2	$+3\sigma_{PCL}^2$		$+6\sigma_{PL}^2$	
Catalyst.Labs	1	σ_{LT}^2	$+3\sigma_{PCL}^2$	$+6\sigma_{CL}^2$		
Pressure.Catalyst.Labs	1	σ_{LT}^2	$+3\sigma_{PCL}^2$			
Residual	16	σ_{LT}^2				

Analyze the data using Genstat, including diagnostic checking and the examination of treatment differences.

NOTE: Genstat will omit BLOCK terms when they occur both in BLOCK and TREAT structures. Because of this you will need to declare two Laboratories factors, Lab and Labs say, so that the BLOCK and TREAT structures would be:

BLOCK Labs/Runs
TREAT Pressure*Catalyst*Lab

Genstat 5 Release 4.1 (PC/Windows NT) 14 April 2000 12:33:26
Copyright 1998, Lawes Agricultural Trust (Rothamsted Experimental Station)

Genstat 5 Fourth Edition - (for Windows)
Genstat 5 Procedure Library Release PL11

```

3  "Data taken from File: D:/ANALYSES/LM/MULTIFAC/FAC2BYPR.GSH"
4  DELETE [redefine=yes] Labs,Runs,Lab,Pressure,Catalyst,Yield
5  FACTOR [modify=yes;nvalues=24;levels=2] Labs
6  READ Labs; frepresentation=ordinal

```

Identifier	Values	Missing	Levels
Labs	24	0	2

```

8  FACTOR [modify=yes;nvalues=24;levels=12] Runs
9  READ Runs; frepresentation=ordinal

```

Identifier	Values	Missing	Levels
Runs	24	0	12

```

11 FACTOR [modify=yes;nvalues=24;levels=2] Lab
12 READ Lab; frepresentation=ordinal

```

```

Identifier    Values    Missing    Levels
Lab           24         0          2

14 FACTOR [modify=yes;nvalues=24;levels=2] Pressure
15 READ Pressure; frepresentation=ordinal

Identifier    Values    Missing    Levels
Pressure      24         0          2

17 FACTOR [modify=yes;nvalues=24;levels=2] Catalyst
18 READ Catalyst; frepresentation=ordinal

Identifier    Values    Missing    Levels
Catalyst      24         0          2

20 VARIATE [nvalues=24] Yield
21 READ Yield

Identifier    Minimum    Mean    Maximum    Values    Missing
Yield         11.00     45.29    98.00     24         0

23
24 PRINT Labs,Runs,Pressure,Catalyst,Yield

Labs    Runs    Pressure    Catalyst    Yield
1        1        1           1           53.00
2        1        1           1           27.00
1        2        1           2           40.00
2        2        1           2           45.00
1        3        1           1           43.00
2        3        1           1           45.00
1        4        1           2           32.00
2        4        1           2           12.00
1        5        1           1           45.00
2        5        1           1           57.00
1        6        1           2           29.00
2        6        1           2           69.00
1        7        2           1           42.00
2        7        2           1           32.00
1        8        2           2           61.00
2        8        2           2           54.00
1        9        2           1           95.00
2        9        2           1           27.00
1       10        2           2           24.00
2       10        2           2           60.00
1       11        2           1           60.00
2       11        2           1           98.00
1       12        2           2           11.00
2       12        2           2           26.00

25 BLOCK Labs/Runs
26 TREAT Pressure*Catalyst*Lab
27 ANOVA [FPROB=Y; PSE=LSD] Yield

```

27.....

***** Analysis of variance *****

Variate: Yield

Source of variation	d.f.	s.s.	m.s.	v.r.	F pr.
Labs stratum					
Lab	1	12.0	12.0		
Labs.Runs stratum					
Pressure	1	360.4	360.4	0.65	0.430
Catalyst	1	1080.0	1080.0	1.96	0.181
Pressure.Catalyst	1	234.4	234.4	0.43	0.523
Pressure.Lab	1	3.4	3.4	0.01	0.939
Catalyst.Lab	1	610.0	610.0	1.11	0.308
Pressure.Catalyst.Lab	1	92.0	92.0	0.17	0.688
Residual	16	8812.7	550.8		
Total	23	11205.0			

* MESSAGE: the following units have large residuals.

Labs 2 Runs 11 45.7 s.e. 19.2

***** Tables of means *****

Variate: Yield

Grand mean 45.3

Pressure	1	2
	41.4	49.2

Catalyst	1	2
	52.0	38.6

Lab	1	2
	44.6	46.0

Pressure Catalyst	1	2
1	45.0	37.8
2	59.0	39.3

Pressure Lab	1	2
1	40.3	42.5
2	48.8	49.5

Catalyst Lab	1	2
1	56.3	47.7
2	32.8	44.3

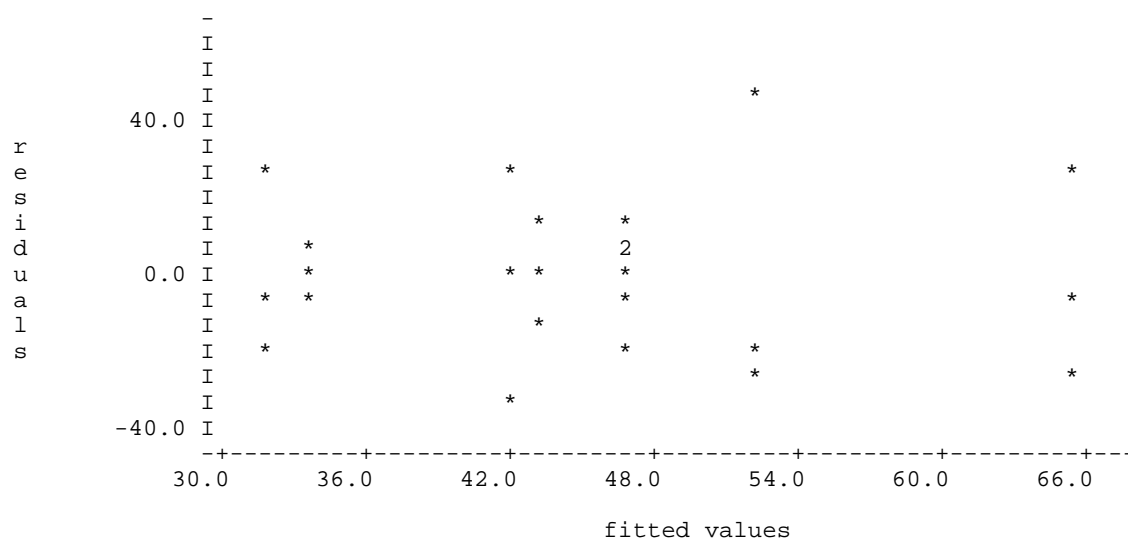
Pressure Catalyst Lab	1	2	1	2
1	47.0	43.0	33.7	42.0
2	65.7	52.3	32.0	46.7

*** Least significant differences of means (5% level) ***

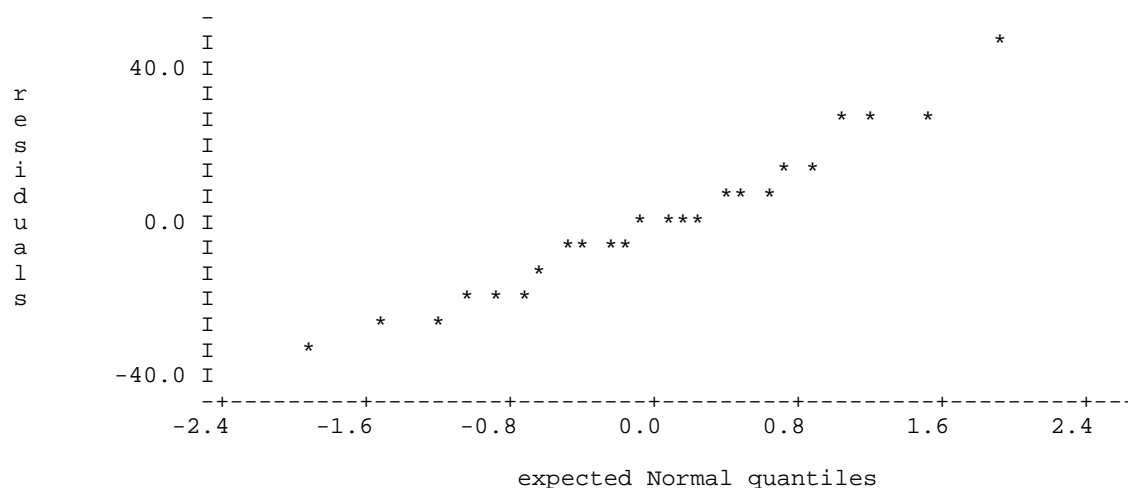
Table	Pressure	Catalyst	Lab	Pressure Catalyst
rep.	12	12	12	6
d.f.	16	16	*	16
l.s.d.	20.31	20.31	*	28.72

Table	Pressure Lab	Catalyst Lab	Pressure Catalyst Lab
rep.	6	6	3
d.f.	*	*	*
l.s.d.	*	*	*
Except when comparing means with the same level(s) of			
Lab	28.72	28.72	40.62
Pressure.Lab			40.62
Catalyst.Lab			40.62

28 APLOT METHOD=fit,normal



Normal plot



```

29 "
-30 **** Tukey's one-degree-of-freedom-for-non-additivity.
-31 **** It is the term designated covariate in the following analysis
-32 "
33 AKEEP [FIT=Fit]
34 CALC ResSq=Fit*Fit
35 ANOVA [PRINT=*] ResSq; RES=ResSq
36 COVAR ResSq "A computational trick"
37 ANOVA [PRINT=A; FPROB=Y] Yield

```

```

***** Warning (Code AN 40). Statement 1 on Line 37
Command: ANOVA [PRINT=A; FPROB=Y] Yield
Stratum variance cannot be estimated
Labs stratum has zero residual sum of squares or degrees of freedom

```

```

37.....

```

```

***** Analysis of variance (adjusted for covariate) *****

```

```

Variate: Yield
Covariate: ResSq

```

Source of variation	d.f.	s.s.	m.s.	v.r.	cov.ef.	F pr.
Labs stratum						
Lab	1	12.0	12.0		1.00	
Labs.Runs stratum						
Pressure	1	360.4	360.4	0.65	1.00	0.430
Catalyst	1	1080.0	1080.0	1.96	1.00	0.181
Pressure.Catalyst	1	234.4	234.4	0.43	1.00	0.523
Pressure.Lab	1	3.4	3.4	0.01	1.00	0.939
Catalyst.Lab	1	610.0	610.0	1.11	1.00	0.308
Pressure.Catalyst.Lab	1	92.0	92.0	0.17	1.00	0.688
Residual	16	8812.7	550.8		1.00	
Total	23	11205.0				

```

38 COVAR

```

Note the warning message about not being to estimated Labs Residual variance — it can be ignored.

Step 1: Set up hypotheses

a) $H_0: \alpha_1 = \alpha_2$
 $H_1: \alpha_1 \neq \alpha_2$

b) $H_0: \beta_1 = \beta_2$
 $H_1: \beta_1 \neq \beta_2$

c) $H_0: \text{there is no interaction between A and B}$

$$\left((\alpha\beta)_{ij} - \overline{(\alpha\beta)}_{i.} - \overline{(\alpha\beta)}_{.j} + \overline{(\alpha\beta)}_{..} \right) = 0 \quad \text{for all } i, j$$

$H_1: \text{there is an interaction between A and B}$

$$\left((\alpha\beta)_{ij} - \overline{(\alpha\beta)}_{i.} - \overline{(\alpha\beta)}_{.j} + \overline{(\alpha\beta)}_{..} \right) \neq 0 \quad \text{for some } i, j$$

d) $H_0: \sigma_{PCL}^2 = 0$
 $H_1: \sigma_{PCL}^2 \neq 0$

e) $H_0: \sigma_{PL}^2 = 0$
 $H_1: \sigma_{PL}^2 \neq 0$

f) $H_0: \sigma_{CL}^2 = 0$
 $H_1: \sigma_{CL}^2 \neq 0$

g) $H_0: \sigma_L^2 = 0$
 $H_1: \sigma_L^2 \neq 0$

Step 2: Calculate test statistics

Source	df	SSq	E[MSq]				MSq	F
Labs	1	12.0	σ_{LT}^2	$+3\sigma_{PCL}^2$	$+6\sigma_{CL}^2$	$+6\sigma_{PL}^2$	$+12\sigma_L^2$	
Labs.Tests	22							
Press	1	360.4	σ_{LT}^2	$+3\sigma_{PCL}^2$		$+6\sigma_{PL}^2$	$+f_P(\psi)$	0.65
Catal	1	1080.0	σ_{LT}^2	$+3\sigma_{PCL}^2$	$+6\sigma_{CL}^2$		$+f_C(\psi)$	1.96
Press.Catal	1	234.4	σ_{LT}^2	$+3\sigma_{PCL}^2$			$+f_{PC}(\psi)$	0.43
Press.Labs	1	3.4	σ_{LT}^2	$+3\sigma_{PCL}^2$		$+6\sigma_{PL}^2$		0.01
Catal.Labs	1	610.0	σ_{LT}^2	$+3\sigma_{PCL}^2$	$+6\sigma_{CL}^2$			1.11
Press.Catal.Labs	1	92.0	σ_{LT}^2	$+3\sigma_{PCL}^2$				0.17
Residual	16	550.8	σ_{LT}^2					0.688

Step 3: Decide between hypotheses

For P.L.C interaction

The P.L.C interaction is not significant.

For P.L, P.C and L.C interactions

The two-factor interactions should be tested against the three factor interaction. However, as the three-factor interaction is not significant and has only one degree of freedom, the two-factor interactions should be tested against the three factor interaction. They are all are not significant.

For P, L and C

For similar reasons as the two-factor interactions, the main effects should be tested against the three factor interaction. The main effects are not significant.