



Statistical Learning II

Lecture 9 - Ridge regression

Bruno Loureiro
@ CSD, DI-ENS & CNRS

brloureiro@gmail.com

Ridge regression

Note the averaged norm of the OLS is given by:

$$\mathbb{E}_{\boldsymbol{\varepsilon}} \left[\|\hat{\boldsymbol{\theta}}_{OLS}\|_2^2 \right] = \|\boldsymbol{\theta}_\star\|_2^2 + \sigma^2 \sum_{j=1}^d \frac{1}{\sigma_j^2}$$

Therefore, small s.v.s lead to larger expected norm.

Ridge regression

Note the averaged norm of the OLS is given by:

$$\mathbb{E}_{\boldsymbol{\varepsilon}} \left[\|\hat{\boldsymbol{\theta}}_{OLS}\|_2^2 \right] = \|\boldsymbol{\theta}_\star\|_2^2 + \sigma^2 \sum_{j=1}^d \frac{1}{\sigma_j^2}$$

Therefore, small s.v.s lead to larger expected norm.



Key idea: penalise the norm.

$$\min_{\boldsymbol{\theta} \in \mathbb{R}^d} \frac{1}{2n} \| \mathbf{y} - \mathbf{X} \boldsymbol{\theta} \|_2^2 + \frac{\lambda}{2} \| \boldsymbol{\theta} \|_2^2$$

Ridge regression

Note the averaged norm of the OLS is given by:

$$\mathbb{E}_{\boldsymbol{\varepsilon}} \left[\|\hat{\boldsymbol{\theta}}_{OLS}\|_2^2 \right] = \|\boldsymbol{\theta}_\star\|_2^2 + \sigma^2 \sum_{j=1}^d \frac{1}{\sigma_j^2}$$

Therefore, small s.v.s lead to larger expected norm.



Key idea: penalise the norm.

$$\min_{\boldsymbol{\theta} \in \mathbb{R}^d} \frac{1}{2n} \|\mathbf{y} - \mathbf{X}\boldsymbol{\theta}\|_2^2 + \frac{\lambda}{2} \|\boldsymbol{\theta}\|_2^2$$

Least squares
empirical risk

Regularisation or
“ridge” penalty

Ridge regression

$$\min_{\theta \in \mathbb{R}^d} \frac{1}{2n} \|y - X\theta\|_2^2 + \frac{\lambda}{2} \|\theta\|_2^2$$

Ridge regression

$$\min_{\theta \in \mathbb{R}^d} \hat{\mathcal{R}}_n^\lambda(\theta) := \frac{1}{2n} ||y - X\theta||_2^2 + \frac{\lambda}{2} ||\theta||_2^2$$

Ridge regression

$$\min_{\theta \in \mathbb{R}^d} \hat{\mathcal{R}}_n^\lambda(\theta) := \frac{1}{2n} \|y - X\theta\|_2^2 + \frac{\lambda}{2} \|\theta\|_2^2$$

Remarks: • The regularised empirical risk is a **strongly convex function** of $\theta \in \mathbb{R}^d$

$$\nabla_{\theta} \hat{\mathcal{R}}_n^\lambda(\theta) = -\frac{1}{n} X^T (y - X\theta) + \lambda \theta$$

$$\nabla_{\theta}^2 \hat{\mathcal{R}}_n^\lambda(\theta) = \frac{1}{n} X^T X + \lambda I_d > 0$$

$$(= \hat{\Sigma}_n + \lambda I_n)$$

Ridge regression

$$\min_{\theta \in \mathbb{R}^d} \hat{\mathcal{R}}_n^\lambda(\theta) := \frac{1}{2n} \|y - X\theta\|_2^2 + \frac{\lambda}{2} \|\theta\|_2^2$$

Remarks: • The regularised empirical risk is a **strongly convex function** of $\theta \in \mathbb{R}^d$

$$\nabla_{\theta} \hat{\mathcal{R}}_n^\lambda(\theta) = -\frac{1}{n} X^T (y - X\theta) + \lambda \theta$$

$$\nabla_{\theta}^2 \hat{\mathcal{R}}_n^\lambda(\theta) = \frac{1}{n} X^T X + \lambda I_d > 0$$

$$(= \hat{\Sigma}_n + \lambda I_n)$$

In other words, **minimiser** always **exist** and is **unique**.

Ridge regression

$$\min_{\theta \in \mathbb{R}^d} \hat{\mathcal{R}}_n^\lambda(\theta) := \frac{1}{2n} ||y - X\theta||_2^2 + \frac{\lambda}{2} ||\theta||_2^2$$

$$\nabla_{\theta} \hat{\mathcal{R}}_n^\lambda(\theta) = -\frac{1}{n} X^\top (y - X\theta) + \lambda \theta \stackrel{!}{=} 0$$

Ridge regression

$$\min_{\theta \in \mathbb{R}^d} \hat{\mathcal{R}}_n^\lambda(\theta) := \frac{1}{2n} \|y - X\theta\|_2^2 + \frac{\lambda}{2} \|\theta\|_2^2$$

$$\nabla_{\theta} \hat{\mathcal{R}}_n^\lambda(\theta) = -\frac{1}{n} X^\top (y - X\theta) + \lambda \theta \stackrel{!}{=} 0$$

↔

$$\left(\frac{1}{n} X^\top X \theta + \lambda I_d \right) \theta \stackrel{!}{=} \frac{1}{n} X^\top y$$

Ridge regression

$$\min_{\theta \in \mathbb{R}^d} \hat{\mathcal{R}}_n^\lambda(\theta) := \frac{1}{2n} \|y - X\theta\|_2^2 + \frac{\lambda}{2} \|\theta\|_2^2$$

$$\nabla_{\theta} \hat{\mathcal{R}}_n^\lambda(\theta) = -\frac{1}{n} X^\top (y - X\theta) + \lambda \theta \stackrel{!}{=} 0$$

↔

$$\left(\frac{1}{n} X^\top X \theta + \lambda I_d \right) \theta \stackrel{!}{=} \frac{1}{n} X^\top y$$

↔

$$\hat{\theta}_\lambda(X, y) = \frac{1}{n} \left(\frac{1}{n} X^\top X + \lambda I_d \right)^{-1} X^\top y$$

Ridge regression

$$\min_{\theta \in \mathbb{R}^d} \hat{\mathcal{R}}_n^\lambda(\theta) := \frac{1}{2n} \|y - X\theta\|_2^2 + \frac{\lambda}{2} \|\theta\|_2^2$$

The unique solution is given by:

$$\hat{\theta}_\lambda(X, y) = \frac{1}{n} \left(\frac{1}{n} X^\top X + \lambda I_d \right)^{-1} X^\top y$$

Ridge regression

$$\min_{\theta \in \mathbb{R}^d} \hat{\mathcal{R}}_n^\lambda(\theta) := \frac{1}{2n} \|y - X\theta\|_2^2 + \frac{\lambda}{2} \|\theta\|_2^2$$

The unique solution is given by:

$$\hat{\theta}_\lambda(X, y) = \frac{1}{n} \left(\frac{1}{n} X^\top X + \lambda I_d \right)^{-1} X^\top y$$



For $\lambda \rightarrow 0^+$, $\hat{\theta}_\lambda \rightarrow \hat{\theta}_{\text{OLS}}$

Ridge regression

$$\hat{\theta}_\lambda(X, \mathbf{y}) = \frac{1}{n} \left(\frac{1}{n} \mathbf{X}^\top \mathbf{X} + \lambda \mathbf{I}_d \right)^{-1} \mathbf{X}^\top \mathbf{y}$$

Remarks: • As before, consider s.v.d. of $\mathbf{X} = \sum_{j=1}^{\text{rank}(\mathbf{X})} \sigma_j \mathbf{u}_j \mathbf{v}_j^\top$

Ridge regression

$$\hat{\theta}_\lambda(X, y) = \frac{1}{n} \left(\frac{1}{n} X^\top X + \lambda I_d \right)^{-1} X^\top y$$

Remarks: • As before, consider s.v.d. of $X = \sum_{j=1}^{\text{rank}(X)} \sigma_j \mathbf{u}_j \mathbf{v}_j^\top$

$$\hat{\theta}_\lambda(X, y) = \sum_{j=1}^{\text{rank}(X)} \frac{\sigma_j}{\sigma_j^2 + n\lambda} \langle \mathbf{u}_j, y \rangle \mathbf{v}_j$$

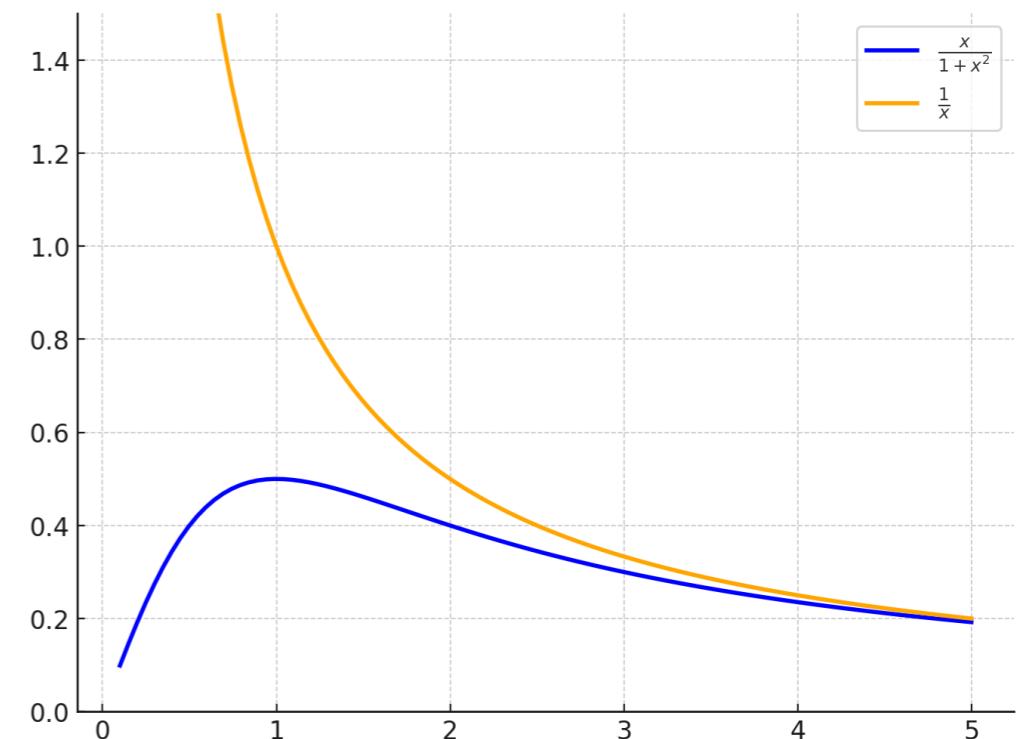
Ridge regression

$$\hat{\theta}_\lambda(X, y) = \frac{1}{n} \left(\frac{1}{n} X^\top X + \lambda I_d \right)^{-1} X^\top y$$

Remarks: • As before, consider s.v.d. of $X = \sum_{j=1}^{\text{rank}(X)} \sigma_j \mathbf{u}_j \mathbf{v}_j^\top$

$$\hat{\theta}_\lambda(X, y) = \sum_{j=1}^{\text{rank}(X)} \frac{\sigma_j}{\sigma_j^2 + n\lambda} \langle \mathbf{u}_j, y \rangle \mathbf{v}_j$$

Ridge performs **shrinkage**:
small s.v.s are suppressed!



Statistical analysis of ridge regression

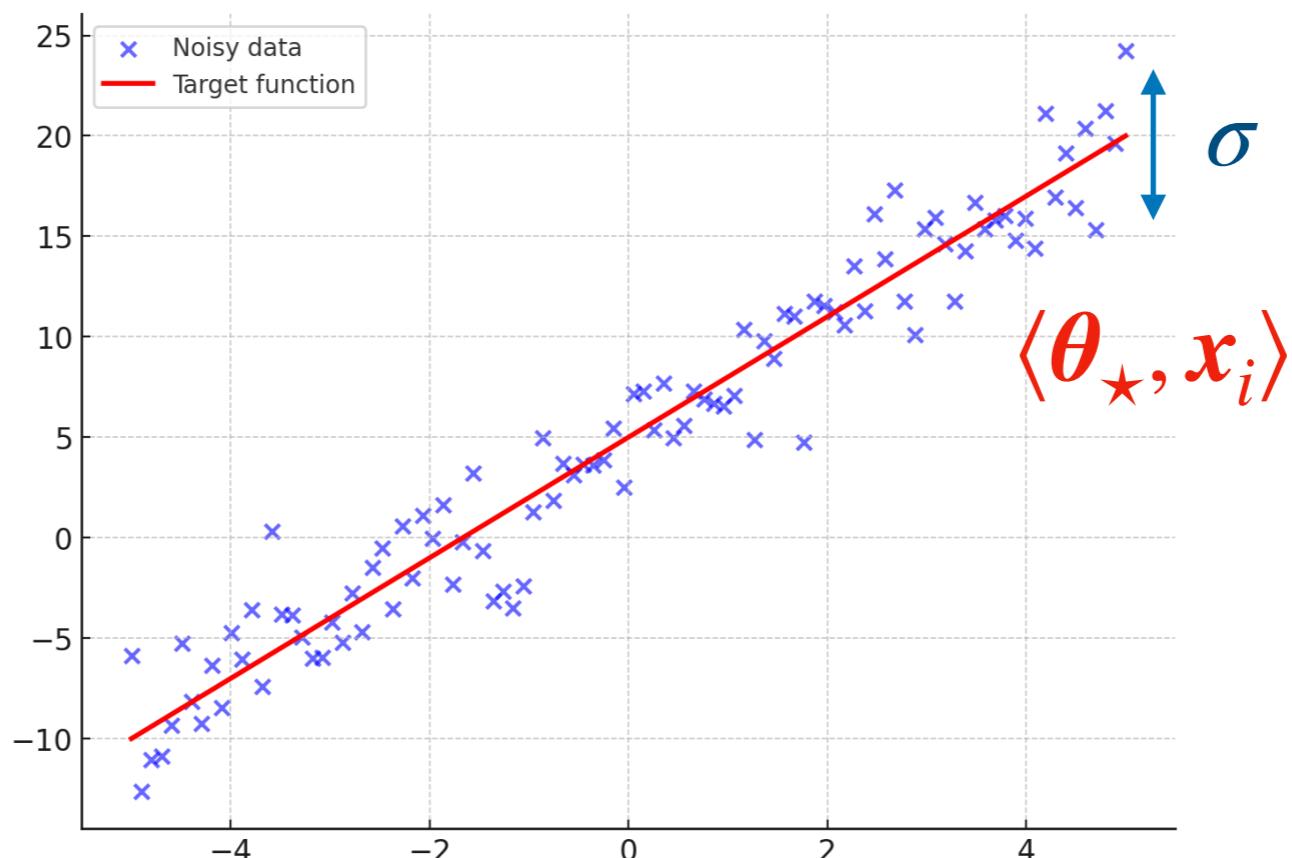
Fixed design assumption

As we did for the OLS, now let's assume:

$$y_i = \langle \theta_\star, x_i \rangle + \varepsilon_i$$

With:

- Fixed $\theta_\star \in \mathbb{R}^d$ and $x_i \in \mathbb{R}^d$ “fixed design”
- $\mathbb{E}[\varepsilon_i] = 0$ and $\mathbb{E}[\varepsilon_i^2] = \sigma^2 < \infty$



Decomposition of ridge

Given a batch of data sampled from this model:

$$\mathbf{y} = \mathbf{X}\boldsymbol{\theta}_\star + \boldsymbol{\varepsilon} \in \mathbb{R}^n$$

The ridge estimator is given by:

$$\hat{\boldsymbol{\theta}}_\lambda(\mathbf{X}, \mathbf{y}) = \frac{1}{n} \left(\hat{\boldsymbol{\Sigma}}_n + \lambda \mathbf{I}_d \right)^{-1} \mathbf{X}^\top \mathbf{y}$$

Decomposition of ridge

Given a batch of data sampled from this model:

$$\mathbf{y} = \mathbf{X}\boldsymbol{\theta}_\star + \boldsymbol{\varepsilon} \in \mathbb{R}^n$$

The ridge estimator is given by:

$$\hat{\boldsymbol{\theta}}_\lambda(\mathbf{X}, \mathbf{y}) = \frac{1}{n} \left(\hat{\boldsymbol{\Sigma}}_n + \lambda \mathbf{I}_d \right)^{-1} \mathbf{X}^\top \mathbf{y}$$

Decomposition of ridge

Given a batch of data sampled from this model:

$$\mathbf{y} = \mathbf{X}\boldsymbol{\theta}_\star + \boldsymbol{\varepsilon} \in \mathbb{R}^n$$

The ridge estimator is given by:

$$\hat{\boldsymbol{\theta}}_\lambda(\mathbf{X}, \mathbf{y}) = \frac{1}{n} \left(\hat{\boldsymbol{\Sigma}}_n + \lambda \mathbf{I}_d \right)^{-1} \mathbf{X}^\top \mathbf{y} = \frac{1}{n} \left(\hat{\boldsymbol{\Sigma}}_n + \lambda \mathbf{I}_d \right)^{-1} \mathbf{X}^\top (\mathbf{X}\boldsymbol{\theta}_\star + \boldsymbol{\varepsilon})$$

Decomposition of ridge

Given a batch of data sampled from this model:

$$\mathbf{y} = \mathbf{X}\boldsymbol{\theta}_\star + \boldsymbol{\varepsilon} \in \mathbb{R}^n$$

The ridge estimator is given by:

$$\begin{aligned}\hat{\boldsymbol{\theta}}_\lambda(\mathbf{X}, \mathbf{y}) &= \frac{1}{n} \left(\hat{\Sigma}_n + \lambda \mathbf{I}_d \right)^{-1} \mathbf{X}^\top \mathbf{y} = \frac{1}{n} \left(\hat{\Sigma}_n + \lambda \mathbf{I}_d \right)^{-1} \mathbf{X}^\top (\mathbf{X}\boldsymbol{\theta}_\star + \boldsymbol{\varepsilon}) \\ &= \left(\hat{\Sigma}_n + \lambda \mathbf{I}_d \right)^{-1} \hat{\Sigma}_n \boldsymbol{\theta}_\star + \frac{1}{n} \left(\hat{\Sigma}_n + \lambda \mathbf{I}_d \right)^{-1} \mathbf{X}^\top \boldsymbol{\varepsilon}\end{aligned}$$

Decomposition of ridge

Given a batch of data sampled from this model:

$$\mathbf{y} = \mathbf{X}\boldsymbol{\theta}_\star + \boldsymbol{\varepsilon} \in \mathbb{R}^n$$

The ridge estimator is given by:

$$\begin{aligned}\hat{\boldsymbol{\theta}}_\lambda(\mathbf{X}, \mathbf{y}) &= \frac{1}{n} \left(\hat{\Sigma}_n + \lambda \mathbf{I}_d \right)^{-1} \mathbf{X}^\top \mathbf{y} = \frac{1}{n} \left(\hat{\Sigma}_n + \lambda \mathbf{I}_d \right)^{-1} \mathbf{X}^\top (\mathbf{X}\boldsymbol{\theta}_\star + \boldsymbol{\varepsilon}) \\ &= \left(\hat{\Sigma}_n + \lambda \mathbf{I}_d \right)^{-1} \hat{\Sigma}_n \boldsymbol{\theta}_\star + \frac{1}{n} \left(\hat{\Sigma}_n + \lambda \mathbf{I}_d \right)^{-1} \mathbf{X}^\top \boldsymbol{\varepsilon} \\ &= \boldsymbol{\theta}_\star - \lambda \left(\hat{\Sigma}_n + \lambda \mathbf{I}_d \right)^{-1} \boldsymbol{\theta}_\star + \frac{1}{n} \left(\hat{\Sigma}_n + \lambda \mathbf{I}_d \right)^{-1} \mathbf{X}^\top \boldsymbol{\varepsilon}\end{aligned}$$

Decomposition of ridge

$$\hat{\theta}_\lambda(X, y) = \theta_\star - \lambda \left(\hat{\Sigma}_n + \lambda I_d \right)^{-1} \theta_\star + \frac{1}{n} \left(\hat{\Sigma}_n + \lambda I_d \right)^{-1} X^\top \boldsymbol{\varepsilon}$$

“signal”

“noise”

Decomposition of ridge

$$\hat{\theta}_\lambda(X, y) = \theta_\star - \lambda (\hat{\Sigma}_n + \lambda I_d)^{-1} \theta_\star + \frac{1}{n} (\hat{\Sigma}_n + \lambda I_d)^{-1} X^\top \boldsymbol{\varepsilon}$$

“signal”

“noise”

In particular:

- Bias: $\mathbb{E}_{\boldsymbol{\varepsilon}} [\hat{\theta}_\lambda(X, y)] = \theta_\star - \lambda (\hat{\Sigma}_n + \lambda I_d)^{-1} \theta_\star$

Decomposition of ridge

$$\hat{\theta}_\lambda(X, y) = \theta_\star - \lambda (\hat{\Sigma}_n + \lambda I_d)^{-1} \theta_\star + \frac{1}{n} (\hat{\Sigma}_n + \lambda I_d)^{-1} X^\top \boldsymbol{\varepsilon}$$

“signal”

“noise”

In particular:

- Bias: $\mathbb{E}_{\boldsymbol{\varepsilon}} [\hat{\theta}_\lambda(X, y)] = \theta_\star - \lambda (\hat{\Sigma}_n + \lambda I_d)^{-1} \theta_\star$
- Variance: $\text{Var}_{\boldsymbol{\varepsilon}} [\hat{\theta}_\lambda(X, y)] = \frac{\sigma^2}{n} (\hat{\Sigma}_n + \lambda I_d)^{-2} \hat{\Sigma}_n$

Decomposition of ridge

$$\hat{\theta}_\lambda(X, y) = \theta_\star - \lambda (\hat{\Sigma}_n + \lambda I_d)^{-1} \theta_\star + \frac{1}{n} (\hat{\Sigma}_n + \lambda I_d)^{-1} X^\top \boldsymbol{\varepsilon}$$

“signal”

“noise”

In particular:

- Bias: $\mathbb{E}_{\boldsymbol{\varepsilon}} [\hat{\theta}_\lambda(X, y)] = \theta_\star - \lambda (\hat{\Sigma}_n + \lambda I_d)^{-1} \theta_\star$
- Variance: $\text{Var}_{\boldsymbol{\varepsilon}} [\hat{\theta}_\lambda(X, y)] = \frac{\sigma^2}{n} (\hat{\Sigma}_n + \lambda I_d)^{-2} \hat{\Sigma}_n$



- Ridge is a **biased** estimator.
- Regularisation shrinks both signal and noise

Risk of ridge

Recall that in Lecture 5 we have shown that for any $\theta \in \mathbb{R}^d$:

$$\mathcal{R}(\theta) - \sigma^2 = (\theta - \theta_\star)^\top \hat{\Sigma}_n (\theta - \theta_\star)$$

Risk of ridge

Recall that in Lecture 5 we have shown that for any $\theta \in \mathbb{R}^d$:

$$\mathcal{R}(\theta) - \sigma^2 = (\theta - \theta_\star)^\top \hat{\Sigma}_n (\theta - \theta_\star)$$

Therefore, inserting the solution $\hat{\theta}_\lambda(X, y)$:

$$\begin{aligned}\mathcal{R}(\hat{\theta}_\lambda) - \sigma^2 &= \lambda^2 \theta_\star^\top (\hat{\Sigma}_n + \lambda I_d)^{-2} \hat{\Sigma}_n \theta_\star \\ &\quad + \frac{1}{n^2} \boldsymbol{\epsilon}^\top X (\hat{\Sigma}_n + \lambda I_d)^{-1} \hat{\Sigma}_n (\hat{\Sigma}_n + \lambda I_d)^{-1} X^\top \boldsymbol{\epsilon} \\ &\quad - \frac{\lambda}{n} \boldsymbol{\epsilon}^\top X (\hat{\Sigma}_n + \lambda I_d)^{-2} \hat{\Sigma}_n \theta_\star \\ &\quad - \frac{\lambda}{n} \theta_\star^\top (\hat{\Sigma}_n + \lambda I_d)^{-2} \hat{\Sigma}_n X^\top \boldsymbol{\epsilon}\end{aligned}$$

Risk of ridge

Taking the expectation with respect to the noise:

$$\mathbb{E}_{\boldsymbol{\varepsilon}}[\mathcal{R}(\hat{\boldsymbol{\theta}}_{\lambda})] - \sigma^2 = \lambda^2 \boldsymbol{\theta}_{\star}^\top (\hat{\boldsymbol{\Sigma}}_n + \lambda \mathbf{I}_d)^{-2} \hat{\boldsymbol{\Sigma}}_n \boldsymbol{\theta}_{\star} + \frac{\sigma^2}{n} \text{Tr } \hat{\boldsymbol{\Sigma}}_n^2 (\hat{\boldsymbol{\Sigma}}_n + \lambda \mathbf{I}_d)^{-2}$$

Risk of ridge

Taking the expectation with respect to the noise:

$$\mathbb{E}_{\boldsymbol{\varepsilon}}[\mathcal{R}(\hat{\boldsymbol{\theta}}_{\lambda})] - \sigma^2 = \lambda^2 \boldsymbol{\theta}_{\star}^\top (\hat{\boldsymbol{\Sigma}}_n + \lambda \mathbf{I}_d)^{-2} \hat{\boldsymbol{\Sigma}}_n \boldsymbol{\theta}_{\star} + \frac{\sigma^2}{n} \text{Tr } \hat{\boldsymbol{\Sigma}}_n^2 (\hat{\boldsymbol{\Sigma}}_n + \lambda \mathbf{I}_d)^{-2}$$

Alternatively, we can also write in terms of a bias-variance decomposition of the risk:

$$\mathbb{E}_{\boldsymbol{\varepsilon}}[\mathcal{R}(\hat{\boldsymbol{\theta}}_{\lambda})] - \sigma^2 = \mathcal{B} + \mathcal{V}$$

Where:

$$\mathcal{B} = \lambda^2 \boldsymbol{\theta}_{\star}^\top (\hat{\boldsymbol{\Sigma}}_n + \lambda \mathbf{I}_d)^{-2} \hat{\boldsymbol{\Sigma}}_n \boldsymbol{\theta}_{\star} \quad \mathcal{V} = \frac{\sigma^2}{n} \text{Tr } \hat{\boldsymbol{\Sigma}}_n^2 (\hat{\boldsymbol{\Sigma}}_n + \lambda \mathbf{I}_d)^{-2}$$

Risk of ridge

Considering the SVD of $X = \sum_{k=1}^{\text{rank}(X)} \sigma_k \mathbf{u}_k \mathbf{v}_k^\top$, we can also write:

$$\mathcal{B} = \frac{1}{n} \sum_{k=1}^{\text{rank}(X)} \frac{(n\lambda)^2 \sigma_k^2 \langle \mathbf{v}_k, \boldsymbol{\theta}_\star \rangle^2}{(\sigma_k^2 + n\lambda)^2} \quad \mathcal{V} = \sigma^2 \sum_{k=1}^{\text{rank}(X)} \frac{\sigma_k^4}{(\sigma_k^2 + n\lambda)^2}$$

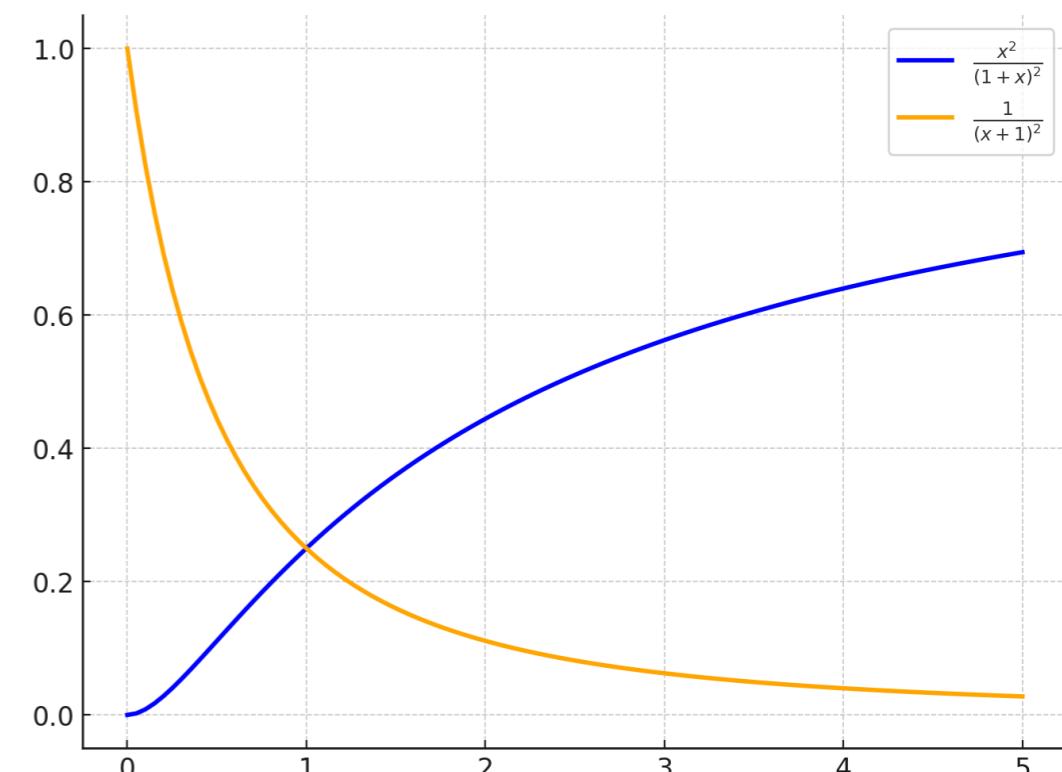
Risk of ridge

Considering the SVD of $X = \sum_{k=1}^{\text{rank}(X)} \sigma_k \mathbf{u}_k \mathbf{v}_k^\top$, we can also write:

$$\mathcal{B} = \frac{1}{n} \sum_{k=1}^{\text{rank}(X)} \frac{(n\lambda)^2 \sigma_k^2 \langle \mathbf{v}_k, \boldsymbol{\theta}_\star \rangle^2}{(\sigma_k^2 + n\lambda)^2} \quad \mathcal{V} = \sigma^2 \sum_{k=1}^{\text{rank}(X)} \frac{\sigma_k^4}{(\sigma_k^2 + n\lambda)^2}$$

Remarks:

- For $\lambda \rightarrow 0^+$, we get the OLS excess risk
- $\mathcal{B}(\lambda)$ is an increasing function of λ
- $\mathcal{V}(\lambda)$ is a decreasing function of λ



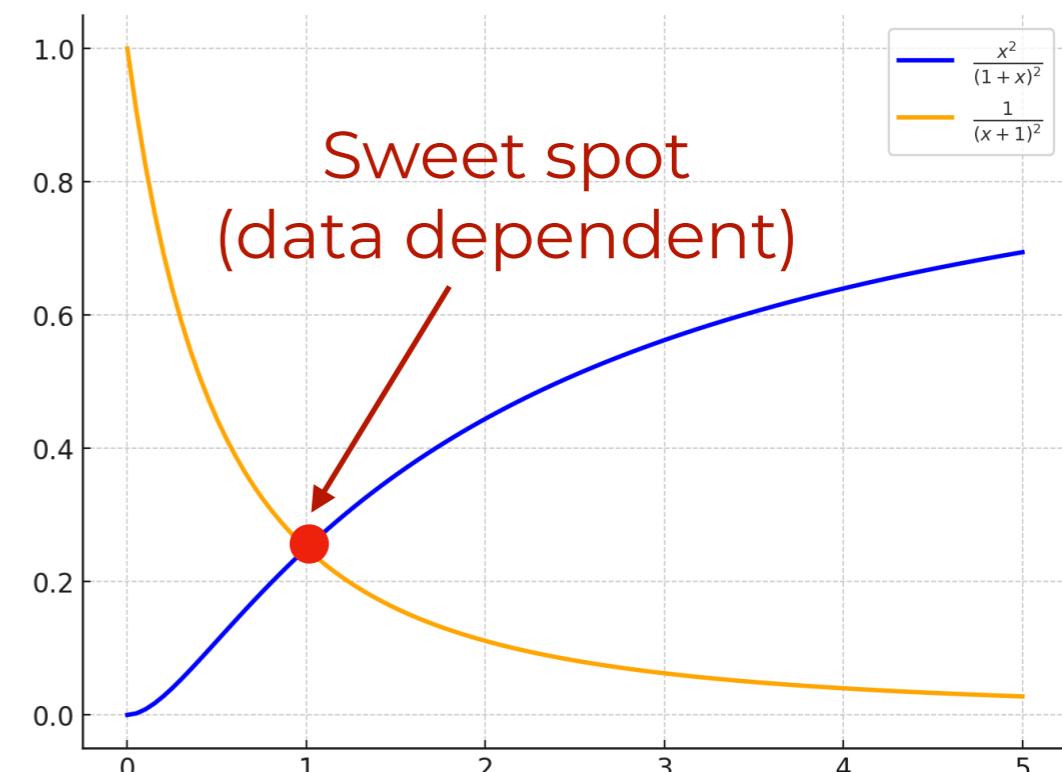
Risk of ridge

Considering the SVD of $X = \sum_{k=1}^{\text{rank}(X)} \sigma_k \mathbf{u}_k \mathbf{v}_k^\top$, we can also write:

$$\mathcal{B} = \frac{1}{n} \sum_{k=1}^{\text{rank}(X)} \frac{(n\lambda)^2 \sigma_k^2 \langle \mathbf{v}_k, \theta_\star \rangle^2}{(\sigma_k^2 + n\lambda)^2} \quad \mathcal{V} = \sigma^2 \sum_{k=1}^{\text{rank}(X)} \frac{\sigma_k^4}{(\sigma_k^2 + n\lambda)^2}$$

Remarks:

- For $\lambda \rightarrow 0^+$, we get the OLS excess risk
- $\mathcal{B}(\lambda)$ is an increasing function of λ
- $\mathcal{V}(\lambda)$ is a decreasing function of λ



Interpretation of variance

Let $A \in \mathbb{R}^{d \times d}$ be a positive definite matrix with decreasing eigenvalues $\text{spec}(A) = \{\lambda_k : k = 1, \dots, d\}$. Define the cumulative:

$$\phi(\lambda) = \#\{k : \lambda_k > \lambda\}$$

“Count eigenvalues
bigger than λ ”

Interpretation of variance

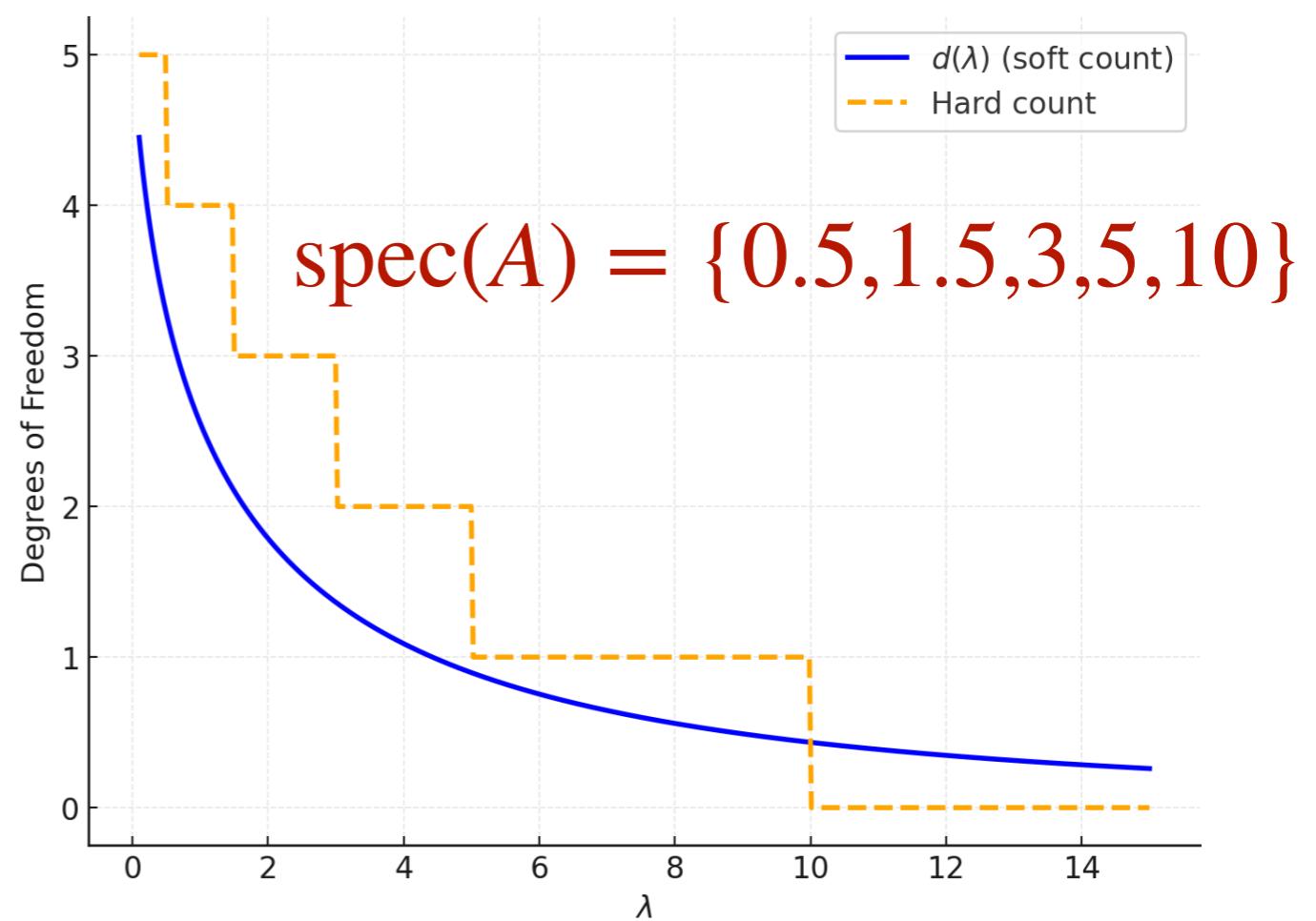
Let $A \in \mathbb{R}^{d \times d}$ be a positive definite matrix with decreasing eigenvalues $\text{spec}(A) = \{\lambda_k : k = 1, \dots, d\}$. Define the cumulative:

$$\phi(\lambda) = \#\{k : \lambda_k > \lambda\}$$

“Count eigenvalues
bigger than λ ”

The variance of the ridge risk can be seen as a soft version:

$$df_2(\lambda) = \sum_{k=1}^d \frac{\lambda_k^2}{(\lambda_k + \lambda)^2}$$



Interpretation of variance

Let $A \in \mathbb{R}^{d \times d}$ be a positive definite matrix with decreasing eigenvalues $\text{spec}(A) = \{\lambda_k : k = 1, \dots, d\}$. Define the cumulative:

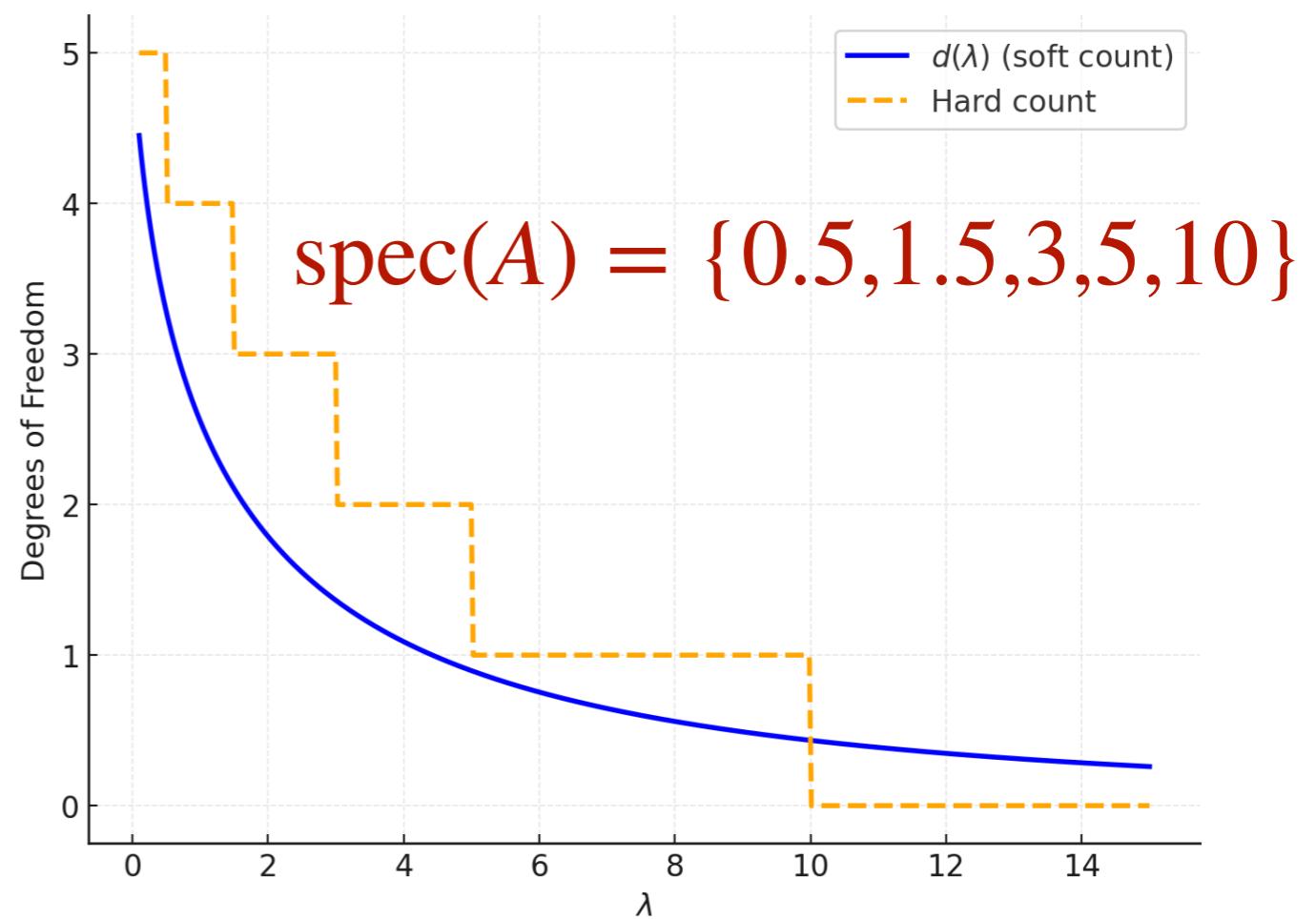
$$\phi(\lambda) = \#\{k : \lambda_k > \lambda\}$$

“Count eigenvalues
bigger than λ ”

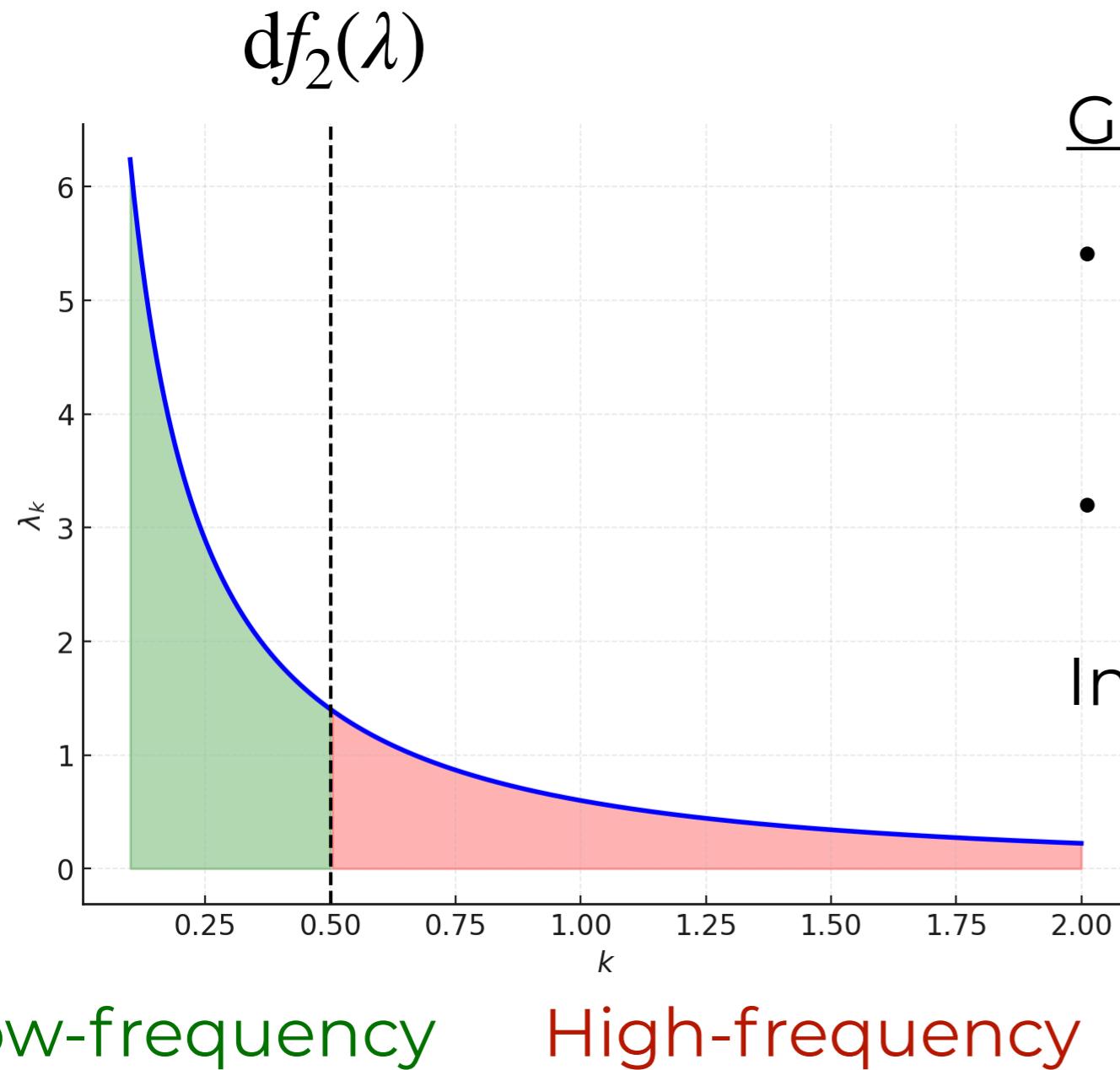
The variance of the ridge risk can be seen as a soft version:

$$df_2(\lambda) = \sum_{k=1}^d \frac{\lambda_k^2}{(\lambda_k + \lambda)^2}$$

- Fast decay: small λ
- Slow decay: large λ



Choosing regularisation



Goal: pick λ such that:

- directions in X that better correlate with θ_\star are retained
- Shrink remaining directions

In practice, **cross-validation...**