



Statistical Learning II

Lecture 10 - BSS

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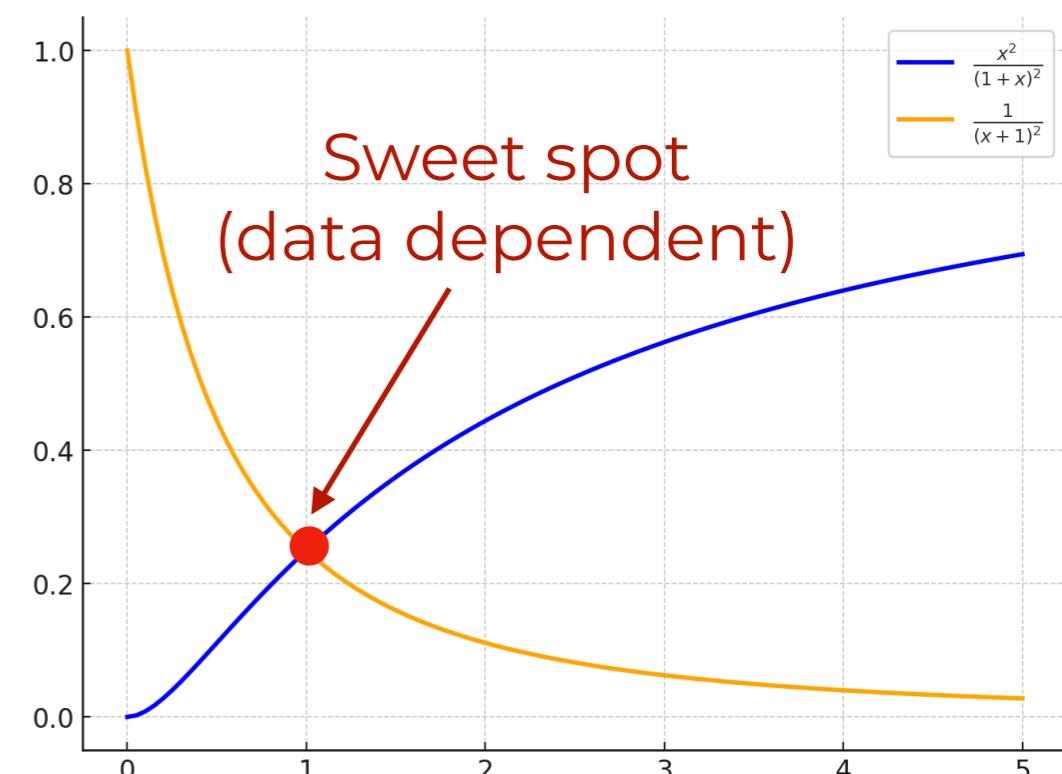
Risk of ridge

Considering the SVD of $X = \sum_{k=1}^{\text{rank}(X)} \sigma_k \mathbf{u}_k \mathbf{v}_k^\top$, we can also write:

$$\mathcal{B} = \frac{1}{n} \sum_{k=1}^{\text{rank}(X)} \frac{(n\lambda)^2 \sigma_k^2 \langle \mathbf{v}_k, \boldsymbol{\theta}_\star \rangle^2}{(\sigma_k^2 + n\lambda)^2} \quad \mathcal{V} = \sigma^2 \sum_{k=1}^{\text{rank}(X)} \frac{\sigma_k^4}{(\sigma_k^2 + n\lambda)^2}$$

Remarks:

- For $\lambda \rightarrow 0^+$, we get the OLS excess risk
- $\mathcal{B}(\lambda)$ is an increasing function of λ
- $\mathcal{V}(\lambda)$ is a decreasing function of λ



Interpretation of variance

Let $A \in \mathbb{R}^{d \times d}$ be a positive definite matrix with decreasing eigenvalues $\text{spec}(A) = \{\lambda_k : k = 1, \dots, d\}$. Define the cumulative:

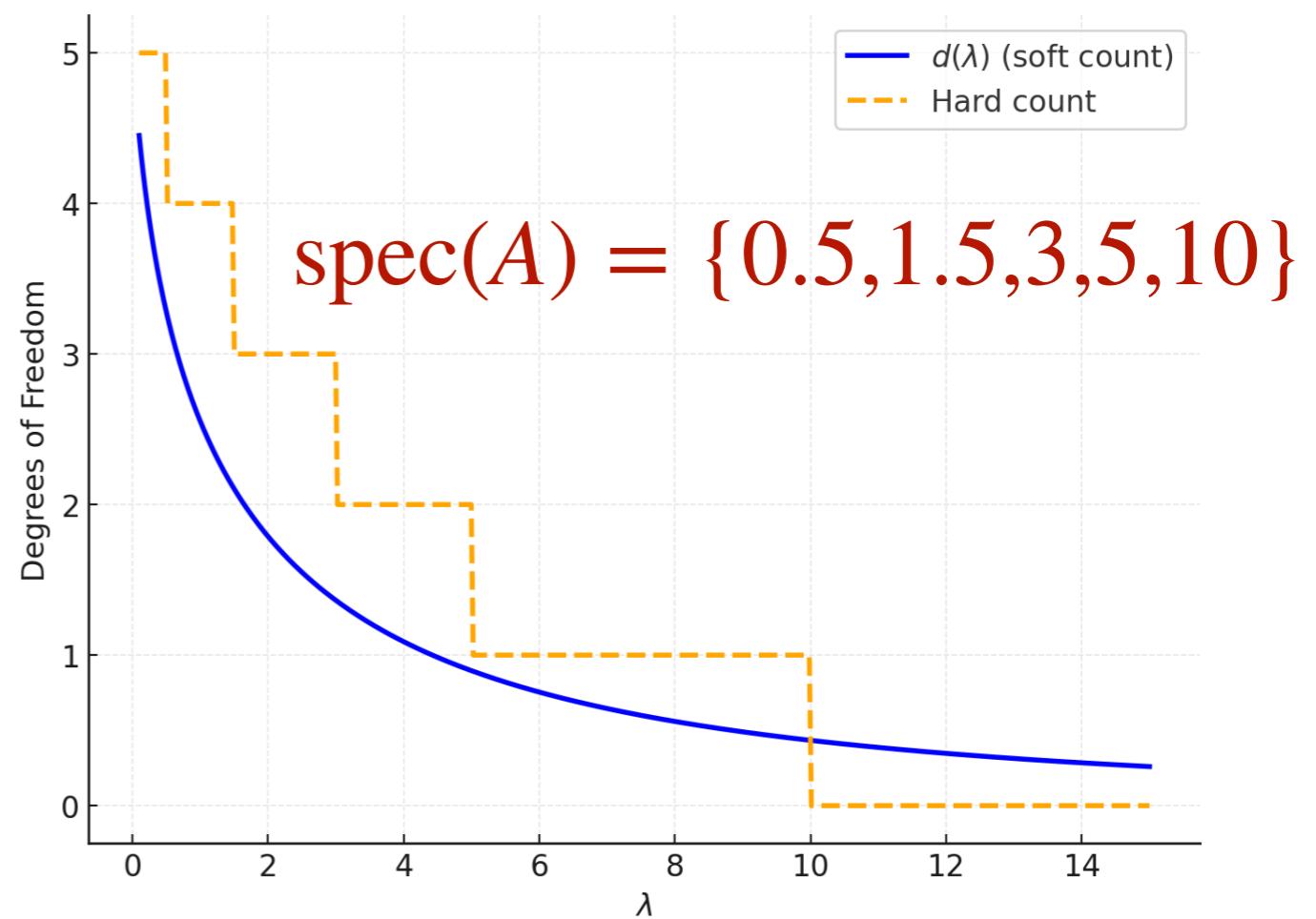
$$\phi(\lambda) = \#\{k : \lambda_k > \lambda\}$$

“Count eigenvalues
bigger than λ ”

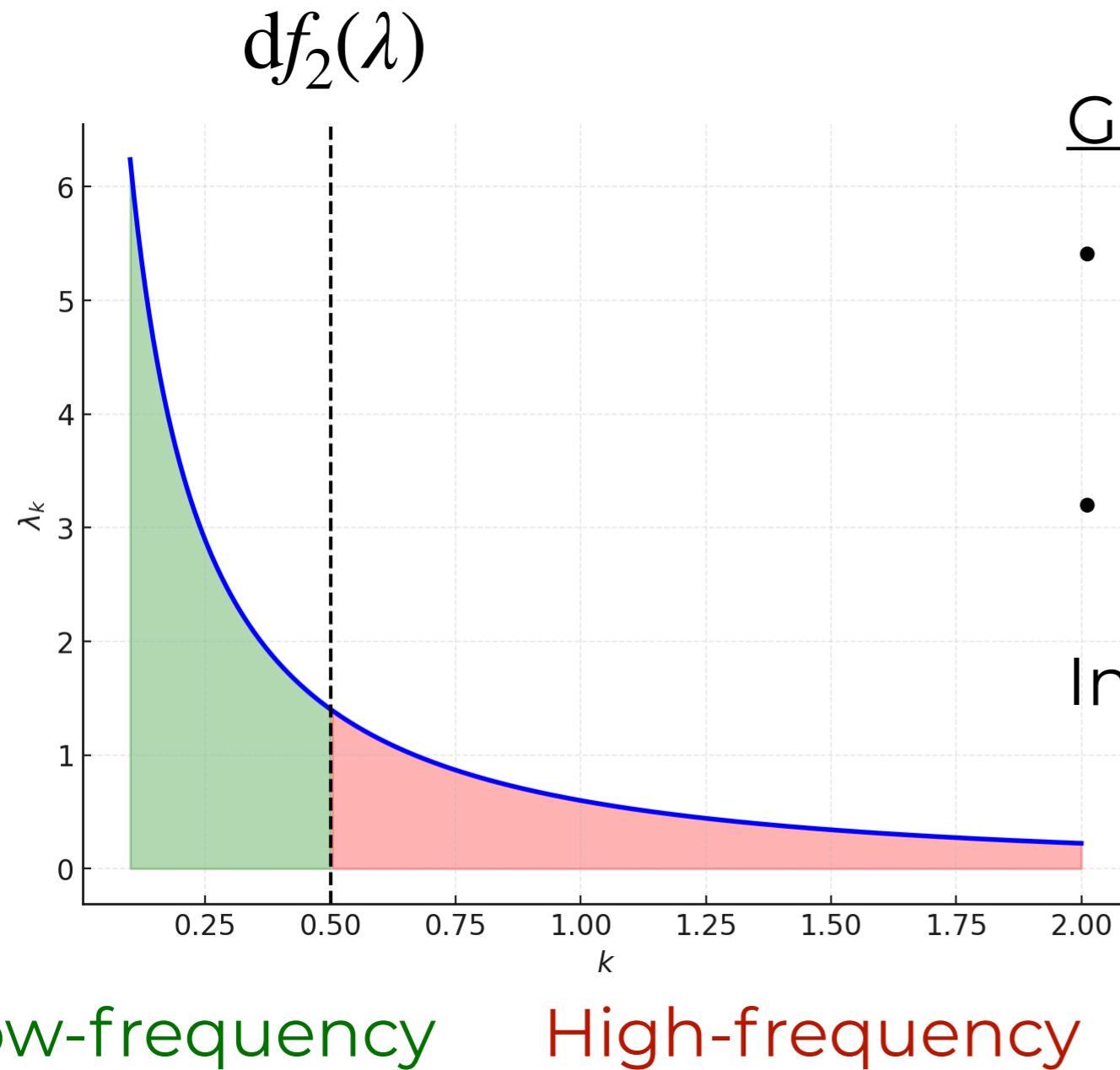
The variance of the ridge risk can be seen as a soft version:

$$df_2(\lambda) = \sum_{k=1}^d \frac{\lambda_k^2}{(\lambda_k + \lambda)^2}$$

- Fast decay: small λ
- Slow decay: large λ



Choosing regularisation



Goal: pick λ such that:

- directions in X that better correlate with θ_* are retained
- Shrink remaining directions

In practice, **cross-validation...**

Best subset selection & the LASSO

Pitfalls of ridge

The ridge estimation performs uniform shrinkage.

$$\hat{\theta}_\lambda(X, y) = \frac{1}{n} \left(\frac{1}{n} X^\top X + \lambda I_d \right)^{-1} X^\top y$$

In other words: ℓ_2 regularisation will control the overall norm $\|\hat{\theta}_\lambda\|_2^2$ by reducing each entry equally

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- Good if θ_\star is a **dense** vector

$$\theta_{\star,j} \neq 0 \quad i = 1, \dots, d$$

$$\theta_\star = \begin{bmatrix} \text{teal} \\ \text{yellow} \\ \text{blue} \\ \vdots \\ \text{red} \end{bmatrix}$$

- Bad if θ_\star is a **sparse** vector

$$\theta_{\star,j} = \begin{cases} 0 & j \in S \subset \{1, \dots, d\} \\ \neq 0 & j \in \{1, \dots, d\} \setminus S \end{cases}$$

$$\theta_\star = \begin{bmatrix} \square \\ \text{yellow} \\ \square \\ \vdots \\ \text{red} \end{bmatrix}$$

Sparsity is everywhere

Many signals of interest admit a sparse representation in a particular basis.

$$f(\mathbf{x}) = \sum_{k \geq 0} f_k \psi_k(\mathbf{x})$$

← basis
← coefficients

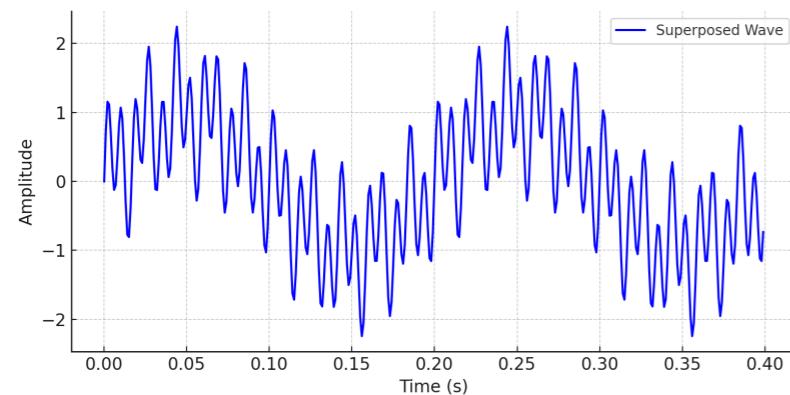
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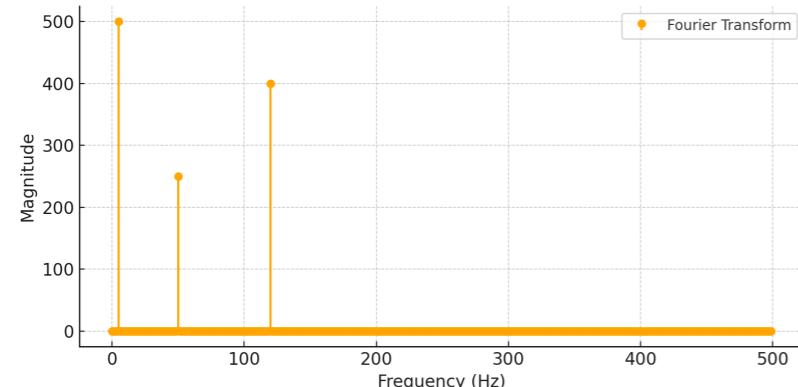
$$f(x) = \sum_{k \geq 0} f_k \psi_k(x)$$

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Example: superposition of sine waves



$$f(t) = \sin(10\pi t) + 0.5 \sin(100\pi t) + 0.8 \sin(240\pi t)$$

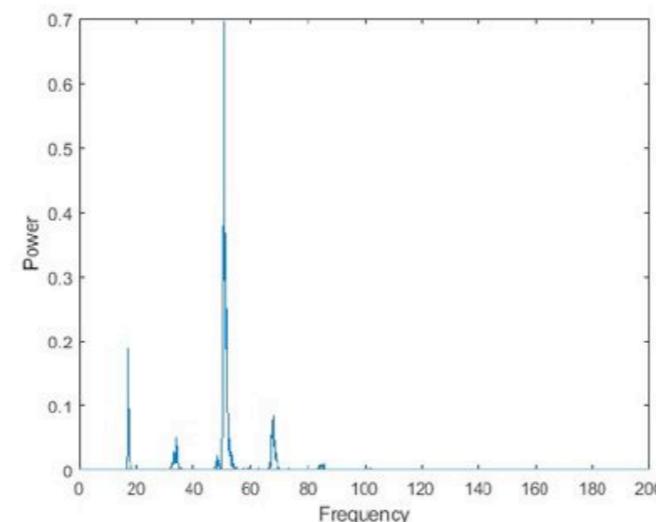
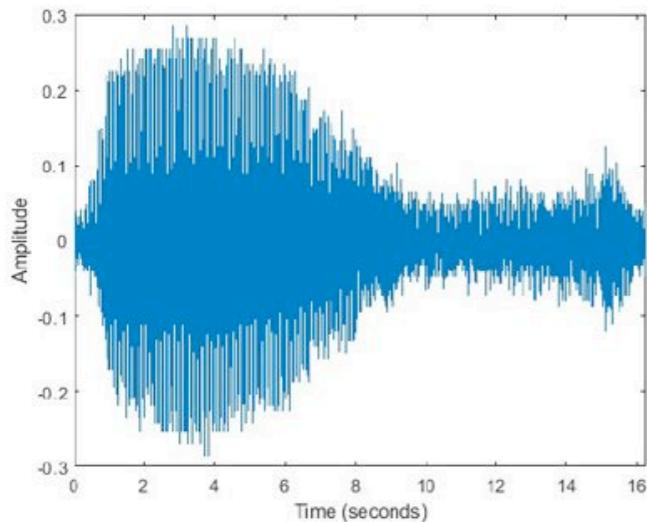


$$\hat{f}(\omega) = \delta_5 + 0.5 \delta_{50} + 0.8 \delta_{120}$$

Sparsity is everywhere

Examples:

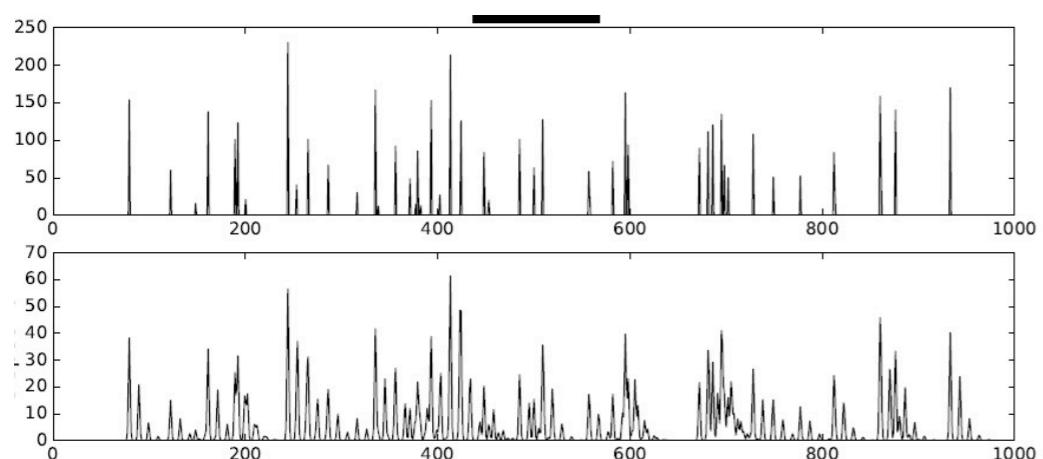
Sound



Images



Scientific signals
(mass spectrography)



And many more...

- Portfolio selection (finance)
- Networks (power grids)
- electroencephalogram
- Etc...

Best subset selection



Idea: encourage solutions which are sparse.

$$\min_{\theta \in \mathbb{R}^d} \frac{1}{2n} \sum_{i=1}^n (y_i - \langle \theta, x_i \rangle)^2 + \lambda \|\theta\|_0$$

where $\|\cdot\|_0 : \mathbb{R}^d \rightarrow \{0, 1, \dots, d\}$ is the ℓ_0 -“norm”:

Strictly not a norm

$$\|\theta\|_0 = \sum_{j=1}^d \mathbb{I}(\theta_j \neq 0) = \# \text{ non-zero entries}$$

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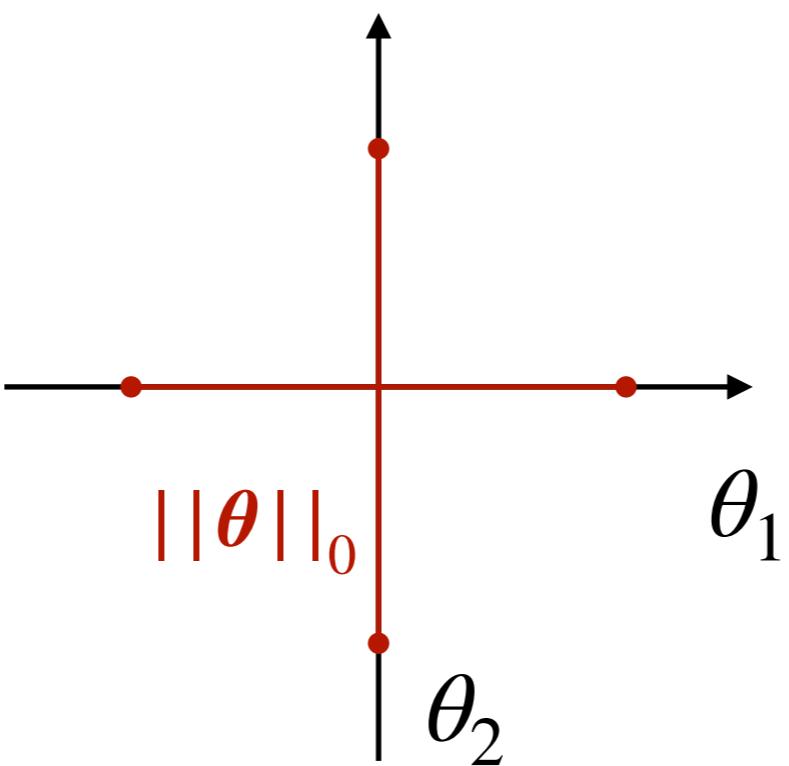
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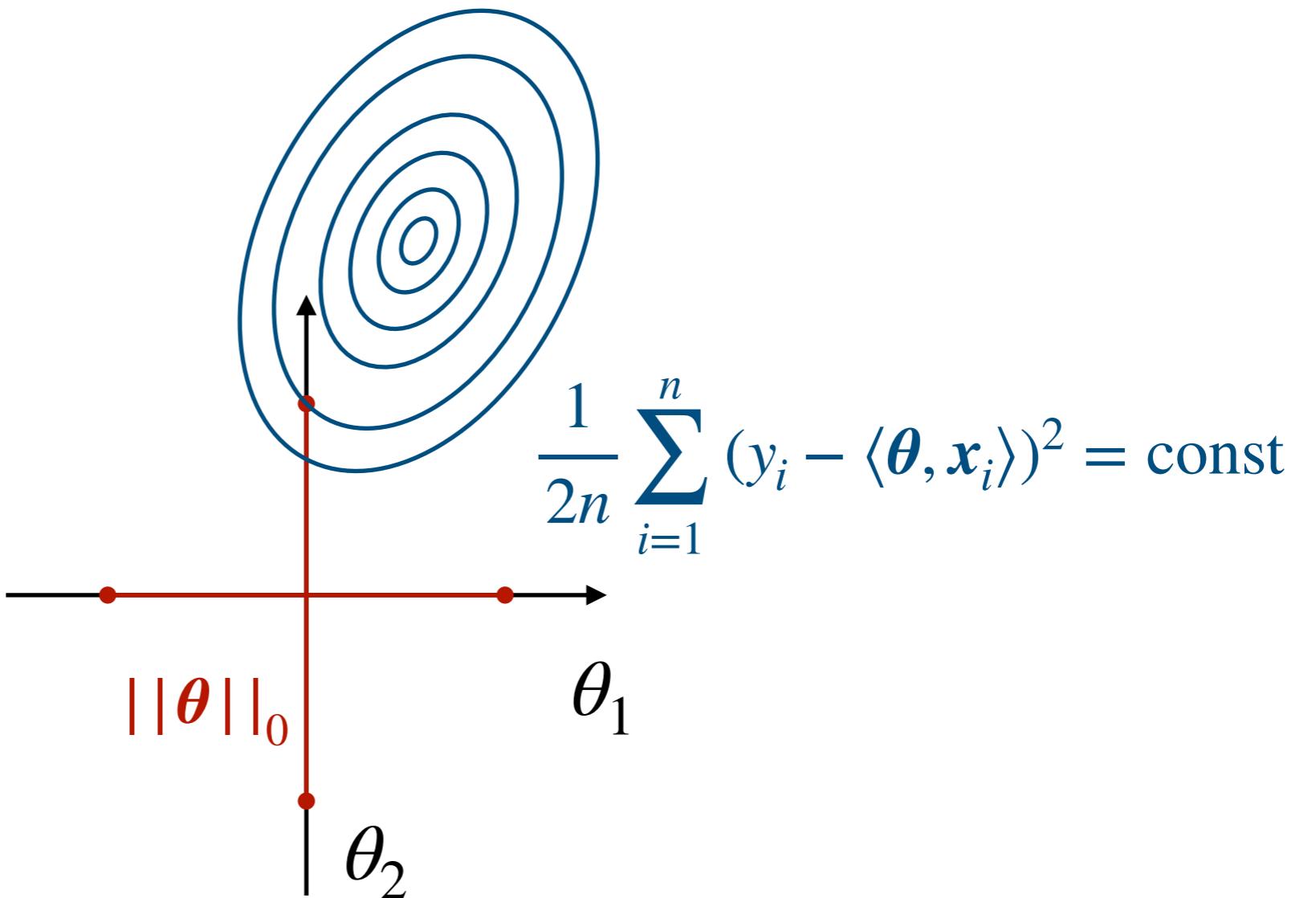
Hence, $\lambda \geq 0$ controls the desired sparsity level

- Large $\lambda \gg 1$: encourage more sparsity
- Small $\lambda \ll 1$: encourage less sparsity

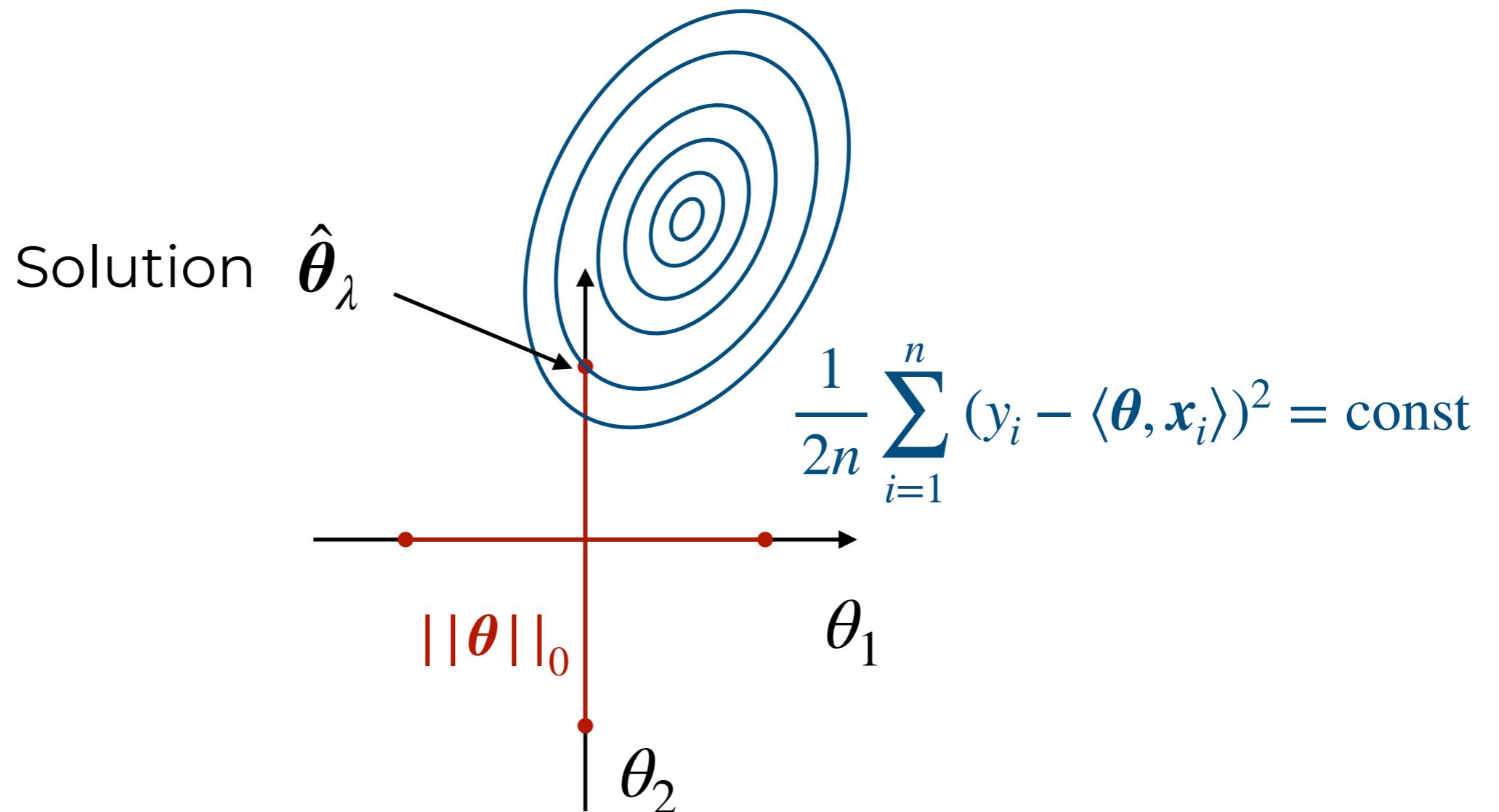
BSS: visualisation



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BSS: orthogonal covariates

To get some intuition about this problem, let's consider a simplified setting: assume the covariates are orthogonal

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$$\|y - X\theta\|_2^2 = \|y\|_2^2 + \theta^\top \textcolor{red}{X^\top X} \theta - 2\theta^\top X^\top y$$

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$$X^\top X = I_d \quad (n \geq d)$$

Therefore, under the above:

$$\min_{\theta \in \mathbb{R}^d} \frac{1}{2n} \sum_{i=1}^n (y_i - \langle \theta, x_i \rangle)^2 + \lambda \|\theta\|_0$$

Is equivalent to:

$$\min_{\theta \in \mathbb{R}^d} \frac{1}{2n} \|z - \theta\|_2^2 + \lambda \|\theta\|_0$$

Which is a simpler problem since it factorises coordinate-wise.

BSS: orthogonal covariates

Coordinate-wise, we need to solve

$$\min_{\theta_j \in \mathbb{R}} L(\theta_j) := \left\{ \frac{1}{2n} (z_j - \theta_j)^2 + \lambda \mathbb{I}(\theta_j \neq 0) \right\}$$

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- In case (a), solution is $\hat{\theta}_{\lambda,j}^{(1)} = 0$
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Hence, the solution is given by:

$$\hat{\theta}_{\lambda,j} = \begin{cases} 0 & \text{if } z_j^2 < 2n\lambda \\ z_j & \text{if } z_j^2 \geq 2n\lambda \end{cases}$$

“Hard threshold”
function

BSS: orthogonal covariates

Putting together, the solution of the BSS problem:

$$\min_{\theta \in \mathbb{R}^d} \frac{1}{2n} \sum_{i=1}^n (y_i - \langle \theta, x_i \rangle)^2 + \lambda \|\theta\|_0$$

Under the assumption of $X^\top X = I_d$ is given by:

$$\hat{\theta}_\lambda = H_{\sqrt{2n\lambda}}(X^\top y)$$

Where:

$$H_\lambda(z) = \begin{cases} 0 & \text{if } |z| < \lambda \\ z & \text{otherwise} \end{cases}$$

