Statistical Learning II

Lecture 5 - Least squares (continued)

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Partiel

Wednesday 22/10/2025

08:30 - 10:00

Least-squares regression

Let $\mathscr{D} = \{(x_i, y_i) \in \mathbb{R}^d \times \mathbb{R} : i = 1,\ldots,n\}$ denote the training data.

Ordinary least-squares (OLS) regression is defined as:

$$\min_{\boldsymbol{\theta} \in \mathbb{R}^d} \hat{\mathscr{R}}_n(\boldsymbol{\theta}) := \frac{1}{2n} \| \boldsymbol{y} - \boldsymbol{X}\boldsymbol{\theta} \|_2^2$$

Where we have defined the data matrix $\boldsymbol{X} \in \mathbb{R}^{n \times d}$ and label vector $\boldsymbol{y} \in \mathbb{R}^n$:

$$\boldsymbol{X} = \begin{bmatrix} - & \boldsymbol{x}_1 & - \\ - & \boldsymbol{x}_2 & - \\ & \vdots & \\ - & \boldsymbol{x}_n & - \end{bmatrix} \in \mathbb{R}^{n \times d} \qquad \boldsymbol{y} = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix}$$

Convexity of OLS

$$\hat{\mathscr{R}}_n(\boldsymbol{\theta}) := \frac{1}{2n} ||\boldsymbol{y} - \boldsymbol{X}\boldsymbol{\theta}||_2^2$$

- Gradient: $$\nabla_{\boldsymbol{\theta}} \hat{\mathscr{R}}_n = -\frac{1}{n}\boldsymbol{X}^\top(\boldsymbol{y} - \boldsymbol{X}\boldsymbol{\theta}) \in \mathbb{R}^d$$
- Hessian: $$\nabla^2_{\boldsymbol{\theta}} \hat{\mathscr{R}}_n = \frac{1}{n}\boldsymbol{X}^\top\boldsymbol{X} \in \mathbb{R}^{d\times d} \quad (:= \hat{\boldsymbol{\Sigma}}_n)$$

Since $\boldsymbol{X}^\top\boldsymbol{X} \succeq 0$, $\hat{\mathscr{R}}_n$ is convex over $\mathbb{R}^d$. This implies that any minimum of $\hat{\mathscr{R}}_n$ is a global minimum.

For $n \geq d$, $\hat{\mathscr{R}}_n$ is strictly convex if and only if $\mathrm{rank}(\boldsymbol{X}^\top\boldsymbol{X}) = d$. This implies that $\hat{\mathscr{R}}_n$ can have at most one global minimum.

Closed-form solution

- Gradient: $$\nabla_{\boldsymbol{\theta}} \hat{\mathscr{R}}_n = -\frac{1}{n} \boldsymbol{X}^\top (\boldsymbol{y} - \boldsymbol{X}\boldsymbol{\theta}) \in \mathbb{R}^d$$

If it exists, a minima must satisfy:

$$\nabla_{\boldsymbol{\theta}} \hat{\mathscr{R}}_n \stackrel{!}{=} 0$$

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$$\nabla_{\boldsymbol{\theta}} \hat{\mathscr{R}}_n \overset{!}{=} 0 \qquad \Leftrightarrow \qquad \boldsymbol{X}^\top \boldsymbol{X}\boldsymbol{\theta} = \boldsymbol{X}^\top \boldsymbol{y}$$

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This is precisely the definition of the pseudo-inverse:

$$\boxed{\hat{\boldsymbol{\theta}}_{OLS} = \boldsymbol{X}^{+}\boldsymbol{y}}$$

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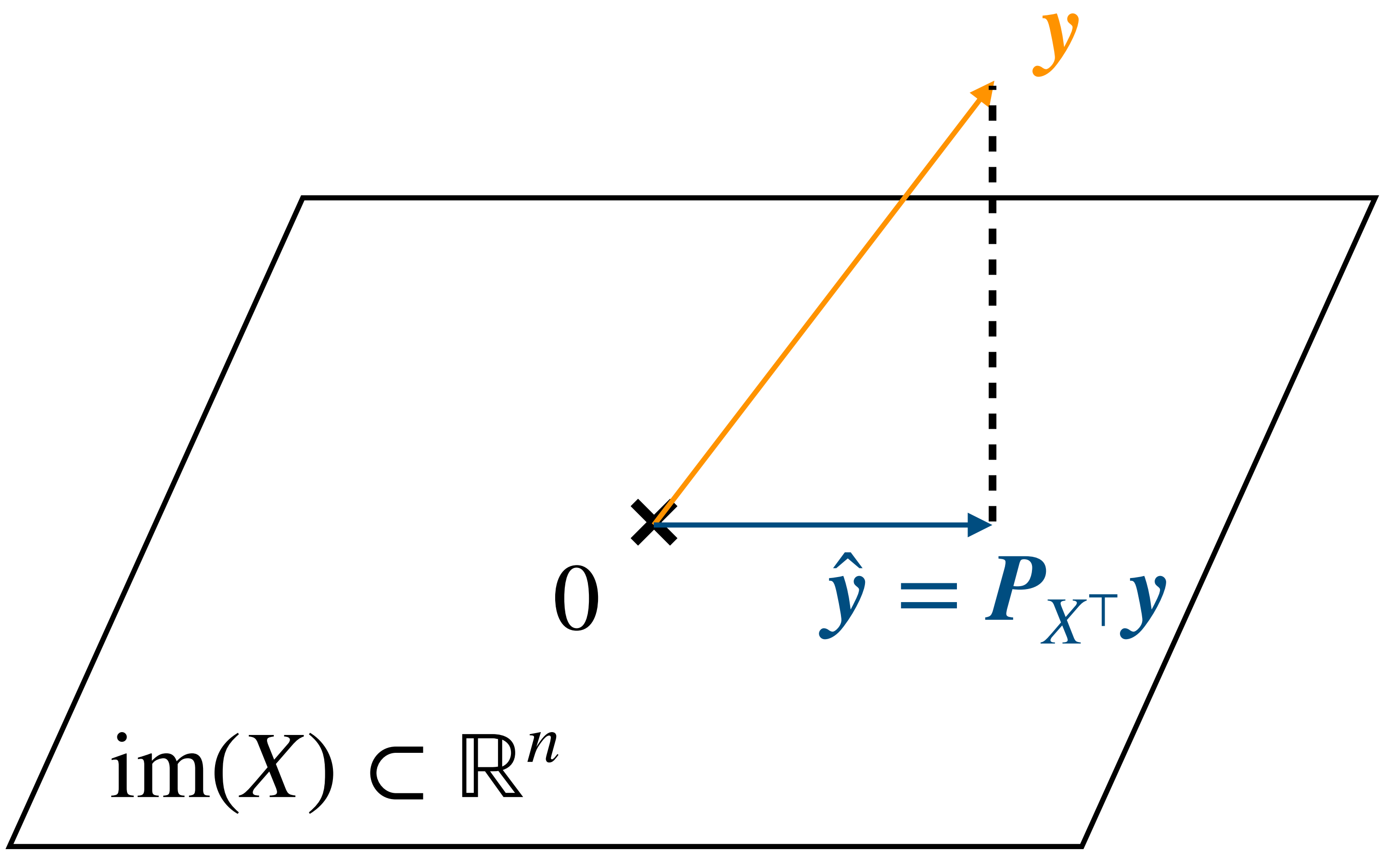
If $\mathrm{rank}(X) = \min(n, d)$: $$\hat{\boldsymbol{\theta}}_{OLS} = \begin{cases} (\boldsymbol{X}^\top \boldsymbol{X})^{-1} \boldsymbol{X}^\top \boldsymbol{y} & \text{if } n \geq d \\ \boldsymbol{X}^\top (\boldsymbol{X}\boldsymbol{X}^\top)^{-1} \boldsymbol{y} & \text{if } n < d \end{cases}$$

Geometrical interpretation

This gives a natural interpretation of the OLS predictor as an orthogonal projection of the labels in the row space of $\boldsymbol{X}$:

$$\hat{\boldsymbol{\theta}}_{OLS} = \boldsymbol{X}^{+}\boldsymbol{y} \quad \Rightarrow \quad \hat{\boldsymbol{y}}_{OLS} = \boldsymbol{X}\hat{\boldsymbol{\theta}}_{OLS} = \boldsymbol{X}\boldsymbol{X}^{+}\boldsymbol{y}$$

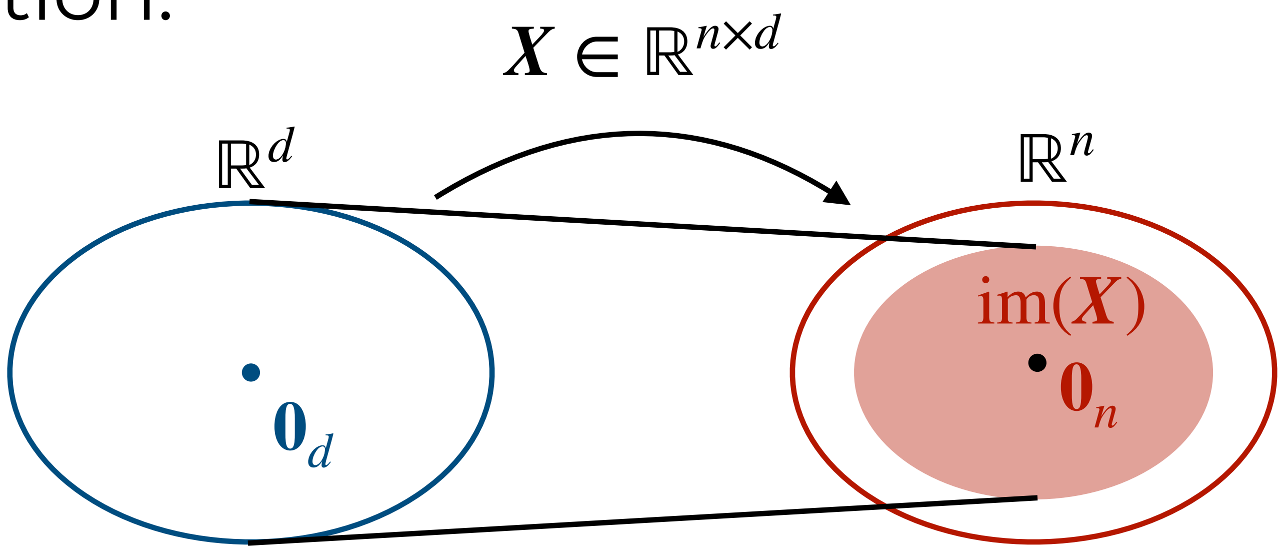


$$\min_{\boldsymbol{z} \in \mathrm{im}(\boldsymbol{X})} ||\boldsymbol{y} - \boldsymbol{z}||_2^2$$

Two scenarios

From now on, let's assume $\boldsymbol{X}$ is full-rank.

- For $n \geq d$: more equations than variables. $\boldsymbol{X\theta} = \boldsymbol{y}$ admits an unique solution.



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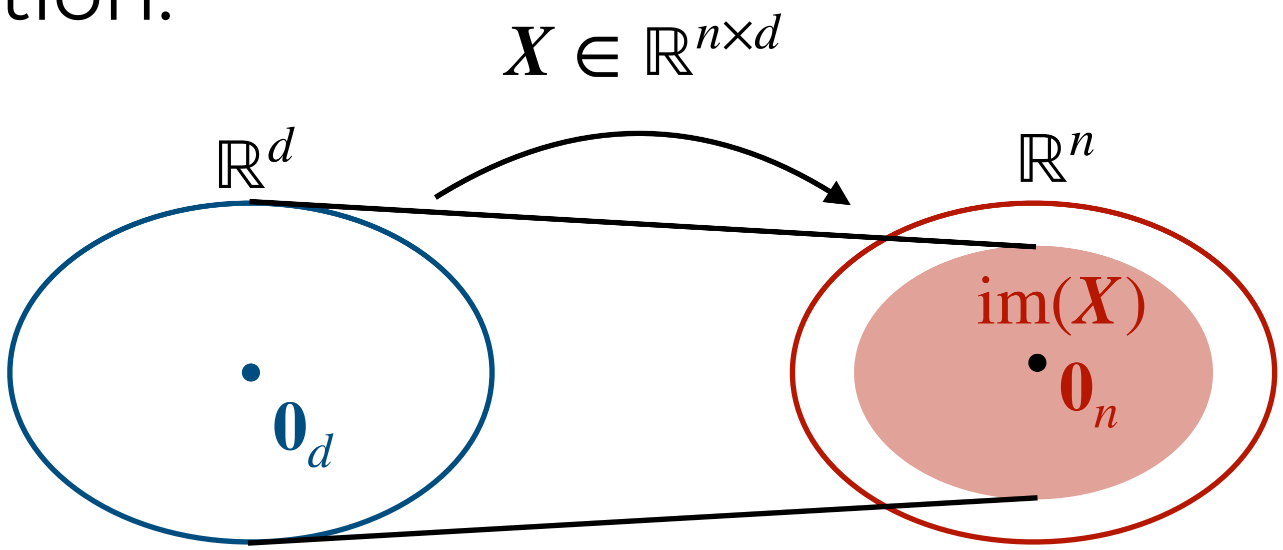
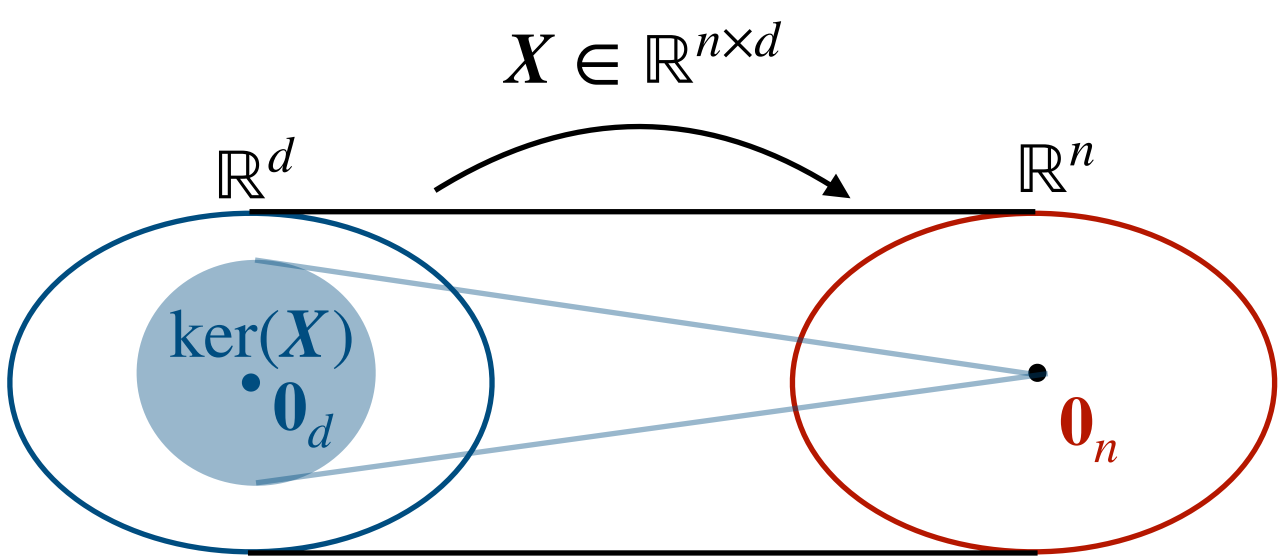


- For $n < d$: more variables than equations. $\boldsymbol{X\theta} = \boldsymbol{y}$ admits several solutions.



OLS as least norm solution

Assume $\mathrm{rank}(\boldsymbol{X}) = n < d$. Then, OLS admits the following interpretation as the minimum ℓ_2-norm solution:

$$\hat{\boldsymbol{\theta}}_{OLS} = \underset{\boldsymbol{\theta} \in \mathbb{R}^d}{\mathrm{argmin}} \; ||\boldsymbol{\theta}||_2$$

$$\text{subject to} \quad \boldsymbol{X\theta} = \boldsymbol{y}$$

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Then: $\langle \hat{\boldsymbol{\theta}} - \hat{\boldsymbol{\theta}}_{OLS}, \hat{\boldsymbol{\theta}}_{OLS} \rangle = \langle \hat{\boldsymbol{\theta}} - \hat{\boldsymbol{\theta}}_{OLS}, \boldsymbol{X}^\top (\boldsymbol{X}\boldsymbol{X}^\top)^{-1} \boldsymbol{y} \rangle$

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$$= \langle \boldsymbol{X}(\hat{\boldsymbol{\theta}} - \hat{\boldsymbol{\theta}}_{OLS}), (\boldsymbol{X}\boldsymbol{X}^\top)^{-1} \boldsymbol{y} \rangle$$

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Therefore $\hat{\boldsymbol{\theta}} - \hat{\boldsymbol{\theta}}_{OLS} \perp \hat{\boldsymbol{\theta}}_{OLS}$.

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Therefore $\hat{\boldsymbol{\theta}} - \hat{\boldsymbol{\theta}}_{OLS} \perp \hat{\boldsymbol{\theta}}_{OLS}$. Hence:

$$||\hat{\boldsymbol{\theta}}||_2^2 = ||\hat{\boldsymbol{\theta}} - \hat{\boldsymbol{\theta}}_{OLS} + \hat{\boldsymbol{\theta}}_{OLS}||_2^2$$

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<u>*Proof:*</u> Let $\hat{\boldsymbol{\theta}} \in \mathbb{R}^d$ denote a different solution from $\hat{\boldsymbol{\theta}}_{OLS}$.

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