

# Speedier Simulations with Quasi-Monte Carlo Methods

## Final Presentation

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## 1 Introduction

## 2 Theoretical Research

- Analysis of Density Estimation Accuracy
- Density Estimation With Conditioning

## 3 Applications

- Combinatorial Objects
- Finance Options

## 4 Conclusion and Final Thoughts



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# Purpose of Quasi-Monte Carlo

Many problems can be written in the form of a multidimensional integral taken over the  $d$ -dimensional unit cube, which can be approximated by point evaluations:

$$\int_{[0,1]^d} f(\mathbf{x}) d\mathbf{x} \approx \frac{1}{n} \sum_{i=1}^n f(\mathbf{x}_i) \quad (1)$$

**Quasi-Monte Carlo** (QMC) methods use point sets

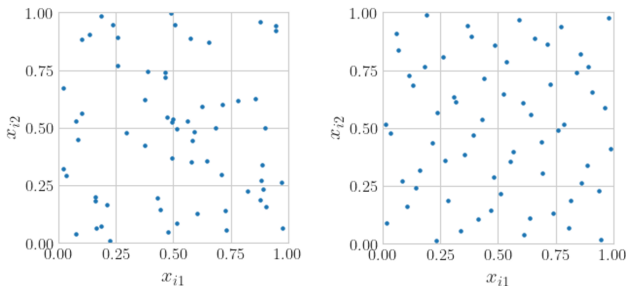
$P_n = \{\mathbf{x}_i\}_{i=1}^n \subset [0, 1]^d$  with **low discrepancy**, while traditional Monte Carlo methods use IID (independent and identically distributed) points.



# Low Discrepancy

**Discrepancy** = a measure of the difference between the empirical distribution of a set of points and the uniform distribution [5].

- In other words, how "evenly" a set of points fills the domain.
- Lower discrepancy = faster convergence and lower-variance estimates
- **Sobol'** points = a type of LD sequence



**Figure 1:** Left: IID, vs. right: low discrepancy (Sobol')



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# Kernel Density Estimation

## Definition

**Kernel Density Estimation** is a non-parametric method for estimating probability distributions using kernels as weights

$$\hat{\rho}_h(y) = \frac{1}{n} \sum_{i=1}^n k_h(y - f(\mathbf{x}_i)) = \frac{1}{n} \sum_{i=1}^n \frac{1}{h} k\left(\frac{y - f(\mathbf{x}_i)}{h}\right)$$

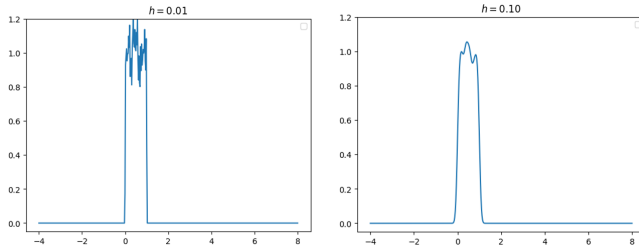
KDE's rely on the **placement** and **distribution** of the data points. Thus, using low discrepancy (LD) sequences allows for density estimates with **reduced sampling bias** and **variance**, especially in higher dimensions [5].



# Bandwidth Parameter

**Bandwidth**, denoted  $h$ , determines the amount of smoothing applied when estimating  $\hat{\rho}(y)$  [6].

- Smaller  $h$  = less smoothing; larger  $h$  more smoothing.



**Figure 2:** KDE for standard uniform distribution with smaller bandwidth (left) vs. larger bandwidth (right).





# Objective of Error Analysis

- Study the effects of various factors on the accuracy of kernel density estimators.
- Factors/parameters include:
  - IID vs. Sobol' (LD) points
  - $K$  = kernel function
  - $n$  = sample size
  - $h$  = bandwidth
  - $d$  = dimension
  - $f(X)$  = function that transforms a set of IID or LD points  $X$  into a random variable with a known distribution



# Parameters and Other Factors

- **Kernels:** Epanechnikov and Gaussian [5]
- **Bandwidths:** 100 different values in the range  $(0, 1]$ , defined on a power scale:  $\{h_i\} = 2^{\{a_i\}}$  where  $\{a_i\}$  is the set of 100 evenly-spaced points between  $-10$  and  $0$ , inclusive.
- **Sample sizes:**  $\{2^{10}, 2^{11}, 2^{12}, 2^{13}, 2^{14}\}$
- **Distributions:**
  - **1-dimensional:** standard Uniform, standard Gaussian, exponential (with  $\lambda = 1$ ), Laplace (with  $\mu = 0$  and  $b = 1$ ), and chi-squared (with  $df = 1$ ).
  - **Multi-dimensional:** triangular (sum of 2 standard uniforms), sum of 3 and 5 standard Gaussians, and chi-squared with  $df = 3$  and  $df = 5$  (i.e., sum of 3 and 5 squared Gaussians).



# Definition of RMSE

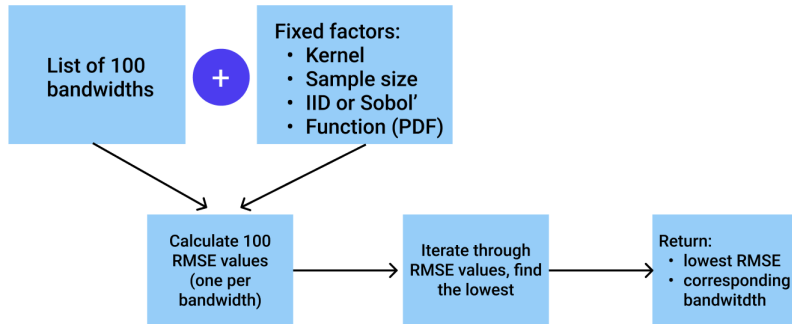
## Definition

The **root mean square error (RMSE)** measures the average difference between the expected values and predicted values of a probability density function:

$$\text{RMSE} = \sqrt{\frac{1}{n} \sum_{i=1}^n (\hat{\rho}(y_i) - \rho(y_i))^2} \quad (2)$$



# Finding Minimum RMSE and Optimal Bandwidth



**Figure 3:** Flowchart to find optimal bandwidth using RMSE calculations. This procedure is performed for each set of factors and parameters that we test.



# Results: RMSE vs. $n$

- We plotted four regression models of the form

$$-\log_{10}(RMSE) = b_0 \cdot \log_{10}(n) + \sum_{j=1}^N b_j \cdot x_j + \varepsilon_j, \quad (3)$$

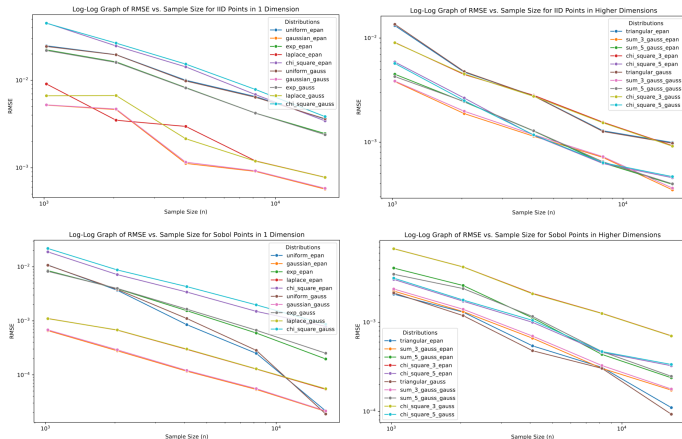
where  $b_0$  is the slope of  $-\log_{10}(RMSE)$  vs.  $\log_{10}(n)$ , and  $x_j$  are indicator variables to separate the  $N$  different distributions we tested.

**Table 1:** Coefficients  $b_0$  for linear regression models of log-scaled RMSE vs. log-scaled sample size.

	1-dim	Multi-dim
IID	0.8379	0.8811
Sobol'	1.0750	0.9465



# Results: RMSE vs. $n$ (con't)



**Figure 4:** Log-log graphs of RMSE vs. sample size ( $n$ ) based on Equation (3). For IID points, slope in 1 dim. is 0.8379 (top left) and in higher dims. is 0.8811 (top right). For Sobol' points, slope in 1 dim. is 1.0750 (bottom left) and in higher dims. is 0.9465 (bottom right).



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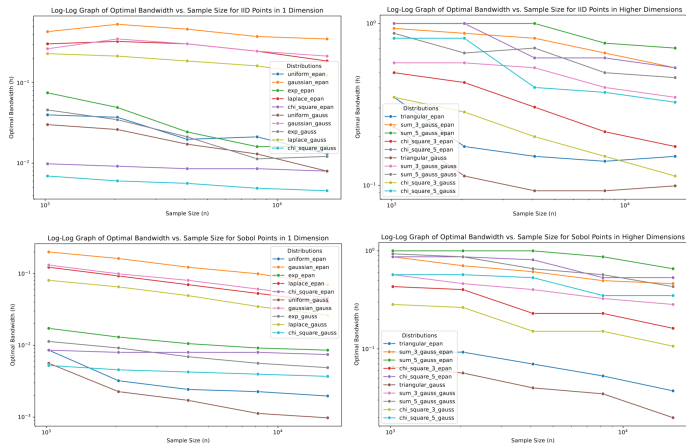
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**Table 2:** Coefficients  $b_0$  for linear regression models of log-scaled optimal bandwidth vs. log-scaled sample size.

	1-dim	Multi-dim
<b>IID</b>	0.2889	0.2747
<b>Sobol'</b>	0.3374	0.2818



# Results: Optimal Bandwidth vs. $n$ (con't)

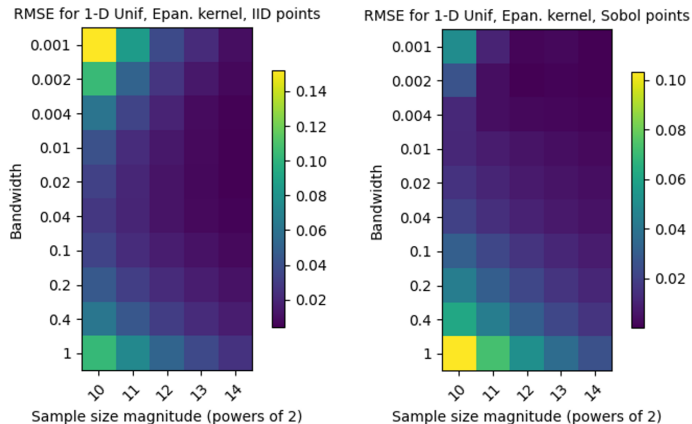


**Figure 5:** Log-log graphs of optimal  $h$  vs. sample size ( $n$ ) based on Eq. (4). For IID points, slope in 1 dim. is 0.2889 (top left) and in higher dims. is 0.2747 (top right). For Sobol' points, slope in 1 dim. is 0.3374 (bottom left) and in higher dims. is 0.2818 (bottom right).





# Results: Optimal Bandwidth vs. n (con't)



**Figure 6:** RMSE values (shaded) for different sample sizes and bandwidths when estimating the 1-dimensional standard uniform distribution using the Epanechnikov kernel.



# Thoughts and Future Work

- Advantages of using Sobol' (LD) points vs. IID points
- Tailoring bandwidth specifically based on distribution & sample size
- *Future work*: testing different distributions, higher dimensions, different sets of bandwidths; theoretical work to illustrate relationships



# Conditional Density Estimation

Let  $Y = f(X_1, \dots, X_d)$  such as

$$Y = X_1^2(X_2 + 3X_3)$$

If all but one of the variables were fixed, then this would be simple.



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Let  $f(X_1, \dots, X_d | \mathcal{G}_{-k}) =: g(X_k; \mathcal{G}_{-k})$

$$F(y | \mathcal{G}_{-k}) = P(Y \leq y | \mathcal{G}_{-k}) = P(X_k \leq g^{-1}(y) | \mathcal{G}_{-k}) = F_k(g^{-1}(y))$$

This gives a conditional density  $\rho(y | \mathcal{G}_{-k}) = \frac{d}{dy} F_k(g^{-1}(y))$



## Theorem

Suppose we have  $\rho(y|\mathcal{G}_{-k})$  that satisfies certain technical assumptions.  
Then

$$\mathbb{E}[\rho(y | \mathcal{G}_{-k})]_{\mathcal{G}_{-k}} = \rho(y) \quad (5)$$

with

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Under slight modification, we can use QMC to help approximate this expectation. The dimension is one less than the dimension of a standard KDE since we only need to realize  $d - 1$  variables!



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If we have the true density  $\rho$ , then we can approximate the 1-dimensional integral with  $n_e$  evenly spaced points:  $\widehat{\text{MISE}} = \frac{b-a}{n_e} \sum_{i=1}^{n_e} (\rho(y_i) - \hat{\rho}(y_i))^2$



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If we don't have the true density, we can generate the conditional density estimator  $n_r$  times and use the unbiased estimator [2]

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$$\widehat{MISE} = \frac{b-a}{n_e} \sum_{i=1}^{n_e} \text{var}[\hat{\rho}_n(y_i)]_{n_r}$$

Thus, as we vary  $n$ , we can get  $\widehat{MISE}(n)$  and fit it to a power law

$$\widehat{MISE}(n) = Kn^{-\nu}$$



# Example 1: Weighted Sum Uniform

$Y = X_1 + X_2 + X_3$  for  $X_j \sim \mathcal{U}([0, 2^{j-1}])$

$$\rho(y \mid \mathcal{G}_{-k}) = \begin{cases} 2^{1-k} & y - \sum_{i \neq k} X_i \in [0, 2^{k-1}] \\ 0 & \text{else} \end{cases}$$

The true PDF is piecewise constant, linear, and quadratic.

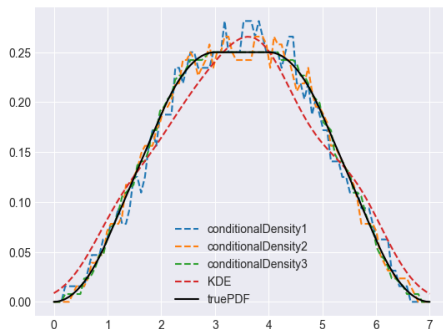


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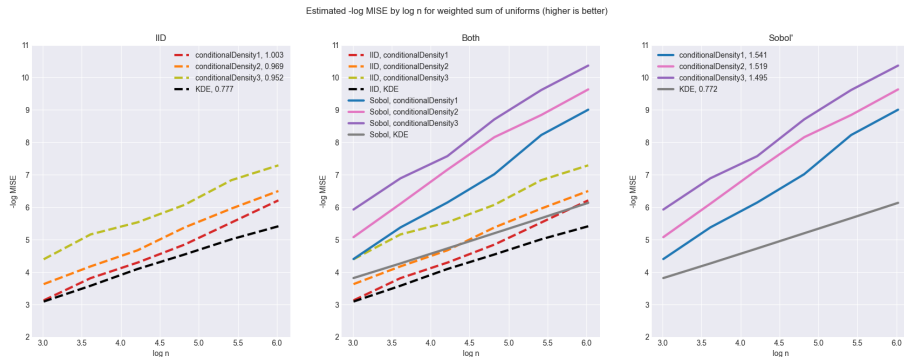
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**Figure 7:** Plotting the conditional densities with Sobol' for  $n = 32$  with a KDE and true PDF



# Error graph



**Figure 8:**  $-\log$ ,  $\log$  plot. The slopes of IID hover around 1, KDE around 0.8, and Sobol' around 1.5. Best conditional density is conditionalDensity3.



## Example 2: Cantilevered Beam

$Y = \frac{\kappa}{X_1} \sqrt{\frac{X_2^2}{256} + \frac{X_3^2}{16}}$  for  $X \sim \mathcal{N}(\mu_j, \sigma_j^2)$  with

$\kappa = 5 \cdot 10^5$ ,  $\mu = (2.9 \cdot 10^7, 500, 1000)$  and  $\sigma = (1.45 \cdot 10^6, 100, 100)$

No true PDF, and the conditional densities have complicated closed-form expressions

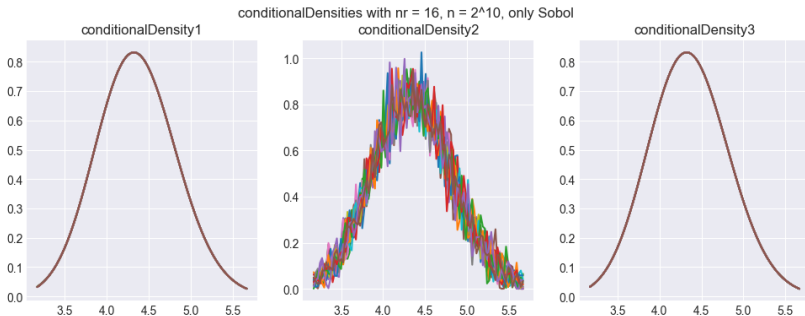


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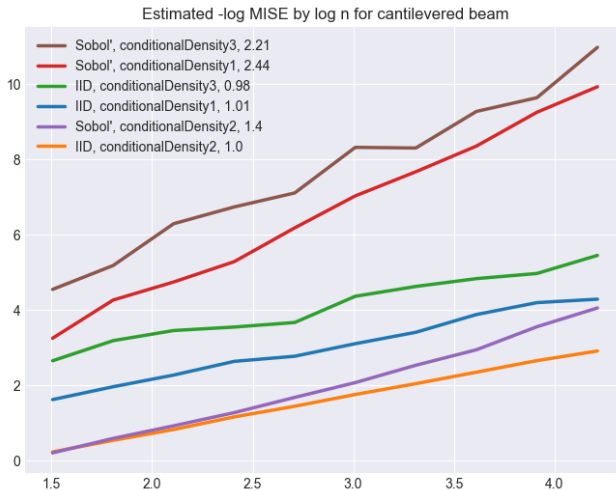


**Figure 9:** Plotting the conditional densities with Sobol' for  $n = 1024$  over  $nr = 16$  repetitions





# Error graph



**Figure 10:** -log, log plot. Slope of IID hover around 1. conditionalDensities vary a lot. Sobol' better than IID



- Conditional density estimation, when applicable, is much less cumbersome and much better than KDE
- Most of these are toy examples; certain real examples (like options pricing) require much more care and attention
- The effect is most noticeable when certain variables contribute to a large portion of the variance
- Interesting work will continue in applying conditional densities estimates to different classes of problems



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- The **Kernel Trick** is a powerful technique for modeling non-linear relationships using linear learning algorithms by implicitly mapping data into a higher-dimensional feature space using a kernel function.
- Why are Kernels Defined on Discrete input Spaces Useful?
  - Various real-world applications: bio-informatics, community detection, manifold learning for deep generative modeling, etc... [3]



# Applying Kernels to Combinatorial Objects

- **Kernel Function:** For discrete objects, like graphs, we define a *graph kernel* (or kernel function) as  $K : \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{R}$  which returns a real number representing the similarity measure between these structures



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- **Gram (Kernel) Matrix:** Kernel evaluations are stored in a Gram Matrix  $\mathbf{K} \in \mathbb{R}^{N \times N}$  whose entries enumerate the pairwise kernel evaluations  $\mathbf{K} := [K(x_i, x_j)]_{i,j=1}^N$ 
  - Entries of this matrix represent similarities between each pair of data points



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    - **Randomised Function:**  $\phi : \mathbb{R}^d \rightarrow \mathbb{R}^s$  maps the original data points into a lower dimensional space
    - **Low-Dimensional Feature Vector:** whose dot product approximates the kernel function:  $\mathbf{K}_{ij} = \mathbb{E}[\phi(x_i)^\top \phi(x_j)]$  meaning the dot product of the transformed vectors is an unbiased estimator of the kernel function



# Graph Random Features

- A recently viable **Graph Random Features** (GRFs) mechanism employs an ensemble of random walkers that deposit a "load" at every vertex they pass, depending on
  - i) the product of weights of edges traversed by the walker, and
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- **Key Limitation:** Only addresses a niche family of graph kernels





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- By parameterizing this modulation function on a **Neural Network**, the algorithm provides an efficient, unbiased approximation of the desired graph kernel



# Quasi-Monte Carlo Graph Random Features

Can we improve the accuracy and efficiency of g-GRFs?

- **Quasi-Monte Carlo Graph Random Features** (q-GRFs) introduces correlated ensembles, or antithetic walkers, to correlate the lengths of random walks with some stopping criterion [4]



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  - In the IID implementation (g-GRFs) two termination random variables (TRVs) are sampled independently from a uniform distribution  $t_{1,2} \sim \text{Unif}(0, 1)$ . Each walker terminates if their respective TRV  $t_{1,2} < p$
  - **Antithetic Walkers:** A pair of walkers are antithetic if their TRVs are marginally distributed as  $t_{1,2} \sim \mathcal{U}(0, 1)$  but are offset by  $\frac{1}{2}$ ,

$$t_2 = \text{mod}_1 \left( t_1 + \frac{1}{2} \right), \quad (6)$$



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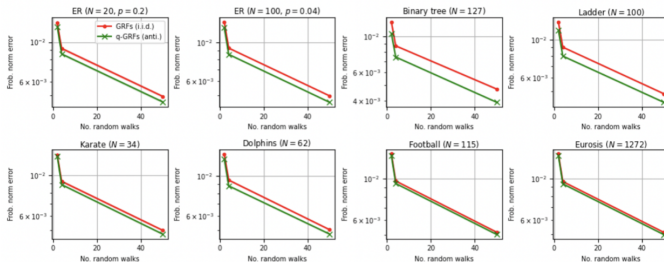
- Diversifying the lengths of random walks prevents clustering; therefore, Choromanski, Reid, and Weller assert that antithetic walkers (q-GRFs) yield lower variance estimators of the 2-regularised Laplacian kernel



- **Motivating Question:** Can we identify an alternative kernel function or graph type where q-GRFs achieve a lower variance estimator like the previously introduced 2-Regularised Laplacian kernel?
  - Diffusion (or Heat), Matérn, Inverse Cosine Kernels
  - Erdős-Rényi and Barabási-Albert random graph models, Binary Trees, and Ladder graphs



# Diffusion (or Heat) Kernel

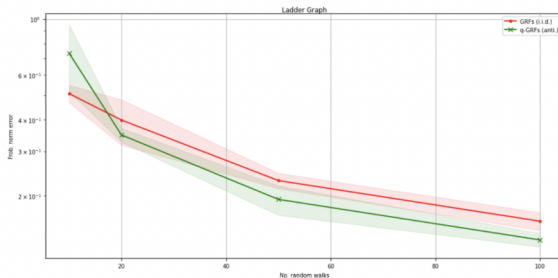
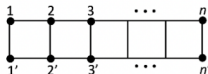


**Figure 11:** Relative Frobenius norm error of estimators of the diffusion (or heat) kernel with  $t = 0.5$  using g-GRFs (red circles) and q-GRFs (green crosses). **Lower is better.**  $N$  gives the number of nodes and  $p$  is the edge-generation probability for the Erdős-Rényi graphs. One standard deviation is shaded.





# Ladder Graphs



**Figure 12:** LEFT: Standard ladder graph. RIGHT: Ladder graph with number of random walks set to (10, 20, 50, 100) and number of rungs set to 9. Q-GRFs yield lower variance estimators of the diffusion kernel than general g-GRFs.



# Results and Future Work

- Despite the inherent randomness in the antithetic procedure, there is strong evidence to suggest that q-GRFs have the potential to yield lower variance estimators of the diffusion kernel – though further experimentation and theoretical work is necessary to make final conjectures
- We assert that the **number of rungs** on a ladder graph impacts q-GRFs ability to achieve a lower variance estimator of the diffusion kernel
- Further theoretical results investigating these phenomena are forthcoming

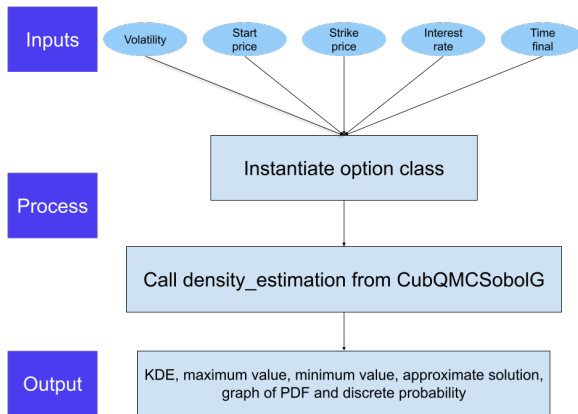


# QMC Application for Finance Options

- Use QMC to estimate the PDF of the payoff of an option.
- QMC allows faster computation than regular MC.
- QMCPy allows someone to compute such estimations.



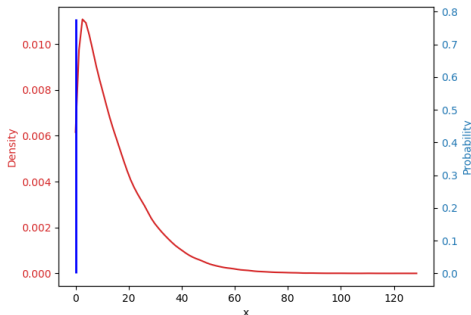
# QMC Application for Finance Options



**Figure 13:** Flowchart of using an option class.



# European Density Estimation



**Figure 14:** PDF of the payoff for a European option with starting price of \$100 and strike price of \$120, with interest rate of %0.5, and volatility of 0.2.



- Finish code that was in development by a graduate student.
- Review code before merging into the main development branch.
- Satisfy requested changes to the code before finally merging.
- Begin merging said code into the main development.



- **Issue:** Code among the option classes have the same code.
- **Solution:** Refactored all option classes to remove duplicate code.



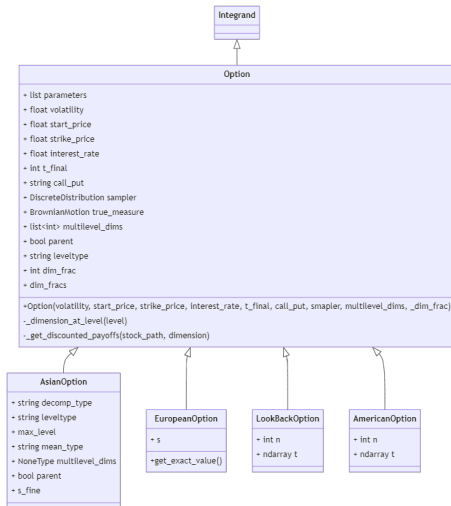
What does refactoring mean?

## Definition

**Refactor:** a change made to the internal structure of software to make it easier to understand and cheaper to modify without changing its observable behavior.[1]







**Figure 15:** UML Class diagram of all option classes.



# Conclusion

Important things learned and encountered when developing robust scientific software:

- Use software engineering principles.
- Structure your code and write everything to be clear.
- Use the documentation of the language and libraries to help build the software.



# Overview

## 1 Introduction

## 2 Theoretical Research

- Analysis of Density Estimation Accuracy
- Density Estimation With Conditioning

## 3 Applications

- Combinatorial Objects
- Finance Options

## 4 Conclusion and Final Thoughts



# Conclusion and Final Thoughts

- Low discrepancy points can provide much lower error than IID uniform points in a variety of situations:
  - Density estimations (using kernels & conditioning)
  - Analyzing combinatoral objects
  - Modeling financial options
- Theoretical analysis, further applications, and robust implementation are key to maximizing the potential of low discrepancy points



# Thank you!



# References I



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# Appendix A: Links

- KDE Error Analysis: Final report and Python notebook
- <https://qmcpy.org/>
- <https://github.com/QMCSoftware/QMCSoftware>
- Conditional QMC: Report and Python notebooks 1 and 2
- QMC with combinatorial objects: Final report and Python notebook





# Appendix B: Epanechnikov vs. Gaussian Kernel

## Definition

The **Gaussian** kernel is defined with the weighting function

$$w(y, h) = \frac{1}{2\pi h} e^{-\frac{y^2}{2h^2}}, \quad -\infty < y < \infty \quad (7)$$

## Definition

The **Epanechnikov** kernel is defined with the weighting function

$$w(y, h) = \begin{cases} \frac{3}{4h} (1 - \frac{y^2}{5h^2}) / \sqrt{5} & \text{for } |y| < \sqrt{h}, \\ 0 & \text{otherwise.} \end{cases} \quad (8)$$



- The above weighting functions,  $w(y, h)$ , are all of the form  $w(y, h) = \frac{1}{h} K(\frac{y}{h})$ , where  $K$  is the *kernel* function [6].



## Appendix B: Epanechnikov vs. Gaussian Kernel (con't)

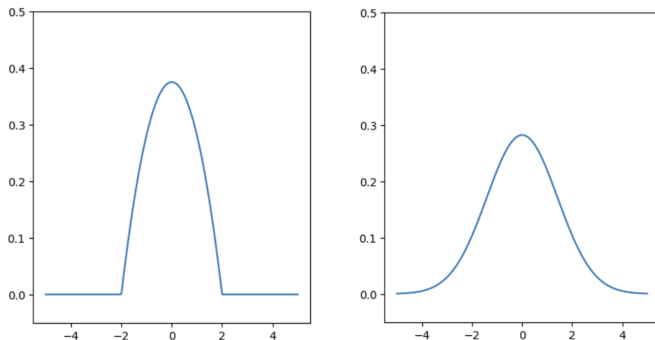
The Epanechnikov kernel has an efficiency of 100%, while the Gaussian kernel has an efficiency of 95.12%, where the *efficiency* of a kernel is defined as:

$$\text{Efficiency}(K) = \left( \frac{RMSE_{opt}(\hat{\rho}) \text{ using } K_{EP}}{RMSE_{opt}(\hat{\rho}) \text{ using } K} \right)^{\frac{5}{4}} \quad (9)$$

This means that the RMSE obtained using an Epanechnikov kernel with  $n \approx 95$  is approximately equal to the RMSE obtained using a Gaussian kernel with  $n = 100$  [5].



## Appendix B: Epanechnikov vs. Gaussian Kernel (con't)



**Figure 16:** Shape of the Epanechnikov (left) vs. the Gaussian kernel (right). The kernel determines the shape of the weighting function.

