Using Random Forest with Proportional Subsetting for Time Series Forecasting

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Abstract

This project explores an alternative approach to time series forecasting by leveraging the Random Forest machine learning algorithm, specifically examining its performance when trained on varying proportions of lagged data. By comparing its predictive accuracy against established methodologies such as Box-Jenkins and Auto ARIMA, we find that Random Forest models trained on smaller lagged proportions, particularly 5% and 10%, exhibit superior performance. These findings establish the importance of challenging traditional assumptions about time series modeling and also highlight the evolving landscape of predictive analytics, where machine learning techniques continue to refine and enhance forecasting accuracy.

1. Introduction

As the ever-changing landscape of the technology market in one of the most capitally-affluent nations begins to grow at rapid speeds beyond comprehension, the world watches and admires in how previous barriers to growth are almost immediately traversed with the ingenuity of emerging topics within machine learning, AI, and big data processes. However, among the many existing challenges within the market, the inability to accurately forecast and predict trends across multiple different latitudes have left high-level researchers and quantitative developers with numerous unanswered questions and temporary solutions. Within this intersection of statistical modeling and machine learning, time series modeling allows users to model a time-correlated process and explore the impact of one of the most elusive, impactful predictors of any model, time itself. For the majority of the discipline's existence, stochastic models have dominated the methodology. Primarily, these include ARIMA and SARIMA models, which model the influence of previous time points on current data. While robust to many troublesome patterns found within the average time series, their performative ability begins to decline with greater data complexity. Furthermore, parametric models, while accurate and reliable in their ideal conditions, always suffer from violated assumptions in modern-day applications. Among these violations is the lack of homoskedasticity, which plagued models with inconsistent variance until Robert F. Engle's ARCH model [1] was born to combat its detrimental influence by building upon autoregressive models by introducing squared terms.

Like Engle, other researchers strive to address these challenges through innovative approaches. As advancements in predictive modeling continue to evolve, we turn our attention to the 2017 study by Hristos Tyralis and Georgia Papacharalampous [2], which demonstrates the superior predictive performance of Random Forest models trained on proportions of lagged data.

In order to advance our understanding of this topic, we will be expanding upon this study, utilizing the methodology of Tyralis and Papacharalampous to inspire our investigation into the effectiveness of Random Forest-engineered time series models in comparison to other contemporary methods. In addition to performing comparative measures of accuracy, we will be focusing our attention towards the unique attributes of the examined BCG Jena Weather Station Dataset from 2017-2024 [3], confronting the potential drawbacks from larger datasets with demonstrated seasonality.

2. Data Processing

```
In [ ]: import kagglehub
        import os
        import pandas as pd
        import numpy as np
        from sklearn.ensemble import RandomForestRegressor
        import matplotlib.pyplot as plt
        from statsmodels.tsa.arima.model import ARIMA
        from sklearn.metrics import mean squared error
        from sklearn.inspection import permutation importance
        from statsmodels.graphics.tsaplots import plot_acf,plot_pacf
        path = kagglehub.dataset download("matthewjansen/bgc-jena-weather-station-dataset-20172024")
        files = os.listdir(path)
        csv file = [file for file in files if file.endswith('.csv')]
        dataset_path = os.path.join(path, csv_file[0])
        df = pd.read csv(dataset path)
        df.head()
```

Out[]:		Date Time	p (mbar)	T (degC)	Tpot (K)		rh (%)	VPmax (mbar)	VPact (mbar)		sh (g/kg)	 wv (m/s)	max. wv (m/s)	wd (deg)	rain (mm)	raining (s)	SWDR (W/m²)	
	0	2017- 01-01 00:10:00	999.77	-4.91	268.27	-8.41	76.3	4.24	3.23	1.00	2.01	 0.78	1.56	184.0	0.0	0.0	0.0	
	1	2017- 01-01 00:20:00	999.63	-5.05	268.13	-8.37	77.4	4.19	3.24	0.95	2.02	 1.52	1.92	202.6	0.0	0.0	0.0	
	2	2017- 01-01 00:30:00	999.54	-4.98	268.21	-8.38	76.9	4.21	3.24	0.97	2.02	 0.98	1.78	227.4	0.0	0.0	0.0	
	3	2017- 01-01 00:40:00	999.40	-4.88	268.33	-8.56	75.2	4.25	3.19	1.05	1.99	 1.16	1.80	212.5	0.0	0.0	0.0	
		2017-																

3 15

1 01

1 96

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2 64 222 1

0.0

0.0

0.0

5 rows × 22 columns

01-01

00:50:00

999.17

-5.17 268.06

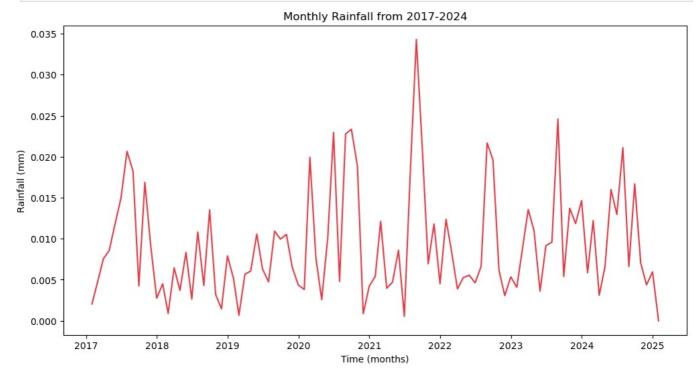
75.8

4 15

The BCG Jena Weather Station Dataset contains 22 measurements of weather related processes such as pressure, temperature, and namely, rainfall. In this project, we will be focusing on the rain measurements and exploring the predictive influence of previous time lags within the dataset. As collectively exhaustive and informative it may be to predicate our models upon the given 10 minute intervals of data from 2017 to 2024, we will be converting to monthly averages in order to reduce the computational expense of our calculations and potential overfitting with the unnecessary collection of insignificant white noise at higher frequencies. The transformed time series is plotted below.

```
In []: df = df[["Date Time", "rain (mm)"]]
    df["datetime"] = pd.to_datetime(df["Date Time"])
    df.set_index("datetime", inplace=True)
    df = df["rain (mm)"].resample("ME").mean().to_frame(name="Rainfall")

plt.figure(figsize=(12, 6))
    plt.plot(df, color="#E63946")
    plt.xlabel("Time (months)")
    plt.ylabel("Rainfall (mm)")
    plt.title("Monthly Rainfall from 2017-2024")
    plt.show()
```



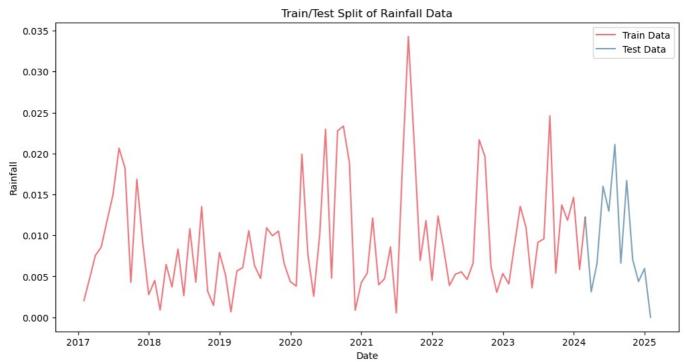
3. Fitting Time Series Models

3.1 Training-Testing Split

We will now partition the data into training and testing sets, with the cutoff set at the beginning of the final 12 months. This approach enables a direct comparison of various predictive models and their accuracy relative to the true data. The following visualization captures

the training-testing split.

```
In [ ]: split date = df.index[-12]
        df_train = df.loc[:split_date]
        df_test = df.loc[split_date:]
        X train, y train = df train.drop(columns=["Rainfall"]), df train["Rainfall"]
        X_test, y_test = df_test.drop(columns=["Rainfall"]), df_test["Rainfall"]
        forecast start = y test.index[0]
        forecast end = y test.index[-1]
        train_index = y_train.index
        test_index = y_test.index
        plt.figure(figsize=(12, 6))
        plt.plot(train_index, y_train, label="Train Data", color="#E63946", alpha=0.7)
        plt.plot(test_index, y_test, label="Test Data", color="#457B9D", alpha=0.7)
        plt.xlabel("Date")
        plt.ylabel("Rainfall")
        plt.title("Train/Test Split of Rainfall Data")
        plt.legend()
        plt.show()
```



3.2 Random Forest Method

For the first method in our analysis, we recreate the results of Tyralis and Papacharalampous. The code below features two functions: $create_lagged_features()$ and $rf_predictor()$. The first of the two creates lagged vectors for the Random Forest model to regress on to. For example, since our training data has 97 observations, the data frame includes the original rainfall measurements with an additional 97 columns with data lagged at $h = 1, h = 2, \dots, h = 97$. Finally, the function outputs the new data frame.

The second of the two functions performs the predictions using Random Forest. After running the *create_lagged_features()* function, *rf_predictor()* fits many different Random Forests with a specific proportion of the lagged columns. These proportions range from 0.05 to 0.5, with 0.05 increments. Tyralis and Papacharalampous suggest using 500 trees for the Random Forest models as a medium between Kuhn and Johnson [4], and Probst and Boulesteix [5], who suggest 1000 and 100 trees, respectively. Within each proportional fitting, the function identifies predictors with importance. We have selected the cutoff to have *any* importance (i.e. above 0). A new Random Forest is then fit with only important predictors. After all Random Forests are fitted, the function outputs a numpy array with the predictions from every iteration. One can easily extract the predictions from each iteration by slicing the array at increments of 12 (our forecast period).

```
In [ ]:
    def create_lagged_features(lags):
        for lag in range(1, lags + 1):
            df[f"lag_{lag}"] = df["Rainfall"].shift(lag)
        return df

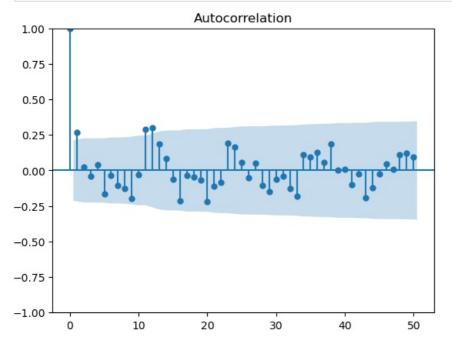
def rf_predictor(df):
    n1 = len(df)
    nfit = n1 - 1
```

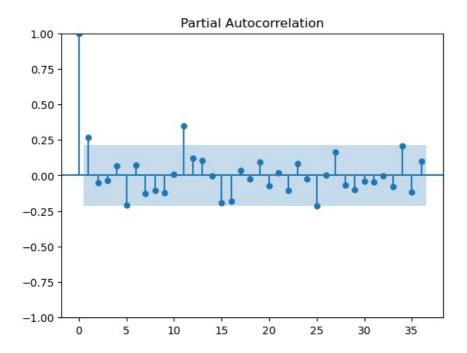
```
test_df = create_lagged_features(nfit)
   rf = RandomForestRegressor(n_estimators=500, random_state=1)
   predictions = []
   num lags = int(proportion * nfit)
       train_df = create_lagged_features(num_lags)
       X_train = train_df.drop(columns=["Rainfall"])
       y_train = train_df["Rainfall"]
       rf.fit(X train, y train)
       X test = test df.drop(columns=["Rainfall"])
       result = permutation_importance(rf, X_train, y_train, n_repeats=10, random_state=1)
       importance = result.importances_mean
       important lags = np.where(importance > 0)[0]
       important_train_df = X_train.iloc[:, important_lags]
       rf.fit(important_train_df, y_train)
       X_test_imp = X_test.iloc[:, important_lags]
       pred imp = rf.predict(X test imp)
       predictions.append(pred_imp)
   predictions = np.concatenate(predictions)
   return predictions
rf pred = rf predictor(y train)
```

3.3 Box-Jenkins Method

For the second method, we fit several ARIMA and SARIMA models using the Box-Jenkins methodology. Furthermore, in examining the ACF and PACF plots to identify model components at the discretion of the viewer, fitting several candidate models, and harnessing AIC to source the best model, Box-Jenkins will remain the most manual approach to weather forecasting in this entire study. Accordingly, we will be using the ACF and PACF to judge stationarity from a glance before using our judgement to fit models, which inherently produces model diagnostics in table format post-fitting procedures.

```
In [ ]: acf_vals = plot_acf(y_train, lags=50)
    pacf_vals = plot_pacf(y_train, lags=36)
```





3.3.1 ACF and PACF Analysis

In an effort to provide equal evaluation of the entire dataset, the above PACF and ACF plots cover the farthest possible range of previous lags in order to evaluate their statistical influence on the latest rainfall metric of the training set. In the assessment of stationarity, our extension of the lag bounds allow us to verify the stationarity and potential seasonality of our data's periodicity. Within this regard, we acknowledge the stationarity of our data with the evident majority of our data residing within the mildly-transparent confidence interval of stationary justification. Thus, our analyses have yielded the following considerations in our ARIMA model-fitting procedures:

- ACF-deduced Autoregressive (AR): 1 or 0
- Differencing (I): 0
- PACF-deduced Moving Average (MA): 1 or 0

However, within these bounds, there appear to be seasonal fluctuations with rainfall influence. Directly attributing this to known weather patterns of rainfall between Spring, Summer, Fall, and Winter of the 7 year examination, we will be incorporating a seasonal autoregressive integrated moving average (SARIMA) model in addition to our standard ARIMA models to account for potential underlying model violations in heteroskedasticity/trend stationarity.

```
In []: import warnings
        from statsmodels.tools.sm exceptions import ConvergenceWarning
        warnings.simplefilter('ignore', ConvergenceWarning)
        # ARIMA Fitting:
        bj_ma_model = ARIMA(y_train, order=(0, 0, 1))
        bj ar_model = ARIMA(y_train, order=(1, 0, 0))
        bj_arma_model = ARIMA(y_train, order=(1, 0, 1))
        print(bj ma model.fit().summary())
        print(bj_ar_model.fit().summary())
        print(bj arma model.fit().summary())
        # SARIMA Fitting:
        sbj_ma_model = ARIMA(y_train, order=(0,0,1), seasonal_order=(1,0,1,12)) \# best model
        sbj_ar_model = ARIMA(y_train, order=(1,0,0), seasonal_order=(1,0,1,12))
        sbj arma model = ARIMA(y train, order=(1,0,1), seasonal order=(1,0,1,12))
        print(sbj_ma_model.fit().summary())
        print(sbj_ar_model.fit().summary())
        print(sbj arma model.fit().summary())
```

```
# For future predictions
bj_pred = sbj_ma_model.fit().predict(start=forecast_start,end=forecast_end)
```

SARIMAX Results

Dep. Variabl	e:	Rainfal	 l No.	Observations:	:	 86	
Model:		ARIMA(0, 0, 1)		Likelihood		312.457	
Date:	Т	hu, 20 Mar 2025	5 AIC			-618.915	
Time:		20:48:29	9 BIC			-611.552	
Sample:		01-31-2017	•	2		-615.951	
Covariance T	vpe:	- 02-29-2024 opg					
:=======	coef	std err	, ======= Z	======== P> z	 [0.025	0.9751	
const na.L1	0.0092 0.2755 4.086e-05	0.001 0.082 6.78e-06	8.343 3.367 6.028	0.000 0.001 0.000	0.007 0.115 2.76e-05	0.011 0.436 5.41e-05	
-======	=======						====
.jung-Box (L Prob(Q):	1) (Q):		0.00 0.98	<pre>Jarque-Bera Prob(JB):</pre>	(JB):		16.46 0.00
leteroskedas		:	1.12	Skew:			1.01
rob(H) (two	-sided):		0.76	Kurtosis:			3.74
Warnings: [1] Covarian	ce matrix	calculated usin	ng the d	•	of gradients	(complex	-step).
ep. Variabl	====== e:	Rainfal	====== l No.	Observations:	:======= :	======= 86	
lodel:		ARIMA(1, 0, 0)		Likelihood		312.493	
ate:	Т	hu, 20 Mar 2025				-618.986	
ime:		20:48:29		-		-611.623	
ample:		01-31-2017 - 02-29-2024	•	_		-616.022	
ovariance T	ype:	op(
=======	coef	std err	z	P> z	[0.025	0.975]	
onst	0.0092	0.001	7.712	0.000	0.007	0.012	
r.L1	0.2684	0.099	2.722	0.006	0.075	0.462	
igma2 	4.083e-05	6.38e-06	6.401	0.000	2.83e-05 	5.33e-05	
jung-Box (L	1) (Q):		0.02	Jarque-Bera	(JB):		 17.61
rob(Q):			0.89	Prob(JB):			0.00
eteroskedas	ticity (H)	:	1.18	Skew:			1.01
rob(H) (two	-sided):		0.66	Kurtosis:			3.92
/arnings: 1] Covarian	ce matrix	calculated usin	ng the d		of gradients	(complex	-step).
ep. Variabl	e:	Rainfal	 l No.	Observations:	:	 86	
lodel:		ARIMA(1, 0, 1)) Log	Likelihood		312.570	
ate:	Т	hu, 20 Mar 2025	5 AIC			-617.140	
ime:		20:48:29				-607.323	
ample:		01-31-2017	•	2		-613.189	
		- 02-29-2024					
ovariance T 		op <u>(</u>	9 =======				
	coef	std err	Z	P> z	[0.025	0.975]	
onst	0.0092	0.001	7.734	0.000	0.007	0.012	
r.L1	0.1013	0.429	0.236	0.813	-0.739	0.942	
a.L1	0.1741	0.389	0.447	0.655	-0.589	0.937	
igma2	4.074e-05	7.01e-06	5.809	0.000	2.7e-05	5.45e-05	
======= jung-Box (L		==========	0.00	Jarque-Bera			===== 16.85
rob(Q):			0.96	Prob(JB):	•		0.00
eteroskedas		:	1.14	Skew:			1.01
rob(H) (two	-sided):		0.73	Kurtosis:			3.80
arnings: 1] Covarian ======ep. Variabl	=======	calculated usin		Results	of gradients	(complex	-step).
ep. variabi Iodel:		$MA(0, 0, 1) \times (1, 1)$			servations: kelihood		315.6
ate:	AUT		, 0, 1, 20 Mar 2		CETHOOG		-621.3
iale:		1114, 4		/\			
ime:		,	20:48				-609.0

01-31-2017

- 02-29-2024

HQIC

Sample:

-616.426

Covariance					========		
	coef	std err	Z	P> z	[0.025	0.975]	
const	0.0093	0.001	6.711	0.000	0.007	0.012	
ma.L1	0.1587	0.099	1.605	0.108	-0.035	0.352	
ar.S.L12 ma.S.L12	0.7579	0.340 0.419	2.230 -1.297	0.026	0.092 -1.366	1.424 0.278	
sigma2	-0.5437 3.706e-05	5.87e-06	6.316	0.195 0.000	2.56e-05	4.86e-05	
Ljung-Box ((L1) (Q):	========	0.03	Jarque-Bera	======== ı (JB):	 1	==== 1.14
Prob(Q):			0.87	Prob(JB):			0.00
	asticity (H)	:	0.86	Skew:			0.84
Prob(H) (tw	wo-sided): ========		0.69	Kurtosis:			3.50
=======		========	SARIMAX	Results		========	=======
Dep. Variat Model: Date: Time:	========= ole:	MA(1, 0, 0)	SARIMAX ======== Rai x(1, 0, [1] Thu, 20 Mar	Results 	 Observations	========	315.57 -621.13
====== Dep. Variat Model: Date: Time:	========= ole:	MA(1, 0, 0)	SARIMAX Rai (1, 0, [1] Thu, 20 Mar 20: 01-31	Results nfall No. , 12) Log 2025 AIC 48:30 BIC -2017 HQIC	Observations Likelihood	========	.=====================================
Dep. Variat Model: Date: Time: Sample:	ole: ARII Type:		SARIMAX 	Results nfall No. , 12) Log 2025 AIC 48:30 BIC -2017 HQIC	Observations Likelihood	========	315.57 -621.13
Dep. Variat Model: Date: Time: Sample:	ole: ARII	MA(1, 0, 0)	SARIMAX 	Results nfall No. , 12) Log 2025 AIC 48:30 BIC -2017 HQIC	Observations Likelihood	========	315.57 -621.13
Dep. Variat Model: Date: Time: Sample: Covariance	Type: coef	MA(1, 0, 0):	SARIMAX	Results	Observations Likelihood	0.975]	315.57 -621.13
Dep. Variat Model: Date: Time: Sample: Covariance const ar.L1	Type: coef 0.0093 0.1610	MA(1, 0, 0): std err 0.001 0.106	SARIMAX Rai (1, 0, [1] Thu, 20 Mar 20: 01-31 - 02-29	Results	Observations Likelihood [0.025 	0.975] 0.012 0.369	315.57 -621.13
======= Dep. Variat Model: Date: Time: Sample: Covariance ======= const ar.L1 ar.S.L12	Type: coef 0.0093 0.1610 0.6588	MA(1, 0, 0): std err 0.001 0.106 0.418	SARIMAX Rai (1, 0, [1] Thu, 20 Mar 20: 01-31 - 02-29	Results	0bservations Likelihood [0.025 	0.975] 0.012 0.369 1.478	315.57 -621.13
Dep. Variat Model: Date: Time: Sample: Covariance const ar.L1 ar.S.L12	Type: coef 0.0093 0.1610	MA(1, 0, 0): std err 0.001 0.106	SARIMAX Rai (1, 0, [1] Thu, 20 Mar 20: 01-31 - 02-29	Results	Observations Likelihood [0.025 	0.975] 0.012 0.369	315.57 -621.13
Dep. Variat Model: Date: Time: Sample: Covariance const ar.L1 ar.S.L12 ma.S.L12 sigma2	Type: coef 0.0093 0.1610 0.6588 -0.4571 3.803e-05	MA(1, 0, 0): std err 0.001 0.106 0.418 0.487 6.1e-06	SARIMAX Rai (1, 0, [1] Thu, 20 Mar 20: 01-31 - 02-29 2 6.928 1.517 1.577 -0.939 6.233	Results nfall No. , 12) Log 2025 AIC 48:30 BIC -2017 HQIC -2024 opg P> z 0.000 0.129 0.115 0.348 0.000	0bservations Likelihood [0.025 0.007 -0.047 -0.160 -1.412 2.61e-05	0.975] 0.012 0.369 1.478 0.497 5e-05	8 315.57 -621.13 -608.86 -616.26
Dep. Variate Model: Date: Fime: Sample: Covariance	Type: coef 0.0093 0.1610 0.6588 -0.4571 3.803e-05	MA(1, 0, 0): std err 0.001 0.106 0.418 0.487 6.1e-06	SARIMAX	Results	0bservations Likelihood [0.025 0.007 -0.047 -0.160 -1.412 2.61e-05	0.975] 0.012 0.369 1.478 0.497 5e-05	315.57 -621.13 -608.86 -616.26
Dep. Variate Model: Date: Time: Sample: Covariance	Type: coef 0.0093 0.1610 0.6588 -0.4571 3.803e-05	MA(1, 0, 0): std err 0.001 0.106 0.418 0.487 6.1e-06	SARIMAX Rai (1, 0, [1] Thu, 20 Mar 20: 01-31 - 02-29 6.928 1.517 1.577 -0.939 6.233 0.06	Results nfall No. , 12) Log 2025 AIC 48:30 BIC -2017 HQIC -2024 opg P> z 0.000 0.129 0.115 0.348 0.000	0bservations Likelihood [0.025 0.007 -0.047 -0.160 -1.412 2.61e-05	0.975] 0.012 0.369 1.478 0.497 5e-05	315.57 -621.13 -608.86 -616.26

Warnings:

[1] Covariance matrix calculated using the outer product of gradients (complex-step). SARIMAX Results

=======================================			===========
Dep. Variable:	Rainfall	No. Observations:	86
Model:	$ARIMA(1, 0, 1) \times (1, 0, 1, 12)$	Log Likelihood	315.014
Date:	Thu, 20 Mar 2025	AIC	-618.028
Time:	20:48:30	BIC	-603.302
Sample:	01-31-2017	HQIC	-612.101
	- 02-29-2024		

Covariance	· Type:		opg ====================================			
	coef	std err	Z	P> z	[0.025	0.975]
const	0.0094	0.001	7.420	0.000	0.007	0.012
ar.L1	0.0575	0.537	0.107	0.915	-0.995	1.110
ma.L1	0.1371	0.504	0.272	0.786	-0.850	1.125
ar.S.L12	0.2028	0.416	0.488	0.626	-0.612	1.017
ma.S.L12	0.0533	0.455	0.117	0.907	-0.838	0.945
sigma2	3.811e-05	6.62e-06	5.753	0.000	2.51e-05	5.11e-05
Ljung-Box	(L1) (Q):		0.00	Jarque-Bera	(JB):	9.08
Prob(Q):			0.97	Prob(JB):		0.03
Heterosked	lasticity (H)	:	0.93	Skew:		0.78
Prob(H) (t	:wo-sided):		0.84	Kurtosis:		3.30
========			=======	========	=========	

Warnings:

[1] Covariance matrix calculated using the outer product of gradients (complex-step).

As expected, our results' accuracy vary heavily with gaps of differing significance in each model. Within our examination of the simpler ARIMA models, a lack of statistical significance from our Jarque-Bera Statistics and Ljung-Box Q-Statistics reveal unreliability in the trends captured. However, as we examine the statistical significance of more complex models accounting for seasonality across periods of 12 months, we grasp trend statistics with greater reliability in forecasting capacity. Specifically, while the SARIMA(1,0,1) \times (1,0,1) $_{12}$ $model \ lacks \ statistical \ significance \ across \ multiple \ underlying \ assumptions, \ SARIMA(1,0,0) \times (1,0,1)_{12} \ and \ SARIMA(0,0,1) \times (1,0,1)_{12}$ retain statistically significant deductions, with a slightly greater edge in accuracy in the latter model observed from AIC. Albeit, a $manually\text{-}conceived \ estimation, \ we \ will \ be \ using \ the \ SARIMA(0,0,1) \times (1,0,1)_{12} \ as \ our \ ideal \ candidate, \ proceeding \ to \ use \ its \ "N-Steps \ our \ ideal \ candidate, \ proceeding \ to \ use \ its \ "N-Steps \ our \ ideal \ candidate, \ proceeding \ to \ use \ its \ "N-Steps \ our \ ideal \ candidate, \ proceeding \ to \ use \ its \ "N-Steps \ our \ ideal \ candidate, \ proceeding \ to \ use \ its \ "N-Steps \ our \ ideal \ candidate, \ proceeding \ to \ use \ its \ "N-Steps \ our \ ideal \ candidate, \ proceeding \ to \ use \ its \ "N-Steps \ our \ ideal \ candidate, \ proceeding \ to \ use \ its \ "N-Steps \ our \ ideal \ our \$ Ahead" prediction as a comparative metric later in our study.

3.4 Auto Arima Method

As our third and final method, we fit an Auto ARIMA model. While Auto ARIMA is widely regarded as highly inconsistent and unpredictable in its autonomous coefficient-selection process, its convenience in comparing a wide range of models allows us to find a reasonable-fitting model within a fraction of the time dedicated to manually-estimating in Box-Jenkins. Our parameter selection reasoning can be confined to the following:

- seasonal = True: confirming the presence of seasonality in our data
- trace = False: hiding pre-model processing for the sake of readability
- suppress_warnings = True: hiding warnings for greater readability
- stepwise=True: As it can be very computationally expensive in fully-evaluating every potential permutational (p,d,q) × (P,D,Q)_s sequence, we apply a stepwise search algorithm, reducing the search time with compromised "greediness" in a smaller possible statespace of evaluated models.
- information_criterion = 'aic': a selection indicating our benchmarking with the Aikage Information Criterion (AIC) metric for comparison
- random_state = 1: setting a seed across our simulation to create consistently comparable results for reproducability

```
# AUTO ARIMA Coding Module
from pmdarima import auto_arima
## Model-Fitting Procedures
monthly model = auto arima(
    y_train,
    seasonal = True,
   trace = False.
    suppress warnings = True,
    stepwise = True,
    stationary = True,
    information_criterion = 'aic',
    random_state = 1
print("Test 1: Monthly Model Summary")
print(monthly_model.summary())
print("\n")
## Forecasting
auto arima pred = monthly model.predict(n periods=len(y test))
```

Test 1: Monthly Model Summary

SARIMAX Results

=========							
Dep. Variab	le:		y No.	Observations	:	86	
Model:		SARIMAX(1, 0, 0) Log	Likelihood		312.493	
Date:		Thu, 20 Mar 202	5 AIC			-618.986	
Time:		20:48:30	0 BIC			-611.623	
Sample:		01-31-201	7 HQI	2		-616.022	
·		- 02-29-202	4				
Covariance ⁻	Type:	ope	g				
========	 coef	std err	====== Z	========= P> z	======= [0.025	0.975]	
intercept	0.0068	0.001	4.803	0.000	0.004	0.010	
ar.L1	0.2685	0.099	2.724	0.006	0.075	0.462	
sigma2	4.082e-05	6.38e-06	6.403	0.000	2.83e-05	5.33e-05	
Ljung-Box (======= L1) (Q):	=========	0.02	======= Jarque-Bera	======= (JB):	 1	7.61
Prob(Q):			0.89	Prob(JB):			0.00
Heteroskedasticity (H):			1.18	Skew:			1.01
Prob(H) (two	o-sided):		0.66	Kurtosis:			3.92
========	=======					========	====

Warnings:

[1] Covariance matrix calculated using the outer product of gradients (complex-step).

Upon calculation and model comparison, the Auto ARIMA has yielded the aformentioned SARIMAX(1,0,0) model, which has an integrate accountancy for exogenous variables, which in this case would keep seasonality and additional confounding influences in consideration. Analyzing the diagnostics, our Auto ARIMA model brings a mixed bag filled with great statistical promise and minor potential gaps. Rather, investigating the diagnostics of our model's table above, we see as how our model's strength resides in its low AIC, high log-likelihood, insignificant Ljung-Box Statistics, and stable variance, further justifying its strong fit. However, in examining the potential drawbacks we inspect the Jarque-Bera Test, which indicates a statistically significant result and potentially non-normally distributed residuals. Thus, while the Auto ARIMA model may have a fairly strong model fit, the potential weakness in its residuals may allow for the presence of non-linear effects in our data.

4. Model Comparison

4.1 Random Forest Method Selection

Now that all models have been fitted, we will be comparing their performance through a visual plot of their predictions against the testing data and a calculation of their test MSE. Note that there were 10 different proportions of lagged predictors from which Random Forests were fit, so the top 3 performers (based on test MSE) are included. These values are calculated below.

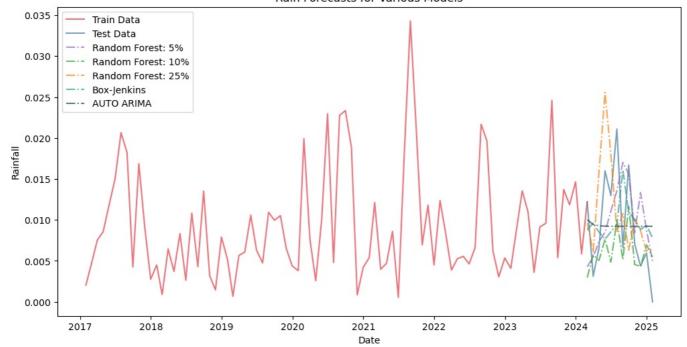
Out[]: Mean Squared Error 5% 0.000034 10% 0.000035 15% 0.000050 20% 0.000050 0.000047 25% 0.000056 30% 35% 0.000051 40% 0.000074 0.000059 45% 50% 0.000056

We can see that in descending order, the Random Forests which used 5%, 10%, and 25% of predictors had the best performance, respectively. These metrics validate the results of Tyralis and Papacharalampous, that Random Forests with a lower proportion of predictor lags exhibit the best performance. Below, we collect the predictions of these top performers along with the models derived from the Box-Jenkins methodology and Auto ARIMA.

4.2 Forecasting Results

```
# Comparable Visualization
plt.figure(figsize=(12, 6))
plt.plot(train_index, y_train, label="Train Data", color="#E63946", alpha=0.7)
plt.plot(test_index, y_test, label="Test Data", color="#45789D", alpha=0.7)
plt.plot(test_index, rf_pred[:12], label="Random Forest: 5%", color="#9467bd", alpha=0.7,ls="-.")
plt.plot(test_index, rf_pred[13:25], label="Random Forest: 10%", color="#2ca02c", alpha=0.7,ls="-.")
plt.plot(test_index, rf_pred[52:64], label="Random Forest: 25%", color="#ff7f0e", alpha=0.7,ls="-.")
plt.plot(test_index, bj_pred, label="Box-Jenkins", color="#2A9D8F", alpha=0.7,ls="-.")
plt.plot(test_index, auto_arima_pred, label="AUTO ARIMA", color="#264653", alpha=0.7,ls="-.")
plt.xlabel("Date")
plt.ylabel("Rainfall")
plt.title("Rain Forecasts for Various Models")
plt.legend()
plt.show()
```

Rain Forecasts for Various Models



The above plot summarizes the effectiveness of the Random Forest models in comparison to the other two methods. Primarily, we find that the Random Forest strikes a fair balance between the overfitting of the Box-Jenkins models and the underfitting of the Auto ARIMA. Additionally, the Random Forests which used 5% and 10% of the engineered lags appear to capture the pattern of the time series. We confirm this claim by calculating the test MSE, found below.

out[]:		Method	Mean Squared Error
	0	Random Forest: 5%	0.000034
	1	Random Forest: 10%	0.000035
	2	Auto Arima	0.000037
	3	Box-Jenkins	0.000042

0.000047

From the above table, we confirm the previous intuition that two of the three Random Forest models outperform the other methods. In particular, the Random Forest model utilizing 5% of the lagged predictors had a significantly lower test MSE, reflecting its ability to properly adapt to the patterns in the data.

5. Conclusion

4 Random Forest: 25%

In this project, we validated the results of Tyralis and Papacharalampous that time series forecasts can be efficiently performed utilizing Random Forests trained on small proportions of lagged data. Further, we compared the prediction accuracy of the Random Forests with other contemporary methods such as Box-Jenkins and Auto ARIMA. In comparing test MSE, we found that the Random Forest models outperformed the other methods.

Of course, our results are not without fault. In the function utilized to fit the Random Forest models, we set the variable importance cutoff to be zero. In doing this, we allowed predictors that may have had minute impact to be included in the final model, which may have contributed to overfitting. Future research may improve upon this by introducing the cutoff as a validation parameter. Further, while we were able to cross validate by fitting several Random Forest models, increased folds with larger data sets may provide more detailed insights.

Overall, these results provide an exciting perspective on time series modeling. While the Box-Jenkins methodology has stood the test of time, new methods continue to emerge, offering fresh approaches to improving accuracy and adaptability in an ever-evolving field.

References

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Processing math: 100%