

# A Comparison of Methods for Estimating $u_*$ From Given $u_z$ and Air-Sea Temperature Differences

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This paper presents an objective assessment of four methods for estimating sea surface friction velocity  $u_*$  from wind speed at height  $z$ ,  $u_z$ , and air-sea temperature difference. The methods are compared by using the computed friction velocity as the normalization factor in parametric correlations of wind-wave parameters with the wind and wave measurements made by NOMAD buoys in the Great Lakes. The results show that (1) wind profile parameters obtained from the four methods are generally comparable, (2) parametric correlations with parameters normalized by  $u_*$  lead to significantly decreased percentage deviations over correlations normalized by  $u_z$ , and (3) correlations based on  $u_*$  derived from the four methods show nearly identical percentage deviations. The conclusions are that (1)  $u_*$  normalization acts to eliminate the effect of atmospheric stability and (2) any of the four methods can be used effectively in practical applications.

## 1. INTRODUCTION

One of the basic tools for studying wind waves, surface currents, mixing, and sea-air interactions in the upper ocean is to nondimensionalize the relevant parameters in order to find universal correlations for use with theoretical models. Wind speed at 10 m above sea surface,  $u_{10}$ , has frequently been used in the nondimensionalization. While it is convenient to use  $u_{10}$ , many routinely available wind speeds are measured at other levels, that is, 19.5 m from a ship or 5 m from a buoy. To reconcile these different wind measurements as well as to eliminate the effects of surface roughness and atmospheric stability, the friction velocity  $u_*$  is used for nondimensionalization. As  $u_*$  is not a readily measured parameter, a number of empirical methods have been developed to obtain  $u_*$  from measurements of  $u_z$  and air-sea temperature differences. In this paper we examine four of these methods and apply them to the wind, wave, and temperature data recorded from NOMAD buoys deployed in the Great Lakes. Since in this case there is no measured  $u_*$  for direct verification of the results, we simply apply the estimated  $u_*$  to correlations of wind-wave parameters. We expect that an effectively estimated  $u_*$  will be free from atmospheric stability effects, and hence its application in the parametric wind-wave correlations will act to reduce the scatter.

## 2. METHODS

The four empirical methods we examine in this paper are (1) the Great Lakes Environmental Research Laboratory (GLERL) method [Schwab, 1978] based on the works of Long and Shaffer [1975], Businger et al. [1971], and Smith and Banke [1975], (2) the Kondo method based on Kondo [1975], (3) the Large and Pond method based on Large and Pond [1981, 1982] and Keller et al. [1985], and (4) the Smith method based on Smith [1981] and Dyer [1975]. All four methods start with the similarity theory of Monin and Obukhov [1954] that expresses wind shear and temperature gradi-

ent in the form

$$\partial u / \partial z = (u_* / kz) \phi_m(\zeta) \quad (1)$$

$$\partial \theta / \partial z = (\theta_* / kz) \phi_h(\zeta) \quad (2)$$

respectively, where  $u_*$  is the friction wind velocity,  $\theta_*$  is the scaling temperature,  $k = 0.4$  is the von Karman constant,  $\zeta = z/L$  is the dimensionless stability height, and  $L = \theta_*^2 / (gk\theta)$  is the Monin-Obukhov stability length.

One difference in the methods tested here is in the definition of functions  $\phi_m(\zeta)$  and  $\phi_h(\zeta)$  by the authors listed in Table 1. The functions are shown graphically in Figure 1. While the differences are almost nonexistent for unstable cases with  $\zeta < 0$ , the functions tend to diverge, in some cases significantly, for the stable case with large  $\zeta$ . Following Paulson [1970] we can integrate (1) and (2) to give the wind and temperature profiles as:

$$u = (u_* / k) [\ln(z/z_0) - \psi_m] \quad (3)$$

$$\theta = \theta_0 + (\theta_* / k) [\ln(z/z_0) - \psi_h] \quad (4)$$

where  $z_0$  is the surface roughness length,  $\theta_0$  is the extrapolated temperature at  $z = 0$ , and the expressions of  $\psi_m(\zeta)$  and  $\psi_h(\zeta)$  corresponding to  $\phi_m$  and  $\phi_h$  are those listed in Table 2. Note in Table 2 that Long and Shaffer [1975] further divide the stable condition into two parts: the upper part for mildly stable cases (that is,  $0 < \zeta < 1$ ) and the lower part for strongly stable cases ( $\zeta \geq 1$ ). These solutions are functions of three unknown parameters:  $u_*$ ,  $z_0$ , and  $L$ . Different authors have devised their own methods for estimating  $L$ . In addition,  $u_*$  can be linked with wind speed  $u_z$  or  $z_0$  through various theoretical or empirical relations. The Charnock [1955] relation, for example, states that  $gz_0/u_*^2$  is a universal constant, where  $g$  is the gravitational acceleration; hence  $z_0$ , characterizing the aerodynamic roughness at the surface, is proportional to  $u_*^2$ . Smith [1981] further postulated that roughness length for a smooth surface should be added to achieve gradual transitions from smooth to rough surface conditions. Thus

$$z_0 = au_*^2 / g + (v/u_*) \exp(-5.5k) \quad (5)$$

where  $a$  is the Charnock constant, and  $v$  is the dynamic viscosity of air ( $\approx 14 \times 10^{-6} \text{ m}^2 \text{ s}^{-1}$ ).

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TABLE 1. Wind Shear and Temperature Gradient Functions as Given by Different Authors

	$\phi_m(\zeta)$		$\phi_h(\zeta)$	
	Stable	Unstable	Stable	Unstable
<i>Businger et al.</i> [1971] (GLERL)	$1 + 4.7\zeta$	$(1 - 15\zeta)^{-1/4}$	$0.74 + 4.7\zeta$	$0.74(1 - 9\zeta)^{-1/2}$
<i>Kondo</i> [1975]	$1 + 6\zeta/(1 + \zeta)$	$(1 - 16\zeta)^{-1/4}$	$1 + 6\zeta/(1 + \zeta)$	$(1 - 16\zeta)^{-1/2}$
<i>Large and Pond</i> [1981]	$1 + 7\zeta$	$(1 - 16\zeta)^{-1/4}$	$1 + 7\zeta$	$(1 - 16\zeta)^{-1/2}$
<i>Dyer</i> [1975], <i>Smith</i> [1981]	$1 + 5\zeta$	$(1 - 16\zeta)^{-1/4}$	$1 + 5\zeta$	$(1 - 16\zeta)^{-1/2}$

In general, all four methods follow a two-stage approach: first, obtain the neutral bulk coefficient and then use air-sea temperature differences to estimate the diabatic conditions. In Smith's method the neutral bulk coefficient is given by

$$C_{zn} = k/\ln(z/z_0)^2 \quad (6)$$

which can be estimated iteratively along with (5). Kondo's method, on the other hand, obtains  $C_{zn}$  from the neutral 10-m-level-bulk coefficient by

$$C_{zn} = k^2[kC_{10n}^{-1/2} + \ln(z/10)]^{-2} \quad (7)$$

In the work by *Kondo* [1975],  $C_{10n}$  is empirically related to the wind speed by:

$$\begin{aligned} 10^3 C_{10n} &= 1.08 u_{10}^{-0.15} & 0 < u_{10} \leq 2.2 \text{ ms}^{-1} \\ 10^3 C_{10n} &= 0.771 + 0.0858 u_{10} & 2.2 < u_{10} \leq 5 \text{ ms}^{-1} \\ 10^3 C_{10n} &= 0.867 + 0.0667 u_{10} & 5 < u_{10} \leq 8 \text{ ms}^{-1} \\ 10^3 C_{10n} &= 1.2 + 0.025 u_{10} & 8 < u_{10} \leq 25 \text{ ms}^{-1} \\ 10^3 C_{10n} &= 0.773 & 25 < u_{10} < 50 \text{ ms}^{-1} \end{aligned} \quad (8)$$

Note here that  $u_{10}$  is the neutral 10-m-level wind speed. To proceed with a nonneutral  $u_z$ , *Kondo* [1975] gave a simple iteration procedure that can be used to obtain an estimate of neutral  $u_{10}$ . *Large and Pond's* [1981] first-stage approach is similar to Kondo's method except that they used a simpler empirical formulation for  $C_{10n}$ :

$$\begin{aligned} 10^3 C_{10n} &= 1.14 & 0 < u_{10} \leq 10 \text{ ms}^{-1} \\ 10^3 C_{10n} &= 0.49 + 0.065 u_{10} & 10 < u_{10} < 50 \text{ ms}^{-1} \end{aligned} \quad (9)$$

The GLERL approach uses the Charnock relation to obtain  $z_0$  with the Charnock constant determined from setting  $u_{10} = 15 \text{ ms}^{-1}$  in the  $C_{10n}$  versus  $u_{10}$  relation given by *Smith and Banke* [1975] as:

$$10^3 C_{10n} = 0.63 + 0.066 u_{10} \quad (10)$$

In the second stage the GLERL method uses the iteration scheme of *Long and Shaffer* [1975] to obtain  $L$  and hence the values of  $\psi$ . *Large and Pond* [1981] developed semitheoretical formulas for directly calculating  $\zeta$  from given wind speed  $u_z$ , air temperature  $T_a$ , and sea temperature  $T_s$ , in which the  $u_z$  calculation is given by *Keller et al.* [1985] as:

$$u_z = u_z(C_{zn})^{1/2}/[1 - (C_{zn})^{1/2}\psi_m/k] \quad (11)$$

The *Smith* [1981] approach requires iterations in which  $u_z$ ,  $\theta_z$ , and  $\psi$  functions are adjusted to yield the desired  $C_{zn}$ . *Kondo* [1975] perhaps provided the most simplified approach by developing approximate formulas for calculating  $C_z$  from a neutral  $C_{zn}$ .

Figure 2 presents plots of  $C_{10}$  versus  $u_{10}$  for different air-sea temperature differences with the four methods. The neutral stratification case is shown by the heavy line. The thin lines correspond to air-sea temperature differences ranging from  $-20^\circ\text{C}$  (curves above the neutral curve) to  $+20^\circ\text{C}$  (curves below the neutral curve) at  $1^\circ\text{C}$  increments. The main differences are at wind speeds less than  $2.5 \text{ ms}^{-1}$ ; the Kondo, Large-Pond, and Smith methods show the unstable drag coefficient increasing as wind speed decreases to zero, while the GLERL method shows it continuously decreasing. This is clearly a result of the different formulations of the methods.

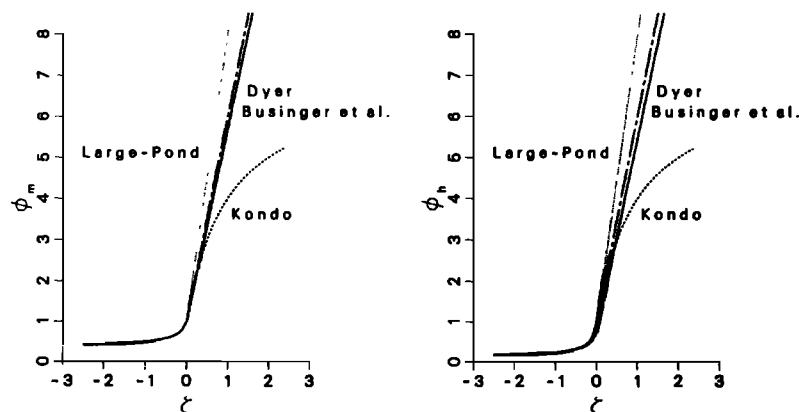


Fig. 1. Comparison of different wind shear and temperature gradient formulas.

TABLE 2. Diabatic Wind and Temperature Profiles as Given by Different Authors

	$\psi_m(\zeta)$	
	Stable	Unstable*
<i>Long and Shaffer</i> [1975] (GLERL)	$0.74 \ln (\zeta/\zeta_0) + 4.7(\zeta - \zeta_0)$	$\ln [(x-1)(1+x_0)/(1+x)(x_0-1)]$ $+ 2[\tan^{-1} x - \tan^{-1} x_0]$
<i>Kondo</i> [1975]	$-\ln \zeta_0 + 4.7/\zeta_0 + 5.7 \ln \zeta$	$\ln [(x_0^2 + 1)(x_0 + 1)^2/(x^2 + 1)(x + 1)^2]$ $+ \ln (\zeta/\zeta_0) + 2[\tan^{-1} x - \tan^{-1} x_0]$
<i>Large and Pond</i> [1981]	$-7.7\zeta$	$2 \ln [(1+x)/2] + \ln [(1+x^2)/2]$ $- 2 \tan^{-1} x + \pi/2$
<i>Smith</i> [1981]	$-5.7\zeta$	$2 \ln [(1+x)/2] + \ln [(1+x^2)/2]$ $- 2 \tan^{-1} x + \pi/2$
	$\psi_h(\zeta)$	
	Stable	Unstable*
<i>Long and Shaffer</i> [1975] (GLERL)	$\ln (\zeta/\zeta_0) + 4.7(\zeta - \zeta_0)$	$0.74 \ln [(x^2 - 1)(x_0^2 + 1)/(x^2 + 1)(x_0^2 - 1)]$
<i>Kondo</i> [1975]	$-0.74 \ln \zeta_0 + 4.7/\zeta_0 + 5.44 \ln \zeta$	$\ln [(x+1)(x_0+1)/(x_0-1)(x-1)]$
<i>Large and Pond</i> [1981]	$-\ln (\zeta/\zeta_0) + 6 \ln [(1+\zeta)/(1+\zeta_0)]$	$\ln [(1+x^2)/2]$
<i>Smith</i> [1981]	$-5.7\zeta$	$2 \ln [(1+x^2)/2]$

\*Where  $x = (1 - a\zeta)^{1/4}$ ,  $x_0 = (1 - a\zeta_0)^{1/4}$ , and  $a = 15$  or  $16$ .

While the methods have different emphases which possibly tend to accentuate each author's own data, the differences are generally no greater than the data point scatter from which they were derived. All four methods can be readily used to estimate a wind profile and hence  $u_*$  from a given wind speed and air-sea temperature difference. In the following we shall compare the results as they are applied to actual wind and wave measurements.

### 3. DATA

There have been eight NOMAD buoys moored in the Great Lakes since 1981 in water depths ranging from 15 to 250 m (Figure 3). These buoys are boat shaped, 6 m in length, with an electronic payload for measuring wind speed, wind direction, barometric pressure, air temperature, sea surface temperature, and surface wave spectral data. Most of the meteorological sensors are located 5 m above the water surface. Data are reported hourly. The wind speed and direction, as well as air and surface water temperatures, are 8.5-min averages of samples obtained at 1-s intervals. The waves are measured with an accelerometer using an on-board Wave Data Analyzer system [Steele and Johnson, 1977] that transmits acceleration spectral data via the UHF GOES satellite to a shore collecting station, where wave frequency spectra with 48 degrees of freedom are calculated from 20 min of measurements. In this study we examined all the data recorded during 1981–1984, an average of over 4000 hourly measurements from each buoy every year. The results from different buoys and from

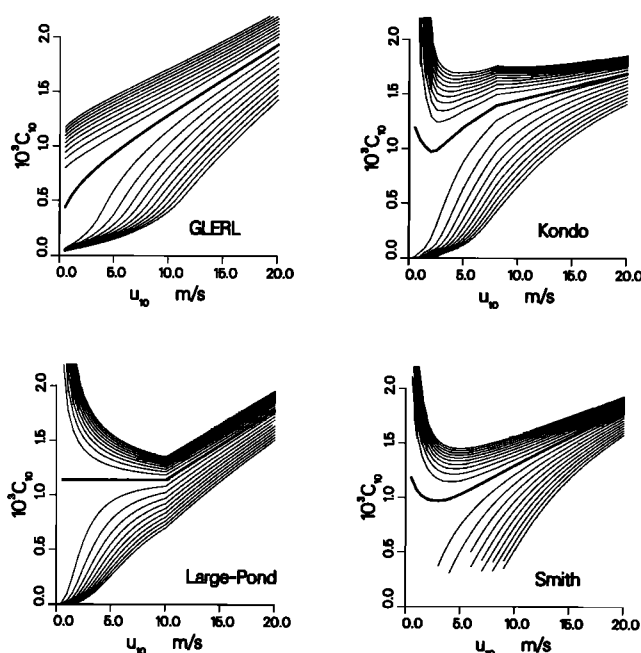


Fig. 2. Comparison of different bulk transfer coefficient formulas as a function of 10-m wind speed for air-sea temperature differences from  $-20^\circ\text{C}$  to  $+20^\circ\text{C}$  at  $1^\circ\text{C}$  increment. The heavy line is the neutral case.

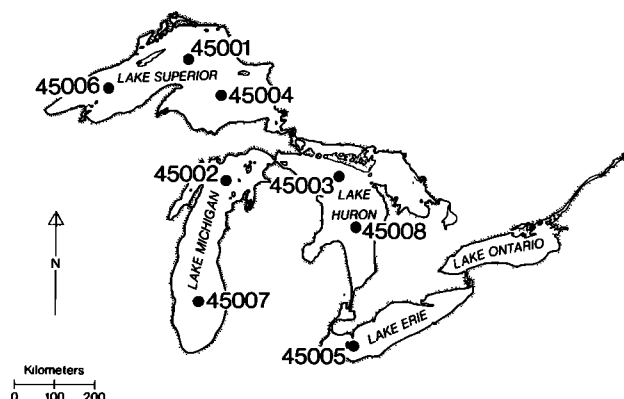


Fig. 3. Great Lakes map showing the locations of the eight National Data Buoy Center NOMAD buoys.

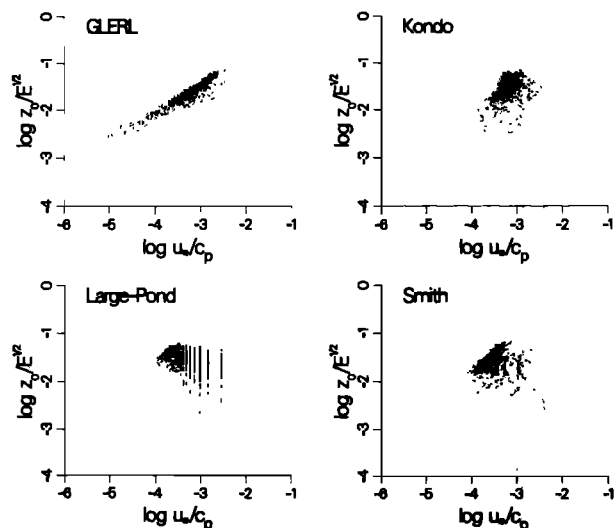


Fig. 4. Comparison of  $z_0/E^{1/2}$  versus  $u_*/c_p$  correlations.

different years are very similar, however. Therefore the detailed analyses presented here use data from the 1984 northern Lake Michigan buoy number 45002 as representative of all the buoys. Buoy 45002 recorded 4212 simultaneous measurements of wind speed, air-sea temperature difference, significant wave height, and peak energy wave period in 1984 with a maximum wind speed of  $15 \text{ ms}^{-1}$  and a maximum significant wave height of 4.5 m. In the following analyses, calculations were based on all of the 4212 data points, while the graphs show only 1056 points (about every fourth point) in order to reduce crowding.

#### 4. RESULTS AND DISCUSSION

##### 4.1. Roughness Length

The purpose of the empirical methods discussed above is to estimate values for the roughness length  $z_0$  and for friction velocity  $u_*$  when these cannot be measured directly. The roughness length  $z_0$  is conceptually an important parameter in the study of wind waves. However, it is difficult to measure or

even define physically. Of the four methods we considered here, only the GLERL and Smith methods use  $z_0$  explicitly; the Kondo and Large-Pond (as interpreted by Keller *et al.* [1985]) methods do not, although a value for  $z_0$  can be estimated from these methods. To examine these  $z_0$  estimations in connection with sea state studies, we follow Huang *et al.* [1986] and plot  $z_0$  normalized by rms wave height  $E^{1/2}$  versus  $u_*$  normalized by peak energy wave speed  $c_p$ . The results are shown in Figure 4. The GLERL method is the only one that shows a consistent correlation between normalized values of  $z_0$  and  $u_*$ ; this merely reflects the use of the Charnock relation in that method. The Smith method uses a nonlinear function for  $z_0$  that approaches the Charnock relation at large  $u_*$ . The Kondo and Large-Pond methods do not use  $z_0$  explicitly. None of these methods show any clear correlation between the normalized  $z_0$  and  $u_*$ . Since all four methods are primarily formulated to estimate the drag coefficient  $C_z$  and the friction velocity  $u_*$  under diabatic conditions,  $z_0$  did not receive a detailed treatment. The results of the correlations reflect this effect. The exact form of the relation between  $z_0$  and  $u_*$  is still an active area of research; Greenaert *et al.* [1986], for example, discuss six different models for the relation.

##### 4.2. $u_*$ Normalizations

To examine the results for  $u_*$ , we apply the estimated  $u_*$  to the normalization of wind wave parameters. For a given wind and wave field with wave spectral energy density  $S(f)$ , total wave energy  $E = \int S(f) df$ , peak energy wave frequency  $f_m$ , and fetch distance  $F$ , the following parameters have been frequently used in the literature [e.g., Hasselmann *et al.*, 1973; Mitsuyasu *et al.*, 1980]:  $\epsilon_* = gE/u_*^2$ , the nondimensional energy;  $\nu_* = f_m u_*/g$ , the nondimensional peak energy frequency; and  $\xi_* = gF/u_*^2$ , the nondimensional fetch. We affix a subscript  $z$  to the parameters which are normalized directly with measured wind speed  $u_z$ .

Correlating these various parameters has led to a number of universal power law relations that played important roles in developing numerical wave prediction models. Hasselmann *et al.* [1973, 1976], Mitsuyasu *et al.* [1980], and Toba [1978] have all deduced similar empirical equations characterizing the correlations among the parameters based on their own

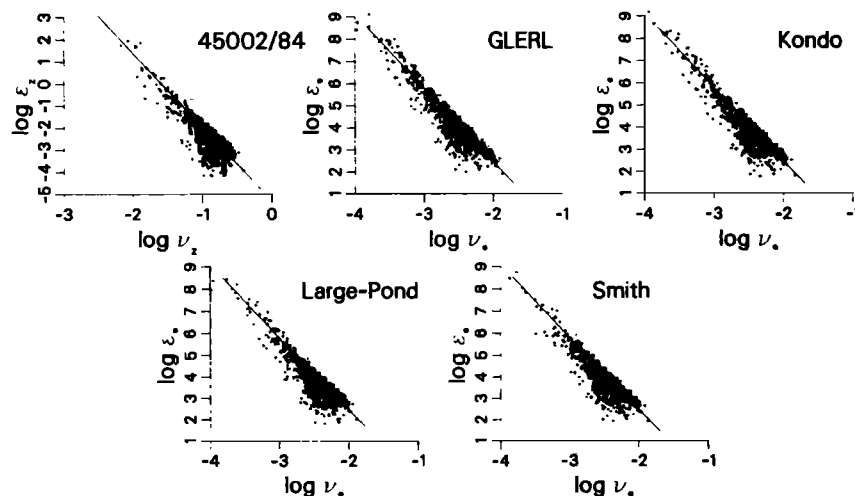


Fig. 5. Comparison of  $\epsilon_z$  versus  $\nu_z$  correlations. The straight lines are given by JONSWAP relations  $\epsilon = 5.3 \times 10^{-6} \nu^{-10/3}$  and  $\epsilon_* = 5.3 \times 10^{-5} \nu_*^{-10/3}$ .

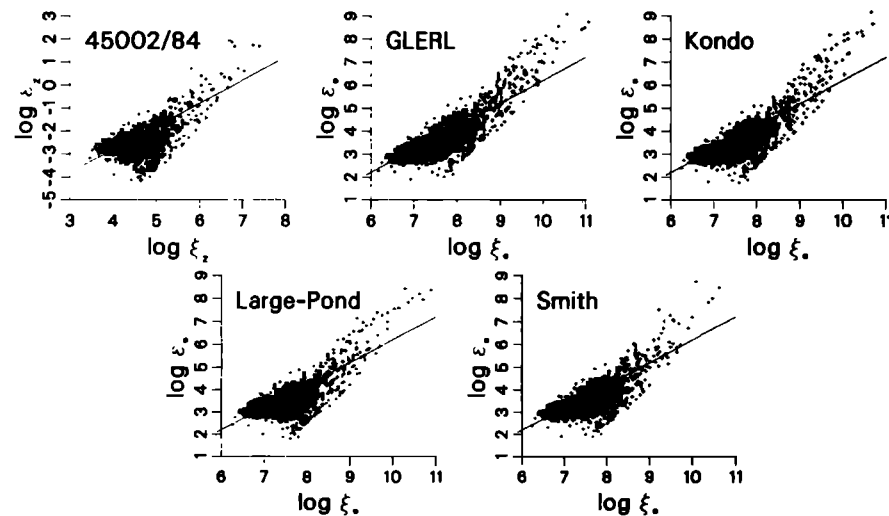


Fig. 6. Comparison of  $\varepsilon$  versus  $\xi$  correlations. The straight lines are given by JONSWAP relations  $\varepsilon = 1.6 \times 10^{-7} \xi$  and  $\varepsilon_s = 1.6 \times 10^{-4} \xi_s$ .

measurements. An examination of these universal correlations was given by Liu [1984, 1985].

Figures 5–7 present the three intercorrelations  $v$  versus  $\xi$ ,  $\varepsilon$  versus  $\xi$ , and  $\varepsilon$  versus  $v$ , respectively. In each figure we present the  $u_z$  normalization graph along with four  $u_x$  normalization graphs corresponding to the four methods. The empirical JONSWAP relations [Hasselmann *et al.*, 1973, 1976] for the respective parameters are plotted as the straight line in each graph for reference. As discussed by Liu [1985], the data points appear to be clustered around galaxylike regions rather than universal lines. The straight lines provide only a crude approximation of the data that can be accurate at times and erroneous at other times. This may be due to failure of the simple power law relations in unsteady conditions or to other processes. The power laws were developed mainly from growing sea states, whereas the data in Figures 5–7 represent both growing and dissipating waves. As shown by Liu [1985], however, even using carefully selected growth episodes does not lead to less scatter in the correlations. Clearly, none of the

four methods of  $u_x$  estimation show any particular advantage in reducing the scatter. The scatter shown in the  $u_x$  normalization graphs is somewhat less than the  $u_z$  normalization graphs.

#### 4.3. Assessments

In order to examine the correlations shown in Figures 5–7 on a quantitative basis we calculated and compared several statistical entities relevant to the correlations. Specifically, we sought the general relation  $y = cx^b$  from our data and analyze the results statistically. Since all the correlations are plotted on the log-log scale, we simply let  $X = \log(x)$ ,  $Y = \log(y)$ , and  $a = \log(c)$ , and by fitting a straight line of the form  $Y = a + bX$  through the data points by least squares method we can calculate the following:

$$\text{standard error} = [\Sigma(Y - a - bX)^2 / (n - 2)]^{1/2} \quad (12)$$

$$\text{percentage deviation} = \Sigma|(Y - a - bX)/Y|(100/n) \quad (13)$$

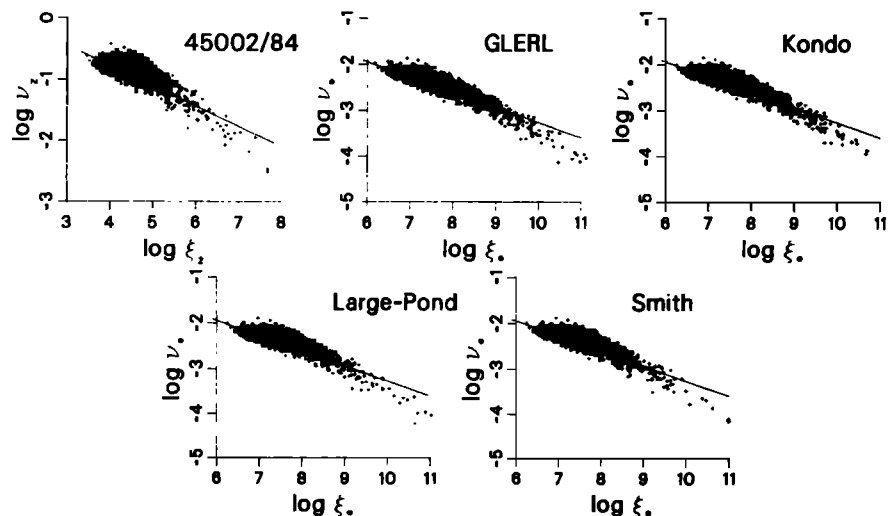


Fig. 7. Comparison of  $v$  versus  $\xi$  correlations. The straight lines are given by JONSWAP relations  $v = 3.5\xi^{-0.33}$  and  $v_s = 1.08\xi_s^{-0.33}$ .

TABLE 3. Summary of Calculated Standard Errors and Percentage Deviations for the Various Correlations

Nomalization	Standard Error			Percentage Deviation		
	$\varepsilon$ versus $v$	$\varepsilon$ versus $\xi$	$v$ versus $\xi$	$\varepsilon$ versus $v$	$\varepsilon$ versus $\xi$	$v$ versus $\xi$
$u_z$	0.384	0.578	0.133	17.24	30.27	11.98
GLERL $u_a$	0.373	0.581	0.135	6.78	11.83	4.49
Kondo $u_a$	0.393	0.629	0.140	7.81	13.53	4.71
Large-Pond $u_a$	0.408	0.655	0.141	8.44	14.08	4.73
Smith $u_a$	0.364	0.540	0.130	7.17	11.52	4.38

where  $n$  is the total number of data points (4212 for our 1984 North Lake Michigan buoy data set). Table 3 lists the calculated standard errors and percentage deviations for each of the 15 correlations shown in Figures 5–7. A well-fitted correlation should provide lower standard error and lower percentage deviation.

An examination of Table 3 shows that based on the calculated standard errors alone we are still unable to make any distinction among the various normalizations. For each of the three correlations,  $\varepsilon$  versus  $v$ ,  $\varepsilon$  versus  $\xi$ , and  $v$  versus  $\xi$ , the standard errors are virtually the same. Since it is applied to  $\log(x)$  and  $\log(y)$ , an average standard error of 0.4, as in the case of  $\varepsilon$  versus  $v$ , indicates that the actual estimates can vary by a factor of 2.5 from the regression line. Similarly, the average standard errors of 0.6 ( $\varepsilon$  versus  $\xi$ ) and 0.135 ( $v$  versus  $\xi$ ) lead to variation factors of 4.0 and 1.4, respectively.

When we examine the calculated percentage deviations, we find that there are significantly lower percentage deviations for the  $u_a$  normalizations. Among the four methods, although there are slight variations in the percentage deviations, none shows any particular advantage in improving the fit significantly. In general, either the GLERL, Kondo, Large-Pond, or Smith method can be used to provide  $u_a$  as well as  $u_{10}$  estimations for practical applications.

In the above analysis we predetermined the functional form of  $y = ax^b$  and chose the dependent and independent variables  $y$  and  $x$ , respectively, according to the familiar JONSWAP formulations (that is,  $\varepsilon$  in terms of  $v$ ,  $\varepsilon$  in terms of  $\xi$ , and  $v$  in terms of  $\xi$ ). We also tried correlating them inversely by interchanging  $y$  with  $x$  as dependent variables as well as estimating the slope coefficient in the linear relationship by the maximum likelihood method [e.g., Kendall and Stuart, 1973]. While both of these analyses resulted in different sets of numbers, the implications are precisely identical to our previous discussions. Hence alternate analysis will not alter the results presented here.

In this analysis we have used all the data available without discriminating between cases representative of swell or wind waves. This is mainly because there is so little swell activity in Great Lakes waves that separation of swell cases from wind wave cases would not improve the results significantly.

In Figure 2 the GLERL method is considerably different from the other three methods at low wind speeds for unstable cases. Approximately 10% of the data we used in the analysis were of unstable cases with wind speed less than  $5 \text{ ms}^{-1}$ . However, we find there is no discernible difference due to this effect in the results. Apparently, this divergence does not have a significant effect on the parametric correlations.

## 5. CONCLUDING REMARKS

From detailed examination of the four methods used to estimate  $u_a$  from a given  $u_z$  and air-sea temperature difference and application of these estimates to wind-wave parametric correlations we expected to distinguish a method which effectively reduced the scatter. We find instead that the four methods examined are virtually indistinguishable. Because all methods follow the same basic approach, with different but comparable empirical formulations, this result is by no means surprising. In practice, any of the four methods will provide similar estimates of  $u_a$ . Until further detailed wind stress and wind profile measurements can be conducted to ascertain or validate the empirical formulations these methods will remain useful tools in providing wind stress, bulk transfer coefficients, and wind profile estimations where only minimal input data are available.

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