### Fundamentals of Mathematical Statistics

Confidence Intervals

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ECN6330 Agenda for 02/04/2011

Brief Review of Some Concepts

Mathematical Statistics

Confidence Intervals

## Agenda for Tonight

- Quiz 3
- Brief Review of Statistical Inference
- Discuss further topics in Mathematical Statistics:
  - Large sample properties of estimators
  - Confidence intervals
  - Hypothesis testing

#### The Normal Distribution

The normal probability density function is defined as follows:

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp{\frac{-(X-\mu)^2}{2\sigma^2}}, \quad \text{for } -\infty < x < \infty$$

where  $E(X) = \mu$  and  $Var(X) = \sigma^2$ . When a random variable is normally distributed we write  $X \sim N(\mu, \sigma^2)$ .

#### The Normal Distribution

A special case is the standard normal distribution, which is defined as follows:

$$\phi(z) = \frac{1}{\sqrt{2\pi}} \exp\left\{\frac{-z^2}{2}\right\}, \quad \text{for } -\infty < z < \infty$$

The standard normal cumulative distribution function is denoted by  $\Phi(z) = P(Z \le z)$ . Using some basic facts from probability we arrive at the following helpful formulas:

$$P(Z > z) = 1 - \Phi(z)$$
  
 $P(Z < -z) = P(Z > z)$   
 $P(a \le Z \le b) = \Phi(b) - \Phi(a)$ 

## The Chi-Square Distribution

The chi–square distribution is obtained directly from independent, standard normal random variables. Let  $Z_i$ ,  $i=1,2,\ldots,n$ , be independent random variables, each distributed as standard normal. Define a new random variable as the sum of the squares of the individual  $Z_i$ :

$$X = \sum_{i=1}^{n} Z_i^2$$

The new random variable X has a chi–square distribution with n degrees of freedom. This is often written as  $X \sim \chi_n^2$ .

#### The Student T Distribution

The t distribution is a workhorse in classical statistics and econometrics. A t distribution is obtained from a standard normal and a chi–square random variable. Let Z have a standard normal distribution and let X have a chi-square distribution with n degrees of freedom. Also assume that Z and X are independent. Then the following random variable

$$T = \frac{Z}{\sqrt{X/n}}$$

has a t distribution with n degrees of freedom. This is denoted by  $T \sim t_n$ . The t distribution gets its degrees of freedom from the chi–square random variable.

#### The F Distribution

Another important distribution for statistics and econometrics is the F distribution. To define an F random variable, let  $X_1 \sim \chi^2_{k_1}$  and  $X_2 \sim \chi^2_{k_2}$  and assume that  $X_1$  and  $X_2$  are independent. Then, the random variable

$$F = \frac{X_1/k_1}{X_2/k_2}$$

has an F distribution with  $(k_1, k_2)$  degrees of freedom. We denote this as  $F \sim F_{k_1, k_2}$ . The order of the degrees of freedom is important.  $k_1$  is the numerator degrees of freedom and  $k_2$  is the denominator degrees of freedom.

## Large Sample Properties of Estimators

We saw with the estimator of  $\mu$ ,  $W=Y_1$  that it was an unbiased, but poor estimator. One notable feature of  $Y_1$  is that its variance is the same no matter what its sample size. It is reasonable to require that as the  $[n \to \infty, \quad \sigma^2 \to 0]$  sample size increases the estimation procedure improves. Example:  $\overline{Y}$  for estimation of population mean  $\mu$ 

$$s^2 = Var(\bar{Y}) = \frac{1}{n-1} \sum_{i=1}^{n} (Y_i - \bar{Y})^2$$

As  $n \to \infty$ ,  $s^2 \to 0$ .

## Consistency

Let  $W_n$  be an estimator of  $\theta$  based on a sample  $Y_1, Y_2, \ldots, Y_n$  of size n. Then  $W_n$  is a *consistent estimator* of  $\theta$  if for every  $\epsilon > 0$ 

$$P(|W_n - \theta| > \epsilon) \to 0$$
 as  $n \to \infty$ 

If  $W_n$  is not consistent for  $\theta$ , we say it is *inconsistent*. When  $W_n$  is consistent, we also say that  $\theta$  is the probability limit of  $W_n$ , written as

$$plim(W_n) = \theta$$

#### Interpretation:

• The distribution of  $W_n$  becomes more and more concentrated about  $\theta$ , which means that for larger sample sizes,  $W_n$  is less and less likely to vary far from  $\theta$ .

## Python Simulation to Show Consistency

- Go to Python, then come back.
- See the code file lln.py

## The Law of Large Numbers

Let  $Y_1, Y_2, \ldots, Y_n$  be independent, identically distributed random variables with mean  $\mu$ . Then

$$plim(\bar{Y}) = \mu$$

#### Interpretation:

• If we want to estimate the population average  $\mu$ , we can get arbitrarily close to  $\mu$  by choosing a sufficiently large sample.

## Properties of the Probability Limit

#### **Propery PLIM1:**

Let  $\theta$  be a parameter and define a new parameter  $\gamma = g(\theta)$ , for some continuous function  $g(\theta)$ . Suppose that  $plim(W_n) = \theta$ .

Define an estimator  $\gamma$  by  $G_n = g(W_n)$ . Then  $plim(G_n) = \gamma$ .

Often stated as:  $plim(g(W_n)) = g(plim(W_n))$  for a continuous function  $g(\theta)$ .

Example:  $g(\theta) = a + b\theta$ ,  $g(\theta) = \sigma^2$ ,  $g(\theta) = \frac{1}{\theta}$ ,  $g(\theta) = \sqrt{\theta}$ ,  $g(\theta) = \exp \theta$ .

## Properties of the Probability Limit

#### **Property PLIM2:**

If  $plim(T_n) = \alpha$  and  $plim(U_n) = \beta$  then

- $plim(T_n + U_n) = \alpha + \beta$
- $plim(T_nU_n) = \alpha\beta$
- $plim\left(\frac{T_n}{U_n}\right) = \frac{\alpha}{\beta}$ , for  $\beta \neq 0$

## An Example

Example: Let  $\{Y_1, Y_2, \dots, Y_n\}$  be a random sample of size n an annual earnings from the population of workers with a high school education with population mean  $\mu_Y$ .

Let  $\{Z_1, Z_2, \ldots, Z_n\}$  be a random sample of size n on annual earnings from the population of workers with a college education with population mean  $\mu_Z$ .

We wish to estimate the percentage difference in annual earnings between the two groups, which is  $\gamma=100\frac{(\mu_Z-\mu_Y)}{\mu_Y}$ , the percentage by which earnings for college grads differs from high school grads.

Because  $\bar{Y}$  is consistent for  $\mu_Y$ , and  $\bar{Z}_n$  is consistent for  $\mu_Z$  it follows that

$$G_n = 100 \frac{(\bar{Z}_n - \bar{Y}_n)}{\bar{Y}_n}$$

is a consistent estimator of  $\gamma$ .

## Asymptotic Normality

Let  $\{Z_n : n = 1, 2, ...\}$  be a sequence of random variables, such that for all numbers Z

$$P(Z_n \le z) \to \Phi(z)$$
 as  $n \to \infty$ 

in which  $\Phi(z)$  is the standard normal CDF.  $Z_n$  is said to have an asymptotic standard normal distribution. Asymptotic normality holds for large n, we have the approximation  $P(Z_n \leq z) \approx \Phi(z)$ . This means that probabilities concerning  $Z_n$  can be approximated by the standard normal probabilities.

### The Central Limit Theorm

Let  $\{Y_1, Y_2, \dots, Y_n\}$  be a random sample with mean  $\mu$  and variance  $\sigma^2$ . Then

$$Z_n = \frac{\bar{Y} - \mu}{\sigma / \sqrt{n}}$$

has an asymptotic standard normal distribution.

A *Confidence Interval* is a range of values, calculated from the sample observations, that are believed, with a particular probability, to contain the true parameter value.

## An Example of Confidence Interval Estimation

Example:

Suppose the population has a  $N(\mu, \sigma = 1)$  distribution and let  $\{Y_1, Y_2, \dots, Y_n\}$  be a random sample from this population (assume  $\sigma = 1$  known).

The sample average  $\overline{Y}$  has a normal distribution with mean  $\mu$  and variance  $\frac{1}{n}$ 

$$\bar{Y} \sim N\left(\mu, \frac{1}{n}\right)$$

Standardizing  $\bar{Y}$  which will give it a standard normal distribution

$$P\left(-1.96 < \frac{\bar{Y} - \mu}{1/\sqrt{n}} < 1.96\right) = 0.95$$

Tells us that the probability that the random interval  $[\bar{Y}-1.96/\sqrt{n},\bar{Y}+1.96/\sqrt{n}]$ , contains the population mean  $\mu$ 

This allows us to construct an interval estimate of  $\boldsymbol{\mu}$ 

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This is called a 95% confidence interval. This is denoted as  $\bar{y} \pm 1.96/\sqrt{n}$ .

Example: suppose 
$$n = 16$$
 and  $\bar{y} = 7.3$  then

$$7.3 \pm 1.96 / \sqrt{16} = 7.3 \pm 0.49 = [6.81, 7.79]$$

Interpretation: the random interval  $[\bar{Y}-1.96/\sqrt{n}, \quad \bar{Y}+1.96/\sqrt{n}]$  contains  $\mu$  with probability 0.95.

In other words, if we sampled from the population and calculated the random interval infinitely many times, it would contain the population parameter  $\mu$  95% of the time.

Note: It does not mean:

$$P(\bar{Y} - 1.96/\sqrt{n} \le \mu \le +1.96/\sqrt{n}) = 0.95$$
 because parameters  $\theta$  are *not* variables.

- Python simulation exercise to show confidence intervals.
- See Python code file interval.py

# The Confidence Interval for the Mean of a Normal Population

Assume  $X \sim \textit{N}(\mu, \sigma)$  and  $\sigma$  is known to be any value, the 95% CI is

## Confidence Interval for the Mean of a Normal Population

To allow for unknown  $\sigma$ , we must use an estimate. Let

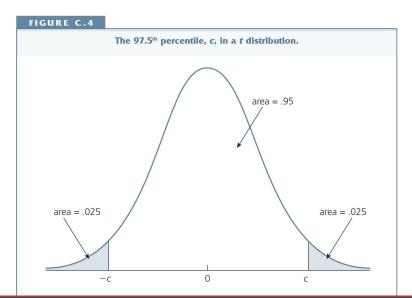
$$s = \left(\frac{1}{n-1}\sum_{i=1}^{n}(y_i - \bar{y})^2\right)^{1/2}$$

denote the sample standard deviation. Using s for  $\sigma$  does not preserve the 95% CI because we must use the sample to calculate s. In this case we will rely on the t distribution

$$\frac{\bar{Y}-\mu}{S/\sqrt{n}}\sim t_{n-1}$$

Where  $\overline{Y}$  is the sample average and s is the sample standard deviation of the random sample  $\{Y_1, Y_2, \ldots, Y_n\}$ . To construct a 95% CI let c denote the 97.5 percentile in the  $t_{n-1}$  distribution.

# Confidence Interval for the Mean of a Normal Population



# Confidence Interval for the Mean of a Normal Population

In other words:

$$P(-c < t_{n-1} < c) = 0.95$$

For a particular sample

$$[\bar{y}-c\frac{s}{\sqrt{n}}, \ \bar{y}+c\frac{s}{\sqrt{n}}]$$

## Confidence Interval for the Mean of a Normal Population: Example

Example: c is chosen from statistical tables for the  $t_{n-1}$  distribution. If n=20 then the df=20-1=19 Then we have c=2.093 and thus

Example: Consider job training grants on worker production. Let

- n = 20
- c = 2.093
- $\bar{y} = 1.15$
- $se(\bar{y}) = 0.54$

# Confidence Interval for the Mean of a Normal Population: Example

Then we have

$$CI_{0.95} = [-2.28, -0.02]$$

Interpretation: Zero is excluded. We conclude with 95% confidence that average change in scrap rates is not zero.

## A Simple Rule of Thumb

A simple rule of thumb for a 95% confidence interval is

- The t distribution approaches the Normal distribution as the degrees of freedom get large.
- In particular for  $\alpha=0.05$ ,  $C_{\frac{\alpha}{2}}\to 1.96$  as  $n\to\infty$ .
- An approximate 95% CI is

# Asymptotic Confidence Intervals for Non-Normal Populations