

Probability & Inference

- Result of tossing a coin is $\in \{H, T\}$

- Random variable $X \in \{1, 0\}$

— Recall the Bernoulli PMF

$$P(X=1) = \theta^x (1-\theta)^{(1-x)}$$

- Sample: $X = \{x^t\}_{t=1}^N$

$$\text{Estimation (MLE)}: \hat{\theta} = \frac{\# H}{\# \text{ Tosses}} = \frac{1}{N} \sum_{t=1}^N x^t$$

- Predict the next toss:

$$1-1 \text{ if } \hat{\theta} > \frac{1}{2}, \quad T \text{ otherwise}$$

NB: Notes based

Alpaydin cup 3

Bayesian Decision Theory

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Conjugate Bayes

$$\theta \sim \text{Beta}(a, b)$$

$$\theta | X \sim \text{Beta}(a^*, b^*)$$

$$a^* = a + N_1$$

$$b^* = b + N_0$$

where $N_1 = \# \text{ heads}$

$N_0 = \# \text{ tails}$
 $N_0 + N_1 = N \text{ trials}$

Classification

- Credit scoring: Inputs are income and savings
- Output: low-risk vs high-risk

• Input: $X = [x_1, x_2]^T$

Output: $C \in \{0, 1\}$

- Prediction

$$\begin{cases} C=1 & \text{if } P(C=1 | x_1, x_2) > 0.5 \\ C=0 & \text{otherwise} \end{cases}$$

choose/select

or

$$\begin{cases} C=1 & \text{if } P(C=1 | x_1, x_2) > P(C=0 | x_1, x_2) \\ C=0 & \text{otherwise} \end{cases}$$

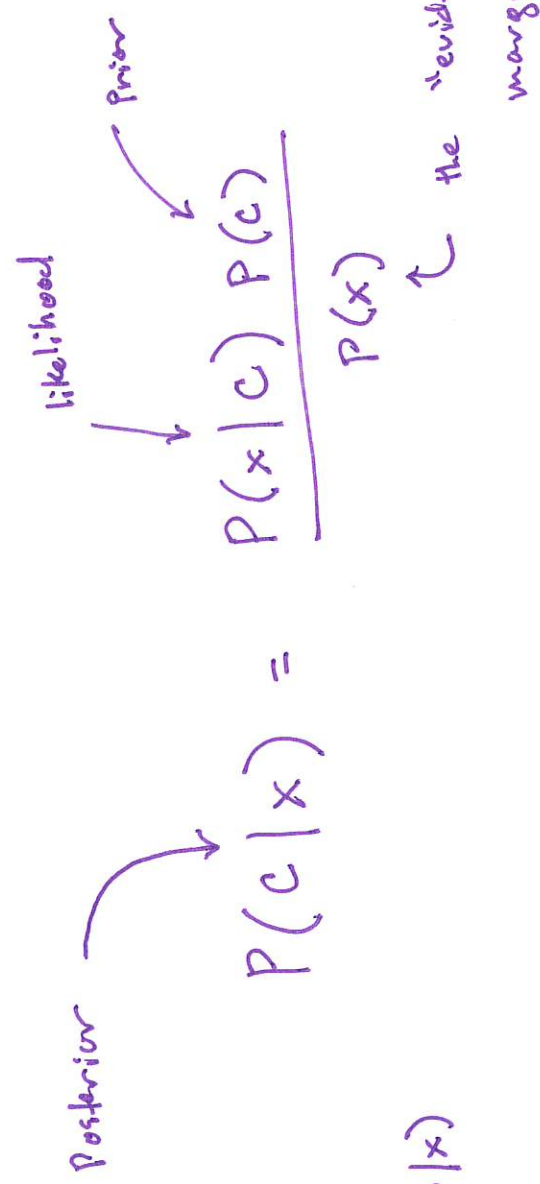
choose/select

Bayes' Rule

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NB:

$$P(c=0|x) + P(c=1|x) = 1$$



Decision
choose $C=1$ if

$$P(c=1|x) > P(c=0|x)$$

Prior prob. $P(c=1)$: prob. of high-risk customer

$$- P(c=0) + P(c=1) = 1$$

Class Likelihood: $P(x|c)$ prob. of event belonging to class C given the observable x

(Q: how ~~likely~~ likely are the data if the true state of the world is C)

- $P(x_1, x_2|c)$ = prob. that high-risk customer will have $[x_1, x_2]^T$

Evidence: $P(x)$: marginal prob. of $[x_1, x_2]^T$ $\left\{ \begin{array}{l} P(x) = P(x|c=1)P(c=1) + P(x|c=0)P(c=0) \end{array} \right.$

Bayes' Rule for $K > 2$ Classes

(mutually exclusive &
exhaustive)

posterior prob
↓

$$P(c_i | x) = \frac{P(x | c_i) P(c_i)}{P(x)}$$

$$= \frac{P(x | c_i) P(c_i)}{\sum_{k=1}^K P(x | c_k) P(c_k)}$$

Decision:

choose as the output the class c_i
that has the maximum posterior prob.

$$P(c_i | x)$$

Prior Probabilities : $P(c_i) \quad \forall i=1, \dots, K$

$$\sum_{i=1}^K P(c_i) = 1$$

Class Likelihood : $P(x | c_i)$

Losses and Risks

- Credit scoring decisions should be made so as to maximize gains / limit losses

- Classes C_1, \dots, C_K

- Let d_i = decision to assign C_i to the input $\forall 1 \leq i \leq K$

- Let λ_{ik} = the loss from assigning C_i to input that belongs to C_k (i.e. misclassification)

- Expected risk for d_i :

$$R(d_i | x) = \sum_{k=1}^K \lambda_{ik} P(C_k | x)$$

- Decision: choose d_i with minimal expected risk

$$R(d_i | x) = \min_k R(d_k | x)$$

0/1 Loss Case

- Correct decisions have 0 loss
- Incorrect decisions have 1 loss

$$\lambda_{ik} = \begin{cases} 0 & \text{if } i=k \\ 1 & \text{if } i \neq k \end{cases}$$

- The risk of taking action d_i is:

$$R(d_i | x) = \sum_{k=1}^K \lambda_{ik} P(c_k | x) = \sum_{k \neq i}^K \underbrace{P(c_k | x)}_{\text{via complement rule}} = 1 - P(c_i | x)$$

- Decision: to minimize risk, assign C_i to the most probable class

NB:

- all equal
- losses are symmetric



Losses & Risks: Reject

NB:

- reject option might require human analysis
- or another ML algorithm

• misclassification: very high cost

• Consider an extra option $(K+1)$ -st class "reject"

• Loss function:

$$\lambda_{ik} = \begin{cases} 0 & \text{if } i = k \\ \lambda & \text{if } i = k+1, \quad 0 < \lambda < 1 \\ 1 & \text{if } i \neq k \text{ and } i \neq k+1 \end{cases}$$

• Risk of reject:

$$R(d_{K+1}|x) = \sum_{k=1}^K \lambda P(c_k|x) = \lambda$$

• Risk of misclassification:

$$R(d_i|x) = \sum_{k \neq i} P(c_k|x) = \underbrace{1 - P(c_i|x)}_{\text{via complement rule (i.e. one minus "doing it correctly")}}$$

Decision:

- choose c_i $\min_{1 \leq i \leq K} R(d_i|x)$

output: $P(c_i|k) > P(c_k|x)$
 - c_i if $P(c_i|x) > 1-\lambda$
 - reject otherwise

Cases:

$\lambda = 0$

"always reject"
 (as good as correct classification)

$\lambda = 1$

"never reject"
 (as bad as incorrect classification)

NB:

$$\lambda \underbrace{\sum_{k=1}^K P(c_k|x)}_1$$

Discriminant functions

- Choose $g_i(x)$, $i=1, \dots, k$ s.t. output C_i if $g_i(x) = \max_k g_k(x)$

Bayes' Classifier:

$$g_i(x) = -R(d_i|x)$$

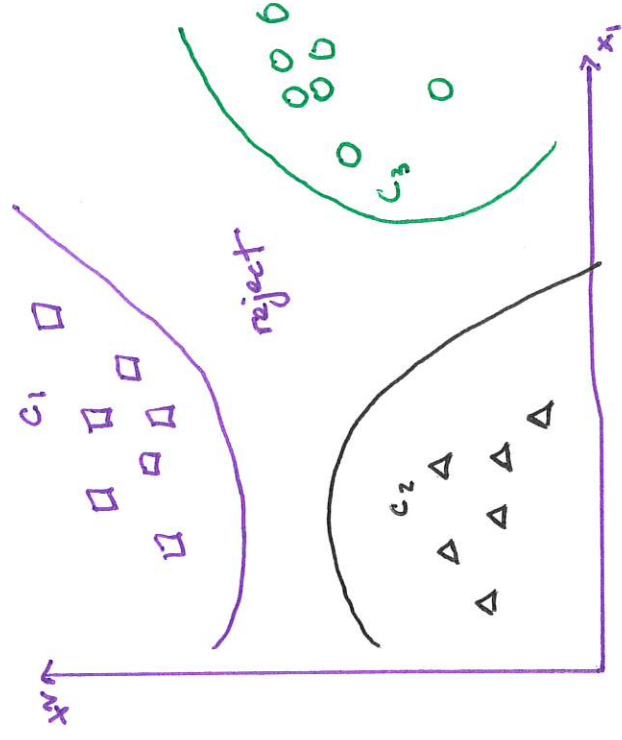
$$\text{-0/1 loss: } g_i(x) = P(C_i|x) = \frac{P(x|C_i)P(C_i)}{P(x)}$$

via Bayes' Rule

Equivalent way: $g_i(x) = P(x|C_i)P(C_i)$ (simpler/easier)

- k decision regions

$$R_i = \{x \mid g_i(x) = g_k(x)\}$$



$K=2$ classes

- Dichotomizer ($k=2$) vs Polychotomizer ($k>2$)

- $g(x) = g_1(x) - g_2(x)$

$$\text{choose } \begin{cases} c_1 & \text{if } g(x) > 0 \\ c_2 & \text{otherwise} \end{cases}$$

- Log odds:

$$\log \frac{P(c_1|x)}{P(c_2|x)} = \underbrace{\log P(c_1|x) - \log P(c_2|x)}_{\text{log difference in probabilities}}$$

Bayesian Thinking Revisited

Introduction to Bayes Factors

• Return to HIV testing

- H_1 : Patient does not have HIV
- H_2 : Patient does have HIV

• Actions/Decisions

- d_1 : Choose H_1
- d_2 : choose H_2

• $L(d)$ the loss associated w/ decision d_i $\forall i \in \{d_1, d_2\}$

- $d = d_1$

— Right: decide H_1 / and they don't $\Rightarrow L(d_1) = 0$
— Wrong: decide H_1 / and they do $\Rightarrow L(d_1) = \omega_1$

- $d = d_2$

— Right: decide H_2 / they do $\Rightarrow L(d_2) = 0$
— Wrong: decide H_2 / they don't $\Rightarrow L(d_2) = \omega_2$

NB: notes based

on chp. 3 Losses

and Decision making

An Introduction to Bayesian

Thinking

Losses

$$L(d_1) = \begin{cases} \emptyset & \text{if } d_1 \text{ correct} \\ \omega_1 = 1000 & \text{else} \end{cases}$$

$$L(d_2) = \begin{cases} \emptyset & \text{if } d_2 \text{ correct} \\ \omega_2 = 10 & \text{else} \end{cases}$$

H_1 : Patient ^{does not} have HIV
 H_2 : Patient has HIV

Posteriors

- (+) stands for a positive result from the ELISA
- $P(H_1 | +) \approx 0.88$ posterior prob. of NOT having HIV given a (+) ELISA result
- $P(H_2 | +) \approx 0.12$ Posterior prob. of Having HIV given the (+) ELISA result (from the complement)

Expected Losses

- $E[L(d_1)] = 0.88(\phi) + 0.12(1000) = 120$

- $E[L(d_2)] = 0.88(10) + 0.12(\phi) = 8.8$

- Since $E[L(d_2)] < E[L(d_1)] \Rightarrow$ decide the patient has HIV

NB:

- Decision highly influenced by losses assigned to d_1 and d_2

- If losses symmetric, say $w_1 = w_2 = 10$

- $E[L(d_1)] = 0.88(\phi) + 0.12(10) = 1.2$

- while $E[L(d_2)]$ would not change

- We would decide that Patient does NOT have HIV!

Bayes factors

- Continue with HIV testing example
- Prior odds = ratio of the prior probabilities of hypotheses

$$O[H_1 : H_2] = \frac{P(H_1)}{P(H_2)}$$

- Posterior odds = ratio of the two posterior probabilities of hypotheses

$$PO[H_1 : H_2] = \frac{P(H_1 | \text{Data})}{P(H_2 | \text{Data})}$$

Using Bayes' Rule, we find that

$$\begin{aligned} P(H_1 : H_2) &= \frac{P(H_1 | Data)}{P(H_2 | Data)} \\ &= \frac{P(data | H_1) P(H_1)}{P(data | H_2) P(H_2)} \cdot \frac{P(data)}{P(data)} \\ &= \frac{P(data | H_1) P(H_1)}{P(data | H_2) P(H_2)} \\ &= \underbrace{\frac{P(data | H_1)}{P(data | H_2)}}_{\text{Bayes Factor}} * \underbrace{\frac{P(H_1)}{P(H_2)}}_{\text{Prior odds}} \end{aligned}$$

So, we have that

$$P_0[H_1 : H_2] = BF[H_1 : H_2] * P[H_1 : H_2]$$

- BF quantifies the evidence of data arising from H_1 versus H_2

- In the discrete case: $BF[H_1 : H_2] = \frac{P(\text{data} | H_1)}{P(\text{data} | H_2)}$
(likelihood ratio)

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- In the continuous case

$$BF[H_1 : H_2] = \frac{\int p(\text{data} | \theta, H_1) d\theta}{\int p(\text{data} | \theta, H_2) d\theta}$$

NB: θ is the index of all possible models/parameters

- Continue w/ HIV case

H_1 : Patient does not have HIV

H_2 : Patient does have HIV

- Priors:

$$P(H_1) = .99852$$

$$P(H_2) = .00148$$

NB: from prevalence of HIV in the population

- Prior odds then is

$$O[H_1 : H_2] = \frac{P(H_1)}{P(H_2)} = \frac{.99852}{.00148} = 674.6757$$

- Posteriors

$$P(H_1 | +) = .8788551$$

$$P(H_2 | +) = .1211449$$

$$PO[H_1 : H_2] = \frac{.8788551}{.1211449} = 7.254578$$

- Bayes Factor

$$\begin{aligned} BF[H_1 : H_2] &= \frac{PO[H_1 : H_2]}{O[H_1 : H_2]} \\ &= \frac{.725457}{674.6757} = 0.00108 \\ &= \frac{P(+ | H_1)}{P(+ | H_2)} = \frac{.01}{.93} \approx .0108 \end{aligned}$$

- So now that we have calculated $BF[H_1 : H_2]$, how should we understand its meaning?

- Jeffrey (1961)

$BF[H_1 : H_2]$	Evidence against H_2
1 to 3	Not worth a bare mention
3 to 20	Positive
20 to 150	Strong
> 150	Very strong

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$2 * \log(BF[H_1 : H_2])$	Evidence against H_1
0 to 2	Not worth a bare mention
2 to 6	Positive
6 to 10	Strong
> 10	Very strong

- NB: notice that for the HN case it doesn't even appear on the scale.

- Let's reader so that it's $BF[H_2 : H_1] = \frac{1}{BF[H_1 : H_2]} = \frac{1}{.0108} = \underline{\underline{92.54259}} \approx 93$

- Hence the evidence against H_1 is ~~strong~~ "very strong"