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Linear Models for Regression

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Material and images in these slides are from (or adapted from): *C. Bishop, Pattern Recognition and Machine Learning, Springer,* 2006

The Regression problem

Predict the value of **continuous** *target* variables *t* given the value of a *D*-dimensional vector **x** of *input* variables.

Given N observations $\{x_n\}$, where n = 1, ..., N, together with corresponding target values $\{t_n\}$, the goal is to predict the value of t for a new value of x.

The Roadmap

We model the predictive distribution p(t | x) expressing our uncertainty about the value of t for each value of x.

- learn $p(t | \mathbf{x})$ by minimizing a loss function.
- common choice of loss function for real-valued variables: Sum of Squared Errors – consequence of maximizing likelihood under the assumption of a Gaussian noise distribution.
 - \rightarrow the optimal solution is given by the conditional expectation of t (from decision theory)

Linear Models: linear functions of the adjustable parameters

IMPORTANT: can be nonlinear with respect to the input variables (e.g. the polynomial we saw earlier)

- simplest form: models are also linear functions of the input variables.
- *basis functions*: nonlinear functions of the input variables, of which we take linear combinations

Regression by linear combination of basis functions

• Simplest linear model for regression (often just called *linear regression*):

$$y(\mathbf{x}, \mathbf{w}) = w_0 + w_1 x_1 + \dots + w_D x_D = w_0 + \sum_{i=1}^{D} w_i x_i$$

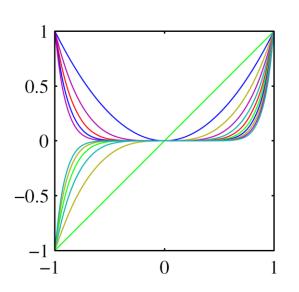
• linear combinations of fixed nonlinear Basis Functions, $\phi_j(\mathbf{x})$:

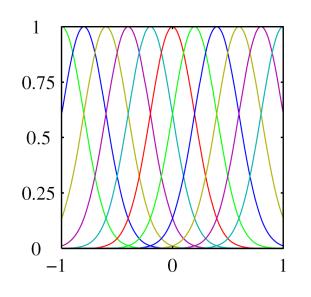
$$y(\mathbf{x}, \mathbf{w}) = w_0 + \sum_{j=1}^{M-1} w_j \phi_j(\mathbf{x})$$

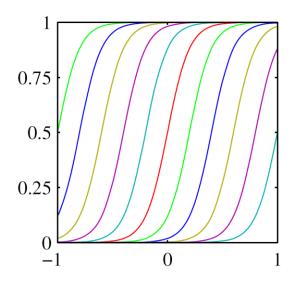
• w_0 , called **bias** parameter, can be included in the sum, by defining one extra basis function $\phi_0(\mathbf{x}) = 1$

$$y(\mathbf{x}, \mathbf{w}) = \sum_{j=0}^{M-1} w_j \phi_j(\mathbf{x}) = \mathbf{w}^{\mathrm{T}} \boldsymbol{\phi}(\mathbf{x})$$

Classic basis functions







Polynomial

$$\phi_j(x) = x^j$$

Gaussian

$$\phi_j(x) = \exp\left\{-\frac{(x-\mu_j)^2}{2s^2}\right\}$$

Sigmoidal

$$\phi_j(x) = \sigma\left(\frac{x - \mu_j}{s}\right)$$

$$\sigma(a) = \frac{1}{1 + \exp(-a)}$$

 μ_i control the locations of the basis functions s controls the scale of the basis function

Maximum likelihood and least squares

(it's similar to what we did earlier...)

Assume *t* is given by:

$$t = y(\mathbf{x}, \mathbf{w}) + \epsilon$$
 ϵ gaussian random variable

$$p(t|\mathbf{x}, \mathbf{w}, \beta) = \mathcal{N}(t|y(\mathbf{x}, \mathbf{w}), \beta^{-1})$$

Given data set of *N* inputs **X**

Likelihood function (a function of w and β)

$$p(\mathbf{t}|\mathbf{X}, \mathbf{w}, \beta) = \prod_{n=1}^{N} \mathcal{N}(t_n|\mathbf{w}^{\mathrm{T}} \boldsymbol{\phi}(\mathbf{x}_n), \beta^{-1})$$

$$\ln p(\mathbf{t}|\mathbf{w}, \beta) = \sum_{n=1}^{N} \ln \mathcal{N}(t_n|\mathbf{w}^{\mathrm{T}} \boldsymbol{\phi}(\mathbf{x}_n), \beta^{-1})$$

$$= \frac{N}{2} \ln \beta - \frac{N}{2} \ln(2\pi) - \beta E_D(\mathbf{w})$$

where: $E_D(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^{N} \{t_n - \mathbf{w}^T \phi(\mathbf{x}_n)\}^2$ is the sum-of-squares error function

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Gradient of the log likelihood function

$$\nabla \ln p(\mathbf{t}|\mathbf{w},\beta) = \sum_{n=1}^{N} \left\{ t_n - \mathbf{w}^{\mathrm{T}} \boldsymbol{\phi}(\mathbf{x}_n) \right\} \boldsymbol{\phi}(\mathbf{x}_n)^{\mathrm{T}}$$

Setting the gradient to 0:

$$0 = \sum_{n=1}^{N} t_n \boldsymbol{\phi}(\mathbf{x}_n)^{\mathrm{T}} - \mathbf{w}^{\mathrm{T}} \left(\sum_{n=1}^{N} \boldsymbol{\phi}(\mathbf{x}_n) \boldsymbol{\phi}(\mathbf{x}_n)^{\mathrm{T}} \right)$$

Solving for **w** we obtain

$$\mathbf{w}_{\mathrm{ML}} = \left(\mathbf{\Phi}^{\mathrm{T}}\mathbf{\Phi}
ight)^{-1}\mathbf{\Phi}^{\mathrm{T}}\mathbf{t}$$

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 $\phi_{0}(\mathbf{x}_{1}) \quad \phi_{1}(\mathbf{x}_{1}) \quad \cdots \quad \phi_{M-1}(\mathbf{x}_{1})$
 $\phi_{0}(\mathbf{x}_{2}) \quad \phi_{1}(\mathbf{x}_{2}) \quad \cdots \quad \phi_{M-1}(\mathbf{x}_{2})$
 $\vdots \quad \vdots \quad \ddots \quad \vdots$
 $\phi_{0}(\mathbf{x}_{N}) \quad \phi_{1}(\mathbf{x}_{N}) \quad \cdots \quad \phi_{M-1}(\mathbf{x}_{N})$
 $\mathbf{\Phi}_{nj} = \mathbf{\varphi}_{j}(\mathbf{x}_{n})$

Moore-Penrose pseudo-inverse of Φ

the normal equations for

$$\Phi_{nj} = \varphi_j(\mathbf{x}_n)$$

Important connection to help us understand...

The Moore-Penrose pseudo-inverse of the matrix Φ ,

Φ[†] can be regarded as a generalization of the notion of matrix inverse to nonsquare matrices

$$oldsymbol{\Phi}^\dagger \equiv \left(oldsymbol{\Phi}^\mathrm{T} oldsymbol{\Phi}
ight)^{-1} oldsymbol{\Phi}^\mathrm{T}$$

If Φ is square and invertible then using the property $(\mathbf{A}\mathbf{B})^{-1} = \mathbf{B}^{-1}\mathbf{A}^{-1}$ we see that $\Phi^{+} \equiv \Phi^{-1}$.

Sequential/online learning

Batch techniques can be computationally costly for large datasets.

Sequential algorithms: data points are considered one at a time. Model parameters updated after each presentation.

Stochastic gradient descent/ sequential gradient descent

$$\mathbf{w}^{(\tau+1)} = \mathbf{w}^{(\tau)} - \eta \nabla E_n$$

 E_n is the error for point n

For sum-of-squares error function:

$$\mathbf{w}^{(\tau+1)} = \mathbf{w}^{(\tau)} + \eta (t_n - \mathbf{w}^{(\tau)\mathrm{T}} \boldsymbol{\phi}_n) \boldsymbol{\phi}_n$$
 Least-Mean-Squares or the LMS algorithm

$$\phi_n = \phi(\mathbf{x}_n)$$

Regularized Least Squares

<u>Idea</u>: adding a regularization term to an error function to control over-fitting

$$E_D(\mathbf{w}) + \lambda E_W(\mathbf{w})$$

1. Weight decay (in machine learning)

$$E_W(\mathbf{w}) = \frac{1}{2} \mathbf{w}^{\mathrm{T}} \mathbf{w}$$

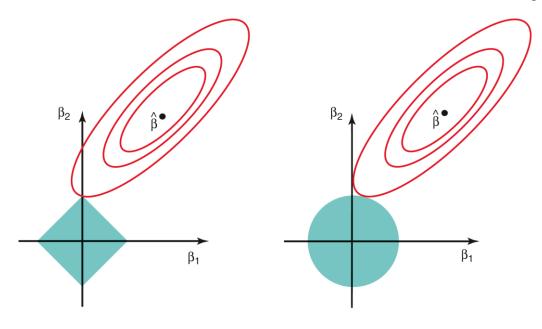
$$\frac{1}{2} \sum_{n=1}^{N} \{t_n - \mathbf{w}^{\mathrm{T}} \boldsymbol{\phi}(\mathbf{x}_n)\}^2 + \frac{\lambda}{2} \mathbf{w}^{\mathrm{T}} \mathbf{w}$$
 (ridge regression in statistics)

$$\mathbf{w} = (\lambda \mathbf{I} + \mathbf{\Phi}^{\mathrm{T}} \mathbf{\Phi})^{-1} \mathbf{\Phi}^{\mathrm{T}} \mathbf{t} \quad \begin{array}{l} \text{normal equations} \\ \text{with weight decay} \end{array}$$

2. Lasso (Least Absolute Shrinkage and Selection Operator)

$$E_W(\mathbf{w}) = \frac{\lambda}{2} \sum_{j=1}^{M} |w_j|$$

$$\frac{1}{2}\sum_{n=1}^{N}\{t_n-\mathbf{w}^{\mathrm{T}}\boldsymbol{\phi}(\mathbf{x}_n)\}^2+\frac{\lambda}{2}\sum_{j=1}^{M}|w_j| \qquad \textit{Lasso regression}$$



some of the coefficients w_i become 0, leading to a sparse model

$$\frac{1}{2}\sum_{n=1}^{N}\{t_n-\mathbf{w}^{\mathrm{T}}\boldsymbol{\phi}(\mathbf{x}_n)\}^2+\frac{\lambda}{2}\sum_{j=1}^{M}|w_j|^q \qquad \textit{General regularization form}$$

3. Elastic Net

$$E_W(\mathbf{w}) = \frac{\lambda_1}{2} \sum_{j=1}^M |w_j| + \frac{\lambda_2}{2} \mathbf{w}^{\mathrm{T}} \mathbf{w}$$

combines Lasso and Ridge regularization

Linear regression for multiple outputs

Predict *K* > 1 target variables – target vector **t**

The generalization is straightforward

$$\mathbf{y}(\mathbf{x}, \mathbf{w}) = \mathbf{W}^{\mathrm{T}} \boldsymbol{\phi}(\mathbf{x})$$
 $p(\mathbf{t}|\mathbf{x}, \mathbf{W}, eta) = \mathcal{N}(\mathbf{t}|\mathbf{W}^{\mathrm{T}} \boldsymbol{\phi}(\mathbf{x}), eta^{-1}\mathbf{I})$
 $\ln p(\mathbf{T}|\mathbf{X}, \mathbf{W}, eta) = \sum_{n=1}^{N} \ln \mathcal{N}(\mathbf{t}_{n}|\mathbf{W}^{\mathrm{T}} \boldsymbol{\phi}(\mathbf{x}_{n}), eta^{-1}\mathbf{I})$
 $= \frac{NK}{2} \ln \left(\frac{\beta}{2\pi}\right) - \frac{\beta}{2} \sum_{n=1}^{N} \left\|\mathbf{t}_{n} - \mathbf{W}^{\mathrm{T}} \boldsymbol{\phi}(\mathbf{x}_{n})\right\|^{2}$
 $\mathbf{W}_{\mathrm{ML}} = \left(\mathbf{\Phi}^{\mathrm{T}} \mathbf{\Phi}\right)^{-1} \mathbf{\Phi}^{\mathrm{T}}\mathbf{T}$
 $\mathbf{w}_{k} = \left(\mathbf{\Phi}^{\mathrm{T}} \mathbf{\Phi}\right)^{-1} \mathbf{\Phi}^{\mathrm{T}}\mathbf{t}_{k} = \mathbf{\Phi}^{\dagger} \mathbf{t}_{k}$

$$\mathbf{W}_{\mathrm{ML}} = \left(\mathbf{\Phi}^{\mathrm{T}}\mathbf{\Phi}
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