Linear Models for Classification

Alberto Paccanaro

EMAp - FGV

www.paccanarolab.org

Material and images in these slides are from (or adapted from): *C. Bishop, Pattern Recognition and Machine Learning, Springer,* 2006

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The Classification problem

Take an input vector \mathbf{x} and to assign it to one of K discrete classes C_k where $k = 1, \ldots, K$.

Classes are disjoint, input space is divided into *decision regions*

Boundaries: decision boundaries (decision surfaces).

Linear models: decision boundaries are linear functions

(D-1)-dimensional hyperplanes in the D dimensional input space.

Linearly separable: data sets whose classes can be separated exactly by linear decision surfaces

Representing target values

Two-classes (K=2)

- one variable, binary representation $t \in \{0, 1\}$ t = 1 represents class C_1 and t = 0 represents class C_2 .
- value of t is the probability that the class is C_1 .

K > 2 classes

• it is convenient to use a 1-of-K coding scheme

$$\mathbf{t} = (0, 1, 0, 0, 0)^{\mathrm{T}}$$

The general model

$$y(\mathbf{x}) = f\left(\mathbf{w}^{\mathrm{T}}\mathbf{x} + w_0\right)$$

To predict class labels we transform the linear function of w using amonlinear activation function $f(\cdot)$

- The model is not linear in the parameters (remember that linear regression was linear in the parameters)
- The decision surfaces are <u>linear functions of x</u>, even if the function $f(\cdot)$ is nonlinear (correspond to $\mathbf{w}^T\mathbf{x} + w_0 = \text{constant}$)

Generalized linear models

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As we discussed earlier, there are 3 approaches

- 1. Discriminant functions
- 2. Model $p(\mathbf{x} | C_k)$ and $P(C_k)$ and then use Bayes theorem to infer $p(C_k | \mathbf{x})$ (the *generative approach*)
- 3. Model $p(C_k|\mathbf{x})$ directly

1. Discriminant Functions

A discriminant is a function that takes an input vector x and assigns it to one of K classes, denoted C_k – no probabilities here \odot

... lets begin with some geometry...

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A. Two Classes (K = 2)

Let us not consider the activation function for now.

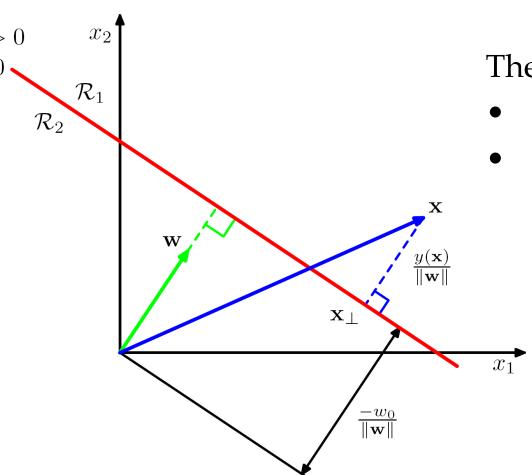
The simplest linear discriminant:

$$y(\mathbf{x}) = \mathbf{w}^{\mathrm{T}}\mathbf{x} + w_0$$
 w: weight vector w_0 : bias

x assigned to C_1 if $y(\mathbf{x}) \ge 0$, C_2 otherwise.

$$y(\mathbf{x}) = \widetilde{\mathbf{w}}^{\mathrm{T}} \widetilde{\mathbf{x}}$$
 including the bias

Decision boundary: y(x) = 0, corresponds to a **(D–1)-dim**. **hyperplane** within the **D-dim**. **input space**.



The decision surface is:

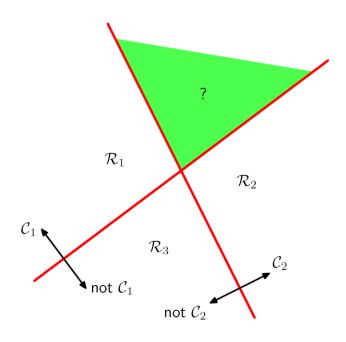
- perpendicular to w
- its displacement from the origin is controlled by the bias parameter w_0 .

• The signed orthogonal distance of a general point \mathbf{x} from the decision surface is given by $\frac{y(\mathbf{x})}{\|\mathbf{w}\|}$

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B. Multiple classes (K > 2)

Combinations of 2-class discriminants do not work!



one-versus-rest

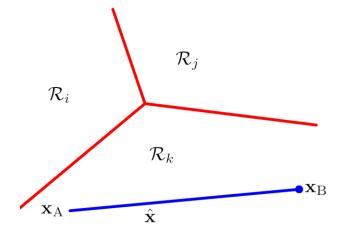
Possible solution: build a single K-class discriminant comprising K linear functions:

$$y_k(\mathbf{x}) = \mathbf{w}_k^{\mathrm{T}} \mathbf{x} + w_{k0}$$

Assign x to class C_k if $y_k(\mathbf{x}) > y_j(\mathbf{x})$ for all $j \neq k$.

• There is no really good solution for this issue... (see Bishop 7.1.3 Multiclass SVMs, page 338)

• The decision regions of such a discriminant are always connected and convex.



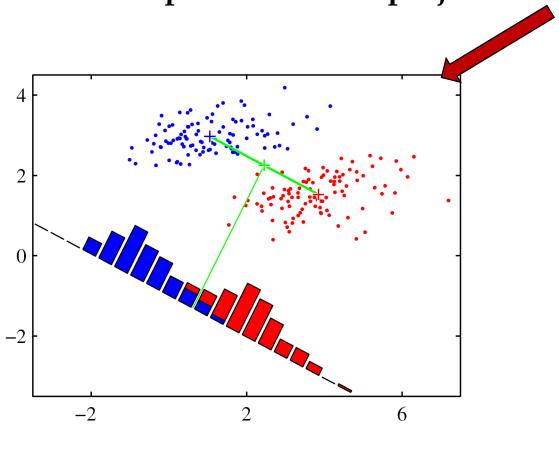
Fisher's Linear Discriminant (2 classes)

- 1) Project the points onto one dimension: $y = \mathbf{w}^T \mathbf{x}$
- 2) If the value (scalar) is larger than w_0 , then C_1 , else C_2 .

Note that this is equivalent to $y(\mathbf{x}) = \mathbf{w}^{\mathrm{T}}\mathbf{x} + w_0$

Question: can I "adjust" the projection so that I have the points that belong to the 2 different classes <u>as separated</u> <u>as possible</u>?

The simplest measure of the separation of the classes is the **separation of the projected class means**.



We need to maximize a large separation between the projected class means while also giving a **small variance within each class**

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Two Classes, N_1 points $\in C_1$, N_2 points $\in C_2$

$$\mathbf{m}_1 = \frac{1}{N_1} \sum_{n \in \mathcal{C}_1} \mathbf{x}_n$$
 $\mathbf{m}_2 = \frac{1}{N_2} \sum_{n \in \mathcal{C}_2} \mathbf{x}_n$ $m_k = \mathbf{w}^{\mathrm{T}} \mathbf{m}_k$

To maximize the separation between the class means, maximize:

$$m_2 - m_1 = \mathbf{w}^{\mathrm{T}}(\mathbf{m}_2 - \mathbf{m}_1)$$

Within variance of the projected data, for class k (here $y_n = \mathbf{w}^T \mathbf{x}_n$).

$$s_k^2 = \sum_{n \in \mathcal{C}_k} (y_n - m_k)^2$$

The Fisher criterion:

$$J(\mathbf{w}) = \frac{(m_2 - m_1)^2}{s_1^2 + s_2^2}$$

Ratio of the between-class variance to the within-class variance

Making the dependence on w explicit:

$$J(\mathbf{w}) = \frac{\mathbf{w}^{\mathrm{T}} \mathbf{S}_{\mathrm{B}} \mathbf{w}}{\mathbf{w}^{\mathrm{T}} \mathbf{S}_{\mathrm{W}} \mathbf{w}}$$

$$\mathbf{S}_{\mathrm{B}} = (\mathbf{m}_2 - \mathbf{m}_1)(\mathbf{m}_2 - \mathbf{m}_1)^{\mathrm{T}}$$
 between-class covariance matrix

$$\mathbf{S}_{\mathrm{W}} = \sum_{n \in \mathcal{C}_{1}} (\mathbf{x}_{n} - \mathbf{m}_{1})(\mathbf{x}_{n} - \mathbf{m}_{1})^{\mathrm{T}} + \sum_{n \in \mathcal{C}_{2}} (\mathbf{x}_{n} - \mathbf{m}_{2})(\mathbf{x}_{n} - \mathbf{m}_{2})^{\mathrm{T}}$$
within-class covariance matrix

Taking the gradient of J(w), it is maximized when:

$$(\mathbf{w}^T\mathbf{S}_B\mathbf{w})\mathbf{S}_W\mathbf{w} = (\mathbf{w}^T\mathbf{S}_W\mathbf{w})\mathbf{S}_B\mathbf{w}$$

We only care about the direction, hence:

$$\mathbf{w} \propto \mathbf{S}_{\mathrm{W}}^{-1}(\mathbf{m}_2 - \mathbf{m}_1)$$
 Fisher's Linear Discriminant

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Note: when the data is "spherical", S_w *is proportional to the identity,* w *is proportional to the difference between the means* ... \odot

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... don't forget ...

Now we have the <u>direction for the projection</u>, <u>but it is</u> <u>not a classification yet</u>

We are left with determining a threshold w_0 for the classification...

One option:

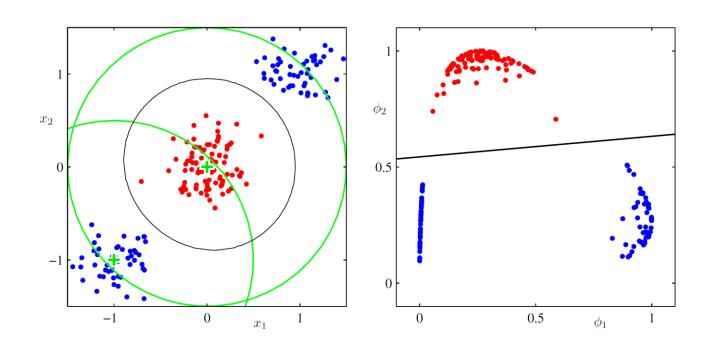
- a) model the class conditional densities $p(y \mid C_k)$ (e.g. using Gaussian distributions, learn the parameters of the gaussians by maximum likelihood)
- b) Calculate the priors $p(C_k)$ from the data
- c) Use Bayes theorem to calculate w_0

This can be generalized to <u>multiple classes</u> but we skip it because the maximization of the criteria is a bit involved.

Note on fixed basis functions

We can first apply a nonlinear transformation of the inputs $\phi(x)$ – *just as we did in linear regression*.

The classification problem can become easier



The Perceptron

A two-class model:

$$y(\mathbf{x}) = f\left(\mathbf{w}^{\mathrm{T}}\boldsymbol{\phi}(\mathbf{x})\right)$$

nonlinear transformation of \mathbf{x} to get $\phi(\mathbf{x})$ $\phi(\mathbf{x})$ used for a generalized linear model Note: $\phi_0(\mathbf{x}) = 1$ (bias)

Nonlinear activation function (step function):

$$f(a) = \begin{cases} +1, & a \geqslant 0 \\ -1, & a < 0 \end{cases}$$

Target values:

$$t = +1$$
 for class C_1
 $t = -1$ for class C_2

We want w such that:

when
$$\mathbf{x}_n \in C_1$$
 $\mathbf{w}^{\mathrm{T}} \phi(\mathbf{x}_n) > 0$
when $\mathbf{x}_n \in C_2$ $\mathbf{w}^{\mathrm{T}} \phi(\mathbf{x}_n) < 0$

$$\mathbf{w}^{\mathrm{T}}\boldsymbol{\phi}(\mathbf{x}_n)t_n > 0$$

Perceptron criterion: no error for patterns that are correctly classified. For a misclassified pattern minimize $-\mathbf{w}^{\mathrm{T}}\phi(\mathbf{x}_n)t_n$

$$E_{\mathrm{P}}(\mathbf{w}) = -\sum_{n \in \mathcal{M}} \mathbf{w}^{\mathrm{T}} \boldsymbol{\phi}_n t_n$$

where M is the set of misclassified patterns

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Applying stochastic gradient descent:

$$\mathbf{w}^{(\tau+1)} = \mathbf{w}^{(\tau)} - \eta \nabla E_{P}(\mathbf{w}) = \mathbf{w}^{(\tau)} + \eta \phi_n t_n$$

Interpretation: we cycle through the training patterns in turn, and for each pattern x_n :

- If the pattern is correctly classified, do nothing
- If it is incorrectly classified:
 - for class C_1 add the vector $\phi(\mathbf{x}_n)$ onto \mathbf{w}
 - for class C_2 we subtract the vector $\phi(\mathbf{x}_n)$ from \mathbf{w}

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Important

- learning rule not guaranteed to reduce the total error function at each stage
- **perceptron convergence theorem**: if the training data set is linearly separable the perceptron will find an exact solution in a finite number of steps
- Solution not unique, and will depend on the initialization of the parameters and on the order of presentation of the data points.
- For data sets that are not linearly separable, the algorithm will never converge.
- Does not generalize readily to K > 2 classes.
- Limited in what it can learn: it is based on linear combinations of fixed basis functions.

2. Probabilistic Generative Models

Models with linear decision boundaries arise from simple assumptions about the distribution of the data...

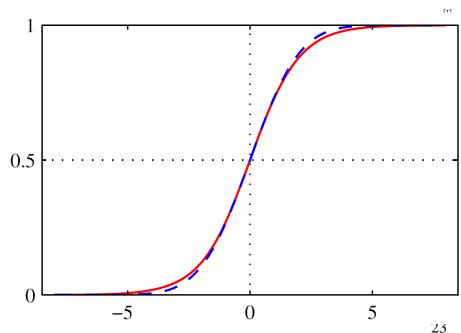
Two classes

$$p(C_1|\mathbf{x}) = \frac{p(\mathbf{x}|C_1)p(C_1)}{p(\mathbf{x}|C_1)p(C_1) + p(\mathbf{x}|C_2)p(C_2)}$$

$$= \frac{1}{1 + \exp(-a)} = \sigma(a)$$
sigmoid function

where we define:

$$a = \ln \frac{p(\mathbf{x}|\mathcal{C}_1)p(\mathcal{C}_1)}{p(\mathbf{x}|\mathcal{C}_2)p(\mathcal{C}_2)}$$



K > 2 classes

$$p(C_k|\mathbf{x}) = \frac{p(\mathbf{x}|C_k)p(C_k)}{\sum_j p(\mathbf{x}|C_j)p(C_j)}$$
$$= \frac{\exp(a_k)}{\sum_j \exp(a_j)} \quad softmax function$$

where we define:

$$a_k = \ln p(\mathbf{x}|\mathcal{C}_k)p(\mathcal{C}_k)$$

IMPORTANT

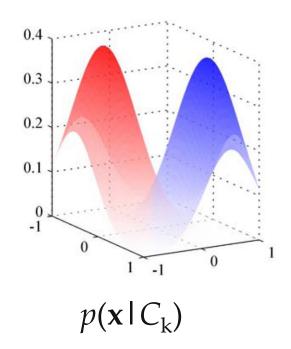
The use of the sigmoid/softmax activation functions allows the outputs to be interpreted as posterior probabilities of class membership.

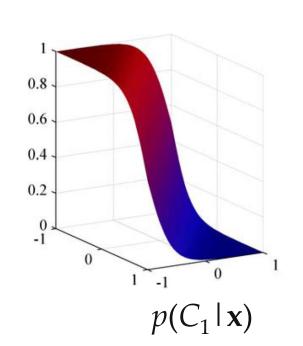
QUESTION

How do the form for the class conditional densities $p(x \mid C_k)$ affects the form of $p(C_k \mid x)$?

2 classes, continuous inputs, Gaussian class-conditional densities, same covariance

We obtain a linear function of x in the argument of the logistic sigmoid.





K classes, continuous inputs, Gaussian class-conditional densities, same covariance

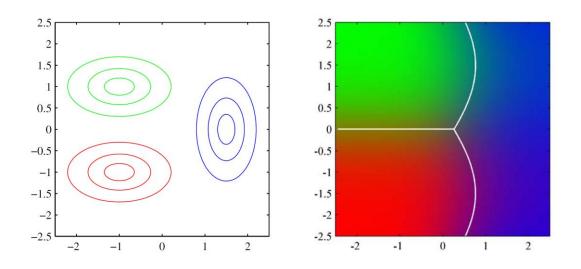
The $a_k(\mathbf{x})$ are linear functions of \mathbf{x}

As before, decision boundaries will be defined by linear functions of \mathbf{x} .

Again we have a generalized linear model.

K classes, continuous inputs, Gaussian class-conditional densities, <u>different covariance</u>

we obtain *quadratic* functions of **x**, giving rise to a quadratic discriminant.



These results can be extended to the class-conditional densities of the exponential family of distributions.

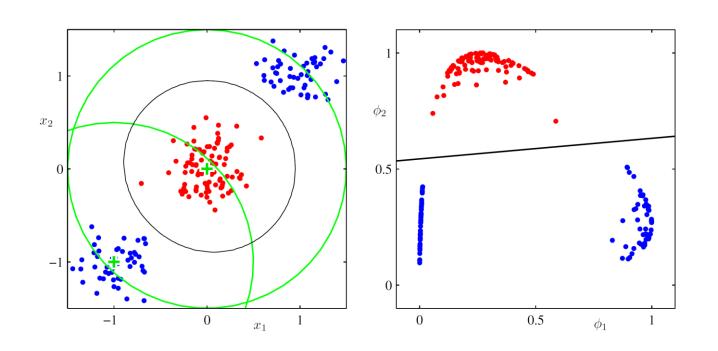
3. Model $p(C_k|x)$ directly – Probabilistic Discriminative Models

The posterior probability of class C_1 can be written as a logistic sigmoid (softmax for multiclass) acting on a linear function of \mathbf{x} .

Idea: use the functional form of the generalized linear model explicitly and determine its parameters directly by using maximum likelihood. That is, we maximize a likelihood function defined through p(Ck|x) (discriminative training).

Again... on fixed basis functions

All these algorithms are equally applicable if we first apply a nonlinear transformation of the inputs $\phi(x)$ Decision boundaries will be linear in the feature space $\phi(x)$ and nonlinear in the original x space



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Logistic Regression (2 classes)

$$p(C_1|\boldsymbol{\phi}) = y(\boldsymbol{\phi}) = \sigma\left(\mathbf{w}^{\mathrm{T}}\boldsymbol{\phi}\right)$$
$$p(C_2|\boldsymbol{\phi}) = 1 - p(C_1|\boldsymbol{\phi})$$

A logistic sigmoid acting on a linear function of the feature vector ϕ

In the terminology of statistics, this is known as logistic regression (a.k.a. Logistic Discriminant, Logistic Discrimination)

This is a model for classification not regression © Only M parameters.

Let us use maximum likelihood to determine the parameters w of the model.

Dataset of N datapoints:
$$\{\phi_n, t_n\}$$
 $\phi_n = \phi(\mathbf{x}_n)$ $t_n \in \{0, 1\}$

Likelihood function:

$$p(\mathbf{t}|\mathbf{w}) = \prod_{n=1}^{N} y_n^{t_n} \left\{ 1 - y_n \right\}^{1 - t_n} \qquad \mathbf{t} = (t_1, \dots, t_N)^{\mathrm{T}}$$
$$y_n = p(\mathcal{C}_1 | \boldsymbol{\phi}_n)$$

Take the negative logarithm:

$$E(\mathbf{w}) = -\ln p(\mathbf{t}|\mathbf{w}) = -\sum_{n=1}^{N} \left\{ t_n \ln y_n + (1 - t_n) \ln(1 - y_n) \right\}$$
Cross-Entropy

$$y_n = \sigma(a_n) \quad a_n = \mathbf{w}^{\mathrm{T}} \boldsymbol{\phi}_n$$

error function

$$\nabla E(\mathbf{w}) = \sum_{n=1}^{N} (y_n - t_n) \phi_n$$

$$y_n = \sigma(a_n)$$

$$a_n = \mathbf{w}^{\mathrm{T}} \phi_n$$

The contribution to the gradient from data point n is given by the 'error' $(y_n - t_n)$ between the target value and the prediction of the model, times the input for point n.

Remember... this is the same form as the gradient of the sum-of-squares error function for the linear regression model....

$$E_D(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^{N} \{t_n - \mathbf{w}^{\mathrm{T}} \boldsymbol{\phi}(\mathbf{x}_n)\}^2$$
$$\nabla \ln p(\mathbf{t}|\mathbf{w}, \beta) = \sum_{n=1}^{N} \{t_n - \mathbf{w}^{\mathrm{T}} \boldsymbol{\phi}(\mathbf{x}_n)\} \boldsymbol{\phi}(\mathbf{x}_n)^{\mathrm{T}}$$

but there is no closed form this time... why?

We use a sequential procedure where patterns are presented one at a time and weight vectors is updated by stochastic gradient descent

IMPORTANT

Cross-entropy error function arises from maximizing likelihood of the data in binary classification problems.

... in a similar way as...

Sum-of-squares error function arises from maximizing likelihood of the data under a Gaussian noise assumption.

Iterative Reweighted Least Squares (IRLS)

The cross-entropy error function is concave → unique minimum.

The Newton-Raphson update:

$$\mathbf{w}^{\text{(new)}} = \mathbf{w}^{\text{(old)}} - \mathbf{H}^{-1} \nabla E(\mathbf{w})$$

Applying it to the cross-entropy error function for the logistic regression model:

$$\nabla E(\mathbf{w}) = \sum_{n=1}^{N} (y_n - t_n) \boldsymbol{\phi}_n = \boldsymbol{\Phi}^{\mathrm{T}}(\mathbf{y} - \mathbf{t})$$

$$\mathbf{H} = \nabla \nabla E(\mathbf{w}) = \sum_{n=1}^{N} y_n (1 - y_n) \boldsymbol{\phi}_n \boldsymbol{\phi}_n^{\mathrm{T}} = \boldsymbol{\Phi}^{\mathrm{T}} \mathbf{R} \boldsymbol{\Phi}$$

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where R is an $N \times N$ diagonal matrix R with elements

$$R_{nn} = y_n(1 - y_n)$$

Iterative algorithm, with update formula:

$$\mathbf{w}^{(\mathrm{new})} = (\mathbf{\Phi}^{\mathrm{T}}\mathbf{R}\mathbf{\Phi})^{-1}\mathbf{\Phi}^{\mathrm{T}}\mathbf{R}\mathbf{z}$$

where
$$\mathbf{z} = \mathbf{\Phi} \mathbf{w}^{(\text{old})} - \mathbf{R}^{-1} (\mathbf{y} - \mathbf{t})$$

Multiclass Logistic Regression

$$p(C_k|\boldsymbol{\phi}) = y_k(\boldsymbol{\phi}) = \frac{\exp(a_k)}{\sum_j \exp(a_j)} \qquad a_k = \mathbf{w}_k^{\mathrm{T}} \boldsymbol{\phi}$$

$$E(\mathbf{w}_1, \dots, \mathbf{w}_K) = -\ln p(\mathbf{T}|\mathbf{w}_1, \dots, \mathbf{w}_K) = -\sum_{n=1}^{\infty} \sum_{k=1}^{\infty} t_{nk} \ln y_{nk}$$

Multiclass cross-entropy error function

$$abla_{\mathbf{w}_j} E(\mathbf{w}_1, \dots, \mathbf{w}_K) = \sum_{n=1}^N \left(y_{nj} - t_{nj}\right) \boldsymbol{\phi}_n$$
Gradient for iterative updates

Extension of IRLS to multiclass is also possible.

IMPORTANT

The derivative of the log likelihood function for a linear regression model takes the form of

'error' $(y_n - t_n)$ times the feature vector ϕ_n

... similarly...

for logistic sigmoid activation function and crossentropy error function &

for softmax activation function and multiclass cross-entropy error function

... we again obtain this same simple form.