## Ec2147, Spring 2021

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Problem Set 1: part (a)

Due: March 16th, 2021

Problem sets are due by 5PM EST. You may work in groups (indeed I encourage you to all work together), but each student should turn in their own write-up.

Helpful reading: Newey (1990), Brown and Newey (1998) and Graham (2011). For an alternative, and highly elegant approach, see Chamberlain (1987).

Let Z = (X', Y)' be a vector of modeling variables and  $\{Z_i\}_{i=1}^{\infty}$  an independent and identically distribution random sequence drawn from unknown distribution  $F_0$ . The sole prior restriction on  $F_0$  is that, for some known function  $\psi(Z, \beta)$  and unknown parameter  $\beta_0 \in \mathbb{B} \subset \mathbb{R}^K$ 

$$\mathbb{E}\left[\psi\left(Z,\beta_{0}\right)\right]=0.$$

Assume that  $\dim (\psi(Z, \beta_0)) = J > \dim (\beta_0) = K$ .

Let  $a(Z,\beta)$  be some known function of Z and  $\beta$ . The goal is to efficiently estimate the mean

$$\mu_0 = \mathbb{E}\left[a\left(Z, \beta_0\right)\right].$$

1. Let  $X = (\tilde{X}', W)$ , with W a binary treatment indicator (also partition  $\beta = (\tilde{\beta}, \alpha')'$ ) and assume that

$$\Pr(Y = 1|X) = \Phi\left(X'\tilde{\beta} + W\alpha\right).$$

Let  $\delta_l = \mathbb{E}[Y|X \in \mathbb{X}_l]$  for l = 1, ..., L. Assume that these conditional means are known to the econometrician (e.g., from register data). The target estimand is

$$\mu_0 = \mathbb{E}\left[\Phi\left(X'\tilde{\beta} + \alpha\right) - \Phi\left(X'\tilde{\beta}\right)\right].$$

Outline a plausible empirical setting to which you could adapt the above components. Show how your setting is accommodated by the general setup outlined above.

- 2. Calculate the semiparametric variance bound, say  $\mathcal{I}(\beta_0)^{-1}$ , for  $\beta_0$ . You may use either the approach outlined in Newey (1990) (and also lecture) or that of Chamberlain (1987).
- 3. Show that the semiparametric variance bound, say  $\mathcal{I}(\mu_0)^{-1}$ , for  $\mu_0$  is

$$\mathcal{I}(\mu_{0})^{-1} = \Sigma_{aa} - \Sigma_{a\psi}\Omega_{0}^{-1}\Sigma_{a\psi}' + \left[\Sigma_{a\psi}\Omega_{0}^{-1}\Gamma_{0} - \Xi_{0}\right]\mathcal{I}(\beta_{0})^{-1}\left[\Gamma_{0}'\Omega_{0}^{-1}\Sigma_{a\psi}' - \Xi_{0}'\right]$$

where  $\Sigma_{aa} = \mathbb{E}\left[\left(a\left(Z,\beta_{0}\right) - \mu_{0}\right)\left(a\left(Z,\beta_{0}\right) - \mu_{0}\right)'\right], \Sigma_{a\psi} = \mathbb{E}\left[\left(a\left(Z,\beta_{0}\right) - \mu_{0}\right)\psi\left(Z,\beta_{0}\right)'\right], \Omega_{0} = \mathbb{E}\left[\psi\left(Z,\beta_{0}\right)\psi\left(Z,\beta_{0}\right)'\right], \Gamma_{0} = \mathbb{E}\left[\frac{\partial\psi(Z,\beta_{0})}{\partial\beta'}\right] \text{ and } \Xi_{0} = \mathbb{E}\left[\frac{\partial a(Z,\beta_{0})}{\partial\beta'}\right]. \text{ As before, you may use either the approach of Chamberlain (1987) or Newey (1990) (in this case I also recommend Brown and Newey (1998)). Interpret each of the three components of the inverse information expression above.$ 

4. Assume that  $\hat{\beta}$  is an efficient estimate of  $\beta_0$ .

(a) Is 
$$\hat{\mu} = \frac{1}{N} \sum_{i=1}^{N} a\left(Z_i, \hat{\beta}\right)$$
 efficient? Explain.

- (b) Let  $\hat{b}_1$  and  $\hat{b}_{\psi}$  be the intercept and slope coefficients associated with the least squares fit of  $a\left(Z_i, \hat{\beta}\right)$  onto a constant and  $\psi\left(Z_i, \hat{\beta}\right)$ . Let  $\hat{\psi}_N = \frac{1}{N} \sum_{i=1}^N \psi\left(Z_i, \hat{\beta}\right)$ .
  - i. Is  $\hat{\mu} = \hat{b}_1 + \hat{\psi}'_N \hat{b}_{\psi}$  consistent for  $\mu_0$ ? Is it efficient? Explain.
  - ii. Is  $\hat{\mu} = \hat{b}_1$  consistent for  $\mu_0$ ? Is it efficient? Explain.
- (c) Assume that  $\beta_0$  is known (such that the analyst computes the least squares fit of  $a(Z_i, \beta_0)$  onto a constant and  $\psi(Z_i, \beta_0)$ ). Show that, in this case,

$$\hat{b}_{1} = \sum_{i=1}^{N} \left[ a(Z_{i}, \beta_{0}) - \sum_{a\psi} \Omega_{0}^{-1} \psi(Z_{i}, \beta_{0}) \right] / N + o_{p} \left( N^{-1/2} \right)$$

and hence that  $\hat{b}_1$  is asymptotically uncorrelated with  $\psi_N = \frac{1}{N} \sum_{i=1}^N \psi(Z_i, \beta_0)$ . Connect this observation to your analysis in part (b) above.

## References

- Brown, B. W. and Newey, W. K. (1998). Efficient semiparametric estimation of expectations. *Econometrica*, 66(2):453 464.
- Chamberlain, G. (1987). Asymptotic efficiency in estimation with conditional moment restrictions. *Journal of Econometrics*, 34(3):305 334.
- Graham, B. S. (2011). Efficiency bounds for missing data models with semiparametric restrictions. *Econometrica*, 79(2):437 452.
- Newey, W. K. (1990). Efficient instrumental variables estimation of nonlinear models. *Econometrica* 58, 58(4):809 837.