

## 2.2 Expectation

Def. The expectation of a random variable  $g(X)$  is

$$\mathbb{E} g(X) := \begin{cases} \int_{-\infty}^{\infty} g(x) f_X(x) dx, & \text{if } X \text{ continuous} \\ \sum_{x \in X} g(x) f_X(x), & \text{if } X \text{ discrete} \end{cases}$$

(\*)

$\nwarrow P(X=x)$

Remark

- Also known as "expected value", "average" of a RV or of the probability distribution of the RV
  - What is the expectation of  $X$ ?
- if  $X \in \mathbb{R}$  then by letting  $g(x) := x$

$$\mathbb{E} X = \begin{cases} \int x f_X(x) dx & \text{if } X \text{ cont} \\ \sum_{x \in X} x f_X(x) & \text{if } X \text{ dis.} \end{cases}$$

(\*)

- if the domain  $X$  is not a subset of  $\mathbb{R}$  (Euclidean space) then (\*) may be invalid but a notation of  $\mathbb{E} X$  may still be defined via the expectations of  $\{\mathbb{E} g(X)\}$  in (\*).

## Examples.

1.  $X \sim \text{Exp}(\lambda)$

Exponential distribution

$$f_X(x) = \frac{1}{\lambda} e^{-x/\lambda}, \quad x \in X = [0, +\infty) \\ \lambda > 0$$

Check

$$\int_X f_X(x) dx = \int_0^\infty \frac{1}{\lambda} e^{-x/\lambda} dx = \left[ e^{-x/\lambda} \right]_{-\infty}^\infty = 1.$$

$$\begin{aligned} \mathbb{E}X &= \int_0^\infty x \frac{1}{\lambda} e^{-x/\lambda} dx \\ &= - \int_0^\infty x d(e^{-x/\lambda}) \\ &\stackrel{\text{integration by part}}{=} -xe^{-x/\lambda} \Big|_0^\infty + \int_0^\infty e^{-x/\lambda} dx \\ &= 0 + \lambda e^{-x/\lambda} \Big|_0^\infty \\ &= \lambda. \end{aligned}$$

## Examples.

2.  $X \sim \text{Binomial}(n, p)$

Binomial distribution

$$f_X(x) = P(X=x) = \binom{n}{x} p^x (1-p)^{n-x}, x \in X = \{0, \dots, n\}$$

$n \in \mathbb{N}, p \in (0, 1).$

Check

$$\sum_{x \in X} f_X(x) = \sum_{x=0} \binom{n}{x} p^x (1-p)^{n-x}$$

$$\begin{aligned} \text{"binomial identity"} &\quad = (p + (1-p))^n \\ &= 1. \end{aligned}$$

$$\mathbb{E}X = \sum_{x=0}^n x \binom{n}{x} p^x (1-p)^{n-x}$$

$$\begin{aligned} \text{see below} &\quad = \sum_{x=1}^n n \binom{n-1}{x-1} p^x (1-p)^{n-x} \end{aligned}$$

$$= np \sum_{x=1}^n \binom{n-1}{x-1} p^{x-1} (1-p)^{n-x}$$

$$y = x-1 \quad = np \sum_{y=0}^{n-1} \binom{n-1}{y} p^y (1-p)^{n-1-y}$$

$$= np$$

Examples.

3.  $X \sim \text{Cauchy}$

if

$$f_X(x) = \frac{1}{\pi} \frac{1}{1+x^2}, \quad x \in X = \mathbb{R}$$

Check

$$\begin{aligned}\int_{-\infty}^{\infty} f_X(x) dx &= \frac{1}{\pi} \int_{-\infty}^{\infty} \frac{1}{1+x^2} dx \\ &= \frac{1}{\pi} \int_{-\pi/2}^{\pi/2} \frac{1}{1+\tan^2 y} d\tan y \\ &= \frac{1}{\pi} \int_{-\pi/2}^{\pi/2} \frac{1}{1+\frac{\sin^2 y}{\cos^2 y}} \frac{1}{\cos^2 y} dy \\ &= \frac{1}{\pi} \int_{-\pi/2}^{\pi/2} 1 dy = 1.\end{aligned}$$

$$\begin{aligned}E(X) &= \int_{-\infty}^{\infty} \frac{1}{\pi} \frac{|x|}{1+x^2} dx \\ &= \int_0^{\infty} \frac{1}{\pi} \frac{|x|}{1+x^2} dx + \int_{-\infty}^0 \frac{1}{\pi} \frac{|x|}{1+x^2} dx \\ &= 2 \int_0^{\infty} \frac{1}{\pi} \frac{x}{1+x^2} dx = \lim_{n \rightarrow \infty} \frac{1}{\pi} \int_0^n \frac{2x}{1+x^2} dx\end{aligned}$$

for any  $M > 0$

$$\int_0^M \frac{2x}{1+x^2} dx = \left. \log(1+x^2) \right|_0^M = \log(1+M^2) \rightarrow \infty \text{ as } M \rightarrow \infty.$$

Hence  $E(X) = \infty$

and  $E(X)$  is undefined (does not exist!)

**NOTE:** Expectation is associated with a distribution

4. Let  $X$  be a Random Variable with dist  $P_X$

Let  $X_1, \dots, X_n$  are  $n$  mutually independent Random variables that have identical distributions as  $X$  i.e.  $X_1, \dots, X_n$  is an  $n$ -iid sample of  $P_X$

The empirical distribution of  $\{X_1, \dots, X_n\}$  is a probability distribution, denoted by  $P_n$ , such that if  $Y \sim P_n$  then  
 $Y \in Y = \{X_1, \dots, X_n\}$  and  
 $P(Y = X_i) = \frac{1}{n}$  for  $i=1, \dots, n$

Thus,  $Y$  is discrete (regardless of  $X$ ), and

$$\begin{aligned} EY &= \sum_{y \in Y} P(Y=y) \times y \\ &= \frac{1}{n}(X_1 + \dots + X_n) \end{aligned}$$

Colloquially, this  $EY$  is also known as the "average" of the (data) sample  $X_1, \dots, X_n$ .

## "LINEARITY of EXPECTATION"

THM. Let  $X$  and  $Y$  be two Real-valued R.V.'s for which the expectations exist. Let  $a, b, c \in \mathbb{R}$ .

- (i)  $\mathbb{E}(ax + by + c) = a\mathbb{E}X + b\mathbb{E}Y + c.$
- (ii) if  $X \geq 0$  almost surely, i.e.  $P(X \geq 0) = 1$  then  $\mathbb{E}X \geq 0.$
- (iii) if  $X \geq Y$  almost surely, i.e.  $P(X \geq Y) = 1$  then  $\mathbb{E}X \geq \mathbb{E}Y.$
- (iv) if  $P(X \in [a, b]) = 1$  then  $\mathbb{E}X \in [a, b].$

Example.

Recall  $X \sim \text{Binomial}(n, p)$

$X$  can be viewed as the number of heads obtained by tossing a coin  $n$  times indep. where  $p = \text{prob. of getting a head.}$

Let  $Y_i = \begin{cases} 1 & \text{if the coin turns at the } i\text{-th toss.} \\ 0 & \text{otherwise} \end{cases}$

Then  $\mathbb{E}Y_i = 1 \cdot p + 0 \cdot (1-p) = p.$

$$X = Y_1 + Y_2 + \dots + Y_n$$

By linearity of expectation  
 $\mathbb{E}X = \mathbb{E}Y_1 + \dots + \mathbb{E}Y_n = np.$

Proof.

(i) This is not a complete proof, but it conveys the main idea.

Assume  $X$  and  $Y$  may be expressed as functions of a random variable  $Z$ , i.e.,

$$\begin{cases} X = g_1(z) \\ Y = g_2(z) \end{cases} \quad (*)$$

Suppose  $Z$  is a continuous RV with  $f_Z$ .  
Then

$$\begin{aligned} & E[aX + bY + c] \\ &= E[ag_1(z) + bg_2(z) + c] \\ &= \int (ag_1(z) + bg_2(z) + c) f_Z(z) dz \\ &= a \underbrace{\int g_1(z) f_Z(z) dz}_{\text{linearity of integration}} + b \underbrace{\int g_2(z) f_Z(z) dz}_{Eg_2(z)} + c \underbrace{\int f_Z(z) dz}_1 \\ &= a E[g_1(z)] + b E[g_2(z)] + c \\ &= a E[X] + b E[Y] + c. \end{aligned}$$

The proof proceeds similarly if  $Z$  is discrete.

[The proof is not complete because we assumed (\*)]  $\square$

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Prog. of (ii) for the case  $X$  is a continuous RV.

$$P(X \geq 0) = 1 \Rightarrow P(X < 0) = 0, \text{ so } f_X(x) = 0 \text{ for } x < 0.$$

Hence,

$$\begin{aligned} \mathbb{E}X &= \int_{-\infty}^{\infty} x f_X(x) dx \\ &= \int_0^{\infty} x f_X(x) dx \geq 0, \text{ since the integrand } \geq 0. \end{aligned}$$

↑ why?

(iii) and (iv) is a direct consequence of (i) and (ii).