

Week 7

* the moments do not capture all info about the dist. in the unbounded support scenario

Q: why do we care?

Answers need to approximate RVs

Theorem (convergence of MGF's leads to convergence of CDF).

Suppose X_1, X_2, \dots is a sequence of random variables, each w.r.t. mgf $M_{X_i}(t)$.

Suppose

$$\lim_{n \rightarrow \infty} M_{X_n}(t) \rightarrow M_X(t)$$

For all t in a neighborhood of 0, and $M_X(t)$ a mgf. (a) \Rightarrow (b) $\exists \delta > 0$ s.t. $\forall t \in (-\delta, \delta)$, $M_{X_n}(t) = M_X(t)$

Then $X_i \rightarrow X$ in distribution

were X is a RV w.r.t. mgf $M_X(t)$

i.e. $F_{X_n}(x) \rightarrow F_X(x)$ at all points x where F_X is continuous

Proof (idea): Beyond scope of class

Def: $X_i \rightarrow X$ in distribution ($X_i \xrightarrow{d} X$)

If $F_{X_n}(x) \rightarrow F_X(x)$ at all continuity points of F_X

- implicit in above item.

EK (Poisson Approx): Let $X \sim \text{Binomial}(n, p)$. We know $\begin{cases} E(X) = np \\ V(X) = np(1-p) \end{cases}$

As n gets large, X "behaves" like a Poisson RV.

$\sim \text{Poisson}(\lambda)$ if

$$f_Y(y) = e^{-\lambda} \frac{\lambda^y}{y!}, y=0, 1, \dots$$

$$E(Y) = V(Y) = \lambda$$

$$M_Y(t) = e^\lambda (e^{t-\lambda})$$

We will establish that

$$P(X \leq x) \approx P(Y \leq x) \text{ as } n \rightarrow \infty \forall x \in \mathbb{N}$$

More precisely, we also need $p \neq 0$

s.t. $np = \text{constant} \in \mathbb{R}$.

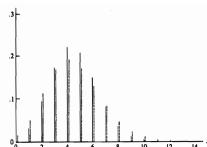


Figure 2.3.3: Poisson (dotted line) approximation to the binomial (solid line), $n = 15, p = .3$

Suppose $\begin{cases} np = \lambda = \text{const} \\ n \rightarrow \infty, p \downarrow 0 \end{cases}$

$$M_X(t) = (pe^t + 1 - p)^n$$

Since $X = Y_1 + \dots + Y_n$ where $Y_i \stackrel{iid}{\sim} \text{Bernoulli}(p)$ s.t. $\sum_{i=1}^n P(Y_i = 1) = p$

$$\Rightarrow M_{Y_i}(t) = E(e^{tY_i}) = E(e^{t(0+1-p+0)}) = E(e^{tp}) \cdots E(e^{tp})$$

and Y_1, \dots, Y_n are independent

$$= \prod_{i=1}^n E(e^{tp}) = (pe^t + 1 - p)^n$$

$$\text{so, } M_{Y_i}(t) = (pe^t + 1 - p)^n$$

$$= (1 + p(e^t - 1))^n$$

$$= (1 + \frac{1}{n} (e^t - 1)n)^n, p = \frac{1}{n} \lambda$$

$$= (1 + \frac{1}{n} (e^t - 1)\lambda)^n$$

$$= (1 + \frac{1}{n} (e^t - 1)\lambda)^{n \rightarrow \infty} e^{t\lambda} \leftarrow \text{you want?}$$

$$= (1 + \frac{1}{n} (e^t - 1)\lambda)^{n \rightarrow \infty} e^{t\lambda}, \lambda = \lim_{n \rightarrow \infty} \frac{1}{n} \lambda^n$$

$$\rightarrow e^{(e^t - 1)\lambda} \text{ as } n \rightarrow \infty$$

Hence, $M_X(t) \rightarrow M_Y(t) \forall t$

By the convergence theorem of mgf,

if $X \sim \text{Binomial}(n, p)$, $n \rightarrow \infty$, then

$$F_X(x) \rightarrow F_Y(x) \text{ as } n \rightarrow \infty \forall x \text{ where } F_Y \text{ is continuous.}$$

For Poisson, F_Y is a step function w.r.t. continuity at $y \in \mathbb{N}$.

For $y \in \mathbb{N}$, $F_Y(y) = F_Y(y - \frac{1}{2})$ continuous at $y \in \mathbb{Z}$.

So,

$$P(X = x) \rightarrow P(Y = x) \text{ as } n \rightarrow \infty \forall x \in \mathbb{N} \text{ (in fact, for } x \in \mathbb{R}).$$

Remark: this holds if $x = \text{Bivariate}(n, p_n)$ and
 $n \rightarrow \infty, p_n \rightarrow 0$ s.t. $p_n \rightarrow x$.

Theorem: $\forall a, b \in \mathbb{R}$,

$$\begin{aligned} \text{Proof (Leibniz's Rule): } M_{a+b}(t) &= E[e^{t(a+bx)}] \\ &= E[e^{taX} e^{tb}] \\ &= e^{tb} E[e^{taX}] \\ &= e^{tb} M_a(t). \quad \square \end{aligned}$$

2.4: TOOLS

Interchange Integral and Differential

Theorem (Leibniz's Rule). If

- $f(x, \theta), a(\theta), b(\theta)$ are differentiable wrt θ
- $\frac{\partial f}{\partial \theta}(x, \theta)$ is continuous on $\mathbb{R} \times (\theta_0, \theta_1)$
 can be measured

Then for $a(\theta), b(\theta) \in (\theta_0, \theta_1)$:

$$\begin{aligned} \frac{d}{d\theta} \int_{a(\theta)}^{b(\theta)} f(x, \theta) dx &= f(b(\theta)) \frac{db}{d\theta}(\theta) - f(a(\theta)) \frac{da}{d\theta}(\theta) \\ &\quad + \int_{a(\theta)}^{b(\theta)} \frac{\partial f}{\partial \theta}(x, \theta) dx \end{aligned}$$

Corollary: If $a(\theta) = a, b(\theta) = b$, then

$$\frac{d}{d\theta} \int_a^b f(x, \theta) dx = \int_a^b \frac{\partial f}{\partial \theta}(x, \theta) dx$$

To strengthen the above result for infinite domain of the integral,
 we need another tool:

Interchanging Limit and Integral

Theorem. Suppose

- $w(x, y)$ is continuous at $y=y_0$ for each fixed x
- There is a function $g(x)$ s.t.

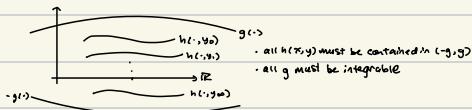
$$\text{envelope condition: } \left\{ \begin{array}{l} |w(x, y)| \leq g(x) \quad \forall x, y \\ \int_{-\infty}^{\infty} g(x) dx < \infty \end{array} \right.$$

Then,

$$\lim_{y \rightarrow y_0} \int_{-\infty}^{\infty} w(x, y) dx = \int_{-\infty}^{\infty} \lim_{y \rightarrow y_0} w(x, y) dx$$

Remarks: - g is called the dominating or envelope function of f

- known as (a version/consequence of) Lebesgue's dominated convergence theorem (DCT)
- think of h as a collection of functions in \mathbb{R} (think y)



Interchange Differential and Integral

Theorem. Suppose

- For each x , $f(x, \theta)$ is differentiable wrt θ , at $\theta \in \Theta$ (1)
- For each $\theta \in \Theta$, there is a function $g(x, \theta)$ and $\delta_0 > 0$ s.t.

$$\left\{ \begin{array}{l} \left| \frac{\partial}{\partial \theta} f(x, \theta) \right|_{\theta=\theta'} \leq g(x, \theta') \quad \forall \theta' \in (\theta - \delta_0, \theta + \delta_0) \\ \int_{-\infty}^{\infty} g(x, \theta) d\theta < \infty \quad \text{for each } \theta \in \Theta \end{array} \right.$$

Then

$$\frac{d}{d\theta} \int_{-\infty}^{\infty} f(x, \theta) dx = \int_{-\infty}^{\infty} \frac{d}{d\theta} f(x, \theta) dx$$

works for each $\theta \in \Theta$

Remark: Leibniz's rule holds under this weaker envelope condition for $f(x, \theta)$.

- if we need only to differentiate at $\theta = \theta_0$, then it is sufficient that the envelope condition be satisfied for a neighborhood of θ_0 i.e. $\exists \delta > 0$ such that $\forall \theta \in (\theta_0 - \delta, \theta_0 + \delta)$ for some $\delta > 0$, and only a dominating function $g(x, \theta)$ (for fixed $\theta = \theta_0$) is required

- this theorem is a direct consequence of Lebesgue's dominated convergence theorem (DCT)

$$\begin{aligned} \text{proof (Leibniz's Rule). LHS: } & \lim_{\delta \rightarrow 0} \frac{1}{\delta} \left(\int_{a(\theta)}^{b(\theta+\delta)} f(x, \theta+\delta) dx - \int_{a(\theta)}^{b(\theta)} f(x, \theta) dx \right) \\ &= \lim_{\delta \rightarrow 0} \frac{1}{\delta} \underbrace{\int_{a(\theta)}^{b(\theta)} [f(x, \theta+\delta) - f(x, \theta)] dx}_{\text{A}(\theta)} + \underbrace{\int_{a(\theta)}^{b(\theta+\delta)} f(x, \theta+\delta) dx}_{\text{B}(\theta)} + \underbrace{\int_{b(\theta)}^{b(\theta+\delta)} f(x, \theta+\delta) dx}_{\text{C}(\theta)} \\ & \text{where } A(\theta) = \lim_{\delta \rightarrow 0} \frac{1}{\delta} \int_{a(\theta)}^{b(\theta)} [f(x, \theta+\delta) - f(x, \theta)] dx \\ & \text{and } C(\theta) = \lim_{\delta \rightarrow 0} \int_{b(\theta)}^{b(\theta+\delta)} f(x, \theta+\delta) dx \quad \text{since } \int_{b(\theta)}^{b(\theta+\delta)} f(x, \theta+\delta) dx \text{ exist because} \\ & \int_{a(\theta)}^{b(\theta)} f(x, \theta+\delta) dx \text{ exist} \\ & B = \lim_{\delta \rightarrow 0} \frac{1}{\delta} \int_{a(\theta)}^{b(\theta+\delta)} f(x, \theta+\delta) dx \\ & = \frac{1}{\delta} (a(\theta+\delta) - a(\theta)) \times \frac{1}{a(\theta+\delta) - a(\theta)} \int_{a(\theta)}^{b(\theta+\delta)} f(x, \theta+\delta) dx \\ & = \frac{1}{\delta} b(\theta) \times f(a(\theta), \theta) \\ & \text{Likewise,} \\ & C = \frac{1}{\delta} b(\theta) \times f(b(\theta), \theta). \quad \square \end{aligned}$$

Exponential: Let $X \sim \text{Exp}(\lambda)$; $F(x) = \frac{1}{\lambda} e^{-x/\lambda}$, $x \in [0, \infty)$.

Moments: $E(x^n) = \int_0^\infty x^n e^{-x/\lambda} dx$, $n=1, 2, \dots$

wish to differentiate

$$\begin{aligned}\frac{d}{dx} E(x^n) &= \frac{d}{dx} \int_0^\infty \dots dx = \int_0^\infty \frac{d}{dx} \dots dx, \text{ need to verify (*)} \\ &= \int_0^\infty x^n \left(-\frac{1}{\lambda^2} e^{-x/\lambda} + \frac{1}{\lambda} e^{-x/\lambda} \cdot \frac{1}{\lambda} \right) dx \\ &= \int_0^\infty x^{n+1} e^{-x/\lambda} \frac{1}{\lambda^2} - \frac{1}{\lambda} x^n e^{-x/\lambda} \frac{1}{\lambda^2} dx \\ &= \frac{1}{\lambda^2} E(x^{n+1}) - \frac{1}{\lambda} E(x^n)\end{aligned}$$

This gives a recursive relation, i.e.

$$E(x^{n+1}) = \lambda E(x^n) + \lambda^2 \frac{d}{d\lambda} E(x^n)$$

(similar identity holds for broad family of dist's)

Now, we justify the interchange of \int and $d/d\lambda$, i.e. $(*)$

$$\text{where } \frac{d}{dx} F(x, \lambda) = x^n e^{-x/\lambda} \frac{1}{\lambda} \left(\frac{x}{\lambda} + 1 \right).$$

Need to find a dominating function $g(x, \lambda)$ s.t.

$$\begin{cases} \sum \left| \frac{d}{dx} F(x, \lambda) \right|_{\lambda=0} < \infty \mid cg(x, \lambda) > 0 \in C([0, \infty), \mathbb{R}) \\ g \text{ integrable wrt } x \\ |\frac{d}{dx} F(x, \lambda)| = x^n e^{-x/\lambda} \frac{1}{\lambda} \left(\frac{x}{\lambda} + 1 \right) \\ \leq x^n e^{-x/\lambda} \frac{1}{\lambda} \left(\frac{x}{\lambda} + 1 \right) \text{ since } x > 0 \\ \leq x^n e^{-x/(0+\epsilon)} \frac{1}{\lambda} \left(\frac{x}{\lambda+\epsilon} + 1 \right) := g(x, \lambda) \end{cases}$$

where the last inequality holds $\forall x > 0, \forall \theta > \theta_0$.

Thus, $(*)$ holds for the chosen $g(x, \lambda)$.

To verify (\dagger) :

$$\begin{aligned} \int g(x, \lambda) dx &= \int x^{n+1} e^{-x/\lambda} \frac{1}{\lambda} \left(\frac{x}{\lambda} + 1 \right) dx + \int x^n e^{-x/\lambda} \frac{1}{\lambda^2} \left(\frac{x}{\lambda} + 1 \right)^2 dx \\ &= \text{multiple of } (n+1)\text{-th moment of an exponential RV} \\ &\quad + \text{multiple of } n\text{-th moment of another exp. RV} \\ &\quad + \dots + \dots \quad " \quad " \quad " \quad \lambda \rightarrow \infty. \end{aligned}$$

Ex(Gaussian): Let $X \sim N(\mu, \sigma^2)$; $F(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$, $x \in \mathbb{R}$.

$$M_X(t) = E(e^{xt}) = \int_{-\infty}^{\infty} e^{xt} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$$

$$\text{where } \frac{d}{dt} M_X(t) = \frac{d}{dt} E(e^{xt}) = E\left[\frac{d}{dt}(e^{xt})\right]$$

Justifying (\dagger) by finding dominating function for

$$\left| \frac{d}{dt} e^{xt} e^{-x^2/2} \right| = |xe^{xt} e^{-x^2/2}| = |x| e^{xt} e^{-x^2/2}$$

Let $O(\delta, t)$:

$$\begin{aligned} \text{if } x \geq 0: |xe^{xt} e^{-x^2/2}| &\leq xe^{(t+\delta)x} e^{-x^2/2} \quad \left\{ \text{gross!} \right. \\ \text{if } x \leq 0: |xe^{xt} e^{-x^2/2}| &\leq (-x)e^{(t+\delta)x} e^{-x^2/2} \end{aligned}$$

Take the max as $g(x, t)$ for $x \geq 0$ and $x \leq 0$.

$$\text{then, } \left| \frac{d}{dt} e^{xt} e^{-x^2/2} \right| \leq g(x, t) \wedge g(-x, t) \wedge e^{(t-\delta)x}, \text{i.e.} \\ \text{the domination holds.}$$

thus, conclude that $g(x, t)$ is integrable.

$$\int_0^\infty xe^{(t+\delta)x} e^{-x^2/2} dx + \int_{-\infty}^0 (-x)e^{(t+\delta)x} e^{-x^2/2} dx$$

$\rightarrow \infty$.

Interchanging Sum and Differential

Suppose $\sum_{n=0}^{\infty} h_n(x, \theta)$ exists (i.e. converges pointwise) $\forall \theta \in C(\mathbb{R})$

Moreover, assume

1) $\frac{\partial}{\partial \theta} h_n(x, \theta)$ is continuous in θ for each x

2) $\sum_{n=0}^{\infty} \frac{\partial}{\partial \theta} h_n(x, \theta)$ converges uniformly for all θ

in a closed subinterval of $C(\mathbb{R})$

then

$$\frac{\partial}{\partial \theta} \sum_{n=0}^{\infty} h_n(x, \theta) = \sum_{n=0}^{\infty} \frac{\partial}{\partial \theta} h_n(x, \theta)$$

Remark: this is a consequence of dominated convergence theorem

Ex(Gaussian): Let $X \sim \text{Normal}(\mu, \sigma^2)$, $\theta \in \mathbb{R}$

$$P(X=x) = \theta C(-\theta)^x, x=0, 1, 2, \dots$$

this is a valid pmf since

$$\sum_{x=0}^{\infty} \theta C(-\theta)^x = 1$$

Differentiating both sides wrt θ , assuming we can interchange \sum and $\frac{\partial}{\partial \theta}$:

$$\sum_{x=0}^{\infty} ((-\theta)^x - \theta x (-\theta)^{x-1}) = 0$$

