样本均值与样本方差相互独立的充要条件

定理 设总体 X 的分布函数 F(x) 具有二阶矩, 即 $EX = \mu < \infty$, $DX = \sigma^2 < \infty$, 若 (X_1, X_2, \cdots, X_n) 为来自总体 X 的一个样本, 则样本均值 $\bar{X} = \frac{1}{n} \sum_{j=1}^n X_j$ 与样本方差 $S^2 = \frac{1}{n-1} \sum_{j=1}^n (X_j - \bar{X})^2$ 相互独立的充要条件是总体 $X \sim N(\mu, \sigma^2)$.

证明. 充分性可参见一般教材, 此处从略, 下面证明必要性. 设总体 X 的特征函数为 $\alpha(t)$, \bar{X} 的特征函数为 $\varphi_1(t_1)$, S^2 的特征函数为 $\varphi_2(t_2)$, (\bar{X}, S^2) 的特征函数为 $\varphi(t_1, t_2)$. 因为 \bar{X} 与 S^2 相互独立, 故 $\varphi(t_1, t_2) = \varphi_1(t_1)\varphi_2(t_2)$, 而由特征函数的性质知:

$$\varphi_1(t_1) = \left(\alpha\left(\frac{t_1}{n}\right)\right)^n, \frac{\varphi_2'(0)}{i} = ES^2 = \sigma^2, \text{ i.e. } \varphi_2'(0) = i\sigma^2.$$

故

$$\left. \frac{\partial \varphi(t_1, t_2)}{\partial t_2} \right|_{t_2 = 0} = \varphi_1(t_1) \varphi_2'(0) = i \sigma^2 \left(\alpha \left(\frac{t_1}{n} \right) \right)^n. \tag{1}$$

另一方面, $\varphi(t_1, t_2) = E\left[e^{i(t_1\bar{X}+t_2S^2)}\right]$, 故

$$\frac{\partial \varphi(t_1, t_2)}{\partial t_2} \bigg|_{t_2 = 0} = E\left(iS^2 e^{i(t_1 \overline{X} + t_2 S^2)}\right) \bigg|_{t_2 = 0} = E\left(iS^2 e^{it_1 \overline{X}}\right)$$

$$= iE\left[\left(\frac{1}{n} \sum_{j=1}^n X_j^2 - \overline{X}^2\right) e^{it_1 \overline{X}}\right]$$

$$= \frac{i}{n} \sum_{j=1}^n E\left(X_j^2 e^{i\frac{t_1}{n} \sum_{j=1}^n X_j}\right) - iE\left(\overline{X}^2 e^{it_1 \overline{X}}\right).$$

又 (X_1, X_2, \cdots, X_n) 的联合分布函数为

$$F^*(x_1, x_2, \dots, x_n) = F(x_1)F(x_2)\cdots F(x_n).$$

故由 Riemann-Stieltjes 积分的性质有

$$\begin{split} E\left(\bar{X}^{2}e^{it_{1}\bar{X}}\right) &= \frac{1}{n^{2}}\left[\left(\sum_{j=1}^{n}X_{j}^{2} + \sum_{j\neq k}X_{j}X_{k}\right)e^{i\frac{t_{1}}{n}\sum_{j=1}^{n}X_{j}}\right] \\ &= \frac{1}{n^{2}}\sum_{j=1}^{n}E\left(X_{j}^{2}e^{i\frac{t_{1}}{n}\sum_{j=1}^{n}X_{j}}\right) + \frac{1}{n^{2}}\sum_{j\neq k}E\left(X_{j}X_{k}e^{i\frac{t_{1}}{n}\sum_{j=1}^{n}X_{j}}\right) \\ &= \left(\alpha\left(\frac{t_{1}}{n}\right)\right)^{n-1}\frac{1}{n^{2}}\sum_{j=1}^{n}\int_{-\infty}^{\infty}x^{2}e^{i\frac{t_{1}}{n}x}\mathrm{d}F(x) + \left(\alpha\left(\frac{t_{1}}{n}\right)\right)^{n-2}\frac{1}{n^{2}}\left[\int_{-\infty}^{\infty}x_{k}e^{i\frac{t_{1}}{n}x_{k}}\mathrm{d}F(x_{k})\right] \\ &= \frac{1}{n}\left(\alpha\left(\frac{t_{1}}{n}\right)\right)^{n-1}\int_{-\infty}^{\infty}x^{2}e^{i\frac{t_{1}}{n}x}\mathrm{d}F(x) + \frac{n-1}{n}\left(\alpha\left(\frac{t_{1}}{n}\right)\right)^{n-2}\left[\int_{-\infty}^{\infty}xe^{i\frac{t_{1}}{n}x}\mathrm{d}F(x)\right]^{2}. \end{split}$$

故

$$\frac{\partial \varphi(t_1, t_2)}{t_2} \bigg|_{t_2 = 0} = \frac{n - 1}{n} i \left\{ \left(\alpha \left(\frac{t_1}{n} \right) \right)^{n - 1} \int_{-\infty}^{\infty} x^2 e^{i \frac{t_1}{n} x} dF(x) + \frac{n - 1}{n} \left(\alpha \left(\frac{t_1}{n} \right) \right)^{n - 2} \left[\int_{-\infty}^{\infty} x e^{i \frac{t_1}{n} x} dF(x) \right]^2 \right\}$$
(2)

由(1)和(2)式得

$$\alpha \left(\frac{t_1}{n}\right) \int_{-\infty}^{\infty} x^2 e^{i\frac{t_1}{n}x} dF(x) - \left[\int_{-\infty}^{\infty} x e^{i\frac{t_1}{n}x} dF(x)\right]^2 = \left(\alpha \left(\frac{t_1}{n}\right)\right)^2 \sigma^2.$$
 (3)

又因为 $\alpha(t) = \int_{-\infty}^{\infty} e^{itx} dF(x)$,故

$$\frac{\mathrm{d}\alpha(t)}{\mathrm{d}t} = i \int_{-\infty}^{\infty} x \mathrm{e}^{itx} \mathrm{d}F(x), \frac{\mathrm{d}^2\alpha(t)}{\mathrm{d}t^2} = i^2 \int_{-\infty}^{\infty} x^2 \mathrm{e}^{itx} \mathrm{d}F(x).$$

则由 (3) 式可知, 关于 $\alpha(t)$ 有

$$-\alpha \frac{\mathrm{d}^2 \alpha}{\mathrm{d}t^2} + \left(\frac{\mathrm{d}\alpha}{\mathrm{d}t}\right)^2 = \sigma^2 \alpha^2. \tag{4}$$

下面解此微分方程. 令

$$\frac{\mathrm{d}\alpha}{\mathrm{d}t} = p, \frac{\mathrm{d}^2\alpha}{\mathrm{d}t^2} = \frac{\mathrm{d}p}{\mathrm{d}t} = \frac{\mathrm{d}p}{\mathrm{d}\alpha}\frac{\mathrm{d}\alpha}{\mathrm{d}t} = p\frac{\mathrm{d}p}{\mathrm{d}\alpha}.$$

故 (4) 化为一阶方程:

$$-\alpha p \frac{\mathrm{d}p}{\mathrm{d}\alpha} + p^2 = -\frac{\alpha}{2} \frac{\mathrm{d}p^2}{\mathrm{d}\alpha} + p^2 = \sigma^2 \alpha^2.$$

令 $Q = p^2$,将上述方程化为

$$-\frac{\alpha}{2}\frac{\mathrm{d}Q}{\mathrm{d}\alpha} + Q = \sigma^2\alpha^2.$$

对此一阶线性微分方程, 不难得到通解为 $Q = (-2\sigma^2 \ln \alpha + C_1)\alpha^2$. 于是可得

$$p = \alpha \sqrt{-2\sigma^2 \ln \alpha + C_1} = \frac{\mathrm{d}\alpha}{\mathrm{d}t}.$$

再一步解得此方程通解为

$$-\frac{1}{\sigma^2}\sqrt{-2\sigma^2\ln\alpha+C_1}=t+C_2,$$

讲而得

$$\ln \alpha = A + Bt - \frac{1}{2}\sigma^2 t^2, \alpha(t) = e^{A + Bt - \frac{1}{2}\sigma^2 t^2}.$$

又根据特征函数的性质有: $1 = \alpha(0) = e^A$, A = 0, 再由 $\mu = \frac{\alpha'(t)}{i}\Big|_{t=0} = \frac{B}{i}$, 故 $B = i\mu$, 即总体 X 的特征函数为 $\alpha(t) = e^{i\mu t - \frac{1}{2}\sigma^2 t^2}$, 故 $X \sim N(\mu, \sigma^2)$, 证毕.

[&]quot;证明方法选自http://www.doc88.com/p-7156270513349.html