# Classification

Machine Learning 2022-2023 - UMONS Souhaib Ben Taieb

# 1 Exercise 1

Suppose we collect data for a group of students in a statistics class with variables:

- $X_1$  = hours studied.
- $X_2$  = undergrad GPA.
- Y = receive an A.

We fit a logistic regression and produce estimated coefficients:

- $\hat{\beta}_0 = -6$
- $\hat{\beta}_1 = 0.05$
- $\hat{\beta}_2 = 1$
- a) How would the model write and how do you interpret its coefficients? **Solution:**

$$p(y|x; \boldsymbol{\beta}) = \begin{cases} \frac{e^{\hat{\beta}_0 + \hat{\beta}_1 x_1 + \hat{\beta}_2 x_2}}{1 + e^{\hat{\beta}_0 + \hat{\beta}_1 x_1 + \hat{\beta}_2 x_2}} & \text{if } y = 1\\ 1 - \frac{e^{\hat{\beta}_0 + \hat{\beta}_1 x_1 + \hat{\beta}_2 x_2}}{1 + e^{\hat{\beta}_0 + \hat{\beta}_1 x_1 + \hat{\beta}_2 x_2}} & \text{if } y = 0 \end{cases}$$

$$= \begin{cases} \frac{e^{-6 + 0.05 x_1 + x_2}}{1 + e^{-6 + 0.05 x_1 + x_2}} & \text{if } y = 1\\ 1 - \frac{e^{-6 + 0.05 x_1 + x_2}}{1 + e^{-6 + 0.05 x_1 + x_2}} & \text{if } y = 0 \end{cases}$$

The log-odds are linear in the input x for the logistic regression model:

$$\log \left( \frac{p(y=1|x; \boldsymbol{\beta})}{1 - p(y=1|x; \boldsymbol{\beta})} \right) = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \hat{\beta}_2 x_2$$
$$= -6 + 0.05 x_1 + x_2.$$

When everything else is held constant, a unit increase in hours studied  $(X_1)$  increases the logodds of a student getting an A by 0.05. In terms of odds, by noting  $\text{odds}_1 = e^{\hat{\beta}_0 + \hat{\beta}_1 x_1 + \hat{\beta}_2 x_2}$  and  $\text{odds}_2 = e^{\hat{\beta}_0 + \hat{\beta}_1 (x_1 + 1) + \hat{\beta}_2 x_2}$ , i.e. when  $x_1$  has been increased by one unit, we find:

odds<sub>2</sub> = 
$$e^{\hat{\beta}_0 + \hat{\beta}_1(x_1+1) + \hat{\beta}_2 x_2}$$
  
=  $e^{\hat{\beta}_0 + \hat{\beta}_1 x_1 + \hat{\beta}_2 x_2} e^{\hat{\beta}_1}$   
= odds<sub>1</sub> $e^{\hat{\beta}_1}$ ,

which yields:

$$\frac{\text{odds}_2 - \text{odds}_1}{\text{odds}_1} = e^{\hat{\beta}_1} - 1 = 0.051.$$

Per extra hours studied, a student increase his odds of getting an A by  $e^{0.05} = 1.051$ , i.e. his odds of getting an A are about  $100 \times (e^{\hat{\beta}_1} - 1) = 5\%$  higher per extra hour studied. Similarly, when everything else is held constant, a unit increase in GPA icreases the log-odds of a student getting a A by 1, which is equivalent to increase his odds of getting an A by  $e^1 = 2.718$ , i.e. his odds of getting an A are about 172% higher per extra GPA score.

b) Estimate the probability that a student who studies for 40h and has an undergrad GPA of 3.5 obtains an A in the class.

## **Solution:**

Using the definition of a logistic regression model, and from the coefficients' estimates, we obtain:

$$p(y=1|x; \boldsymbol{\beta}) = \frac{e^{\hat{\beta}_0 + \hat{\beta}_1 x_1 + \hat{\beta}_2 x_2}}{1 + e^{\hat{\beta}_0 + \hat{\beta}_1 x_1 + \hat{\beta}_2 x_2}}$$

$$= \frac{e^{-6 + 0.05 \times 40 + 1 \times 3.5}}{1 + e^{-6 + 0.05 \times 40 + 1 \times 3.5}}$$

$$= \frac{e^{-0.5}}{1 + e^{-0.5}}$$

$$\approx 0.378$$

b) How many hours would the above student need to study to have a 50% chance of getting an A in the class ?

# **Solution:**

$$p(x) = \frac{e^{-6+0.05 \times x_1 + 1 \times 3.5}}{1 + e^{-6+0.05x_1 + 1 \times 3.5}}$$
$$= \frac{e^{0.05x_1 - 2.5}}{1 + e^{0.05x_1 - 2.5}}$$
$$= 0.5$$

$$\Rightarrow e^{0.05x_1 - 2.5} = 0.5 + 0.5e^{0.05x_1 - 2.5}$$
$$\Rightarrow e^{0.05x_1 - 2.5} = 1$$
$$\Rightarrow x_1 = \frac{\log(1) + 2.5}{0.05} = 50$$

Consider the following dataset with n = 8 observations, three binary input features and a binary response.

$X_1$	$X_2$	$X_3$	Y
1	0	1	1
1	1	1	1
0	1	1	0
1	1	0	0
1	0	1	0
0	0	0	1
0	0	0	1
0	0	1	0

Assume we are using a naive Bayes classifier to predict the value of Y from the values of the other variables.

• a) What is 
$$P(Y = 1|X_1 = 1, X_2 = 1, X_3 = 0)$$
?

## **Solution:**

You've seen in the lecture that in a Naïve Bayes classifier, we make the assumption that the covariance matrix is diagonal, i.e. if p = 2,  $\sum = \begin{pmatrix} \sigma_{11}^2 & 0 \\ 0 & \sigma_{22}^2 \end{pmatrix}$ , which implies  $\sigma_{12}^2 = \sigma_{21}^2 = \text{Cov}(X_1, X_2) = 0$ . In fact, this property results from an even stronger assumption : the variables  $X_i$  are mutually **conditionally independent** given Y.

Under this assumption, we have  $P(X_1 = x_1, X_2 = x_2 | Y = y) = P(X_1 = x_1 | Y = y)P(X_2 = x_2 | Y = y)$ 

$$\begin{split} &P\Big(Y=1|X_1=1,X_2=1,X_3=0\Big)\\ &=\frac{P\Big(X_1=1,X_2=1,X_3=0|Y=1\Big)P\Big(Y=1\Big)}{P\Big(X_1=1,X_2=1,X_3=0\Big)}\\ &=\frac{P\Big(X_1=1|Y=1\Big)P\Big(X_2=1|Y=1\Big)P\Big(X_3=0|Y=1\Big)P\Big(Y=1\Big)}{P\Big(X_1=1,X_2=1,X_3=0|Y=1\Big)P\Big(Y=1\Big)}\\ &=\frac{P\Big(X_1=1|Y=1\Big)P\Big(X_2=1|Y=1\Big)P\Big(X_2=1|Y=1\Big)P\Big(Y=1\Big)}{P\Big(X_1=1|Y=1\Big)P\Big(X_2=1|Y=1\Big)P\Big(X_3=0|Y=1\Big)P\Big(Y=1\Big)}\\ &=\frac{P\Big(X_1=1|Y=0\Big)P\Big(X_2=1|Y=0\Big)P\Big(X_3=0|Y=0\Big)P\Big(Y=0\Big)+P\Big(X_1=1|Y=1\Big)P\Big(X_2=1|Y+1\Big)P\Big(X_3=0|Y=1\Big)P\Big(Y=1\Big)}{P\Big(X_1=1|Y=0\Big)P\Big(X_2=1|Y=0\Big)P\Big(X_3=0|Y=0\Big)P\Big(Y=0\Big)+P\Big(X_1=1|Y=1\Big)P\Big(X_2=1|Y+1\Big)P\Big(X_3=0|Y=1\Big)P\Big(Y=1\Big)}\\ &=\frac{0.5\times0.25\times0.5\times0.5}{0.5\times0.25\times0.5+0.5\times0.25\times0.5\times0.5}\\ &=0.5\end{split}$$

• b) What is 
$$P(Y = 0|X_1 = 1, X_2 = 1)$$
?

# **Solution:**

$$\begin{split} P\Big(Y &= 0 | X_1 = 1, X_2 = 1\Big) \\ &= \frac{P\Big(X_1 = 1 | Y = 0\Big) P\Big(X_2 = 1 | Y = 0\Big) P\Big(Y = 0\Big)}{P\Big(X_1 = 1 | Y = 0\Big) P\Big(X_2 = 1 | Y = 0\Big) P\Big(Y = 0\Big) + P\Big(X_1 = 1 | Y = 1\Big) P\Big(X_2 = 1 | Y = 1\Big) P\Big(Y = 1\Big)} \\ &= \frac{0.5 \times 0.5 \times 0.5}{0.5 \times 0.5 \times 0.5 \times 0.5 \times 0.5 \times 0.5} \\ &= 2/3 \end{split}$$

Now, suppose that we are using a joint Bayes classifier to predict the value of Y from the values of the other variables.

• c) What is 
$$P(Y = 1|X_1 = 1, X_2 = 1, X_3 = 0)$$
?

#### **Solution:**

In a joint Bayes classifier, we do not make the above assumption of conditional independence, meaning that  $P(X_1 = x_1, X_2 = x_2 | Y = y) \neq P(X_1 = x_1 | Y = y)P(X_2 = x_2 | Y = y)$ .

$$P(Y = 1|X_1 = 1, X_2 = 1, X_3 = 0)$$

$$= \frac{P(X_1 = 1, X_2 = 1, X_3 = 0|Y = 1)P(Y = 1)}{P(X_1 = 1, X_2 = 1, X_3 = 0)}$$

$$= \frac{0 \times 0.5}{0.125} = 0$$

As  $P(X_1 = 1, X_2 = 1, X_3 = 0 | Y = 1) = 0 \neq \frac{1}{16} = P(X_1 = 1 | Y = 1)P(X_2 = 1 | Y = 1)P(X_3 = 0 | Y = 1)$ , the variables  $X_1, X_2$  and  $X_3$  are not mutually conditionally independent given Y, which means that the assumption that we made when using Naïve Bayes is in reality not valid.

• d) What is 
$$P(Y = 0|X_1 = 1, X_2 = 1)$$
?

### **Solution:**

$$P(Y = 0|X_1 = 1, X_2 = 1)$$

$$= \frac{P(X_1 = 1, X_2 = 1|Y = 0)P(Y = 0)}{P(X_1 = 1, X_2 = 1)}$$

$$= \frac{0.25 \times 0.5}{0.25}$$

$$= 0.5$$

This problem relates to the QDA model, in which the observations within each class are drawn from a normal distribution with a class specific mean vector and a class specific covariance matrix. We consider the simple case where p = 1; i.e. there is only one feature.

Suppose that we have K classes, and that if an observation belongs to the  $k^{th}$  class, then X comes from a one-dimensional normal distribution,  $X \sim \mathcal{N}(\mu_k, \sigma_k^2)$ . Prove that, in that case, the Bayes' classifier is not linear. Argue that it is in fact quadratic.

#### **Solution:**

For a QDA model, we don't make the assumption of equal covariance matrices (or equal variances here as p=1) across the classes. Therefore, we have that  $\sigma_1^2 \neq \sigma_2^2 \neq ... \neq \sigma_K^2$ , and thus:

$$f_k(x) = \frac{1}{\sigma_k \sqrt{2\pi}} \exp\left(-\frac{1}{2\sigma_k^2} (x - \mu_k)^2\right)$$

And therefore:

$$\begin{aligned} p_k(x) &= \frac{\pi_k f_k(x)}{\sum_l^K \pi_l f_l(x)} \\ &= \frac{\pi_k \frac{1}{\sqrt{2\pi\sigma_k}} \exp\left(-\frac{1}{2\sigma_k^2} (x - \mu_k)^2\right)}{\sum_l^K \pi_l \frac{1}{\sqrt{2\pi\sigma_l}} \exp\left(-\frac{1}{2\sigma_l^2} (x - \mu_l)^2\right)} \\ &= \frac{\frac{\pi_k}{\sigma_k} e^{\gamma_k(x)}}{\sum_l^K \frac{\pi_l}{\sigma_l} e^{\gamma_l(x)}} \quad \text{By posing}: \ \gamma_l(x) = -\frac{1}{2\sigma_l^2} (x - \mu_l)^2 \end{aligned}$$

In QDA, we want to find the value k that maximizes  $p_k(x)$ , i.e. we want to solve the following problem :

$$\begin{aligned} \underset{k}{\operatorname{argmax}} \ p_k(x) &= \underset{k}{\operatorname{argmax}} \ \frac{\frac{\pi_k}{\sigma_k} e^{\eta_k(x)}}{\sum_l^K \frac{\pi_l}{\sigma_l} e^{\eta_l(x)}} \\ &= \underset{k}{\operatorname{argmax}} \log \left( \frac{\frac{\pi_k}{\sigma_k} e^{\eta_k(x)}}{\sum_l^K \frac{\pi_l}{\sigma_l} e^{\eta_l(x)}} \right) \\ &= \underset{k}{\operatorname{argmax}} \log \left( \frac{\pi_k}{\sigma_k} e^{\eta_k(x)} \right) - \log \left( \sum_l^K \pi_l e^{\eta_l(x)} \right) \\ &= \underset{k}{\operatorname{argmax}} \log(\pi_k) + \gamma_k(x) - \log(\sigma_k) - \log \left( \sum_l^K \pi_l e^{\eta_l(x)} \right) \\ &= \underset{k}{\operatorname{argmax}} \log(\pi_k) + \gamma_k(x) - \log(\sigma_k) \qquad \text{As } \sum_l^K \pi_l e^{\eta_l(x)} \text{ is constant } \forall k \\ &= \underset{k}{\operatorname{argmax}} \log(\pi_k) + \frac{1}{2\sigma_k^2} (x - \mu_k)^2 - \log(\sigma_k) \\ &= \underset{k}{\operatorname{argmax}} \log(\pi_k) + \frac{(x^2 + \mu_k^2 - 2\mu_k x}{\sigma_k^2} - \log(\sigma_k) \\ &= \underset{k}{\operatorname{argmax}} \log(\pi_k) + \frac{1}{2\sigma_k^2} x^2 + \frac{\mu_k}{\sigma_k^2} x + (\log(\pi_k) - \log(\sigma_k) - \frac{\mu_k^2}{2\sigma_k} \end{aligned}$$

Which is quadratic in x, hence the name Quadratic Discriminant Analysis.

Bob is playing a bar game, for which the principle is the following: While being blindfolded, Bob has to throw a dart at random on a target that only contains number between 0 and 1. Once he has thrown, he can take off the blindfold, and look at the target value x he got. Based on this value, the bartender secretly pours a beer with probability 0.2 + 0.4x, a mojito with probability 0.6 - 0.4x, and a glass of wine with probability 0.2. If Bob correctly guesses the beverage that has been served, he gets it for free, otherwise he is obliged to pay for it.

- 1. Depending on the target value obtained, what could be the optimal prediction that Bob could make and what would be the name of such a classifier? You will need to derive the boundary decisions of the classifier.
- 2. What would be the misclassification error rate of this classifier? Your answer should be a scalar.

#### **Solution**

(1)

Given  $\mathscr{Y} = \{b, m, w\}$ , where b, m, w stand for "beer", "mojito", and "wine" respectively, the optimal prediction that Bob could make is by creating a Bayes optimal classifier, which is defined as:

$$f_{BOC} = \underset{y \in \mathscr{Y}}{\operatorname{argmax}} P(Y = y|x)$$

Let's note P(Y = b|x), P(Y = m|x) and P(Y = w|x) the probabilities to have a beer, a mojito, and a glass of wine respectively. The data generating process described above can be resumed as such:

$$X \sim \mathrm{U}[0,1]$$
  $Y|X = x \sim \begin{cases} b, & \text{with probability } 0.2 + 0.4x \\ m, & \text{with probability } 0.6 - 0.4x \\ w, & \text{with probability } 0.2, \end{cases}$ 

with U[0,1] being the uniform distribution defined in the interval [0,1]. We have:

$$\begin{cases} P(Y = b|x) = 0.2 + 0.4x \\ P(Y = m|x) = 06. - 0.4x \\ P(Y = w|x) = 0.2. \end{cases}$$
 (1)

We have that:

$$P(Y = b|x) \ge P(Y = m|x) \iff 0.2 + 0.4x \ge 0.6 - 0.4x$$
 (2)

$$\iff 0.8x \ge 0.4 \tag{3}$$

$$\iff x > 0.5$$
 (4)

$$P(Y = b|x) \ge P(Y = w|x) \iff 0.2 + 0.4x \ge 0.2$$
$$\iff 0.6x \ge 0$$
$$\iff x \ge 0$$

$$P(Y = m|x) \ge P(Y = w|x) \iff 0.6 - 0.4x \ge 0.2$$
$$\iff 0.4x \le 0.4$$
$$\iff x < 1,$$

which leads to:

$$f_{BOC}(x) = \begin{cases} b & x \in [0.5; 1] \\ m & x \in [0; 0.5] \end{cases}$$
 (5)

(2)

The Bayes error rate is given by:

BER = 
$$\mathbb{E}_{X} \left[ 1 - \max_{k \in \mathscr{Y}} P(Y = k|x) \right]$$
  
=  $1 - \int_{0}^{1} \max_{k \in \mathscr{Y}} P(Y = k|x) f(x) dx$   
=  $1 - \int_{0}^{1} \max_{k \in \mathscr{Y}} P(Y = k|x) dx$   
=  $1 - \int_{0}^{0.5} P(Y = m|x) dx - \int_{0.5}^{1} P(Y = b|x) dx$   
=  $1 - \int_{0}^{0.5} (0.6 - 0.4x) dx - \int_{0.5}^{1} (0.2 + 0.4x) dx$   
=  $1 - \left[ 0.6x - 0.2x^{2} \right]_{0}^{0.5} - \left[ 0.2x + 0.2x^{2} \right]_{0.5}^{1}$ 

Suppose that you are given a set of n i.i.d. observations  $\mathcal{D} = \{(x_i, y_i)\}_{i=1}^n$  where  $y_i$  is a categorical variable belonging to K categories,  $\mathcal{Y} = \{C_1, ..., C_K\}$ . You wish to fit a multiclass logistic regression model to  $\mathcal{D}$ , i.e.

$$\mathbb{P}(y_i = C_k | x_i) = p_k(x_i; \boldsymbol{\beta}) = \frac{e^{\beta^{(k)} x_i}}{\sum_{l=1}^{K} e^{\beta^{(l)} x_i}}$$

Write the expression of the conditional log-likelihood as a function of the data and the unknown coefficients  $\beta$ .

#### **Solution:**

$$\mathcal{L}(\boldsymbol{\beta}; \mathcal{D}) = p(y_1, ..., y_n | x_i, ..., x_n; \boldsymbol{\beta})$$

$$= \prod_{i=1}^n p(y_i | x_i; \boldsymbol{\beta}) \quad \text{The } y_i \text{ are conditionally independent given the } x_i.$$

$$= \prod_{i:y_i = C_1} p_1(x_i; \boldsymbol{\beta}) ... \prod_{i:y_i = C_K} p_K(x_i; \boldsymbol{\beta}).$$

$$\log \mathcal{L}(\boldsymbol{\beta}; \mathcal{D}) = \sum_{i:y_i = C_1} \log p_1(x_i; \boldsymbol{\beta}) \dots \sum_{i:y_K = C_K} \log p_K(x_i; \boldsymbol{\beta})$$

$$= \sum_{i=1}^n \sum_{k=1}^K \mathbb{1}_{y_i = C_k} \log p_k(x_i; \boldsymbol{\beta})$$

$$= \sum_{i=1}^n \sum_{k=1}^K \mathbb{1}_{y_i = C_k} \log \left( \frac{e^{\boldsymbol{\beta}^{(k)} x_i}}{\sum_{l=1}^K e^{\boldsymbol{\beta}^{(l)} x_i}} \right)$$

where  $\mathbb{1}_{y_i=C_k}$  is an indicator function that equals 1 when  $y_i=C_k$ , 0 otherwise. This is the definition of the categorical cross-entropy, sometimes referred to as the "log-loss".