

TSKS04 Digital Communication Continuation Course

Solutions for the exam 2014-06-09

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We are given impulse responses

$$\begin{aligned} g_{\text{TX}}(t) &= \text{rect}\left(\frac{t}{T}\right), \\ g_{\text{C}}(t) &= \text{rect}\left(\frac{t}{2T}\right), \\ g_{\text{RX}}(t) &= \text{rect}\left(\frac{t}{T/2}\right), \end{aligned}$$

of the transmit filter, channel and receive filter, respectively. Let $p(t)$ be the total impulse response of the cascade of the sender filter and the channel i.e.

$$p(t) = (g_{\text{TX}} * g_{\text{C}})(t) = \begin{cases} 0, & t < -3T/2, \\ 3T/2 + t, & -3T/2 \leq t < -T/2, \\ T, & |t| \leq T/2, \\ 3T/2 - t, & T/2 < t \leq 3T/2, \\ 0, & t > 3T/2. \end{cases}$$

We notice that $p(t)$ is real and even. Then we have the matched filter

$$p_{\text{MF}}(t) = p^*(-t) = p(t).$$

Furthermore, let $y(t)$ denote the output from the channel.

- The optimal (ML) case is if we use $p_{\text{MF}}(t)$ as our receiver filter. ML detection based on samples from the output of the receiver filter is possible if $p_{\text{MF}}(t)$ can be written as a linear combination of shifted versions of $g_{\text{RX}}(t)$. We observe that $g_{\text{RX}}(t)$ is piecewise constant, while $p_{\text{MF}}(t)$ is not. It is therefore impossible to write $p_{\text{MF}}(t)$ as a linear combination of shifted versions of $g_{\text{RX}}(t)$. So, ML detection based on samples from the output of the receiver filter is in this case impossible.
- We observe that matched filter of the channel is $g_{\text{C}}(t)$ itself, since it is real and even. We also observe that $g_{\text{C}}(t)$ can be written as

$$\begin{aligned} g_{\text{C}}(t) &= g_{\text{RX}}\left(t + \frac{3T}{4}\right) + g_{\text{RX}}\left(t + \frac{T}{4}\right) + \\ &\quad + g_{\text{RX}}\left(t - \frac{T}{4}\right) + g_{\text{RX}}\left(t - \frac{3T}{4}\right). \end{aligned}$$

Thus, the alternative sender filter

$$g_{\text{TX}}(t) = \delta(t),$$

would give us

$$p(t) = (g_{\text{TX}} * g_{\text{C}})(t) = g_{\text{C}}(t),$$

and its matched filter is as observed

$$\begin{aligned} p_{\text{MF}}(t) &= p^*(-t) = p(t) = g_{\text{C}}(t) \\ &= g_{\text{RX}}\left(t + \frac{3T}{4}\right) + g_{\text{RX}}\left(t + \frac{T}{4}\right) + \\ &\quad + g_{\text{RX}}\left(t - \frac{T}{4}\right) + g_{\text{RX}}\left(t - \frac{3T}{4}\right). \end{aligned}$$

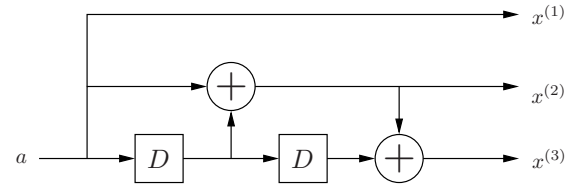
So, in this situation, we can perform ML detection based on samples of the output of the receiver filter.

Answer:

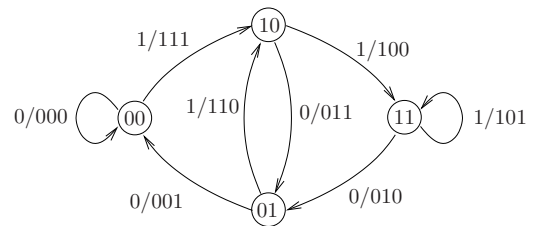
- ML detection is impossible for the given situation.
- ML detection is possible if we change the sender filter to $g_{\text{TX}}(t) = \delta(t)$.

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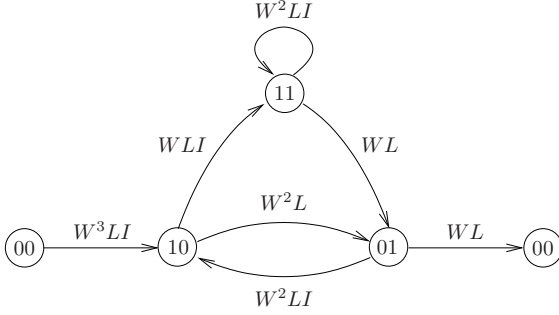
We are given the encoder below.



- The state diagram of the encoder is given below.



To determine the Extended Path Enumerator $T(W, L, I)$ of the encoder, we assign $W^w L$ to a branch with weight w (among output bits) and information bit 0, and $W^w LI$ to a branch with weight w (among output bits) and information bit 1. This gives us the graph below.



Let F_S be the generating function of state S . This gives us the equations

$$\begin{aligned} F_{10} &= 1 \cdot W^3LI + F_{01} \cdot W^2LI, \\ F_{11} &= F_{11} \cdot W^2LI + F_{10} \cdot WLI, \\ F_{01} &= F_{11} \cdot WL + F_{10} \cdot W^2L, \\ T(W, L, I) &= F_{01} \cdot WL \end{aligned}$$

Solving this equation system gives us

$$T(W, L, I) = \frac{W^6 L^3 I (1 + LI - W^2 LI)}{1 - W^2 LI (1 + LI + W^2 LI - W^4 L^2 I)}$$

- b. The dimension, k , of the block code is the number of information bits, i.e. $k = 8$. The convolutional code is a rate $1/3$, which means that every information bit (including trailing dummy zeros) produces 3 codeword bits. The shift register has length 2, which means that we need 2 dummy zeros to force the encoder to the all-zero state. So, the length, n , of the block code is given by $n = 3(k + 2) = 30$. Finally, the minimum distance d of the block code is the free distance of the convolutional code, which is $d = 6$ since the smallest power of W in $T(W, L, I)$ is 6. Totally, the parameters of the block code are

$$[n, k, d] = [30, 8, 6].$$

Answer:

- a. $T(W, L, I) = \frac{W^6 L^3 I (1 + LI - W^2 LI)}{1 - W^2 LI (1 + LI + W^2 LI - W^4 L^2 I)}$
b. $[n, k, d] = [30, 8, 6]$

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The probability that a Poisson variable Y with mean m is y is given by

$$\Pr\{Y=y\} = \frac{m^y e^{-m}}{y!},$$

according to Madhow, page 477. Notice that Y takes only non-negative integer values. Here, it is a photon count.

- a. The likelihood ratio in this situation is therefore

$$\begin{aligned} L(y) &= \frac{\Pr\{Y=y|1 \text{ sent}\}}{\Pr\{Y=y|0 \text{ sent}\}} \\ &= \frac{m_1^y e^{-m_1}/y!}{m_0^y e^{-m_0}/y!} = \left(\frac{m_1}{m_0}\right)^y e^{m_0 - m_1} \end{aligned}$$

The ML criterion compares the likelihood ratio to one (or the log-likelihood function to zero). Thus, we should compare y to the threshold

$$y_0 = \frac{m_1 - m_0}{\ln(m_1/m_0)}$$

For the special case $m_1 = 10m_0 = 100$, we have $y_0 \approx 39.1$. Since y is a photon count and only takes integer values, the estimate is 1 when we have $y > 39$ and 0 otherwise.

- b. According to the above, the estimate is 0 if we have $y \leq 39$, which is an error if 1 was sent. Similarly, the estimate is 1 if we have $y > 39$, which is an error if 0 was sent.

The conditional error probabilities are therefore

$$\begin{aligned} P_{e|1} &= \Pr\{Y \leq 39 | 1 \text{ sent}\} = \sum_{y=0}^{39} \Pr\{Y=y | 1 \text{ sent}\} \\ &= \sum_{y=0}^{39} \frac{m_1^y e^{-m_1}}{y!} = \sum_{y=0}^{39} \frac{100^y e^{-100}}{y!} \\ P_{e|0} &= \Pr\{Y > 39 | 0 \text{ sent}\} = \sum_{y=40}^{\infty} \Pr\{Y=y | 0 \text{ sent}\} \\ &= 1 - \sum_{y=0}^{39} \Pr\{Y=y | 0 \text{ sent}\} \\ &= 1 - \sum_{y=0}^{39} \frac{m_0^y e^{-m_0}}{y!} = 1 - \sum_{y=0}^{39} \frac{10^y e^{-10}}{y!} \end{aligned}$$

Answer:

- a. Output 1 when we have $y > 39$ and 0 otherwise.
b. $P_{e|1} = \sum_{y=0}^{39} \frac{100^y e^{-100}}{y!}$ och $P_{e|0} = 1 - \sum_{y=0}^{39} \frac{10^y e^{-10}}{y!}$

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We have the two hypotheses, H_0 and H_1 , meaning that 0 and 1 was sent, respectively. Under H_0 , the received variable, Y , is $CN(0, N_0)$, i.e. Complex Normal with mean zero and variance N_0 (which means that the real part and the imaginary part are independent, and each with mean 0 and variance $N_0/2$). Under H_1 , Y is $CN(0, A^2 + N_0)$.

- a. The ML rule compares the two PDFs and we therefore get

$$\frac{1}{\pi(A^2 + N_0)} e^{-\frac{|y|^2}{A^2 + N_0}} \underset{H_0}{\overset{H_1}{\geq}} \frac{1}{\pi N_0} e^{-\frac{|y|^2}{N_0}}$$

The constant π can be removed from both sides. After that, taking the natural logarithm of both sides, leaves us with

$$-\frac{|y|^2}{A^2 + N_0} - \ln(A^2 + N_0) \underset{H_0}{\overset{H_1}{\geq}} -\frac{|y|^2}{N_0} - \ln(N_0),$$

which we rewrite as

$$|y|^2 \underset{H_0}{\overset{H_1}{\geq}} N_0 \left(1 + \frac{N_0}{A^2}\right) \ln \left(1 + \frac{A^2}{N_0}\right)$$

- b. Let y_0 be the threshold, i.e. we have

$$y_0 = N_0 \left(1 + \frac{N_0}{A^2}\right) \ln \left(1 + \frac{A^2}{N_0}\right)$$

Assuming that 0 and 1 are equally probable to be sent, then the bit energy is given by

$$E_b = \frac{A^2}{2},$$

since we send A or 0, both with probability 1/2. So, the SNR is given by

$$\frac{E_b}{N_0} = \frac{A^2}{N_0 2},$$

Under the hypothesis H_0 , $|Y|$ has PDF

$$f_{|Y||H_0}(a) = \frac{2a}{N_0} e^{-\frac{a^2}{N_0}} I_{\{a>0\}}(a),$$

i.e. $|Y|$ is rayleigh distributed with mean $\sqrt{\pi N_0}/2$. The conditional error probability is then

$$\begin{aligned} P_{e|0} &= \Pr\{|Y| > \sqrt{y_0} | H_0\} = \int_{\sqrt{y_0}}^{\infty} f_{|Y||H_0}(a) da \\ &= \int_{\sqrt{y_0}}^{\infty} \frac{2a}{N_0} e^{-\frac{a^2}{N_0}} da = \left[-e^{-\frac{a^2}{N_0}} \right]_{\sqrt{y_0}}^{\infty} = e^{-\frac{y_0}{N_0}} \end{aligned}$$

We plug in y_0 from above, and get

$$P_{e|0} = \left(1 + \frac{A^2}{N_0}\right)^{1 + \frac{N_0}{A^2}} = \left(1 + 2\frac{E_b}{N_0}\right)^{1 + \frac{N_0}{2E_b}}$$

- c. The above assumes that we do not know the channel gain h , except that we know its distribution. Now, we know h and therefore we have coherent detection. According to Eq. 3.37 in Madhow, we then have

$$\operatorname{Re}\{y \cdot (hA)^*\} - \frac{|hA|^2}{2} \underset{H_0}{\overset{H_1}{\geq}} 0$$

Notice that the variance is not part of this expression. Plugging in the given numbers, we find

$$\begin{aligned} \operatorname{Re}\{(1+j) \cdot (j\frac{3}{2})^*\} - \frac{|j\frac{3}{2}|^2}{2} &= \\ &= \operatorname{Re}\{(1+j) \cdot (-j\frac{3}{2})\} - \frac{|3/2|^2}{2} \\ &= \frac{3}{2} - \frac{9}{8} = \frac{3}{8} > 0. \end{aligned}$$

Thus, the ML detector announces the estimate 1.

Answer:

- a. $|y|^2 \underset{H_0}{\overset{H_1}{\geq}} N_0 \left(1 + \frac{N_0}{A^2}\right) \ln \left(1 + \frac{A^2}{N_0}\right)$
b. $P_{e|0} = \left(1 + 2\frac{E_b}{N_0}\right)^{1 + \frac{N_0}{2E_b}}$
c. The ML detector announces the estimate 1.

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Assumption according to the problem formulation: ± 1 BPSK symbols.

- a. Let b_n be the n 'th sent symbol and let \bar{r}_n be the n 'th received vector of samples. Let \bar{c} be the correlator. The mean square error is given by

$$\begin{aligned} \text{MSE} &= \mathbb{E}\{(\bar{c}^T \bar{r}_n - b_n)^2\} \\ &= \bar{c}^T \mathbb{E}\{\bar{r}_n \bar{r}_n^T\} \bar{c} - 2\bar{c}^T \mathbb{E}\{b_n \bar{r}_n\} + \mathbb{E}\{b_n^2\} \\ &= \bar{c}^T R \bar{c} - 2\bar{c}^T p + 1, \end{aligned}$$

where $R = \mathbb{E}\{\bar{r}_n \bar{r}_n^T\}$ and $p = \mathbb{E}\{b_n \bar{r}_n\}$ are given in Eq. 5.38 in Madhow. Eq. 5.37 in Madhow gives us $\bar{p} = R \bar{c}_{\text{MMSE}}$, which we plug into the above with $\bar{c} = \bar{c}_{\text{MMSE}}$, and we get

$$\begin{aligned} \text{MMSE} &= \bar{c}_{\text{MMSE}}^T \bar{p} - 2\bar{c}_{\text{MMSE}}^T p + 1 \\ &= 1 - \bar{c}_{\text{MMSE}}^T p = 1 - \bar{p}^T R^{-1} p \end{aligned}$$

- b. Given model (Eq. 5.25 in Madhow):

$$\bar{r}_n = b_n \bar{u}_0 + \sum_{i \neq 0} b_{n+1} \bar{u}_i + \bar{w}_n,$$

The MSE is given by

$$\text{MSE} = \mathbb{E}\left\{\left|\langle \bar{c}, \bar{r}_n \rangle - b_n\right|^2\right\} = \mathbb{E}\left\{\left|b_n \langle \bar{c}, \bar{r}_n \rangle - 1\right|^2\right\},$$

where we have used the given assumption $b_n = \pm 1$. Plugging in the model in this expression gives us

$$\begin{aligned} \text{MSE} &= \mathbb{E}\left\{\left|b_n \langle \bar{c}, b_n \bar{u}_0 + \sum_{i \neq 0} b_{n+1} \bar{u}_i + \bar{w}_n \rangle - 1\right|^2\right\} \\ &= \mathbb{E}\left\{\left|\langle \bar{c}, \bar{u}_0 + b_n \sum_{i \neq 0} b_{n+1} \bar{u}_i + b_n \bar{w}_n \rangle - 1\right|^2\right\} \end{aligned}$$

Assuming that the sequence of BPSK symbols are IID and the two outcomes are equally probable, then the above simplifies to

$$\begin{aligned} \text{MSE} &= \mathbb{E}\left\{\left|\langle \bar{c}, \bar{u}_0 + b_n \sum_{i \neq 0} b_{n+1} \bar{u}_i + b_n \bar{w}_n \rangle - 1\right|^2\right\} \\ &= \left|\langle \bar{c}, \bar{u}_0 \rangle - 1\right|^2 + \sum_{i \neq 0} \left|\langle \bar{c}, \bar{u}_i \rangle\right|^2 + \left|\langle \bar{c}, \bar{w}_n \rangle\right|^2 \end{aligned}$$

since the expectation of all mixed terms become zero. We notice that we have $\left|\langle \bar{a}, \bar{b} \rangle\right|^2 = \bar{a}^T \bar{b} \bar{b}^T \bar{a}$ for vectors \bar{a} and \bar{b} , where $\bar{b} \bar{b}^T$ is a square matrix. Using that observation, we rewrite our expression as

$$\text{MSE} = \left|\langle \bar{c}, \bar{u}_0 \rangle - 1\right|^2 + \bar{c}^T A \bar{c},$$

where we have $A = \sum_{i \neq 0} \bar{u}_i \bar{u}_i^T + C_{\bar{w}_n}$, and where $C_{\bar{w}_n} = \mathbb{E}\{\bar{w}_n \bar{w}_n^T\}$ is the correlation matrix of the noise. The signal-to-interference ratio (SIR) according to equation 5.43 in Madhow, and specialized for the assumed situation is

$$\text{SIR} = \frac{\left|\langle \bar{c}, \bar{u}_0 \rangle\right|^2}{\bar{c}^T A \bar{c}}.$$

We observe that a mere scaling of \bar{c} does not affect SIR, since that results in the same scaling of $\left|\langle \bar{c}, \bar{u}_0 \rangle\right|^2$ and of $\bar{c}^T A \bar{c}$. Thus, maximizing SIR is equivalent to minimizing MSE, which is exactly what the MMSE solution does, as proved on pages 221 and 222 in Madhow.

Answer: Proven above.