Random forest versus logit models: Which offers better early warning of fiscal stress?

replication paper

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1 Introduction and Literature Review

Since the latest Great Recession with its corresponding deterioration of public finances, the monitoring and prevention of fiscal crises has become increasingly prominent in the political debate, leading to an increasing demand for the development of reliable and early indicators that signal possible fiscal stress. In order to be able to assess countries' vulnerability to fiscal distress ex-ante, the literature is increasingly devoted to the development of early warning systems for fiscal stress, which build on early warning systems for banking and currency crises (Honda, Tapsoba, and Issifou 2022). The standard tool used in the literature for early warning systems are the signalling approach as well as discrete dependent variable models, such as logistic regression (Jarmulska 2020).

As an alternative to the traditional methods, early warning models based on machine learning techniques are proposed, claiming a possible improvement of prediction accuracy (Beutel, List, and Schweinitz 2019). As an example, when predicting the build-up of banking crises, the early warning system developed by Casabianca et al. (2019) which builds upon a supervised machine learning algorithm, i.e. Adaptive Boosting, outperforms the traditional approach of using a logit model. Another example which shows that using machine learning could drastically improve prediction accuracy is the early warning system developed by Samitas, Kampouris, and Kenourgios (2020), which reaches an accuracy of 98.8% when predicting the risk of contagion inside a financial network using a quadratic support vector machine.

However, a disadvantage many of these machine learning methods compared to more traditional approaches is the difficulty of understanding how a result was obtained and are therefore often referred to as a black box (Ghoddusi, Creamer, and Rafizadeh 2019). Consequently, the researcher is confronted with a trade-off between prediction and interpretation. Ghoddusi, Creamer, and Rafizadeh (2019) argue, that emphasis should be in interpretation for scientific research or policy decisions, since understanding the relationship and behavior among different variables is more important than prediction accuracy. In contrast, more emphasis is to be placed on prediction accuracy in specific industrial applications.

This paper aims to replicate some of the work done by Jarmulska (2020). Particular emphasis is placed on the comparison of the traditional method, i.e. a logit model with a least absolute shrinkage and selection operator, and a model based on machine learning, i.e. an implementation of the random forest algorithm. As the results are often criticized for being difficult to interpret, ways of interpreting the developed early warning model based on random forest are presented.

2 Model Describtion

2.1 Performance Metrics

Jarmulska (2020) uses sensitivity, specificity, their average as well as the area under receiver operating curve (AUROC) as measures to assess the effectiveness of the early warning models. Since sensitivity corresponds to the proportion of stress episodes correctly classified whereby specificity corresponds to the proportion of tranquil episodes correctly classified, these metrics are dependent on the threshold, which determines whether a period is classified as stress or tranquil episode (Jarmulska 2020). In this paper, this threshold is specified by maximizing the weighted sum of sensitivity and specificity. In contrast, the AUROC is a robust measure, as all possible thresholds are taken into account in their calculation. This measure represents the area under the receiver operating curve, which displays the trade-off between the true positive rate (i.e. sensitivity) and the false positive rate (i.e. 1 - specificity). Theoretically, the AUROC can be between 0 and 1 (perfect classifier), whereby random guessing would result in to a value of 0.5 (Fawcett 2006).

2.2 Logit Model with LASSO penalisation

Jarmulska (2020) implemented two version of discrete dependent variable models (logit regression), first a standard logit model with ordinary least squares estimates and second a logit model with a least absolute shrinkage and selection operator (LASSO) penalization. These models are often used as the standard econometric approach, which is why they are used as the benchmark in this study.

Ordinary least squares estimates often have low bias but large variance, reducing prediction accuracy—the prediction accuracy can sometimes be improved by shrinking some coefficients towards zero to sacrifice bias in order to reduce variance of the predicted values and possibly improve overall prediction accuracy (Tibshirani 1996). To do so, LASSO penalization as proposed in Tibshirani (1996) can be applied. Because of this characteristics, only the logit LASSO model is considered in this replication.

Following Hastie et al. (2009), the LASSO problem in the Lagrangian form is given as follows:

$$\hat{\beta}^{lasso} = \underset{\beta}{\operatorname{argmin}} \left\{ \frac{1}{2} \sum_{i=1}^{N} (y_i - \beta_0 - \sum_{j=1}^{p} x_{ij} \beta_j)^2 + \lambda \sum_{j=1}^{p} |\beta_j| \right\}$$
 (1)

whereby λ corresponds to the penalization parameter. As can be seen in Equation (1), the higher λ , the higher the number of coefficients shrunk to zero. Here, λ is chosen by 5-fold cross-validation, maximizing the AUROC.

2.3 Random Forest

As an alternative to the logit model, Jarmulska (2020) applied the random forest algorithm following Breiman (2001) as an ensemble of multiple classification and regression trees for binary classification. A single tree divides the predictor space into distinct and non-overlapping regions, whereby an observation is classified depending on the region it falls into, i.e. at the terminal node. Each non-terminal node corresponds to a question with a binary response, what determines the structure of the tree. Since each terminal node assigns the same class to all observations within this node, minimizing the classification error at the level of the terminal nodes results in a minimization of the tree's overall classification error (Jarmulska 2020). To measure the precision of the fit, the Gini Index, which is used as a loss function in classification and regression trees,

can be utilised (Jarmulska 2020):

$$g(w) = \sum_{k \neq j} p_{wk} p_{wj} = \sum_{k} p_{wk} (1 - p_{wk})$$
(2)

whereby p_{wj} corresponds to the probability distribution of class j in node w.

However, such decision trees might suffer from overfitting, what can be counteracted by bootstrapping the training set and averaging all the predictions (James et al. 2013). If these so called bagged trees are all influenced by some strong predictors, the single trees might be correlated resulting again in overfitting. To decorrelate the single trees, the non-terminal nodes can be forced to consider only a subset of the predictors, increasing the difference between the single trees, through which the overfitment problem can be reduced (James et al. 2013). This ensemble method corresponds to the random forest algorithm.

hier noch gewählte Parameter für Random Forest beschreiben

3 Data Describtion

3.1 Dependent Variable

The binary dependent variable to be predicted follows the definition in Dobrescu et al. (2011) and takes the value of 1 in the case of a fiscal stress event in the next period and 0 otherwise. According to the definition provided by Dobrescu et al. (2011), an economy faces a fiscal stress event, if at least one of the following four conditions are fulfilled: (1) the economy fails to service debt as payments come due as well as if debt exchanges are distressed (2) the economy receives a large support program by the International Monetary Fund (3) hyperinflation is prevalent in the economy (inflation exceeding 35% for advanced economies, 500% for emerging economies) (4) the economy faces extreme financing constraints, i.e. the sovereign spread exceeds 1000 basis points or 2 standard deviations from the country average.

For the analysis, Jarmulska (2020) considers 43 economies (19 developing economies, 24 developed economies) for an observation period for years 1994-2018. 16.5% of the recorded observation are classified as fiscal stress events, whereby these stress events are not equally distributed across country groups and over time. Figure 1 displays the distribution of the recorded stress events over time for developing and developed economies. The data shows that developing economies are more prone to fiscal stress events with 29% of all observations classified as fiscal stress events compared to 7.1% for developed economies.

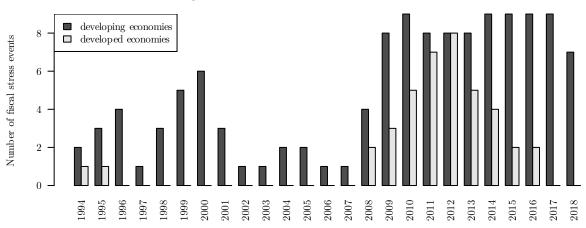


Figure 1: Distribution of Stress Periods

3.2 Explanatory Variables

Annual frequency data lagged 2 years in regard to the dependent variable, hence for an observation period for years 1992-2016, are used to train and test the early warning models. Jarmulska (2020) chose a lag of two years to simulate the reality of how the early warning models could be used in practice, as the data becomes available with a delay of up to one and a half years, leaving the decision-makers half a year time to react on the results of the models.

All variables used for building the early warning model are listed in Table 1 including their means with the distinction whether the economy is in a fiscal stress period or not. In addition, using a Wilcoxon test with statistical significance at 0.05%, Table 1 shows if the means in tranquil periods and in stress periods are statistically different and therefore indicating that the observed variables behave differently in stress periods, hence giving the variables potentially explanatory power.

Table 1: Means of Explanatory Variables

Variable	All periods	Tranquil periods	Stress periods	P-value	Significance			
Competitiveness and domestic demand								
Current account balance	-0.52	0.28	-4.57	0.00	yes			
CPI	4.26	3.68	7.18	0.00	yes			
Credit to GDP change	1.42	1.58	0.58	0.40	no			
Unemployment change	-0.04	-0.14	0.48	0.00	yes			
Consumption dynamics	-4.21	-3.80	-6.31	0.00	yes			
Export share dynamics	0.60	0.78	-0.36	0.08	no			
Financial								
Fixed capital formation dynamics	7.82	8.00	6.96	0.34	no			
FX rate dynamics	1.87	0.69	7.82	0.00	yes			
Fiscal								
GDP dynamics	2.90	3.14	1.71	0.00	yes			
China ĞDP dynamics	9.62	9.63	9.59	0.55	no			
US GDP dynamics	2.46	2.58	1.82	0.00	yes			
Labor market					-			
Labor productivity dynamics	1.76	1.90	1.05	0.00	yes			
GDP per capita	26.13	28.08	16.28	0.00	yes			
Macroeconomic and global economy								
Interest on debt	3.58	3.32	4.92	0.00	yes			
US interest rates	4.27	4.40	3.64	0.00	yes			
Net lending	-2.47	-2.06	-4.55	0.00	yes			
Oil price dynamics	5.05	5.87	0.89	0.03	yes			
Currency overvaluation	-33.76	-31.59	-44.73	0.00	yes			
Public debt	58.51	56.69	67.71	0.00	yes			
VIX	20.17	20.07	20.70	0.74	no			

Pairwise correlation is also examined and visualized in Figure 2. Accordingly, the pairwise correlations are relatively low in most cases, which is why it can be presumed that the variables contain different information and could thus be relevant for the early warning model. However, some variables are highly correlated, which is problematic for econometric models such as the logit model and therefore should be excluded in the model specification, while high pairwise correlations is unproblematic for the random forest model and the variables concerned can be kept in the model specification (Jarmulska 2020).

4 Empirical Results

4.1 Performance

For each year in the interval from 2006 to 2016, both the logit model and the random forest model are implemented recursively, trying to classify the state of the economy two years later. For the logit model,

Currency averalluation Then portrent dang CIR bet capita Currency overvaluation Public debt Current account GDP per capita Greatit to GIR strange Unemployment change Public debt Current account balance China CIR dynamid M VIX Of Price dynamics FX rate dynamics Credit to GDP change Interest on debt Net lending China GDP dynamics Oil price dynamics US CITY DYTHAMICS Interest on debt Export share dynamics CPI

Figure 2: Pairwise Correlation

all variables with pairwise correlations exceeding 65% are excluded to avoid multicollinearity and biases. The models are fitted twice using either a binary variable as a explanatory variable to indicate whether an economy is developed or developing or GDP per capita as a continuous measure for the state of development of an economy. If the binary variable is used, interaction terms are added to the logit model, while the random forest model takes interactions into account by construction thus making further adjustments redundant (Jarmulska 2020).

 $0.2 \quad 0.4 \quad 0.6 \quad 0.8$

-1 -0.8 -0.6 -0.4 -0.2 0

US interest rates US GDP dynamics Labor productivity dynamics

GDP dynamics

Fixed capital formation dynamics

The performance of the models measured using sensitivity (% of correctly classified stress episodes) and specificity (% of correctly classified tranquil episodes) depends on the threshold chosen, whereby when the predicted dependent variable exceeds this threshold, a period is classified as a stress episode. Following Jarmulska (2020), this threshold is determined by maximizing the weighted average of sensitivity and specificity, weighting sensitivity by factor 1, 1.5 and 2 relatively to sensitivity. Table 2 compares the performance of the various models summarized as the average performance of the recursively trained models by year on the corresponding test data, using a threshold determined by the maximized weighted average using factor 1.5 (note that the AUROC is not dependent on the threshold chosen (Jarmulska 2020)). Thereby it can be seen that the model using random forest outperforms the logit LASSO approach in both specifications and in all performance metrics.

Table 2: Average prediction accuracy of early warning models for years 2009-2018

	Logit LASSO		Random Forest	
	advanced	GDP	advanced	GDP
	dummy	per capita	dummy	per capita
sensitivity (% of correctly classified stress episodes)	89.76	76.23	87.69	87.9
specificity (% of correctly classified tranquil episodes)	53.39	69.85	68	69.58
Average of sensitivity and specificity	71.58	73.04	77.84	78.74
AUROC	0.85	0.86	0.88	0.89

4.2 Interpretability of Random Forest Algorithm

4.2.1 Interpretation of logit LASSO vs Random Forest

As seen in Table 2, the model using random forest outperforms the more traditionally used logit model. Depending on the application, however, performance alone is not the only relevant factor when choosing a model. Another central criterion in many use cases is the interpretability of the model, whereby the logit model benefits from the possibility of applying classical econometric methods (marginal effects, Wald test). However, there are also various methods that can be used to interpret the random forest model. These methods are discussed in the following chapters.

4.2.2 Variable Importance (Breiman 2001) and Shapley Values (Shapley 1953)

According to Jarmulska (2020) following Breiman (2001), variable importance corresponds to the average improvement of a model's performance caused by the addition of the respective variable. This improvement can be quantified either as the mean increase in performance or the mean decrease in impurity measured by the Gini index (see Equation (2)). When correlation is present between the different variables considered, it is important to be aware that the variable importance might be biased due to unrealistic observations caused by permutation when calculating the variable importance. In addition, the magnitude of the results will be decreased when correlation is present, since the actual importance is distributed between the correlated variables.

As Štrumbelj and Kononenko (2014) showed, Shapley values (see Shapley (1953)) - originally a solution concept in cooperative game theory - can be used to interpret the results of a machine learning algorithm by determining which variables are influential in a prediction. By assigning a quantitative value to each variable, Shapley values show the importance of a variable on the result of the prediction over a collection of observations (Ma and Tourani 2020). Further technical information can be found for example in Štrumbelj and Kononenko (2014) or in Ghorbani and Zou (2019).

Figure 3 shows the variable importance and the Shapley values for all variables considered in the model, whereby a higher value indicates a higher influence in the prediction. In line with the results obtained by Jarmulska (2020), both Breiman's variable importance and Shapley values similarly rank the importance of the variables.

Current account balance Current account balance GDP per capita US interest rates CPI GDP per capita nterest rates Net lending Net lending Consumption dynamics Consumption dynamics change GDP dynamics Unemployment change Currency overvaluation Public debt Interest on debt Fixed capital formation dynamics Interest on debt Public debt GDP dynamics GDP dynamics Currency overvaluation productivity dynamics Fixed capital formation dynamics Export share dynamics Credit to GDP change China GDP dynamics Labor productivity dynamics US GDP dynamics Export share dynamics FX rate dynamics FX rate uy nou... Credit to GDP change VIX China GDP dynamics Oil price dynamics Oil price dynamics 10 20 25 0.02 0.03 0.04 15 0.01 Breiman's Variable Importance Shapley Values

Figure 3: Variable Importance and Shapley Values of Predictors used

4.2.3 Partial dependence plots (Friedman 2001) and Accumulated local effects plots (Apley and Zhu 2020)

In this application of predicting probabilities, partial dependence plots according to Friedman (2001) show the effect of a variable at different values on the predicted probability. To do this, all other independent variables are averaged and a model is fitted, depending on the varying variable to be considered. Certain artificially created observations could bias the results, especially in the case of correlated variables, since these observations might be implausible in reality (Jarmulska 2020).

As an alternative to partial dependence plots, accumulated local effects plots following Apley and Zhu (2020) can be used, for which correlation among the independent variables does not lead to biases. Reason for this is that accumulated local effects plots display the effect of a change in the variable to be considered on the predicted probability only in a small interval, accumulated over a grid of such small intervals, hence making pure effect of changes in this variable visible (Jarmulska 2020).

The partial dependence plots (left-hand side) and the local effects plots (right-hand side) are visible in Figure 4 for the variables current account balance and unemployment changes, whereby a similar shape of the curves is recognisable. When examining these plots for the independent variables used in this model, for many variables a U-shaped curve can be detected, indicating that both up- and downshifts in the corresponding variable contribute to higher probability of fiscal stress. However, these U-shaped curves are strongly influenced by the tails of the distribution, potentially problematic for both partial dependence and accumulated local effects plots and should therefore be interpreted with caution (Jarmulska 2020).

hier ein plot ersetzen mit u-shaped da interessanter, zB unemployment

5 Conclusion

reputation and credibility (p19)?

see slide 22 ML VL 10 -> RF outperformed by other methods?

Predicted probability of stress Predicted probability of stress 0.10 -1.0 0.05 0.00 -1.5 -10 -20 -10 10 20 -20 10 20 Current account balance Current account balance Predicted probability of stress Predicted probability of stress -0. -1.0 0.06 -1.20.04-0.02--1.6-5 10 -5 10

Figure 4: Partial Dependence (LHS) and Accumulated Local Effects (RHS) Plots

6 References

The code and data used for this project can be found in the corresponding GitHub-repository: https://github.com/bt-koch/ML-in-Economics.

Change in unemployment

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