

Assignment 3

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$$\begin{aligned}
 & 1. \quad P(X_n \leq t \mid x_1 < x_2 < \dots < x_n) \\
 &= \frac{P(X_n \leq t, x_1 < x_2 < \dots < x_n)}{P(x_1 < x_2 < \dots < x_n)} \\
 &= \frac{\int_0^t f(x_n) dx_n \int_0^{x_n} f(x_{n-1}) dx_{n-1} \dots \int_0^{x_1} f(x_1) dx_1}{\int_0^\theta f(x_n) dx_n \int_0^{x_n} f(x_{n-1}) dx_{n-1} \dots \int_0^{x_1} f(x_1) dx_1} \\
 &= \frac{\frac{x_n^{3n}}{n! \cdot \theta^{3n}} \Big|_0^t}{\frac{x_n^{3n}}{n! \cdot \theta^{3n}} \Big|_0^\theta} = \frac{t^{3n}}{\theta^{3n}}
 \end{aligned}$$

$$\begin{aligned}
 \text{hence, } & P\left(c < \frac{x_n}{\theta} < 1 \mid x_1 < x_2 < \dots < x_n\right) \\
 &= 1 - P\left(x_n \leq c \cdot \theta \mid x_1 < x_2 < \dots < x_n\right) \\
 &= 1 - \frac{(c\theta)^{3n}}{\theta^{3n}} = 1 - c^{3n}. \quad \square
 \end{aligned}$$

$$2. (a) \quad H_0: \theta = 1/2 \quad H_1: \theta < 1/2$$

$$\text{set } \theta' = 1/2, \quad \theta'' \in H_1$$

$$\frac{L(\theta', x_1, \dots, x_k)}{L(\theta'', x_1, \dots, x_k)} = \frac{\prod_{i=1}^k \theta'^{x_i} (1-\theta')^{1-x_i}}{\prod_{i=1}^k \theta''^{x_i} (1-\theta'')^{1-x_i}}$$

$$= \left(\frac{\theta' (1-\theta'')}{\theta'' (1-\theta')} \right)^{\sum_{i=1}^k x_i} \cdot \left(\frac{1-\theta'}{1-\theta''} \right)^k \leq k$$

$$\Leftrightarrow \sum_{i=1}^k x_i \cdot \ln \left(\frac{\theta' (1-\theta'')}{\theta'' (1-\theta')} \right) \leq \ln k$$

$$\Leftrightarrow \sum_{i=1}^k x_i \leq c, \quad \text{set } c \triangleq \frac{\ln k}{\ln \left(\frac{\theta' (1-\theta'')}{\theta'' (1-\theta')} \right)}$$

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so $C \triangleq \left\{ \sum_{i=1}^5 x_i \leq c \right\}$ is critical region.

$Y \leq c$ is uniformly most powerful test.

$$\begin{aligned} \text{(b) significance level } \alpha &= P\left(\sum_{i=1}^5 x_i \leq 1 \mid H^0\right) \\ &= P(Y \leq 1 \mid \theta = 1/2) \end{aligned}$$

$$Y \sim \text{Binomial}(n, \theta)$$

$$\alpha = \binom{5}{0} \cdot \left(\frac{1}{2}\right)^5 + \binom{5}{1} \left(\frac{1}{2}\right)^5 = \frac{3}{16}$$

$$\text{(c) } \alpha = P(Y = 0 \mid \theta = 1/2) = \binom{5}{0} \cdot \left(\frac{1}{2}\right)^5 = \frac{1}{32} \quad \square$$

$$3. \text{ (a) } E(\bar{X}) = E(k\bar{X}_1 + (1-k)\bar{X}_2)$$

$$= k \cdot \mu + (1-k)\mu$$

$$= \mu,$$

so it's unbiased for μ .

$$\text{(b) } \lim_{n \rightarrow \infty} P(|\bar{X} - \mu| \geq \epsilon) = 0 \text{ is what we need to prove.}$$

First, prove $\bar{X}_1 \xrightarrow{P} \mu$,

$$MSE(\bar{X}_1) = \text{Var}(\bar{X}_1) + \text{bias}(\bar{X}_1)^2 = \frac{1}{n} \text{Var}(X_1) + 0 \rightarrow 0 \text{ as } n \rightarrow \infty$$

so we have $\bar{X}_1 \xrightarrow{P} \mu$.

Similarly, $\bar{X}_2 \xrightarrow{P} \mu$.

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$$\bar{x}_1 \xrightarrow{P} \mu, \bar{x}_2 \xrightarrow{P} \mu \Rightarrow \bar{x} = k\bar{x}_1 + (1-k)\bar{x}_2 \xrightarrow{P} \mu.$$

$$(c) \text{Var}(\bar{x}) = k^2 \text{Var}(\bar{x}_1) + (1-k)^2 \text{Var}(\bar{x}_2)$$

$$= k^2 \cdot \frac{\sigma_1^2}{n} + (1-k)^2 \cdot \frac{\sigma_2^2}{n}$$

$$\frac{\partial \text{Var}(\bar{x})}{\partial k} = 2k \cdot \frac{\sigma_1^2}{n} + 2(k-1) \cdot \frac{\sigma_2^2}{n} = 0$$

$$\Rightarrow k = \frac{\sigma_2^2}{\sigma_1^2 + \sigma_2^2}$$

$$\frac{\partial^2 \text{Var}(\bar{x})}{\partial k^2} = 2 \left(\frac{\sigma_1^2 + \sigma_2^2}{n} \right) > 0, \text{ so } k = \frac{\sigma_2^2}{\sigma_1^2 + \sigma_2^2} \text{ is minimizer. } \square$$