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Net ID: 2411 24569

Name: Zhong Qiaoyang

1. Let  $X_1, \ldots, X_n$  be a random sample from a geometric distribution that has pmf  $f(x|\theta) =$  $(1-\theta)^x\theta$ ,  $x=0,1,2\ldots,0<\theta<1$ , zero elsewhere. Show that  $\sum_{i=1}^n X_i$  is a sufficient statistic for  $\theta$ .

**2.** Let  $X_1, \ldots, X_n$  be a random sample from a  $\text{Beta}(\theta, 5)$ . Show that the product  $X_1 \times \cdots \times X_n$ is a sufficient statistic for  $\theta$ .

$$f(\infty \mid \theta, \xi) = \frac{x^{\theta-1} (1-x)^{\alpha}}{15(\theta, \xi)}$$

3. Write the pdf

$$f(x|\theta) = \frac{1}{604}x^3e^{-x/\theta}, \quad 0 < x < \infty, \ 0 < \theta < \infty,$$

zero elsewhere, in the exponential form. If  $X_1, X_2, ..., X_n$  is a random sample from this distribution, find a complete sufficient statistic  $Y_1$  for  $\theta$  and the unique function  $\phi(Y_1)$  of this statistic that is the MVUE of  $\theta$ . Is  $\phi(Y_1)$  itself a complete sufficient statistic?

@ E(x) = ( +00 1 604 x4.e-x10 dx

$$= \int_{0}^{+\infty} -\frac{4}{60} \times x^{4} de^{-x/\theta} = \left( \frac{+x}{60} + \frac{4}{60} + x^{5} e^{-x/\theta} dx \right)$$

$$= \int_0^{+\infty} \frac{4a!x!x!}{6} e^{-\pi/\theta} d\pi = 40$$

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$$E(\frac{h}{2} \times i) = 0$$
, is unbiased estimator.

女文 is unique MVUL.

3 since  $T_i = \sum_{i=1}^{n} x_i$  is complete sufficient statistics,

\$ I i) complete sufficient statistics as well. □

**4.** Let  $X_1, \ldots, X_n$  be a random sample from the density function  $f(x|\theta) = \frac{1}{\theta}e^{x/\theta}, \ x < 0, \ \theta > 0$ . Determine the MVUE of  $\theta$ .

4.

(x) 2 xi 13 complete sufficient statistics

 $E(x) = -\theta$ ,  $E(x) = -\theta$ .

-I is unbiased estimator of 0,

 $\phi(\vec{x}) = E(-\vec{x}|\vec{x}) = -\vec{x}$  is the unique MVUE for  $\theta$ .  $\Box$ 

5. Let  $X_1, \dots, X_n$  be a random sample from a uniform distribution  $[0, \theta]$ , for some unknown parameter  $\theta > 0$ . Is  $X_{(n)}$ , the largest order statistic among all samples, a sufficient statistic for  $\theta$ ?

5. Xin = max { 1, -- , xn}

$$P(x_{(n)} \in t) = P(x_1 \in t). P(x_2 \in t) - P(x_n \in t)$$
  
=  $(\frac{t}{\Theta})^n$ , for  $0 \in t \in \Theta$ .

$$f_{\times n}(t|\theta) = n \cdot \frac{t^{n-1}}{\theta n}, \quad 0 \le t \le \theta.$$

$$\frac{f(x_{1}, |\theta) - f(x_{1}, |\theta)}{f_{X_{1}(x_{1})}(t|\theta)} = \frac{(\frac{1}{\theta})^{h}}{n \cdot \frac{t^{n+1}}{\theta^{n}}} = \frac{1}{n \cdot t^{n+1}} \quad \text{is independent with } \theta,$$
hence  $X_{1}(x_{1}) = \frac{1}{n \cdot t^{n+1}} = \frac{1}{n \cdot t^{n+1}}$ 

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